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The Great Unequalizer: Initial Health Effects of COVID-19 in the United States
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ABSTRACT

We measure inequities from the COVID-19 pandemic on mortality and hospitalizations in the United States during the early months of the outbreak. We discuss challenges in measuring health outcomes and health inequality, some of which are specific to COVID-19 and others that complicate attribution during most large health shocks. As in past epidemics, pre-existing biological and social vulnerabilities profoundly influenced the distribution of disease. In addition to the elderly, Hispanic, Black and Native American communities were disproportionately affected by the virus, particularly when assessed using the years of potential life lost metric. For example, Hispanic and Black Americans in 2020 saw 39.5 and 25 percent increases in excess mortality relative to trend, compared to a less than 15 percent increase for Whites; we find losses in potential years of life three to four times larger among Hispanic and Black compared to White Americans. Individual-level data from a commercially insured population show that otherwise similar Black and Hispanic enrollees were hospitalized due to COVID-19 at a higher rate than White enrollees. We provide a conceptual framework and initial empirical analysis which seek to shed light on contributors to pandemic-related health inequality, and suggest areas for future research.

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Introduction

When SARS-CoV-2 first appeared it was discussed as an equal opportunity pathogen: no one was immune, and therefore all potentially affected (Mein 2020, Krishnan et al. 2020). Early cases of COVID-19 disease among elites in entertainment, politics, and industry such as actor Tom Hanks, U.K. Prime Minister Boris Johnson and Morgan Stanley chief executive officer James Gorman gave credence to this view.

Yet historical episodes of infectious disease are generally not experienced evenly across social strata. Rudolf Virchow, the founder of modern cellular pathology and a proponent of medicine as a social science, noted that "statistics will be our standard of measurement: we will weigh life for life and see where the dead lie thicker, among the workers or the privileged" (Virchow, 1848). Figure 1 demonstrates that for prominent novel infectious disease threats across the last two centuries in the United States, the dead indeed "lie thicker" among those less privileged. The upper-left panel shows that when cholera struck Boston in 1849, the mortality gap between native-born Bostonians and Irish immigrants and their children was about 50 percent larger than during a non-epidemic year like 1854. The upper-right panel illustrates how the age-adjusted mortality gap between White and non-White Americans rose during the influenza pandemic of 1918. The bottom-left panel displays how the HIV/AIDS pandemic worsened disparities in mortality between Black and White Americans. Finally, the lower-right panel shows that the age-adjusted difference in all-cause mortality rates between non-Hispanic White Americans and non-Hispanic Black Americans widened by over 80 percent during the first year of the COVID-19 pandemic.¹

This paper discusses the initial health effects of COVID-19 in the United States. During 2020, the first year of the pandemic, COVID-19 was recorded as the underlying or contributing cause of 378,000 deaths nationwide (Ahmad et al., 2021). The impact of COVID-19 on health, however, extends beyond its direct toll on mortality. We begin by discussing the various ways by which COVID-19's health effects have been measured, as well as the role pandemic-induced changes in the demand and supply sides of health care may have played in affecting mortality from causes other than the virus itself. Comparing the overall increase in mortality from 2019-20 to the number of recorded COVID-19 deaths in 2020 indicates that the pandemic has had substantial indirect effects on health.

We next turn to examining inequality in the effect of COVID-19 on the health of different population groups. Infectious pathogens exploit both biological and social vulnerabilities, and the presentation of medical statistics can render gradients more or less conspicuous. Older age groups are particularly vulnerable to COVID-19, raising health risks for population groups with higher average ages, such as non-Hispanic Whites. Mortality rates conditional on age, however, are considerably higher for historically disadvantaged groups such as Black, Hispanic, and American Indians – especially due to their young average ages, these groups account for a disproportionate share of COVID-19 deaths.

¹See Appendix Figure 1 for a comparison of 2019 and 2020 mortality rates including other races and ethnicities.
Panel A: Cholera (1849)

Baseline Mortality

Pandemic Mortality

Δ = 15.97

Panel B: Influenza (1918)

Baseline Mortality

Pandemic Mortality

Δ = 9.32


Baseline Mortality

Pandemic Mortality

Δ = 1.77

Panel D: COVID-19 (2020)

Baseline Mortality

Pandemic Mortality

Δ = 1.51

Source: Panel A: Authors’ calculations from Shattuck (1846), Simonds (1850), Clark et al. (1850), Wright (1855), Curtis (1856), and Bushee (1899); Panel B: Authors’ calculations from Bureau of the Census (1919, 1920, 1922) and National Center for Health Statistics (2020a); Panel C: Authors’ calculations from National Center for Health Statistics (2020c); Panel D: Authors’ calculations from National Center for Health Statistics (2020a, 2020c, 2021a) and United States Census Bureau (2020a).

Note: Figure reports mortality rate estimates by group during selected pandemics in U.S. history and for proximate non-pandemic years. Area shaded blue denotes recorded mortality from the pandemic disease (and in the case of influenza/pneumonia, includes seasonal baseline mortality patterns from influenza and/or pneumonia), while area shaded grey denotes mortality from other causes. See Data Appendix for computational details.

Figure 1: Changes in Mortality for Different Groups during Pandemics, 1849-2020
The third section of the paper examines different explanations for why COVID-19 has had such unequal health effects, with a focus on racial and ethnic disparities. We provide a simple framework for organizing factors that contribute to the observed gradients, and consider whether they are driven by pre-existing differences in health risks and socioeconomic factors, or by differential impacts from the same across advantaged versus disadvantaged groups. We conclude by pointing out that the patterns of health inequality seen during the pandemic mirrored those that existed in the United States prior to COVID-19, and offer thoughts about how the evolution of these gradients and resilience for the next pandemic will depend on technology, health policy, and broader social policy.

**Measuring the Overall Initial Health Impact of COVID-19**

The first case of SARS-CoV-2 infection in the United States was reported on January 20, 2020 (Holshue et al., 2020). From then to the end of the year, there were an additional 20.4 million confirmed infections nationwide (Centers for Disease Control and Prevention, 2021c). Of these cases, 378,000 resulted in death from COVID-19, the disease caused by SARS-CoV-2. By the measure of confirmed deaths from disease, the COVID-19 pandemic ranks among the deadliest in United States history, comparable to the 1918 influenza and HIV pandemics (Goldstein and Lee, 2020).

While the health effects of the COVID-19 pandemic have clearly been significant, quantifying them is complex. A first challenge in measurement is data quality, which varies substantially by outcome of interest. Case reports of COVID-19 are often submitted with little information on patient demographics or their severity; 37.5 percent of cases in the Centers for Disease Control and Prevention’s (CDC) COVID-19 surveillance system are missing race/ethnicity data and 88.4 percent lack information on underlying health conditions. Although reports of "long COVID” indicate that the disease may have persistent health effects among some of those infected, systematic data on the morbidity impacts of COVID-19 are scarce (COMEBAC Study Group, 2021). In light of these data constraints, we focus on COVID-19’s effects on mortality, which is a key health outcome of interest and reported by law to the CDC (National Research Council, 2009).

Quantifying the impact of COVID-19 on mortality still has challenges. While the number of deaths attributed to COVID-19 disease in 2020 indicates the effects have been substantial, this figure may be an underestimate of the pandemic’s mortality impacts. An estimated 3.4 million deaths occurred in the United States during 2020, an increase of 504,000 from the 2.9 million deaths which occurred during 2019. Evolving and variable clinical presentations alongside failures in testing, both of which characterized the early pandemic, may have resulted in deaths from COVID-19 go-

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2See Appendix Table 1 for details.

3We also do not compute period life expectancy at birth. As discussed in Goldstein and Lee (2020), “in the context of epidemic mortality, life expectancy at birth is a misleading indicator, because it implicitly assumes the epidemic is experienced each year over and over again as a person gets older.” Estimates of reductions in US life expectancy during 2020 are nevertheless striking: Andrasfiy and Goldman (2021) project a 1.13 year decline in 2020 life expectancy at birth compared to a scenario without COVID-19.
Effects of COVID-19 on the Demand for Health Care

Widespread avoidance of health care facilities early in the pandemic has been documented. The CDC’s Morbidity and Mortality Weekly Report from September 2020 estimated that 41 percent of U.S. adults had delayed or avoided medical care due to concerns about COVID-19, including 12 percent who had avoided urgent or emergency care (Czeisler et al., 2020b). In a nationally representative survey of 1337 adults conducted in July 2020 by researchers at Johns Hopkins, 29 percent of respondents who reported needing care forwent it due to fear of viral transmission, with 7 percent forgoing care due to financial repercussions of the pandemic (Anderson et al., 2021).

Similar results are found in the Census Bureau’s Household Pulse Survey, a repeated cross-section of 1.8 million U.S. adults from the Bureau’s Master Address File (United States Census Bureau, 2020b). Of those surveyed between April and December 2020, 37 percent reported having delayed medical care over the previous four weeks due to the ongoing pandemic. The Pulse data indicate that delay of care followed the general contours of national COVID-19 prevalence, reaching a peak through the spring of 2020, declining in late summer, and plateauing at a lower level in early autumn before rising again. The share delaying care, however, topped 30 percent in every week the Pulse survey was fielded, and stood at 35 percent in December — nine months after the nation’s national emergency began.

Effects of COVID-19 on the Supply of Health Care

On the supply side, many non-emergency interventions were suspended due to the pandemic. Shortly after President Donald Trump declared the COVID-19 outbreak a national emergency in March 2020, the Centers for Medicare and Medicaid Services (CMS) recommended the cancellation or delay of most elective surgeries and non-emergency medical, surgical, and dental procedures (Centers for Medicare and Medicaid Services 2020). National and state-level policies sought to curtail patient volume in order to conserve scarce personal protective equipment, free up beds and personnel for COVID-19 patients, and reduce SARS-CoV-2 transmission. These changes may have elevated non-COVID-19 morbidity and mortality (Chen and McGeorge, 2020).

The sharp reductions in volume and increased costs providers faced during the pandemic resulted in financial distress for many. The $187 billion in federal aid allocated to providers during the crisis exhibited little relationship to COVID-19 disease burden or hospital financial health, and failed to save many struggling providers even as well-resourced hospital networks, their losses cushioned with aid, engaged in a renewed wave of consolidation (Kakani et al. 2020, Abelson 2021). The closure of an estimated eight percent of physician practices and a record number of

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4Symptomatic COVID-19 disease varies from mild to severe and can lead to death in a variety of ways. Pneumonia and respiratory failure are prominent final pathways, but cardiac conditions, embolic events and systemic inflammation are also possible (Gupta et al. 2020, Malas et al. 2020, Jose and Manuel 2020, Long et al. 2020).
rural hospitals, along with the higher prices and lower quality of care generally accompanying provider consolidation, may affect patient outcomes in the longer term (Physicians Foundation 2020, Basu et al. 2019, Gaynor 2018).

On net, the pandemic’s impact on the supply and demand channels described above resulted in extraordinary declines in health care utilization. Non-COVID-19 medical admissions fell by 40 percent during the first wave of COVID-19 and remained depressed nearly a year later (Birkmeyer et al. 2020, Heist et al. 2021). The implementation of policies such as stay at home or business closure orders may have contributed to the fall in outpatient visits (Ziedan et al., 2020). While substitution to telemedicine partly offset the drop in in-person care, important preventive services such as vaccinations and screenings could not be shifted online and saw precipitous declines – one study using data on over 5 million individuals with employer-sponsored insurance found decreases of 22 percent in vaccinations among children aged 0-2, 67 percent in mammograms among women 46-64, and 70 percent in colonoscopies among individuals aged 46-64 (Patel et al. 2021, Whaley et al. 2020). The consequences of these delays in care will likely reverberate in the form of delayed diagnosis of non-communicable disease, preventable cases of infectious disease, and strain on providers, long after the pandemic ends.

Additional Spillover Effects

The COVID-19 health crisis is also an economic crisis. Based on prior recessions, Ruhm (2000) has noted that mortality tends to be procyclical. Since the opportunity cost of leisure declines, individuals have more time to exercise, prepare healthy food and, in non-pandemic times, seek medical care. The quality of healthcare, particularly in nursing homes, may also display cyclical fluctuations (Stevens et al., 2015). Declines in economic activity and mobility during the pandemic recession may have led to reductions in non COVID-19 deaths, compensating in part for the rise in mortality from infectious disease and delayed care. While there is some suggestive evidence of declines in air pollution and motor vehicle deaths from the early months of the pandemic, as well as a decrease in seasonal flu deaths resulting from reduced social interaction, averted deaths are likely small in number (Cicala et al. 2020, Centers for Disease Control and Prevention 2021a).

Unlike in most recessions, protective measures taken by governments and individuals to limit disease transmission during the pandemic resulted in unprecedented levels of social isolation. Disruptions in daily routines, community life, and support systems were accompanied by a troubling increase in substance use disorder, with the CDC’s Morbidity and Mortality Weekly Report from August 2020 estimating that 13 percent of US adults had started or increased substance use to cope with the pandemic’s effects (Czeisler et al., 2020a). Following three years of relative stability, drug overdose deaths nationwide sharply rose beginning in April 2020, the first full month of the COVID-19 national emergency, and grew through October 2020 (National Center for Health Statistics, 2021b). The rise in substance abuse deaths concomitant with the pandemic suggests another avenue through which the COVID-19 crisis may have indirectly elevated mortality.
Excess Deaths: A Summary of COVID-19’s Mortality Effects

The many ways through which the COVID-19 pandemic affected mortality renders precise attribution to any one cause challenging. Indeed, as individuals are often at risk for more than one type of death, some deaths recorded as due to COVID-19 disease would have occurred even in the absence of the pandemic (Gichangi and Vach, 2005). Figure 2 illustrates how the process of assessing the mortality toll of COVID-19 is complicated by the phenomena of substitution between different causes of death (competing risks), indirect deaths (spillovers), and averted deaths. The intractability of individually ascertaining the number of deaths resulting from each possible cause has motivated the use of "excess deaths" to capture the overall effect of the pandemic.

Excess deaths refer to differences between observed deaths in a particular time period and historical or expected deaths in a similar time period (National Center for Health Statistics, 2020d). As Figure 1 suggests, years in which the United States experienced an infectious disease epidemic demonstrate higher death rates than contemporaneous non-epidemic years, with the increase in deaths attributable to both mortality among infected individuals and a net increase in deaths from other causes. The sum of recorded deaths from the disease and the difference in deaths from all other causes compared to proximate time periods provides a summary statistic of the epidemic’s total effect on mortality. This number of excess deaths, which can be adjusted to account for pre-existing mortality trends, is typically then divided by the size of the population to yield a rate of excess mortality.

Following Polyakova et al. (2020, 2021), we estimate excess mortality as the deviation from a linear mortality trend. Figure 3 plots all-cause mortality rates among all Americans for each year from 2011 to 2020, using death data from the National Center for Health Statistics and population estimates from the American Community Survey (National Center for Health Statistics 2021a, United States Census Bureau 2020a). After declining throughout much of the 20th century, mortality rates in the United States have generally risen since 2010, in part due to the nation’s aging
population. The number of deaths recorded in 2020, however, was far above the number expected based on prior trends. The deviation in the 2020 all-cause mortality rate from the 2011-19 trend was 1.39 deaths per 1000 population, or a 15.4 percent increase relative to trend. By comparison, the mortality rate from COVID-19 disease in 2020 was 1.08 per 1000 population, suggesting that spillovers contributed to elevated mortality during the pandemic.

![Figure 3: All-Cause Mortality Rates in the United States, 2011-2020](image)

Source: Authors’ calculations from National Center for Health Statistics (2020a, 2020e, 2021a) and United States Census Bureau (2020a).

Note: Figure plots mortality rates from all causes for the United States from 2011 through 2020. The difference in the 2020 mortality rate compared to the 2011-19 linear trendline is labeled. Mortality rates are not adjusted for age.

### Measuring COVID-19 Health Inequality

The health consequences of the COVID-19 crisis affected Americans of all backgrounds, with over half of respondents in a Pew Research Center survey reporting that they knew someone personally who had died or been hospitalized due to the disease (Pew Research Center, 2020). The toll of the pandemic, however, fell most heavily on Black, American Indian, and Hispanic individuals, who disproportionately bore the total mortality burden of COVID-19 in at least two ways: they died almost always at greater rates, and they died at younger ages. We examine inequality in pandemic-related mortality, with a focus on inequality by race and ethnicity, in the section below.
Excess Mortality by Race/Ethnicity and Age

Having estimated all-cause excess mortality during the first year of the pandemic for all Americans, we replicate this approach by race/ethnicity and age. Panel A of Figure 4 reports excess mortality rates in 2020 for non-Hispanic White, non-Hispanic Black, Hispanic, and non-Hispanic Asian Americans within six age groups (0-44, 45-54, 55-64, 65-74, 75-84, and 85 and over).5

Panel A: Excess Mortality Rates

![Excess Mortality Rates Graph](image)

Panel B: Ratio of Excess Mortality Rates to non-Hispanic Whites

![Ratio of Excess Mortality Rates Graph](image)

Source: Authors' calculations from National Center for Health Statistics (2020e, 2021a) and United States Census Bureau (2020a).

Note: The rate ratio for non-Hispanic Whites is 1 by construction within each age group, and is shown for reference. The rate ratio for Asians age 0-44 is omitted as this group did not experience excess mortality.

Figure 4: All-Cause Excess Mortality in 2020 by Race/Ethnicity and Age Group

5While American Indians and Alaska Natives (AIAN) experienced high rates of COVID-19 infection during the pandemic, we do not assess excess mortality among these populations due to known data quality issues (Yellow Horse and Huyser 2021, National Center for Health Statistics 2021c). Using data from a limited set of states, a CDC Morbidity and Mortality Weekly Report from December 2020 found that disparities in COVID-19 death rates between AIAN and non-Hispanic White individuals were large and particularly so at younger ages (Arrazola et al., 2020).
Pandemic all-cause excess mortality rises sharply with age, largely because age is the strongest single determinant of mortality from COVID-19 (Petrilli et al., 2020) and because avoided non-COVID-19 healthcare is more likely detrimental to the oldest adults. Appendix Table 2 shows how Americans age 70 and above experience case fatality rates (rates of death conditional on diagnosis) about 200 times higher than those below age 40. Rates of excess mortality at any given age, however, vary sharply by race and ethnicity. Panel B of Figure 4 plots the ratio of non-Hispanic Black, Hispanic, and non-Hispanic Asian excess mortality rates to the non-Hispanic White excess mortality rate for each age group.

The rate ratios presented in Panel B point to two dimensions of mortality disadvantage that Black and Hispanic Americans faced. First, the Black-White and Hispanic-White ratios are above one at every age, indicating that Blacks and Hispanics experienced elevated rates of excess death compared to non-Hispanic Whites. Indeed, when all age groups are pooled and excess mortality is computed for each race and ethnicity as a whole, it is evident that Black and Hispanic populations suffered the highest rates of excess death in 2020. Panel A of Figure 5 shows that Black Americans experienced excess mortality of 2.14 deaths per 1000 population in 2020, or a 25.0 percent increase in mortality relative to trend, while Hispanic Americans in 2020 saw excess mortality of 1.44 deaths per 1000 population in 2020, or a startling 39.5 percent rise relative to trend. Non-Hispanic Whites and Asians experienced increases in mortality of 1.29 and 0.58 deaths per 1000 population respectively, which are both increases of less than 15 percent relative to trend. As Panel B of Figure 5 indicates, these disparities widen further when excess mortality rates are age-adjusted to account for differences in age distributions between races and ethnicities — namely, the younger Hispanic and older non-Hispanic White age structures.6

Second, the ratios in Panel B of Figure 4 are strikingly higher at younger ages compared to older groups. The Black-White ratio in excess mortality is above four for the youngest age group of 0-44, and above three for all age groups through 65-74. Similarly, the Hispanic-White ratio in excess mortality is above 2.5 for all age groups from 0-44 to 65-74. By contrast, the Black-White and Hispanic-White excess mortality ratios for individuals age 85 and over are a comparatively low 1.5. The steep age gradient in excess mortality disparities indicates that the already high number of Black and Hispanic pandemic-related deaths disproportionately occurred among the young.

Mortality rates, whether age-adjusted or unadjusted, do not differentiate between deaths at younger or older ages. Examining death rates alone therefore overlooks an important dimension of inequality: differences in the number of years individuals in each group would likely have lived had they not died due to pandemic-related causes.

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6Age-adjusted statistics are computed by weighting deaths in different age groups among a given race or ethnicity in a manner that matches the share of each age group in the general population. Under age-adjustment, races and ethnicities with an age distribution younger than the general population have deaths at older ages weighted relatively more and deaths at younger ages weighted relatively less, whereas the converse would be true for races or ethnicities with an age distribution older than the general population.
Excess mortality rate in 2020, unadjusted for age (per 1000 population)
White Black Hispanic Asian
Panel A: Crude Excess Mortality

Excess mortality rate in 2020, adjusted for age (per 1000 population)
White Black Hispanic Asian
Panel B: Age-Adjusted Excess Mortality

YPLL-65 (per 100,000 population age 0-64)
White Black Hispanic Asian
Panel C: Years of Potential Life Lost by Age 65

Source: Authors’ calculations from National Center for Health Statistics (2020a, 2020e, 2021a) and United States Census Bureau (2020a).

Note: Figure reports excess mortality rates unadjusted for age (Panel A), excess mortality rates adjusted for age (Panel B), and excess rates of years of potential life lost before age 65 (Panel C) during 2020.

Figure 5: Measures of Racial and Ethnic Disparities in COVID-19 Pandemic Mortality

Disparities in Lost Years of Life

The concept of years of potential life lost, or YPLL, is used in the public health literature to quantify premature mortality (Gardner and Sanborn, 1990). As premature death is typically defined as a death occurring before age 65, the years of potential life lost for an individual who died prematurely is calculated by subtracting their age at death from 65, with those dying at age 65 or older assigned a years of potential life lost value of zero. The total years of potential life lost in a population is then computed by summing up the years of potential life lost among all individuals in the population who died early. This sum is usually normalized by dividing it by the number of individuals in the population under age 65. The aim of the years of potential life lost statistic is to measure life lost in years of life foregone as opposed to deaths incurred, providing a complementary measure to mortality rates.

As with excess death rates, we estimate "excess" years of potential life lost during the COVID-19 pandemic as the deviation in 2020 from the 2011-19 linear trend. Panel C of Figure 5 plots excess years of potential life lost from all causes per 100,000 individuals under age 65 by race and ethnicity during 2020. Among all Americans, rates of YPLL rose by 12.8 percent in 2020 compared to trend. Similarly with mortality rates, however, this increase was again concentrated among Black and Hispanic Americans. Black rates of YPLL rose in 2020 by 1350.3 per 100,000 individuals under 65, or an increase of 19.5 percent relative to the 2011-19 trend. Among Hispanics, YPLL rose in 2020 by 925 per 100,000 individuals under 65, or 29.2 percent relative to trend. By comparison, YPLL rates among non-Hispanic Whites increased by 316.2 per 100,000 individuals under 65, or 8.4 percent, and among Asian Americans YPLL rates fell slightly relative to trend.

Examining ratios of Black and Hispanic excess YPLL rates to the non-Hispanic White YPLL rate, in a similar manner to Figure 4, underscores the mortality disadvantage Blacks and Hispanics have faced during the pandemic. The Black-White ratio in excess years of potential life lost is 4.2,
whereas the Hispanic-White ratio in excess years of potential life lost is 2.9. In contrast, the Black-
White ratio for age-adjusted excess mortality is 2.6, and the Hispanic-White age-adjusted excess
mortality ratio is 2.5. The elevated YPLL ratios suggest that not only have Black and Hispanic
Americans died at greater rates during the pandemic, but those who died on average had many
more years of life left to live. Far from being an equal opportunity pathogen, SARS-CoV-2 has
exposed societal cleavages between less privileged and more advantaged groups.

Understanding the Unequal Health Effects of COVID-19

The COVID-19 pandemic affected some groups, particularly Black and Hispanic Americans, more
than others. Why was this the case? Our goal in this section is twofold: to provide a simple frame-
work for organizing the main factors that contribute to the observed disparities, and to present the
results of a decomposition that examines the relative importance of some of these factors.

A Simple Framework for Understanding COVID-19 Health Inequality

We begin by focusing on deaths directly associated with COVID-19. Inequality in COVID-19 dis-
ease may be due to social determinants (such as differences in occupation, income or education),
medical determinants (including differences in comorbidities, health care quality and insurance),
and long-standing institutional features that perpetuate systemic racism and intergenerational
poverty (Snowden and Graaf, 2020). These factors are not exhaustive nor are they mutually exclu-
sive. They can, however, be mapped into an expanded model of disease transmission.

The probability of death from COVID-19 is the product of the probability of SARS-CoV-2 in-
fection and death from COVID-19 conditional on infection. Following standard epidemiological
models, the probability of SARS-CoV-2 infection can be written as $P(\text{SARS-CoV-2 infection}) = 1 - (1 - p)n_i(1 - m_i)$, where $i$ refers to an individual, $p$ represents prevalence, $n$ is the number of
contacts, and $m$ is the proportion of mitigated contacts (Halloran, 2009). All else equal, the prob-
bility of infection rises as prevalence in the community increases, and also increases if one has
more contact with others. An individual’s infection probability, however, declines if more mitigat-
ing measures, such as mask-wearing, physical distancing, and vaccination, are taken.

Prevalence ($p$) in an individual’s community and their number of contacts ($n$) can be affected
by social inequality, population density, and local policies. Black and Hispanic populations both
live in areas with higher COVID-19 prevalence and face higher costs of reducing their number of
contacts. Black and Hispanic Americans live in larger households that are often multi-generational,
and often have poor housing conditions (Pew Research Center, 2018). They are also more likely
to be frontline workers who must work in-person despite the risk of infection, and cannot stop
working or cut back on hours due to relatively low wealth levels or, particularly in the case of
Hispanics, challenges in accessing federal benefits (Blau et al., 2020).

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7 Prevalence can be expressed as a function of the behavior of others around the individual ($p(n_{-i}(1 - m_{-i}))$), which
in turn is affected by the factors noted above.
Mitigation \( m \) can help offset the risk associated with labor or leisure-related interactions. However, ability to follow public health guidance depends on access to public health information, complementary tools such as vaccines and masks, and beliefs in the credibility of such information. A survey of approximately 5000 Americans conducted early in the pandemic showed that Black and Hispanic individuals, younger people, and men were less likely to have accurate information about COVID-19 transmission than other groups (Alsan et al., 2020a). Mitigation behavior by individuals during the pandemic has also been shaped by the dissemination of misinformation and features of the messenger, such as whether they are of the same race or ethnicity as the individual or whether they are an expert or peer (Simonov et al. 2020, Alsan et al. 2020b, Alsan and Eichmeyer 2021). Although communication with Black physicians has been shown to raise take-up of preventive health among Black Americans, just 4 percent of physicians in the United States are Black (Alsan et al., 2019).

As discussed above, mortality during the COVID-19 pandemic can be conditional on infection (i.e. direct) or due to spillover effects. Access to quality health care is important in determining mortality, direct and indirect, from COVID-19. While higher-quality hospitals are associated with lower mortality rates, low-income, Black, and Hispanic Americans obtain health care from lower-quality facilities (Jha et al. 2011, Doyle et al. 2019, Chandra et al. 2020). Elevated COVID-19 caseloads in Black and Hispanic communities also contribute to non-COVID-19 excess deaths by reducing the ability of health care facilities to treat non-COVID-19 patients and causing individuals to avoid or delay necessary care due to fear of contagion. Black and Hispanic adults were also less likely at the start of the pandemic to be covered by health insurance, potentially contributing to care delays (Cohen et al., 2020).

The distribution of pre-existing conditions differs by race and ethnicity, raising the probability of death conditional on COVID-19 infection and the inability to receive needed care for other chronic illnesses. Relative to non-Hispanic Whites, rates of diabetes are 1.7 and 1.8 times higher among Black and Hispanic populations (Centers for Disease Control and Prevention, 2020). Rates of obesity and hypertension are similarly elevated among Black and Hispanic individuals as well (Reeves and Smith, 2020). Diabetes, obesity, and hypertension are conditions that increase the risk of death from COVID-19 (Centers for Disease Control and Prevention, 2021b).

**COVID-19 Health Inequality: A Decomposition**

The potential drivers of health disparities seen during the COVID-19 pandemic are manifold. Black and Hispanic Americans are disadvantaged socioeconomically relative to non-Hispanic Whites and tend to have a greater number of comorbidities that heighten the risk of severe COVID-19 disease. To what extent are racial and ethnic disparities in COVID-19 health outcomes driven by differences in these characteristics? Or, are Black, Hispanic, and non-Hispanic White individuals differentially impacted by COVID-19 even when they possess the same attributes? We aim to examine the relative importance of each of these factors through a decomposition exercise. As CDC datasets largely lack detailed comorbidity data or information on individuals who have *not*...
contracted COVID-19, we obtain the necessary data from the Optum Clinformatics® Data Mart (CDM), a comprehensive commercial and Medicare Advantage claims database. In addition, we shift focus from the relatively rare outcome of mortality to COVID-19 hospitalizations.

The Optum database includes approximately 67 million unique lives of all ages across 2007-2020 and is broadly representative geographically. We include adults age 21 or older who identify as either non-Hispanic White, non-Hispanic Black, or Hispanic, enrolled prior to July 2019, and filed a medical claim at least once during 2019 (thus avoiding cases where comorbidities have been undiagnosed and allowing us to have three quarters of data prior to the peak of the pandemic’s first wave). Our analysis sample includes all enrollees who were hospitalized for COVID-19 during the first three quarters of 2020, along with a five percent random sample of those not hospitalized for COVID-19 as a control group.⁸ Our final sample includes approximately 322,000 non-Hispanic White, 50,000 non-Hispanic Black, and 61,000 Hispanic enrollees.

We first measure whether sampled enrollees were previously diagnosed with medical conditions that increase the risk of severe illness from COVID-19.⁹ We also extract social and demographic information including age, sex, average educational attainment in the enrollee’s Census block of residence, and Census division of residence. We conduct a “threefold” Oaxaca-Blinder decomposition that parcels out racial and ethnic differences in the likelihood of hospitalization due to COVID-19 into three components (Jann, 2008). The first component captures how much of the gap is from group differences in the predictors (the “endowments”). The second component captures the part due to differences in the coefficients (the “return to endowments”). The third component is the interaction between endowments and returns to endowments.

Furthermore, we perform a “detailed” decomposition, as we track two sets of predictors: comorbidities, which are indicators for the medical conditions we measure, and sociodemographics.¹⁰ The decomposition is formulated from the viewpoint of Black or Hispanic enrollees. Although it is well-known that the reverse decomposition from the White perspective may provide different results, our approach is designed to produce two relevant counterfactuals. First, what is the expected change in hospitalization rates for Black and Hispanic Americans if the relevant minority group had the majority group’s predictor levels? Second, what would be the expected change if the minority group had the majority group’s coefficients?¹¹

Table 1 presents the decomposition results. The left panel displays results for non-Hispanic Black vs. non-Hispanic White enrollees and the right panel displays results for Hispanic vs. non-Hispanic White enrollees. The hospitalization rate for COVID-19 in our constructed sample is

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⁸COVID-19 testing and results are not reliably included in claims data; moreover, testing was not necessarily evenly distributed across groups (Rubin-Miller et al., 2020).

⁹We extract information on hypertension, diabetes, obesity, cancer, heart disease, and chronic obstructive pulmonary disease based on diagnosis codes in claims filed between January 1 and December 31, 2019.

¹⁰For categorical variables, such as census block educational attainment and census division, we follow Yun (2005) to normalize the effects of categorical variables to avoid the issue of varying coefficients due to the choice of omitted group.

¹¹We estimate with a linear probability model following Montenovo et al. (2020), but note that a logistic model provides similar results in terms of the importance of coefficients. See Data Appendix for additional data and methodological details.
7 percentage points higher for Black than White enrollees and 4.6 percentage points higher for Hispanic than White enrollees. For both groups, coefficients contribute much more to the overall difference than predictors. Perhaps surprisingly, the presence of comorbidities explains a much smaller share of the overall difference than the return to sociodemographics for both Black and Hispanic groups. Indeed, the return to sociodemographics is the single largest contributor to the overall gaps in hospitalization, accounting for 28.6 percent of the raw gap for Black compared to White enrollees and 56.7 percent of the raw gap for Hispanic compared to White enrollees. The results indicate that the same predictors exert a more damaging impact on Black and Hispanic sampled enrollees. As an example, Black or Hispanic male enrollees might be more likely than White male enrollees in similar situations to be engaged in work-related activities that place them at higher risk of contracting the disease and/or have less access to care and therefore present at a later stage, thus requiring hospitalization.

In sum, the decomposition results suggest that the stark differences in COVID-19 health outcomes for Black and Hispanic Americans compared to non-Hispanic Whites cannot be attributed to a greater prevalence of pre-existing conditions, lower neighborhood levels of educational attainment, or (broad) geographical disadvantage alone. Rather, otherwise similar Black and Hispanic individuals, all of whom are insured in our sample, are hospitalized due to COVID-19 at a higher rate than non-Hispanic Whites. These results are specific to our sample and decomposition decisions we have taken, but they are consistent with the broader narrative that Black and Hispanic individuals face institutional disadvantages including inconsistent providers, lower-quality care, and systemic racism, that worsen their returns to similar endowments and contribute to COVID-19 health inequality.

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12The difference in intercept contributes negatively to the differences in returns to sociodemographic variables.
<table>
<thead>
<tr>
<th>Overall Gap in Sample</th>
<th>Black vs. White</th>
<th>Hispanic vs. White</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Comorbidities</td>
<td>Sociodemographics</td>
</tr>
<tr>
<td><strong>Endowments</strong></td>
<td>0.011</td>
<td>0.007</td>
</tr>
<tr>
<td>Percent of Total Difference</td>
<td>16.2%</td>
<td>10.1%</td>
</tr>
<tr>
<td><strong>Returns to Endowments</strong></td>
<td>0.016</td>
<td>0.020</td>
</tr>
<tr>
<td>Percent of Total Difference</td>
<td>22.8%</td>
<td>28.6%</td>
</tr>
<tr>
<td><strong>Interaction</strong></td>
<td>0.009</td>
<td>0.007</td>
</tr>
<tr>
<td>Percent of Total Difference</td>
<td>12.4%</td>
<td>9.9%</td>
</tr>
<tr>
<td><strong>Number of Obs.</strong></td>
<td>371,483</td>
<td></td>
</tr>
</tbody>
</table>

**Source:** Authors’ calculations from Optum (2021).

**Notes:** Table reports results from a threefold Oaxaca-Blinder decomposition from the perspective of the minority group. Sociodemographics includes age and age squared, gender, education dummies, and census division fixed effects. Effects of education and census division are normalized. Comorbidities represent a series of dummy variables for hypertension, diabetes, obesity, heart disease, cancer and chronic obstructive pulmonary disease. The difference in intercepts is included in the difference in returns to endowments of sociodemographics.
Conclusion

This paper has drawn on history, medicine and economics to place the initial health effects of the current pandemic into broader context. That COVID-19 disproportionately killed the frail and disadvantaged could be expected based on viral dynamics, past epidemics and marked differences in the ability of individuals to protect their health during the crisis. Yet the heightened salience of these patterns, along with the stark mortality burden of the pandemic, may serve as a catalyst for change — in particular, for changing how Americans view the importance of public health and the social safety net (Rees-Jones et al., 2020).

The initial health effects we outlined may have consequences for years to come. Most directly, a growing body of evidence suggests that a substantial proportion of individuals infected with COVID-19 suffer a range of long-term health consequences, including cognitive dysfunction, fatigue, and injury to the heart and lungs (COMEBAC Study Group 2021, del Rio et al. 2020). The pandemic’s long-term effects, however, will likely reach past those who contracted the disease and extend beyond health alone. Disruptions in screenings and routine health care may beget future premature morbidity and mortality from other communicable and non-communicable diseases (Chen and McGeorge, 2020). Scarring in utero exhibited in the influenza pandemic of 1918 and other epidemics of infectious disease may emerge with consequences for disability, educational attainment, and earnings (Almond, 2006). The disruptive effects of the COVID-19 crisis on education could widen inequality in income and health for future generations. These factors may exacerbate immediate economic disparities already experienced through labor markets as a consequence of the pandemic (Montenovo et al., 2020).

With the introduction of recently approved vaccines against COVID-19 has come hope that the disruption the disease has wrought on health and society will soon cease. The same health gradients seen during the country’s descent into the pandemic, however, are likely to be observed as we emerge from it. Reports suggest that vaccination distribution by race and ethnicity has not been aligned with who has been affected most by the virus, placing vulnerable individuals at risk of adverse outcomes during a time in which SARS-CoV-2 continues to mutate (Ndugga et al. 2021). The medium and long-run health effects of COVID-19 — as well as the consequences of future novel infectious disease outbreaks which will assuredly emerge — will be shaped by how effectively and equitably policymakers respond to these formidable, yet not wholly unprecedented, challenges.
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Appendix

Appendix Figure 1: All-Cause Mortality by Race/Ethnicity, 2019 and 2020

Source: Authors’ calculations from National Center for Health Statistics (2020a, 2020e, 2021a) and United States Census Bureau (2020a).

Note: Figure plots age-standardized mortality rates from COVID-19 and other causes for non-Hispanic Asian, Hispanic, non-Hispanic White, non-Hispanic American Indian and Alaska Native (AIAN), and non-Hispanic Black Americans during 2019 and 2020. This figure expands on Figure 1, which displays age-standardized mortality rates in 2019 and 2020 for non-Hispanic White and Black individuals. Labels above each bar indicate the difference in the age-standardized rate of mortality in the given year between the racial/ethnic group of interest and non-Hispanic Whites. Area shaded blue denotes recorded mortality from the pandemic disease, while area shaded grey denotes mortality from other causes.
### Appendix Table 1: Missing Elements in the CDC Case Surveillance Datafile

<table>
<thead>
<tr>
<th></th>
<th>(1) Age</th>
<th>(2) Sex</th>
<th>(3) Race/Ethnicity</th>
<th>(4) Comorbidities</th>
<th>(5) Any</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Cases (%)</strong></td>
<td>114,238</td>
<td>180,080</td>
<td>6,198,330</td>
<td>14,603,487</td>
<td>14,938,556</td>
</tr>
<tr>
<td><strong>Hospitalizations (%)</strong></td>
<td>400</td>
<td>2,070</td>
<td>125,967</td>
<td>572,170</td>
<td>598,599</td>
</tr>
<tr>
<td><strong>Deaths (%)</strong></td>
<td>21</td>
<td>774</td>
<td>52,573</td>
<td>233,414</td>
<td>239,672</td>
</tr>
</tbody>
</table>

Source: Authors' calculations from Centers for Disease Control and Prevention, COVID-19 Response (2020).

Note: Table shows the number and percent of missing elements indicated as column headers in the CDC COVID-19 case surveillance database. Of cases, hospitalization, and deaths reported to the CDC, Columns 1 to 4 show the number and percent of missing age, sex, race/ethnicity, and comorbidities data, respectively. Column 5 shows the number and percent of missing data on any of the four elements. Percentages are reported in parentheses. Note that we do not investigate the extent to which cases are missing entirely from this data source, relative to the counts present in other standard sources of COVID-19 data.
### Appendix Table 2: COVID-19 Case-Fatality Rates (CFRs)

<table>
<thead>
<tr>
<th>Panel A: Age Groups</th>
<th>Age 0 to 39</th>
<th>Age 40 to 49</th>
<th>Age 50 to 59</th>
<th>Age 60 to 69</th>
<th>Age 70+</th>
</tr>
</thead>
<tbody>
<tr>
<td>Case Fatality Rate</td>
<td>0.126</td>
<td>0.705</td>
<td>1.885</td>
<td>5.759</td>
<td>24.307</td>
</tr>
<tr>
<td>Percent of Deaths</td>
<td>1.81%</td>
<td>2.92%</td>
<td>7.63%</td>
<td>16.68%</td>
<td>70.97%</td>
</tr>
<tr>
<td>Percent of Population</td>
<td>52.08%</td>
<td>12.28%</td>
<td>12.90%</td>
<td>11.58%</td>
<td>11.15%</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Panel B: Race/Ethnicity</th>
</tr>
</thead>
<tbody>
<tr>
<td>White</td>
</tr>
<tr>
<td>Case Fatality Rate (Crude)</td>
</tr>
<tr>
<td>Case Fatality Rate (Age-adjusted)</td>
</tr>
<tr>
<td>Percent of Deaths</td>
</tr>
<tr>
<td>Percent of Population</td>
</tr>
</tbody>
</table>


*Note:* Table reports COVID-19 related case-fatality rates (CFRs) by age groups (Panel A) and race/ethnicity (Panel B). Population and death shares by race/ethnicity are computed as shares of population and deaths with race/ethnicity recorded as non-Hispanic White, non-Hispanic Black, Hispanic/Latino, non-Hispanic Asian, or non-Hispanic American Indian. For clarity of presentation, Alaska Natives are included in the American Indian category, whereas Pacific Islanders are included in the Asian category. Of the 303,211 deaths recorded in the CDC COVID-19 case surveillance database, 3.20% are among individuals of races/ethnicities other than those noted above, while 17.34% do not contain information on race/ethnicity.