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URBAN FORESTS: ENVIRONMENTAL HEALTH VALUES AND RISKS

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## **ABSTRACT**

Urban forests are ubiquitous, yet their impacts and values remain largely unknown. We study a massive urban afforestation policy in Beijing that planted 1/3 of a million acres of greenery in less than a decade. We conduct a remote-sensing audit of the program, finding that it contributes to a substantial greening up of the city. This causes significant downwind air quality improvement, reducing average PM2.5 concentration at city population hubs by 4.2%. Rapid vegetation growth unexpectedly led to a 7.4% increase in pollen exposure. Analysis of medical claims data shows that increased aeroallergens triggered emergency room visits, mirroring pollution effects though much less severe. Monetized net health benefits of the program amount to 1.5% of the city's GDP. Urban forests are only partially capitalized in housing values, with buyers mainly appreciating proximity to green spaces but not the air quality improvements they bring.

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# 1. Introduction

Forests accompany the cities we build. There are an estimated 5.5 billion urban trees in the United States alone. Worldwide, urban tree cover – the proportion of the total urban land area covered by tree canopy – is estimated to be around 25 percent ([Nowak and Greenfield, 2018; 2020](#)). Representing cities' most salient and widespread natural capital, urban forests are expected to deliver a multitude of environmental benefits, such as pollution reduction, climate moderation, and aesthetic improvements. Beyond environmental gains, urban forests also contribute to landscapes, property values, and employment opportunities, which often makes them a politically viable policy tool.

Despite their relevance for residents' economic and environmental life, the real-world impacts of urban forests – and the economics of urban afforestation programs – are poorly understood. This paper provides an empirical evaluation of a large-scale urban afforestation policy. Our study setting is Beijing, the capital city of China, which has a population of over 20 million people. In response to the central government's national campaign on environmental recovery, Beijing implemented a large-scale urban afforestation effort called the Million *Mu* Project (henceforth MMP), which began in 2012 and involves planting trees and shrubs throughout the city. In a decade, the project has added 2 million *mu* of greenery – equivalent to 515 square miles, roughly the size of Los Angeles – spread throughout the city, with many forest patches situated near the city's most densely populated areas.

We evaluate this policy in five steps. First, we conduct a remote-sensing analysis to document how the MMP has contributed to a notable green up of the city over the past decade. Second, we develop a quasi-experimental research design to quantify the impact of the urban forests on downwind air pollution reduction – the city's most canonical challenge and the primary aim of the policy. Third, we provide an empirical assessment of a prevalent concern about urban afforestation – increased exposure to vegetation's pollen emissions – using novel data on pollen counts monitoring. Fourth, we use medical claims data to provide direct estimates of the health burdens associated with both pollution reduction and aeroallergens. Fifth, we take the local effect estimates, each identified off its own credible source of variation, and then use

them jointly in a unified cost-and-benefit framework to quantify health trade-offs from afforestation.<sup>1</sup>

We begin with a remote-sensing audit of Beijing's greening progress. Using a satellite measurement of vegetation growth, we document a 30 percent increase in greenery within the city between 2001 and 2020. This trend accelerated notably after 2012, with numerous densely populated areas experiencing annual vegetation index growth of over 10 percent. By cross-referencing government maps of MMP planting sites with high-resolution satellite data, we demonstrate that much of this green growth in hotspot areas is indeed driven by the MMP program. The vegetation growth is highly relevant to city residents, as many planting sites are located right on the outskirts of densely populated areas. We estimate that over 1.3 million people (7 percent of the city population) live in areas directly influenced by MMP planting, and the vast majority of the city's population resides within a few kilometers of these sites – which, as we will show later, are significantly impacted by the environmental effects of the new greenery through spatial spillover.

These new forests serve as natural absorbers and filters that reduce pollution that would otherwise reach population centers. We construct a quasi-experimental estimation equation that uses pollution monitoring data from *in-situ* real-time air quality monitoring sites installed by the government near population hubs, combined with prevailing wind and satellite vegetation data, to estimate the impact of urban forests on air quality. Exploiting plausibly exogenous daily changes in wind patterns, we find that vegetation located upwind of monitors significantly reduces air pollution detected at these sites. Forests are particularly effective in reducing concentrations of fine particulate matter (PM<sub>2.5</sub>) – widely considered among the most harmful pollutants to human health – and they also reduce levels of trace gas pollutants, such as sulfur dioxide and carbon monoxide. Over the first eight-year period since MMP started, the project reduces the average population PM<sub>2.5</sub> exposure in the city by 4.2 percent (about 2.9 ug/m<sup>3</sup> from 2012 baseline). This effect is significant: the city of Beijing has achieved a 40 percent reduction in since China's War on Pollution campaign ([Greenstone et al., 2021](#)). Our findings suggest a sizable share of that reduction is contributed by urban afforestation.

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<sup>1</sup> We also report a supplementary analysis on housing market capitalization, which provides an alternative lens to gauge whether and how the environmental values of the MMP are – or are not – captured through changes in home purchasing prices.

A prevalent concern about urban afforestation is the resulting increases in pollen emissions and potential adverse health consequences such as allergy attacks. We assess this empirical possibility using *in-situ* monitoring data from pollen monitors that record daily pollen counts at city population hubs. Using the same wind-direction-based estimation method described earlier, we document that upwind vegetation indeed leads to a sharp increase in pollen levels. The impact is sizable, exhibiting greater magnitude and statistical precision in elasticity compared to the pollution reduction impact. For example, given an upwind vegetation shock of the same size, the response of the pollen spike in log scale is over twice as large as that of the reduction in fine particulate matter (PM<sub>2.5</sub>). Overall, the MMP program increased the average population pollen exposure by 7.4 percent.

We proceed to estimate the health effects associated with these environmental changes using administrative medical claims data covering all city residents. We highlight two facts about pollen and pollution variation: First, while the timing of pollen seasons is relatively fixed, there is substantial variation in pollen exposure across pollen seasons of different years; Second, there is a significant amount of independent variation in daily fluctuations of pollen and air pollution. These variations allow us to use OLS regressions with high-dimensional fixed effects to provide estimates on the effect of day-to-day pollen fluctuations on healthcare utilization, and to compare them with the same outcomes but looking at pollution fluctuations. We find that daily pollen shocks lead to significant increases in the number of emergency room (ER) visits, driven mostly by those due to respiratory or sensory system emergencies. The elasticity is on par with the effect of daily PM<sub>2.5</sub>, both using estimates from our own data and those borrowed from the prior literature.

The health evidence suggests that the negative health externality associated with pollen exposure should not be overlooked: it is significant enough to trigger ER visits, with the magnitude of the elasticity similar to the effects of pollution on ER visits. However, several characteristics of the effects indicate that the overall health cost of pollen exposure is likely small compared to the benefits of pollution reduction: First, although both pollen exposure and PM<sub>2.5</sub> contribute to an increase in the *number* of ER visits to a similar extent, the impact on ER *spending* is three times smaller for pollen exposure. This suggests that ER visits due to pollen exposure tend to be less severe compared to those triggered by PM<sub>2.5</sub>. Second, pollen primarily leads to increased ER visits for patients who do not require further inpatient care. In contrast, it is well-

documented that PM<sub>2.5</sub> can result in severe health emergencies and can lead to death even at relatively low exposure levels ([Deschenes, Greenstone, and Shapiro, 2017](#); [Deryugina et al., 2019](#); [Huang, Xing, and Zou, 2023](#)). Third, unlike industrial pollution, which tends to disproportionately affect the elderly population, pollen has an equal or potentially stronger effect on the non-elderly population. We also find that pollen's impact appears to be concentrated among individuals with prior allergy histories, which may facilitate targeted prevention efforts.

Using both our own empirical findings and relevant estimates from the prior literature, we calculate that the annual healthcare cost savings resulting from the reduction of PM<sub>2.5</sub> due to the MMP project range from 229 to 916 million CNY (32 to 88 million USD). This corresponds to about 0.25 percent of Beijing's annual total reported health spending. Additionally, the annual mortality benefits, measured in terms of the value of statistical life (VSL), amount to an estimated value of 5 billion CNY (710 million USD). The healthcare costs associated with the increased pollen exposure related to the MMP project are about one-ninth of the magnitude of the pollution benefits, totaling 25 to 102 million CNY (3.5 to 14 million USD). The health value of the MMP program attributable to the pollution change is thus about 55 billion CNY over the course of the decade since its initiation. This is a significant number even compared to the magnitude of Beijing's GDP which is about 4,000 billion CNY. Per government reports, the total cost of the MMP project is 75 billion CNY. Our analysis thus concludes that these costs will likely be recouped via health gains alone over the next decade.

Our study shows the importance of understanding the impacts of urban forests: even when considering only changes in air quality, the identified benefits and risks are substantial. At the same time, however, it is difficult to fully quantify all possible channels of benefits and costs. An alternative way to gauge the overall value of the afforestation program is through housing market capitalization ([Chay and Greenstone, 2005](#); [Bayer, Keohane, and Timmins, 2009](#)). The air quality data used in this research were available in real-time to the public over the study period. It is thus tempting to ask whether changes in home transaction value in Beijing reflect MMP-induced air quality improvements. Using home transaction data in a difference-in-differences design, we find an increase in home prices by 3.2 percent for every kilometer closer to MMP forests. This price appreciation is primarily observed within a 2 km radius of MMP sites but does *not* align with the pollution reduction spillover attributed to the forests in our earlier findings. This indicates that the price capitalization is more likely attributable to localized amenity effects,

such as aesthetic improvements, rather than the air quality improvements, which may be more subtle and not as readily perceived by residents.

To the best of our knowledge, our study provides the first thorough empirical assessment of the environmental benefits, risks, and health values of a large-scale urban afforestation program. This paper joins a nascent economics literature on the causal impacts of forests, such as the environmental and economic consequences of tree mortality ([Druckenmiller, 2020](#)), the impacts of large-scale deforestation or afforestation on downwind precipitation and agricultural outcomes ([Araujo, 2023](#); [Grosset, Papp, and Taylor, 2023](#)), the effects of urban trees on microclimate and property values ([Han et al., 2021](#); [Li, 2023](#)), the effect of urban afforestation on infant health improvement ([Jones and Goodkind, 2019](#)), and the welfare effects of urban greenbelts ([Koster, 2024](#)). The dearth of economic studies contrasts with a vast, multi-disciplinary effort to understand urban forests and their potentials in air quality improvement, climate moderation, energy conservation, carbon sequestration, noise reduction, stormwater management, wildlife habitat provision, among many other ecosystem services that are critical for sustainable urban development and climate change policies more broadly ([Kahn and Walsh, 2015](#)). While the natural sciences offer valuable insights into the interactions between forests and the environment, numerous questions remain open regarding the real-world effectiveness, efficiency, and equity of urban afforestation policies from an econometric standpoint.<sup>2</sup>

Our study also shows that urban afforestation is an integral part of urban pollution control – a universal challenge for cities worldwide. Cities have to deal with pollution from various sources, such as transportation, construction, waste, energy generation, and industrial activities ([Glaeser and Kahn, 2010](#); [Currie and Walker, 2011](#); [Zheng and Kahn, 2013](#); [Gendron-Carrier et al., 2022](#)). Traditional pollution policies focus on reducing emissions from these individual pollution *sources*. Another approach is to mitigate pollution exposure at *receptors* – the people. Our study establishes urban forests as an effective – and one of the most prevalent – policy instruments in that vein. The political feasibility in fact seems a unique feature of urban

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<sup>2</sup> For example, on the topic of urban forestry and air quality, [Nowak, Crane, and Stevens \(2006\)](#) uses computational modeling approach combined with parameter calibration to predict impact of urban trees on air pollution removal. In atmospheric science, [Abhijith et al. \(2017\)](#) provides a detailed review of the atmospheric interactions between vegetation and surrounding, and the implications for pollution exposure. In epidemiology, [Rojas-Rueda et al. \(2019\)](#) conducts a meta-analysis of recent cohort studies linking NDVI to mortality.

afforestation policies. Besides Beijing, there are numerous examples of cities having been able to implement urban afforestation at mega scale, such as Los Angeles' Million Trees Initiative, New York City's Million Trees NYC initiative, Toronto's Every Tree Counts program, and Singapore's City in a Garden project. This level of implementation is noteworthy when considering urban afforestation within the context of environmental protection policies, which often encounter feasibility challenges, enforcement and compliance issues, and political resistance ([Gray and Shimshack, 2011](#); [Meng 2017](#); [Mastini et al., 2021](#); [Giles, 2022](#); [Browne et al., 2023](#)). We expect that similar urban afforestation programs and other forms of green infrastructure will play an increasingly important role in future sustainable urban development ([Thacker et al., 2019](#)).

Our analysis also provides among the first estimates on the impacts of pollen, adding to a small health economics literature on seasonal allergy and the economics costs in terms of cognitive outcomes such as test performance, crimes, and accidents ([Marcotte, 2015](#); [Bensnes, 2016](#); [Chalfin, Danagoulian, and Deza, 2019](#); [Akesaka and Shigeoka, 2022](#); [Danagoulian and Deza, 2024](#)).

On the method front, we join a growing econometric literature on treatment effect with unit interference ([Sävje, Aronow, and Hudgens, 2021](#)): although the MMP policy causes forests to be planted in a given area, the environmental (and thus health) effects of those forests are not confined locally to those areas. We leverage structural knowledge on the nature of the interference – that the spatial spillover effects are likely driven by wind transport – and build a wind directivity design to capture the phenomenon. We then combine this design with causal estimates on two other marginal effect estimates that can be obtained through local treatment effect estimation – the impact of policy on local vegetation growth, and the impact of local pollen on health – to deliver the overall causal effect of the program on the environment and health. Our analysis is facilitated by a comprehensive dataset we compiled combining government surveys, high-resolution remote sensing data, real-time ground monitoring, and administrative medical claims data, which allows us to directly estimate each link on the causal chain from policy implementation to endpoint health outcomes.

The rest of the paper is organized as follows: Section 2 discusses policy background and data. Section 3 outlines our empirical framework. Section 4 documents Beijing's greening up and the role of the urban afforestation policy. Section 5 studies environmental effects of urban forests. Section 6 studies health effects. Section 7 quantifies the benefits of the program by monetizing

health values and risks, and by empirically estimating housing market capitalization. These numbers are then compared to the program's costs. Section 8 concludes the paper.

## 2. Background and Data

### 2.1 Background

**A Brief Chronicle of Urban Afforestation Policies in Beijing.** The 1977 World Conference on Desertification Control in Nairobi, Kenya, identified Beijing as a city on the brink of desertification, raising serious ecological alarms to the Chinese government. In reaction, China inaugurated Arbor Day in 1979 and the National Capital's Obligatory Tree Planting Day in 1985, encouraging public participation in the country's green transformation; the country also introduced numerous initiatives in the two decades between 1991 and 2010 including the Three-North Shelter Forest program, the Taihang Mountain afforestation project, among others. Many cities followed suit with local programs focused on sand control, afforestation, water source protection, and the establishment of green corridors and green separation zones.

Beijing, the capital city, has been particularly proactive in this movement. This culminated when the city committed to a "Green Olympics" as the host of the 2008 Beijing Olympics, propelling major regional afforestation and ecological initiatives such as the Beijing-Tianjin Sand Source Control Project and the Beijing-Hebei Ecological Water Source Protection Forest Construction Project.

Historically, the plain area of Beijing, the hub for the city's population and industries (Figure A.3, areas with <100m altitude), has been ecologically under-resourced, with lower forest coverage compared to the mountainous regions. In an effort to remedy this, in 2012, the Beijing Municipal Government introduced the "Million Mu Afforestation Project in Plain Areas." We call this project the Million Mu Project (MMP) in this paper. This ambitious initiative aimed to augment the forest coverage in these plain areas by an additional one million mu (approximately 165,000 acres) over five years, with a goal to attain over 25% forest coverage rate. This marked a significant milestone in Beijing's ongoing journey of urban afforestation and ecological restoration.

**The Million Mu Project.** The MMP project was discussed in a January 2012 meeting of the Beijing Municipal Government as part of the *Beijing Municipal Air Pollution Control Plan 2012-2020*. The stated goal of the MMP project is to improve the ecological environment of the capital city, reduce PM<sub>2.5</sub> pollution, and promote green and ecological development. The MMP project involves both the conversion of existing construction land and the reclamation of abandoned sand and gravel pits for afforestation purposes. It focused on key areas such as ecologically sensitive zones, the peripheries of roads and rivers, water source protection areas, and agriculturally non-viable land for focused forest creation.

The spatial layout of these urban forests was planned in alignment with the broader urban and green space system planning. The guidelines are known as “Two Rings, Three Belts, Nine Wedges, and Multiple Corridors.”<sup>3</sup> Specifically, “Two Rings” refers to the creation of continuous green belts alongside the Fifth Ring Road, each extending 100 meters in width. These green belts act as the first line of ecological defense in the plain area. Further out, beyond the Sixth Ring Road, two additional green belts have been established – an outer one extending 1,000 meters in width and an inner one spanning 500 meters, jointly serving as the second ecological protection layer in the plain area. “Three Belts” denotes the establishment of permanent green belts along the banks of the Yongding River, Beiyun River, and Chaobai River, each stretching at least 200 meters wide. These belts function as vital ecological preservation zones. “Nine Wedges” is the creation of four functionally-defined, moderately-sized suburban parks within nine wedge-shaped areas of limited development. Together with large, contiguous forested areas, these parks provide crucial green spaces that bridge the urban core with the city’s outskirts. “Multiple Corridors” encompasses the development of green passages along key roads, riversides, and railways, as well as health-oriented greenways that interconnect different regional forest landscapes and park green spaces.

The MMP shapefiles we use in the analysis (Figure 2, for example) represent actual planting sites that fall under these forest creation guidelines. See Section 2.2 for more discussion.

**Program Financing.** Funding for the MMP construction was sourced from both city and sub-city (district) levels. Data published by the city government indicate that the initial investment during the first three years of the project (2012 to 2015) amounted to 34.3 billion CNY,

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<sup>3</sup> [http://yllhj.beijing.gov.cn/zwgk/cwgg/jbcwgg/202103/t20210319\\_2311459.html](http://yllhj.beijing.gov.cn/zwgk/cwgg/jbcwgg/202103/t20210319_2311459.html)

of which 25.5 billion CNY was contributed by the city government and the rest by district governments. Investment and construction policies varied across regions, but major cost items included purchasing fixed assets, building conservation areas, and performing maintenance and management duties post-planting. As the program expanded between 2015-2022, it incurred an additional cost of 40.6 billion CNY. We estimate the total recorded cost of the program to be around **75 billion CNY** (about 10 billion USD) which is about 1.8 percent of Beijing's annual GDP. We will compare this figure with potential health benefits in Section 7.

**Program Outcomes.** From 2012 to 2017, a total of 1.17 million *mu* (780,000 hectares) of afforestation and greening was completed in the city. Following the success of these efforts, a new phase of the million *mu* project was launched by the municipal government in late 2017 and concluded in 2022. This second phase of the project had a more extensive coverage, spanning the central urban area, new towns, and low mountainous regions. Through two consecutive five-year rounds of the MMP project, a total of **2.07 million mu (135,000 hectares)** of new afforestation and greening has been added to the city. In the plain areas of Beijing, the forest coverage rate increased from 15% in 2011 to 31% in 2022, while the overall forest coverage rate in the city improved from 38% to 45% during the same period.<sup>4</sup>

Due to the lack of data regarding the exact planting sites established in the first five years versus the second five-year phase of the project, our analysis considers the entire decade starting from 2011 as the MMP policy period.

**The Science of Urban Forests and Pollution.** Vegetation leaves reduce pollution through a process known as phytoremediation. This happens in two main ways: First, small openings on leaves (stomata) absorb trace gas pollutants such as nitrogen dioxide and sulfur dioxide (Harris and Manning, 2010; Yin et al., 2011); Second, leaves "filter out" particle pollutants through dry deposition (McDonald et al., 2007; Nowak, Crane, and Sevens, 2006). In atmospheric science and urban planning literature, forests and urban greenery are widely recognized as an effective tool for pollution reduction (e.g., Baldauf, 2017; Kumar et al., 2019).

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<sup>4</sup> Some local surveys indicate that tree mortality rates in MMP areas could be as high as 50% (e.g., <https://news.sciencenet.cn/htmlnews/2014/7/299222.shtml>). In both our environmental and health impact analyses, rather than exploiting variation based on the intended / nominal afforestation from government-reported planting records, we look at realized vegetation coverage as measured by satellite-based NDVI. Hence implicitly, we are evaluating surviving trees in our estimation.

There is one notable exception: the implications for ground-level ozone, where forests and urban greenery can have both positive and negative effects. On one hand, planting trees can increase the rate of ozone deposition and absorption, reducing near-ground ozone concentrations. On the other hand, certain tree species, such as poplar, willow, and oak, release volatile organic compounds (VOCs) to repel pests and attract pollinating insects. These VOCs are precursors to ozone formation – that is, they react with nitrate compounds and sunlight to form ozone. Our empirical results below indeed show that the impact of urban forests on ozone is not as clearcut as their impacts on other criteria air pollutants.

**Clinical Evidence on Pollen and Health.** In Section 6, we econometrically quantifies the link between ambient pollen concentration and emergency room visits. Here we review some of the clinical foundation of that link.

Exposure to allergenic pollen has become an increasingly concerning environmental health issue in urban areas in recent decades ([Biedermann et al., 2019](#); [D'Amato et al., 2015](#)). The prevalence rates of seasonal allergies caused by pollen range from 10% to 40% in developed countries, with an estimated 400 million sufferers worldwide ([Greiner et al., 2011](#); [Meltzer et al., 2012](#)). The rise in allergies is particularly concerning in the context of global warming, as it is expected to prolong the plant growing season and increase the overall pollen production per season ([Ziska et al., 2019](#)).

Inhalation of airborne pollen can lead to seasonal allergies, often referred to as hay fever or pollinosis. This condition is a widespread chronic issue and a global health concern. Unlike the year-round threat posed by air pollution, such as PM<sub>2.5</sub>, which impacts the respiratory and cardiovascular health of a large portion of the population, exposure to pollen exhibits strong seasonal variations and differentially affects individuals with allergy histories. Allergic reactions are triggered by the immune system. When a person who is allergic to substances like dust, mold, or pollen encounters these substances, their immune system might react excessively, producing antibodies that attack the allergen aggressively. A number of typical allergic reactions are linked with the production of a specific antibody known as immunoglobulin E (IgE) by the body. Allergens can be introduced into the body through inhalation, consumption, or contact with the skin.

Symptoms of seasonal allergies caused by pollen exposure involves the sensory system, including allergic rhinitis (sneezing, runny and stuffy nose) and allergic conjunctivitis (itchy eyes and tears). In rare cases, asthma and atopy may also occur (Sun et al., 2016). Severe allergic reactions can lead to bronchitis, bronchial asthma, pulmonary heart disease, and even life-threatening situations (Brunekreef et al., 2000). Pollen has a particularly noticeable impact on individuals with respiratory allergies, which affect approximately 10-30% of the global population (Sierra-Heredia et al., 2018). Seasonal allergies not only worsen physical and mental health but also decrease productivity, increase medical expenses, and reduce daily activities, thereby impacting people's quality of life. Higher pollen counts lead to increased visits to asthma emergency departments and more sales of over-the-counter allergy medications. Allergy medications used to alleviate symptoms can have side effects such as drowsiness, dry mouth, lethargy (Jáuregui et al., 2009; Meltzer et al., 2012), which may negatively impact cognitive performance and productivity.

Different types of pollen vary in allergenicity, which depends on the strength of allergenic pollen antigens and the pollen concentration. In Beijing, herbaceous plants like *Artemisia* exhibit relatively high allergenicity but low pollen levels during the autumn season, while deciduous trees like *Cupressaceae* have relatively weak allergenicity but high pollen levels during the spring season.

In the 1980s, the total amount of airborne pollen, primarily from herbaceous plants, was higher during summer and autumn than during the spring peak period. Since 2000, the pollen content from deciduous tree species has been increasing each spring, making spring the season with the largest proportion of the total annual pollen content. This shift became even more pronounced since 2010, with a significant increase in pollen from spring-flowering deciduous trees like *Ginkgo*, *Platanus*, and *Cupressaceae* (Zhao et al., 2021). Currently, the predominant pollen types in Beijing are from cypress and poplar trees. Chinese juniper pollen has also become an important allergen for spring pollen allergies, with the peak allergy season running from March to May.<sup>5</sup>

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<sup>5</sup> See [http://bj.cma.gov.cn/xwzx/mtjj/202103/t20210326\\_3023644.html](http://bj.cma.gov.cn/xwzx/mtjj/202103/t20210326_3023644.html)

## 2.2 Data Sources

This project's data sources are tabulated in Table 1. Here we provide more details about each source.

**Planting Sites.** We get location information on the universe of MMP planting sites from a policy document published in 2022 by the Beijing Municipal Commission of Development and Reform. To be clear, the map shows where trees and other greenery ended up being planted (rather than where planting was planned).<sup>6</sup> We digitize and geo-reference these maps to create polygon files that represent locations of MMP sites. The total area of MMP sites according to our digitized data is 2.38 million *mu* (about 1,589 km<sup>2</sup>), which is close to the official number of 2.07 million *mu* (1,380 km<sup>2</sup>). Our analysis Section 4 provides further cross-reference checks using remote-sensing based vegetation index data, showing sharp increase in vegetation growth immediately starting the MMP boundaries.<sup>7</sup>

**Land Use.** We obtain land-use information from China's Land-Use/Cover Datasets (CLUD) from the Institute of Geographic Sciences and Natural Resources Research at the Chinese Academy of Sciences. CLUD are mainly based on Landsat images, and generated using a human-computer interaction (HCI) interpretation process. The data are available at 30-meter spatial resolution and five-year intervals, and we obtain data for the 2000, 2010, and 2020 cross sections. Our analysis uses level-1 classifications which contain six categories: cropland, forest, grassland, water bodies, construction land, and unutilized land. CLUD's level-1 classification is estimated to have an accuracy rate of over 94 percent ([Yang and Huang, 2021](#)). We use this data to characterize the type of places where MMP plantings areas were sited.

**Population.** We use population estimates data for the city of Beijing from the WorldPop 100-meter resolution product for year 2020. WorldPop combines satellite data, census, and machine learning to predict population distribution at fine spatial scale. We use WorldPop mainly to demonstrate the location of population hubs in the city.

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<sup>6</sup> *14th Five-Year Period Land Resources Protection and Utilization Plan in Beijing* (京政发 [2022] 26号: 北京市“十四五”时期土地资源保护利用规划, last accessed January 5<sup>th</sup>, 2023.) This document outlines land resources protection and utilization planning for the five-year period of 2021-2025. The document provides MMP maps as a part of its summary of achievement in the previous decade.

<sup>7</sup> To alleviate further concerns about measurement errors, particularly those associated with the digitization of small forest patches, in Appendix Figure A.4 we report a robustness check where we restrict our analysis to planting areas that exceed 1km<sup>2</sup> in size.

**Vegetation Index.** Normalized Difference Vegetation Index (NDVI) is a widely-used remote-sensing measure of the density of vegetation (Pettorelli et al., 2005). NDVI is a function of the reflectance of two specific wavelengths of light: near-infrared (NIR) and visible red (RED). Healthy vegetation reflects a large amount of NIR light and absorbs most of the RED light. In contrast, non-vegetated surfaces, such as bare soil or water, reflect less NIR light and more RED light. The formula for NDVI for a given location  $i$  is:

$$\text{NDVI}_i = \frac{\text{NIR}_i - \text{RED}_i}{\text{NIR}_i + \text{RED}_i}$$

The formula normalizes the values to a range of -1 to 1, with higher values indicating more vegetation. Negative NDVI values occur when an area reflects more red light than near-infrared light. This mostly occurs for water bodies in our sample, which we drop from our analysis.

The NDVI data are obtained from NASA Moderate Resolution Imaging Spectroradiometer (MODIS) product MOD13Q1.061, which provides 16-day composite of vegetation indices at a spatial resolution of 250 meters. We obtain data from January 2001 to December 2020 for the city of Beijing, and we aggregate up the temporal frequency from 16-day to monthly.

**Remote-Sensed Air Pollution.** A part of our analysis uses a satellite measure of Aerosol Optical Depth (AOD) to measure air quality. The data are from MODIS Multi-Angle Implementation of Atmospheric Correction (MAIAC) algorithm which provides AOD at daily frequency and 1-km spatial resolution from 2001-2020 (product MCD19A2.006; Lyapustin, 2018). AOD is an index measure of sunlight scattering and absorbance, which provides a good proxy for the concentration of particulate pollution in the atmosphere. A downside of AOD is that it is technically a measure of column (ground to top-of-atmosphere) pollution instead of ground-level pollution. The upside is that, being a remote-sensing measurement, the data allow us to look at pollution changes where in-situ monitoring data are not available, such as places near or inside planting sites.

**In-Situ Air Pollution.** Our main air pollution analysis uses data from 35 ground-level monitoring stations in the city from 2014 to 2019. Each station performs real-time monitoring of six criteria pollutants: fine particulate matter (PM<sub>2.5</sub>), coarse particulate matter (PM<sub>10</sub>), ozone (O<sub>3</sub>), sulfur dioxide (SO<sub>2</sub>), nitrogen dioxide (NO<sub>2</sub>), and carbon monoxide (CO). Our data are at the

station-daily level. The data are sourced from the Beijing Municipal Ecological and Environmental Monitoring Center.<sup>8</sup>

**Pollen Counts.** We have access to four years (2013-2016) of daily pollen counts data from 20 stations from the Beijing Meteorological Service Center. The pollen monitoring season spans from March 1 to September 30 each year, with some station-years extending to October 15. Pollen counts are measured using a filter-based method in which the pollen deposited on a filter is observed and the concentration is calculated by dividing the counts by the amount of air that passes through (with the unit of measurement being counts per 1,000 mm<sup>3</sup> of air).

Each station monitors six pollen species: *Cupressaceae* (Cypress), *Salicaceae* (Willow), *Pinaceae* (Pine), *Moraceae* (Mulberry), *Artemisia* (Sagebrush/Wormwood), *Chenopodiaceae* (Goosefoot). In our main analysis, we define pollen count as the sum of pollen counts across all six species. We also provide disaggregated analysis in Section 6 to examine species-specific health effect estimates, and compare those to clinical evidence. For example, pine tree pollen allergy is known to be relatively uncommon, whereas other species such as cypress tree pollen are known to be more allergenic ([Charpin et al., 2005](#); [AAAI, 2018](#)).

A part of our analysis in Section 6 compares the ER impacts of ambient PM<sub>2.5</sub> and pollen. One technical issue worth addressing head-on: because pollens are also particles, are they picked up by PM<sub>2.5</sub> monitors as well? Most pollen grain diameters range from about 10 to 100 micrometers, which are too large to be picked up by PM<sub>2.5</sub> monitors which use a size-selective inlet to only allow particles of 2.5 micrometers in diameters or smaller to pass through the inlet before the monitor measures particle concentration. This is often realized using a cyclone inlet design where air sample is forced through a spiral pathway, during which the larger particles are left out because they cannot change direction as quickly as smaller particles due to greater inertia. Most pollen particles will be left out at this stage.

**Wind.** We obtain wind direction data from 20 weather stations from 2014-2019 maintained by the Institute of Geographic Sciences and Natural Resources Research at the Chinese Academy

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<sup>8</sup> In 2013, China implemented a pollution-monitoring reform which replaces its old monitoring system that was both sparse and shown to have reliability issues ([Ghanem and Zhang, 2014](#)). The post-reform system adopts real-time monitoring technologies used by the U.S. EPA, and the data are believed to be much more reliable ([Greenstone et al., 2022](#)).

of Sciences. For each air pollution and pollen monitoring station, we assign daily wind direction based on what was recorded at that station's nearest weather station.

**Emergency Room Visits.** To measure ER uses, we use anonymized medical claims data maintained by the Beijing Municipal Medical Insurance Bureau. This data contains the universe of institutional (hospital) medical service transactions for all city residents enrolled in the Urban Employee Basic Medical Insurance, a state-run mandatory medical insurance program that covers all urban residents employed in the formal sector. The data covers about 91 percent of all permanent residents of Beijing, and we have access to four years of data from 2013 to 2016.

It is worth noting that our claims data are *insurance* records maintained for reimbursement purposes, rather than hospital *medical* records. Therefore, although the data allows us to observe most information about an ER visit such as patient demographics, hospital location, date, and spending, it does not contain primary diagnosis made by the ER attending physician.<sup>9</sup>

But identifying the medical condition underlying the ER visit is important for this research. We do so by leveraging information on the type of medications associated with the ER visit, including drugs administered during the visit and prescribed medications for post-care use. Our data cover more than 20,000 distinct medications. We categorize these into one of the 14 disease treatment categories (such as respiratory, sensory, among others) using the Beijing Reimbursement Drug List for Basic Medical Insurance crosswalk file. In principle, an ER visit may involve medications from multiple categories. However, for simplicity, we define three primary categories: respiratory ER visits (involving any respiratory drugs), sensory ER visits (involving any sensory drugs), and other ER visits (not involving any respiratory or sensory drugs). Appendix Table A.1 tabulates the top ten most frequently prescribed medications in our respiratory and sensory ER categories.

**Housing Transactions.** We draw housing transaction records of new and resale properties from the Beijing Municipal Commission of Housing and Urban-Rural Development and Lianjia, China's leading online platform for real estate information and transactions. This dataset, first constructed by [Jerch et al. \(2023\)](#), contains 851,741 transactions from 2006 to 2016, representing about half of all home purchases in Beijing throughout our study period. The dataset includes

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<sup>9</sup> We likewise do not observe patient death, which is a commonly studied outcome in the environmental economics literature.

details for each transaction, such as the latitude and longitude of the apartment unit, the transaction price, and characteristics of both the apartment unit and its parent complex. Following [Jerch et al. \(2023\)](#), we remove transactions with missing or zero prices, and those in the lowest and highest 1 percent range of price per square meter.

Table 1 summarizes data sources and where they are primarily used in the paper. Generally speaking, the remote-sensing data are used in Section 4 to document the effect of MMP planting on local vegetation growth and air quality change. The monitor data are used in Section 5 to document the impact of MMP planting on air quality and pollen changes at population hubs. Both the monitor and the medical claims data are used in Section 6 to estimate health implications. Housing transactions data are used in Section 7 to look at capitalization effects.

### 3. Research Framework

Before proceeding, we outline the overall structure of the empirical analysis, and clarify several decisions we have made and why. The goal of this paper is to study the environmental and health effects of urban forests. We organize our econometric exercises around two conceptual equations. The first is an environment equation, which links the afforestation policy (MMP) to changes in city greenery, and to changes in environmental agents (ambient air pollutants and pollen):

$$\Delta \text{Agent}_i = \frac{\partial \text{Green}}{\partial \text{Policy}} \cdot \frac{\partial \text{Agent}_i}{\partial \text{Green}} \quad (1)$$

The second is a health equation that quantifies the health consequences of the environmental changes:

$$\Delta \text{Health} = \sum_i \frac{\partial \text{Health}}{\partial \text{Agent}_i} \cdot \Delta \text{Agent}_i \quad (2)$$

Our empirical analysis revolves around econometrically identifying the three set of causal parameters highlighted in brackets in these equations. These parameters are the focus of Section 4, 5, and 6, respectively. Several features of this framework worth discussing:

A notable feature of our framework is that it takes a *marginal* approach: instead of directly estimating the causal effect of the MMP policy on environmental conditions and on health in one pass, we estimate three separate links on the causal chain (Policy → Greenery → Environment → Health). The primary motivation underlying this design is the violation of the Stable Unit Treatment Value Assumption (SUTVA) in our study context. Whereas the MMP policy causes forests to be planted in a given area, the environmental (and thus health) effects of those forests are not confined locally to those areas. For example, a location's air quality is not only influenced by the presence of greenery in that exact location, but the prevailing wind direction and how much greenery is present in the upwind direction that altered the air before it reaches the location. Conventional instrumental variables model that assumes no cross-unit interference – that is, a study unit's outcome depends *only* on policy treatment status of the unit itself – is not applicable in our context when studying environmental and health changes.

Our idea of addressing this issue is to produce treatment effect estimates in separation, and in each part, we use the most appropriate design to causally estimate the corresponding marginalized effect, given the available data. Though identified from different sources of variations, these estimates are likely compatible as they pertain to nearly the same population basis (e.g., monitors in population hubs; near-universal hospital claims coverage) and are often estimated using variation and specifications of similar nature (e.g., using day-to-day fluctuations in environmental conditions with high-dimensional fixed effects in regression specifications).

Section 4 is the most straightforward, where we combine government-provided maps of MMP planting area and satellite data on vegetation index to document the impact of planting on local greenery growth.

Section 5 is where we need to address the SUTVA violation: the environmental condition at any given location can be influenced by nearby MMP forests. Two features help us build the research design: First, we know the nature of unit interference is likely driven by wind transport. This gives us structural knowledge of where spatial spillover comes from, and we use a up/downwind directivity design – one that is in fact quite familiar with environmental economists – to capture that. Another helpful feature is that, most city population cluster in a small number of areas, and those areas all have ground-based monitors for both pollution and pollen. To the extent that we primarily care about changes in environmental agents experienced

by the typical citizen, we only need to run the estimation at the monitor level, which makes the task computationally feasible.

Once we quantified how local environmental conditions changed due to MMP, Section 6 is a standard environmental-health analysis, linking local variation in environmental conditions with health outcomes. Our medical claims data allow us to build healthcare use variables at the district-by-day level, which we merge with air pollution and pollen data to conduct a standard panel data analysis.

Our main contribution using the economic approach is to provide a unified empirical framework that traces – using research designs as close to causal as possible – the full chain of links from a policy initiative, to urban vegetation growth, to air pollution reduction and pollen changes, and ultimately to endpoint health outcomes. In Appendix Table A.2, we provide an additional survey of existing studies examining individual links in the atmospheric science, ecology, or medical literatures. However, because these studies come from different disciplines, they often use very different methods, study different contexts, and focus on different underlying populations. As a result, it is difficult to simply piece together those estimates into a coherent and credible policy evaluation. In contrast, we assemble data that represent the same underlying population, and we adopt research designs that are methodologically compatible with each other. This coherence allows us to combine the estimates in a meaningful way and directly inform the policy question.

## 4. The Greening Up

### 4.1 Raw Statistics

We begin by characterizing Beijing's greening up using satellite NDVI data. For each 250m-by-250m grid, we calculate a “rate of vegetation growth” variation defined as the annual change in NDVI of the grid, i.e., the slope of a linear time trend fit across annual NDVI observations of that grid. We do so separately for the pre-MMP period (2001-2011) and the post-MMP period (2012-2020).

Figure 1 plots the results for the pre-MMP period (panel a) and the post-MMP period (panel b). In each panel, we provide a whole-city view on the left, and a zoomed-in, city-center

view on the right. In the city-center view, we overlay a map of major road networks to provide a sense of the geography of economic activities.

These maps show that the city was experiencing vegetation growth even before 2012, but the growth has accelerated after 2012. The magnitude of this acceleration is substantial. Panel (b) of Figure 1 shows numerous hotspots (grids with blue color) popping up with an annual growth rate of over 0.02 units of NDVI per year. Based on a sample average NDVI of 0.34 in year 2010, these hotspots have experienced an increase of more than 53% over the period of 2012-2020.

In the Online Appendix, we provide more companion statistics on the greening up. Appendix Figure A.1, panel (a) shows annual NDVI grew by 11.7% between 2001 and 2011, and by another 16.7% between 2012-2020. The speed of growth is therefore 28% higher over the latter period. Figure A.2, panel (a) shows satellite-based land use categorization data that are available in years 2000, 2010, and 2020. Consistent with acceleration of vegetation growth, the land use data suggests a clear deacceleration of urbanization after 2010.

## 4.2 The Role of the Million *Mu* Project

In Figure 2, we use polygons to highlight areas corresponding to the Million *Mu* Project (MMP) planting areas, and then we overlay the map with NDVI growth rate 2012-2020. Panel (a) provides a whole-city view, and on the right hand side of the panel we zoom in two six example areas with large MMP patches. Panel (b) once again provides the city-center view. The results suggest a high spatial correspondence between MMP planting areas and post-MMP greening up.

Next, we provide more systematic evidence on this spatial correspondence. For each grid, we calculate its distance to the nearest MMP planting area boundary. Because many planting areas are large and contain multiple grids, for those grids that fall within the MMP boundary, we assign them negative distances. About 16.3% of grids fall within MMP planting areas. Among all grids that are outside of the MMP areas, about 90% are within 2 km to the nearest MMP boundary. Appendix Figure A.2, panel (b) shows the distribution of grids by distance. We then estimate the relationship between NDVI growth and the grid's distance to MMP area by fitting the following regression equation:

$$\text{NDVI growth}_i = \sum_j \beta_j \cdot 1(\text{distance bin} = j)_i + \varepsilon_i \quad (3)$$

where  $NDVI_{growth_i}$  is rate of NDVI growth for grid  $i$  over either the 2001-2011 or 2012-2020 period,  $1(distance\ bin = j)_i$  is a series of dummies indicating bins of distance-to-MMP-areas in 200-meter increment spanning -1km to 2km omitting the furthest bin 1.8km to 2km, and  $\varepsilon_i$  is the error term. We are interested in the  $\beta_j$  estimates, which shows the difference in NDVI growth rate of grids in a given distance bin  $j$  relative to grids in the reference bin 1.8km to 2km. We report confidence intervals calculated from 1-km gridded cluster bootstrap standard errors.<sup>10</sup> Note that equation (3) is simply a non-parametric expression of NDVI growth and distance-to-MMP.

The city of Beijing comprises the plain area in the southeast where the vast majority of the people live, and the mountainous area in the northwest. Appendix Figures A.2 and A.3 provide illustration. Given our objective is to measure the impact of the policy on citizens' experiences, our analysis will focus on the plain area, which we define as grids with elevation lower than 100 meters. This area contains 94 percent of the city's population. We report robustness checks using alternative elevation cutoffs, such as 50 meters or 200 meters, in Appendix Figure A.4.

Figure 3, panel (a) shows the results. The solid line, showing data from the post-MMP period, shows a sharp gradient of NDVI growth with respect to distance. This suggests the spatial correspondence between NDVI growth and MMP polygon that one can eyeball from Figure 2 before indeed holds true systematically in the data. The satellite measure detects differential change (that is, faster NDVI growth relative to the reference bin of 1.8-2km) outside of the MMP boundary up to 400 meters, which can either reflect errors evolved in government-provided MMP boundary map or in our digitization. However, it is reassuring that overall the MMP boundary corresponds to NDVI growth very well.

One potential concern when interpreting the distance gradient as the causal effect of the MMP policy is whether the MMP areas are more likely to be located in places that would have experienced faster NDVI growth anyway.<sup>11</sup> To speak to this concern, panel (a) of Figure 3 also shows the NDVI growth-distance relationship for the *pre-MMP* period, where no sharp change

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<sup>10</sup> That is, we project all individual grids  $i$ 's onto a 1-km gridded map of the city. We call each 1-km grid a "cluster", and calculate standard errors using a cluster bootstrap. We go with this approach because other commonly-used methods, such as Conley standard errors, took infeasibly long to compute given the size of our data (over 120,000 grids).

<sup>11</sup> For example, suppose MMP areas are disproportionately located in high-vegetation regions, and high-vegetation regions tend to feature higher speed of vegetation growth, then the observed relationship may not purely reflect the effect of the MMP policy but rather just site selection.

in NDVI growth is observed around MMP boundary. In Appendix Figure A.2, panel (b), we show that land use categorization at baseline (year 2010) also change smoothly around MMP boundary. Together, the evidence suggests that the sharp, nonlinear increase in NDVI growth around MMP boundary, and only for the post-MMP period, is likely a consequence of the MMP policy.

### 4.3 Change in Air Quality near MMP Planting Areas

The estimation framework of equation (3) gives us an opportunity to take an initial look into the effect of vegetation growth on air quality. We repeat equation (3) but replace the outcome variable with the grid's growth rate of aerosol pollution – in fact, it is the rate of *decline* as most of the grids saw decline of pollution during the 2012-2020 period ([Greenstone et al., 2021](#)).

Figure 3, panel (b) shows a mirror image of panel (a), where the rate of decline of pollution is much faster within MMP areas. The fact that the gradient of pollution change starts almost exactly where the gradient of NDVI change starts adds to the confidence of the causal effect of MMP. Once again, no similar pattern is shown for the pre-MMP period of 2001-2011.

Three points worth noting. First, we have made two causal claims. The first causal claim is that plantation under MMP causes NDVI growth to accelerate after 2012. The second causal claim is that increased vegetation growth causes the reduction in air pollution near the planting area. As we discussed above, these claims are supported by the fact that, spatially, NDVI and pollution gradient changes exactly where MMP areas start, and, temporally, such relationships are only observed after MMP was in place.

One remaining confounder in making the latter causal claim (vegetation causes pollution decline) which we have not addressed is the fact that both the NDVI and aerosol pollution data are drawn from satellite sources, which are ultimately derived from shared underlying data. So suppose for some reason it is easier to detect pollution change in areas with faster NDVI change, then panel (b) of Figure 3 could reflect a mechanical relationship due to natures of satellite detection, rather than genuine changes in pollution. We are not aware of a clear pathway that can lead to such mechanical relationship, though we are not able to fully rule it out either. However, our analysis in Section 5 using ground level pollution monitoring data will not be subject to this concern.

Second, the graphical pattern of panel (b) of Figure 3 appears to suggest a lack of effect on NDVI or pollution beyond the MMP boundary. But notice that equation (3) can only speak to changes in the outcome *relative to* the baseline group. In other words, while it is possible that MMP vegetation has spillover effect on air quality beyond the MMP boundary – for example, through wind transport – equation (3) is not suited to estimate that “general equilibrium” effect. We will directly estimate spillover effect in Section 5 next.

Third, in terms of effect sizes, panel (a) of Figure 3 suggests MMP grids grew about 8.1 percentage points faster than the growth rate of grids outside MMP; panel (b) suggests the rate of pollution reduction is 1.0 percentage point faster for MMP grids than non-MMP grids. This implies a long-term growth rate elasticity between NDVI and pollution of -0.125. We will show later that our short-term estimate of NDVI-pollution elasticity at the daily level using ground-based monitor data has larger but same order-of-magnitude elasticity (about -0.40).

## 5. Environmental Effects

Section 4 documents the substantial greening up of the city and demonstrates the role of the MMP policy. A limitation of the analysis, as we briefly discussed, is that the cross-sectional approach focused on fine-grained comparison of environmental conditions in areas in- and outside of MMP planting areas, and is not suited to identify the potential spillover effect of vegetation growth on the broader population, which might have missed the bulk of the policy. We estimate that about 1.3 million people live within MMP planting areas, while the city’s population is 20 million people. The goal of this section is to estimate the causal effect of the vegetation growth on environmental conditions felt by the “typical resident”, who lives in city population centers.

### 5.1 Estimation

The key insight of the empirical analysis of this section is this: to estimate environmental changes experienced by the typical resident, we will simply use air quality monitor as a unit of analysis, and estimate how changes in vegetation near the monitor has caused a response in air

quality values as captured by the monitor. There are two main rationales behind this decision.<sup>12</sup> First, monitors are installed by the government to record population exposure to environmental conditions, and therefore most monitors are sited in areas with high population density. In panel (a) of Figure 4, we overlay the location of air pollution (or pollen) monitors on a map of grid-level population estimates. Evidently, monitors are placed near most densely populated areas. Conducting analysis at the monitor level therefore will give rise to estimates that reflect conditions that are representative to population exposure. Second, monitors capture environmental conditions that happen at the ground level and therefore more accurately reflect what people experience. This is in contrast with satellite data of pollution, which measures pollution concentration for the entire ground-to-top-of-atmospheric column of air. Using data directly collected by the monitors also helps avoid potential mechanical link between our satellite-based vegetation measure and pollution measure, which we discussed earlier in Section 4.

Recall that the objective of our analysis is to speak to how MMP-led vegetation growth may cause changes in air quality outside of the immediate planting area. In other words, for a given monitor, we are interested in not only how vegetation condition at that exact location matters for air quality recording of the monitor, but also potentially the impact of vegetation condition farther away. The particular mechanism we have in mind for this to render is through wind transport. To see this, let  $P_{it}$  denote the air quality condition (pollution or pollen) captured by monitor  $i$  on date  $t$ . A monitor can capture air quality because there is air movement, and therefore  $P_{it}$  is a measure of the (unobservable) true air quality condition upwind of the monitor, which we denote  $P_{it}^U$ . Hence, one can write:

$$P_{it} = \gamma^U \cdot P_{it}^U + \varepsilon_{it} \quad (4)$$

where  $\gamma^U$  is a positive coefficient, and  $\varepsilon_{it}$  is the error term. Now consider the role of (observable) vegetation in the upwind area, denoted  $Green_{it}^U$ , which we assume interact with pollution in the area following:

$$P_{it}^U = \tau \cdot Green_{it}^U + \epsilon_{it} \quad (5)$$

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<sup>12</sup> For the sake of brevity, we will henceforth use “air quality” to refer to either air pollution concentration or pollen counts. We will be more specific which one we refer to when necessary.

where  $\tau$  is an emission factor. For pollution, the emission factor of greenery is negative (vegetation absorbs or filters out pollution); for pollen, the emission factor is positive (vegetation emits pollen).  $e_{it}$  captures remaining randomness in the relationships. Substitute (5) into (4), one gets an equation that relates air quality to vegetation condition in the upwind area:

$$P_{it} = \gamma^U \tau \cdot \text{Green}_{it}^U + e_{it} \quad (6)$$

This is the equation we take to the data. Specifically, we estimate:

$$\text{Log Air Quality}_{it} = \beta^U \cdot \text{Log NDVI}_{it}^U + \alpha_i + \alpha_t + e_{it} \quad (7)$$

In this equation,  $\text{Log Air Quality}_{it}$  is logged air quality measured at monitor  $i$  on date  $t$ .  $\text{Log NDVI}_{it}^U$  is logged NDVI at the “upwind area,” defined as a 135-degree cone of a 10km-radius area in the upwind direction of the monitor. See panel (b) of Figure 4 for an illustration. In practice, because NDVI is available at the monthly frequency, the upwind NDVI variable is constructed using daily wind direction data but monthly NDVI value. For simplicity, in our primary specification, we calculate upwind NDVI using all grids that fall within the upwind cone. In the appendix, we report an alternative version where we restrict only to MMP grids within the upwind cone.<sup>13</sup> Our results turn out to be more precise if we use MMP grids only. However, we stick with all grids because we do not see a strong reason to assume that local air quality changes are driven only by vegetation growth in the MMP boundary.  $\alpha_i$  and  $\alpha_t$  are unit and time fixed effects. Our baseline model uses monitor, year, month, and day-of-week fixed effects.  $e_{it}$  is an error term, and we cluster standard errors two-way at both the monitor and the date levels.<sup>14</sup>

The coefficient of interest is  $\beta^U$ . Following our previous discussion, we expect this coefficient to be negative for air pollution, and positive for pollen.

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<sup>13</sup> Our cone-based approach implicitly assumes a Gaussian process for pollution dispersion, in which atmospheric conditions such as wind speed, wind direction, and atmospheric stability at each source location are treated as locally constant over time and space during the release and dispersion. The Gaussian framework is commonly used to model pollution dispersion at small, neighborhood-level geographic scales (e.g., [Jones et al., 2013](#)). This contrasts with large-scale applications, such as modeling long-range pollution transport across city, state, or national borders, where Lagrangian models like HYSPLIT are more appropriate ([Hernandez-Cortes and Meng, 2023](#); [Heo, Ito, and Kotamarthi, 2025](#)).

<sup>14</sup> There is abundant day-to-day variation in wind directions in our study setting. Appendix Figure A.5 shows variation in wind direction, both across seasons and from day to day within a season.

We augment the baseline estimation equation (7) in the following ways to assess its validity and robustness. First, we estimate a version of equation (7) adding 20 leads and 20 lags of  $\text{Log NDVI}_{it}^U$ , giving rise to an event study representation of the impact of vegetation shocks on air quality outcomes. To be precise, we estimate:

$$\text{Log Air Quality}_{it} = \sum_{k \in [-20, 20]} \beta_k^U \cdot \text{Log NDVI}_{i(t+k)}^U + \alpha_i + \alpha_t + e_{it} \quad (8)$$

and we plot the  $\{\beta_k^U\}_{k \in [-20, 20]}$  coefficient estimates. Second, we conduct a placebo test, replacing the upwind shock variable to measured vegetation at the *downwind* direction of the monitor, which is expected to have no impact on air quality readings capture at the monitor's location.<sup>15</sup> See panel (b) of Figure 4 for an illustration. Third, we report sensitivity checks along three dimensions where we had to make some arbitrary specification decisions: the radian of the cone that defines upwind and downwind area, the radius around the monitor where we calculate vegetation exposure, and fixed effects choices. We will give more details to these sensitivity checks when we come to discuss those results.

In addition to these checks, we further examine the potential for confounding factors that may have coincided with the afforestation program that even the upwind / downwind strategy cannot fully address. A particular example is correlated emission reduction: Suppose factories near MMP sites were subject to disproportionate regulation (or even relocation) at the same time that trees were planted, our design would not be able to fully disentangle the two effects because both the siting and the timing of emission reductions and tree planting would be highly correlated. We provide several empirical tests in the Online Appendix. First, using China's Environmental Survey and Reporting (ESR) database that cover 85% of city's total industrial emissions between 2001-2014, we show that the share of major polluting plants within the MMP area remained largely stable over time (Appendix Figure A.6, panel a). Second, using China's continuous emissions monitoring systems (CEMS) data from 2014-2020, we show that no statistically significant difference in emissions between plants located within and outside the

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<sup>15</sup> This is, of course, subject to potential autocorrelation of NDVI at the upwind and downwind direction. In practice, our results are similar if we include both upwind and downwind NDVI in the same specification, that is:

$$\text{Log Air Quality}_{it} = \beta^U \cdot \text{Log NDVI}_{it}^U + \beta^D \cdot \text{Log NDVI}_{it}^D + \alpha_i + \alpha_t + e_{it} \quad (9)$$

where  $\text{Log NDVI}_{it}^D$  is logged NDVI at the downwind area. We report these results in Appendix Table A.4.

MMP area in any year, nor do we observe any significant trends over time (Appendix Figure A.6, panel b).<sup>16</sup> Both these patterns suggest that the MMP policy, both in its initial siting and its subsequent implementation, does not appear to be correlated with any significant redistribution of industrial activities.

We further test sensitivity of our main pollution (and pollen) regressions to the inclusion of fixed effects for citywide or monitor-specific wind direction bins. Specifically, based on each day's prevailing wind direction, we assign a wind direction bin corresponding to a 135-degree cone (consistent with how we define upwind in the baseline specification), and include these fixed effects in equation (9). Wind direction fixed effects are highly predictive of pollution exposure from stationary sources (Deryugina et al., 2019). For example, if most polluting activities are located to the north of a city, a northerly wind will result in higher pollution concentrations at the monitor compared to winds from other directions. Therefore, if there are indeed pollution regulations that differentially target industrial emission reduction near the MMP sites, the inclusion of wind-direction fixed effects would attenuate the estimated NDVI effect. By contrast, if plantation sites and pollution source changes are uncorrelated, then after controlling for general wind-direction effects, it remains possible to isolate the causal impact of upwind afforestation. Appendix Table A.6 shows that including wind-direction controls slightly reduces the magnitude and precision of the estimates. But overall, the evidence indicates that confounding from changes in emissions from fixed sources is unlikely to be driving our results.

## 5.2 Results

Figure 5 summarizes the main results using equation (8). To recap, the structure of the underlying data is a daily panel dataset of pollution or pollen monitors. On the horizontal axis, we order coefficients so that negative event days are coefficients of the lead terms of  $\text{Log NDVI}_{it}^U$ ,

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<sup>16</sup> Specifically, we estimate the following regression models:

$$Y_{ijp} = \beta \times MMP_i + \delta_j + \theta_p + \varepsilon_{ijp}$$

where  $Y_{ijp}$  is the log of emission of pollutant  $p$  from unit  $j$  of plant  $i$ .  $MMP_i$  equals 1 if the plant falls inside MMP planting areas, and 0 otherwise.  $\delta_j$  and  $\theta_p$  are unit fixed effects and pollutant fixed effects, respectively.  $\beta$  represents the cross-sectional difference between plants located in the MMP area and those outside the MMP area. We estimate  $\beta$ 's for each year of our data, from 2014 to 2020. Standard errors are clustered at the plant level.

positive event days are coefficients of the lag terms, and event day zero represents the day-of term – hence the axis title “Days since shock.” We plot the coefficient estimates of  $\{\beta_k^U\}$  and the associated 95% confidence intervals. On the same chart, we overlay the placebo test where we run the exact same regression, but using *downwind* NDVI as the right-hand-side variables instead.

Panel (a) shows results for fine particulate matter (PM<sub>2.5</sub>), which shows a decline of pollution of about 0.5 log units for two days upon a log increase in the upwind NDVI shock. Panel (b) shows results for pollen counts, showing an increase of about 1 log unit per log increase in upwind NDVI. The pollen effect shows no lagged effect, concentrating on the day of the shock. For both the pollution and the pollen outcome, we find no effects for the placebo, downwind NDVI shocks.

In Appendix Figure A.7, we further report the event study analysis for five other air pollutants we have data on: coarse particulate matter (PM<sub>10</sub>), ozone (O<sub>3</sub>), nitrogen dioxide (NO<sub>2</sub>), sulfur dioxide (SO<sub>2</sub>), and carbon monoxide (CO). We find clear evidence of a pollution reduction effect across these pollutants except for O<sub>3</sub>, which shows no response to vegetation shocks. As discussed in Section 2, ozone is a secondary pollutant that forms due to reactions of precursor pollutants, including nitrogen oxides (NOx) and volatile organic compounds (VOCs); while urban forests absorb precursors of ozone such as NO<sub>2</sub>, they generate VOCs, which may cause the net effect on ozone ambiguous.

The Online Appendix reports a series of sensitivity checks. Appendix Figure A.8 and Appendix Tables A.3 through A.5 show results when we alter the definition of upwind cone using alternative radian degrees (0.25π, 0.75π, or 0.875π), alternative fixed effects (the inclusion of more flexible, station-by-month fixed effects and/or year-by-month fixed effects, alternative radius area around the monitor (5 km, 10 km, or 15 km), and whether we calculate NDVI using all grids or grids fall within the MMP planting areas.<sup>17</sup> Our results are robust to these arguably significant specification changes. One exception is panel II of Appendix Table A.3, where we lose power in the pollution reduction effect estimation when we confine the definition of upwind NDVI to areas within a 5km radius of the pollution monitor.

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<sup>17</sup> More stringent fixed effects, such as monitor-by-month and year-by-month fixed effects, may help capture important background trends in pollution over the study period, including those driven by China’s nationwide “War on Pollution” (e.g., [Greenstone et al., 2021](#)).

It is worth noting that this pattern does not necessarily conflict with our cross-sectional findings in Section 4, where we find that satellite-measured pollution has declined faster within the MMP area, where more NDVI growth is observed. The estimation here diverges from the cross-sectional exercise in that it links the day-to-day fluctuations in pollution *within a location* to NDVI from a specific wind direction on that day. Our results suggest that the impact of urban forestry on pollution reduction requires a fairly expansive area to become apparent. In other words, the reduction in pollution isn't solely dependent on the area immediately surrounding the point of observation, but instead is determined by a larger area of forests upwind. The cross-sectional exercise in Section 4 cannot tease out this relationship.

Appendix Figure A.9 reports an analysis of heterogeneous effects of upwind NDVI's pollution reduction effects by season. We augment equation (9) by interacting the upwind NDVI term with dummies for Spring (March-May), Summer (June-August), Autumn (September-November), and Winter (December-February the next year), along with the corresponding seasonal main effects. We omit the main upwind NDVI term so that the interactive effects can be separately identified for all four seasons. As a placebo check, we also include analogous interactions for the downwind NDVI term. We find that PM<sub>2.5</sub> reduction effects are weakest in Summer, while otherwise statistically similar across the other seasons.<sup>18</sup> Appendix Figure A.9 also shows that the pollution-reduction effect is slightly larger in the latter half of the study period, consistent with the idea that trees planted under the MMP were initially juveniles and grew into maturity with better pollution-reduction capacity over time.

Circling back to Figure 5. To the best of our knowledge, this is the first incident where researchers were able to estimate – both qualitatively and quantitatively – the impact of vegetation shock on air pollution reduction and pollen increase within the same empirical framework. Taken at the face value – that is, if we just judge the results from a pollution-vegetation and pollen-vegetation elasticity perspective – the pollen response appears large and

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<sup>18</sup> Beijing's recent afforestation initiatives include a mix of deciduous and evergreen species, with planting guidelines recommending a roughly 4:6 ratio (evergreen to deciduous) to provide both seasonal variety and year-round cover. At the same time, Beijing experiences its highest pollution levels in winter, driven by heating-related emissions and frequent temperature inversions that trap smog. These factors may explain why we still observe a significant pollution-reduction effect during winter, even when many deciduous trees are not in leaf.

potentially important. The natural next step is to quantify what are the magnitudes of pollution and pollen shocks in terms of health costs, and how do they compare with each other.

## 6. Health

This section aims to quantify the health parameters in equation (2). We begin by providing new estimates on the health effects of pollen exposure, linking day-to-day variation in pollen counts to emergency room visits and spending. We next discuss estimates on the health effects of PM<sub>2.5</sub> pollution from the existing literature, and a direct estimation using our own data.

### 6.1 Estimation

We do not observe the residential location of the patient from the medical claims data, but rather the hospital at which the healthcare was delivered. Similarly, as we discussed above, the city monitors pollen and pollution at 12 and 35 monitoring locations, respectively. To streamline these measurements, in our analysis we use “district” as the cross-sectional unit. There are 16 districts in Beijing, and Panel (a) of Figure 4 shows the district delineation. We make this choice because every district has at least one hospital, one pollen monitor, and one pollution monitor, and therefore it is intuitive to aggregate the medical claims data, the pollen data, and the pollution data to the district-daily level.<sup>19</sup>

We estimate the health effects of pollen using a fixed effects panel estimation model:

$$\text{Log Emergency Room Uses}_{it}^g = \beta \cdot \text{Log Pollen Counts}_{it} + \alpha_i + \alpha_t + X_{it}\gamma + e_{it} \quad (10)$$

where  $\text{Log Emergency Room Visits}_{it}^g$  denotes the log of total emergency room (ER) utilization – total number of visits or total medical spending associated with the visits – for district  $i$  and date  $t$ . In the main specification, we measure ER outcomes using a three-day lookahead window,

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<sup>19</sup> Implicitly, we assume here that a monitor’s measurement is a good proxy for the average population exposure in the district, and that people seek healthcare in hospitals in their home district. The former assumption is reasonable because monitors are placed in densely-populated areas. District-based healthcare seeking is a reasonable assumption to the first order, especially in the context of emergency room visits – our focal outcome measure of healthcare use – where people are unlikely to travel long distances to seek emergency care.

counting ER uses for day  $t$ ,  $t + 1$ , and  $t + 2$ . This is intended to capture the effect of a transient pollen shock on health outcomes that might take some days to manifest. In Appendix Tables A.5 and A.6, we repeat the estimation using same-day outcome (no lookahead) or seven-day outcomes (even longer lookahead) as robustness checks. For simplicity, our main analysis uses a log-log specification. In Appendix Table A.9, we report robustness checks that instead use the rate of ER visits (daily visits divided by district population) as the outcome variable, with regressions weighted by district population.

In heterogeneity analysis, we will look at ER outcomes for different subgroups, and hence the superscript  $g$ . The subgroup categorization we are most interested in is the underlying medical conditions that led to the ER. We will report coefficients when the outcomes are ER for all causes, respiratory causes, sensory causes, and other non-respiratory and non-sensory causes. We will show that the effects of pollen exposure are driven by respiratory and/or sensory-related ER, which helps provide a sense of plausibility of the results. In other heterogeneity analyses we will also look at subgroups by, for example, patient's age.

Regarding the rest of the equation (10):  $\text{Log Pollen Counts}_{it}$  is the log of pollen counts at district  $i$  on date  $t$ . We include district-by-month, year, month, day-of-week, and holiday fixed effects, which we denote  $\alpha_i$  and  $\alpha_t$  for simplicity. In some robustness specifications we further control for factors that can potentially correlate with both pollen and influence health outcomes (such as particulate pollution or air temperature), which is captured by the matrix  $X_{it}$ . Standard errors are two-way clustered at the district and the date levels.

Equation (10) does not have a causal research design, and so whether  $\beta$  reflects the causal impact of pollen depends on how much we believe the day-to-day variation in pollen, conditional on the fixed effects controls, is as good as random. We first inspect the nature of pollen variation. Panel (a) of Figure 6 shows the daily time series of city-average pollen counts. The city has two pollen seasons, one around March and the other around September. This seasonality holds consistently across the four years (2013-2016) of our data, and therefore any strategy that relies on within-year, cross-season comparison will miss out confounders that are correlated with such seasonality. Instead, we note that there seems to be abundant pollen variation *within* a given season, but *across* different years, which is more likely to be driven by idiosyncratic variation. This underlies our choice to include the *district-by-month* fixed effects in

equation (10), which allows us to compare conditions at a given district *and* in a given month-of-year, but across different years when pollen counts are high versus low.

For the rest of the discussion, we will simply refer to  $\beta$  as the *effect* of pollen exposure on ER uses.

## 6.2 Results

Panel (b) of Figure 6 provides an initial look of the main results. It shows decile bin scatterplots of fixed effects-residualized ER visits and pollen counts, done separately for respiratory/sensory-related ER visits and all other ER visits. The two dashed lines are superimposed linear fit, whose slopes correspond to the  $\beta$  coefficient from equation (10). The results reveal that the effect of pollen concentrates mainly for respiratory-sensory causes, but not others. In a way, this finding provides a degree of ex-post confidence that our estimates do capture the health effects of pollen – in lieu of a spurious correlation – because a priori clinical knowledge suggests the potential effect of pollen should operate through respiratory and sensory channels (Section 2).

Table 2 provides more details of the estimation results. Each cell in this table represents the  $\beta$  coefficient estimate from a separate regression. Given the log-log specification, all numbers in the table represents elasticity. Columns 1-4 look at number of ER visits as the outcome variable, and columns 5-8 look at ER spending as the outcome variable. For both ER visits and spending, we examine all-cause ERs, those due to respiratory or sensory causes, and all other non-respiratory/sensory ERs.

Panel A reports the baseline results. We find an ER-pollen elasticity of 0.021 (SE = 0.004) for respiratory ER visits and 0.042 (SE = 0.006) for sensory visits (columns 2 and 3). We found similar effect sizes for ER spending (columns 6 and 7). The effects on non-respiratory/sensory ER uses are an order of magnitude smaller and, in the case of spending, statistically insignificant (column 4 and 8).

Panels B, C, and D of Table 2 documents three dimensions of heterogeneity. In panel B, we stratify ERs by whether the patient ended up being hospitalized or not. We find that the effect comes through almost entirely by ERs that do not need further inpatient care.

In panel C, we stratify ERs by age of the patient. We find that the effect sizes (in elasticity) are statistically indistinguishable for elderly (aged over 60) and non-elderly. If anything, some of the effect sizes – for example, sensory-related ER visits – are larger for the non-elderly.

These heterogeneity results in panels B and C are worth noting because they suggest the characteristics of pollen's health effects differ from what we know about the health effects of industrial pollution. For example, PM<sub>2.5</sub> pollution is well known to disproportionately affect the elderly population and can cause severe conditions including hospitalization and even death ([Deschenes, Greenstone, and Shapiro, 2017](#); [Deryugina et al., 2019](#)). This is something to keep in mind when we compare the health costs of pollen versus PM<sub>2.5</sub>: while we argue that pollen exposure imposes non-negligible health risks, they are unlikely to *exceed* the health benefits of industrial pollution reduction. We will come back to this point in Section 7.

In panel D, we further document that pollen's effects are larger for patients with prior respiratory-sensory conditions, defined as those who had respiratory and sensory visits in the past month. From a health management perspective, this result suggests the health effects of pollen might be more manageable than those of pollution, as one may target the protection of easily-identifiable vulnerable group during pollen seasons, which are also relatively fixed and predictable than industrial pollution episodes.

In Table 3, to cast light on the type of pollen species that are particularly damaging, we replace the right hand side variable **Log Pollen Counts<sub>it</sub>** of equation (10) to be the log of pollen counts of a specific species. We find that both tree (*Cupressaceae*, *Salicaceae*, *Moraceae*) and weed pollen (*Artemisia*, *Chenopodiaceae*) seem to trigger increased ER visits. Once again, for all of these species, respiratory and sensory ERs are driving the effects. The only exception where we observe no significant ER effect is *Pinaceae* pollen that is mostly associated with pine trees. This is consistent with clinical evidence that pine tree pollen-caused allergy is [uncommon](#) ([AAAI, 2018](#)). The takeaway of Table 3 is that, except for pine tree pollen, the other five species appear almost equally bad as ER triggers. If we view ER spending as a measure of severity, then it seems reasonable to conclude that *Cupressaceae* (Cypress) and *Moraceae* (Mulberry) have the largest impacts.

One potential characteristic of pollen's health effects is that they may evolve over time. There are several countervailing forces. First, allergic sensitization often requires multiple seasons of exposure, and once sensitized, subsequent exposures can accelerate clinical reactivity

and symptom severity (e.g., [Greiner et al., 2011](#)). Second, some individuals may develop stronger regulatory immune responses or desensitization after prolonged exposure, and in some cases even achieve full remission (e.g., [Bodtger et al., 2004](#)). Third, allergy sufferers often stay indoors on high-pollen days, use air purifiers, wear masks, or use medications, which mitigate exposure over time. Appendix Table A.10 shows that the effect of pollen on ER visits remains stable across our study period. This stability is important for policy cost-benefit analysis, as it suggests the effect reflects a persistent burden rather than a temporary surge.

### 6.3 Effect Sizes

The analysis so far finds a statistically significant increase in ERs due to pollen exposure. For example, panel (a) of Figure 6 suggests an ER visits-pollen elasticity of 0.0061 (SE = 0.0013) and an ER spending-pollen elasticity of 0.0037 (SE = 0.0012). How important are these effects? A simple way to answer this question is to estimate the same sets of ER-elasticities, but for industrial pollution such as PM<sub>2.5</sub>. The prior literature on the health effects of PM<sub>2.5</sub> also has established estimates that we can benchmark against, which we discuss towards the end of this section.

To estimate the effect of PM<sub>2.5</sub> exposure on ER, one replaces the log pollen variable on the right hand side of equation (10) with  $\text{Log PM2.5}_{it}$ , i.e., log PM<sub>2.5</sub> concentration at district  $i$  on date  $t$ . Before doing that, it worth considering whether one actually has independent variation between pollen and PM<sub>2.5</sub> concentration to pick up effect separately for these two environmental agents. We discussed in Section 2 that, in theory, most pollen grain diameters range from about 10 to 100 micrometers, which are too large to be picked up by PM<sub>2.5</sub> monitors which are calibrated to measure particles of 2.5 micrometers in diameter or smaller. To assess this fact empirically, in panel (a) of Figure 7, we provide scatterplots of daily PM<sub>2.5</sub> against pollen counts at the district level, separately for each month of the pollen monitoring season (March to September). We find that in most month of the year, PM<sub>2.5</sub> and pollen exhibit small and statistically insignificant correlation, with the direction of correlation showing no particularly consistent pattern (positive in March and May, and negative for other months). The exception is the month of August where the two variables exhibit a strongly *negative* correlation (elasticity = -0.240, SE = 0.045). Given the overall abundant independent variation between the two variables, it seems reasonable to

proceed with estimating the separate impacts of the two environmental agents on health outcomes.

Panel (b) of Figure 7 reports the estimates. The solid lines repeat Table 2a, columns (1) and (5) on the effects of pollen on ER visits and ER spending. The dashed lines show the effects of  $PM_{2.5}$  using the exact same specifications. We find that a log increase in  $PM_{2.5}$  leads to an increase of three-day ER visits by 0.007 log points ( $SE = 0.0023$ ), which is similar to the pollen-ER visits elasticity estimate. The  $PM_{2.5}$  elasticity for ER spending is estimated to be 0.0095 ( $SE = 0.0027$ ), over two-fold of the pollen-ER spending elasticity. This is consistent with our prior observation that pollen exposure tends to cause less severe conditions, and thus the effect of  $PM_{2.5}$  on ER spending is larger, even if the effect on visitation rate is similar.

Are these reasonable estimates of the effect size of  $PM_{2.5}$ ? We compare our estimates to published estimates from the literature that are similar in approach. Deryugina et al. (2019) links daily variation of  $PM_{2.5}$  driven by changes in wind directions to health outcomes in the U.S. Medicare population. In their primary specification, the effects of  $PM_{2.5}$  on three-day ER visits and spending converts to an elasticity of 0.0071 and 0.0126, respectively. A potential concern with comparing to is that the background  $PM_{2.5}$  level in Deryugina et al. (11  $\mu g/m^3$ ) is much lower than that Beijing during our study period (67  $\mu g/m^3$ ), and so the elasticities may not be comparable. Barwick et al. (2022) links daily  $PM_{2.5}$  in China to frequency of bank card transactions in healthcare categories. Their primary specification converts to a daily  $PM_{2.5}$ -transaction elasticity of 0.0364. Xia et al. (2019) studies the effect of  $PM_{2.5}$  on medical spending using the same source of medical claims data from Beijing. Their most precise estimate linking daily  $PM_{2.5}$  to three-day medical spending converts to an elasticity of 0.0132. We reiterate that these estimates should be interpreted with caution, given that they rely on OLS with high-frequency data and fixed effects rather than a quasi-experimental design (which tends to produce larger effect sizes from the prior literature). Our pollution estimates are smaller but generally in line with prior evidence.<sup>20</sup>

It is worth emphasizing that the point of this subsection is to put our pollen-ER elasticity estimates in perspective by comparing them with  $PM_{2.5}$ -ER elasticities, both estimated using our

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<sup>20</sup> There is a large literature on the mortality effects of air pollution, but the availability of morbidity and healthcare costs estimates is more limited. We will discuss the mortality literature in more detail in Section 7.

own data and from the literature. Our headline conclusion – that the healthcare externality of pollen allergy is nonnegligible – is based on our finding that pollen exposure is qualitatively bad enough to trigger ER visits, and quantitatively similar with ER effects of pollution. However, this is not to say that the *overall* health costs of pollen are necessarily on par with those from industrial pollution exposure. In particular, various studies have documented the mortality effect of PM<sub>2.5</sub>, which is believed to be an overwhelming component of the overall health costs of pollution when compared to morbidity and healthcare costs (e.g., [Landrigan et al., 2018](#)). Though our data do not allow us to observe mortality rate and so we cannot directly estimate the pollen-mortality coefficient, we expect the effect to be small, if any, given that most pollen-related ER does not lead to hospitalization, let alone more severe, life-threatening complications.

## 7. Program Benefits and Costs

### 7.1 Monetizing Health Benefits and Risks

We are now ready to plug in the marginal parameter estimates from Sections 4, 5, and 6 to the environment and health equations of Section 3. We have three key sets of parameters: First, the MMP increases planting site NDVI by 0.032 units or about 10.5 percent over the 2012-2020 period. Second, for the typical resident, a log unit increase in upwind NDVI in MMP areas causes -0.4 log point reduction in PM<sub>2.5</sub> and +0.7 log point increase in pollen. Third, a log unit decrease in a district's PM<sub>2.5</sub> leads to 0.007 log points decrease in ER visits; a log unit increase in pollen leads to 0.0061 log points increase.

Together, we estimate that, by 2020, the MMP policy reduces the average population PM<sub>2.5</sub> exposure in the city by 4.2 percent (about 2.9 ug/m<sup>3</sup> from 2012 baseline).<sup>21</sup> This effect is significant: the city of Beijing has achieved a 40 percent reduction in pollution since China's "War on Pollution" campaign. Our estimates therefore suggest about 10% of that reduction is contributed by urban afforestation.

We can calculate the health value of such air quality improvement. There are many established estimates we can choose from the literature. Here we focus on quasi-experimental,

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<sup>21</sup> To get baseline 2012 PM<sub>2.5</sub> value, we use annual average 69.5 ug/m<sup>3</sup> as reported by U.S. Embassy in Beijing.

economics studies that are conducted using health data from China, so that the study contexts are more compatible. Starting with healthcare costs. [Xia et al. \(2019\)](#) uses medical claims data from Beijing, and finds that a 10  $\mu\text{g}/\text{m}^3$  increase in  $\text{PM}_{2.5}$  concentration increases city-wide medical spending by 791 million CNY (112 million USD) per year. [Xia et al. \(2019\)](#) focuses on short-term impact of  $\text{PM}_{2.5}$  shock on spending over a three-day window. [Barwick et al. \(2022\)](#) introduces a finite B-splines method to capture lagged effects of pollution for up to three-months. Using national data on credit and debit card transactions took place in healthcare institutions, the study finds that the cumulative effect of pollution at the medium run (three-month) is about four times larger than the effect of a contemporaneous, one-day shock. If we scale up the [Xia et al. \(2019\)](#) estimate accordingly, it implies an annualized cost of 3.2 billion CNY (448 million USD) per 10  $\mu\text{g}/\text{m}^3$  increase in  $\text{PM}_{2.5}$ . Therefore, the annual healthcare benefits of a 2.9  $\mu\text{g}/\text{m}^3$  reduction in  $\text{PM}_{2.5}$  due to the MMP project is estimated to be between **229 to 916 million CNY** (32 to 88 million USD). This amounts to 0.1 to 0.4 percent of Beijing's reported annual total health spending of 219 billion CNY.<sup>22, 23</sup>

Another important component of the health benefit of pollution reduction is through a mortality effect. [Ebenstein et al. \(2017\)](#) exploits quasi-experimental differences in long-term  $\text{PM}_{10}$  exposure across cities in China. The study finds that a 10  $\mu\text{g}/\text{m}^3$  increase in  $\text{PM}_{10}$  (about 6  $\mu\text{g}/\text{m}^3$  in  $\text{PM}_{2.5}$ ) causes an 8 percent increase in the cardiovascular mortality. Combining these estimates with an average mortality rate of 4.3 per 1,000 residents (per Beijing city government) and a population of 20 million people, we estimate that a 2.9  $\mu\text{g}/\text{m}^3$  reduction in  $\text{PM}_{2.5}$  due to the MMP project contributes to an annual reduction of 3,325 deaths. Following [Barwick et al. \(2022\)](#), we use a VSL of 1.5 million CNY and this gives us an annual mortality benefits of **5 billion CNY** (710 million USD). Alternatively, we use the Air Quality Life Index ([Greenstone and Fan, 2018](#)) and calculate that the potential gain in life expectancy is **0.28 years** for the average city resident if the reduced pollution effect of the MMP is to sustain in the long run.

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<sup>22</sup> We focus on the pollution-health mapping on  $\text{PM}_{2.5}$  as the focal pollutant which is both the canonical health-relevant pollutant in prior work and, in our data, the pollutant with the largest elasticity to upwind greening (Appendix Figure A.7).

<sup>23</sup> A caveat worth noting is that some atmospheric science studies suggest the capacity of trees to absorb or filter pollution may be influenced by the proximity of nearby industrial emission sources (e.g., [Grylls and Reeuwijk, 2022](#)). To the extent that this effect matters, the pollution-reduction benefits of afforestation are context-dependent, and the magnitude we estimate for Beijing may not directly generalize to cities with very different pollution source profiles.

The health benefit from pollution reduction likely overwhelms the cost of pollen increases. This is for two reasons. First, we find no evidence that pollen increases the need for inpatient care (Table 2), whereas an inpatient visit costs 40 more than the average hospital visit, and total inpatient spending accounts for 70 percent of overall medical spending in Beijing. Combining this fact with our estimate in Figure 7 that the cost elasticity of pollen is about a third of that of pollution exposure, we estimate the magnitude of the healthcare costs of pollen increased related to MMP project to be about one ninth of the magnitude of the pollution benefit, or **25 to 102 million CNY** (3.5 to 14 million USD). Second, the lack of a hospitalization effect, combined with the fact that we observe a major share of the effect comes through the non-elderly population, makes it reasonable to think that pollen is unlikely to lead to more severe, life-threatening complications that cause deaths. Therefore, if we were to take into account the potential life years saving of pollution reduction, the relative health costs of the pollen increase may be even lower.

We reiterate that it is not the intention of this calculation to downplay the health cost of pollen. Quite the opposite, our empirical analysis points out that pollen allergy is important enough to cause significant increases in emergency room usage. Our findings indicate that pollen triggers emergency room visits to a similar extent as PM<sub>2.5</sub> does. The smaller health cost numbers are driven by the observation that pollen-related emergency room visits are generally less severe, with a lower likelihood of requiring expensive inpatient care or resulting in fatalities, compared to pollution-related visits. However, a comprehensive understanding of the full range of effects of pollen exposure on cognitive and physical health, as well as other aspects of well-being, requires further research and investigation.

The health value of the MMP program per the pollution change is thus about **55 billion CNY** over the course of the decade since its initiation. This is a significant number compared to the magnitude of Beijing's GDP which is about 4,000 billion CNY. Recall that in Section 2.1, we mentioned that the total cost of the MMP project is **75 billion CNY**. Our health calculation thus leads to a conclusion that these costs will likely be recouped via health gains over the next decade.

In Online Appendix A, we further examine the housing market capitalization effect of the MMP policy, showing that homes located closer to MMP plantation areas experienced greater

price appreciation following the program's launch. This provides an alternative way to quantify the value of the program.

## 7.2 Additional Policy Implications

Our main analysis demonstrates a substantial net benefit of the program for environmental health. Here, we outline two additional policy recommendations that follow from our empirical evidence.

**Afforestation along the Winter-Wind Corridor.** Our results show that urban forests filter air pollutants and reduce downstream pollution along prevailing wind paths. In Beijing, pollution peaks during the winter months, when northwestern winds dominate (Appendix Figure A.12). By contrast, pollen-related allergies are most prevalent in spring and autumn, when southern winds prevail (Appendix Figure A.5). To maximize pollution mitigation while minimizing pollen-related health risks, a sensible policy recommendation is to prioritize afforestation efforts in the city's northwestern corridor.

**Tree Species Choices.** Another policy implication arises from our pollen species analysis. In Beijing, airborne pollen primarily originates from six plant families: Cupressaceae (cypress), Salicaceae (willow), Pinaceae (pine), Moraceae (mulberry), Artemisia (wormwood), and Chenopodiaceae (goose-foot). Our heterogeneous-effects analysis in Table 3 indicates that all of these families—except for Pinaceae—significantly increase ER visits. Where feasible, policymakers should therefore prioritize planting pine trees or other low-allergen species to mitigate adverse health impacts.

These said, urban afforestation is a complex ecological task. Policymakers must also consider factors such as local water availability, the role of biodiversity in promoting ecosystem resilience, and the opportunity cost of land use ([Holl and Brancalion, 2020](#)). Our cost-benefit analysis is one valuable input rather than a prescriptive rule for tailoring afforestation policy to local conditions in order to maximize overall social welfare.

## 8. Conclusion

Urban forests have a ubiquitous presence in cities worldwide, but few economic research studies them. This paper examines the effect of urban forests on air quality and health, leveraging the greening up of the city of Beijing and its policy experiment over the past decade. We begin by documenting a substantial greening up of a mega city and examine the contribution by a government-led mass afforestation policy. We then quantify the impact of urban forests on downwind air quality improvement using a quasi-experimental research design. Our paper also investigates increased pollen exposure as a source of negative health shocks. We conclude that Beijing's afforestation brings enormous health benefits: by our calculation, the policy is expected to cause a 0.28 years increase in life expectancy for the average city resident if the reduced pollution effect is to sustain in the long run. This benefit is large relative to the cost of the policy. We show that the environmental value of the policy is only partially capitalized in housing prices, likely because long-range pollution spillover benefits are subtle and not as readily perceived by residents as more localized, aesthetic benefits.

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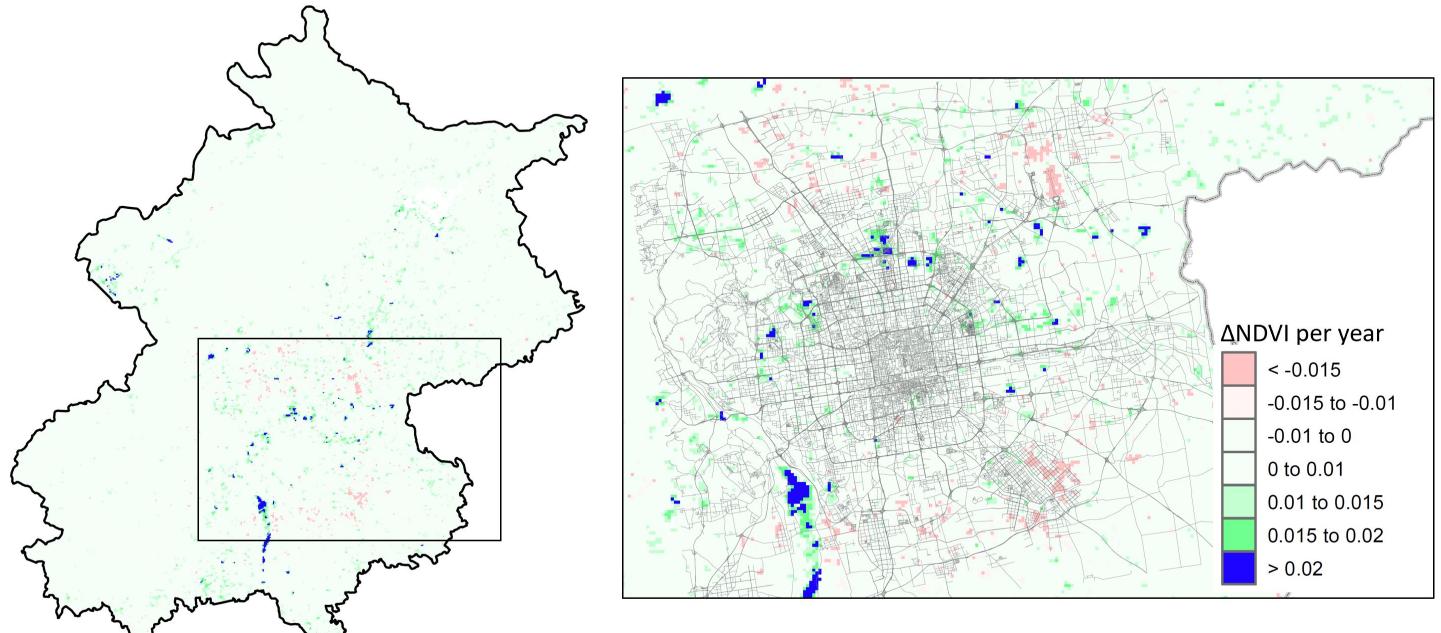
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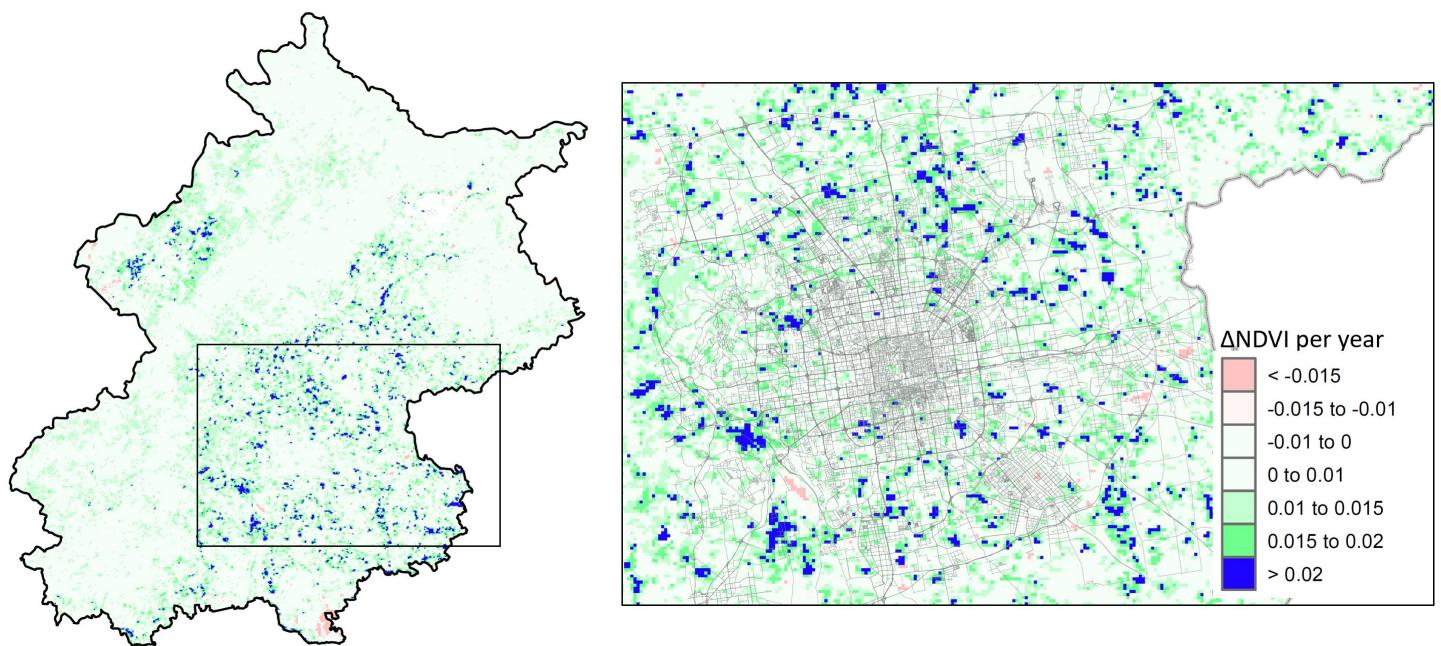
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Figure 1. The Greening Up of Beijing

(a) Rate of vegetation growth, 2001-2011



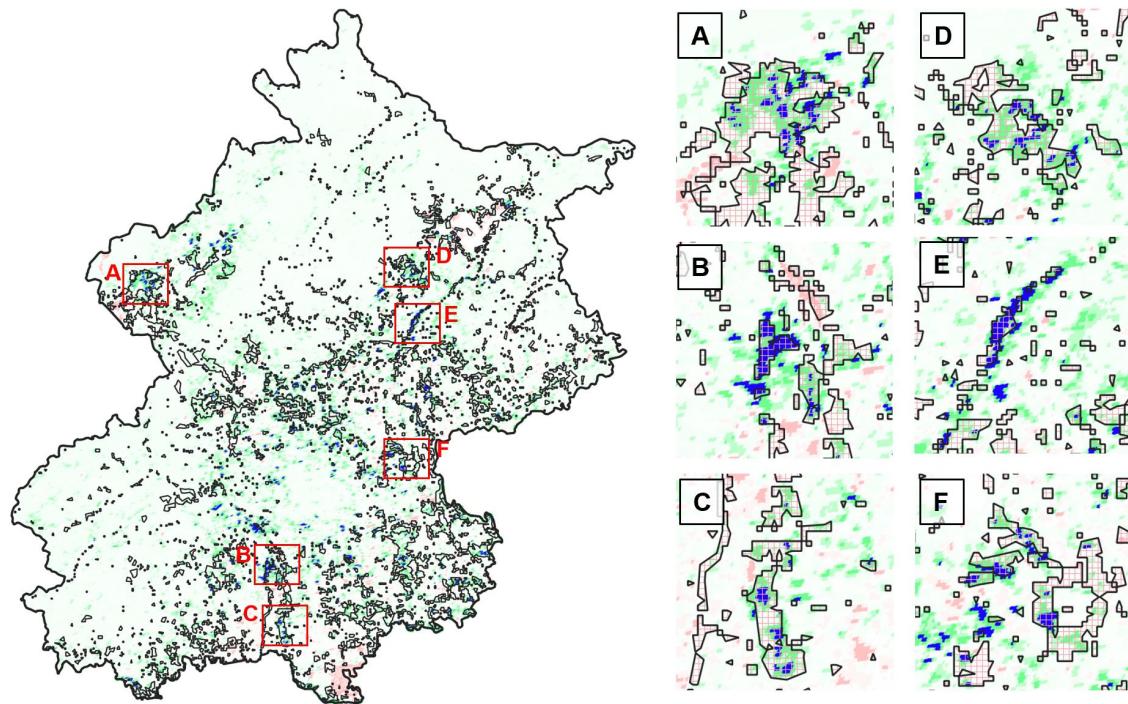
(b) Rate of vegetation growth, 2012-2020



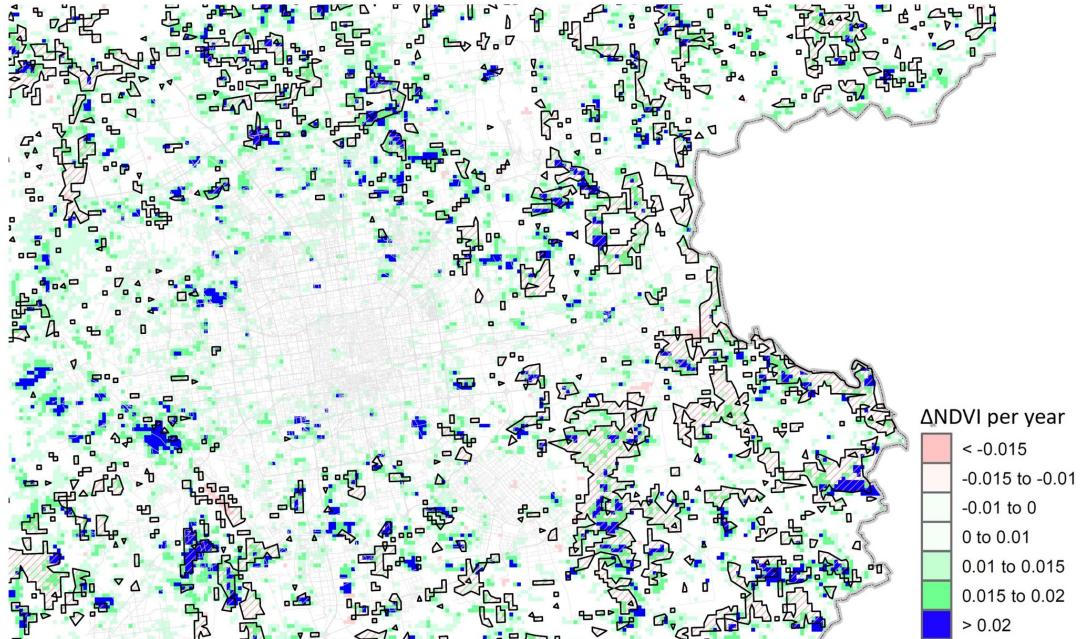
Notes: This figure shows the grid-level annual rate of NDVI growth for the 2001-2011 period (panel a) and the 2012-2020 period (panel b). The left side of each panel shows the conditions for the entire prefecture city of Beijing. A zoomed-in view of the city center is provided on the right side of the panel. Gray lines within the zoomed-in view represent the road networks.

Figure 2. The Million Mu Project (MMP) and Vegetation Growth

(a) Location of MMP planting sites

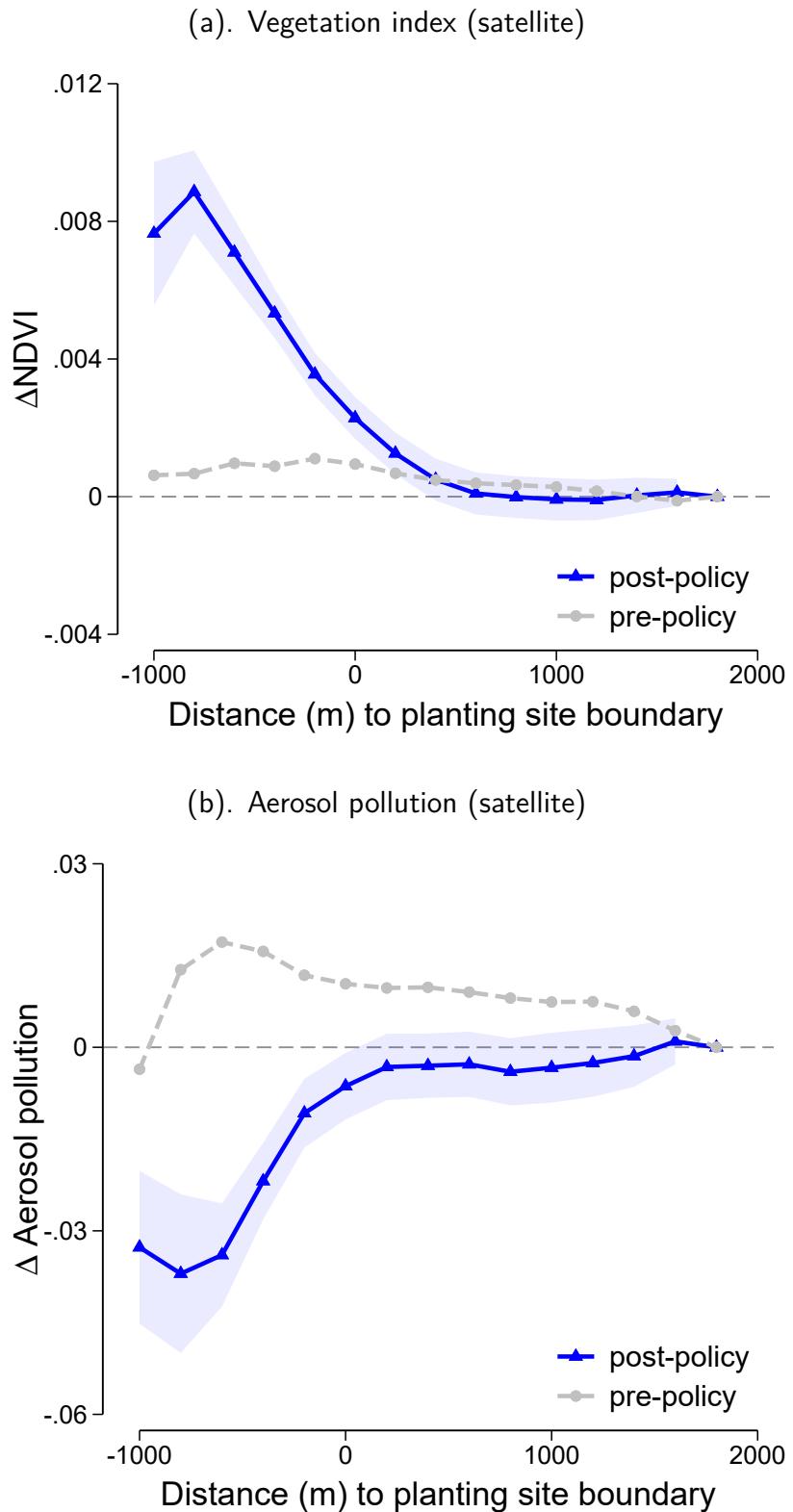


(b) Location of MMP planting sites (city center)



Notes: Color indicates the annual rate of growth of NDVI between 2012-2020, with blue indicating strong growth. Polygons represent MMP planting sites. Panel (a) shows the condition for the entire prefecture city of Beijing, and six example zoom-in areas. Panel (b) provides a zoomed-in view of the city center.

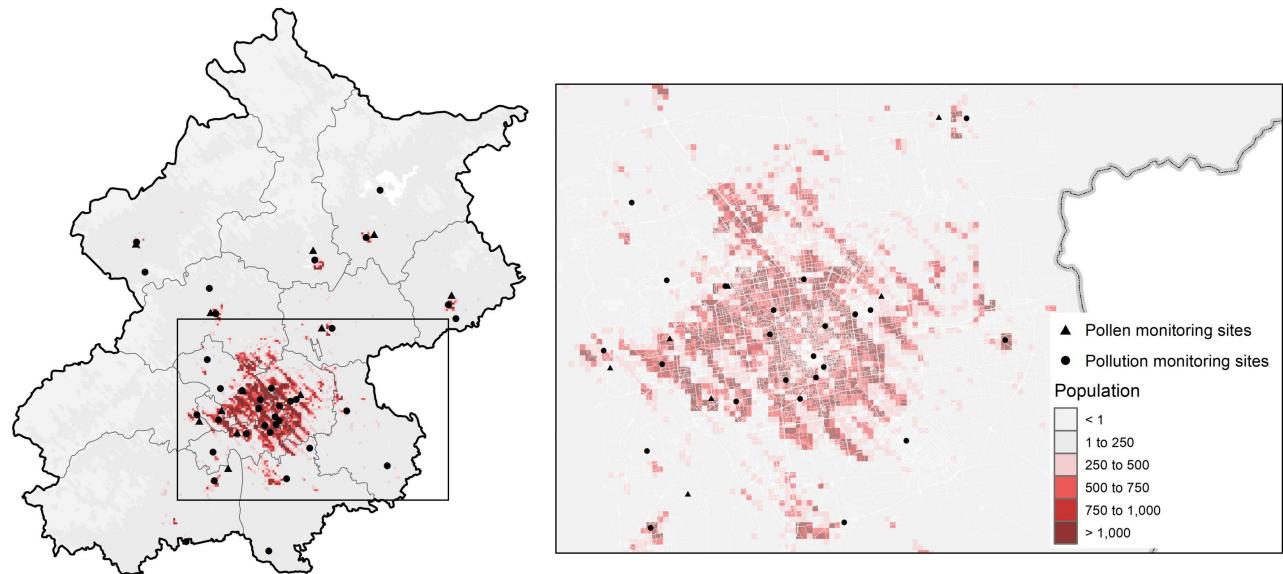
Figure 3. Changes in Vegetation and Air Quality Near Planting Sites



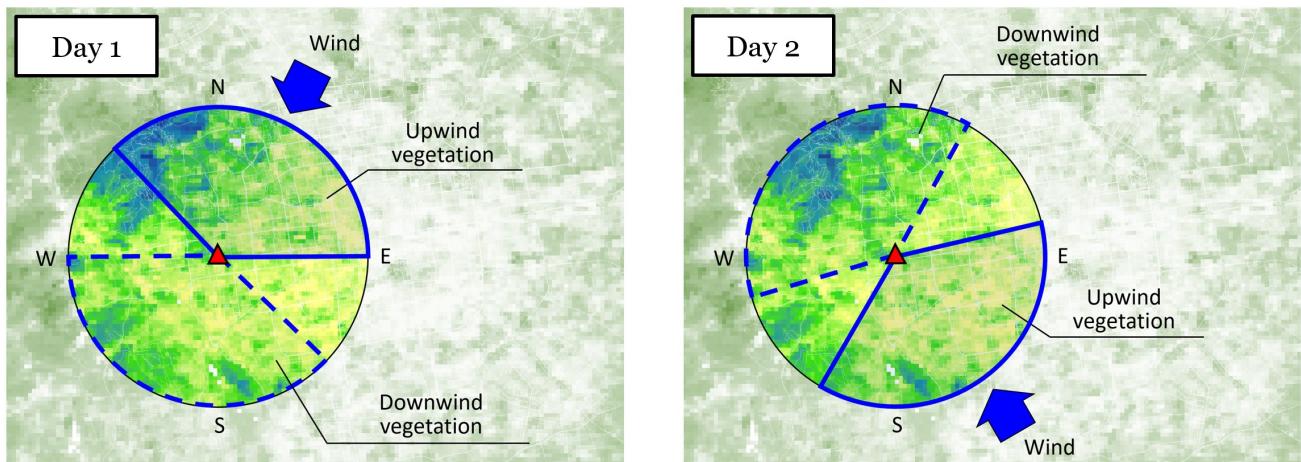
*Notes:* This figure shows the annual rate of change in NDVI (panel a) and AOD (panel b) as a function of the distance to the nearest MMP planting site boundary. The dashed line represents the pre-MMP period of 2001-2011, and the solid line corresponds to the post-MMP period of 2012-2020. These estimates are derived from grid-level, cross-sectional regressions of the annual rate of change on a series of dummy variables indicating distance bins (200-meter increments), with the 1800-2000 meter bin as the omitted category. The estimation is done separately for the pre- and post-MMP periods. For the post-MMP estimates, range plots display the 95% confidence interval, constructed using 1-km grid cluster bootstrap standard errors.

Figure 4. Urban Center Effects: Illustration of Empirical Strategy

(a) Population representation of pollution and pollen monitors

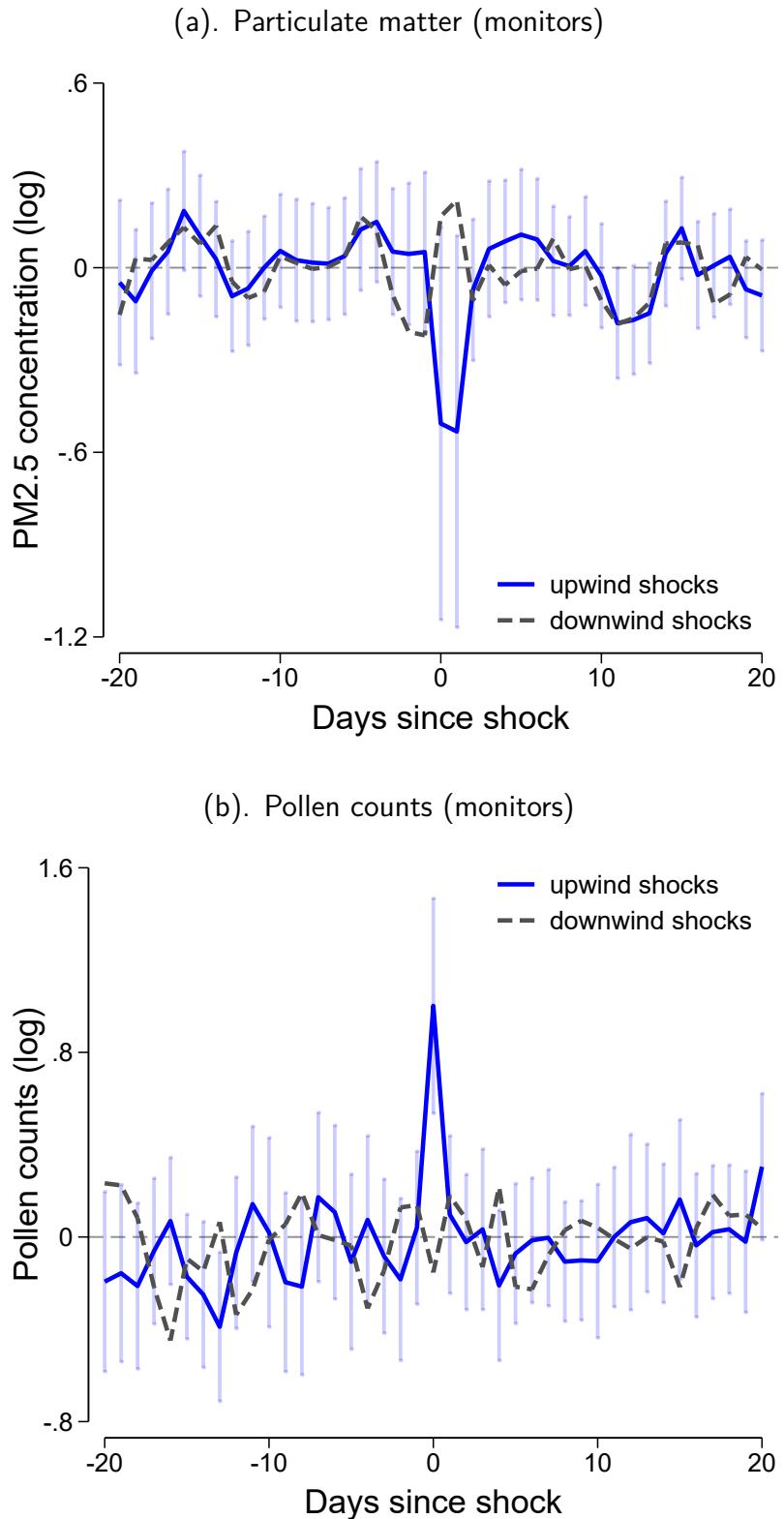


(b) Illustration of Empirical Strategy



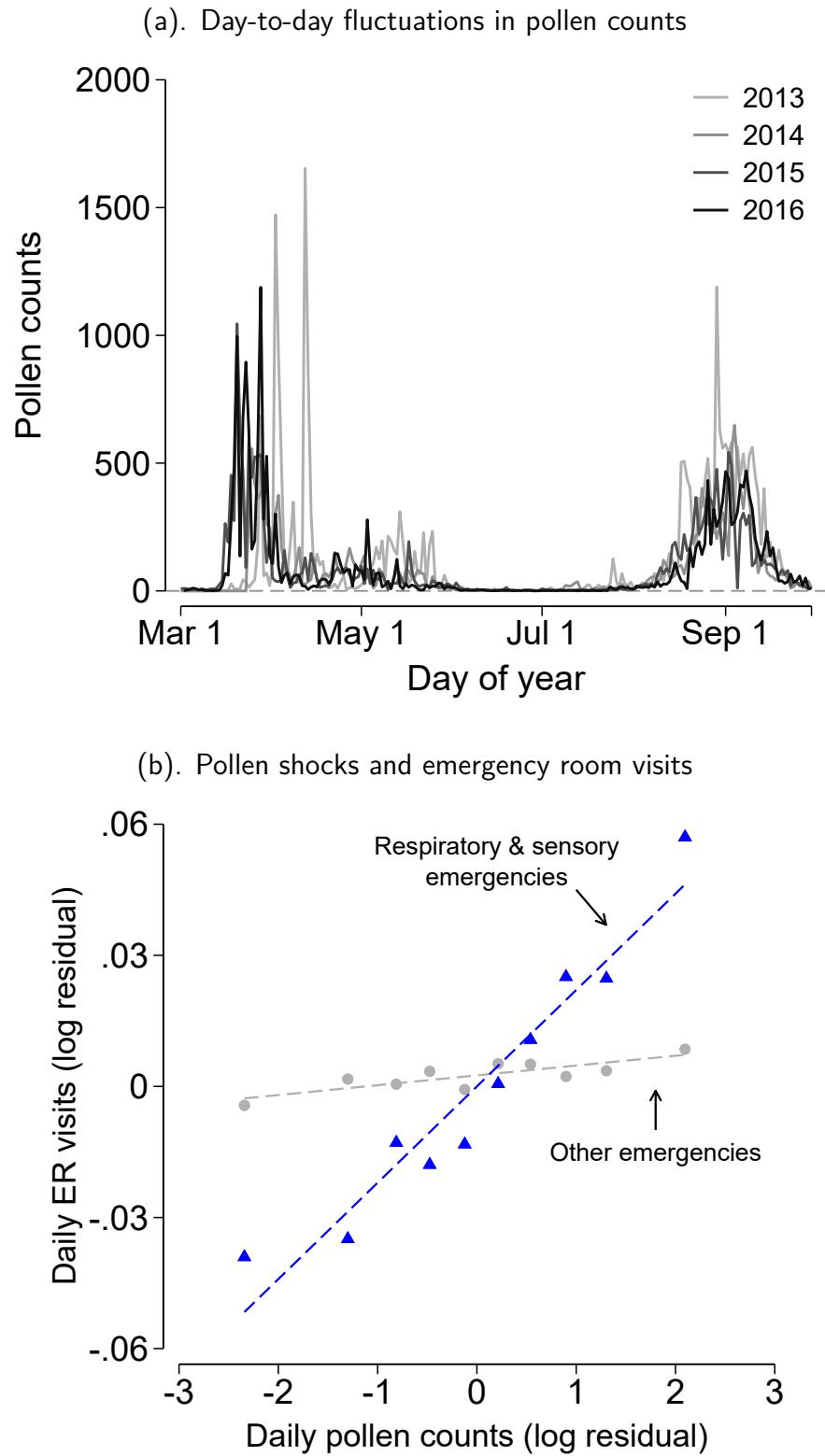
Notes: Panel (a) shows the location of in-situ air pollution monitors (dots) and pollen monitors (triangles), overlaid with grid-level estimates of the population as of the year 2010. Panel (b) provides an illustration of the identification strategy used to estimate the causal effect of vegetation on pollution and pollen. Consider a specific location of a pollution or pollen monitor, represented by the red triangle. The blue arrow indicates the prevailing wind direction for a given day. "Upwind vegetation" is defined as the average NDVI index across all grids that fall within a 135-degree cone in the upwind direction of the monitor on that day. "Downwind vegetation" is defined similarly but using the downwind direction.

Figure 5. Urban Center Effects: Upwind Shocks, Particulates Pollution, and Pollen Counts



*Notes:* This figure shows regression coefficients of air pollution (panel a) and pollen counts (panel b) on 20 lead, contemporaneous, and 20 lag terms of upwind vegetation shocks. We focus on fine particulate matter pollution ( $PM_{2.5}$ ) in panel (a), and we report results for other criteria air pollutants ( $PM_{10}$ ,  $O_3$ ,  $NO_2$ ,  $SO_2$ , and  $CO$ ) in the Online Appendix. In each panel, a set of placebo coefficients is also displayed, obtained by running the same regression but replacing upwind vegetation shocks with downwind shocks. Range bars show 95% confidence intervals constructed using standard errors two-way clustered at both the monitor and the day-of-sample level.

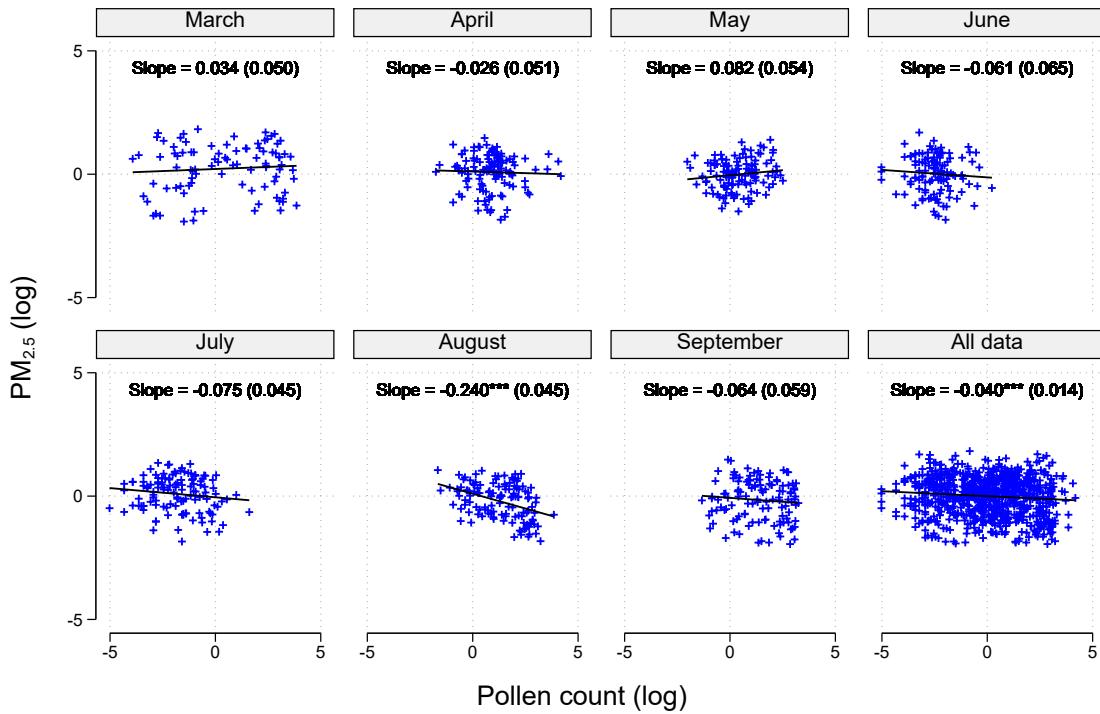
Figure 6. Health Effects: Daily Pollen Exposure and Emergency Room Visits



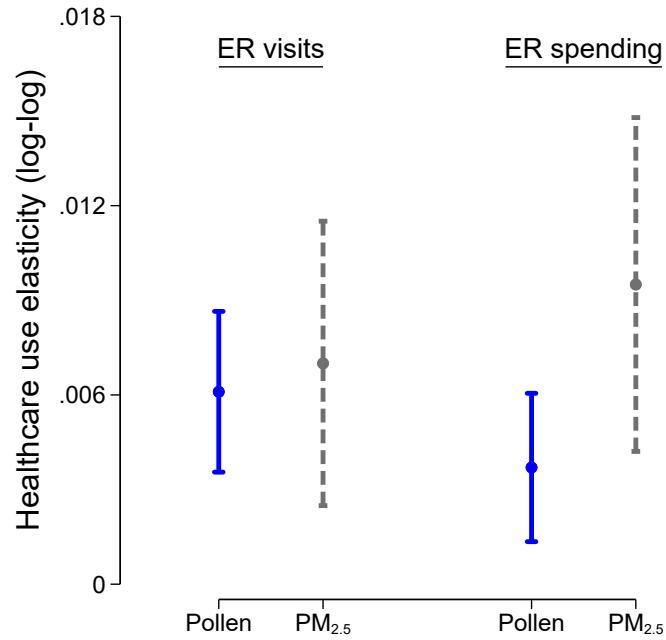
Notes: Panel (a) shows the average pollen counts per day-of-year at the site level for each year of data during the pollen monitoring season (March 1 to October 15). Panel (b) is a bin scatterplot on the relationship between fixed effects-residualized log daily emergency room visits and residualized log daily pollen counts. The patterns are separated for respiratory and sensory emergencies versus all other emergencies.

Figure 7. Comparing Health Effects: Pollen vs. Fine Particulate Matter (PM<sub>2.5</sub>)

(a). Independent variation between pollen and PM<sub>2.5</sub>



(b). ER effects of Pollen vs. PM<sub>2.5</sub>



Notes: Panel (a) shows the correlation between log PM<sub>2.5</sub> concentration and log pollen count at the district-day level. Each cross represents a district-day observation, and the black line is the superimposed OLS regression line, with slope and standard error estimates reported. Regressions are done separately for each month-of-year and for the pooled sample. Notice that data are only shown for March-September which is the pollen-monitoring season. Panel (b) shows coefficient estimates from four separate regression of an ER outcome (log visits on the left and log spending on the right) on log pollen or log PM<sub>2.5</sub>. All regressions control for district-by-month fixed effects, year-by-month fixed effects, day-of-week fixed effects, and holiday fixed effects. Range plots show 95% confidence intervals constructed using standard errors two-way clustered at the district and day-of-sample levels.

Table 1. Summary of Data Sources

Variable	(1) Data source	(2) Spatial frequency	(3) Temporal frequency	(4) Collection methods	(5) Used in
MMP planting site	City government	N/A	N/A	Survey	Section 4
Land use	RESDC	30 m	2000, 2010, 2020	Remote-sensing	Section 4
Population	WorldPop	100m	2010	Remote-sensing	Section 4
AOD	NASA MODIS	1 km	Daily, 2001-2020	Remote-sensing	Section 4
NDVI	NASA MODIS	250 m	16-day, 2001-2020	Remote-sensing	Sections 4 & 5
Wind	City weather bureau	20 stations	Daily, 2001-2019	In-situ sampling	Section 5
Pollution	City environmental agency	35 stations	Daily, 2014-2019	In-situ sampling	Sections 5 & 6
Pollen counts	City weather bureau	12 station	Daily, 2013-2016	In-situ sampling	Sections 5 & 6
Emergency room visits	City medical bureau	All hospitals	Records, 2013-2016	Administrative	Section 6
Home transactions	City housing bureau	Latitude/longitude	Records, 2006-2016	Administrative	Section 7

Notes: This table tabulates key variables used, source of data (column 1), spatial and temporal resolution of the raw data (columns 2-3), collection methods of the raw data (column 4), and where the variables are mainly used in the paper (column 5).

Table 2. Pollen Exposure and Emergency Room (ER) Utilization

	(1) All causes	(2) Respiratory	(3) Sensory	(4) Others	(5) All causes	(6) Respiratory	(7) Sensory	(8) Others
<b>A. Baseline specification</b>								
All ER	0.0061*** (0.0013)	0.0214*** (0.0038)	0.0424*** (0.0056)	0.0034** (0.0012)	0.0037*** (0.0012)	0.0161*** (0.0033)	0.0429*** (0.0117)	0.0009 (0.0013)
<b>B. Effects by severity</b>								
ER → Not hospitalized	0.0059*** (0.0013)	0.0214*** (0.0038)	0.0424*** (0.0056)	0.0031** (0.0013)	0.0036** (0.0012)	0.0164*** (0.0036)	0.0432*** (0.0114)	0.0008 (0.0013)
ER → Hospitalized	0.0019 (0.0063)	0.0043 (0.0057)	-0.0025 (0.0031)	-0.0002 (0.0056)	0.0029 (0.0310)	0.0207 (0.0470)	-0.0314 (0.0403)	-0.0006 (0.0318)
<b>C. Effects by age</b>								
Age < 60	0.0067*** (0.0015)	0.0242*** (0.0045)	0.0440*** (0.0060)	0.0037** (0.0013)	0.0041* (0.0019)	0.0211*** (0.0043)	0.0450*** (0.0118)	0.0013 (0.0017)
Age ≥ 60	0.0044** (0.0014)	0.0149*** (0.0042)	0.0176** (0.0059)	0.0023 (0.0015)	0.0038* (0.0020)	0.0141** (0.0056)	0.0594* (0.0295)	0.0006 (0.0023)
<b>D. Effects by prior condition</b>								
Prior respiratory-sensory visits	0.0310*** (0.0042)	0.0410*** (0.0082)	0.0395*** (0.0080)	0.0258*** (0.0048)	0.0231*** (0.0070)	0.0688** (0.0297)	0.1609*** (0.0317)	0.0249** (0.0083)
No prior visits	0.0057*** (0.0012)	0.0191*** (0.0036)	0.0353*** (0.0059)	0.0033** (0.0012)	0.0036*** (0.0011)	0.0140*** (0.0030)	0.0386*** (0.0119)	0.0012 (0.0013)

Notes: Each cell represents a separate regression using district-day level data. Each column presents ER records corresponding to different diagnoses. All outcomes are measured using a three-day look-ahead window (e.g., total number of ER visits today, tomorrow, and the day after tomorrow). Panel (a) uses all ER records. Panel (b) stratifies by visits that did and did not end up with hospital admissions. Panel (c) stratifies by age of the patient. Panel (d) stratifies by whether the patient had respiratory and sensory visits in the previous 30 days. All regressions control for district-by-month fixed effects, year-by-month fixed effects, day-of-week fixed effects, and holiday fixed effects. Standard errors are two-way clustered at the district and day-of-sample levels. Number of observation for each regression is 8,394. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table 3. Pollen Exposure and Emergency Room (ER) Utilization: Effects by Pollen Species

	(1) All causes	(2) Respiratory	(3) Sensory	(4) Others	(5) All causes	(6) Respiratory	(7) Sensory	(8) Others
	ER visits					ER spending		
<i>Cupressaceae</i> (Cypress)	0.0069*** (0.0019)	0.0216*** (0.0034)	0.0519*** (0.0065)	0.0042* (0.0019)	0.0022 (0.0018)	0.0134*** (0.0038)	0.0552*** (0.0098)	-0.0010 (0.0015)
<i>Salicaceae</i> (Willow)	0.0061** (0.0026)	0.0179*** (0.0034)	0.0330*** (0.0090)	0.0039 (0.0027)	0.0001 (0.0033)	0.0085 (0.0065)	0.0310 (0.0244)	-0.0018 (0.0031)
<i>Pinaceae</i> (Pine)	-0.0017 (0.0022)	0.0024 (0.0053)	0.0025 (0.0071)	-0.0023 (0.0020)	-0.0008 (0.0026)	-0.0025 (0.0060)	0.0044 (0.0169)	-0.0005 (0.0026)
<i>Moraceae</i> (Mulberry)	0.0059** (0.0024)	0.0225*** (0.0054)	0.0549*** (0.0081)	0.0029 (0.0024)	0.0058** (0.0025)	0.0225** (0.0076)	0.0491* (0.0238)	0.0023 (0.0026)
<i>Artemisia</i> (Sagebrush/Wormwood)	0.0041* (0.0019)	0.0252*** (0.0056)	0.0422*** (0.0082)	0.0006 (0.0017)	0.0034* (0.0016)	0.0267*** (0.0074)	0.0254 (0.0217)	-0.0012 (0.0015)
<i>Chenopodiaceae</i> (Goosefoot)	0.0010 (0.0022)	0.0216*** (0.0059)	0.0488*** (0.0088)	-0.0024 (0.0024)	-0.0001 (0.0017)	0.0165** (0.0070)	0.0172 (0.0218)	-0.0033 (0.0024)

Notes: Each cell represents a separate regression using district-day level data. Each column presents ER records corresponding to different diagnoses. All outcomes are measured using a three-day look-ahead window (e.g., total number of ER visits today, tomorrow, and the day after tomorrow). Each row corresponds to regression using a different right-hand-side measure of pollen species. All regressions control for district-by-month fixed effects, year-by-month fixed effects, day-of-week fixed effects, and holiday fixed effects. Standard errors are two-way clustered at the district and day-of-sample levels. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

## Appendix A. Supplementary Analysis on Housing Market Capitalization

Here we present an alternative way to put the program cost and benefits in perspective, which is to assess the degree of housing market capitalization. We estimate the following equation:

$$\begin{aligned} \text{Log Price}_{ijt} = & \beta \cdot \text{Near}_i \times \text{Post}_t \\ & + \alpha \cdot \text{Near}_i + \theta_j + \tau_t + X_i \cdot \gamma + \epsilon_{ijt} \quad (11) \end{aligned}$$

where the dependent variable is the log of transaction price per square meter for unit  $i$ , street  $j$ , and year-month  $t$ .<sup>1</sup>  $\text{Near}_i$  is the distance from the unit's geographical coordinates to the centroid of the closest MMP area (multiplied by negative one).  $\text{Post}_t$  is a dummy for periods after 2012 when the MMP program began.  $X_{ijt}$  includes a rich set of control variables describing (time-invariant) characteristics of the unit and (time-variant) characteristics of the apartment complex that we can observe from the transaction data.<sup>2</sup>  $\theta_j$  and  $\tau_t$  are apartment complex and district-by-year-month fixed effects. We cluster standard errors at the street level.

The key capitalization effect coefficient is  $\beta$ , which quantifies how the price-proximity gradient changed after the MMP policy took place in 2012. The identification assumption is that there are no other contemporaneous shocks that affect the price gradient. This assumption is plausible because the MMP areas are extensively distributed throughout the city, rather than localized to a few areas. It is thus improbable that there are confounding, time-varying policies that would specifically impact such widely dispersed areas.

Appendix Figure A.10 shows estimation results from two versions of equation (11). First, we replace  $\text{Near}_i$  with a series of dummies indicating distance bins in 500-meter increments. The  $\beta$ 's coefficients represent the shift in the price-distance relationship post-2012, compared to 2012, for distances up to 7 kilometers from the nearest MMP area. These findings indicate an approximate 10 percent rise in house prices within a 2km radius of an MMP area, diminishing to

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<sup>1</sup> Streets, or “jiedao”, are administrative units similar to census tracts.

<sup>2</sup> These include fixed effects dummies for transaction type (new sale versus resale), floor level, facing direction, bedroom count, decoration standard, ownership type, and the total floors in the building; continuous variables such as the unit's age and its square, size, floor-area ratio, green space within the complex, property management and parking fees, and the count of housing units and buildings in the complex.

zero as the distance approaches 5km. In a second version of equation (11), we replace the  $\text{Post}_t$  dummy with a series of year dummies. Here the  $\beta$ 's represent an event study of the (log) housing price – distance gradient as a function of time. Results suggest an absence of statistically significant gradient until year 2012 – exactly when the MMP policy took place – when the gradient turns positive and significant. Therefore, both from a spatial and temporal perspective, these result suggest the MMP policy has a positive impact on home prices.

Appendix Table A.11, columns 1, 4, and 7 report the  $\beta$  estimates from the original equation (11), with increasingly stringent controls. The analysis reveals an average increase in home prices of about 3.2 percent for every kilometer closer to an MMP area.

The overall appreciation of housing prices near MMP areas may encompass multiple mechanisms including aesthetic and/or environmental changes. We try to tease out the impact of pollution reduction in two ways. First, in columns 2, 5, and 7, we augment equation (11) by fully interacting the model with a dummy variable indicating whether a housing unit has more MMP areas upwind than downwind, based on the sample-average prevailing wind direction.<sup>3</sup> Given that the influence of urban forests on air pollution predominantly occurs through wind-driven transport (as shown in Section 5), the absence of a notable effect in the downwind direction implies that the observed rise in housing prices is to a larger degree influenced by the localized benefits, such as aesthetic enhancements, than pollution reductions.

Second, in columns 3, 6, and 9, we augment equation (11) by introducing an additional interaction term between the  $\text{Post}_t$  dummy and the predicted percentage reduction in  $\text{PM}_{2.5}$  pollution at housing unit  $i$ 's location (" $\% \Delta \widehat{\text{PM}}_{2.5,i}$ ").<sup>4</sup> The latter term is computed based on the spatial distribution of the MMP planting site, wind directions, and the causal estimates on the impact of upwind vegetation on  $\text{PM}_{2.5}$  from Section 5. Appendix Figure A.11 shows the spatial distribution of this variable. The augmented specification therefore tries to test if, conditional on the impact of being close to MMP sites (" $\text{Near}_i \times \text{Post}_t$ "), having more MMP-induced  $\text{PM}_{2.5}$  pollution reduction at the housing unit's location (due to variability in wind) has additional

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<sup>3</sup> The exact estimation equation is  $\text{Log Price}_{ijt} = \beta \cdot \text{Near}_i \# \text{Post}_t \# \text{Downwind}_i + \theta_j + \tau_t + X_i \cdot \gamma + \epsilon_{ijt}$  where " $\# \#$ " is understood to be the full factorial operator.

<sup>4</sup> The exact estimation equation is  $\text{Log Price}_{ijt} = \beta_1 \cdot \text{Near}_i \times \text{Post}_t + \beta_2 \cdot \% \Delta \widehat{\text{PM}}_{2.5,i} \times \text{Post}_t + \alpha \cdot \text{Near}_i + \theta_j + \tau_t + X_i \cdot \gamma + \epsilon_{ijt}$ .

explanatory power on housing price changes (“ $\% \Delta \hat{P}M_{2.5,i} \times Post_t$ ”). Once again, we do not find statistically apparent evidence on an additional impact from the pollution reduction.

It is important to note other potential explanations of the lack of capitalization of local PM<sub>2.5</sub> benefits. First, the absence of capitalization may reflect fluctuations in wind direction. Even if residents experience a meaningful average reduction in pollution due to upwind tree cover, the day-to-day variation in wind could make it difficult for them to perceive and attribute these improvements. This is consistent with the broader hedonic housing literature, which tends to identify capitalization of medium- or long-run changes in pollution, rather than highly transitory fluctuations (e.g., [Chay and Greenstone, 2005](#)). This provides a compelling explanation for why PM2.5 capitalization might not emerge despite genuine improvements in air quality.

Second, actual PM<sub>2.5</sub> exposure may be less pronounced than predicted reductions due to behavioral or secondary effects, and hence the lack of capitalization. For example, households might reduce their reliance on air purifiers once ambient air quality improves, thereby dampening any perceived benefit. Alternatively, the creation of new green belts or parks could attract tourism and related congestion, introducing additional pollution that offsets the gains from afforestation. Prior studies in other contexts have documented both channels as relevant behavioral responses to pollution (e.g., [Keiser, Lade, and Rudik, 2018](#) ; [Ito and Zhang, 2020](#)).

## Appendix B. Additional Literature Reviewed in Appendix Table A.1

### Panel A: Vegetation and Pollution

Nowak D J, Crane D E, Stevens J C. Air pollution removal by urban trees and shrubs in the United States. *Urban forestry & urban greening*, 2006, 4(3-4): 115-123.

Nowak D J, Hirabayashi S, Bodine A, et al. Modeled PM<sub>2.5</sub> removal by trees in ten US cities and associated health effects. *Environmental pollution*, 2013, 178: 395-402.

Jin S, Guo J, Wheeler S, et al. Evaluation of impacts of trees on PM<sub>2.5</sub> dispersion in urban streets. *Atmospheric Environment*, 2014, 99: 277-287.

Ai H, Zhang X, Zhou Z. The impact of greenspace on air pollution: Empirical evidence from China. *Ecological indicators*, 2023, 146: 109881.

Ren F, Qiu Z, Liu Z, et al. Trees help reduce street-side air pollution: A focus on cyclist and pedestrian exposure risk. *Building and Environment*, 2023, 229: 109923.

Islam A, Pattnaik N, Moula M M, et al. Impact of urban green spaces on air quality: A study of PM10 reduction across diverse climates. *Science of the Total Environment*, 2024, 955: 176770.

Kofel D, Bourgeois I, Paganini R, et al. Quantifying the impact of urban trees on air quality in Geneva, Switzerland. *Urban Forestry & Urban Greening*, 2024, 101: 128513.

Venter Z S, Hassani A, Stange E, et al. Reassessing the role of urban green space in air pollution control. *Proceedings of the National Academy of Sciences*, 2024, 121(6): e2306200121.

### Panel B: Vegetation and Pollen

Katz D S W, Carey T S. Heterogeneity in ragweed pollen exposure is determined by plant composition at small spatial scales. *Science of the Total Environment*, 2014, 485: 435-440.

Maya-Manzano J M, Fernández-Rodríguez S, Smith M, et al. Airborne Quercus pollen in SW Spain: Identifying favourable conditions for atmospheric transport and potential source areas. *Science of the Total Environment*, 2016, 571: 1037-1047.

Charalampopoulos A, Lazarina M, Tsiripidis I, et al. Quantifying the relationship between airborne pollen and vegetation in the urban environment. *Aerobiologia*, 2018, 34(3): 285-300.

Weinberger K R, Kinney P L, Robinson G S, et al. Levels and determinants of tree pollen in New York City. *Journal of Exposure Science & Environmental Epidemiology*, 2018, 28(2): 119-124.

Kasprzyk I, Ćwik A, Kluska K, et al. Allergenic pollen concentrations in the air of urban parks in relation to their vegetation. *Urban Forestry & Urban Greening*, 2019, 46: 126486.

Ciani F, Dell'Olmo L, Foggi B, et al. The effect of urban green areas on pollen concentrations at ground level: A study in the city of Florence (Italy). *Urban Forestry & Urban Greening*, 2021, 60: 127045.

Stas M, Aerts R, Hendrickx M, et al. Association between local airborne tree pollen composition and surrounding land cover across different spatial scales in Northern Belgium. *Urban Forestry & Urban Greening*, 2021, 61: 127082.

Katz D S W, Morris J R, Batterman S A. Pollen production for 13 urban North American tree species: allometric equations for tree trunk diameter and crown area. *Aerobiologia*, 2020, 36(3): 401-415.

### **Panel C: Pollution and ER Visits**

Xu Q, Wang S, Guo Y, et al. Acute exposure to fine particulate matter and cardiovascular hospital emergency room visits in Beijing, China. *Environmental Pollution*, 2017, 220: 317-327.

Aguilera R, Corringham T, Gershunov A, et al. Wildfire smoke impacts respiratory health more than fine particles from other sources: observational evidence from Southern California. *Nature Communications*, 2021, 12(1): 1493.

Ren Z, Liu X, Liu T, et al. Effect of ambient fine particulates (PM2. 5) on hospital admissions for respiratory and cardiovascular diseases in Wuhan, China. *Respiratory Research*, 2021, 22(1): 128.

Ugalde-Resano R, Riojas-Rodríguez H, Texcalac-Sangrador J L, et al. Short term exposure to ambient air pollutants and cardiovascular emergency department visits in Mexico city. *Environmental Research*, 2022, 207: 112600.

Zhou Y, Jin Y, Zhang Z. Short-term exposure to various ambient air pollutants and emergency department visits for cause-stable ischemic heart disease: a time-series study in Shanghai, China. *Scientific Reports*, 2023, 13(1): 16989.

Tian Y, Ma Y, Xu R, et al. Landscape fire PM2. 5 and hospital admissions for cause-specific cardiovascular disease in urban China. *Nature Communications*, 2024, 15(1): 9604.

Zhang Y, Xu R, Huang W, et al. Respiratory risks from wildfire-specific PM2. 5 across multiple countries and territories. *Nature Sustainability*, 2025, 8(5): 474-484.

### **Panel D: Pollen and ER Visits**

Héguy L, Garneau M, Goldberg M S, et al. Associations between grass and weed pollen and emergency department visits for asthma among children in Montreal. *Environmental Research*, 2008, 106(2): 203-211.

Darrow L A, Hess J, Rogers C A, et al. Ambient pollen concentrations and emergency department visits for asthma and wheeze. *Journal of Allergy and Clinical Immunology*, 2012, 130(3): 630-638. e4.

Caillaud D, Thibaudon M, Martin S, et al. Short-term effects of airborne ragweed pollen on clinical symptoms of hay fever in a panel of 30 patients. *J Investig Allergol Clin Immunol*, 2014, 24(4): 249-256.

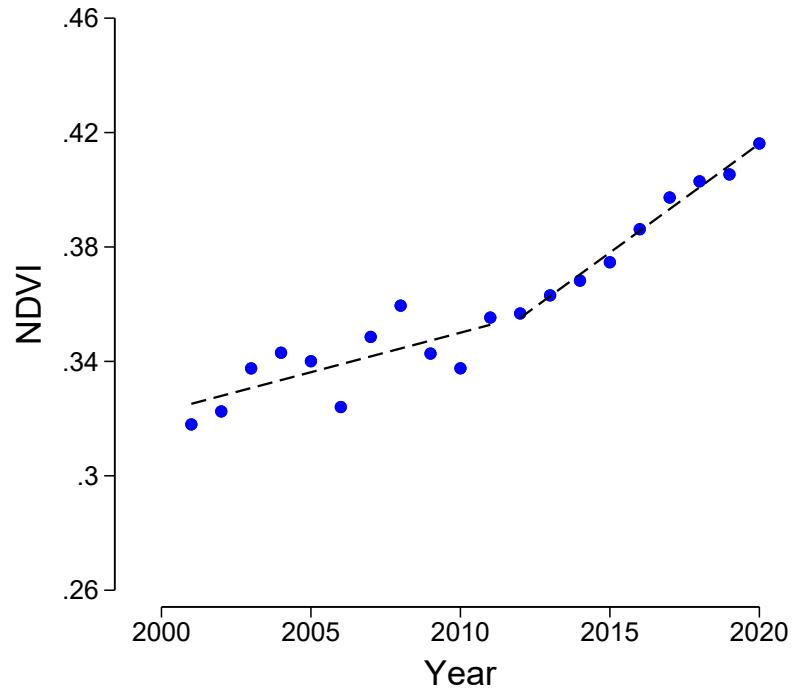
Roblin D W, Jones J W, Fuller C H. Pollen exposure and associated healthcare utilization: a population-based study using health maintenance organization data in the Washington, DC, area. *Annals of the American Thoracic Society*, 2021, 18(10): 1642-1649.

Lappe B L, Ebelt S, D'Souza R R, et al. Pollen and asthma morbidity in Atlanta: a 26-year time-series study. *Environment international*, 2023, 177: 107998.

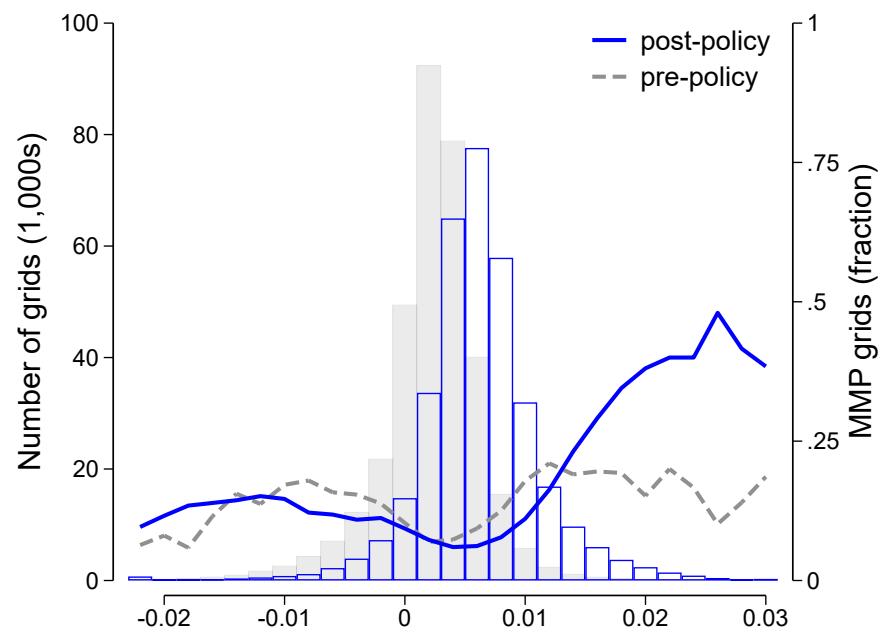
## **Online Appendix Figures and Tables**

Figure A.1. Summary Statistics of Beijing's Vegetation Growth

(a). Annual trends



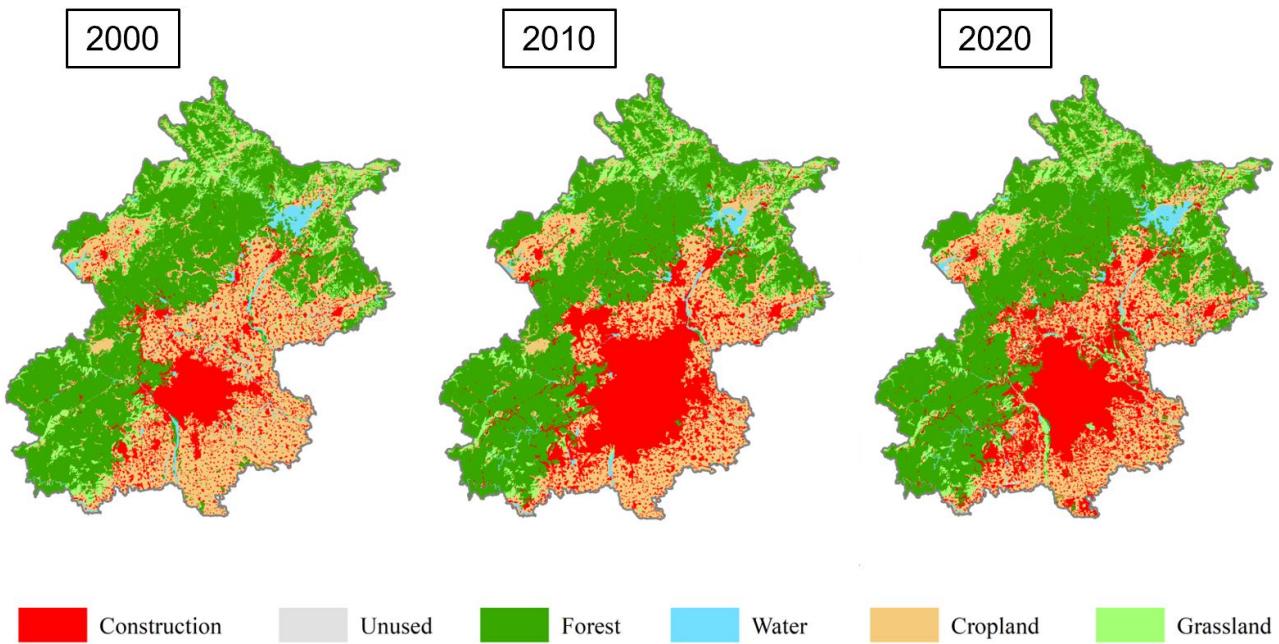
(b). Distribution of grid-level vegetation growth



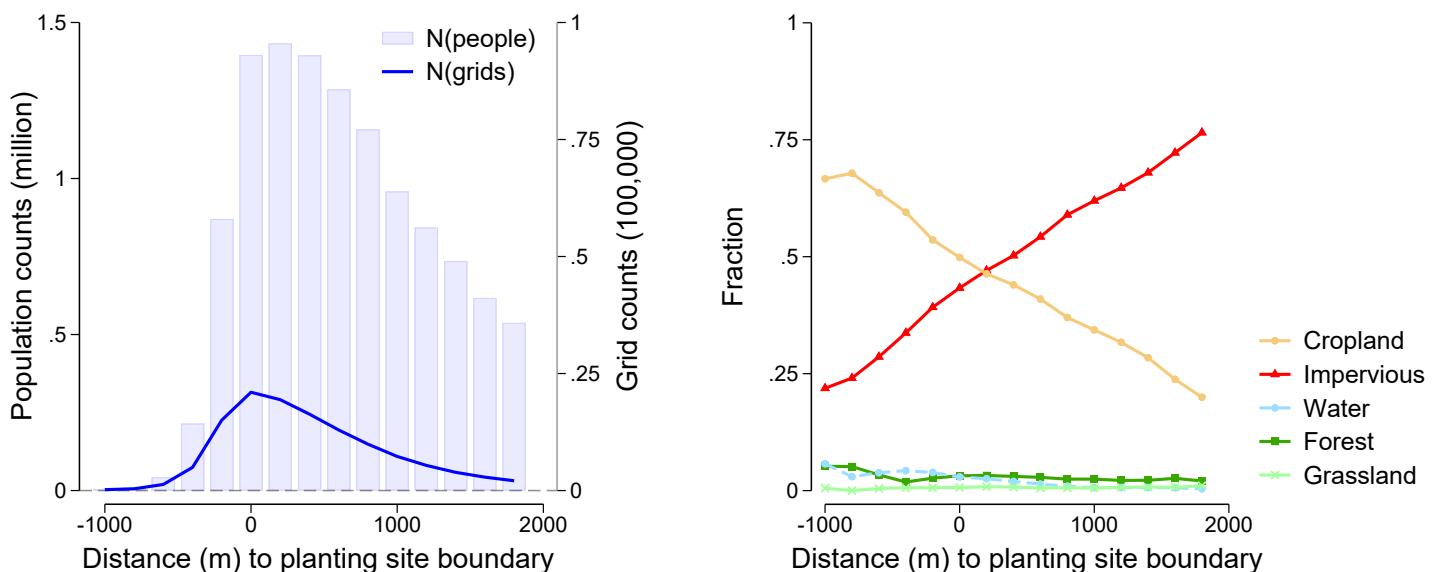
Notes: Panel (a) shows trends in annual average NDVI for the city of Beijing. Panel (b) shows a grid-level distribution of NDVI growth rate in the pre-policy (gray) and post-policy (blue) period. Lines show fraction of grids in the corresponding bins that are MMP during the pre-policy (dashed) and post-policy (solid) period.

Figure A.2. Summary Statistics of Land Use

### (a). Land use changes

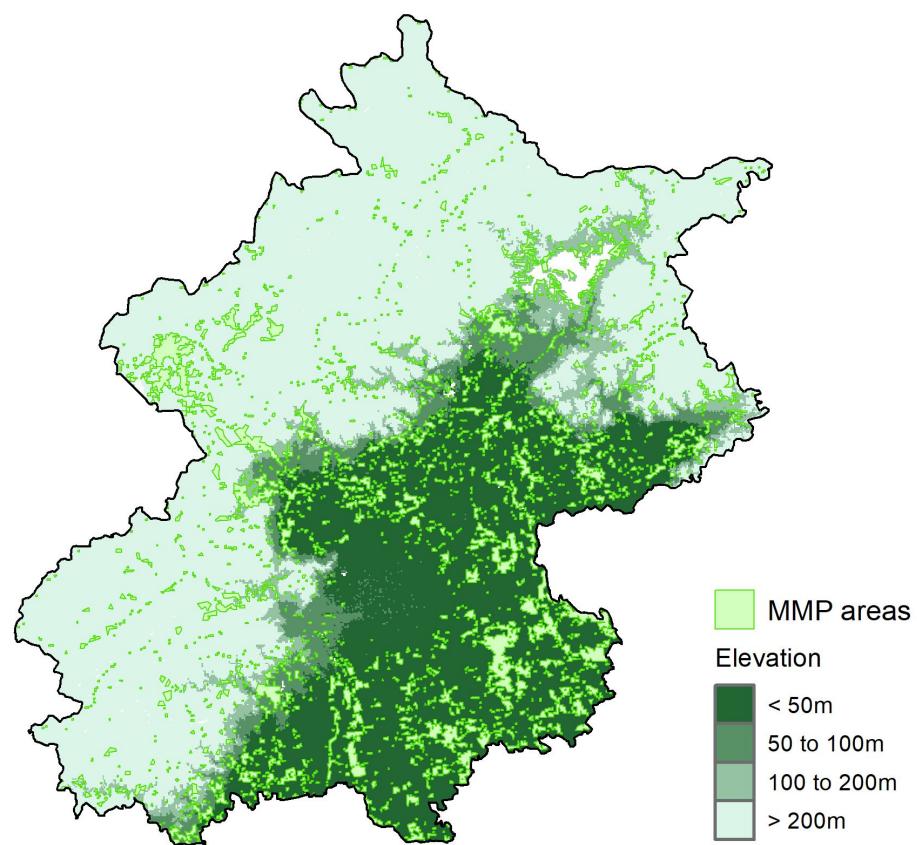


(b). MMP Planting Sites and Baseline (2010) Land Use Categorization



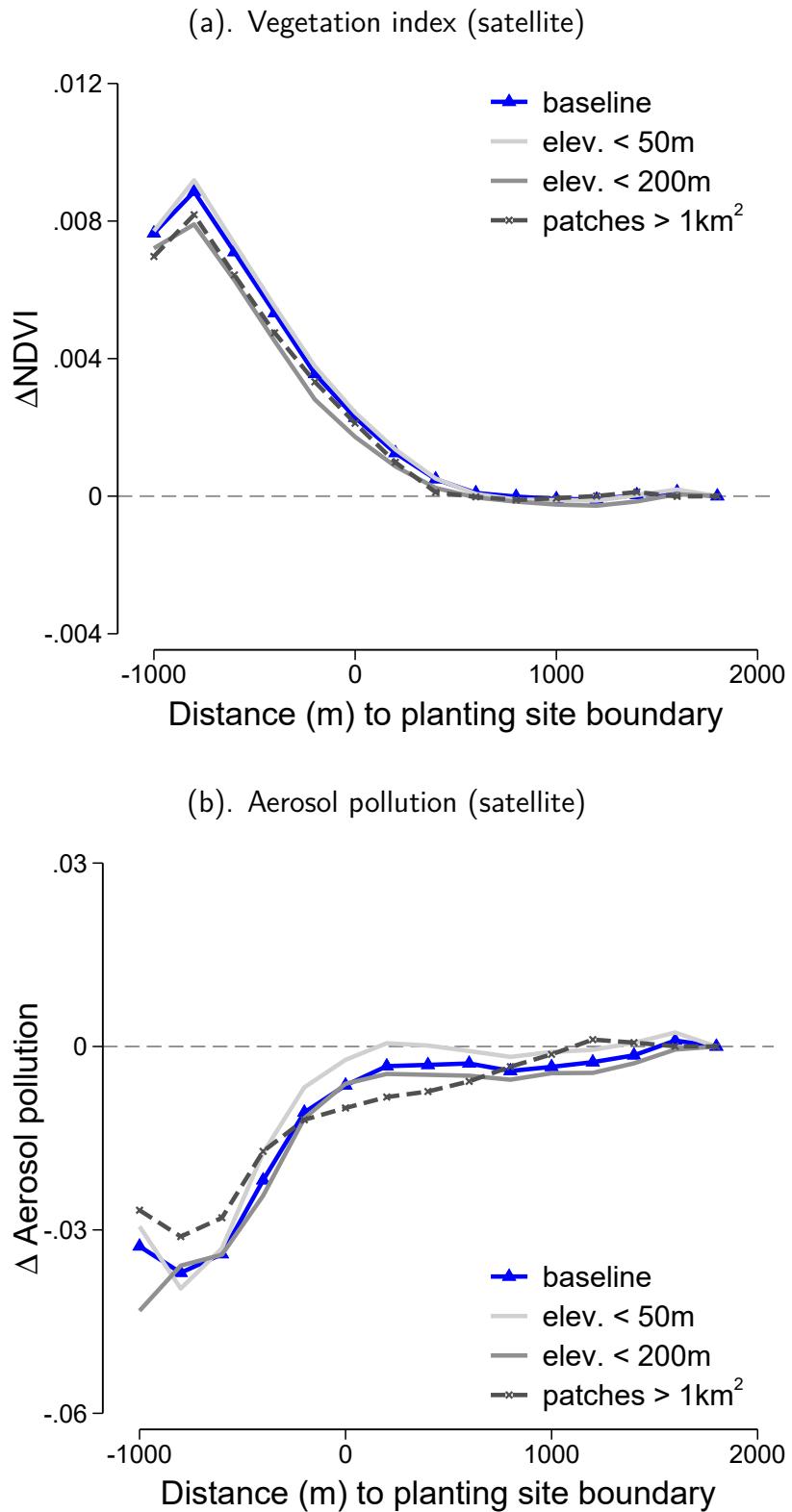
Notes: Panel (a) shows land use categorization in 2000, 2010, and 2020. Panel B shows the distribution of population and grids (left) and the distribution of land use as of year 2010 (right) as a function of distance to the nearest MMP planting site boundary. The graph excludes a very small proportion of the “unused” category.

Figure A.3. Beijing Elevation and MMP Planting Sites



*Notes:* This map shows elevation of Beijing, overlaid with the location of the Million Mu Project (MMP) planting sites.

Figure A.4. Robustness: Changes in Vegetation and Air Quality Near Planting Sites



Notes: This figure shows the annual rate of change in NDVI (panel a) and AOD (panel b) as a function of the distance to the nearest MMP planting site boundary for the post-MMP period of 2012-2020. Each line is from a separate robustness check. These estimates are derived from grid-level, cross-sectional regressions of the annual rate of change on a series of dummy variables indicating distance bins (200-meter increments), with the 1800-2000 meter bin as the omitted category.

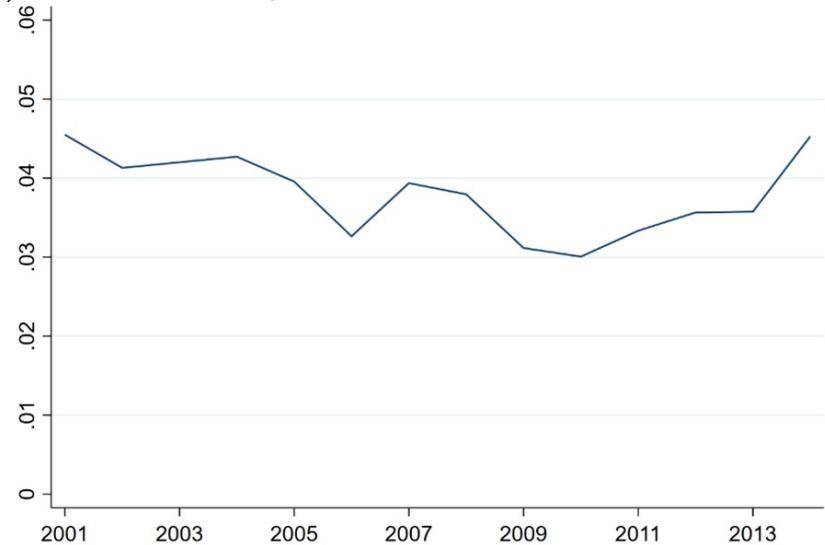
Figure A.5. Variation in Wind Directions in Beijing



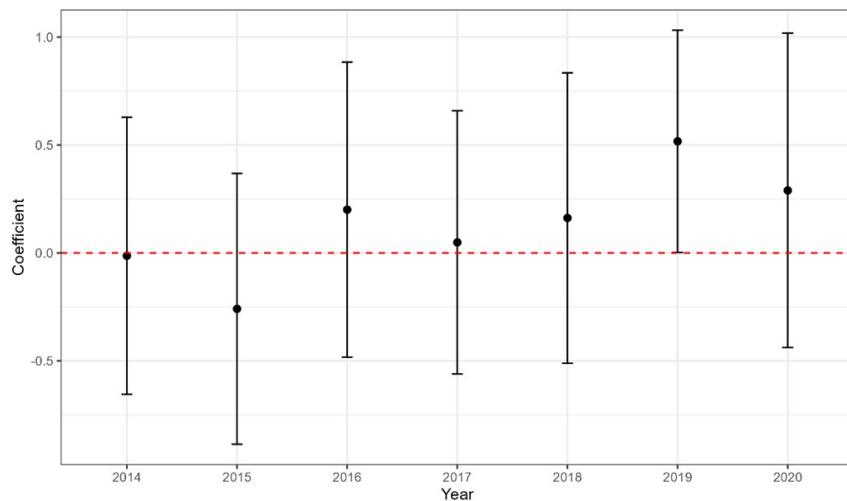
*Notes:* This figure shows the monthly distribution of wind directions in Beijing. Each panel represents a different month, showing the frequency of winds originating from eight cardinal and intercardinal directions (N, NE, E, SE, S, SW, W, NW). The bars are plotted in polar coordinates, with the bar length indicating the relative frequency of winds from each direction. At the day-to-day level, the probability of a change in wind direction on the second day is 70%, and the probability of a change on at least one of the subsequent three days is 95%.

Figure A.6. Changes in Industrial Emission Activities Near MMP Planting Sites

(a). Trends in the Proportion of Pollution Units Within MMP Sites

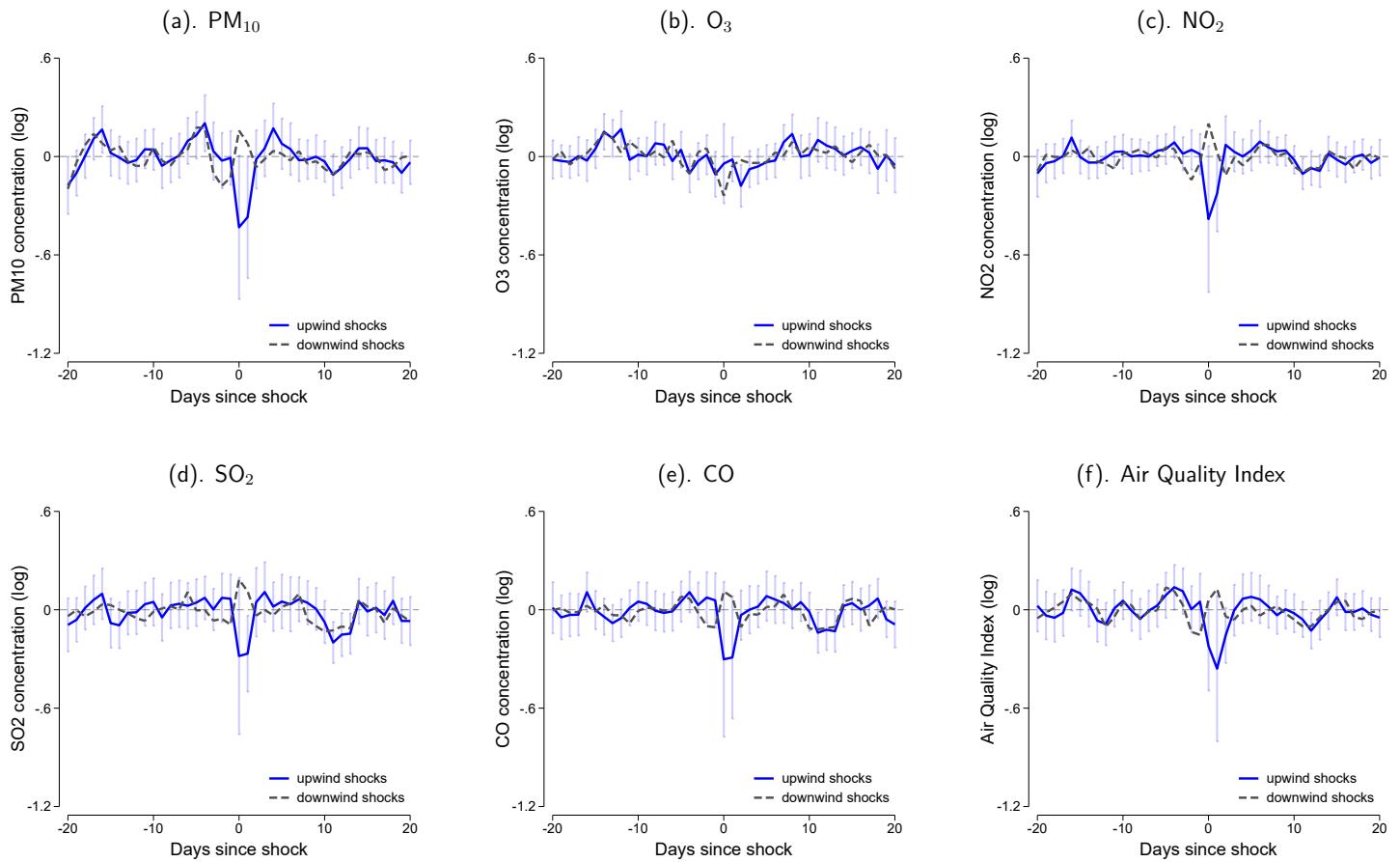


(b). Estimated Differences between Plant Emissions Within and Outside MMP Sites



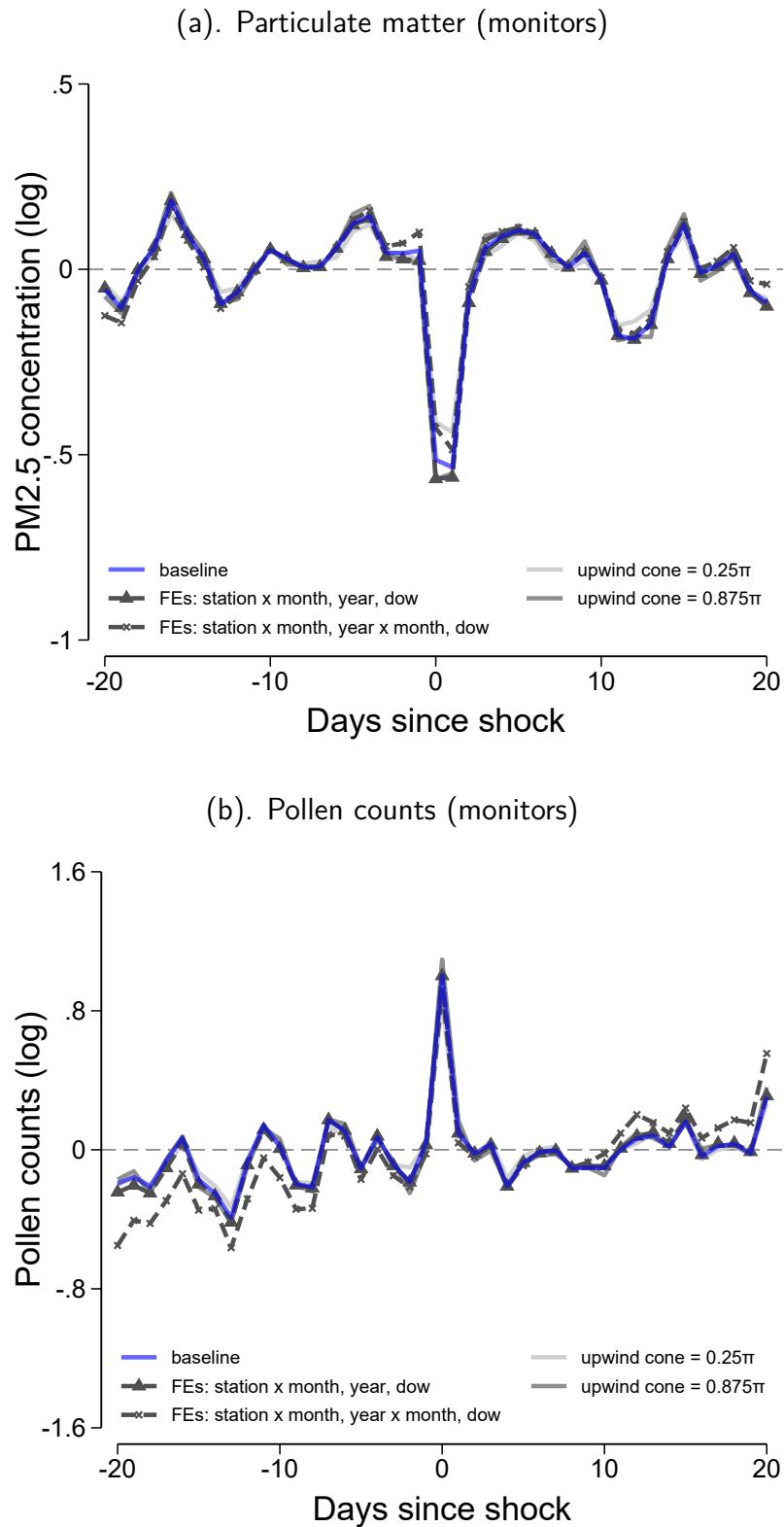
Notes: Panel (a) reports fraction of industrial emission sources in the Environmental Survey and Reporting (ESR) database that locate inside the MMP area from 2001-2014. Panel (b) presents the estimated annual differences in pollution emissions between plants located within and outside the MMP area in Beijing, based on data from the Continuous Emission Monitoring System (CEMS) spanning 2014 to 2020. Each point represents the estimated coefficient from a cross-sectional regression for a given year, with vertical lines indicating 95% confidence intervals constructed using standard errors clustered at the plant level.

Figure A.7. The Effect of Upwind Shocks on Urban Center Pollution: Other Pollutants



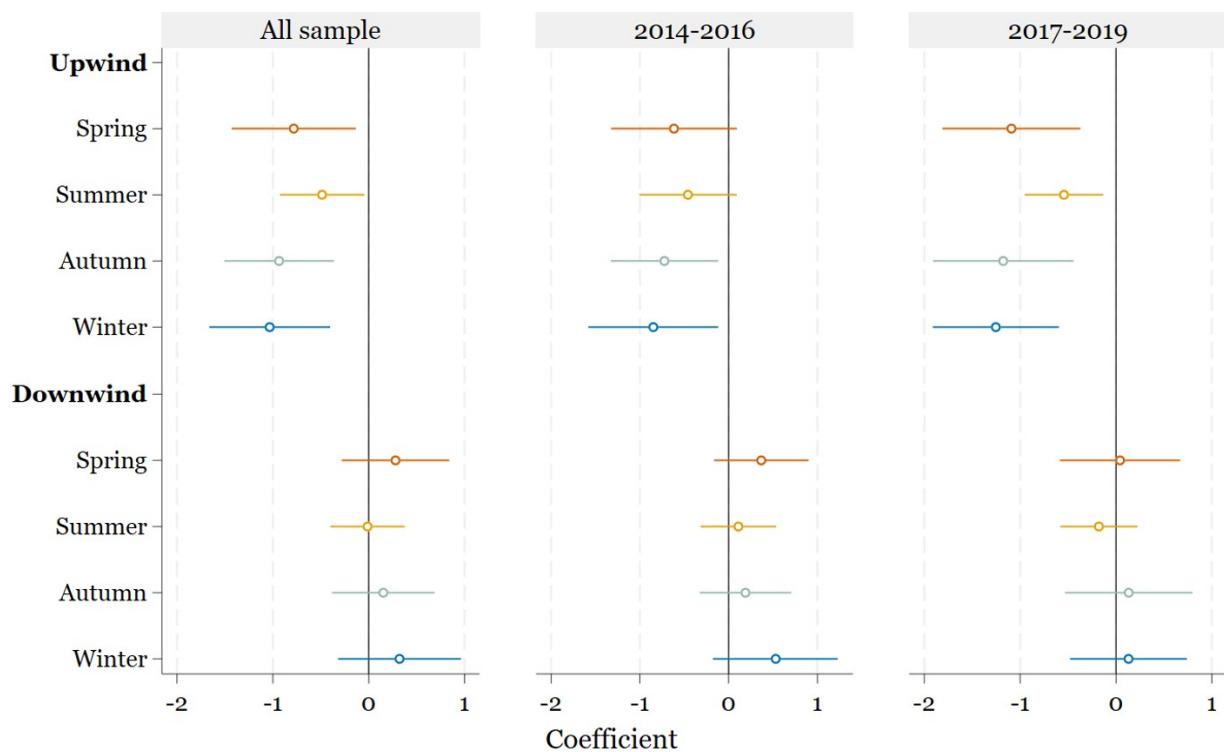
*Notes:* This figure shows regression coefficients of air pollution on 20 lead, contemporaneous, and 20 lag terms of upwind vegetation shocks. In each panel, a set of placebo coefficients is also displayed, obtained by running the same regression but replacing upwind vegetation shocks with downwind shocks. Range bars show 95% confidence intervals constructed using standard errors two-way clustered at both the monitor and the day-of-sample level.

Figure A.8. Urban Center Effects: Upwind Shocks, Particulates Pollution, and Pollen Counts



Notes: This figure shows regression coefficients of air pollution (panel a) and pollen counts (panel b) on 20 lead, contemporaneous, and 20 lag terms of upwind vegetation shocks. Each line is from a separate robustness check.

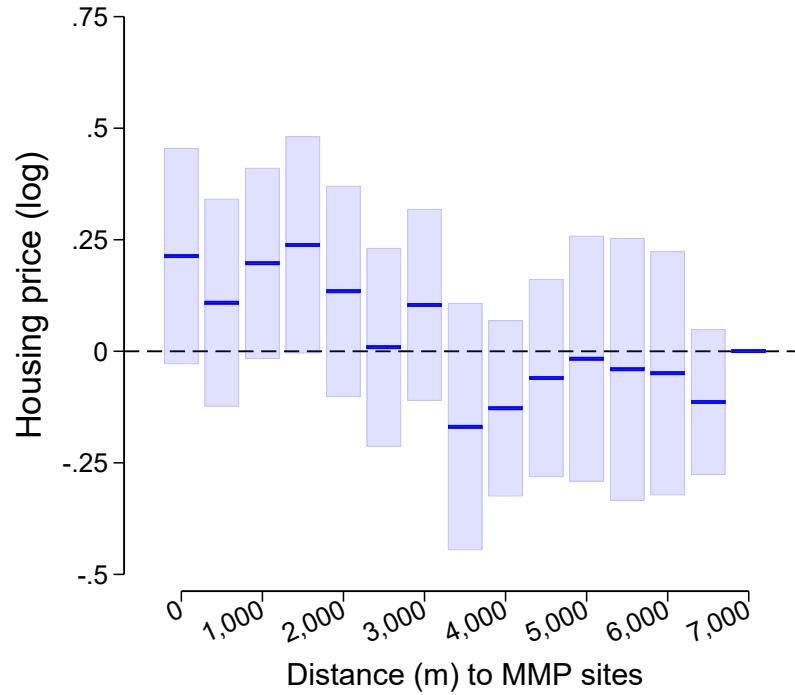
Figure A.9. Upwind Shocks and Particulates Pollution: Effects by Season and Over Time



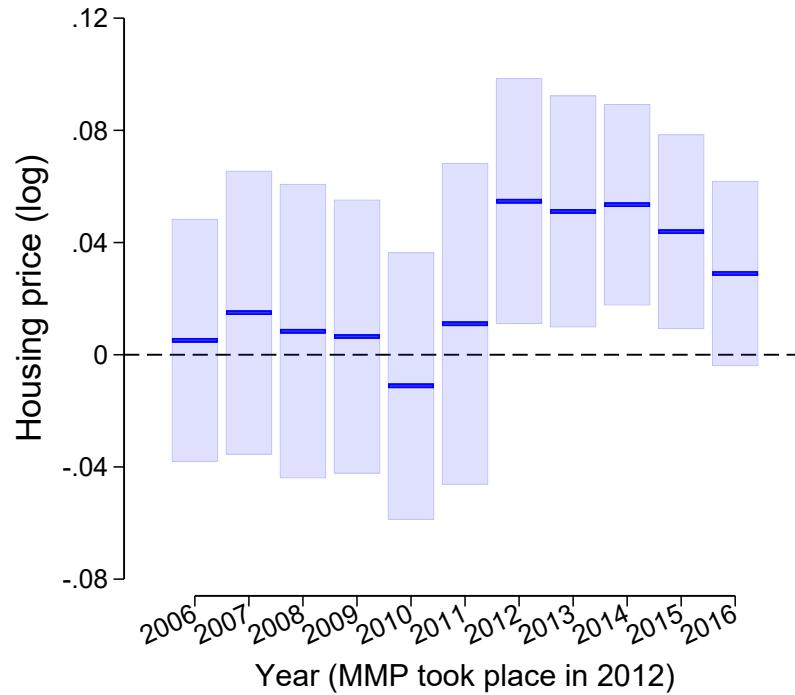
Notes: This figure shows the seasonal effects of upwind and downwind NDVI on log PM<sub>2.5</sub> concentration. All regressions control for monitor fixed effects, year fixed effects, month fixed effects and day-of-week fixed effect. Bars show 95% confidence intervals constructed using standard errors two-way clustered at both the monitor and the day-of-sample levels.

Figure A.10. Housing Price Effects

(a). Post-MMP  $\times$  Distance bins coefficients

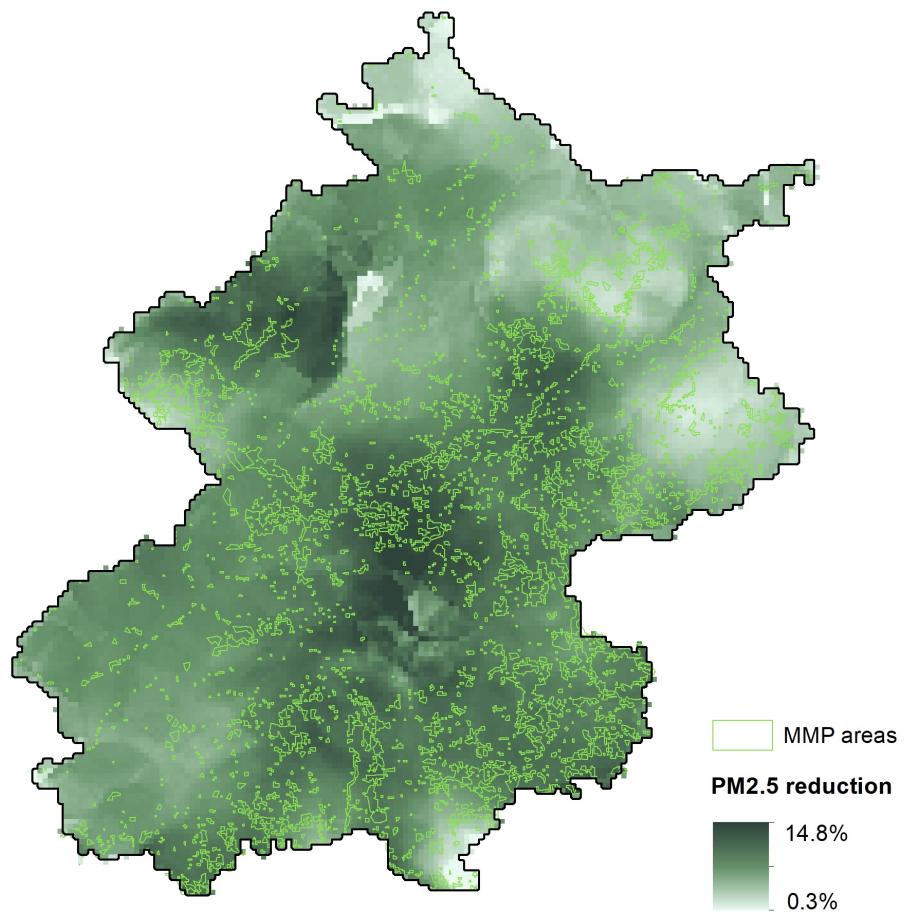


(b). Near-MMP  $\times$  Year dummies coefficients



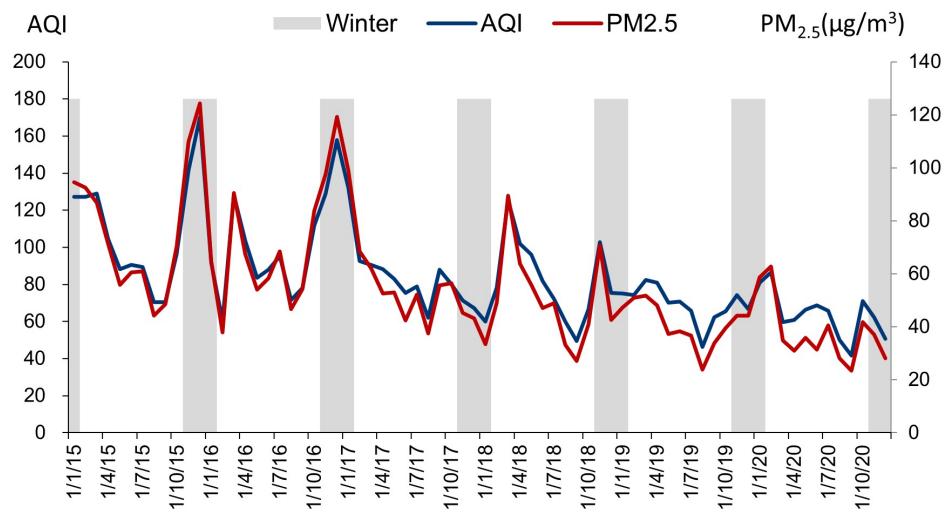
Notes: This figure reports version of the difference-in-differences estimation, where the proximity to MMP variable is grouped into distance bins (panel a) and where the post MMP dummy is separated into years (panel b). Horizontal bars show point estimates and range bars show 95% confidence intervals.

Figure A.11. Predicted PM<sub>2.5</sub> Reduction due to the MMP Policy



*Notes:* This map shows predicted reduction in PM<sub>2.5</sub> concentration due to the Million *Mu* Project (MMP), taking into account the distribution of MMP planting areas (also shown on the map), wind variability, and the causal estimates on the impact of upwind vegetation on PM<sub>2.5</sub>.

Figure A.12. Trends in AQI and PM<sub>2.5</sub> Concentrations in Beijing, 2015-2020



Notes: This figure presents the monthly trends of the Air Quality Index (AQI) and PM<sub>2.5</sub> concentrations in Beijing from January 2015 to December 2020. Both indicators exhibit noticeable seasonal fluctuations and an overall downward trend over the period. The x-axis represents time in years, and the y-axis indicates the level of each respective indicator. Shaded areas represent winters (from November to January next year).

Table A.1. Top Ten Medications Most Frequently Prescribed for Respiratory and Sensory ER Visits

Rank	(1) Respiratory	(2) Sensory
1	ambroxol hydrochloride	levofloxacin eye drops
2	promethazine hydrochloride injection	ofloxacin eye ointment
3	ambroxol hydrochloride injection	rb-bFGF
4	acetaminophen tablets (Tylenol)	tobramycin eye drops
5	doxophylline injection	levofloxacin hydrochloride eye gel
6	compound liquorice tablets	pranoprofen eye drops
7	pseudoephedrine hydrochloride tablets (Sudafed)	erythromycin eye ointment
8	diprophylline injection	emedastine difumarate eye drops
9	ambroxol hydrochloride tablets	levofloxacin hydrochloride eye drops
10	budesonide inhaler (Pulmicort)	sodium hyaluronate eye drops

Notes: This table shows the top ten medications that are most commonly prescribed for respiratory emergencies (column 1) and sensory emergencies (column 2) in the emergency room.

Table A.2. Additional Survey of the Scientific Literature

Publication	Method	Main Finding
<b>Panel A: Vegetation and Pollution</b>		
Nowak et al. (2006)	Modelling	US urban trees remove 711,000 tons pollution yearly, worth \$3.8B, improving air quality.
Nowak et al. (2013)	Modelling	Urban trees remove 4.7–64.5 tons PM <sub>2.5</sub> yearly, worth \$1.1M–\$60.1M, saving up to 7.6 lives.
Jin et al. (2014)	Experiment	Optimal canopy density 50–60% and LAI 1.5–2.0 minimizes PM <sub>2.5</sub> in street canyons.
Ai et al. (2023)	Estimation	NDVI > 0.3 significantly cuts PM <sub>2.5</sub> in China, with stronger effects in southern, higher-level cities.
Ren et al. (2023)	Experiment	Medium tree spacing cuts sidewalk inhalable particles by 59.05% and bicycle lane by 55.85% in Xi'an.
Islam et al. (2024)	Modelling	In 2014, Geneva's trees removed up to 150t O <sub>3</sub> , 66t PM <sub>10</sub> , but emitted 130t BVOCs.
Kofel et al. (2024)	Modelling	≥27% green space reduces PM <sub>10</sub> .
Venter et al. (2024)	Estimation	2010–2019, more green space cut pollution ~0.8% in US, effects varied by scale and vegetation type.
<b>Panel B: Vegetation and Pollen</b>		
Katz and Carey (2014)	Experiment	Vacant lots drive ragweed pollen; 10 m & 1 km land use explains 48% variation.
Maya-Manzano et al. (2016)	Modelling	Wind speed and the distance from monitoring sites to oak woodlands jointly determine pollen concentration.
Charalampopoulos (2018)	Modelling	In Thessaloniki, each <i>Platanus</i> tree increases airborne pollen by 6%, <i>Olea</i> by 5%, <i>Cupressaceae</i> by 0.7%.
Weinberger et al. (2018)	Modelling	A 1% increase in tree canopy cover within 500 m leads to a 1.9% increase in tree pollen.
Kasprzyk et al. (2019)	Experiment	Tree pollen concentration values were not dependent on total canopy volume, but the downtown park exhibited the highest pollen concentration and allergenic potential.
Ciani et al. (2021)	Modelling	Study of 16 Florence sites links urban tree cover to varying pollen concentrations.
Stas et al. (2021)	Modelling	14 sites show meso-scale (1–5 km) land cover drives airborne tree pollen composition.
Katz et al. (2024)	Modelling	Tree planting decisions have a significant impact on pollen production, but only for species that are primarily planted and high producers of allergenic pollen.
<b>Panel C: Pollution and ER Visits</b>		
Xu et al. (2017)	Estimation	A 10 µg/m <sup>3</sup> rise in PM <sub>2.5</sub> increased cardiovascular ER visits by 0.14%, with heart failure visits rising up to 1.21%.
Aguilera et al. (2021)	Estimation	A 10 µg/m <sup>3</sup> rise in wildfire PM <sub>2.5</sub> increases respiratory hospitalizations by 1.3–10%, compared to 0.67–1.3% for non-wildfire PM <sub>2.5</sub> .
Ren et al. (2021)	Estimation	A 10 µg/m <sup>3</sup> rise in PM <sub>2.5</sub> increased hospital admissions by 1.23% for cardiovascular and 1.95% for respiratory diseases in Wuhan.
Ugalde-Resano et al. (2022)	Estimation	A 10 µg/m <sup>3</sup> rise in PM <sub>2.5</sub> increased cardiovascular ER visits by 3.7%, with up to 11% of visits attributable to air pollutants.
Zhou et al. (2023)	Estimation	A 10 µg/m <sup>3</sup> increase in PM <sub>2.5</sub> raised SIHD ED visits by 0.58%, while PM <sub>10</sub> increased visits by 0.50%.
Tian et al. (2024)	Estimation	A 1 µg/m <sup>3</sup> rise in landscape fire PM <sub>2.5</sub> is linked to same-day increases in cardiovascular (0.16%), ischemic heart disease (0.28%), and ischemic stroke (0.25%) hospitalizations.
Zhang et al. (2025)	Estimation	A 1 µg/m <sup>3</sup> rise in wildfire PM <sub>2.5</sub> increases respiratory hospitalization risk by up to 0.79% and accounts for 42.4% of PM <sub>2.5</sub> -related cases.
<b>Panel D: Pollen and ER Visits</b>		
Héguy et al. (2008)	Estimation	Grass pollen increases pediatric asthma ER visits by 1.73% per 10 grains/m <sup>3</sup> at lag 3; first visits up to 2.08–1.91% and readmissions 2.37%.
Darrow et al. (2012)	Estimation	<i>Quercus</i> and <i>Poaceae</i> pollen increased asthma ED visits by 2–3% per SD, with 10–15% higher risk on peak pollen days in Atlanta.
Caillaud et al. (2014)	Experiment	A 10 grains/m <sup>3</sup> increase in ragweed pollen raised odds of nasal symptoms to 1.43, ocular 1.32, and bronchial 1.14.
Roblin et al. (2021)	Estimation	A one SD increase in pollen raised respiratory healthcare use by 1.4–1.8% for visits and up to 12.8% for calls/emails.
Lappe et al. (2023)	Estimation	Each SD increase in pollen linked to 1–8% higher asthma/wheeze ED visits, strongest in 1993–2000, younger, and Black patients.

Notes: Each row represents a study. See Online Appendix for full citations.

Table A.3. Robustness: Wind Shocks, Particulates Pollution, and Pollen Counts

Source of vegetation shock:	(1)	(2)	(3)	(4)	(5)	(6)
	MMP grids (5km rad.)			All grids (5km rad.)		
<b>I. Outcome = Log pollen counts</b>						
Log upwind NDVI	0.640** (0.232)	0.735** (0.288)	0.613*** (0.160)	0.688** (0.253)	0.745** (0.323)	0.473** (0.177)
Log downwind NDVI	0.097 (0.274)	0.082 (0.250)	-0.014 (0.247)	-0.043 (0.258)	-0.067 (0.298)	-0.339 (0.249)
Observations	6,912	6,912	6,912	8,394	8,394	8,394
<b>II. Outcome = Log PM<sub>2.5</sub></b>						
Log upwind NDVI	-0.050 (0.255)	-0.005 (0.262)	0.113 (0.235)	-0.491 (0.338)	-0.555 (0.339)	-0.197 (0.304)
Log downwind NDVI	-0.098 (0.255)	-0.042 (0.269)	0.098 (0.236)	0.064 (0.310)	0.063 (0.327)	0.405 (0.289)
Observations	47,974	47,974	47,974	74,022	74,022	74,022
FEs: monitor	✓			✓		
FEs: month	✓			✓		
FEs: year	✓	✓		✓	✓	
FEs: monitor×month		✓	✓		✓	✓
FEs: year×month			✓			✓
FEs: day-of-week	✓	✓	✓	✓	✓	✓

*Notes:* Each panel-column is a separate regression. Each panel looks at a different outcome variable. Each column uses a different set of fixed effects controls. In columns 1-3, upwind and downwind NDVIs are defined using MMP grids within 5km radius of the monitor. In columns 4-6, upwind and downwind NDVIs are defined using all grids within 5km radius of the monitor. Standard errors are two-way clustered at both the monitor and the day-of-sample levels. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table A.4. Alternative Specifications: Wind Shocks, Particulates Pollution, and Pollen Counts

Source of vegetation shock:	(1)	(2)	(3)	(4)	(5)	(6)
	MMP grids (10km rad.)			All grids (10km rad.)		
<b>I. Outcome = Log pollen counts</b>						
Log upwind NDVI	0.727*** (0.234)	0.730** (0.279)	0.601** (0.210)	0.612** (0.261)	0.660* (0.333)	0.374* (0.194)
Log downwind NDVI	-0.079 (0.247)	-0.063 (0.272)	-0.164 (0.247)	-0.011 (0.252)	0.007 (0.312)	-0.292 (0.186)
Observations	8,385	8,385	8,385	8,394	8,394	8,394
<b>II. Outcome = Log PM<sub>2.5</sub></b>						
Log upwind NDVI	-0.416* (0.230)	-0.425* (0.234)	-0.277 (0.221)	-0.561* (0.294)	-0.649** (0.288)	-0.299 (0.253)
Log downwind NDVI	0.030 (0.222)	0.022 (0.238)	0.127 (0.215)	0.020 (0.262)	-0.033 (0.294)	0.298 (0.271)
Observations	64,854	64,854	64,854	74,022	74,022	74,022
FEs: monitor	✓			✓		
FEs: month	✓			✓		
FEs: year	✓	✓		✓	✓	
FEs: monitor×month		✓	✓		✓	✓
FEs: year×month			✓			✓
FEs: day-of-week	✓	✓	✓	✓	✓	✓

*Notes:* Each panel-column is a separate regression. Each panel looks at a different outcome variable. Each column uses a different set of fixed effects controls. In columns 1-3, upwind and downwind NDVIs are defined using MMP grids within 10km radius of the monitor. In columns 4-6, upwind and downwind NDVIs are defined using all grids within 10km radius of the monitor. Standard errors are two-way clustered at both the monitor and the day-of-sample levels. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table A.5. Alternative Specifications: Wind Shocks, Particulates Pollution, and Pollen Counts

Source of vegetation shock:	(1) MMP grids only (15km rad.)	(2) MMP grids only (15km rad.)	(3) MMP grids only (15km rad.)	(4) All grids (15km rad.)	(5) All grids (15km rad.)	(6) All grids (15km rad.)
<b>I. Outcome = Log pollen counts</b>						
Log upwind NDVI	0.788** (0.303)	0.863** (0.343)	0.799** (0.289)	0.580** (0.261)	0.599* (0.323)	0.239 (0.208)
Log downwind NDVI	-0.215 (0.274)	-0.160 (0.311)	-0.244 (0.238)	0.001 (0.262)	0.020 (0.315)	-0.368* (0.183)
Observations	8,394	8,394	8,394	8,394	8,394	8,394
<b>II. Outcome = Log PM<sub>2.5</sub></b>						
Log upwind NDVI	-0.835*** (0.292)	-0.881*** (0.291)	-0.671** (0.261)	-0.783*** (0.274)	-0.870*** (0.260)	-0.512** (0.235)
Log downwind NDVI	0.384 (0.275)	0.374 (0.296)	0.463* (0.263)	0.202 (0.237)	0.130 (0.289)	0.435 (0.287)
Observations	74,022	74,022	74,022	74,022	74,022	74,022
FEs: monitor	✓			✓		
FEs: month	✓			✓		
FEs: year	✓	✓		✓	✓	
FEs: monitor×month		✓	✓	✓	✓	✓
FEs: year×month			✓			✓
FEs: day-of-week	✓	✓	✓	✓	✓	✓

*Notes:* Each panel-column is a separate regression. Each panel looks at a different outcome variable. Each column uses a different set of fixed effects controls. In columns 1-3, upwind and downwind NDVIs are defined using MMP grids within 15km radius of the monitor. In columns 4-6, upwind and downwind NDVIs are defined using all grids within 15km radius of the monitor. Standard errors are two-way clustered at both the monitor and the day-of-sample levels. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table A.6. Upwind Vegetation, Pollution, and Pollen: Robustness to Wind Direction Bin Fixed Effects

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	Baseline			+ Wind direction FEs			+ Monitor $\times$ wind direction FEs		
<b>A. Outcome = Log pollen counts</b>									
Log upwind NDVI	0.788** (0.303)	0.863** (0.343)	0.799** (0.289)	0.685* (0.338)	0.727* (0.383)	0.703** (0.276)	0.730* (0.353)	0.769* (0.395)	0.733** (0.313)
Log downwind NDVI	-0.215 (0.274)	-0.160 (0.311)	-0.244 (0.238)	-0.052 (0.286)	-0.025 (0.340)	-0.055 (0.244)	-0.106 (0.268)	-0.067 (0.330)	-0.094 (0.218)
<b>B. Outcome = Log PM<sub>2.5</sub></b>									
Log upwind NDVI	-0.835*** (0.292)	-0.881*** (0.291)	-0.671** (0.261)	-0.674** (0.248)	-0.683*** (0.250)	-0.594*** (0.211)	-0.724** (0.280)	-0.737** (0.276)	-0.640** (0.236)
Log downwind NDVI	0.384 (0.275)	0.374 (0.296)	0.463* (0.263)	0.306 (0.237)	0.316 (0.260)	0.266 (0.221)	0.360 (0.265)	0.370 (0.286)	0.311 (0.243)
FEs: monitor	✓				✓			✓	
FEs: month	✓				✓			✓	
FEs: year	✓	✓			✓	✓		✓	
FEs: monitor $\times$ month		✓	✓		✓	✓		✓	✓
FEs: year $\times$ month			✓			✓			✓
FEs: day-of-week	✓	✓	✓	✓	✓	✓	✓	✓	✓
FEs: wind				✓	✓	✓			
FEs: monitor $\times$ wind							✓	✓	✓

Notes: Each panel-column is a separate regression. Each panel looks at a different outcome variable indicated by panel name. Each column uses a different set of fixed effects controls. Standard errors are two-way clustered at both the monitor and the day-of-sample levels. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table A.7. Robustness: Pollen Exposure and Emergency Room (ER) Utilization, Same-Day Outcomes

	(1) All causes	(2) Respiratory	(3) Sensory	(4) Others	(5) All causes	(6) Respiratory	(7) Sensory	(8) Others
<b>A. Baseline specification</b>								
ER visits							ER spending	
All ER	0.0079*** (0.0017)	0.0197*** (0.0043)	0.0411*** (0.0066)	0.0055*** (0.0016)	0.0048*** (0.0015)	0.0141*** (0.0040)	0.0727*** (0.0202)	0.0028 (0.0017)
<b>B. Effects by severity</b>								
ER → Not hospitalized	0.0076*** (0.0017)	0.0197*** (0.0043)	0.0412*** (0.0066)	0.0051*** (0.0016)	0.0048*** (0.0015)	0.0146*** (0.0042)	0.0731*** (0.0199)	0.0027 (0.0017)
ER → Hospitalized	-0.0044 (0.0044)	-0.0012 (0.0033)	-0.0015 (0.0017)	-0.0036 (0.0047)	-0.0164 (0.0279)	-0.0078 (0.0402)	-0.0211 (0.0244)	-0.0020 (0.0315)
<b>C. Effects by age</b>								
Age < 60	0.0081*** (0.0019)	0.0218*** (0.0047)	0.0423*** (0.0066)	0.0053*** (0.0017)	0.0050** (0.0021)	0.0198*** (0.0050)	0.0851*** (0.0194)	0.0028 (0.0021)
Age ≥ 60	0.0069*** (0.0014)	0.0137*** (0.0041)	0.0121** (0.0043)	0.0054*** (0.0016)	0.0055* (0.0025)	0.0211* (0.0100)	0.0466* (0.0216)	0.0042 (0.0032)
<b>D. Effects by prior condition</b>								
Prior respiratory-sensory visits	0.0330*** (0.0038)	0.0313*** (0.0051)	0.0193*** (0.0047)	0.0278*** (0.0047)	0.0322*** (0.0084)	0.0883*** (0.0196)	0.1122*** (0.0274)	0.0429*** (0.0112)
No prior visits	0.0073*** (0.0015)	0.0169*** (0.0037)	0.0328*** (0.0060)	0.0052*** (0.0015)	0.0047*** (0.0013)	0.0133*** (0.0033)	0.0638*** (0.0183)	0.0030 (0.0017)

*Notes:* Each cell represents a separate regression using district-day level data. Each column presents ER records corresponding to different diagnoses. Panel (a) uses all ER records. Panel (b) stratifies by visits that did and did not end up with hospital admissions. Panel (c) stratifies by age of the patient. Panel (d) stratifies by whether the patient had respiratory and sensory visits in the previous 30 days. All regressions control for district-by-month fixed effects, year-by-month fixed effects, day-of-week fixed effects, and holiday fixed effects. Standard errors are two-way clustered at the district and day-of-sample levels. Number of observation for each regression is 8,394.

\*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table A.8. Robustness: Pollen Exposure and Emergency Room (ER) Utilization, Seven-Day Outcomes

	(1) All causes	(2) Respiratory	(3) Sensory	(4) Others	(5) All causes	(6) Respiratory	(7) Sensory	(8) Others
ER visits								
<b>A. Baseline specification</b>								
All ER	0.0042** (0.0014)	0.0149*** (0.0034)	0.0353*** (0.0052)	0.0023 (0.0014)	0.0030** (0.0011)	0.0134*** (0.0023)	0.0288*** (0.0072)	0.0006 (0.0013)
ER spending								
<b>B. Effects by severity</b>								
ER → Not hospitalized	0.0041** (0.0014)	0.0149*** (0.0034)	0.0354*** (0.0053)	0.0021 (0.0014)	0.0030** (0.0011)	0.0138*** (0.0025)	0.0296*** (0.0073)	0.0005 (0.0014)
ER → Hospitalized	0.0036 (0.0084)	0.0042 (0.0087)	-0.0054 (0.0047)	0.0032 (0.0071)	0.0021 (0.0237)	-0.0044 (0.0550)	-0.0939 (0.0566)	0.0074 (0.0249)
<b>C. Effects by age</b>								
Age < 60	0.0049*** (0.0015)	0.0173*** (0.0041)	0.0391*** (0.0054)	0.0027* (0.0014)	0.0036* (0.0017)	0.0173*** (0.0038)	0.0295** (0.0095)	0.0014 (0.0016)
Age ≥ 60	0.0027 (0.0018)	0.0095* (0.0044)	0.0161 (0.0092)	0.0013 (0.0017)	0.0029 (0.0022)	0.0115** (0.0044)	0.0423 (0.0351)	-0.0000 (0.0022)
<b>D. Effects by prior condition</b>								
Prior respiratory-sensory visits	0.0301*** (0.0049)	0.0453*** (0.0091)	0.0480*** (0.0093)	0.0254*** (0.0051)	0.0218*** (0.0059)	0.0430** (0.0172)	0.1204*** (0.0332)	0.0229*** (0.0067)
No prior visits	0.0039** (0.0014)	0.0130*** (0.0034)	0.0297*** (0.0068)	0.0022 (0.0014)	0.0030** (0.0012)	0.0124*** (0.0024)	0.0197* (0.0090)	0.0007 (0.0015)

*Notes:* Each cell represents a separate regression using district-day level data. Each column presents ER records corresponding to different diagnoses. All outcomes are measured using a seven-day look-ahead window (e.g., total number of ER visits today and the next six days). Panel (a) uses all ER records. Panel (b) stratifies by visits that did and did not end up with hospital admissions. Panel (c) stratifies by age of the patient. Panel (d) stratifies by whether the patient had respiratory and sensory visits in the previous 30 days. All regressions control for district-by-month fixed effects, year-by-month fixed effects, day-of-week fixed effects, and holiday fixed effects. Standard errors are two-way clustered at the district and day-of-sample levels. Number of observation for each regression is 8,394. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table A.9. Robustness: Pollen Exposure and Emergency Room (ER) Utilization, Rate Specification

	(1) All	(2) Respiratory	(3) Sensory	(4) Others	(5) All	(6) Respiratory	(7) Sensory	(8) Others
ER visits					ER spending			
<b>A. Baseline specification</b>								
All ER	0.0534*** (0.0100)	0.0270*** (0.0047)	0.0089*** (0.0015)	0.0199** (0.0079)	0.5087 (7.3422)	6.0547** (2.6736)	1.0388** (0.3606)	-6.0384 (5.6884)
Mean dep. var.	11.913	1.4312	0.1505	10.343	4371.1	813.37	30.200	3531.5
<b>B. Effects by severity</b>								
ER Not hospitalized	0.0508*** (0.0102)	0.0269*** (0.0047)	0.0088*** (0.0015)	0.0175* (0.0080)	0.3381 (7.4626)	6.1211** (2.6549)	1.0760** (0.3604)	-6.2872 (5.8025)
Mean dep. var. within group	11.735	1.4279	0.1503	10.168	4381.5	830.55	30.380	3511.7
ER Hospitalized	0.0001 (0.0002)	0.0001 (0.0001)	-0.0000 (0.0000)	-0.0000 (0.0002)	-1.6644 (3.4968)	0.7158 (3.2626)	-0.4406 (1.6700)	-1.5172 (2.0355)
Mean dep. var. within group	0.0315	0.0102	0.00106	0.0207	405.65	225.25	32.555	172.78
<b>C. Effects by age</b>								
Age < 60	0.0514*** (0.0108)	0.0238*** (0.0044)	0.0084*** (0.0015)	0.0215** (0.0072)	7.6124 (4.2846)	6.1653*** (1.4167)	1.1227*** (0.2889)	0.8836 (3.2458)
Mean dep. var. within group	8.6026	0.8806	0.1361	7.5962	2449.7	326.46	21.439	2104.5
Age ≥ 60	0.0020 (0.0066)	0.0032 (0.0020)	0.0005** (0.0002)	-0.0015 (0.0054)	-7.1038 (6.7886)	-0.1106 (2.7854)	-0.0383 (0.1988)	-6.9220 (4.4783)
Mean dep. var. within group	3.3401	0.4140	0.0301	2.8982	1865.4	411.05	10.417	1445.4
<b>D. Effects by prior condition</b>								
Prior respiratory-sensory visits	0.0093*** (0.0014)	0.0029*** (0.0004)	0.0008*** (0.0001)	0.0058*** (0.0012)	2.8152*** (0.7755)	0.8581** (0.2843)	0.1365 (0.0788)	1.8817*** (0.5688)
Mean dep. var. within group	0.2970	0.0585	0.0075	0.2322	124.57	38.135	2.0519	84.898
No prior visits	0.0357*** (0.0067)	0.0167*** (0.0029)	0.0043*** (0.0007)	0.0160** (0.0060)	-1.1262 (5.5287)	3.1161 (2.4200)	0.3724 (0.2235)	-4.3293 (4.0018)
Mean dep. var. within group	8.2686	1.0190	0.0823	7.1741	3217.4	635.90	18.659	2565.5

*Notes:* Each panel-column represents a separate regression using district-day level data. All outcomes are measured in “visits per 10,000 residents” scale using a three-day look-ahead window (e.g., total number of ER visits today, tomorrow, and the day after tomorrow). All regressions control for district-by-month fixed effects, year-by-month fixed effects, day-of-week fixed effects, and holiday fixed effects. All regressions are weighted by district-level population. Standard errors are two-way clustered at the district and day-of-sample levels. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table A.10. Pollen Exposure and Emergency Room (ER) Utilization: Effects over Time

	(1) All causes	(2) Respiratory	(3) Sensory	(4) Others	(5) All causes	(6) Respiratory	(7) Sensory	(8) Others
	ER visits					ER spending		
Log (Pollen)	0.0122*** (0.0038)	0.0304*** (0.0078)	0.0277* (0.0144)	0.0093** (0.0033)	0.0078* (0.0039)	0.0233*** (0.0070)	0.0284 (0.0325)	0.0044 (0.0038)
Log (Pollen) $\times$ 2014	-0.0069 (0.0050)	-0.0081 (0.0083)	-0.0038 (0.0223)	-0.0066 (0.0049)	-0.0088** (0.0034)	-0.0053 (0.0065)	-0.0253 (0.0674)	-0.0090* (0.0041)
Log (Pollen) $\times$ 2015	-0.0121** (0.0050)	-0.0166 (0.0096)	0.0209 (0.0168)	-0.0119** (0.0046)	-0.0067 (0.0053)	-0.0138 (0.0085)	0.0350 (0.0405)	-0.0054 (0.0057)
Log (Pollen) $\times$ 2016	-0.0066 (0.0042)	-0.0123 (0.0075)	0.0377* (0.0173)	-0.0064 (0.0038)	-0.0030 (0.0051)	-0.0102 (0.0087)	0.0371 (0.0381)	-0.0019 (0.0050)

Notes: Each column presents ER records corresponding to different diagnoses in different years compared with 2013. All regressions control for district-by-month fixed effects, year-by-month fixed effects, day-of-week fixed effects, and holiday fixed effects. Standard errors are two-way clustered at the district and day-of-sample levels. Number of the observation for each regression is 8,394. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .

Table A.11. Housing Market Capitalization of the MMP Policy

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Near × Post	0.0320*** (0.0117)	0.0308** (0.0119)	0.0325*** (0.0115)	0.0412** (0.0183)	0.0425** (0.0182)	0.0387** (0.0178)	0.0241** (0.0109)	0.0241** (0.0106)	0.0229** (0.0110)
Near × Post × Downwind		0.0072 (0.0078)			0.0082 (0.0073)			0.0008 (0.0055)	
% $\Delta\hat{PM}_{2.5}$ × Post			0.0149 (0.0095)			0.0250* (0.0132)			0.0082 (0.0077)
FEs: street	✓	✓	✓	✓	✓	✓	✓	✓	✓
FEs: year×month	✓	✓	✓						
FEs: district×year×month				✓	✓	✓	✓	✓	✓
Unit characteristics controls							✓	✓	✓

Notes: Each column is a separate regression. "Near" is housing unit's distance to the nearest MMP boundary times negative one. "Post" indicates years since 2012 when the MMP took place. "Downwind" is MMP areas upwind of the unit minus MMP areas downwind of the unit within 5km radius (and hence the higher the index, the larger the degree the unit sits downwind MMP areas). "% $\Delta\hat{PM}_{2.5}$ " is predicted percentage change in PM<sub>2.5</sub> pollution due to the MMP policy at the housing unit's location. "Unit characteristics controls" include log of unit's age, size, floor-to-area ratio, management fees, and full interactions of these terms, and indicators of floor. Standard errors are two way clustered at the street and the year-by-month levels. \*:  $p < 0.10$ ; \*\*:  $p < 0.05$ ; \*\*\*:  $p < 0.01$ .