

NBER WORKING PAPER SERIES

MISFORTUNE AND MISTAKE:  
THE FINANCIAL CONDITIONS AND DECISION-MAKING ABILITY  
OF HIGH-COST LOAN BORROWERS

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Working Paper 26328  
<http://www.nber.org/papers/w26328>

NATIONAL BUREAU OF ECONOMIC RESEARCH  
1050 Massachusetts Avenue  
Cambridge, MA 02138  
September 2019

This research is carried out in cooperation with Meniga, a financial aggregation software and smartphone application provider. We are grateful to the executives and employees who have made this research possible. We thank Nicola Persico, John Gathergood, and participants in several seminars and conferences for their many helpful comments on the paper. Special thanks to Jon Zinman for his detailed and insightful comments on an earlier draft. The views expressed herein are those of the authors and do not necessarily reflect the views of the National Bureau of Economic Research.

At least one co-author has disclosed a financial relationship of potential relevance for this research. Further information is available online at <http://www.nber.org/papers/w26328.ack>

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Misfortune and Mistake: The Financial Conditions and Decision-making Ability of High-cost Loan Borrowers

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NBER Working Paper No. 26328

September 2019

JEL No. D14,G2

**ABSTRACT**

This paper studies the relationship between adverse financial conditions (“misfortunes”), imperfect decision-making (“mistakes”), and the demand for high-cost credit. Imperfect choices are hard to identify; unobserved factors can justify many behaviors as optimal. We address this by linking administrative and experimental data. Bank records from Iceland detail the financial conditions associated with high-cost loan demand. The experiments manipulate constraints, while holding preferences and beliefs constant, to identify choice imperfections and measure decision-making ability. We then relate loan demand to decision-making ability and measures of constraints and preferences. High-cost borrowers are especially illiquid in the days leading up to getting a loan. They also have much lower decision-making ability: 28% of loan dollars are lent to the bottom 10% of the decision-making ability distribution, and 53% are lent to the bottom 20%. The relationship between decision-making ability and loan demand is not explained by demographic characteristics, granular information on financial conditions, or measures of preferences from the experiments, and is mirrored by the relationship between decision-making ability and an unambiguous “mistake,” the accrual of insufficient fund fees. Estimates from U.S. survey data are quantitatively similar. The results thus provide evidence that both misfortune and mistake are important for high-cost loan demand.

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## I. Introduction

Several forms of consumer credit, including payday loans, deposit advance products, and vehicle title loans, are controversial because they are used disproportionately by low-income households and involve high fees. In 2015, lower income U.S. households spent an estimated \$62.7 billion in interest and fees on short-term loan products like these (Schmall and Wolkowitz, 2016). Critics call the loans usurious and warn that they take advantage of financially unsophisticated borrowers who end up in damaging cycles of debt. Proponents describe the high costs of the loans as necessary given the risk to the lender, and note that the costs to the borrower of defaulting on other obligations can be much greater. They argue that these forms of credit provide valuable liquidity to those who struggle to find it elsewhere.

The controversy surrounding high-cost credit has spurred both regulation aimed at protecting unsophisticated borrowers, and concern about that regulation.<sup>1</sup> The costs and benefits of this regulation depend on the extent to which demand for high-cost credit is due to “misfortune” and “mistake.” By “misfortune” we mean adverse financial conditions that cause borrowers to place high value on a loan but limit its availability at low cost. These circumstances include income, liquidity, and expenditure shocks. By “mistake” we mean an imperfect choice. A choice that, given the same information, the person would make differently if he attended to it more carefully or had greater ability to assess the factors that determine its payoff.<sup>2</sup> If borrowers turn to high-cost credit because of “misfortune,” policy is justified if it reduces market imperfections that elevate prices and limit trade in credit. If borrowers use high-cost credit because they do not properly balance its costs and benefits, policy should also work to reduce this harm to consumers.

This paper links rich administrative data with information from surveys and experiments, all at the individual level, to assess the influence of “misfortune” and “mistake” in determining the demand for high-cost credit. Evaluating the role of “mistakes” is especially challenging because

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<sup>1</sup> In 2017, for example, the U.S. Consumer Finance Protection Bureau approved rules mandating that lenders underwrite loans to ensure the borrower can pay back while meeting basic needs, and limiting the number of times lenders can attempt unsuccessfully to withdraw loan payments from a borrower’s bank account. The implementation of those rules has been placed on hold, however, as opponents raise concerns that the regulations impose important burdens on lenders and will reduce the availability of valuable credit.

<sup>2</sup> This notion of a choice imperfection relates to Gilboa’s (2012) definition of rational behavior. In this view, a person’s choices are irrational or, in our words, imperfect if he or she thinks of them erroneous after careful explanation, analysis, and consideration of their costs and benefits.

imperfect choices are hard to identify. On the one hand, unobserved constraints, preferences, or beliefs can justify many behaviors as optimal, and caution dictates respect for consumer choice. On the other hand, evidence points to the potential for “mistakes.” Prior studies show the choice to use a payday loan is sometimes ill-informed (Bertrand and Morse, 2011), may be dominated by cheaper forms of credit (Agarwal et al. 2009), and is often followed by undesirable consequences (Melzer 2011, 2018, Carrell and Zinman, 2014, Gathergood et al. 2018).

In this paper, we address this identification problem in two complementary ways. First, using daily records drawn from individual bank and credit card balances and transactions in Iceland, we describe the (changing) financial conditions associated with payday loan demand. These observational data are derived from a financial aggregation app, serving approximately 20 percent of the Icelandic adult population, that links administrative records from its users’ various financial accounts. In that analysis, we document the extent to which the individual circumstances of payday borrowers differ from that of others in the data, and how those circumstances change in the days leading up to and following the receipt of a payday loan.

Second, using the results of experiments conducted via online survey with 1,700 users of the financial aggregator, we capture measures of both economic preferences and decision-making ability. The experiments involve multiple incentivized choices under risk and uncertainty, and about the intertemporal allocation of money. The price variation in these experiments is sufficiently rich to permit well-powered tests of consistency with utility maximization and related, normative properties of choice.

Following Choi et al. (2014) and Carvalho and Silverman (2019), we interpret consistency with these normative properties of choice as a measure of financial decision-making ability. In the context of the experiments, consistency with (normatively appealing) utility maximization means the participant has revealed a single, stable, and sensible objective of the several financial choices he made while facing varying incentives over a short period of time (Afriat, 1967). We interpret revealing such an objective of choice as reflecting an ability to adequately attend to financial decisions, understand their relevant tradeoffs, and map available choices into objectives. This interpretation of consistency with utility maximization as a measure of decision-making ability is supported by evidence in studies showing these measures, as derived from experiments, are

positively correlated with financial success in both experiments and in the field, even conditional on many measures of the standard drivers of economic outcomes. See Choi et al. (2014), Stango and Zinman (2019), Carvalho and Silverman (2019).

The results from the administrative data alone show, as we would expect from prior studies, that payday borrowers have on average substantially less income than the rest of the sample and, more important, they tend to have limited access to other forms of liquidity. Over the nearly six years of observation, payday borrowers maintain on average virtually no liquid assets, and carry an average of about a month's salary in debt in the form of overdrafts on their checking accounts.

Levels of liquidity for payday borrowers are especially low in the days leading up to receiving a loan. Looking back over the 30 days prior to getting a loan, the average of a borrower's checking and savings balances, net of credit card balances, declines steadily until the day the loan arrives and then slowly recovers over the next two weeks to levels close to, but somewhat short of, the levels 30 days prior to the loan. From that point, the liquidity declines again arriving, within another two weeks, at the same critical level associated with the day before receiving a loan.

Some prior research has provided evidence on the extent to which payday loan demand is attributable to "mistake" by testing whether borrowers have access to cheaper credit at the time they take the payday loan. Results have been mixed, with some finding large fractions of payday borrowers with access to substantial amounts of credit at lower cost (Agarwal et al., 2009) and others finding that the bulk of payday borrowers have virtually no other cheaper form of market credit available when they take the loan (Bhutta et al., 2015).

In the data we study, which integrate available credit from multiple sources, a majority of payday borrowers have little if any cheaper credit available through market sources at the time they take the loan. The median borrower in the data has access to \$244 of cheaper credit when they take out the loan. There is, however, substantial heterogeneity and 25% of payday loan borrowers have more than \$1,149 in cheaper credit available on the day they receive their payday loan. In this way, the analysis of the observational data suggests a substantial but hardly a dominant role for "mistake" in driving demand for payday loans.

On their own, the administrative data results are not dispositive. Even among those who have no access to cheaper market credit, the choice to take a payday loan may not be well-

considered or best. To further examine the role of “mistake” in the demand for high-cost loans, we relate decision-making ability as measured in the experiments to demand for payday loans.

Payday loan borrowers exhibit substantially lower decision-making ability in the experiments and those with low ability play an outsized role in the market for payday loans. In these data, 28% of payday loan dollars are lent to the bottom 10% of the decision-making ability distribution, and 53% are lent to the bottom 20% of the distribution. In individual-level regression analysis, the relationship between decision-making ability and high-cost loan demand is not explained by demographic characteristics, granular information on economic circumstances, or measures of preferences from the experiment.

The negative conditional correlation between decision-making ability measures and high-cost loan demand is consistent with the view that “mistakes” are quantitatively important drivers of demand for these loans. Such inference would be misguided, however, if these measures of decision-making ability were simply capturing a “type” of consumer whose unmeasured constraints, preferences, or beliefs rationalize demand for high-cost loans.<sup>3</sup> To further evaluate this possibility, we study the relationship between measures of decision-making ability from the experiment and an unambiguous “mistake” in the administrative data. The “mistake” is the accrual of non-sufficient funds (NSF) fees. These fees obtain when, in the process of using a debit card to make a purchase, an individual exceeds his or her checking account overdraft limit. Different from costly overdrafts in markets like the U.S., there is no benefit to exceeding the limit because the purchase will not be authorized. In this way, a choice that results in an NSF fee appears clearly imperfect; it is dominated by the decision not to try to make the purchase. NSF fees can thus provide further evidence on the validity of using experimental measures of consistency with (normatively appealing) utility maximization as measures of decision-making ability.

The results on NSF fees are qualitatively similar to those for high-cost loan demand. Conditional on demographic characteristics, economic preferences, and financial conditions, those with lower decision-making ability incur significantly more NSF fees. These results are consistent

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<sup>3</sup> The distinction may be blurred if some of those constraints could, themselves, be produced by prior “mistakes.” An obvious example is the level of liquidity that results from having taken out a payday loan in the past. If an earlier decision to take a payday loan was a “mistake” an individual’s current level of liquidity, treated as a constraint in our analysis, would in fact be a consequence of an earlier “mistake.”

with the view that decision-making ability, as captured in the experiments, measures a set of skills useful for avoiding financial mistakes in field settings. This evidence bolsters the view that high-cost credit is, holding financial circumstances fixed, disproportionately taken up by those who struggle to make financial decisions that are consistent with their objectives.

Last, we evaluate the external relevance of the Iceland findings and the potential for relying on survey data alone to do similar analyses, by comparing, to the extent possible, the relationships estimated there with those estimated from a survey of U.S. consumers. The U.S. survey data on economic outcomes are self-reported and the measures of high-cost credit take-up, preferences, and decision-making ability, are relatively coarse. Nevertheless, we find that the relationship between decision-making ability and the probability of receiving a payday loan is very similar in these U.S. data and in the Icelandic data.

Taken together, the linked administrative and experimental measures from Iceland, augmented by U.S. survey data, indicate that both misfortune and mistake are relevant for high-cost loan demand. These results suggest that policy aimed at these markets should be concerned both with the possibility that market imperfections limit trade and, at the same time, that “mistakes” lead to excess trade in these kinds of loans.

## **II. Related Literature**

This paper contributes to a literature on high-cost credit, the financial conditions of borrowers in those markets, and the consequences of access to these loans. Prominent examples from that literature include Agarwal et al. (2009), Zinman (2010), Melzer (2011, 2018), Morse (2011), Bertrand and Morse (2011), Bhutta et al. (2015), Bhutta et al. (2016), Gathergood et al. (2018), and Skiba and Tobacman (2018a,b). Our paper is distinguished from the bulk of that literature by its use of comprehensive, high-frequency, administrative data on the balances and transactions of the study sample. Prior studies with access to administrative data have used credit files to observe debt and the availability of other sources of credit, but not the entire balance sheet of the consumer over time. In this way, we obtain a granular view of the financial circumstances of high-cost credit borrowers and thus understand better the extent to which “misfortunes” drive their demand for these loans.

Our reliance on administrative records from a financial aggregator relates to a growing literature that uses these kinds of data to study a variety of phenomena. Examples include Gelman et al. (2014, 2018), Kueng (2018), and Baker (2018). In particular, these Icelandic data have been used to study the dynamics of liquid asset holdings and spending in response to income (Olafsson and Pagel, 2018), how different generations use financial products to manage their personal financials (Carlin, Olafsson, and Pagel, 2019), and how consumers use credit lines in response to transitory income shocks (Hundtofte, Olafsson, and Pagel, 2019).

Our interest in measuring consistency with utility maximization, and relating it to observable characteristics and behavior, connects our work to the literature that has developed different measures of economic rationality (Dean and Martin, 2016; Halevy et al. 2018, Polisson et al. 2019; Echenique et al. 2019) and a literature that has used such measures to study the correlates and determinants of rationality (Carvalho and Silverman 2019, Banks et al. 2019, Kim et al. 2018). Our analysis draws on elements of this literature in its use of recent advances in revealed preference tests of (the degree of) consistency with different axioms of choice.

The paper contributes to these literatures, and to a broader stream of research that seeks to understand the fundamentals of economic behavior through granular financial data, by linking comprehensive administrative records to survey measures. To our knowledge, similar linkages are rare (an exception is Epper et al. (2018), which links experiments to yearly snapshots of assets and liabilities) and no other study has linked experimental economic data to comprehensive and high-frequency bank data at the individual level. This link between survey and comprehensive bank data is key to understanding the role of “misfortunes” and “mistakes” in driving demand for high-cost credit. It allows both an assessment of the extent of “misfortune” and isolation of the roles of preferences and decision-making ability in determining demand for high-cost credit.

### **III. Background – Consumer Credit in Iceland**

In many other countries, credit cards are a leading source of revolving credit to consumers. In Iceland, however, overdrafts on checking accounts are the most common form of revolving consumer debt. Virtually all checking accounts in Iceland offer an overdraft facility, the size of which is based on initial balances and credit history. This overdraft facility can be used at any time

without consulting the bank and overdraft status can be maintained indefinitely (subject to ad hoc reviews). Overdrafts dominate the unsecured consumer credit market, representing approximately 10% of all household loans during 2011-2017, and they charge average annual percentage rates (APRs) of around 12%.<sup>4</sup>

While overdraft facilities on checking accounts are the primary source of revolving consumer credit in Iceland, access to high-cost, short-term loans has grown substantially in recent years. Payday loans were first offered in Iceland in 2009. They require only a minimal credit assessment, are for short terms, and are available almost immediately after application in potentially substantial amounts. During the period covered by our sample, five payday lenders were active in Iceland. These firms operate only online or through short message services (SMS). To obtain a loan, individuals need to (i) affirm their legal competence to manage their own financial affairs, (ii) provide the Icelandic equivalent of Social Security Number, (iii) be formally registered as living in Iceland, (iv) supply an active email address/phone number and an active debit card number, and (v) not be undergoing debt mitigation. While they are called “payday loans,” obtaining this form of credit in Iceland does not require documentation of employment or the timing of paydays. Lending periods are flexible, individuals can choose durations between 1 and 90 days. Upon successful application, loans are deposited in the borrower’s bank account within a few minutes. The total borrowing limit of the five payday loan providers active in the sample period was approximately \$6,000.

Oversight of Iceland’s payday loan market is weak. For regulatory purposes, payday lenders are not classified as financial institutions, they do not need an operating license, they are all headquartered abroad, and government supervision of their activities is limited. Indeed, payday lending was effectively unregulated in Iceland prior to 2013.

In part due to the lack of government oversight, systematic evidence about the costs of payday loans in Iceland is limited. Kristjánsdóttir (2013) documents the costs of payday loans by all the Icelandic payday providers in 2013 and compares the costs of payday loans to those in other

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<sup>4</sup> Statistics, Central Bank of Iceland  
[www.sedlabanki.is/library/Fylgiskjol/Hagtolur/Fjarmalafyrirtaeki/2019/1013\20INN\\_Utlan\\_052019.xlsx](http://www.sedlabanki.is/library/Fylgiskjol/Hagtolur/Fjarmalafyrirtaeki/2019/1013\20INN_Utlan_052019.xlsx)

Nordic countries and the UK. This comparison shows that in 2013 the APR of payday loans was somewhat higher in Iceland than in the countries, with APRs starting at approximately 2,800%.

In November 2013, Iceland's Consumer Loans Act no. 33/2013, capped the APR on consumer debt at 50 percentage points above the Central Bank of Iceland's key interest rate. There is no evidence, however, that this regulation was binding on the costs of payday loans. Payday lenders appear to have circumvented or ignored the regulation. Some lenders skirted the law by, for example, having borrowers purchase e-books in exchange for expedited loan processing. Such fees are not included in the calculation of the APR. Others either ignored the law or interpreted their fees as exempt from it. To illustrate, Figure A1 in the Appendix shows an example of a payday loan contract and a screenshot from the homepage of one of the payday loan providers. These examples were collected by Iceland's Ministry of Tourism, Industry, and Innovation in 2018. The figure shows that the APR charged on a 30-day loan was 3,448.8%, very similar to the APRs documented by Kristjánsdóttir (2013) prior to the act. Consistent with the view that the regulation was not binding for payday lenders, we find no evidence in the Meniga data of a break in the number or size of loans around November 2013.

Information on the size of the payday lending market in Iceland is limited. To the best of our knowledge, ours is the first study to compare the use of payday loans to the use of other sources of consumer credit and relate it to other financial behavior in Iceland. Approximately 5.6% of the consumers in our data used payday loans at least once during a period of 6 years. Thus, as in other developed economies, payday borrowing is relatively uncommon; but the magnitude of borrowing among those who use payday loans users is substantial and seems likely to have an important influence on their financial circumstances.<sup>5</sup>

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<sup>5</sup> This is consistent with statistics from the Debtors' Ombudsman of Iceland for debt mitigation which shows that the share of people aged 18-29 who have applied for debt relief has increased sharply in recent years, and payday loans account for a much larger proportion of these troubled borrowers' total obligations. By 2017, 70% of debt mitigation applicants aged 18-29 owed payday loans. Among applicants who had payday debt, it accounted for about 20% of their total debt (Central Bank of Iceland, 2018).

## IV. Administrative Data

We use data from Iceland generated by Meniga, a financial aggregation software provider to European banks and financial institutions. Its account aggregation platform allows bank customers to view and manage all their bank accounts and credit cards across multiple banks in one place. Each day, the software automatically records all the bank and credit card transactions, including descriptions as well as balances, overdraft, and credit limits. Additionally, the data contain demographic information, such as age and gender.

Anyone who has an online bank account in Iceland can register at [meniga.is](http://meniga.is) to access the personal financial management platform. All who sign up agree to be a part of a sample for analytical purposes. In January 2017, the Icelandic population was 338,349 individuals, of whom 262,846 were older than 16. At the same time, Meniga had 50,573 users, which is about 20 percent of the population above age 16. Because their service is marketed through banks, the sample of users is fairly representative—see Table 4 in section V.

We restrict our analysis sample to users for whom we observe income and demographic information and whose expenditure data is credible.<sup>6</sup> In our analysis, we use four different types of information from the administrative data. First, we use the amounts and dates of payday loans. Second, we use the daily balances of checking accounts, savings accounts, and credit cards and on overdraft and credit card limits. Third, we use transaction-level information on income receipts, including the date of receipt and the income source, which we use to calculate monthly salary and monthly income. Finally, we use information on the number of non-sufficient funds charges each month. The different pieces of information are available for different periods. Data on payday loans are available from January 1, 2011 to January 31, 2017. Information on daily balances is available from September 1, 2014 to February 13, 2017. Income is available from January 6, 2011 to

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<sup>6</sup> The credibility of expenditure data depends on how well integrated a user is with the Meniga platform. When a user signs up, he agrees to import two years of transaction history into the Meniga database. If a user does not import all of his accounts in use, his financial activity cannot be accurately captured and we exclude such users. As a sampling criterion, we use minimum data activity captured by the following requirements: (1) The user must be active for at least 23 out of 24 months; (2) have been active for the past 3 months; and (3) have at least 5 transactions in food, (groceries or eat out). After applying these filters, comparison with the Statistics Iceland's consumption index and with visa credit card turnover indicates that the spending captured by the platform is close to those in other sources.

February 19, 2017 and non-sufficient funds charges from January 2011 to February 2017 (these are reported on a monthly basis).

After applying the filters, we have data for 12,747 Meniga users, of whom 717 have taken at least one payday loan during the 6 years of observations.

### *Preliminary Statistics*

Table 1 shows summary statistics of the Meniga sample regarding payday loans. All monetary figures shown in the paper are in hundreds of Icelandic króna (kr.). In 2017, 100 kr. corresponded approximately to 1 US dollar. Therefore, the reader can safely treat the monetary figures as US dollars. The mean and the median amount of loans are approximately \$244 and \$200. During the 6-year period of the data, payday loan borrowers took on average 18 loans. The median borrower took 8 loans, and borrowed \$1,800.

**Table 1: Summary Statistics of Payday Loans**

	Mean	Percentiles				
		10th	25th	50th	75th	90th
Amount Individual Loans	244	80	100	200	300	400
<i>Among Payday Loan Borrowers</i>						
Number of Payday Loans	18	1	2	8	24	47
Total Amount Borrowed	4,359	0.04	300	1,800	5,000	11,450

*Note:* This table shows summary statistics for 12,794 individual payday loans taken by 717 borrowers.

Table 2 compares payday loan borrowers to non-borrowers. Borrowers earn less and have less money in their checking and savings accounts. Some borrowers have relatively high incomes, however. The 90th percentile of the distribution of monthly income after taxes among borrowers is approximately \$5,000. The typical borrower has no money in her savings account and is

overdrafted by \$1,291. Borrowers also have lower credit card balances, which partly reflects that they have lower credit card limits (not shown in the table).

**Table 2: Summary Statistics of Income, Checking, Savings, and Credit Cards**

		Percentiles				
	Mean	10th	25th	50th	75th	90th
<i>Monthly Salary</i>						
Non-Borrowers	5,013	853	1,652	2,973	4,627	6,718
Borrowers	2,378	855	1,332	2,133	3,012	4,091
<i>Monthly Income</i>						
Non-Borrowers	5,931	1,295	2,219	3,635	5,386	7,654
Borrowers	3,074	1,315	1,919	2,799	3,872	5,092
<i>Checking Balance</i>						
Non-Borrowers	-226	-6,246	-1,662	187	1,233	3,618
Borrowers	-3,121	-8,867	-5,087	-1,291	29	334
<i>Savings Balance</i>						
Non-Borrowers	3,744	0	0	3	929	7,250
Borrowers	456	0	0	0	21	617
<i>Credit Card Balance</i>						
Non-Borrowers	1,531	0	168	1,147	2,267	3,540
Borrowers	748	0	0	0	971	2,262

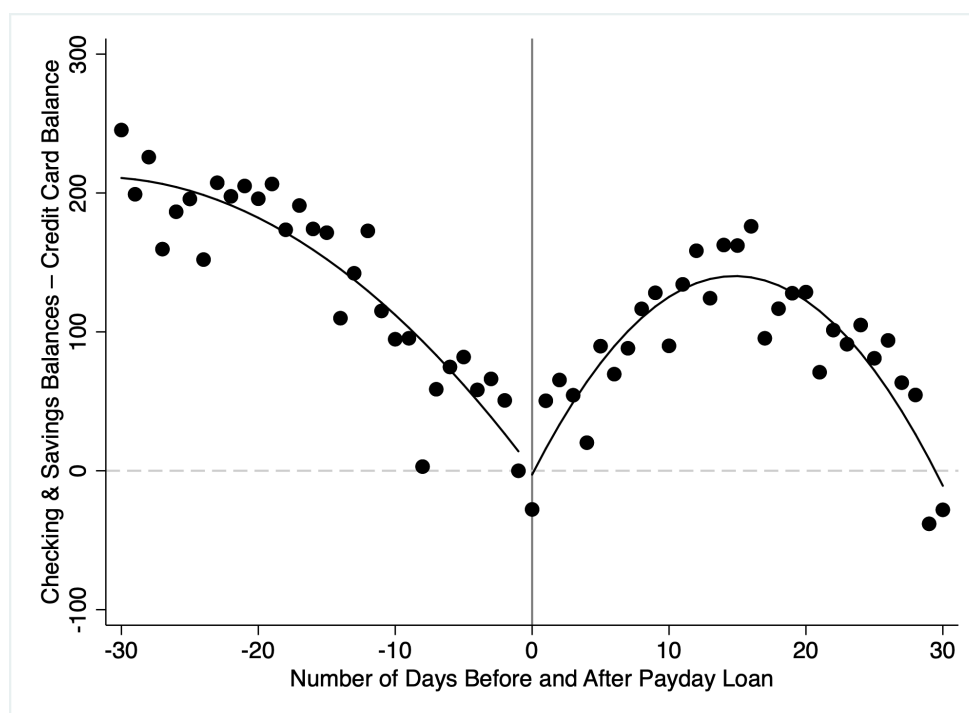
*Note:* This table shows summary statistics for payday loan borrowers (N = 596 for salary and income and N = 594 for balances) and for non-borrowers (N = 12,006 for salary and income and N = 11,074 for balances). Monthly salary and monthly income correspond to the individual's average monthly salary and average monthly income between February 2011 and January 2017. The balances correspond to the individual's median daily balances between September 1, 2014 and February 13, 2017.

The typical payday borrower has little liquidity, on average, and Figure 1 shows that he or she is especially illiquid in the days leading up to getting the loan. It shows the changes in the average amount borrowers had in their checking and savings account (minus how much they owed in credit card debt) before and after the loan was taken. This analysis is concerned with changes in liquidity, not levels. The level one day before the loan was taken (i.e., at -1) is therefore normalized to zero and, to reduce measurement error which disproportionately affects credit limits,

we assume here that overdraft and credit card limits are constant over the 60 days surrounding the day a payday loan is taken.

Figure 1 shows that liquidity gradually reduces by an average of about \$200 until the loan is taken. Liquidity then temporarily bounces back, but to a lower level than originally. After the recovery, liquidity falls again, such that 30 days after the loan liquidity returns to levels comparable to when the loan was taken.

**Figure 1: Patterns in Liquidity Around Payday Loans**



*Note:* This figure shows a time event study of liquidity – measured as the checking account balance plus the savings account balance minus the credit card balance – 30 days before and 30 days after a payday loan was taken. The analysis adjusts for when the interval between two loans is shorter than 30 days. It also includes day-of-the-week and calendar-day-of-the-month effects (see Appendix). Liquidity one day before a loan was taken was normalized to 0. The curves show pre-loan and post-loan quadratic trends. The analysis uses data on payday loans taken during October 1, 2014 and January 14, 2017 because the data on checking, savings, and credit card accounts is available for the period between September 1, 2014 and February 13, 2017. 3,453 payday loans were taken between October 1, 2014 and January 14, 2017 by 311 participants.

Focussing in on the day before the loan is taken, Table 3 shows that most payday borrowers have no access to cheaper liquidity on the day they take loans. The typical borrower can borrow

only \$32 via an overdraft, has no savings, and has virtually maxed out his credit card at the time he takes the loan. For these participants, a payday loan may be the only alternative available. There is, however, substantial heterogeneity. Some payday borrowers have cheaper alternatives. In particular, those in the 75<sup>th</sup> and 90<sup>th</sup> percentiles have about \$384 and \$1,276, respectively, available to them if they borrow via overdraft and tap into their savings.

**Table 3: Liquidity One Day Before Taking Payday Loan**

	Mean	Percentiles				
		10th	25th	50th	75th	90th
Checking Balance + Overdraft Limit (1)	273	0	0	32	190	745
Savings Balance (2)	466	0	0	0	1	531
Credit Card Limit – Credit Card Bal. (3)	541	0	0	7	352	1,750
(1) + (2)	740	0	2	58	384	1,276
(1) + (2) + (3)	1,280	0	28	244	1,149	3,323

*Note:* This table shows summary statistics of daily balances individuals had one day before they took payday loans. The number of payday loan borrowers is 322 and the number of loans is 3,672.

## V. Experimental Protocols

Analysis of the observational data, alone, suggests that “misfortune” is the major force driving demand for high-cost credit. Payday borrowers in Iceland tend to have lower income and low liquidity, on average. They also tend to be especially illiquid in the days just before receiving the loan and we find some evidence they are heading into a downward liquidity spiral. While some of this illiquidity might be the product of earlier “mistakes,” perhaps even an earlier decision to take a payday loan, a simple test for dominated choices suggests this is not the case. Only a relatively small fraction of borrowers have access to substantial amounts of cheaper liquidity and thus reveal a clear role for “mistake” in driving demand for payday loans.

The observational results are not, however, decisive. Even among those who have no access to cheaper market credit, the choice to take a payday loan may not be best. Indeed, many people with low income and liquidity choose not to borrow from payday lenders, and even those who turn to payday loans when they are especially illiquid do not always do so. To further examine what underlies the heterogeneity in decisions to take payday loans, and evaluate the role of

“mistakes,” we relate preferences and decision-making ability as measured in the experiments to demand for payday loans.

### *Recruitment & Survey Design*

Meniga sent a subset of its clients in Iceland an email with an invitation and a link to an online survey that we designed and programmed. We successfully delivered 8,913 e-mail invitations to users with complete records. Of those, 1,701 (19.8%) completed the survey. Compared with similar studies, this is a relatively high response rate. Epper et al. (2018), e.g., report 13% and Andersson et al. (2016) report 11%.

The survey contained three experimental tasks – a risk, an ambiguity, and an intertemporal choice task – and a brief questionnaire with questions about education, household composition, assets and debt. We discuss our sample and then the experimental tasks in detail.

### *Sample*

Table 4 compares the survey sample to a nationally representative sample. Statistics Iceland reports that in 2017 the average age among those above age 15 was 45.3 and that women constituted 50% of the population. The average age in the survey sample is 43.5 and the share of women is 47%. The share of singles in our survey is lower and the share of individuals living with a spouse and children is higher than in the overall population. Besides selection, this discrepancy may also be explained by the fact that individuals who live with a spouse (and possibly children), but are not registered as such, are counted by Statistics Iceland as living alone.

Table 4 also compares the education of our sample to the education of the Icelandic population. The largest difference is in the share of individuals who have only completed mandatory education. The difference may be partly explained by differences in measurement. Statistics Iceland receives information on graduates directly from the educational institutions. This means that degrees obtained abroad are not registered. Icelanders who get university degrees abroad, which is common, would be registered as having only completed mandatory education.

**Table 4: Comparison of Survey Sample to Iceland Population**

	Survey Participants	Icelandic Population
Female	47%	50%
Age	43.5	45.3
Labor Income	4,343	4,153
<i>Family Composition</i>		
Spouse	29%	28%
Single	23%	42%
Spouse and children	43%	25%
Single and children	6%	5%
<i>Highest Degree Obtained</i>		
Mandatory education	9%	39%
Journeyman's examination	4%	5%
Master of a certified trade	3%	6%
Matriculation examination	11%	8%
Tertiary education	8%	15%
Technical degree	5%	2%
Bachelor	30%	15%
Master	30%	7%
Ph.D.	2%	1%
Other	0%	3%

*Note:* This table compares survey participants to the general Icelandic population.

Appendix Table 9 compares the survey sample to Meniga users with complete records.

### *Experimental Tasks*

#### *Risk Task*

Participants allocated an experimental endowment of 500 kr. (appr. \$5) across 2 or 5 assets. The assets paid different amounts depending on whether a ball drawn from an urn was black or white. Participants were informed that the urn had 5 black balls and 5 white balls. Their decisions involved choosing how much to invest in each asset. Participants were presented with 15 investment problems (one of the 15 problems was randomly selected for payment). In the first 8 investment problems, there were 2 assets. In the last 7 investment problems, there were 5 assets. We varied the asset returns across the investment problems.

To illustrate, Online Appendix Figure 1 shows a screenshot of the interface for problems with two assets. The table at the top of the screen shows the returns of assets A and B per 1 kr. invested. The participant was then prompted to make her investment choices. The graph below the table displays two bars: the first bar shows the amount invested in asset A; the second bar shows the amount invested in asset B. Participants made their investments by either dragging the bars up and down or by clicking on the + and – buttons. The interface was such that participants always invested 100% of their experimental endowment. A similar interface was used in the investment problems with 5 assets (see Online Appendix Figure 2). The only distinction is that they were shown information about 5 assets – A, B, C, D, and E – and the graph displayed 5 bars.

Half of the participants were randomly assigned to be offered the option of avoiding the investment problem. In particular, these participants were offered the choice between making the investment decision or taking an outside option of –50 kr., 0 kr., or 100 kr. The amount of the outside option was varied across the investment problems. The participant was paid the outside option if in the problem selected for payment she chose to avoid. Appendix Table 1 shows the parameters of the 15 decision problems.

The interfaces for the participants with the outside option were slightly different. Online Appendix Figure 3 shows a screenshot of the interface they used. It differs from the interface used by other participants (Online Appendix Figure 1) in two ways. First, the graph with the bars is not shown. Second, the prompt to invest (“You will choose the amount you want to invest on each asset.”) is replaced by a prompt for the participant to choose between investing the experimental endowment (button “Invest Y kr.”) and taking the outside option (button “Receive X kr.”). If she clicked on the first button, the bars were unveiled and she could make her investment choices using the same interface used by other participants. If she clicked on the second button, she was presented with the next decision problem.

### *Ambiguity Task*

The ambiguity task was similar to the risk task with 3 distinctions. First, participants were informed that the urn now had 8 balls of one color and 2 balls of the other. However, they did not know whether the urn had 8 black balls and 2 white or if it had 2 black and 8 white. Second, in all

15 investment problems there were just 2 assets. Third, participants were not offered the option of avoiding the investment problem. Appendix Table 2 shows the parameters of the 15 investment problems. As in the risk task, 1 of the 15 problems was randomly selected for payment.

### *Intertemporal Choice Task*

Participants had to allocate their experimental endowment across a sooner date and a later date. The amount allocated to the later date accrued an experimental interest rate. Participants were presented with 12 intertemporal allocation problems (1 of the 12 problems was randomly selected for payment). We varied the experimental endowment, the experimental interest rate, and the sooner date across the problems. In the first 6 problems, the sooner date was today. In the last 6 problems, the sooner date was one year away. The time interval between the sooner and later dates was always one month. Within a time frame, the interest rate increased monotonically. Appendix Table 3 shows the parameters of the 12 intertemporal allocation problems.

Online Appendix Figure 4 shows a screenshot of the interface participants used in the intertemporal task. Two calendar sheets at the top of the screen show the sooner date (calendar sheet on the left) and the later date (calendar sheet on the right). The graph below the calendar sheets displays two bars: the bar on the left shows the amount to be received at the sooner date; the bar on the right shows the amount to be received at the later date (including the interest accrued).

### *Measuring Decision-making Ability*

Our main measure of decision-making ability is a composite measure that reflects the internal consistency of the participant's decision-making in the risk, ambiguity, and intertemporal choice tasks. We exploit the within-subject variation in asset returns (in the risk and ambiguity tasks) and in the endowment and interest rate (in the intertemporal choice task) to construct individual-specific measures of decision-making ability for each task. In the ambiguity and intertemporal choice tasks, we study whether choices violate the General Axiom of Revealed Preference (GARP).<sup>7</sup> In the risk task, we use different measures for those who had the option of

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<sup>7</sup> In the intertemporal choice task, we calculated CCEI separately using the choices for a given time frame and then took the minimum of the CCEI across the two time frames.

avoiding the investment problem and those who did not. We study whether the choices of the first group violate monotonicity with respect to first-order stochastic dominance (FOSD) and whether the choices of the second group violate GARP and FOSD (Polisson et al. 2019).

Choi et al. (2014) and Kariv and Silverman (2013) argue that consistency with GARP is a necessary condition for high quality decision-making. This view draws on Afriat (1967), which shows that if an individual's choices satisfy GARP in a setting like the one we study, then those choices can be rationalized by a well-behaved utility function. Consistency with GARP thus implies that the choices can be reconciled with a single, stable objective. We assess how nearly individual choice behavior complies with GARP using Afriat's (1972) Critical Cost Efficiency Index (CCEI). The CCEI is a number between zero and one, where one indicates perfect consistency with GARP. The degree to which the index falls below one may be viewed as a measure of the severity of the GARP violations.

Consistency with GARP may be seen as too low a standard of decision-making ability because it treats all stable objectives of choice as equally high-quality. A more stringent requirement would also require monotonicity of preferences. In particular, violations of monotonicity with respect to first-order stochastic dominance (FOSD) – that is, choices that yield payoff distributions with unambiguously lower payoffs than other available options – may be regarded as errors and provide a compelling criterion for decision-making quality.

We use the distribution of possible payoffs to assess how closely individual choice behavior complies with this dominance principle. To illustrate a violation of FOSD, consider a simplified case with two assets and no outside option. Asset 1 pays  $B$  if a black ball is drawn and 0 if a white ball is drawn. Asset 2 pays 0 if a black ball is drawn and  $W$  if a white ball is drawn. Let  $a$  be the amount invested on asset 1. The remaining  $500 - a$  are invested on asset 2. Investing  $a^{RF}$  on asset 1 and  $500 - a^{RF}$  on asset 2 is the risk-free allocation that pays the same amount irrespective of the color of the ball drawn, i.e.,  $a^{RF}B = (500 - a^{RF})W$ .

Suppose that asset 1 has a higher return than asset 2, i.e.,  $B > W$ , and that a participant chooses to invest less on asset 1 than the amount invested in the risk-free allocation, i.e.,  $a < a^{RF}$ . In this case, investing  $a' = 500 - aB/W$  on asset 1 yields an unambiguously higher payoff distribution than investing  $a$  on asset 1. First, notice that the minimum payout when investing  $a$

(black ball is drawn) is equal to the minimum payout when investing  $a'$  (white ball is drawn):  $aB$ . Second, the expected return of investing  $a'$ ,  $250 + a'(B - W)/2$ , is higher than the expected return of investing  $a$ ,  $250 + a(B - W)/2$ , because  $B > W$ .

Following Choi et al. (2014), we calculated a FOSD score as follows. If the selected investment portfolio was dominated as in the example above, the FOSD score was calculated as  $\frac{250W + a(B - W)/2}{250W + a'(B - W)/2}$ , which equals the expected return of the selected allocation as a fraction of the maximal expected return. The availability of the outside option introduces more opportunities for violating FOSD. First, if the participant invests  $a < a^{RF}$  on asset 1 and the outside option is greater than  $250 + a(B - W)/2$ , then investing  $a$  is dominated both by investing  $a' = 500 - aB/W$  and by the outside option, in which case we calculated the FOSD score as  $\frac{250W + a(B - W)/2}{\max\{250W + a'(B - W)/2, \text{outside option}\}}$ . Second, the participant violates FOSD by investing  $a > a^{RF}$  if the outside option is greater than  $250 + a(B - W)/2$ —in this case we calculated the FOSD score as  $\frac{250W + a(B - W)/2}{\text{outside option}}$ . Finally, one violates FOSD by taking the outside option if it is lower than the risk-free return,  $250 + a^{RF}(B - W)/2$ , in which case the FOSD was calculated as  $\frac{\text{outside option}}{250 + a^{RF}(B - W)/2}$ . The FOSD score was assigned a value of 1 if there was no FOSD violation.

We also calculate a unified measure of violations of GARP and violations of monotonicity with respect to FOSD, following Polisson et al. (2019). This revealed preference index, like the CCEI, lies between 0 and 1 where 1 represents perfect consistency with both GARP and monotonicity with respect to FOSD.

To reduce the influence of measurement error on estimates of the relationship between decision-making ability and demand for high-cost credit, we constructed a composite measure of decision-making ability derived from multiple decision tasks. We first calculated participants' percentile ranks in the distribution of decision-making ability in each task. For the risk task, we calculated separate percentile ranks for those participants who had the option of avoiding the investment problem and those who did not. For the first group, we calculated their percentile ranks in the distribution of the measure of FOSD violations. For the second group, we calculated their percentile ranks in the distribution of the unified measure of GARP and FOSD violations. Finally,

we constructed a decision-making ability index as the first component of a principal component analysis of the measures of decision-making ability in each one of the three tasks.

In Section VI, we assess the validity of this index by evaluating its ability to predict an unambiguous mistake revealed in the administrative data. That mistake, the accrual of insufficient fund fees which produce no benefit to the consumer, are strongly correlated with this principal component index of consistency with utility maximization.

As an alternative to this principal component approach to measurement error, we also adopt an instrumental variables approach (Cf. Gillen et al., 2019). That approach treats decision-making ability derived from one task as an instrument, in a two-stage least squares framework, for decision-making ability derived in another task. See Appendix Table 6 for details. These two approaches to measurement error provide qualitatively similar results.

### *Measuring Time and Risk Preferences*

Let  $s_{i,r}^0$  be the fraction of the endowment allocated by participant  $i$  in the intertemporal choice task to the sooner date when the sooner date is today and the interest rate is  $r$  and  $s_{i,r}^1$  be the fraction allocated by  $i$  to sooner when the interest rate is  $r$  and the sooner date is one year away. We measured impatience of  $i$  as the average of  $s_{i,r}^1$  across the 5 different positive interest rates. Define  $\Delta_{i,r} \equiv s_{i,r}^0 - s_{i,r}^1$ . We measured the present bias of  $i$  as the average of  $\Delta_{i,r}$  across all 6 interest rates. Participant  $i$  was classified as present-biased if this average was positive, time consistent if it was zero, and future-biased if it was negative.

The measure of small-stakes risk aversion depends on whether the participants had the outside option. Returning to our example above, suppose that  $B_j > W_j$  in choice  $j$ . For a participant  $i$  without the outside option, let  $\phi_{i,j} = \min \left\{ 0.5, \frac{(500 - a_{i,j})W_j}{a_{i,j}B_j + (500 - a_{i,j})W_j} \right\}$ , where  $a_{i,j}$  is the amount invested by participant  $i$  on asset 1 in choice  $j$ . The second term on the right-hand side is the payout in the low-return state of the world as a fraction of the sum of the payouts in both states. A risk neutral participant would invest  $a_{i,j} = 500$ , such that  $\phi_{i,j}$  would be equal to zero. A participant with infinite risk aversion would invest  $a_{i,j}^{RF} = 500 \frac{W_j}{B_j + W_j}$ , such that  $\phi_{i,j}$  would be equal to 0.5. For

a participant  $i$  with the outside option, let  $\varphi_{i,j} = \min\left\{0.5, \frac{(500-a_{i,j})W_j}{a_{i,j}B_j+(500-a_{i,j})W_j}\right\}$  if he chose in choice  $j$  to invest and as  $\varphi_{i,j} = 0.5$  if he chose to avoid the investment decision. We then averaged  $\phi_{i,j}$  or  $\varphi_{i,j}$  over the 8 investment problems with the overall lowest avoidance rates, i.e.,  $\phi_i = \frac{1}{8}\sum_{j=1}^8 \phi_{i,j}$  and  $\varphi_i = \frac{1}{8}\sum_{j=1}^8 \varphi_{i,j}$ . Finally, we calculated separate percentile ranks for participants with and without the outside option – percentile ranks in the distribution of  $\phi_i$  for those without the outside option and ranks in the distribution of  $\varphi_i$  for those with the outside option.

In Appendix Table 10, we assess the validity of these measures of economic preferences by evaluating whether we can reproduce associations documented in previous work. The table shows that they predict the relevant outcomes in expected ways: impatience predicts wealth (as in Epper et al. 2018); present bias predicts consumer debt (as in Meier & Sprenger 2010); and risk aversion predicts stock market participation (as in Barsky et al. 1997).<sup>8</sup>

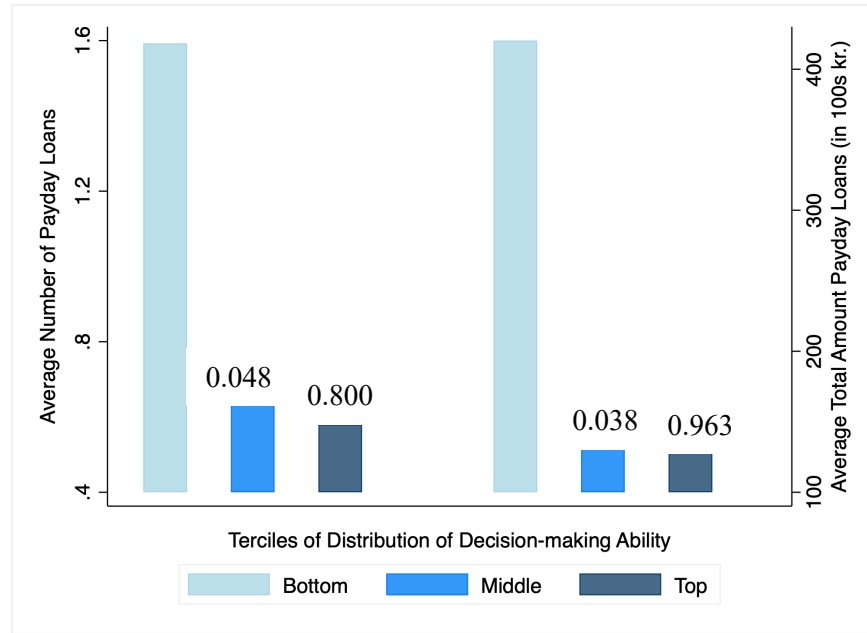
## VI. Experimental Results

We begin with preliminary evidence that individuals who exhibit lower decision-making ability in the experimental tasks make greater use of payday loans. Figure 2 shows averages of the number of payday loans (left y-axis) and of the total amount borrowed (right y-axis), separately by terciles of the distribution of the decision-making ability. The number above a bar reports the p-value of a test of differences in means between the bar and the bar to its left. For example, the 0.048 above the second bar is the p-value of a test of the difference between the middle and bottom terciles of the decision-making ability distribution in the number of payday loans. Individuals in the bottom tercile of the distribution of the decision-making ability index have on average approximately 1 payday loan more than individuals in the top tercile of the distribution of the decision-making ability. They borrowed on average 3 times more.

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<sup>8</sup> Data on wealth and stock market participation come from the survey. Participants reported the value of different types of assets, including stocks. Information about overdraft balances come from the administrative data.

**Figure 2: Payday Loans and Decision-Making Ability**



*Note:* This figure shows the use of payday loan services by decision-making ability. The three bars on the left show the average number of payday loans per individual for individuals in the bottom, middle, and top terciles of the decision-making ability distribution. The three bars on the right show the average amount per individual of all payday loans for individuals in the bottom, middle, and top terciles of the decision-making ability distribution. *Number of participants* = 567 in each tercile for a total of 1,701. The number above a bar reports the p-value of a test of the difference between the bar and the bar to its left.

**Table 5: Cumulative Share of Total Amount Borrowed by Percentile of Decision-Making Ability Distribution**

<i>Percentile of Decision-Making Quality Distribution</i>								
10th	20th	30th	40th	50th	60th	70th	80th	90th
28%	53%	56%	62%	69%	78%	81%	90%	99%

*Note:* This table shows the share of the total amount of payday loans borrowed by individuals in the bottom X<sup>th</sup> percentile of the decision-making ability distribution as a fraction of the total amount of all payday loans taken by survey participants. For example, together the payday loan borrowers borrowed a total of \$ 387,832. Those individuals in the bottom 20<sup>th</sup> percentile of the decision-making ability distribution borrowed collectively a total of \$ 206,220. *Number of participants* = 1,701.

Payday loans are rare in the population, so the average level differences in borrowing by quantiles of the decision-making ability distribution may understate the importance of those with

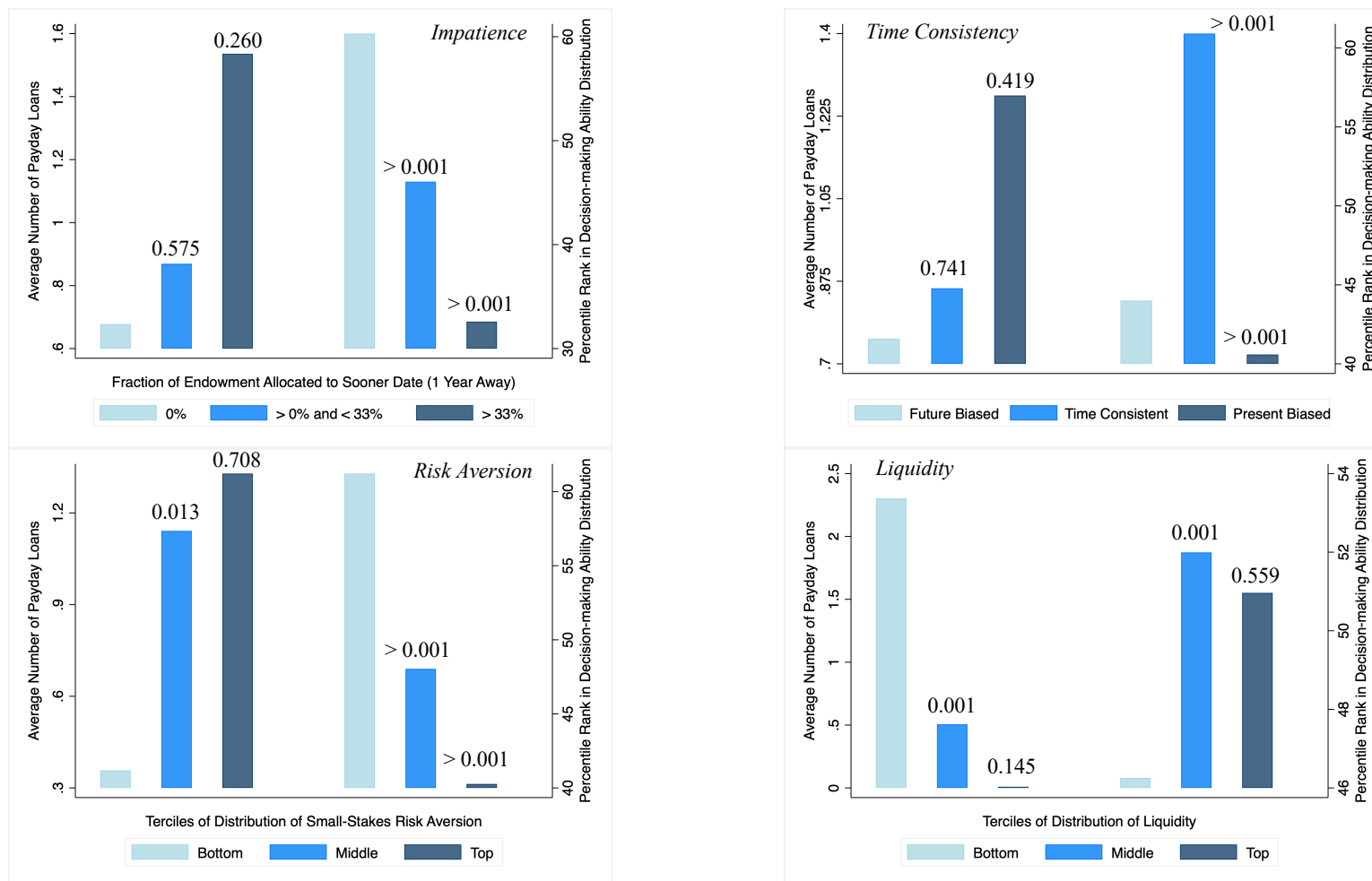
low ability in the payday loan market. Indeed, lower decision-making ability people appear to play an outsized role in this market. Table 5 shows the share of the total amount borrowed by percentile of the decision-making ability distribution. Those in the bottom 10% borrowed 28% of the total. The bottom 20% of the decision-making ability distribution borrowed more than half of the total amount borrowed. In contrast, those at the top 10% borrowed 1% of the total amount.

The strong association between payday loans and decision-making ability may partly reflect individual differences in preferences or liquidity. Figure 3 documents the association of payday loans and of decision-making ability with these potential confounders. The panels show averages of the number of payday loans (left y-axis) and of the percentile rank in the distribution of decision-making ability (right y-axis), separately by impatience, present bias, small-stakes risk aversion, and by liquidity. A participant's liquidity is the median, across all days, of the daily sum of savings and checking account balances, overdraft limit, and credit card limit minus balance.

The relationships between payday loan demand, preferences, and liquidity all go in the expected direction. Individuals who are more impatient, more present-biased, more risk averse, or have lower liquidity take on average more payday loans. The relationships between decision-making ability and impatience and small-scale risk aversion are monotonic: the more impatient and more risk averse exhibit lower decision-making ability. We also find that the present-biased exhibit substantially lower decision-making ability than the time-consistent. Those in the bottom tercile of the distribution of liquidity have lower decision-making ability than those in the middle and top terciles.

The results in Figure 3 suggest that the relationship between decision-making ability and payday loans may be confounded by both economic preferences and liquidity. In Table 6, we use regression analysis to estimate the relationship between payday loan borrowing and decision-making ability conditioning on these potential confounders. The dependent variable is the number of payday loans. The independent variables shown in the first 5 rows – decision-making ability, liquidity, impatience, present bias, and risk aversion – are measured in percentile ranks divided by 10, such that the coefficients can be interpreted as the effects of increasing these variables 10 percentiles. The liquidity measure here is the median of the daily sum of checking and savings

**Figure 3: Association of Payday Loans and of Decision-making Ability with Economic Preferences and with Liquidity**



*Note:* This figure investigates the relationship between number of payday loans and the percentile rank in the distribution of decision-making ability, on one hand, and impatience, time consistency, risk aversion, and liquidity, on the other. The left y-axis in the figures shows the average number of payday loans. The right y-axis shows the average percentile rank in the distribution of decision-making ability. The top left figure shows separate numbers for those who allocated 0% to the sooner date ( $N = 863$ ); those who allocated more than 0% and less than 33% ( $N = 405$ ); and those who allocated more than 33% ( $N = 433$ ). The top right figure shows separate numbers for those who exhibited future-biased behavior ( $N = 535$ ), time consistency ( $N = 686$ ), and present-biased behavior ( $N = 480$ ) – see section for description of how we constructed these groups. The bottom figures show numbers for those in the bottom ( $N = 568$  and  $526$ ), middle ( $N = 567$  and  $525$ ), and top ( $N = 566$  and  $525$ ) terciles of the distribution of risk aversion and of the distribution of liquidity respectively. The number above a bar reports the p-value of a test of the difference between the bar and the bar to its left.

balances plus overdraft and credit card limits minus the credit card balance. All regressions include controls for the log of average monthly income, years of schooling, gender, age, and age squared.<sup>9</sup>

The results in Table 6 indicate that both “misfortune” and “mistake” are important in determining payday loan borrowing. Individuals in worse financial circumstances and with lower decision-making ability take more payday loans. The relationship between payday loans and decision-making ability and the relationship between payday loans and liquidity are robust to controlling for demographics, education, income, and time and risk preferences. In the first three specifications, decision-making ability is statistically significant at the 5% confidence level (the p-value in the last specification is 0.052). Liquidity is always significant at the 1% level. Improving decision-making ability in 10 percentiles reduces the number of payday loans by 0.12-0.21 loans depending on the specification. Increasing liquidity in 10 percentiles reduces the number of payday loans by 0.47-0.49 loans. These estimates are not small given that the average number of payday loans is 0.94 and that as shown in Table 6 these loans are concentrated among the individuals with lower decision-making ability.<sup>10</sup>

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<sup>9</sup> In Appendix Tables 4 and 5, we present the results of alternative specifications that allow for non-linear effects of liquidity or alternative measures of the potentially confounding variables. The point estimate of the relationship between decision-making ability and loan demand is stable across specifications.

<sup>10</sup> Applying an instrumental variables (IV) approach to measurement error in decision-making ability (Gillen et al., 2019), suggests these estimates may be understating the magnitude of the relationship between decision-making ability and payday loans. The IV approach treats decision-making ability derived from one task as an instrument, in a two-stage least squares framework, for decision-making ability derived in another task. The results of Appendix Table 6 show that the IV point estimate of the relationship between decision-making ability and the number of loans is more than twice the OLS estimate.

Impatient, present-biased, and risk averse individuals take more payday loans, but these point estimates are relatively imprecise. Income and gender also have substantial, independent relationships with demand for payday loans. Women take, on average, one less loan than men and the point estimate indicates a 10% increase in average income is associated with a 0.07 increase in the number of payday loans received. The counterintuitive, positive relationship with income derives from conditioning on liquidity. The coefficient on income is not statistically distinguishable from zero with conventional levels of confidence when we do not condition on measures of liquidity.

**Table 6: Independent Effects of Liquidity and Decision-making Ability on Payday Loans**

	<i>Number of Payday Loans</i>				
DMA	-0.21 (0.08)		-0.16 (0.08)	-0.15 (0.07)	-0.12 (0.06)
Liquidity		-0.49 (0.09)	-0.48 (0.09)	-0.47 (0.09)	-0.47 (0.08)
Impatience				0.02 (0.04)	0.02 (0.04)
Present Bias				0.06 (0.07)	0.06 (0.07)
Risk Aversion					0.07 (0.06)
Log Income	0.06 (0.15)	0.67 (0.20)	0.69 (0.21)	0.69 (0.21)	0.70 (0.21)
Years of Schooling	-0.08 (0.05)	-0.03 (0.05)	-0.02 (0.05)	-0.02 (0.05)	-0.02 (0.05)
Female	-0.73 (0.35)	-0.82 (0.35)	-0.88 (0.36)	-0.91 (0.38)	-0.94 (0.38)
Age	-1.53E-02 (0.02)	1.25E-02 (0.02)	2.70E-03 (0.02)	1.60E-03 (0.02)	4.26E-04 (0.02)
Age <sup>2</sup>	-9.29E-04 (6.69E-04)	-2.98E-04 (6.82E-04)	-4.38E-04 (6.63E-04)	-4.90E-04 (6.56E-04)	-4.90E-04 (6.56E-04)

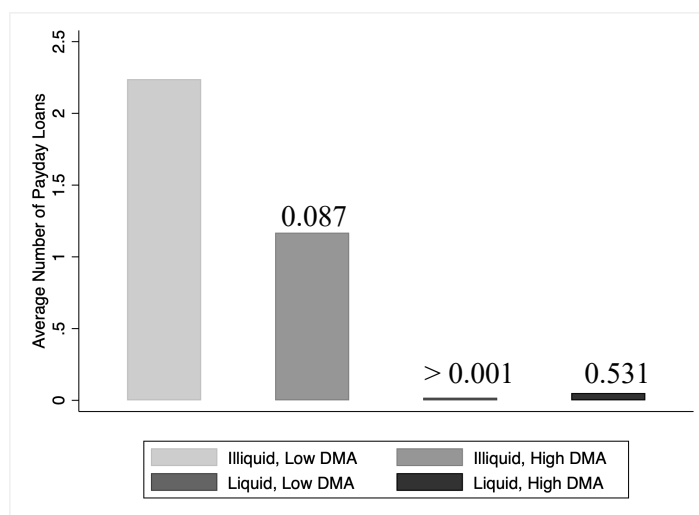
Notes: This table investigates the relationship between payday loan borrowing, decision-making ability, and liquidity. The mean of the dependent variable is 0.94. Decision-making ability, liquidity, time preferences and risk preferences are measured in percentile ranks divided by 10. *Number of observations* = 1,573.

Appendix Table 5 shows that the estimated relationship between decision-making ability and payday loans is robust to using alternative measures of demographics, education, liquidity, income, time and risk preferences. Appendix Table 6 in turn shows that the relationship is also robust to controlling for liquidity more flexibly.

#### *Interactions Between “Misfortune” and “Mistake”*

The results in Table 6 assume that liquidity and decision-making ability have separable effects on payday loan borrowing, but it is plausible that the influence of one is affected by the level of the other. Figure 4 provides preliminary evidence that this is the case. It divides the sample roughly into quarters by high and low liquidity and by high and low decision-making ability. It then displays the average number of payday loans for each quarter of the sample.

**Figure 4: Average Number of Payday Loans by Liquidity × Decision-Making Ability**



*Note:* This figure shows the average number of payday loans for four different groups: 1) those in the bottom half of the liquidity distribution and the bottom half of the decision-making ability distribution (“Illiquid, Low DMA”); 2) those in the bottom half of the liquidity distribution and the top half of the decision-making ability distribution (“Illiquid, High DMA”); 3) those in the top half of the liquidity distribution and the bottom half of the decision-making ability distribution (“Liquid, Low DMA”); and 4) those in the top half of the liquidity distribution and the top half of the decision-making ability distribution (“Liquid, High DMA”). The number of participants in each group is respectively 474, 439, 377, and 411 for a total of 1,701 participants. The number above a bar reports the p-value of a test of the difference between the bar and the bar to its left.

Figure 4 shows that – regardless of decision-making ability – those in the top half of the liquidity distribution virtually never take payday loans. Among the bottom half of the liquidity

distribution, however, those with lower decision-making ability take three times as many loans as those with higher decision-making ability. Table 7 further investigates these results in a regression framework that controls for demographics, income, and economic preferences.

In particular, Table 7 presents the results of a regression of the number of payday loans on liquidity, decision-making ability, and the interaction of the two (both are demeaned). The effect of an increase in ability of 10 percentiles is equal to the coefficient on the interaction term times liquidity plus the coefficient on decision-making ability. Similarly, the effect of an increase in liquidity of 10 percentiles is equal to the coefficient on the interaction term times decision-making ability plus the coefficient on liquidity. To illustrate, if an individual is at the 60<sup>th</sup> percentile of the decision-making ability distribution, the effect of a reduction in liquidity of 10 percentiles is equal to the coefficient on the interaction term minus the coefficient on liquidity.

The first column of Table 7 reproduces the middle column of Table 6 for comparison. In the second column, we add the interaction term. The coefficients on decision-making ability and on liquidity barely change. Time preferences are included in the third column while risk preferences are included in the fourth column. These results confirm that higher decision-making ability protects against the negative effects of illiquidity. The coefficient on liquidity in the fourth column is  $-0.46$ . The coefficient on the interaction term, which is statistically significant at the 1% level, is 0.06. This implies that a 10 percent reduction in liquidity increases the number of payday loans by 0.71 for someone in the 10<sup>th</sup> percentile of the distribution of decision-making ability, by 0.46 for someone with median decision-making ability, and by 0.21 for someone in the 90<sup>th</sup> percentile. Similarly, the effect of lower decision-making ability is decreasing in liquidity. A 10 percent reduction in decision-making ability increases the number of payday loans by 0.38 for someone in the 10<sup>th</sup> percentile of the distribution of liquidity and has no effect on the number of payday loans of someone in the 70<sup>th</sup> percentile.

**Table 7: Interactive Effects of Liquidity and Decision-making Ability on Payday Loans**

		<i>Number of Payday Loans</i>			
DMA * Liquidity		0.06 (0.02)	0.06 (0.02)	0.06 (0.02)	
	DMA	-0.16 (0.08)	-0.16 (0.08)	-0.15 (0.07)	-0.12 (0.06)
	Liquidity	-0.48 (0.09)	-0.47 (0.09)	-0.46 (0.08)	-0.46 (0.08)
	Impatience			0.02 (0.04)	0.02 (0.04)
	Present Bias			0.06 (0.07)	0.07 (0.07)
	Risk Aversion				0.07 (0.05)
	Log Income	0.69 (0.21)	0.72 (0.21)	0.73 (0.21)	0.73 (0.21)
	Years of Schooling	-0.02 (0.05)	-0.02 (0.05)	-0.02 (0.05)	-0.02 (0.05)
	Female	-0.88 (0.36)	-0.82 (0.36)	-0.85 (0.37)	-0.88 (0.38)
	Age	2.70E-03 (0.02)	3.13E-03 (0.02)	2.12E-03 (0.02)	9.06E-04 (0.02)
	Age <sup>2</sup>	-4.38E-04 (6.63E-04)	-3.45E-04 (6.69E-04)	-3.98E-04 (6.62E-04)	-3.98E-04 (6.62E-04)

Notes: This table investigates the relationship between payday loan borrowing, decision-making ability, and liquidity. The mean of the dependent variable is 0.94. Decision-making ability, liquidity, time preferences and risk preferences are measured in percentile ranks divided by 10. *Number of observations* = 1,573.

### *Decision-making Ability and High-Frequency Variation in Liquidity*

The prior results indicate decision-making ability plays a meaningful role in determining demand for payday loans, especially for those with low average liquidity. These results may, however, overstate the relative importance of decision-making ability and, by implication, “mistakes” because they account only for an individual’s average financial circumstances over a relatively long period. While decision-making ability may be quite stable over time, liquidity often

is not and averaging over the sample period may gloss over the key liquidity events that drive high-cost credit demand.

To investigate this possibility, we estimate analogous relationships between financial circumstances and high-cost loan demand at the daily level, conditional on demographics, decision-making ability and preferences. Table 8 presents the results, where the unit of observation is now the individual-day, the dependent variable is an indicator for whether the individual received a payday loan that day, and liquidity is measured on the day before the loan was received. Standard errors on the point estimates are clustered at the level of the individual.

In specification (1) of Table 8 we find, as in the low-frequency specifications, a negative relationship between decision-making ability and payday loan demand, conditional on average income, education, and demographics. Given the low probability of taking a loan on any given day, the magnitude of the point estimate is correspondingly smaller, but is again statistically distinguishable from zero with high confidence. In specification (2), we also condition on liquidity levels the day before, and find, as expected, a significant negative relationship. Importantly, however, adding this daily measure of the level liquidity has no meaningful impact on the point estimate of the relationship between decision-making ability and payday loan demand. As in the low-frequency specifications, adding controls for preferences in specifications (3) and (4) alters the estimated relationship between decision-making ability and payday loan demand only modestly. To account for the frequency of zeros and outliers in the liquidity distribution, specification (5) replaces the level measure of liquidity with its inverse hyperbolic sign and the point estimate on decision-making ability is little changed.

**Table 8: Decision-making Ability and High-Frequency Variation in Liquidity**

	1 if Took a Payday Loan					
	(1)	(2)	(3)	(4)	(5)	(6)
DMA	-1.03 (0.46)	-1.01 (0.46)	-0.90 (0.42)	-0.85 (0.43)	-0.78 (0.42)	-0.86 (0.43)
Liquidity in \$10,000s		-0.78 (0.25)	-0.76 (0.24)	-0.77 (0.24)		
IHS of Liquidity					-2.49 (0.66)	
Pctile Rank of Liquidity						-0.37 (0.18)
Impatience			0.08 (0.23)	0.08 (0.23)	0.03 (0.23)	0.11 (0.23)
Present Bias			0.53 (0.37)	0.54 (0.37)	0.48 (0.36)	0.55 (0.37)
Risk Aversion				0.14 (0.26)	0.10 (0.26)	0.13 (0.26)
Log Income	0.42 (0.84)	1.02 (0.90)	1.08 (0.91)	1.10 (0.91)	2.58 (1.07)	0.62 (0.87)
Years of Schooling	-0.29 (0.20)	-0.27 (0.20)	-0.28 (0.20)	-0.27 (0.20)	-0.09 (0.18)	-0.30 (0.20)
Female	-3.25 (1.91)	-3.40 (1.94)	-3.66 (2.02)	-3.72 (2.01)	-3.62 (1.97)	-3.57 (1.99)
Age	-0.05 (0.10)	-0.04 (0.10)	-0.05 (0.10)	-0.05 (0.10)	-1.19E-03 (0.09)	-0.06 (0.10)
Age <sup>2</sup>	-5.70E-03 (4.32E-03)	-4.57E-03 (4.23E-03)	-4.98E-03 (4.18E-03)	-4.97E-03 (4.18E-03)	-4.79E-03 (4.18E-03)	-6.02E-03 (4.29E-03)

*Note:* This table controls for more flexible forms of liquidity. It shows results from regressions at the individual-daily level. The dependent variable is an indicator for whether participant  $i$  took a payday loan on day  $d$ . We multiplied it by 10,000 so the coefficients can be interpreted as the effect on a hundredth of a percentage point. Its mean is 3.79. Liquidity refers to the liquidity on the previous day, i.e.,  $d - 1$ . Columns (2)-(4) include liquidity in levels as a control. Column (5) controls for the inverse hyperbolic sine of liquidity. Column (6) adds a *within*-participant percentile rank measure of liquidity. In particular, the liquidity of participant  $i$  on day  $d - 1$  was ranked relative to the liquidity of participant  $i$  in all other days in the individual time series of the participant. Decision-making ability, time and risk preferences are measured in percentile ranks divided by 10, such that the coefficient gives the effect of an increase of the independent variable in 10 percentiles. The regressions include dummies for day of the week and for calendar day of the month. Number of observations = 1,388,959. Number of participants = 1,573. Number of days = 883. Standard errors clustered at the individual level.

Finally, specification (6) evaluates the possibility that the circumstances which represent a liquidity “crisis” depend on an individual’s typical liquidity. While each specification has, so far, conditioned on measures of average income and education, the situations that trigger an individual’s demand for a payday loan may depend on the extent to which liquidity has fallen below its usual levels. In this last specification, therefore, we replace the daily liquidity level with its within-individual percentile rank. The results show, that the relative level of financial circumstances is a significant predictor of payday loan demand, but conditioning on it has little influence on the estimated relationship between decision-making ability and the likelihood of taking a payday loan.<sup>11</sup>

### *Scarcity Effects and Reverse Causation*

In the preceding analysis, we have treated the association between decision-making ability and liquidity as potentially confounding estimates of the relationship between decision-making ability and the demand for payday loans. Theories of scarcity (Mullainathan and Shafir, 2013; Carvalho et al. 2016) suggest, however, the potential for reverse causation. In this view, the many challenges associated with a lack of financial resources may impede cognitive function and degrade decision-making ability.

The scarcity view of the relationship between financial circumstances and consistency with utility maximization accords with our interpretation of the latter as decision-making ability. Nevertheless, if financial scarcity to some extent causes lower decision-making ability, then the regression analyses above will attribute to decision-making ability some of the gross effects of misfortune on payday loan demand.

To evaluate the potential for scarcity effects, Appendix Figure 5 presents estimates of the relationship between within-participant variation in liquidity and decision-making ability. More

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<sup>11</sup> Appendix Table 7 presents results that allow for interactive effects of liquidity and decision-making ability at the daily level. Results are qualitatively similar to those in the low-frequency specification of Table 7. By estimating the relationship between payday loan demand and liquidity measured at both individual average and individual daily frequencies, we assess the role of both highly persistent and immediate financial circumstances in the decision to take a high cost loan. In Appendix Table 8, we evaluate a role for intermediate financial circumstances by repeating the analysis in Table 8, but at a monthly frequency. The results are qualitatively similar to those at the daily frequency.

specifically, for each calendar day of the month we calculated the within-participant median liquidity across months. For each participant, calendar days were then sorted from lowest median liquidity to highest median liquidity. We then estimated via OLS a linear relationship between the rank of the day in which the participant took the survey and the average decision-making ability rank of participants surveyed on that day. This analysis provides no evidence of scarcity effects. Those surveyed on especially low liquidity days have, if anything, slightly higher average decision-making ability.

### *Decision-making Ability and NSF Fees*

The preceding results are consistent with the view that “mistakes” are quantitatively important drivers of demand for payday loans. This interpretation would be misguided, however, if decision-making ability were simply capturing a “type” whose unmeasured constraints, preferences, or beliefs rationalize demand for high-cost loans. It may be, for example, that consistency with utility maximization in the experiment is correlated with unmeasured access to friends or family on whom to rely when misfortune strikes. If so, we would improperly attribute to “mistake” what is actually a lack access to informal credit.

To evaluate this and related possibilities, we study the relationship between measures of decision-making ability from the experiment and an unambiguous “mistake” in the administrative data: the accrual of non-sufficient funds (NSF) fees. NSF fees obtain when, in the process of using a debit card to make a purchase, an individual exceeds his or her checking account overdraft limit.<sup>12</sup> Note this may occur even to an individual who has liquidity in the form of another checking account, a savings account, or a credit card. Different from costly overdrafts in markets like the U.S., there is no benefit to exceeding the limit because the purchase will not be authorized. A choice that results in an NSF fee is dominated by the decision not to try to make the purchase. Like an American looking left (but not right) before crossing the street in the United Kingdom, incurring an NSF fee may be understandable, but would almost universally be viewed as a “mistake.”

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<sup>12</sup> The conditional mean and conditional median of monthly NSF charges were \$15.48 and \$9.05. Each month, about 1.8% of participants accrued an NSF fee. More than a quarter of them accrued an NSF fee at least once during the 74 months. Among those, they accrued NSF fees in, on average, in 9.3 months for an average total cost of \$79.47.

The relationship between decision-making ability and NSF fees can thus provide evidence on the validity of using measures of consistency with utility maximization in the experiments as measures of decision-making ability. It can, in particular, inform the hypothesis that it is a correlation with unmeasured constraints, preferences, or beliefs, rather than with decision-making ability, that drives the estimated relationship between consistency with utility maximization and demand for payday loans. If consistency with utility maximization in the experiments largely proxies for unmeasured constraints, preferences, or beliefs that rationalize payday loans, it should not predict the unambiguous “mistake” of NSF fees.

**Table 9: Non-Sufficient Funds Charges and Decision-making Ability**

		<i>Number of Non-Sufficient Funds Charges</i>				
DMA * Liquidity					0.04 (0.03)	0.04 (0.03)
DMA	-0.26 (0.09)		-0.22 (0.08)	-0.24 (0.09)	-0.22 (0.08)	-0.24 (0.09)
Liquidity		-0.57 (0.09)	-0.55 (0.09)	-0.56 (0.09)	-0.55 (0.09)	-0.55 (0.09)
Impatience				0.01 (0.05)		0.01 (0.05)
Present Bias				-0.08 (0.08)		-0.08 (0.08)
Risk Aversion				-0.06 (0.06)		-0.06 (0.06)
Log Income	0.25 (0.30)	1.04 (0.32)	1.08 (0.33)	1.07 (0.33)	1.11 (0.33)	1.10 (0.33)
Years of Schooling	-0.09 (0.07)	-0.04 (0.07)	-0.02 (0.07)	-0.02 (0.07)	-0.03 (0.07)	-0.02 (0.07)
Female	-0.01 (0.41)	-0.08 (0.40)	-0.16 (0.41)	-0.11 (0.40)	-0.13 (0.40)	-0.08 (0.39)
Age	0.05 (0.02)	0.08 (0.02)	0.07 (0.02)	0.07 (0.02)	0.07 (0.02)	0.07 (0.02)
Age <sup>2</sup>	-4.74E-03 (9.60E-04)	-3.98E-03 (9.24E-04)	-4.14E-03 (9.25E-04)	-4.10E-03 (9.37E-04)	-4.08E-03 (9.28E-04)	-4.03E-03 (9.40E-04)

Notes: This table investigates the relationship between non-sufficient funds (NSF) charges, decision-making ability, and liquidity. The mean of the dependent variable is 2.43. Decision-making ability, liquidity, time preferences and risk preferences are measured in percentile ranks divided by 10. *Number of observations* = 1,542.

Table 9 shows that the measure of decision-making ability is indeed predictive of this unambiguous “mistake.” A reduction of 10 percentiles in decision-making ability increases the number of NSF charges by 0.22-0.26 (relative to a base rate of 2.43). Liquidity is also associated with NSF charges: A reduction of 10 percentiles in liquidity increases the number of NSF charges by 0.55-0.57. Interestingly, the interaction between decision-making ability and liquidity is not statistically distinguishable from zero at conventional levels of confidence, which is consistent with the interpretation that NSF charges are simple mistakes. Even among those with liquidity, individuals with lower decision-making ability are more likely to engage in this imperfection.

## **VII. External Relevance: Results from US Survey Data**

The combination of administrative and survey data from Iceland has advantages for studying the relationships between economic circumstances, decision-making ability, and demand for high-cost credit. The administrative data offer high-frequency, accurate measures of economic circumstances and demand for a relatively large, long, and balanced panel. The survey data provide rich measures of preferences and decision-making ability derived from multiple decision domains. One potential concern, though, is that Iceland is a small economy and its people and markets may have distinctive characteristics that limit the external relevance of findings derived from them.

To assess external relevance, we turn to survey data from the U.S. and compare, to the extent possible, the relationships between economic circumstances, decision-making ability, and demand for high-cost credit in those data with the analogous evidence from Iceland. The U.S. data are drawn from the Understanding America Study (UAS), an Internet panel with respondents aged 18 and older living in the U.S.<sup>13</sup> About twice a month, respondents receive an email with a request to visit the UAS site and complete questionnaires. Regular questionnaires collect self-reported economic and demographic information. Two supplementary UAS questionnaires first fielded in 2015 have asked respondents whether they have a payday loan or have had one in the past year. A third survey administered Choi et al.’s (2014) choice under risk experiment. See Appendix for more details.

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<sup>13</sup> Respondents are recruited by address-based sampling. Those without Internet access at the time of recruitment are provided tablets and Internet access.

Combining responses from these three supplements with information from regular UAS questionnaires, we can estimate the relationship between self-reported economic circumstances, preferences and decision-making ability as revealed in the choice under risk experiment,<sup>14</sup> and self-reported demand for this kind of high-cost credit. Table 10 presents the results from the UAS alongside analogous estimates from the Icelandic data.

**Table 10: Decision-making Ability and Payday Loan Demand in Iceland and U.S.**

	<i>1 if Had Payday Loan</i>							
	Iceland	UAS	Iceland	UAS	Iceland	UAS	Iceland	UAS
DMA	-0.41 (0.19)	-0.53 (0.11)	-0.59 (0.20)	-0.55 (0.12)	-0.52 (0.20)	-0.39 (0.12)	-0.41 (0.20)	-0.42 (0.11)
Risk Aversion							0.33 (0.17)	0.16 (0.11)
Log Income					-0.14 (0.63)	-0.68 (0.26)	-0.09 (0.63)	-0.66 (0.26)
Years of Schooling					-0.52 (0.19)	-0.80 (0.16)	-0.51 (0.19)	-0.79 (0.16)
Female			-1.44 (1.11)	2.47 (0.71)	-1.24 (1.14)	2.20 (0.71)	-1.38 (1.15)	2.10 (0.71)
Age			-0.15 (0.05)	-0.16 (0.03)	-0.14 (0.06)	-0.16 (0.03)	-0.15 (0.06)	-0.16 (0.03)
Age <sup>2</sup>			-1.59E-03 (2.50E-03)	-1.64E-04 (1.49E-03)	-3.48E-03 (2.80E-03)	-2.07E-04 (1.49E-03)	-3.46E-03 (2.80E-03)	-1.93E-04 (1.49E-03)

Notes: This table compares the relationship between decision-making ability and payday loan demand in U.S. data from the Understanding America Study (UAS) with the analogous evidence from Iceland. The dependent variable in the former is an indicator variable for whether the participant had had a payday loan in the past year. In the latter, the dependent variable is an indicator for whether the participant had a payday loan during a 6-year period. The dependent variables were multiplied by 100, such that the coefficients are in percentage points. The mean of the dependent variable is 4.89 in Iceland and 5.05 in the UAS. Decision-making ability and risk aversion are measured in percentile ranks divided by 10, such that the coefficient gives the effect of increasing decision-making ability or risk aversion in 10 percentiles. For example, the coefficient in the first column implies that an increase in decision-making ability in 10 percentiles is associated with a reduction in 0.41 percentage points in the probability of having a payday loan. In Iceland, the number of participants is 1,573. The UAS data is longitudinal with two waves. The number of observations is 5,243 and the number of participants is 2,954 (not all participants were surveyed in both waves). Standard errors are clustered at the individual level in the UAS. Robust standard errors are estimated for Iceland.

<sup>14</sup> Derived just from the risk experiment, we are limited in the UAS to measuring preferences with risk tolerance and decision-making ability with consistency with maximization of a utility function that satisfies a dominance principle.

The point estimates are similar in the Icelandic and U.S. data. The unconditional correlation between the percentile rank of the decision-making ability distribution and the probability of taking a loan is -0.41 in Iceland and -0.53 in the US data. Conditioning on several economic and demographic variables, and on a measure of risk aversion, leaves the point estimate in the Icelandic data effectively unchanged and brings the point estimate in the U.S. data to -0.42. Coefficients on the other variables are qualitatively similar in the two data sets, with the exception of gender. In the US, women are approximately two percentage points more likely to report they have or have had a payday loan while the (relatively imprecise) point estimate in Iceland is approximately -1.4. Taken together, we view the similarities of the two sets of estimates as evidence of the external relevance of the richer set of results derived from the Icelandic data.

### **VIII. Conclusion**

Motivated by the debate on regulation of the high-cost credit market, this paper evaluated the relationship between adverse financial conditions (“misfortune”), imperfect decision-making (“mistakes”), and the demand for high-cost credit. The policy debate revolves around efforts to restrict the circumstances under which individuals may obtain high-cost credit, and the possibility that many choices to take such loans are imperfect. Advocates of regulation see high-cost credit as too often exploiting unsophisticated borrowers who would be better off without the loans. Opponents of the regulation see this form of credit as serving those who are in acute need of liquidity and who find it difficult to obtain elsewhere.

Advancing the debate is difficult in part because “mistakes” are typically hard to identify. Unobserved constraints, preferences, or beliefs can justify many behaviors as optimal, including the demand for high-cost credit. We addressed this identification problem by combining high-quality administrative and experimental data from Iceland. The administrative data describe in detail the financial conditions associated with high-cost loan demand. In the experimental data, we manipulated constraints while holding preferences and beliefs constant, which allowed us to identify choice imperfections that provide a measure of decision-making ability.

We then related high-cost loan demand to measures of decision-making ability along with measures of constraints and preferences. The results show that borrowers of high-cost loans have

less income and liquidity, and are especially illiquid in the days leading up to getting the loan. They also exhibit substantially lower decision-making ability in the experiments; 28% of payday loan dollars are lent to the bottom 10% of the decision-making ability distribution, and 53% are lent to the bottom 20%.

In a regression framework, the relationship between decision-making ability and high-cost loan demand is not explained by demographic characteristics, economic circumstances, or measures of preferences from the experiment, and is mirrored by the relationship between decision-making ability and an unambiguous “mistake,” the accrual of not sufficient fund fees. The external relevance of the Iceland findings is supported by results from a survey of U.S. consumers where the relationship between decision-making ability and the probability of receiving a payday loan is very similar.

Taken together, the results of this study indicate that both misfortune and mistake are important for high-cost loan demand. It follows that policy may be justified if it works to address market imperfections that make this credit market incomplete, or if it better equips consumers to consider carefully the consequences of taking up high-cost loans.

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## APPENDIX

### Examples of Recent Payday Loan Contracts in Iceland

**Figure A1: Recent Payday Loan Contract**

9/25/2018 Loan Agreement

Heim Sport og vaxið Hættu þess egi umf. Um þessu Daga tilfæla

# Lánssamningur

1. SAMNINGADAGAR 1909 577-1909

Lánveitandi: Ecommerce 2020 ApS  
Havnegade 39  
1058 København K

Lánþegi: Cvr. 38073532  
Hákon Stefánsson  
Ólafsgaia 53, 113

Kt. 0507724709

## 2. LÁNAFYRIRGREIÐSLA

Lánafyrirgreiðsla:	20.000 ISK
Stofnað:	
Utlánsveitur	Vaxtalaust
Gildistími lánasamnings	95 daga
Heildar kostnaður	

Lán	20.000 ISK
Lántökugjald	0 ISK
Heildarkostnaður vaxta o.fl. á gildistíma	0 ISK
Mánaðarleg umslúgjöld	6.816 ISK

Heildar lántökukostnaður ásamt lánsfjárhæð sem skal endurgreiða	26.816
Árlegur hlutfallslegur kostnaður í prósentum (ÁHK)	3.444,8 %
Afborganir ásamt vöxtum og kostnaði.	Lánið endurgreiðist í þrem afborgunum (greiðslum): Fyrsta afborgun að upphæð 26.814 ISK greiðist eftir 30 daga. Af þessari upphæð er 1 endurgreiðsla. Önnur afborgun (endurgreiðsla) að upphæð 1 ISK endurgreiðist eftir 60 daga. Þriðja afborgun (loka endurgreiðsla) að upphæð 1 ISK endurgreiðist eftir 95 daga.

https://1909.dk/IS/Default/Agreement?cost=6816&amount=20000&ahk=3444.8&days=30 1/5

1909

1 Upphæð láns: Upphæð láns: 20.000 kr. lán

2 Niðurstaða

Lengd láns: 15 dagar 22 dagar 30 dagar

Kostnaður: 6.816 kr.

ÁHK: 3.444,8 %

Samtals: 26.816 kr.

☐ Eg hef lesið Lánssamningur e/Commerce 2020 ApS og samþykki þá.

☐ Eg hef lesið EU Lán - info og samþykki þá.

SENDU UMKSÖN

*Note:* A screen shot of an Iceland payday loan contract and associated user interface from 2018. The loan is for 20,000 ISK (approximately \$200) with a term of 30 days. The finance charge is 6,816 ISK (approximately \$68) which, as declared on the contract and user interface amounts to an APR of 3,444.8%. Source: Starfsumhverfi smálánafyrirtækja á Íslandi, The Ministry of Tourism, Industry and Innovation, January (2019).

### *Time Event Study*

We first restricted the sample to participants who received a payday loan between October 1, 2014 and January 14, 2017 to ensure we could observe liquidity in the 30 days before loans taken on October 1, 2014 and liquidity in the 30 days after loans taken on January 14, 2017.

First, we “detrended” participants’ time series of liquidity, controlling for day-of-the-week and calendar-day-of-the-month effects. In particular, using the entire time series of those participants (from September 1, 2014 to February 13, 2017), we ran a regression of liquidity on individual-specific fixed effects, dummies for day of the week, and dummies for calendar day of the month (Monday and the first of the month were the omitted categories). Liquidity is the sum of checking account balance plus the savings account balance minus the credit card balance. Our main outcome of interest is the residual from this regression,  $\hat{u}_{it}$ , where  $i$  indexes the participant and  $t$  the day.

Let  $Loan_{it}$  be equal to 1 if participant  $i$  received a loan on day  $t$  and 0 otherwise. In a second step, we restricted the sample to participant-days that fell up to 30 days before or up to 30 days after the participant received a loan:

$$\{it \mid Loan_{it-30} = 1 \text{ or } Loan_{it-29} = 1 \text{ or } \dots \text{ or } Loan_{it+29} = 1 \text{ or } Loan_{it+30} = 1 \}$$

Finally, we ran the following regression:

$$\hat{u}_{it} = \alpha + \sum_{k=1}^{30} \beta_k Before_{it}^k + \sum_{k=1}^{30} \gamma_k After_{it}^k + \varepsilon_{it}$$

where  $Before_{it}^k$  is equal to 1 if participant  $i$  took a loan on  $t + k$  (and 0 otherwise). Similarly,  $After_{it}^k$  is equal to 1 if participant  $i$  took a loan on  $t - k$  (and 0 otherwise).

### *Understanding America Study*

The Understanding America Study (UAS) is an Internet panel with respondents aged 18 and older living in the U.S.<sup>15</sup> About twice a month, respondents receive an email with a request to visit the UAS site and complete questionnaires.

In UAS survey modules 18 and 119, participants were asked the following question:

Payday loans are small, short-term loans that must be paid in full when the borrowers receive their next pay check or other regular deposit (such as a Social Security payment). These loans are often paid with a post-dated check. Please select the following statement that best describes your situation regarding these products.

- 1 I have never considered getting a payday loan from a payday lender
- 2 I currently have a payday loan
- 3 I have had a payday loan in the past year
- 4 I currently have a payday loan and I have had one in the past year
- 5 I considered getting a payday loan but was rejected
- 6 I have considered getting a payday loan but decided not to get it

We coded those that answered (2), (3), and (4) as having had a payday loan.

In UAS survey module 5, participants were administered Choi et al. (2014)'s risk choice experiment. Each participant made 25 choices. There were ten different sets, each with 25 budget lines. Participants were randomly assigned to one of these sets. The measure of decision-making ability is a unified measure of violations of GARP and FOSD—as with these participants in Iceland who did not have the outside option. We calculated the percentile rank of decision-making ability, separately for each one of the ten sets. We also used the choices from the risk choice experiment to calculate risk aversion, following the same procedure described in the paper (with the exception that in this case we used all 25 budget lines).

The demographic information (i.e., gender and age) and the information about education (which was converted from highest degree to years of schooling) come from the “my household” survey module that is administered on a quarterly basis. The information on total household income comes from a survey module (in particular survey module 24) based on the 2014 wave of the Health and Retirement Study (HRS) that was administered to all UAS participants.

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<sup>15</sup> Respondents are recruited by address-based sampling. Those without Internet access at the time of recruitment are provided tablets and Internet access.

## *Experimental Tasks*

### *Risk Task*

Participants allocated an experimental endowment of 500 kr. (appr. \$5) across two or five assets. The assets paid different amounts depending on whether a ball drawn from an urn was black or white. Participants were informed that the urn had five black balls and five white balls. Their decisions involved choosing how much to invest on each asset. Participants were presented with 15 investment problems (one of the 15 problems was randomly selected for payment). We varied the asset returns across the investment problems.

In the first eight investment problems, there were two assets. In the last seven investment problems, there were five assets. This design, which follows Carvalho & Silverman (2019), permits identifying the effects of complexity on each participant.

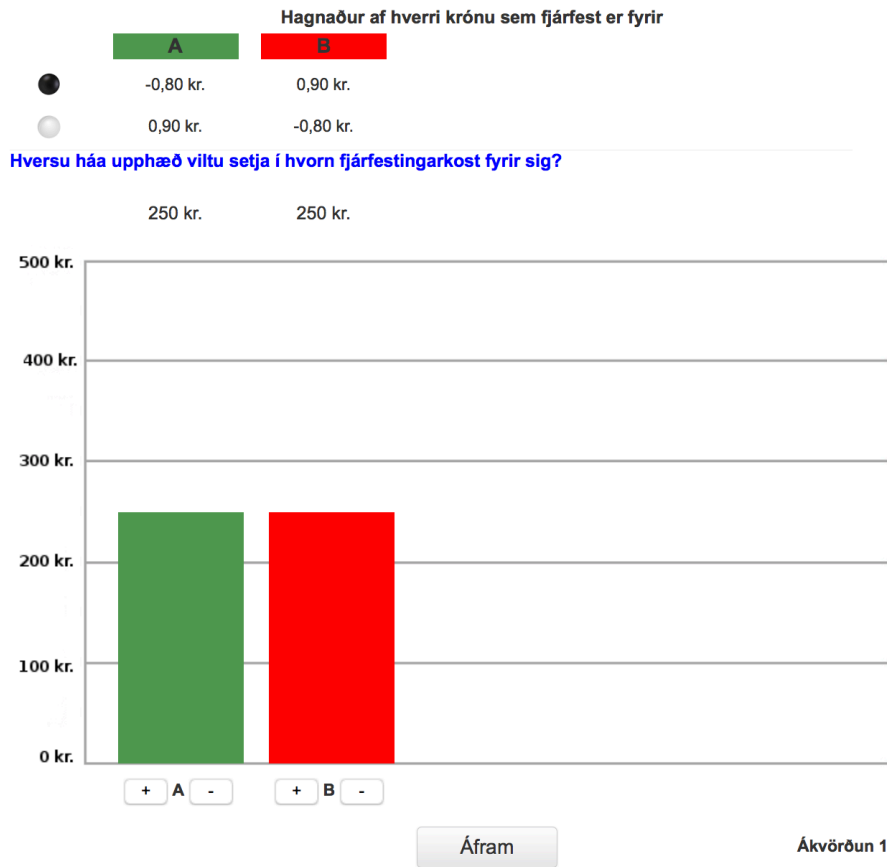
Appendix Figure 1 shows a screenshot of the interface for problems with two assets. The table at the top of the screen shows the returns of assets A and B per 1 kr. invested. The participant was then prompted to make her investment choices. The graph below the table displays two bars: the first bar shows the amount invested on asset A; the second bar shows the amount invested on asset B. Participants made their investments by either dragging the bars up and down or by clicking on the + and – buttons.<sup>16</sup> We originally included negative contingent returns in order to be able to estimate loss aversion. Participants were given a show up fee of 800kr. or more to ensure that the show up fee plus the earnings in all three tasks would be positive.<sup>17</sup>

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<sup>16</sup> The interface was such that participants always invested 100% of their experimental endowment.

<sup>17</sup> Some participants were randomly assigned to have a show up fee of 4,000kr.

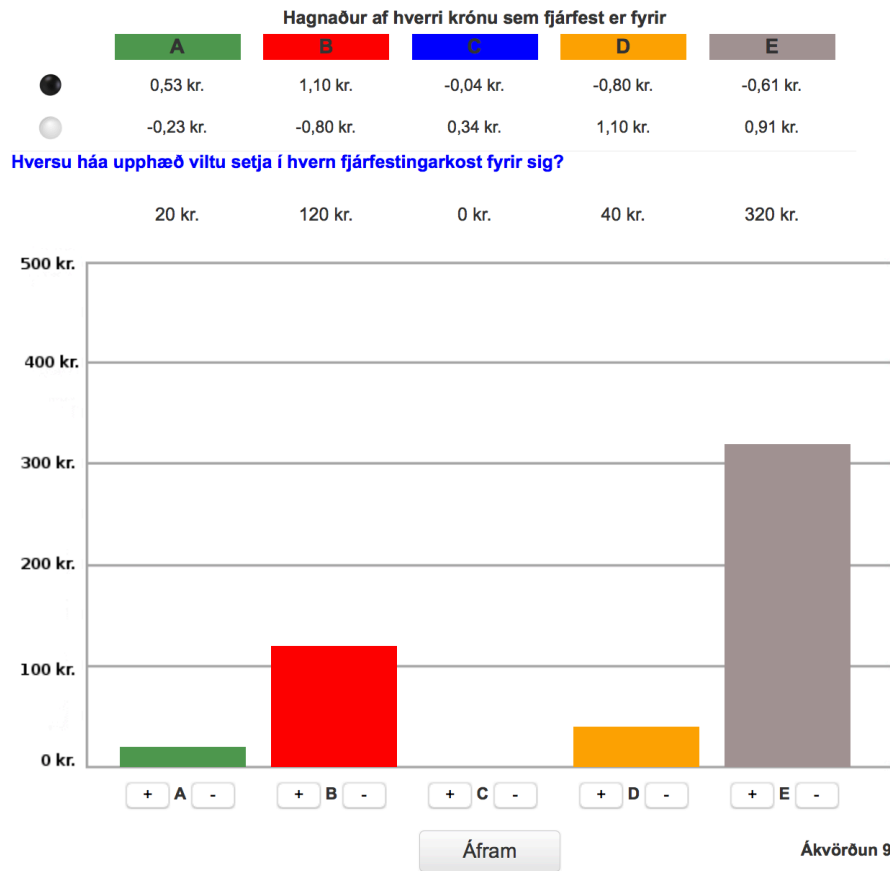
**Appendix Figure 1: Interface Risk Task with Two Assets and without Outside Option**



*Note:* This figure shows the interface participants without the outside option used to make investment choices in the risk task when two assets were available (i.e., choices 1-8). That is also the interface that all participants used to make their investment choices in the ambiguity task.

Appendix Figure 2 shows that a similar interface was used in the investment problems with five assets. The only distinction is that they were shown information about 5 assets – A, B, C, D, and E – and the graph displayed 5 bars.

**Appendix Figure 2: Interface Risk Task with Five Assets and without Outside Option**



*Note:* This figure shows the interface participants without the outside option used to make investment choices in the risk task when five assets were available (i.e., choices 9-15).

In order to study choice avoidance, half of the participants were randomly assigned to be offered the option of avoiding the investment problem (Carvalho & Silverman 2019). In particular, these participants were offered the choice between making the investment decision or taking an outside option of –50 kr., 0 kr., or 100 kr. The amount of the outside option was varied across the investment problems. The participant was paid the outside option if in the problem selected for payment she chose to avoid. Appendix Table 1 shows the parameters of the 15 decision problems.

**Appendix Table 1: Parameters Risk Task**

	Outside Option	<b>A</b>		<b>B</b>		<b>C</b>		<b>D</b>		<b>E</b>	
		Heads	Tails	Heads	Tails	Heads	Tails	Heads	Tails	Heads	Tails
1	100	-0.80	0.90	0.90	-0.80						
2	-50	0.95	-0.80	-0.80	1.13						
3	-50	0.70	-0.80	-0.80	1.20						
4	100	-0.80	0.35	1.50	-0.80						
5	-50	-0.80	0.45	1.40	-0.80						
6	100	-0.80	0.80	1.20	-0.80						
7	0	0.46	-0.80	-0.80	1.30						
8	0	-0.80	0.60	1.30	-0.80						
9	0	0.53	-0.23	1.10	-0.80	-0.04	0.34	-0.80	1.10	-0.61	0.91
10	100	-0.80	1.60	-0.71	1.36	0.10	-0.80	-0.17	-0.08	-0.44	0.64
11	0	-0.55	0.10	0.20	-0.20	0.95	-0.50	1.70	-0.80	-0.80	0.20
12	-50	-0.69	1.18	-0.36	0.52	-0.03	-0.14	0.30	-0.80	-0.80	1.40
13	-50	-0.20	0.40	-0.80	1.20	-0.65	1.00	0.70	-0.80	0.25	-0.20
14	100	-0.60	0.64	0.60	-0.32	0.00	0.16	-0.80	0.80	1.20	-0.80
15	0	1.30	-0.80	-0.80	0.60	-0.59	0.46	0.04	0.04	0.67	-0.38

*Note:* This table shows the parameters of the 15 decisions in the risk task. The first column shows the outside option. The other columns show for each asset the return per 1 kr. invested depending on the outcome of the coin toss.

Appendix Figure 3A and Appendix Figure 3B show screenshots of the interfaces used by participants with the outside option. It differs from the interface used by other participants (Appendix Figure 1 and Appendix Figure 2) in two ways. First, the graph with the bars is not shown. Second, the prompt to invest (“You will choose the amount you want to invest on each asset.”) is replaced by a prompt for the participant to choose between investing the experimental endowment (button “Invest Y kr.”) and taking the outside option (button “Receive X kr.”). If she clicked on the first button, the bars were unveiled and she could make her investment choices using the same interface used by other participants. If she clicked on the second button, she was presented with the next decision problem.

### Appendix Figure 3A: Interface Risk Task with Two Assets and with Outside Option

Hagnaður af hverri krónu sem fjárfest er fyrir

	A	B
<input checked="" type="radio"/>	-0,80 kr.	0,90 kr.
<input type="radio"/>	0,90 kr.	-0,80 kr.

Hvort kýst þú að:

Ákvörðun 1

*Note:* This figure shows the screen in which participants with the outside option were prompted to choose between investing in two assets and taking the outside option (in this example 100 kr.). If the participant chose to invest, s/he used the interface shown in Appendix Figure 1 to make her investment choices. If s/he chose the outside option, she would move to the next decision (in this example move from decision 1 to decision 2).

### Appendix Figure 3B: Interface Risk Task with Five Assets and with Outside Option

Hagnaður af hverri krónu sem fjárfest er fyrir

	A	B	C	D	E
●	0,53 kr.	1,10 kr.	-0,04 kr.	-0,80 kr.	-0,61 kr.
○	-0,23 kr.	-0,80 kr.	0,34 kr.	1,10 kr.	0,91 kr.

Hvort kýst þú að:

Ákvörðun 9

*Note:* This figure shows the screen in which participants with the outside option were prompted to choose between investing in five assets and taking the outside option (in this example 0 kr.). If the participant chose to invest, s/he used the interface shown in Appendix Figure 2 to make her investment choices. If s/he chose the outside option, she would move to the next decision (in this example move from decision 9 to decision 10).

### *Ambiguity Task*

The ambiguity task was similar to the risk task with three distinctions. First, participants were informed that the urn now had eight balls of one color and two balls of the other color. However, they did not know whether the urn had eight black balls and two white balls or if it had two black balls and eight white balls. Second, in all 15 investment problems there were just two assets. Third, participants were not offered the option of avoiding the investment problem. Appendix Table 2 shows the parameters of the 15 investment problems. As in the risk task, one of the 15 problems was randomly selected for payment.

We had a slightly different research idea at the time of the data collection that required having measures of ambiguity aversion. Hence why we administered the ambiguity task. We think there is no clear prediction as to how ambiguity aversion may affect the demand for payday loans so we opted for excluding ambiguity aversion from the analyses reported in this paper.

**Appendix Table 2: Parameters Ambiguity Task**

	<b>A</b>		<b>B</b>	
	Heads	Tails	Heads	Tails
1	-0.80	0.90	0.90	-0.80
2	0.95	-0.80	-0.80	1.13
3	0.70	-0.80	-0.80	1.20
4	-0.80	0.35	1.50	-0.80
5	-0.80	0.45	1.40	-0.80
6	-0.80	0.80	1.20	-0.80
7	0.46	-0.80	-0.80	1.30
8	-0.80	0.60	1.30	-0.80
9	1.10	-0.80	-0.80	1.10
10	0.10	-0.80	-0.80	1.60
11	-0.80	0.20	1.70	-0.80
12	-0.80	1.40	0.30	-0.80
13	1.14	-0.80	-0.80	0.96
14	0.19	-0.80	-0.80	1.50
15	1.60	-0.80	-0.80	0.27

*Note:* This table shows the parameters of the 15 decisions in the ambiguity task. It shows for assets A and B the return per 1 kr. invested depending on the outcome of the coin toss.

### *Intertemporal Choice Task*

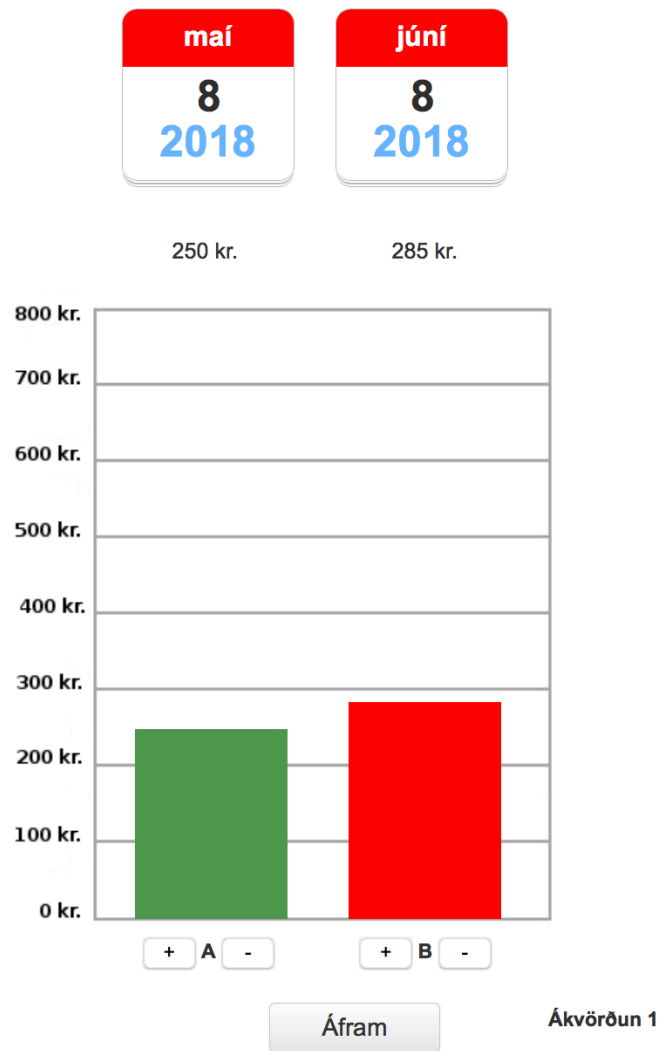
Participants had to allocate their experimental endowment across a sooner date and a later date. The amount allocated to the later date accrued an experimental interest rate. Participants were presented with 12 intertemporal allocation problems (one of the 12 problems was randomly selected for payment). We varied the experimental endowment, the experimental interest rate, and the sooner date across the problems. In the first six problems, the sooner date was today. In the last six problems, the sooner date was one year away. The time interval between the sooner and later dates was always one month. Within a time frame, the interest rate increased monotonically. Appendix Table 3 shows the parameters of the 12 intertemporal allocation problems.

Appendix Figure 4 shows a screenshot of the interface participants used in the intertemporal allocation task. Two calendar sheets at the top of the screen show the sooner date (calendar sheet on the left) and the later date (calendar sheet on the right). The graph below the calendar sheets displays two bars: the bar on the left shows the amount to be received at the sooner date; the bar on the right shows the amount to be received at the later date (including the interest accrued). Participants made their intertemporal allocations by either dragging the bars up and down or by clicking on the + and – buttons.<sup>18</sup>

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<sup>18</sup> The interface was such that participants always invested 100% of their experimental endowment.

**Appendix Figure 4: Intertemporal Choice Task**



*Note:* This figure shows the interface participants used to make their intertemporal choices. The calendar on the left showed the earlier date. The calendar on the right showed the later date. The green bar on the left showed the amount received sooner. The red bar on the right showed the amount received later.

**Appendix Table 3: Parameters Intertemporal Choice Task**

	Sooner	Later	Endowment	Interest Rate
1	Today	In 1 Month	550	-5%
2	Today	In 1 Month	550	10%
3	Today	In 1 Month	475	25%
4	Today	In 1 Month	475	50%
5	Today	In 1 Month	400	75%
6	Today	In 1 Month	350	125%
7	In 12 Months	In 13 Months	550	-5%
8	In 12 Months	In 13 Months	550	10%
9	In 12 Months	In 13 Months	475	25%
10	In 12 Months	In 13 Months	475	50%
11	In 12 Months	In 13 Months	400	75%
12	In 12 Months	In 13 Months	350	125%

*Note:* This table shows the parameters of the 12 decisions in the intertemporal choice task.

**Appendix Table 4: Non-linear Effects of Liquidity**

	<i>Number of Payday Loans</i>					
	(1)	(2)	(3)	(4)	(5)	(6)
DMA	-0.12 (0.06)	-0.11 (0.06)	-0.11 (0.06)	-0.12 (0.06)	-0.13 (0.07)	-0.12 (0.07)
Liquidity	-0.47 (0.08)	-0.46 (0.08)	-0.94 (0.35)	-2.01 (0.45)	-3.08 (0.78)	-3.09 (1.29)
Liquidity <sup>2</sup>		0.09 (0.03)				
Liquidity * Top Half			0.77 (0.38)			
Top Half			-3.26 (1.06)			
Liquidity * Middle Tercile				1.64 (0.49)		
Liquidity * Top Tercile				1.82 (0.45)		
Middle Tercile				-3.43 (1.65)		
Top Tercile				-4.51 (1.19)		
Liquidity * 2nd Quartile					3.40 (1.13)	
Liquidity * 3rd Quartile					2.97 (0.78)	
Liquidity * Top Quartile					2.91 (0.77)	
2nd Quartile					-7.05 (2.98)	
3rd Quartile					-6.14 (1.52)	
Top Quartile					-5.71 (1.45)	
Liquidity * 2nd Vintile						3.18 (1.32)
Liquidity * 3rd Vintile						2.57 (1.36)
Liquidity * 4th Vintile						2.97 (1.28)
Liquidity * Top Vintile						2.86 (1.29)
2nd Vintile						-6.54 (1.81)
3rd Vintile						-3.56 (3.17)
4th Vintile						-6.11 (1.71)
Top Vintile						-5.23 (1.84)

Notes: This table investigates whether the effects of liquidity are nonlinear. Decision-making ability, liquidity, and preferences are measured in percentile ranks divided by 10. The mean of the dependent variable is 0.94. The regressions include controls for impatience, present bias, risk aversion, log income, years of schooling, gender, age and age squared. *Number of observations* = 1,573.

**Appendix Table 5: Alternative Measures of Right-Hand Side Variables**

	Number of Loans							
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
DMA	-0.12 (0.06)	-0.12 (0.06)	-0.10 (0.06)	-0.12 (0.06)	-0.13 (0.06)	-0.15 (0.07)	-0.14 (0.07)	-0.17 (0.08)
Pctile Rank Liquidity	-0.47 (0.08)	-0.47 (0.09)	-0.44 (0.08)	-0.41 (0.08)	-0.47 (0.09)			-0.47 (0.08)
Liquidity in \$10,000s						-0.26 (0.09)		
5th Percentile of Liquidity							-0.43 (0.22)	
10th Percentile of Liquidity							0.07 (0.20)	
20th Percentile of Liquidity							-0.18 (0.10)	
Impatience	0.02 (0.04)	0.02 (0.04)	0.01 (0.04)	0.02 (0.04)	0.02 (0.04)	0.04 (0.04)	0.04 (0.04)	
Discount Rate								-0.06 (0.05)
Present Bias	0.06 (0.07)	0.06 (0.07)	0.07 (0.07)	0.06 (0.07)	0.06 (0.07)	0.08 (0.07)	0.08 (0.07)	
Present Bias Parameter								-0.04 (0.07)
Risk Aversion	0.07 (0.06)	0.07 (0.06)	0.06 (0.05)	0.07 (0.05)	0.07 (0.06)	0.09 (0.06)	0.09 (0.06)	
1/Intertemp. Elas. of Subs.								0.06 (0.07)
Log Income	0.70 (0.21)	0.77 (0.21)	0.73 (0.23)			0.25 (0.18)	0.29 (0.18)	0.69 (0.21)
Income in \$10,000s				2.55E-03 (2.00E-03)				
Pctile Rank Income					0.19 (0.08)			
Years of Schooling	-0.02 (0.05)	-0.01 (0.05)		-0.02 (0.05)	-0.02 (0.05)	-0.07 (0.05)	-0.07 (0.05)	-0.02 (0.05)
High School Graduate			-0.97 (1.09)					
College Degree			-1.92 (0.90)					
Postgraduate Degree			-1.66 (0.89)					
Female	-0.94 (0.39)	-0.93 (0.39)	-0.81 (0.36)	-1.05 (0.40)	-0.92 (0.38)	-0.86 (0.38)	-0.89 (0.39)	-0.90 (0.38)
Age	4.26E-04 (1.65E-02)		-3.32E-03 (1.69E-02)	2.14E-02 (1.68E-02)	-1.98E-03 (1.58E-02)	-1.47E-02 (1.67E-02)	-1.39E-02 (1.68E-02)	2.42E-03 (1.70E-02)
Age <sup>2</sup>	-4.90E-04 (6.56E-04)		-9.06E-04 (6.97E-04)	-1.39E-03 (6.42E-04)	-4.62E-04 (6.76E-04)	-6.98E-04 (6.75E-04)	-7.13E-04 (6.71E-04)	-5.18E-04 (6.45E-04)
Age between 35 and 49		-0.35 (0.35)						
Age between 50 and 64		0.35 (0.73)						
65 and Older		-1.18 (0.48)						

Notes: This table investigates the robustness of the results. “5<sup>th</sup> Percentile of Liquidity”, “10<sup>th</sup> Percentile of Liquidity”, and “20<sup>th</sup> Percentile of Liquidity” were constructed by calculating the within-participant 5<sup>th</sup>, 10<sup>th</sup>, and 20<sup>th</sup> percentiles of liquidity over time. The discount rate, the present bias parameter, and the intertemporal elasticity of substitute are estimated using the intertemporal choices and then ranked across participants. The mean of the dependent variable is 0.94. *Number of observations* = 1,573.

**Appendix Table 6: Measurement Error in Decision-Making Ability**

	<i>Amb. DMA</i> 1st Stage	<i>Number of Loans</i>			
		OLS	2SLS	OLS	2SLS
Risk DMA	0.19 (0.02)				
Ambiguity DMA		-0.19 (0.09)	-0.48 (0.28)	-0.19 (0.09)	-0.48 (0.28)
Ambiguity DMA * Liquidity				0.06 (0.02)	0.05 (0.03)
Liquidity	0.00 (0.02)	-0.47 (0.09)	-0.46 (0.08)	-0.46 (0.08)	-0.46 (0.08)
Impatience	-0.06 (0.01)	0.03 (0.04)	0.01 (0.04)	0.03 (0.04)	0.01 (0.04)
Present Bias	-0.02 (0.02)	0.07 (0.07)	0.06 (0.07)	0.07 (0.07)	0.07 (0.07)
Risk Aversion	-0.06 (0.02)	0.08 (0.06)	0.05 (0.05)	0.08 (0.06)	0.05 (0.05)
Log Income	0.10 (0.08)	0.70 (0.21)	0.73 (0.22)	0.74 (0.22)	0.76 (0.23)
Years of Schooling	0.01 (0.02)	-0.02 (0.05)	-0.02 (0.05)	-0.02 (0.05)	-0.02 (0.05)
Female	0.08 (0.10)	-0.92 (0.38)	-0.90 (0.37)	-0.86 (0.37)	-0.84 (0.36)
Age	-6.25E-03 (0.01)	2.36E-03 (0.02)	-1.69E-03 (0.02)	2.84E-03 (0.02)	-1.22E-03 (0.02)
Age <sup>2</sup>	7.98e-05 (3.09E-04)	-4.36E-04 (6.62E-04)	-4.59E-04 (6.71E-04)	-3.43E-04 (6.69E-04)	-3.78E-04 (6.73E-04)

Notes: This table investigates the extent to which measurement in decision-making ability biases the estimates of the relationship between payday loan demand and decision-making ability. We use the measure of decision-making ability in the risk task to instrument for decision-making ability in the ambiguity task. The first column shows the first stage. The second and fourth columns show OLS estimates. The third and fifth columns show two stages least squares estimates. The mean of the dependent variable is 0.94. Decision-making ability, liquidity, and preferences are measured in percentile ranks divided by 10. *Number of observations* = 1,573.

**Appendix Table 7: Decision-making Ability and High-Frequency Variation in Liquidity, Interactive Effects**

	<i>I if Took a Payday Loan</i>					
	(1)	(2)	(3)	(4)	(5)	(6)
DMA	-0.85 (0.43)	-0.85 (0.43)	-0.78 (0.42)	-0.80 (0.42)	-0.86 (0.43)	-0.86 (0.43)
Liquidity in Levels	-76.57 (24.37)	-69.64 (22.99)				
DMA * Liquidity		11.55 (7.89)				
IHS of Liquidity			-2.49 (0.66)	-2.43 (0.62)		
DMA * IHS of Liquidity				0.57 (0.25)		
PR of Liquidity					-0.37 (0.18)	-0.37 (0.18)
DMA * PR of Liquidity						0.05 (0.08)
Impatience	0.08 (0.23)	0.08 (0.23)	0.03 (0.23)	-0.01 (0.23)	0.10 (0.23)	0.10 (0.23)
Present Bias	0.54 (0.37)	0.54 (0.37)	0.48 (0.36)	0.48 (0.36)	0.55 (0.37)	0.55 (0.37)
Risk Aversion	0.14 (0.26)	0.14 (0.26)	0.10 (0.26)	0.12 (0.25)	0.13 (0.26)	0.13 (0.26)
Log Income	1.10 (0.91)	1.06 (0.91)	2.58 (1.07)	2.80 (1.11)	0.62 (0.87)	0.62 (0.87)
Years of Schooling	-0.27 (0.20)	-0.28 (0.20)	-0.09 (0.18)	-0.09 (0.18)	-0.29 (0.20)	-0.29 (0.20)
Female	-3.72 (2.01)	-3.67 (2.00)	-3.62 (1.97)	-3.29 (1.91)	-3.57 (1.98)	-3.57 (1.98)
Age	-0.05 (0.10)	-0.05 (0.10)	-1.19E-03 (0.09)	-3.37E-03 (0.09)	-0.06 (0.10)	-0.06 (0.10)
Age <sup>2</sup>	-4.97E-03 (4.18E-03)	-5.11E-03 (4.18E-03)	-4.79E-03 (4.18E-03)	-4.23E-03 (4.18E-03)	-6.02E-03 (4.29E-03)	-6.01E-03 (4.29E-03)

*Note:* This table controls for more flexible forms of liquidity. It shows results from regressions at the individual-daily level. The dependent variable is an indicator for whether participant  $i$  took a payday loan on day  $d$ . We multiplied it by 10,000 so the coefficients can be interpreted as the effect on a hundredth of a percentage point. Its mean is 3.79. Liquidity refers to the liquidity on the previous day, i.e.,  $d - 1$ . Columns (1)-(2) include liquidity in levels as a control. Columns (3)-(4) control for the inverse hyperbolic sine of liquidity. Columns (5)-(6) add a *within*-participant percentile rank measure of liquidity. In particular, the liquidity of participant  $i$  on day  $d - 1$  was ranked relative to the liquidity of participant  $i$  in all other days in the individual time series of the participant. Decision-making ability, time and risk preferences are measured in percentile ranks divided by 10, such that the coefficient gives the effect of an increase of the independent variable in 10 percentiles. The regressions include dummies for day of the week and for calendar day

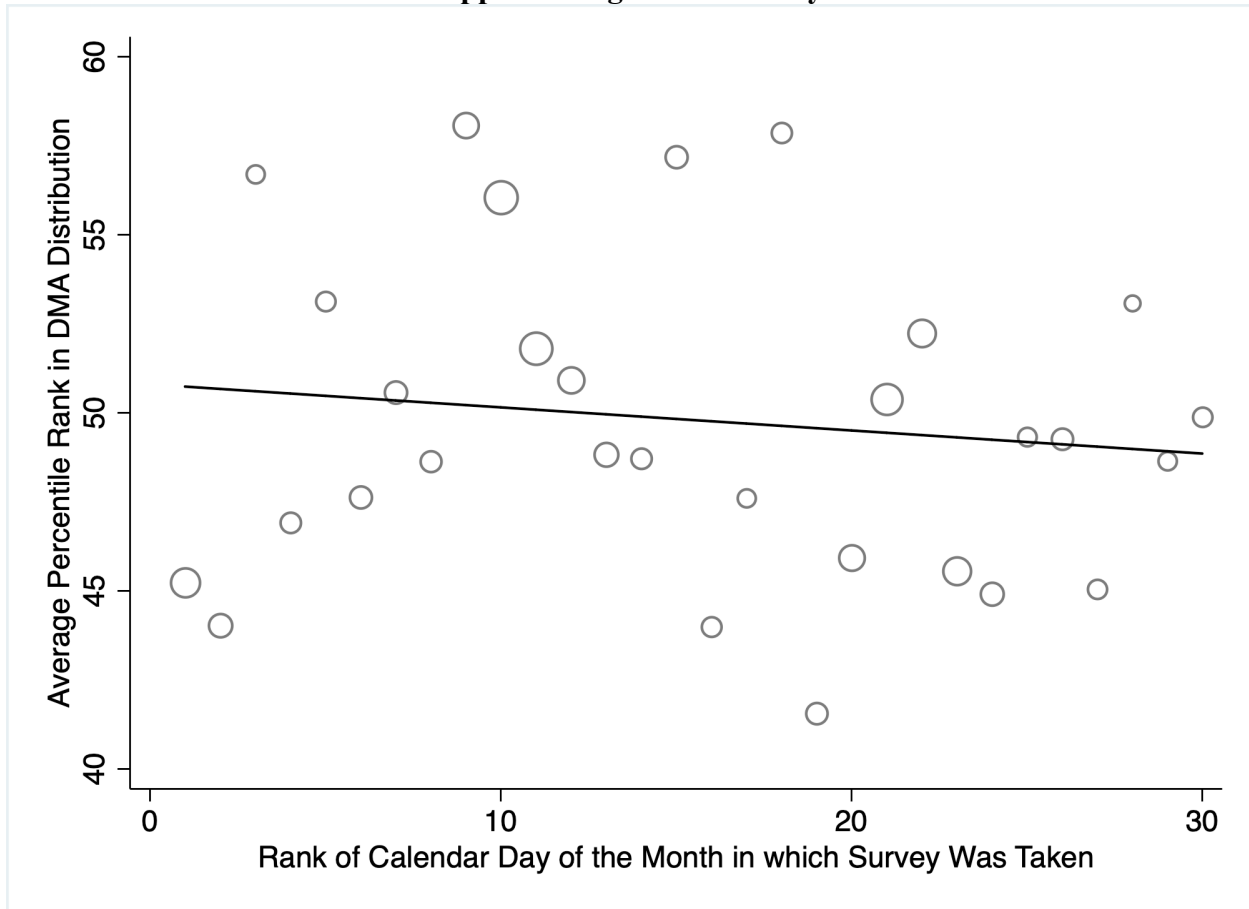
of the month. Number of observations = 1,388,959. Number of participants = 1,573. Number of days = 883. Standard errors clustered at the individual level.

**Appendix Table 8: Decision-making Ability and Monthly Variation in Liquidity**

	<i>Number of Loans in the Month</i>					
	(1)	(2)	(3)	(4)	(5)	(6)
DMA	-30.70 (13.91)	-30.14 (13.80)	-25.33 (13.08)	-25.30 (13.08)	-23.11 (12.63)	-23.20 (12.65)
Liquidity in Levels		-2,582.42 (816.78)	-2,527.88 (786.92)			
Lagged Liquidity in Levels				-2,420.37 (776.50)		
IHS of Liquidity					-79.65 (21.20)	
IHS of Lagged Liquidity						-74.53 (19.92)
Impatience			2.00 (7.13)	2.05 (7.14)	0.38 (7.14)	0.56 (7.13)
Present Bias			16.56 (11.13)	16.61 (11.14)	14.67 (10.87)	14.88 (10.89)
Risk Aversion			4.80 (7.79)	4.81 (7.79)	3.44 (7.75)	3.59 (7.75)
Log Income	11.73 (25.80)	31.04 (27.66)	33.75 (27.93)	32.86 (27.85)	79.78 (32.32)	76.57 (31.90)
Years of Schooling	-9.03 (6.06)	-8.25 (6.00)	-8.47 (5.98)	-8.49 (5.98)	-2.94 (5.60)	-3.28 (5.61)
Female	-97.67 (58.34)	-102.53 (59.19)	-112.35 (61.53)	-111.94 (61.45)	-109.26 (60.17)	-108.47 (60.09)
Age	-1.22 (2.99)	-0.96 (2.99)	-1.20 (2.94)	-1.22 (2.94)	0.32 (2.83)	0.21 (2.82)
Age <sup>2</sup>	-0.18 (0.13)	-0.15 (0.13)	-0.16 (0.13)	-0.16 (0.13)	-0.16 (0.13)	-0.16 (0.13)

*Note:* This table controls for more flexible forms of liquidity. It shows results from regressions at the individual-daily level. The dependent variable is the number of payday loans participant  $i$  took in month  $m$ . We multiplied it by 100 so the coefficients can be interpreted as the effect in percentage points. Its mean is 115.3. Liquidity refers to the liquidity on the same month, i.e.,  $m - 1$ . Lagged liquidity refers to the liquidity in the previous month, i.e.,  $m - 1$ . Columns (2)-(3) include liquidity in levels as a control. Column (4) controls for lagged liquidity in levels, column (5) for the inverse hyperbolic sine (IHS) of liquidity and column (6) for the IHS of lagged liquidity. Decision-making ability, time and risk preferences are measured in percentile ranks divided by 10, such that the coefficient gives the effect of an increase of the independent variable in 10 percentiles. Number of observations = 44,044. Number of participants = 1,573. Number of months = 28. Standard errors clustered at the individual level.

**Appendix Figure 5: Scarcity**



*Note:* This figure investigates the hypothesis that financial circumstances at the time of the survey may drive the measure of decision-making ability. For each calendar day of the month, it was calculated the within-participant median liquidity across months. For each participant, calendar days were then sorted from lowest median liquidity to highest median liquidity. The variable in the horizontal axis corresponds to the rank of the day in which the participant took the survey. The variable in the vertical corresponds to the average percentile rank of participants surveyed on a day of a given rank. The size of the circumference corresponds to the number of participants surveyed on this day. The correlation between the two variables is 0.0084. Number of observations = 1,576.

**Appendix Table 9: Comparison of Survey Sample to Meniga Users**

	Mean		P-value Test Diff in means	Frequency	Number of Observations
	Admin	Survey			
Female	49%	47%	0.172	—	11,630
Age	38.8	37.4	9.29E-06	—	11,630
1 if Took Payday Loan	0.04%	0.04%	0.795	Daily	25,798,650
Amount Payday Loan	0.10	0.11	0.746	Daily	25,798,650
Income	6,019	5,490	0.797	Monthly	837,360
Wages	5,300	3,230	0.117	Monthly	837,360
Checking Balance	366	377	0.961	Daily	10,432,110
Savings Balance	4,340	4,386	0.922	Daily	10,432,110
Credit Card Balance	1,735	1,823	0.467	Daily	10,432,110
Number of NSF Charges	0.05	0.03	5.40E-05	Monthly	846,856
Cost NSF Charges	0.36	0.28	0.002	Monthly	846,856

*Note:* This table compares survey participants ( $N = 1,573$ ) to Meniga users ( $N = 11,630$ ). Both samples are restricted to individuals for whom there are complete data on demographics, payday loans, income, and balances (they may have incomplete data on NSF charges). For the NSF outcomes, the number of individuals are respectively 1,542 and 11,444. The test of difference in means cluster standard errors at the individual level.

**Appendix Table 10: Economic Preferences and Real-Life Outcomes**

	Rank Percentile of Wealth				100 if Participates in Stock Market			Rank Percentile of Overdraft Balance		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
Impatience	-0.46 (0.18)	-0.45 (0.18)	-0.40 (0.19)	-0.31 (0.20)		-0.56 (0.25)	-0.61 (0.27)	0.99 (0.21)	0.95 (0.21)	0.90 (0.22)
Present Bias		-0.24 (0.24)	-0.25 (0.24)	-0.21 (0.24)		-0.36 (0.31)	-0.38 (0.31)	0.89 (0.28)	0.89 (0.28)	0.87 (0.29)
Risk Aversion			-0.48 (0.24)	-0.37 (0.26)	-1.01 (0.35)	-0.90 (0.35)	-0.96 (0.37)		0.37 (0.27)	0.31 (0.28)
DMA				0.39 (0.28)			-0.22 (0.39)			-0.22 (0.33)
Log Income	8.55 (1.34)	8.50 (1.34)	8.40 (1.34)	8.34 (1.33)	5.27 (1.53)	5.09 (1.53)	5.13 (1.53)	2.16 (1.15)	2.24 (1.14)	2.28 (1.14)
Years of Schooling	0.91 (0.23)	0.92 (0.23)	0.89 (0.23)	0.87 (0.23)	1.02 (0.28)	0.98 (0.28)	0.99 (0.28)	-0.73 (0.23)	-0.72 (0.24)	-0.71 (0.24)
Female	-6.44 (1.40)	-6.33 (1.40)	-6.12 (1.40)	-6.08 (1.40)	-11.09 (1.89)	-10.58 (1.88)	-10.59 (1.88)	1.88 (1.58)	1.71 (1.58)	1.70 (1.58)
Age	0.36 (0.08)	0.36 (0.08)	0.38 (0.08)	0.39 (0.08)	0.13 (0.10)	0.16 (0.10)	0.15 (0.10)	0.34 (0.08)	0.33 (0.08)	0.32 (0.08)
Age <sup>2</sup>	-4.08E-03 (4.81E-03)	-3.94E-03 (4.79E-03)	-3.93E-03 (4.78E-03)	-3.75E-03 (4.76E-03)	0.01 (5.68E-03)	0.01 (5.66E-03)	0.01 (5.67E-03)	-0.03 (4.28E-03)	-0.03 (4.27E-03)	-0.03 (4.27E-03)
Number of Obs	1,672	1,672	1,672	1,672	1,682	1,682	1,682	1,410,981	1,410,981	1,410,981
Mean of Y	48.86	48.86	48.86	48.86	18.91	18.91	18.91	27.98	27.98	27.98

*Note:* This table investigates whether the experimental measures of economic preferences predicted real-life outcomes that have been documented by previous work (Epper et al. 2018; Barsky et al. 1997; Meier & Sprenger 2010). In columns (8)-(10), standard errors clustered at the individual level. Robust standard errors in all other columns.