### NBER WORKING PAPER SERIES

## THE COVID-19 BABY BUMP: THE UNEXPECTED INCREASE IN U.S. FERTILITY RATES IN RESPONSE TO THE PANDEMIC

Martha J. Bailey Janet Currie Hannes Schwandt

Working Paper 30569 http://www.nber.org/papers/w30569

NATIONAL BUREAU OF ECONOMIC RESEARCH 1050 Massachusetts Avenue Cambridge, MA 02138 October 2022, Revised August 2023

The analyses, interpretations, and conclusions in this manuscript are solely the responsibility of the authors and do not necessarily represent the official views of the National Center for Health Statistics, the California Department of Public Health, or National Bureau of Economic Research. We gratefully acknowledge the generous support of the Institute of Policy Research at Northwestern University in acquiring the restricted California natality data, the Northwestern SESP Dean's Venture Research Fund, and the use of the services and facilities of the California Center for Population Research at UCLA (P2CHD041022). We are grateful to Ron Lee, two anonymous referees, and the editor for helpful comments. We are also indebted to Beatrice Chao, Deniz Gorgulu, and Kelsey Figone for excellent research support.

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The Covid-19 Baby Bump: The Unexpected Increase in U.S. Fertility Rates in Response to the Pandemic Martha J. Bailey, Janet Currie, and Hannes Schwandt NBER Working Paper No. 30569 October 2022, Revised August 2023 JEL No. J0

### **ABSTRACT**

We use natality microdata covering the universe of U.S. births for 2015-2021 and California births from 2015 through February 2023 to examine childbearing responses to the COVID-19 pandemic. We find that 60% of the 2020 decline in U.S. fertility rates was driven by sharp reductions in births to foreign-born mothers although births to this group comprised only 22% of all U.S. births in 2019. This decline started in January 2020. In contrast, the COVID-19 recession resulted in an overall "baby bump" among U.S.-born mothers which marked the first reversal in declining fertility rates since the Great Recession. Births to U.S.-born mothers fell by 31,000 in 2020 relative to a pre-pandemic trend but increased by 71,000 in 2021. The data for California suggest that U.S. births remained elevated through February 2023. The baby bump was most pronounced for first births and women under age 25, suggesting that the pandemic led some women to start families earlier. Above age 25, the baby bump was most pronounced for women ages 30-34 and women with a college education. The 2021-2022 baby bump is especially remarkable given the large declines in fertility rates that would have been projected by standard statistical models.

Martha J. Bailey University of California, Los Angeles Department of Economics Bunche Hall 315 Portola Plaza Los Angeles, CA 90095 and NBER marthabailey@ucla.edu

Janet Currie Department of Economics Center for Health and Wellbeing 185A Julis Romo Rabinowitz Building Princeton University Princeton, NJ 08544 and NBER jcurrie@princeton.edu Hannes Schwandt School of Education and Social Policy Northwestern University 2120 Campus Drive Evanston, Il 60208 and NBER schwandt@northwestern.edu

### I. Introduction

Between 2007 and 2020, the U.S. total fertility rate (TFR), a commonly used measure of the expected number of children over a woman's lifetime, declined from 2.1 to 1.6 (Hamilton et al. 2021).<sup>1</sup> The 2020 figure set a new record for the lowest U.S. fertility rate and prompted widespread concerns about the future of the American family, the strength of the labor force, and the solvency of public programs that rely on the contributions of younger generations.

The COVID-19 pandemic and skyrocketing unemployment rates served to heighten these concerns (Figure 1A). Based on over a century of research describing the negative relationship between fertility rates and unemployment (1-11), demographers expected the pandemic and unemployment to reduce fertility rates significantly. For example, a calculation based on Currie and Schwandt's (8) analysis of the previous four recessions suggests that U.S. births should have fallen by 199,481 more than the pre-pandemic trend (Figure 1B). The red dotted line shows that the 9-month lagged unemployment rate predicts that the decline in birth counts should have been on trend for most of 2020, followed by a dramatic decline in 2021.<sup>2</sup> Early in the pandemic, prominent projections by Kearney and Levine (12, 13) found that there might be as many as 300,000-500,000 fewer births in 2021, if unemployment stayed high. Wilde (14) predicted a decline in childbearing of 15% by 2021—a decline twice as large as during the Great Recession.

By early 2021, data from the Centers for Disease Control appeared to confirm that a baby bust was in progress. These figures were widely reported (15). The *New York Times* noted that,

<sup>&</sup>lt;sup>1</sup> The TFR is the total number of children that a woman would have if she experienced the current period's age-specific fertility rates for the entirety of her childbearing years. While this measure does not reflect the experience of any cohort, it provides a succinct and easily interpretable measure of age-adjusted fertility rates at a point in time.

 $<sup>^2</sup>$  See Materials and Methods for additional details. An alternative version of Figure 1, which plots the predicted changes in births net of the 2015-2019 linear trend and changes in unemployment is shown in Appendix Figures B1 and B2.

"The U.S. Birthrate Has Dropped Again. The Pandemic May Be Accelerating the Decline" (16) and *FiveThirtyEight.com* provocatively asked, "How Low Can America's Birth Rate Go Before It's a Problem?" (17).

But the baby bust quickly evaporated. Kearney and Levine later found that an initial reduction in births of 62,000 in 2021 was followed by a rebound and that birth declines were larger in places with greater economic contractions due to pandemic lockdowns and in places with more COVID deaths (18, 19). Consistent with findings by Cohen (2021), one of their more puzzling findings was that over half of the missing births in 2020 appeared to reflect missing conceptions from the period before the COVID pandemic and lockdowns began (20).<sup>3</sup>

This paper uses microdata covering the universe of childbirths in the U.S. for 2015 to 2021 and from the state of California through February 2023 to examine fertility changes for different demographic groups in the period leading up to the pandemic and up to two years (in the case of California) after the pandemic began. Although California is only one state, it comprises approximately an eighth of the U.S. population and we show that where they can be compared, trends in California match those of the overall U.S. closely. Our main outcomes are seasonally adjusted birth counts and birth rates, and the absolute and percent deviations of these outcomes from their pre-pandemic trends. Trends are estimated separately for each sub-group considered.

Our paper provides two reasons why U.S. birth rates deviated from Figure 1B's projection. First, actual births declined before the unemployment rate could have affected birth rates due in large part to the sharp reduction births in early 2020 to foreign-born women. We find that 60% of the 2020 decline in U.S. fertility rates was driven by sharp reductions in births to foreign-born

<sup>&</sup>lt;sup>3</sup> Cohen noted early in the pandemic that births had begun to decline in California and Florida during the summer of 2020, which is too early to have reflected changes in conceptions due to the onset of the pandemic. (See <u>https://osf.io/z5b46</u>, accessed April 1, 2023).

mothers although births to this group only comprised 22% of all U.S. births in 2019. U.S. births to foreign-born mothers fell by 45,000 more than expected in 2020 and by 39,000 more than expected in 2021, for a total of 85,000 missing births—a statistically significant 5.2% decrease for 2020-2021 relative to a linear projection of the pre-pandemic trend. These declines in fertility rates among foreign-born women can be contrasted with a substantial increase in fertility rates among U.S.-born women.

A second reason for the deviation of actual birth rates from predicted birth rates is that births to U.S-born women surged much more than predicted by the historical relationship between U.S. fertility rates and changes in unemployment rates. The number of births to U.S.-born mothers fell by 31,000 in 2020 relative to a pre-pandemic linear trend. This decline represents only a 1% fall in births in this group and is not statistically different from the pre-pandemic linear groupspecific trend. In contrast to the declining fertility rates observed after previous recessions, births then increased by 71,000 among U.S.-born mothers in 2021, for a net gain of 40,000 births in 2020-2021.

The TFR for U.S.-born mothers was a statistically significant 5.1% higher than a prepandemic linear trend by the end of the 2021. This increase is in sharp contrast to the large decline that models based on previous recessions would have predicted (see the discussion of Figure 1B above). This 2021 "baby bump" among U.S.-born women represents the first substantial reversal in U.S. fertility rates since the 2007 Great Recession and is large enough to reverse two years of post-2007 fertility decline. The baby bump was most pronounced for first births and among women under age 25, suggesting that the pandemic led many women to start families sooner. The baby bump was also pronounced among women 30-34 and among those ages 25-49 with a college degree or more. While it is difficult to say what the long-term effect of these changes will be, data through February 2023 in California are potentially informative. Although the levels are lower, trends in the TFR for California track those in the rest of the U.S. through 2021, and the increase in births to U.S.-born women continued in California through February 2023. Using the sustained increase in California to project U.S. fertility rates through February 2023 suggests that there may have been an additional increase of 130,000 births to U.S.-born mothers between January 2022 and February 2023 relative to the pre-pandemic trend. Births to foreign-born mothers had returned to trend levels by the end of 2022.

### II. Results: The COVID-19 Baby Bump

Figures 2A and 2B plot seasonally adjusted TFRs and birth counts for all U.S.-born and foreign-born mothers from 2015 to 2021 as well as TFRs for both groups giving birth in California through February 2023 (2022 data for the entire U.S. are not yet available).<sup>4</sup> The solid blue line presents the pre-pandemic linear trend for all U.S. births and the dashed line shows its projection through December 2021. For U.S.-born mothers, the patterns in the TFR (left vertical axis) and birth counts (right vertical axis) track one another closely. The TFR for U.S.-born mothers in California has been falling more quickly and was at a lower level in January 2021 than in the U.S. overall.

Contrary to the well-known statistical relationship between unemployment and fertility rates described above, neither the TFR nor total births among U.S.-born women show the strong decline predicted by the spike in unemployment seen in Figure 1. Despite the substantial increase

<sup>&</sup>lt;sup>4</sup> Appendix Figure A1 shows patterns for all U.S. (not stratified by mothers' nativity). All U.S. births began to decline around March 2020 and continued to further deviate from trend through January 2021. The pattern for the TFR is very similar. In addition, the TFR for California which is available through February 2023 shows a similar time-series pattern.

in unemployment, Figure 2A shows that births to U.S.-born women remained mostly within a 95percent confidence interval around the pre-pandemic linear trend throughout the first nine months of 2020, falling outside that interval only briefly in August 2020 and again in November 2020 and January 2021.

After January 2021, the TFR and birth rates to U.S.-born women increased. By mid-2021, the measures exceeded the level implied by their pre-pandemic trend. By the end of 2021, both the TFR and birth counts exceeded their pre-pandemic trend by around 5% (Figure 2C). This pattern suggests that conceptions increased sharply as early as May 2020 and that this increase in conceptions persisted until at least March 2021.

The 2021 increase in births among U.S.-born mothers above pre-pandemic trends exceeded the 2020 decline. While births to U.S.-born women in 2020 were 31,104 below the pre-pandemic trend, they exceeded the trend in 2021 by more than twice this amount—70,865 births—resulting in a net increase of around 40,000 births over the entire 2020-2021 period (Table 1B).

Appendix Figure A2 shows that using a linear pre-pandemic trend that is estimated from 2015 through the end of 2019 is a conservative choice when measuring the pandemic baby bump. Estimating non-linear trends over a longer period going back further in time or including births up to November 2020 (nine months after the latest pre-pandemic conceptions) yields a smaller shortfall in births to U.S.-born women in 2020 and a larger baby bump in 2021. Alternative trend choices are further discussed below.

As shown in Figures 2C and 2E, the deviations from trend tend to be similar in California and the U.S. though the California deviations are higher at the end of 2021. California's TFR remained higher by an average of 4.15% from January 2022 and through the end of February 2023. If patterns among U.S.-born mothers in the U.S. remain as elevated as in California—as they did before, during, and after the pandemic—then the national baby bump among U.S.-born mothers will be seen to have extended through February 2023. A continued positive deviation of 4.15% in 2022 and through February 2023 would imply around 134,000 additional births among U.S.-born mothers over that time period relative to a linear pre-pandemic trend.<sup>5</sup> The total estimated net increase among U.S.-born mothers since the beginning of the pandemic through the first quarter of 2023 would then amount to an excess of up to 174,000 births relative to pre-pandemic trends (134,000 projected additional births from January 2022 to February 2023 plus the 40,000 births for 2020 to 2021 from Table 1).

In contrast to births among U.S.-born mothers, Figure 2B shows that both the TFR and births to foreign-born women fell sharply beginning in the early months of 2020. Births to foreign-born women continued to fall until early 2021, after which they recovered, bringing births among foreign-born women back to the pre-pandemic trend by the end of 2021. The TFR for foreign-born women also reverted to close to its pre-pandemic trend.<sup>6</sup> Data for California suggest that fertility for foreign-born women may have begun to rise above trend in 2022, although not by enough to offset declines in 2020 and 2021.

Table 1B shows that births to foreign-born women fell by more than 45,000 in 2020. Consequently, around 60% (45,000/76,000) of the reduction in overall U.S. births during the first nine months of the pandemic in 2020 was due to a decline in births to foreign-born women even though they accounted for 22% of total births in 2019. In 2021, births to foreign-born women

<sup>&</sup>lt;sup>5</sup>Note that the estimated 1/2022-2/2023 deviation of around 134,000 births is substantially larger than the estimated 2021 deviation of around 71,000. This is because the 12 months of 2021 included two months with negative birth deviations while the later period consists of 14 months which all experienced positive deviations.

<sup>&</sup>lt;sup>6</sup> The TFR for foreign-born women is rising while the birth counts are falling which suggests that the number of births to foreign-born women is falling more slowly than changes in the population counts of foreign-born women would suggest.

continued to remain 4.7% lower, amounting to 39,000 fewer births than predicted by the prepandemic linear trend.

Figures 2D and 2F break these patterns down by country of origin for foreign-born mothers. Although births to women born in China typically make up less than 5% of births in the U.S., births to this group fell precipitously after January 2020, falling by proportionally more than among other groups of foreign-born mothers. This steep decline exceeded 40% by January 2021, and unlike births to other foreign-born mothers, births to women born in China remained consistently low through December 2021.

Although the percent decline in births was largest among mothers born in China, the largest absolute change in births to foreign-born women is due to reductions in childbearing among women born in Mexico and Latin America (Table 1C). These women accounted for 54% of all births to foreign-born mothers in 2019, and for 12% of total U.S. births. These births began to fall early in the pandemic and then started to rise again in January 2021, exceeding the pre-pandemic trend by mid-year.

Table 2 (and Appendix Figures A3-A7) breaks down the baby bump for U.S.-born women by demographic group to uncover clues about the mechanisms driving these patterns. Prepandemic linear trends are estimated for each of the subgroups separately, and deviations from linear trends are calculated relative to each group's pre-pandemic trend (see details in the Material and Methods section below). Appendix Figures A3-A7 plot birth rates by month and their percent deviations from trend by age group, birth parity, race/ethnicity group, marital status, and educational attainment. The panels showing deviations from trend also plot the data for California through February 2023. One can see that in most cases, the patterns for 2021 continued through 2023 in California, suggesting that they may also have continued for the U.S. as a whole. Childbearing responses to the pandemic among U.S.-born women varied sharply by age group. Table 2B shows that the two youngest age groups did not experience a baby bust at all in 2020: Births in these groups exceeded the pre-pandemic trend by 9,000, or 4%, in 2020. In 2021, births to 15–24-year-olds exceeded the trend prediction by 28,000. For 25-29- and 30–34-year-olds, births fell insignificantly in 2020 relative to the pre-pandemic trend. In 2021, women ages 25-29 saw a decrease of 1.7%, whereas those aged 30-34 saw an increase of 6% relative to trend. Aggregating over both 2020 and 2021, women aged 25-34 had around 5,000 more births than would have been expected based on linear pre-pandemic trends.

Births to women aged 35 and older show a modest decline relative to pre-pandemic trends in 2020. These dips, however, quickly reversed, and older women's childbearing showed large jumps in 2021—for women 40 and older, the birth rate peaked at 13% above trend by mid-2021. Table 2B summarizes these changes in 2020 and in the first post-pandemic year, showing that net decreases in childbearing for 25-29- and 35–39-year-olds were offset by gains for other groups, resulting in a net 40,000 births relative to the pre-pandemic trend by the end of 2021. The California data suggest that the age groups that gained births continued to do so in 2022.

Table 2C shows that childbearing responses to the pandemic vary by parity. The COVID-19 baby bump was largely driven by first and, to some extent, second births rather than by higherorder births. Although second and higher-order births among U.S.-born women declined in 2020, first births rose slightly relative to the pre-pandemic trend. In 2021, first births rose by 5.3% and second births by 3.0%, whereas higher-order births fell by 1.4%. Appendix Figure A4 shows a spike in births in February 2021, suggesting that large increases in the number of these conceptions had already begun by May 2020. In contrast, births of parity three and higher declined through December 2020 before returning to their pre-pandemic trend by mid-2021. The California data suggest that second order births may also have returned to trend by the end of 2022, though first births remained elevated.

Table 2D explores differences in childbearing responses to the pandemic by race/ethnicity for U.S. born women. The results show that the COVID-19 baby bust in February 2021 was largest among non-Hispanic Black and Hispanic/Latina women (see Appendix Figure A5), and that the baby bump over the rest of 2021 occurred primarily among non-Hispanic White and Hispanic/Latina women as well as Asian/Pacific Islander women. In 2021, non-Hispanic White, Non-Hispanic Asian/Pacific Islanders, and Hispanic/Latina women gave birth to 3.9%, 5.1%, and 3.5% more children than expected relative to the pre-pandemic trend, and these groups' birth rates remained elevated at the end of 2021. In contrast, births among Non-Hispanic Black women declined by -3.3% more than the pre-pandemic trend through 2021 and remained below the pre-pandemic trend through 2021. The California data suggest that births to non-Hispanic White women, Hispanic/Latina women, and Asian/Pacific Islander women remained elevated in 2022.

Table 2E breaks down childbearing responses to the pandemic by marital status. California no longer includes marital status on the birth certificate, so these calculations are missing approximately one eighth of the population and we have no data for 2022. Over our sample period, approximately 40% of U.S. births were to unmarried mothers. Married women show both a larger (and earlier) baby bust in 2020 as well as a larger baby bump in 2021.<sup>7</sup> In contrast, unmarried women show little evidence of a baby bust (except in January 2021) and also less of a baby bump. One important caveat to these results is that marriage rates also changed in response to the pandemic. One study estimates that marriage rates fell by 12% in 2020 relative to what would have

<sup>&</sup>lt;sup>7</sup> We also examined marital and non-marital births by age and found that within five-year age groups, trends were similar for married and unmarried women except for 25–29-year-olds. In this age group, births to married women fall during 2020 and then rise above trend. However, unmarried women in the 25–29-year-old age group saw declines in births in 2020 which then flattened out and had not returned to trend by December 2021.

been expected (21). Many of these marriages appear to have been postponed until they could be safely held. If marriages were postponed but childbearing was not, this could cause a dip in marital births and an increase in non-marital births which could offset a downturn in the non-marital fertility rates among women who were not postponing weddings.

In the last section, Table 2F reveals striking differences by mother's educational attainment. Because college education was itself affected by the COVID-19 pandemic, the sample is restricted to women aged 25-49. Appendix Figure A7 shows that women with four or more years of college showed little evidence of a baby bust in 2020 relative to their pre-pandemic trend, but their birth rates surged in 2021, exceeding the trend by 6.9%. California data suggest that these rates continued to be elevated in 2022. In contrast, women aged 25 and over with less than four years of college education saw declines in birth rates relative to trend in 2020 but returned to their pre-pandemic trend by the middle of 2021. Women with less than a college education did not experience a baby bump.

### **III. Discussion**

This paper starts from the fact that in 2020 U.S. fertility rates declined by much less than would have been predicted by standard statistical models given the large changes in unemployment and uncertainty created by the pandemic. This paper's first contribution is to show that the declines in childbearing in 2020 reflect disproportionally large declines in births among foreign-born mothers. A second contribution is to show that there was a large and unexpected increase in childbearing in 2021 among U.S.-born women, which we break down by population subgroups. Data for California suggests that the increases in U.S. births to U.S.-born women are likely to have continued into February 2023. A third contribution is that we show that the timing of the 2020 baby bust was different for foreign-born and U.S. born mothers. For the former, births began to

decline in January 2020, while for the U.S.-born declines, in addition to being small in percentage terms, did not become apparent until many months after the pandemic began.

The COVID-19 recession is the first in recent history that was not followed by a baby bust. In a resounding rejection of this well-documented statistical relationship, the 2020 COVID-19 recession was followed by a surge in births, which marked the first substantial reversal in declining fertility rates since the Great Recession. The U.S. experience does not appear to be unusual. A review of 22 countries' fertility rates after the COVID-19 pandemic finds Finland and South Korea show evidence of increased fertility rates, 14 countries show no significant change, and six countries (primarily in the Mediterranean and East Asia) exhibit reduced fertility rates (22). Using the data for California, which has trends that track the overall U.S. data closely, we project that U.S. births to U.S.-born women remained elevated through February 2023. Importantly, the increases in fertility among U.S.-born women not only made up for declines in the initial stages of the pandemic but resulted in a net increase of 40,000 births for the 2020-2021 period, and possibly with an additional 134,000 births from January 2022 to February 2023. If and when fertility rates return to their long-term trend, there is likely to be a larger pandemic-era birth cohort going forward.

One reason for the reversal of the standard relationship between unemployment and fertility is that the COVID-19 recession was unlike previous downturns. In an unprecedented response to job losses affecting 22 million workers, the federal government spent \$650 billion in federal pandemic unemployment benefits between March 2020 and September 2021 (23). These programs might have mitigated the fertility reductions that are usually observed in response to job losses (24). The U.S. Census Bureau reports that, as a result of these programs, poverty fell in 2020 in every race and age group (25). Moreover, the increase in unemployment was short-lived relative

to previous recessions. By December 2020 unemployment had fallen from a high of 14.7 percent to a still elevated 6.7 percent. Hall and Kudlyak note that the "unprecedented burst of temporary layoffs early in the pandemic [was] followed by their rapid reversal from April to November 2020"... [however] We show that, after we account for the unusual surge in temporary layoffs, the unemployment pattern in the current recovery is actually similar to the past" (26). In addition to the unusual number of temporary layoffs, other changes included the fact that more affluent families saw increases in the value of their assets as both stock markets and home prices soared (27). And employment opportunities for women dried up as the "shecession" affected the services industries employing women more than workers in traditionally male employments like manufacturing (28).

The standard economic model of fertility emphasizes both the income effects of recessions and the impacts of changes in the opportunity cost of childbearing (29). In the case of COVID-19, both reduced employment and greater work from home as well as pandemic aid may have increased childbearing for certain groups. For example, the reduction in in-person employment may have reduced the opportunity cost of having a child while additional income from pandemic aid loosened budget constraints.

Working backward from the timing of births implies that increases in conceptions among U.S.-born women started about the time that the first economic stimulus checks arrived in Americans' bank accounts. More generally, many women saw little income loss and even income gains due to pandemic support programs, which *ceteris paribus* would tend to raise fertility rates.

Research into changes in fertility rates after the pandemic suggests a number of additional reasons for changes in childbearing, including changes in pregnancy desires, the use of contraception, intimacy patterns, relationship status, and maternal stress. Survey data show that

many U.S. women planned to delay or avoid childbearing due to the pandemic and that these planned delays were especially pronounced among lower-wage women and racial minority women, who experienced greater economic hardship and more COVID-19 cases (30, 31). On the other hand, the college-educated women we examine are older (over 25) and may have had less scope to delay childbearing in the face of continuing pandemic-related uncertainty. On the other hand, the effect of pandemic lockdowns themselves may have encouraged childbearing, similar to research that finds that severe storm warnings induced spikes in births nine months later (32, 33).

Researchers have also used GoogleTrends data to understand changes in pregnancy desires with mixed results. One study finds little evidence of changes in Google searches related to fertility (34), but another predicts based on Google keyword searches relating to conception and pregnancy, that Black women and women with less than a college education were more likely to experience declines in childbearing—a prediction this paper substantiates (14). There may also have been biological factors associated with reduced fertility conditional on people's desires. For example, some evidence shows that maternal stress reduces the likelihood of childbirth nine months later by reducing the release of pregnancy supporting hormones in the early months of gestation (35).

Offsetting declines in desired childbearing and biological effects of stress is the reduction in access to reproductive health care services and abortion. In some cases, reproductive health services completely shut down (36), which may have increased unintended pregnancies at younger ages by reducing access to contraception (37) as well as reducing fertility rates at older ages by affecting access to ART (38). Appendix Figure A8 shows a reduction in the share of births to mothers using ART and in the number of multiple births (which are associated with the use of ART). Breakdowns for U.S.-born women by demographic group provide clues about which mechanisms may have been operating. The baby bump was most pronounced for first births and women under age 25, which suggests the pandemic led some women to start their families earlier. Above age 25, the baby bump was pronounced for women ages 30-34 and women with a college education. In contrast, we find larger declines among Black women, who were hard hit by both the recession and pandemic.

Understanding why the baby bump was concentrated in some groups of U.S.-born mothers is complicated as the explanation is likely to be multi-faceted. The pandemic led to a historic rise in remote work, particularly for more educated workers, and 40% of days worked were still at home by spring 2021 (39). This may be one reason that the baby bump was stronger among college-educated women since working from home could reduce the opportunity costs of having a child. Consistent with this idea, there were much larger gains in fertility among childless women (i.e., mothers with first births) who, in addition to increased flexibility, did not have to manage the loss of day care and schooling opportunities for older children.

These observations highlight the potentially significant role that the organization of labor markets and childcare markets may be playing in driving the long-term U.S. fertility decline. For example, one study suggests that institutions such as stable public sector employment and maternity leave are important determinants of fertility (6). Another focuses on the burden of childcare for working mothers, showing that it is a major determinant of the "motherhood penalty" in women's wages, which in turn may also discourage childbearing (40). To the extent that workfrom-home trends begun during COVID-19 lockdowns helped mitigate these barriers to motherhood, childbearing may continue to be elevated for groups who can take advantage of them. Black women are the only race/ethnic group who showed a pandemic decline in fertility *without* a subsequent recovery. Black people suffered the largest economic losses (41) as well as a disproportionate burden of COVID morbidity and mortality (42, 43). By the end of 2022 in California, fertility rates for this group still had not rebounded. The trauma of the pandemic may continue to depress fertility rates for the worst affected groups. In short, different experiences with the pandemic may have led to very different fertility responses across groups, which are still playing out in 2023.

Turning to the timing of the downturn in births, one explanation for declining birth rates in early 2020 is that the COVID-19 recession was anticipated. Research shows that conceptions are typically a "leading indicator" of recessions (44). However, unlike other recessions in recent memory, the COVID-19 recession did not emerge from deteriorating economic circumstances (which could have been felt by individuals planning their families) but from an unexpected pandemic and ensuing lockdowns to contain its spread. In June 2019, the U.S. unemployment rate was very low at 3.6% and it remained at 3.5% until February 2020. The 2019 GDP growth rate of 2.3% was the same as in 2015 and higher than in 2016 or 2017 though it was lower than in 2018 when it was 2.9%. To look for signs of fears about the economy, we used GoogleTrends to search for discussions about recession in the summer and fall of 2019 (results shown in Appendix Figure A9). There was little discussion of impending recession except for a one-week spike in mid-August. In short, there is little evidence that the economy was sliding into recession prior to the pandemic or that the COVID-19 recession, brought on by rapid lockdowns, was expected.

For the shortfall in births beginning in early 2020 to reflect falling conceptions, behavioral changes would have had to begin by June 2019—well before COVID-19 had been identified and before there was any evidence or even talk of future recession. Moreover, these changes in

conceptions in 2019 would have had to occur among foreign-born women, but not among U.S.born women, to match the distinct patterns in these populations. All things considered, changes in conceptions starting in mid-2019 are unlikely to explain the patterns.

Turning to the results for foreign-born women, one explanation for the sharp reduction in births to foreign-born mothers in early 2020 relates to the Trump Administration's restrictions on travel from foreign countries to contain the spread of COVID-19. Restrictions on individuals traveling from China were enacted in January 2020 and then restrictions were extended to individuals traveling from Mexico, Canada, and most European countries in March 2020. Births to Chinese-born women began to decline in January 2020, and births to Mexican and Latin American women began to decline in March 2020—changes which correspond closely to the implementation of these travel restrictions.

The Trump administration also issued a controversial January 2020 ban on pregnant women entering the country on tourist visas.<sup>8</sup> This ban was explicitly aimed at preventing "birth tourism," a term referring to foreign nationals who enter the U.S. on tourist visas, give birth, and then return to their home countries, presumably so that their children can gain U.S. citizenship. Although it is extremely difficult to quantify, this phenomenon has received increasing attention in recent years, with much of that attention focused on Chinese nationals. Estimates of the number of pregnant Chinese-born women entering the U.S. to give birth before returning home range from 10,000 to 50,000 per year (45, 46). There have also been recent immigration crackdowns on agencies and organizations facilitating such travel (47).

One piece of evidence consistent with the idea that reductions in U.S. births to foreign nationals drove at least some of the sharp decline in births starting in January 2020 is that the

<sup>&</sup>lt;sup>8</sup> See <u>https://help.cbp.gov/s/article/Article1838</u> (accessed April 1, 2023) for more information about the policy change.

number of U.S. births that were self-paid (rather than paid by insurance) fell sharply starting in January 2020 (Appendix Figure A10). Anecdotal reports suggest that many women entering the country to give birth pay out of pocket.<sup>9</sup>

The COVID-19 recession and lockdowns may also have reduced the number of migrants coming to the U.S. to seek or hold jobs. Cherlin and coauthors argues that the substantial drop in fertility seen among Mexican-born women during the Great Recession can be explained by the recession's effect on the inflow of new Mexican migrants, since these new migrants historically have had higher fertility than Mexican-born women who have been in the U.S. for some time. He notes that, "the reasons for this sudden drop in net migration include the decline in employment opportunities during the Great Recession, along with heightened border enforcement and the difficulty and danger of border crossings...The consequences for fertility are straightforward" (p. 222) (48).

While most foreign-born mothers are U.S. residents, the declines in births to foreign-born mothers began in January 2020, coinciding with this policy change and with more general pandemic travel restrictions, and suggesting that such policy changes may have been important on the margin. The decline in births to foreign-born mothers may have also reflected some foreign-born pregnant women who are usually U.S. residents returning to their country of origin during lockdowns to be with family.

It is also important to note that with the important exception of Black U.S.-born women, the 2020 reductions in childbearing are concentrated among the foreign-born per se, rather than in a particular race or ethnic group. Appendix Figure A11 shows that if we break Hispanic/Latina women and Asian/Pacific Islander women into the U.S.-born and the foreign-born, U.S.-born

<sup>&</sup>lt;sup>9</sup> Some women apparently receive advice about how to claim Medicaid eligibility suggesting that the decline in the number of self-paid births may understate the fall in the number of foreign nationals giving birth in the U.S. (37).

mothers in both groups look quite similar to other U.S.-born women, whereas births to foreignborn Asian/Pacific Islander and Hispanic/Latina women fall starting in January 2020 and March 2020, respectively.

Births to foreign-born mothers recovered but did not exceed pre-pandemic trends in 2021. The V-shaped pattern in births to foreign-born women shown in Figure 2B may reflect two countervailing trends. If foreign-born mothers who are U.S. residents experienced the same surge in childbearing as U.S.-born mothers in 2021 (as Appendix Figure A11 suggests), then this could have offset declines in childbearing among foreign-born women who were also foreign nationals. U.S. birth certificates do not record citizenship making it impossible to investigate this hypothesis statistically in our data.<sup>10</sup>

These estimates highlight the important role that foreign-born women play in bolstering U.S. fertility rates. Without these births, the U.S. TFR would be 0.12 births lower and closer to European countries such as Norway or Germany. Moreover, the fact that the fall in births to foreign-born women occurred contemporaneously with disruptions to international travel and the recession suggests that the phenomenon of foreign nationals giving birth in the U.S. may be an important one that is worthy of further study.

Our results should be interpreted with some limitations in mind. One important limitation of a deviations-from-trends analysis is that the size of the deviation depends on how the trend is measured. There is no theoretical reason, for example, that trends must be linear over time or that pre-pandemic trends should continue indefinitely. In addition to considering the shape of the trend, one might also choose alternative periods over which to estimate pre-pandemic trends.

<sup>&</sup>lt;sup>10</sup> Birth certificates contain information on place of residence, but 99.73% of certificates report a U.S. residence suggesting that births to foreign nationals are not well captured.

To evaluate the sensitivity of our results to alternative choices, Appendix Figure A2 shows pre-pandemic trends estimated over different time windows. Whereas trends in birth rates appear linear in the short run between 2015 and 2019, a quadratic fits the data better when extending the period to 2011. However, our conclusion that foreign-born women drove the 2020 baby bust and that U.S.-born women drove the 2021 baby bump is robust to multiple alternative specifications. For example, a comparison of the alternative trend lines in Figure A2.E and A2.F shows that trends calculated over a longer period suggest a larger baby bump among U.S.-born mothers in 2021 compared to our baseline trend specification, while still indicating a large negative trend deviation for foreign-born mothers in 2020.

### **IV.** Materials and Methods

This analysis draws on natality data and population estimates to create seasonally adjusted estimates of childbearing between 2015 and 2021 for the U.S. and between 2015 and February 2023 for California. The following sections discuss data and methods. See Appendix Section C for additional information on data access and replication.

#### A. Natality Data

Natality data are from the National Center for Health Statistics (NCHS) for 2011-2021 and the California Department of Health (CDPH) for January 2015 to February 2023 (49, 50). Although California mothers are slightly older and more educated than U.S. mothers on average, the state is home to one eighth of the U.S. population and allows us to provide a rich description of multiple subgroups through February 2023. In contrast, national microdata are currently only available through December 2021.<sup>11</sup> Natality data contain detailed information about every birth in the U.S.

<sup>&</sup>lt;sup>11</sup> We have also searched state vital statistics offices in search of additional state data to inform this study. At the time of our writing, other large states tended to release data aggregate data on birth rates only through 2020 or 2021. For

and California, respectively, including the month and year of the birth and birth order (e.g., first birth, second birth). The data also include the nativity of the mother, mother's race and ethnicity, age, and education. The subsample of U.S.-born mothers includes those born in the 50 U.S. states, the District of Columbia, and U.S. territories. Although this analysis focuses on births that occurred in 2020 through 2021 (or February 2023 for California), births in 2015-2019 and alternatively 2011-2019 were also examined to estimate the pre-pandemic natality trends and assess deviations from them.

#### B. Construction of Fertility Rates

The analysis focuses on birth counts, total fertility rates, and birth rates to allow comparisons to other estimates in the literature and to aid in interpretation. The total fertility rate (TFR), which projects the number of children a woman would have over her lifetime if she experienced the age-specific birth rates in a particular period, is calculated for month m as follows:

$$TFR_m = 5 \times \sum_a \frac{12 \times births_{a,m}}{population_{a,m}},$$

where *births* captures the number of births to women in age group *a* in month *m*, and *population* is the number of women in age group *a* in month *m*. Age is grouped in five-year age intervals for women ages 15 to 49. Rates are multiplied by 12 to translate monthly birth rates into an annual equivalent to facilitate comparisons with more typically reported annual TFRs. Multiplying by five relates to the fact that birth rates are measured for five-year groups, and hence women are expected to experience an age-group specific birth rate for five years.

example, Florida has data through 2021, Illinois only posts birth counts through 2020, Pennsylvania has data available through 2020, New York has data through 2020, Texas has data through 2020, and Wisconsin only has data available through 2020. To the best of our knowledge, California is unique in making its microdata available for as recently as February 2023.

In addition, birth rates by month m are computed for specific demographic groups g as follows:

$$B_m^g = \frac{12,000 \times births_m^g}{population_m^g}$$

where  $births_m^g$  is the number of births to women in group g in month m, and population is for the same group in the same month. It is standard to report birth rates per 1,000 in the population, and multiplying by 12,000 translates monthly birth rates into an annual equivalent which facilitates comparisons with annual birth rates.

Groups *g* are defined as five-year age groups<sup>12</sup>, race groups<sup>13</sup> (White Non-Hispanic, Black Non-Hispanic, Hispanic/Latina, Asian/Pacific Islander Non-Hispanic), marital status, education groups (high school or less, some college, college attainment), and birth parity (first birth, second births, or higher order births). One complication relates to the rise of multi-race reporting, because the incidence of people reporting that they are multi-race has risen dramatically over the last decade. For transparency, our main results by race classify only individuals who report a single race group. Appendix Figure A12 reports supplemental figures for women reporting multiple races or without a race/ethnicity specified.

Information on mothers' education and their country of birth is missing from birth certificates for Connecticut in 2015 and for New Jersey in the third and fourth quarter of 2015. This missing information induced discontinuous changes in birth counts and fertility rates, which affected the estimation of pre-pandemic trends. To correct this problem, the missing information

 $<sup>^{12}</sup>$  Due to small sample sizes, we group 40-44- and 45–49-year-old women into a 40+ category for the age-group analysis.

<sup>&</sup>lt;sup>13</sup> In the Vital Statistics Natality data, the six race classifications are American Indian or Alaska Native; Asian; Black or African American; more than one race; Native Hawaiian or Other Pacific Islander; White; Unknown or Not Stated. In the ACS, individuals can choose White, Black or African American, American Indian or Alaska Native, Asian, Native Hawaiian or Other Pacific Islander, or Other Race, and can check more than one box. In our primary analysis of natality and ACS data, individuals reporting multiple races are omitted from the estimates by race.

is extrapolated assuming that these characteristics followed the trend observed in 2016-2019—an assumption that was verified in states without missing data. In particular, the share of births in each subcategory is first calculated at the state-year level. Then the predicted shares are calculated for each reported group in each subcategory by estimating a regression for each group-subcategory and state with the share as the dependent variable and the date as a linear predictor, using non-missing information across all years prior to 2020. Birth counts in missing categories and years are then replaced for the states of interest by calculating the total monthly birth counts and multiplying that number by the predicted shares for each month in the missing year.

#### C. Population Estimates

Fertility rates are calculated to aid in interpretation of the results and to facilitate comparisons with other population estimates. Calculations of fertility rates require population estimates, which are difficult to estimate during periods like the pandemic, when individuals may change their country of residence. For instance, the exact number of foreign-born women in the U.S. in any month in 2020 is unknowable, because migration flows experienced changes due to the pandemic. In addition, well-documented problems with the 2020 Census complicate the direct calculation of populations in 2020 and 2021 (51). Consequently, indirect methods were used to compute populations by group, which helps avoid some issues with measurement error. Moreover, all results are also presented using birth counts, so that the results are not driven by the construction of the population estimates.

Overall population counts (and by age group, nativity, marital status, race and ethnicity group, and education group) were calculated by mother's individual birth-year cohort based on the mean population in the 5-year American Community Surveys (ACS) for 2015-2019 (52). To avoid problems with the 2020 Census and 2021 ACS enumeration, the cohort population was assumed

to remain the same in 2020 and 2021 (for California through February 2023). This number will overstate the population denominators in these years slightly due to deaths and emigration, but these outcomes are rare for women of childbearing age. Denominators may be less reliable for foreign-born women. Annual population counts are assigned to December of the corresponding calendar year and smoothed across months using cubic interpolations to avoid inducing discontinuities or kinks in population denominators. This procedure results in population estimates by month.

For U.S.-born mothers, cohort counts are summed up to correspond to the groups *g* defined previously. For race and parity birth rates, women 15-49 are used to construct the population denominators.<sup>14</sup> Education birth rates are based on a narrower sample of women aged 25-49, because they are more likely than younger women to have completed their education. Younger women may have had their education disrupted by the pandemic. To account for the fact that education increases with age, education shares in 2015-2019 are regressed on age-specific linear trends to predict each cohort's educational distribution in 2020 and 2021. Population estimates by subgroup are plotted in Appendix Figure A13.

#### D. Seasonal Adjustment, Pre-Pandemic Trends and Deviations from Trends

Childbearing is highly seasonal, and seasonality differs across population subgroups (53). To adjust for seasonality, monthly birth counts or rates from 2015-2019—overall and for each specific subgroup—are regressed on calendar month fixed effects. The regression is estimated using the pre-pandemic period. Residuals are then computed for the entire period, either for January 2015 to December 2021 in the U.S. or January 2015 to February 2023 in California. These

<sup>&</sup>lt;sup>14</sup> For comparability with the natality data, our denominators for race/ethnicity groups only count individuals in single race categories.

residuals are then added to the mean of the outcome for 2015-2019, so that the level (either the count or rate) is easily interpretable. We refer to these constructs as "seasonality adjusted" birth or fertility rates, and this is what is used in the analysis.

Trends in seasonally adjusted births or birth rates are computed separately for each group, *g*, by regressing these outcomes on a linear trend in months using data for the pre-pandemic period, January 2015 to December 2019, which is plotted as a solid line in the figures.<sup>15</sup> The figures present the projection of this trend as a dashed line in January 2020 to December 2021 for the U.S. overall and through February 2023 for California. In addition, 95-percent confidence intervals for the linear prediction are displayed as a shaded region on the figures. Total deviations from prepandemic trends are computed by taking the difference in monthly births or fertility rates from the trend at a particular point in time or by summing over all months in a period. Percent deviations are computed by dividing the sum of total deviations by the average trend level in the same period.

### E. Using Unemployment Rates to Predict Changes in Fertility Rates

We use Currie and Schwandt's 2014 methodology to calculate how the TFR for the U.S. would evolve if conceptions had responded similarly to the changes in the national unemployment rate as they did in the previous four recessions (8). We use their estimates of the association between live births (lagged by 9 months) and changes in the unemployment rate over the previous 12 months. Note that Figure 1B plots these predictions in calendar time even though the association is estimated using lagged birth rates. See the Methods Appendix for more information on this methodology.

<sup>&</sup>lt;sup>15</sup> Kearney and Levine (2022) seasonally adjust monthly birth data using month fixed effects and estimate prepandemic trends using all U.S. births from October 2016 to September 2020 (16). The use of all 2020 data in the pretrend estimation means their pre-pandemic trend will be more sharply downward sloped due to changes occurring in early 2020 in response to the pandemic as we show in Appendix Figure A2.

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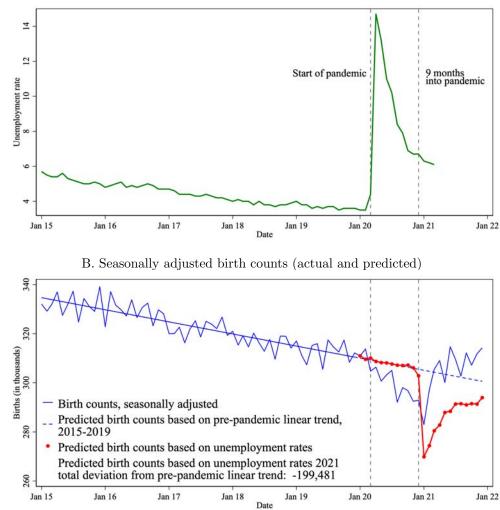
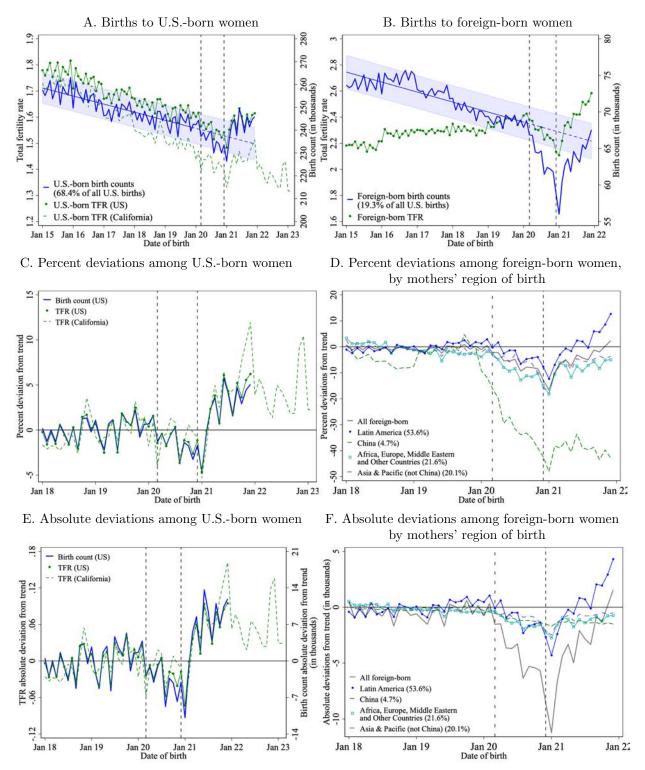


Figure 1: Changes in the U.S. Unemployment Rate and Predicted Changes in Births

*Notes*: Panel A plots changes in the U.S. unemployment rate from March 2015 to March 2021. Panel B plots three series: the seasonally adjusted U.S. birth counts from January 2015 to December 2021; a linear trend fit on prepandemic seasonally adjusted birth counts from 2015 to 2019 (solid line) and projected to 2020 to 2021 (dashed line); and the predicted U.S. fertility rate based on changes in unemployment rates using estimates from Currie and Schwandt (2014). Using information from the previous four recessions, Currie and Schwandt (2014) estimate that a 1 percentage point increase in the unemployment rate is associated with changes in live births for different age groups as detailed in the text. See Materials and Methods for more details about seasonal adjustments and our Methods Appendix for more details regarding the Currie and Schwandt's (2014) prediction.

Figure 2: Seasonally Adjusted Births and Deviations from Pre-Pandemic Trends, by Mothers' Nativity



*Notes*: Calculations use all births occurring in the U.S. and stratify by whether the mother was born in or outside the U.S. The legends in panels A and B indicate in parentheses the percent of all births in the U.S. in 2019 that occurred to U.S.-born women (panel A) or foreign-born women (panel B). The legends in panels D and F indicate in parentheses the percent of all births in the U.S. to foreign-born women in 2019 occurring to women in the indicated region of birth grouping.

	$\frac{2019}{\text{All births}}$ or TFR	2020			2021			2020-2021
		All births or TFR	Total dev.	% dev.	All births or TFR	Total dev.	% dev.	Net births relative to trend
A. Births in US to women	ages 15-49							
Number	3,754,688	$3,\!617,\!034$	-76,304	-2.03	3,666,993	32,777	1.00	-43,528
Total fertility rate (TFR)	1.74	1.68	,	-1.86	1.70	,	1.38	,
B. Births in US to women a	ages 15-49, b	y nativity of	mother					
Number, U.Sborn	2,919,763	2,839,025	-31,104	-1.06	$2,\!898,\!521$	70,865	2.58	39,760
TFR, U.Sborn	1.62	1.56		-0.95	1.58		3.00	
Number, foreign-born	826,991	770,561	$-45,\!338$	-5.41	760,240	-39,257	-4.67	-84,595
Number, nativity not specified	7,934	7,448	138	2.16	8,232	1,169	18.88	1,307
C. Number of births in US	to foreign-bo	orn women, b	by mothers	' region o	of birth			
Asia & Pacific	166,181	157,383	-9,314	-5.43	153,412	-11,213	-6.65	-20,527
(not China)								
China	42,569	31,033	-11,674	-25.87	25,496	-16,917	-39.68	-28,592
Latin America	439,270	415,680	-9,327	-2.09	416,913	4,154	1.43	-5,173
Africa, Europe, Middle	178,971	166,465	-15,022	-8.10	164,419	-15,281	-8.37	-30,303
Eastern and Other								
Countries								

Table 1: U.S. Birth Counts, Birth Rates, and Deviations from Pre-Pandemic Trends, 2019-2021

*Notes:* Columns labeled "All births or TFR" present the number of births in the U.S. or the total fertility rate (TFR) for the indicated calendar year. "Total deviation" is computed by summing over the difference between the seasonally adjusted, monthly births or fertility rates and the pre-pandemic trend for the indicated year. "% deviation" is calculated by dividing the total deviation by the average level of the trend in the same year. "Net births relative to trend" is calculated by adding the total deviations in 2020 and 2021 together. "Number" indicates the total count of births in a calendar year. "Total fertility rate (TFR)" is constructed as described in the paper. Restricted data for all births in the U.S. for January 2015 to December 2021 come from natality data from the National Center for Health Statistics (NCHS). \* dev. stands for "deviation".

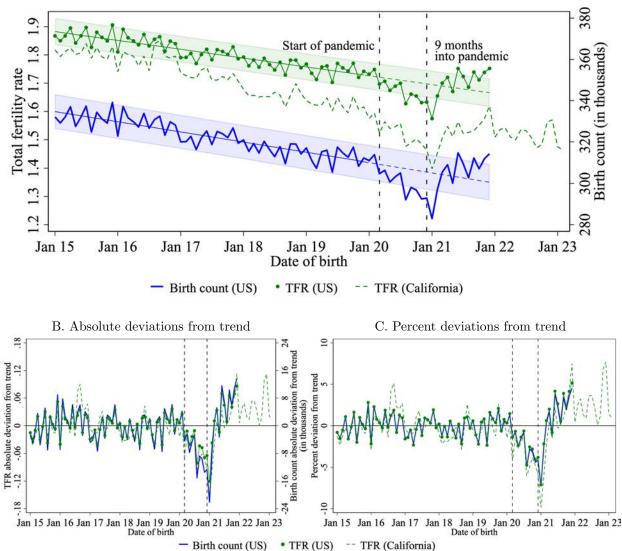
Table 2: U.S. Birth Counts, Birth Rates, and Deviations from Pre-Pandemic Trends among US-Born Women, 2019-2021 by Subgroup

	2019	2020			2021			2020-2021
	All births or TFR	All births or TFR	Total dev.	% dev.	All births or TFR	Total dev.	% dev.	Net births relative to trend
A. Number of births in US	5 to U.Sbor	n women age	es 15-49					
Number Total fertility rate (TFR)	2,919,763 1.62	2,839,025 1.56	-31,104	-1.06 -0.95	2,898,521 1.58	70,865	$2.58 \\ 3.00$	39,760
B. Number of births in US	to U.Sbori	n women, by	age group	)				
Age 15-19	150,807	139,456	3,682	2.79	128,298	5,082	4.50	8,764
Age 20-24	$602,\!605$	572,030	5,124	0.93	558,700	22,541	4.34	27,666
Age 25-29	871,810	835,379	-29,370	-3.35	840,517	-14,588	-1.66	-43,958
Age 30-34	$818,\!576$	817,066	937	0.14	864,276	48,048	5.95	48,986
Age 35-39	$397,\!945$	395,514	-12,515	-2.96	421,357	4,348	1.12	-8,167
Age 40-44	$73,\!552$	$75,\!195$	1,028	1.55	80,712	5,065	7.02	6,093
Age 45-49	4,468	4,385	9	1.55	4,661	368	10.21	377
C. Number of births in US	to U.Sbori	n women, by	parity					
First birth	1,128,390	1,113,401	12,862	1.18	1,133,890	$55,\!664$	5.26	68,526
Second birth	933,970	900,826	-13,718	-1.47	926,074	25,932	2.96	12,214
Third or higher	850,677	818,531	-30,757	-3.57	832,138	-12,243	-1.40	-43,000
Not specified	6,726	6,267	509	13.61	6,420	1,512	35.42	2,022
D. Number of births in US	5 to U.Sbor	n women, by	race/ethn	licity				
White Non-Hispanic	1,785,360	1,721,009	-20,449	-1.15	1,765,038	$65,\!357$	3.92	44,908
Hispanic/Latina	495,350	494,367	-311	-0.04	513,161	16,566	3.47	16,255
Black Non-Hispanic	456,768	443,650	-7,686	-1.67	432,160	-15,130	-3.32	-22,816
Asian/Pacific Islander	$52,\!517$	52,471	-476	-0.83	56,515	2,646	5.10	2,170
Other or not specified	129,768	$127,\!528$	-2,181	-1.65	$131,\!647$	1,426	1.19	-756
E. Number of births in US	to U.Sbori	n women, by	marital st	atus				
Married	1,508,948	1,459,346	-25,972	-1.71	1,520,111	58,288	4.07	32,316
Unmarried	1,121,216	1,098,742	-2,272	-0.19	1,089,939	2,897	0.34	625
Not specified	289,599	280,937	-2,861	-0.99	288,471	9,680	3.61	6,819
F. Number of births in US	to U.Sborn	n women age	s 25-49, by	v educat	ion			
Number, ages 25-49	2,166,351	2,127,539	-39,911	-1.80	2,211,523	43,242	2.05	3,331
< 4 years of college	1,217,391	1,185,819	-34,400	-2.77	$1,\!196,\!528$	-23,077	-1.84	-57,477
$\geq 4$ years of college	$924,\!956$	$918,\!175$	-5,267	-0.53	987,695	$63,\!435$	6.94	58,168
Not specified	24,004	23,545	-242	-0.86	27,300	2,886	12.64	2,644

*Notes:* See Table 1 notes. Restricted data for all births in the U.S. to U.S.-born women ages 15 to 49 for January 2015 to December 2021 come from natality data from the National Center for Health Statistics (NCHS). "Other or not specified women" for race/ethnicity groups the following women together: women identifying as American Indian/Alaskan Native, women identifying as more than one race, and women who do not indicate their race or ethnic group. Some states are missing information at points in time. Country of birth and education is missing for Connecticut in 2015 and in New Jersey for Q3 2014-Q4 2015. We project outcomes for these states as defined in our Materials and Methods. Marital status is missing for California for 2017-2021, so California is in the "not specified" category in all years. \* dev. stands for "deviation".

#### A. Appendix Figures

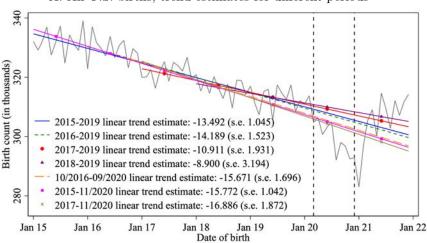


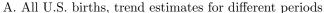


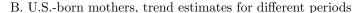
A. Number of births in the U.S. and total fertility rate

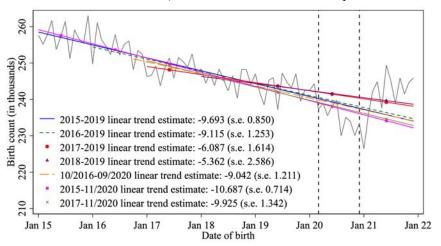
*Notes*: Calculations use all births occurring in the U.S. Births and fertility rates have been seasonally adjusted as indicated in the paper. Panel A presents the annualized total fertility rate (TFR) constructed as described in the paper (left vertical axis) alongside birth counts (right vertical axis). Linear pre-pandemic trends are fit for each series using January 2015 to December 2019 (presented in a solid line, projection in dashed line). The shaded region is the 95-percent confidence interval constructed using the standard error of the linear forecast. Panel B presents the percent deviations from trend for the same series. In both panels, the first dashed line from the left indicates the start of the pandemic in March 2020 and the second dashed line indicates nine months after the pandemic began in January 2021. Restricted data for all births in the U.S. for January 2015 to December 2021 come from natality data from the National Center for Health Statistics (NCHS). Restricted data for all births in California for January 2015 to February 2023 come from the California Department of Health.

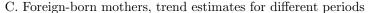
#### Appendix Figure A2: Birth Count Trend Sensitivity

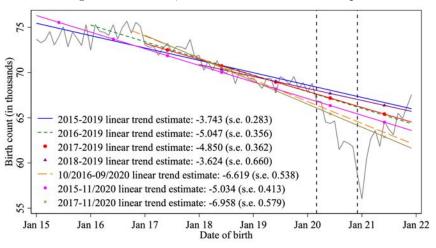




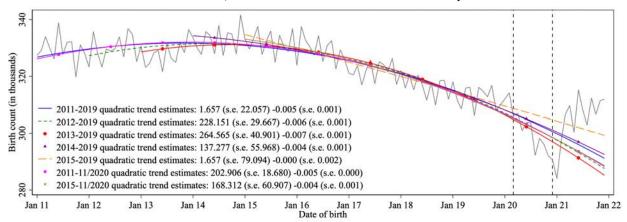




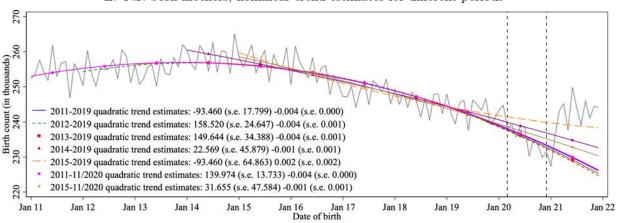




#### Appendix Figure A2: Birth Count Trend Sensitivity (continued)

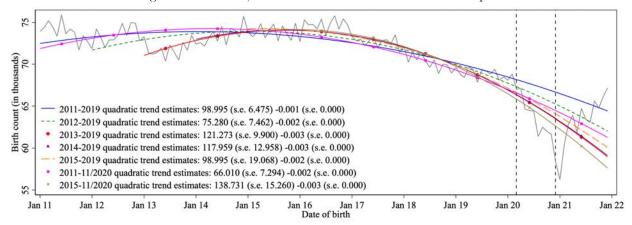






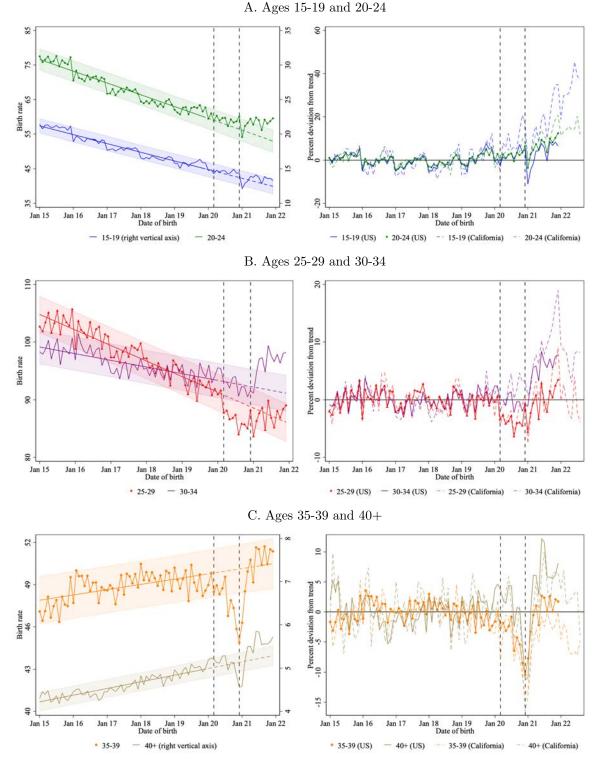
E. U.S.-born mothers, nonlinear trend estimates for different periods

F. Foreign-born mothers, nonlinear trend estimates for different periods



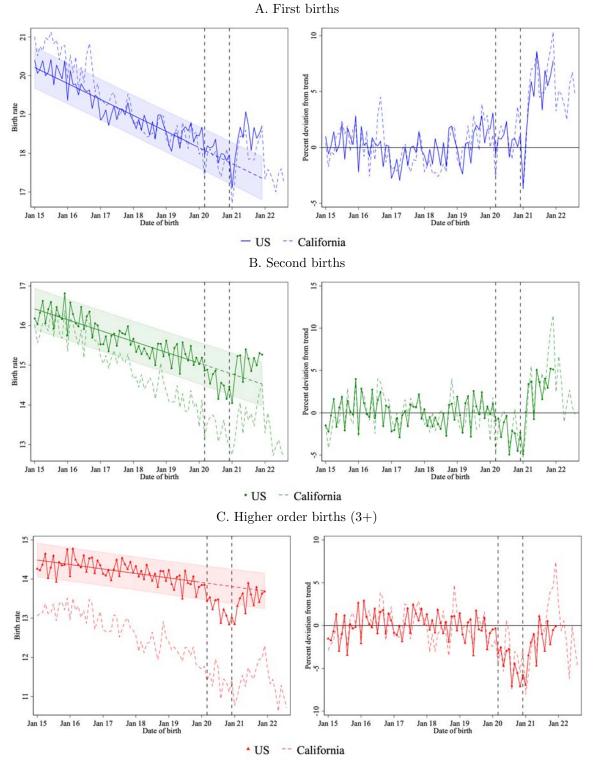
*Notes:* Panels A-C show the sensitivity of our findings to shortening the period for the linear trend from 2015-2019, to 2016-2019, 2017-2019 and 2018-2019 and also using births in 2020 to calculate the trend. In addition, we present alternative periods appearing in other papers, including Kearney and Levine (1) who use a linear trend based on data for births occurring from October 2016 to September 2020. The legend includes the estimate of the slope of the linear trend after the colon with the standard error of the slope estimate in parentheses. Panels D-F show the sensitivity of our findings to using a longer period of natality data as indicated in the legend and a quadratic trend estimate. The first term after the colon is the coefficient on the linear component and the second term is the coefficient on the squared component. Both coefficient estimates are followed by their standard errors in parentheses.

Appendix Figure A3: Seasonally Adjusted Birth Rates and Percent Deviation from Trend among U.S.-Born Women, by Five-Year Age Group



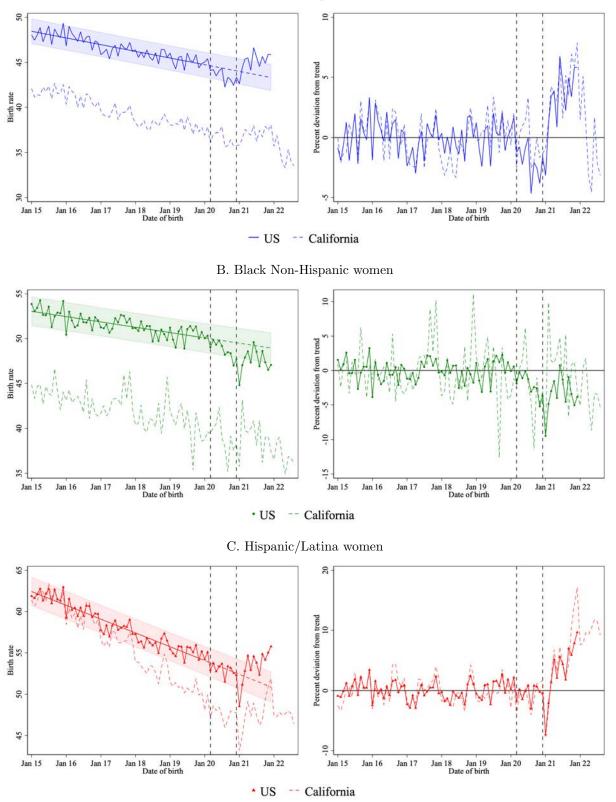
*Notes*: The figure presents seasonally adjusted, monthly birth rates for five-year age groups among U.S.-born women and linear pre-pandemic trends in birth rates (solid line, projection in dashed line) on the left side. On the right side, figures show percent deviations from pre-pandemic trends. Panels A and C present birth rates very different in levels, which motivates the use of separate vertical axes. Population denominators use U.S.-born women in the same five-year age groups. See also Figure 2 notes.





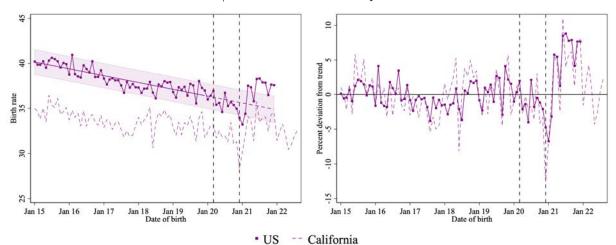
*Notes*: The figure presents seasonally adjusted, monthly birth rates among U.S.-born women by parity and linear pre-pandemic trends in birth rates (solid line, projection in dashed line) on the left side. On the right side, figures show percent deviations from pre-pandemic trends. Population denominators use U.S.-born women ages 15 to 49. Appendix Figure A12 presents estimates for women with no information on parity on the birth certificate. See also Figure 2 notes.

Appendix Figure A5: Seasonally Adjusted Birth Rates and Percent Deviation from Trend among U.S.-Born Women, by Race/Ethnicity



A. White Non-Hispanic women

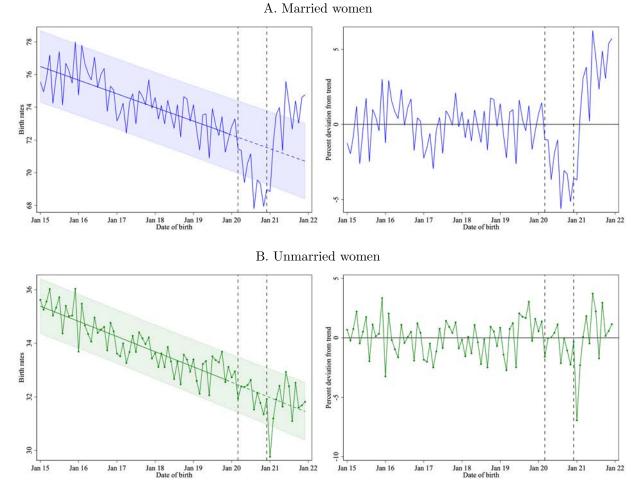
Appendix Figure A5: Seasonally Adjusted Birth Counts and Percent Deviation from Trend among U.S.-Born Women, by Race/Ethnicity (continued)



G. Asian/Pacific Islander Non-Hispanic women

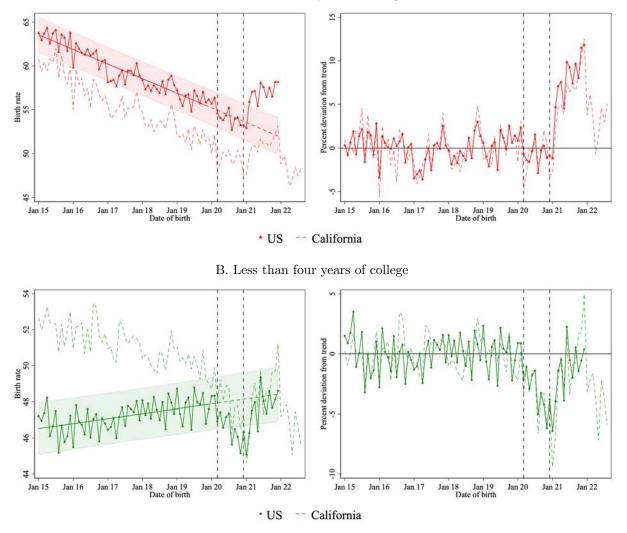
*Notes*: The figure presents seasonally adjusted, monthly birth rates among U.S.-born women by race/ethnicity and linear pre-pandemic trends in birth rates (solid line, projection in dashed line) on the left side. On the right side, figures show percent deviations from pre-pandemic trends. Panels A, B, and D are restricted to women who do not report Hispanic/Latina origin and only report a single race/ethnicity category, whereas Panel C includes all women who report Hispanic/Latina origin. Population denominators use U.S.-born women ages 15 to 49 in the same subgroups. Appendix Figure A12 presents estimates for women reporting multiple races or other race/ethnicity categories. See also Figure 2 notes.

Appendix Figure A6: Seasonally Adjusted Birth Rates and Percent Deviation from Trend among U.S.-Born Women, by Marital Status



*Notes:* The figure presents seasonally adjusted, monthly birth rates among U.S.-born women by marital status and linear pre-pandemic trends in birth rates (solid line, projection in dashed line) on the left side. On the right side, figures show percent deviations from pre-pandemic trends. Data exclude California, because California does not report marital status in the natality microdata. Appendix Figure A12E and A12F present estimates for women with no information on marital status on the birth certificate. See also Figure 2 notes.

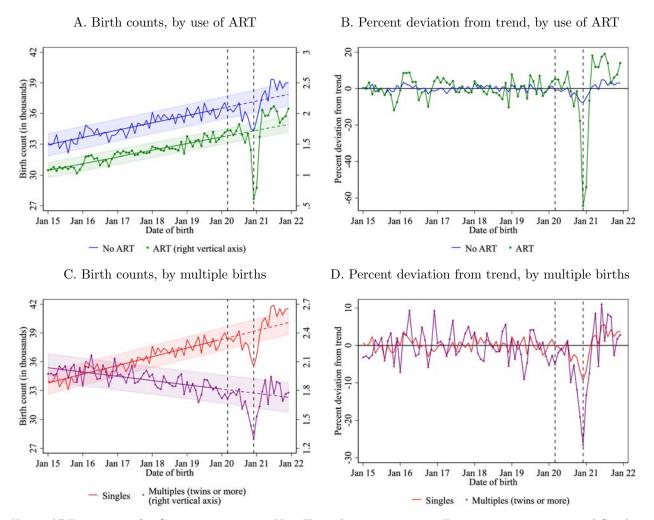
Appendix Figure A7: Seasonally Adjusted Birth Rates and Percent Deviation from Trend among U.S.-Born Women, by Education



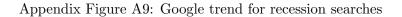
A. Four or more years of college

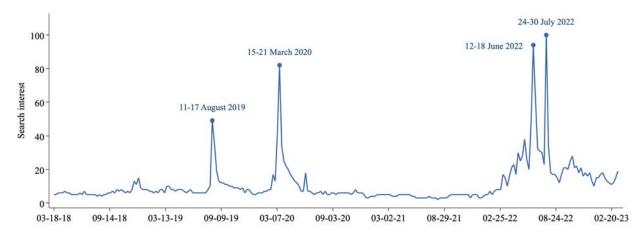
*Notes*: The figure presents seasonally adjusted, monthly birth rates among U.S.-born women by education and linear pre-pandemic trends in birth rates (solid line, projection in dashed line) on the left side. On the right side, figures show percent deviations from pre-pandemic trends. Population denominators use U.S.-born women ages 25 to 49 in the same educational groups. Appendix Figure A12 presents estimates for women with no information on education on the birth certificate. See also Figure 2 notes.

Appendix Figure A8: Seasonally Adjusted Birth Counts and Percent Deviations by use of Assisted Resproductive Technologies (ART) and Multiple Births among U.S.-born Women ages 35-49, 2015-2021



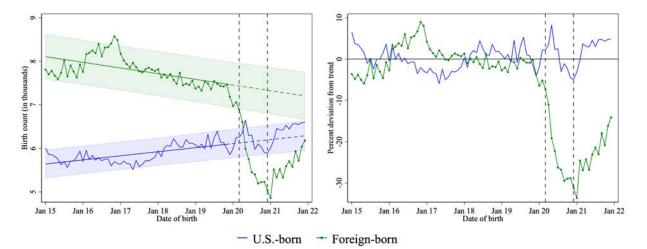
*Notes:* ART is missing for Connecticut in 2015, New Hampshire in 2015-2016, Tennessee in 2015-2018, and South Carolina in 2015-2020, so we exclude these states.





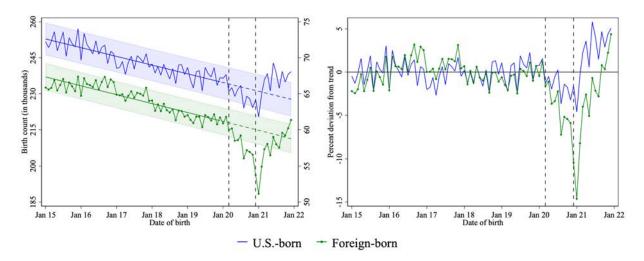
*Notes:* Numbers represent search interest relative to the highest point on the chart for the given region and time. A value of 100 is the peak popularity for the term. A value of 50 means that the term is half as popular. A score of 0 means there was not enough data for this term.

Appendix Figure A10: Seasonally Adjusted Birth Counts and Percent Deviations by Payment Source and Nativity, 2015-2021



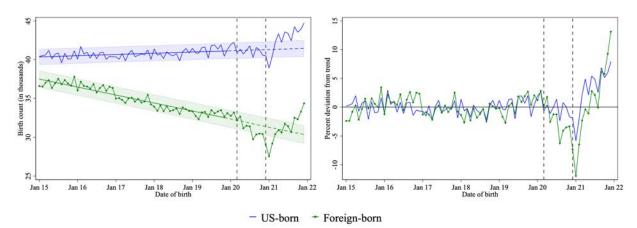
A. Self-pay, by mother's nativity

B. Other payment source, by mother's nativity



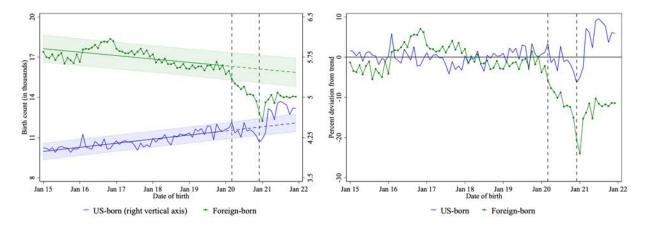
*Notes:* Other payment source category includes Medicaid, private insurance, unknown payment source, Indian Health Services, CHAMPUS/TRICARE, other government (federal, state, local) insurance and other payment sources. In Panel B, foreign-born birth counts are on the right vertical axis.

Appendix Figure A11: Seasonally Adjusted Birth Counts and Percent Deviations by Race/Ethnicity and Nativity, 2015-2021



A. Births to Hispanic/Latina women, by mother's nativity

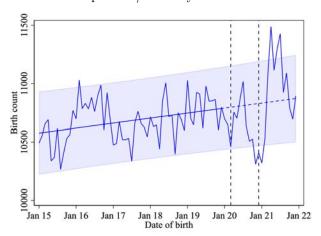
B. Births to Asian Pacific Islander Non-Hispanic women, by mother's nativity



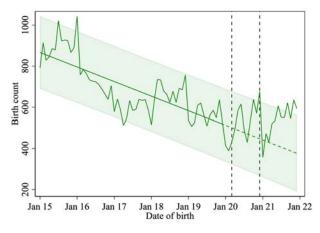
*Notes:* The figure presents seasonally adjusted, monthly birth rates by race/ethnicity and mother's nativity and linear pre-pandemic trends in birth counts (solid line, projection in dashed line) on the left side. On the right side, figures show percent deviations from pre-pandemic trends. See also Figure 2 notes.

Appendix Figure A12: Seasonally Adjusted Birth Counts and Percent Deviation from Trend among U.S.-Born Women for Not-Reported or Other Categories

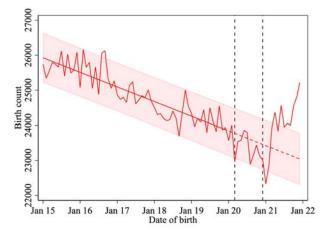
A. Seasonally adjusted birth counts among U.S.born women, race not specified or multiple race/ethnicity indicated



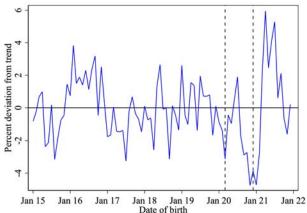
C. Seasonally adjusted birth counts among U.S.born women, parity not specified



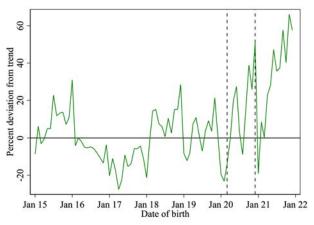
E. Seasonally adjusted birth counts among U.S.born women, marital status not specified



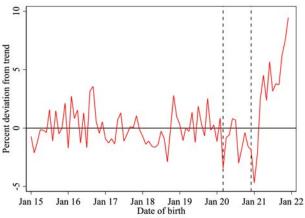
B. Seasonally adjusted percent deviation from trend among U.S.-born women, race not specified or multiple race/ethnicity indicated



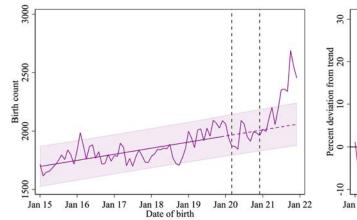
D. Seasonally adjusted percent deviation from trend among U.S.-born women, parity not specified



F. Seasonally adjusted percent deviation from trend among U.S.-born women, marital status not specified

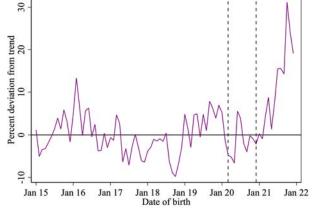


Appendix Figure A12: Seasonally Adjusted Birth Counts and Percent Deviation from Trend among U.S.-Born Women for Not-Reported or Other Categories (continued)

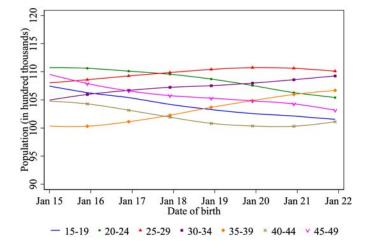


G. Seasonally adjusted birth counts among U.S.born women, education not specified

H. Seasonally adjusted percent deviation from trend among U.S.-born women, education not specified

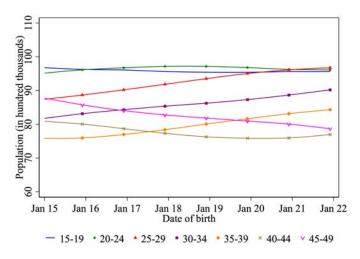


Appendix Figure A13: Population Estimates, by Month and Population Subgroup

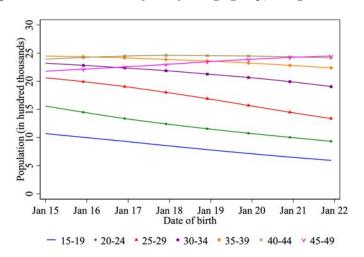


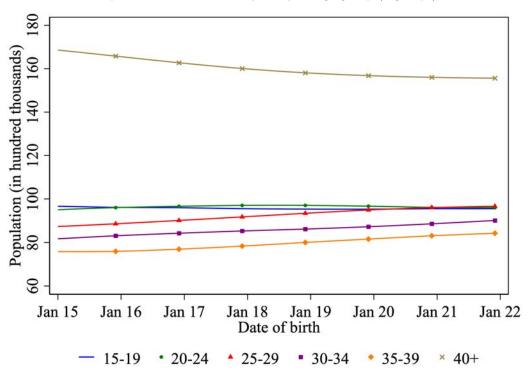
A. Population denominators by five-year age group, overall U.S.

B. Population denominators by five-year age group, U.S.-born women



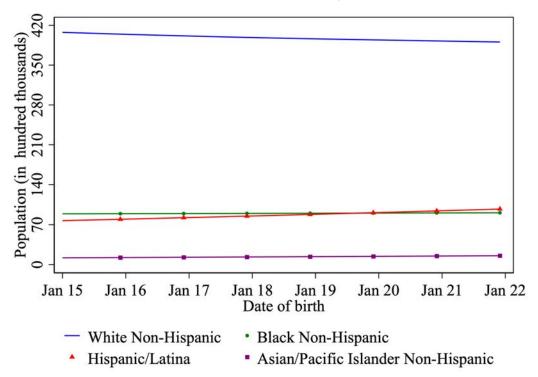
C. Population denominators by five-year age group, foreign-born women

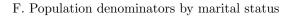


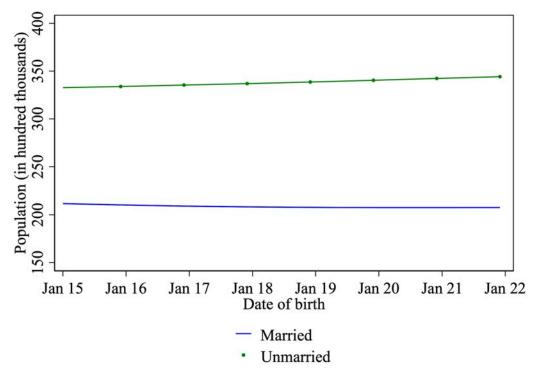


D. Population denominators by five-year age group (6 groups)

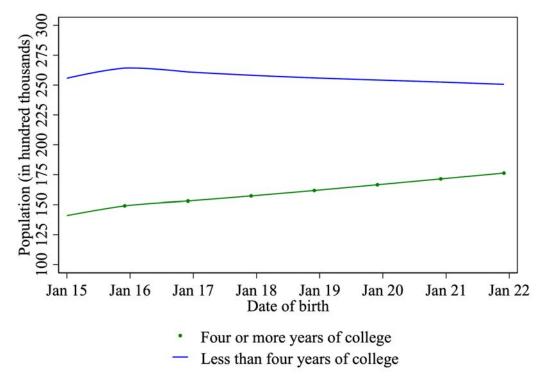
E. Population denominators by race/ethnicity







G. Population denominators by education



## B. Methods Appendix

## Using Unemployment Rates to Predict Changes in Fertility Rates

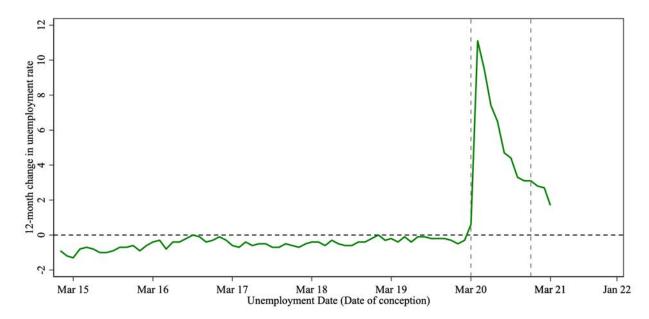
In Figure 1B, we use Currie and Schwandt's (2) methodology to calculate how the TFR for the U.S. would have evolved if conceptions had responded to the national unemployment rate in the same way as in the previous four recessions. Currie and Schwandt (2) find that conception rates (restricted to conceptions resulting in live births) are negatively associated with changes in the unemployment rate. Note that Figure 1B plots these predictions in calendar time even though the association is estimated using the approximate time of conception 9 months earlier. Because of the 9-month lag and the fact that the unemployment rate prior to March 2020 exhibited normal variation, the model does not predict a drop in births below trend for most of 2020.

As reported in Figure 2 of Currie and Schwandt (2), a 1 percentage point increase in the unemployment rate is associated with the following changes in the number of live births per 1,000 women nine months later: -0.59 for women ages 15 to 19, -1.27 for women ages 20 to 24, -0.9 for women ages 25 to 29, -0.48 for women ages 30 to 34, -0.24 for women ages 35 to 39, and by 0.002 for women ages 40 to 44. Currie and Schwandt (2) show in Appendix Figure S2 that the association of changes in the unemployment rate and subsequent fertility rates is very similar across all four recession periods covered by their analysis, although the association was somewhat stronger during the Great Recession. We then use these estimates to calculate a counterfactual fertility response as follows.

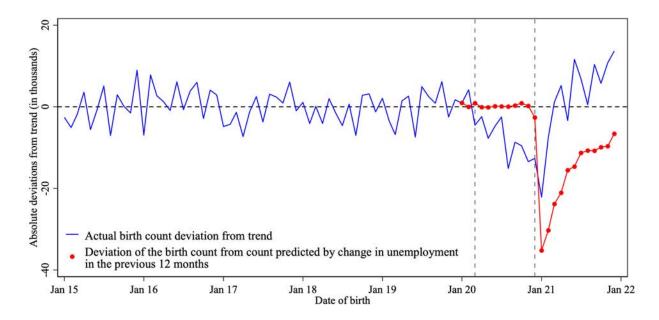
We calculated the 12-month change in the national unemployment rate for each month between January 2000 and March 2021. Then we indexed seasonally adjusted monthly U.S. birth counts from 2015 to 2021 for each age group to be nine months before they occurred (which corresponds to the approximate timing of conceptions). Next, separately for each age group, we estimated a linear trend for these birth rates between January 2015 and December 2019 and projected this trend through December 2021. For each month between 3/2020 and 3/2021 and each age group, we then multiplied the 12-month change in the national unemployment rate with the age-group specific coefficients reported in Figure 2 of Currie and Schwandt (2). This predicted change in conceptions was then added to the extrapolated trend nine months later.

Two figures (see next page) provide more description of the calculations underlying Figure 1. The first plots 12-month changes in unemployment from 2015-2021. The second plots deviations in actual birth counts from the trend as predicted by the 12-month change in unemployment (red line with o-markers). This figure makes it clear that changes in unemployment rate (or their absence) predict changes in fertility rates fairly well between 2016 and March 2020, but this relationship (which held in prior recessions) breaks down during the pandemic. The pattern of actual deviations differed from what would have been predicted on the basis of previous recessions. Moreover, it continued to be quite different even after the initial large spike in unemployment (caused by temporary layoffs) was passed. Our paper provides two reasons for this deviation from projection. First, actual births decline before the unemployment rate should have affected birth rates due in large part to the sharp reduction in births to foreign-born women. Second, births to U.S-born women surge much more than predicted by changes in unemployment rates.

Appendix Figure B1: Twelve-Month Changes in the U.S. Unemployment Rate, 2015-2021



Appendix Figure B2: Deviations in Actual Birth Counts from Trend and Deviations Predicted Using Twelve-Month Changes in Unemployment, 2015-2022



## C. Data Access and Replication

Readers wishing to replicate our analysis can access individual-level national vital statistics natality data at https://www.cdc.gov/nchs/data\_access/VitalStatsOnline.htm. In addition, users wishing to obtain restricted data containing geographical information can be requested at https://www.cdc.gov/nchs/nvss/nvss-restricted-data.htm. The individual-level California vital statistics natality data is available with an approved application only. Data access to the California vital statistics can be requested following the process outlined here: https://www.cdph.ca.gov/Programs/CHSI/Pages/Data-Applications.aspx.

Replication programs can be downloaded from OpenICPSR at https://doi.org/10.3886/E192846.

## Appendix References

- 1. M. S. Kearney, P. Levine, The US Covid-19 Baby Bust and Rebound. *NBER Working Paper 30000* (2022).
- 2. J. Currie, H. Schwandt, Short-and long-term effects of unemployment on fertility. *Proceedings* of the National Academy of Sciences **111**, 14724-14739 (2014).