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HOW DO TRAVEL COSTS SHAPE COLLABORATION?

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## How Do Travel Costs Shape Collaboration?

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### **ABSTRACT**

We develop a simple theoretical framework for thinking about how geographic frictions, and in particular travel costs, shape scientists' collaboration decisions and the types of projects that are developed locally versus over distance. We then take advantage of a quasi-experiment — the introduction of new routes by a low-cost airline — to test the predictions of the theory. Results show that travel costs constitute an important friction to collaboration: after a low-cost airline enters, the number of collaborations increases between 0.3 and 1.1 times, a result that is robust to multiple falsification tests and causal in nature. The reduction in geographic frictions is particularly beneficial for high quality scientists that are otherwise embedded in worse local environments. Consistent with the theory, lower travel costs also endogenously change the types of projects scientists engage in at different levels of distance. After the shock, we observe an increase in higher quality and novel projects, as well as projects that take advantage of complementary knowledge and skills between sub-fields, and that rely on specialized equipment. We test the generalizability of our findings from chemistry to a broader dataset of scientific publications, and to a different field where specialized equipment is less likely to be relevant, mathematics. Last, we discuss implications for the formation of collaborative R&D teams over distance.

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# 1 Introduction

The drastic reduction in communication costs brought by the diffusion of the internet initially led to claims about a future in which technology could overcome geographic frictions and facilitate the rapid exchange of ideas, goods and services independent of distance (Cairncross, 1997; Friedman, 2005). Empirically, this “death of distance” hypothesis has found limited support, as most evidence points to agglomeration mattering more, not less, than before across a variety of settings (Leamer & Levinsohn, 1995; Blum & Goldfarb, 2006; Forman et al., 2005; Agrawal et al., 2015). Instead of substituting for co-location, digital interactions often complement it (Agrawal & Goldfarb, 2008),<sup>1</sup> resulting in non-obvious changes in how teams and organizations structure collaborations and develop new ideas when communication costs are low, but teamwork and R&D require specialized expertise and resources that are geographically dispersed (Adams et al. 2005; Jones, Wuchty & Uzzi 2008; Wuchty, Jones & Uzzi 2007).<sup>2</sup>

Moreover, not all types of interactions have benefited in the same way from improvements in communication technology. Co-location plays a disproportionate role in the serendipitous discovery of new collaborators and ideas (Catalini, 2017), and in the absence of offline opportunities for interaction, search frictions can prevent individuals from finding ideal collaborators even within the boundaries of the same institution (Boudreau et al., 2017). Similarly, exchanges that require the transfer of complex information and tacit knowledge (Polanyi, 1958; Von Hippel, 1994) still heavily rely on face-to-face interactions (Rosenthal & Strange, 2001; Gaspar & Glaeser, 1998; Storper & Venables 2004). As a result, firms, communities of experts and teams invest substantial amounts of time, effort and resources to ensure that the right individuals can be co-located – even if only temporarily – to discuss ideas, make

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<sup>1</sup>Agrawal & Goldfarb’s (2008) study of Bitnet, an internet predecessor, finds that as more academic institutions joined the network, collaboration among affected scientists increased. Interestingly, their results hint at the technology being a complement to offline interactions, as co-authorship increases disproportionately among university pairs that are co-located. Other studies have found an effect of Bitnet on collaborations in the academic life sciences (Ding et al. 2010), and of the internet on cooperative R&D between firms (Forman & Zeebroeck 2012).

<sup>2</sup>By 2000, less than 20% of papers in science and engineering were single authored. Similar patterns, and in particular the rise of coauthorship and distant coauthorship, have been documented in economics. See Gaspar & Glaeser (1998), Hamermesh & Oster (2004), Rosenblatt & Mobius (2004).

progress on projects, and develop the relationships that can later support more effective interactions over distance. Such temporary forms of co-location have been shown to foster both idea diffusion and the formation of new collaborations (Chai & Freeman, 2018).<sup>3</sup>

If face-to-face interactions are instrumental in finding and evaluating new collaborators, establishing trust, and advancing joint work, then as communication costs drop, if they are a complement and not a substitute to remote interactions, they should become more valuable. Furthermore, their absence would likely constitute the key remaining friction in the formation and operation of geographically distributed teams. Ironically, by making online communication extremely efficient, the internet may have enhanced the role that travel technology plays in the economy.

The objective of this paper is to develop and test a simple theoretical framework for thinking about how geographic frictions, and in particular travel costs, shape collaboration decisions and the types of projects that are developed locally versus over distance. The model highlights a key trade-off individuals face when deciding if they should work with a local versus a distant collaborator: whereas the global pool of potential collaborators is often deeper and may therefore offer an ideal match, collaboration over distance incurs additional communication and travel costs. We build on this basic tension in a context where individuals endogenously allocate effort to projects based on their potential, and where a project's variance in outcomes or the need for complementary expertise, equipment or resources can influence with whom a project is pursued. The simple framework captures an increasingly relevant challenge: to be able to solve problems of rising complexity, teams of specialized experts have to be put together (Jones, 2009), but this often involves collaboration over distance.

We take advantage of a quasi-experiment – the introduction of new routes by a major low-cost airline – to test the predictions of the theory within the context of collaborations

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<sup>3</sup>Chai & Freeman (2018) compare collaboration patterns among attendees of the Gordon conference before and after the event in a difference-in-differences framework using a carefully constructed control group of qualitatively similar non-participants. They find that attendees are more likely to be cited by, and collaborate with other participants, especially if they were new to this community of experts. In a related paper, Campos, de Leon & Mcquilin (2018) document that a conference cancellation led to a decrease in individuals likelihood of co-authoring together.

between scientific labs. The setting allows us to observe the full set of scientists at risk of collaboration in any given year as well as important characteristics about them such as their age, career stage, past productivity, area of specialization, and departmental funding.

The cheaper fares brought by the expansion of the low-cost airline (Southwest Airlines)<sup>4</sup> are part of a broader, 50% reduction in the cost of air travel that took place in the United States over the last 30 years (Perry 2014).<sup>5</sup> Furthermore, they provide a source of plausibly exogenous variation in the cost of conducting research between scientists at the affected airports.

Using a difference-in-differences empirical strategy we are able to recover a causal estimate of the effect of a reduction in travel costs not only on the rate of collaboration, but more importantly on the type of projects scientists pursue. Results show that travel costs are an important friction to collaboration: after Southwest entry, the number of collaborations increases between 0.3 and 1.1 times, a result that is robust to multiple falsification tests and causal in nature. The reduction in geographic frictions is particularly beneficial for high quality scientists that are otherwise embedded in worse local environments, although women scientists do not seem to benefit. Consistent with the theory, lower travel costs also endogenously change the types of projects scientists engage in locally versus over distance. After the shock, we observe an increase in higher quality and more novel projects, as well as projects that take advantage of complementary knowledge and skills between sub-fields, or that rely on specialized equipment. We test the generalizability of our findings within chemistry to a broader dataset of scientific publications, and to mathematics, a field where specialized equipment is less likely to be relevant. Last, we discuss implications for the formation of collaborative R&D teams in the presence of geographic frictions.

The rest of the paper is as follows: in Section 2 we provide additional institutional details about scientific collaboration, on how chemistry differs from others fields, and the data we

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<sup>4</sup>Southwest has been described as the most significant development in the market structure of the U.S. airline industry by the Transportation Research Board (1999) and by industrial economists (Morrison 2001, Borenstein & Rose 2007, Goolsbee & Syverson 2008).

<sup>5</sup>Kim, Morse & Zingales (2009) and Freeman, Ganguli & Murciano-Goroff (2014) note that secular declines in both communication costs and air travel costs may have facilitated long distance collaborations.

use. Section 3 introduces our empirical strategy and main results, together with a series of robustness tests and extensions targeted at assessing the generalizability of our findings to different samples. Section 4 develops a model to guide the interpretation of the findings, as well as the exploration of more nuanced hypotheses about the type of projects pursued in response to a reduction in geographic frictions. Section 5 tests these additional predictions, and Section 6 concludes.

The reason why we first explore the difference-in-differences results on the rate of collaboration (Section 3) and then present the model and test hypotheses about the type of collaborations that emerge (Sections 4 and 5) is because we follow the natural evolution of the project. We started from an empirical assessment of the presence of an effect on the rate of collaboration, moved to theory development to form predictions about the types of projects affected and to identify additional data to be collected (e.g. on novelty, equipment intensity), and then finally brought these new predictions to the data.

## 2 Scientific Collaboration

Scientific research is an increasingly collaborative endeavour, as reflected in the growing number of authors per paper over time (Wuchty, Jones & Uzzi, B., 2007). Collaborations are typically formed to combine skills and knowledge (Freeman, Ganguli, and Murciano-Goroff 2014), to access complementary generalist or specialist talent and resources (Teodoridis, 2015; Sauermann and Haeussler, 2017), and to expand the knowledge frontier when information is tacit and difficult to transfer or recombine without extensive, direct interactions (Stephan 2012). Increasing complexity has also been linked to a rising need for interdisciplinary teams (Falk-Krzesinski, 2011; Milojevi, 2014), with Wu et al. (2019) showing that both small and large teams play an important, but different role in pushing the knowledge frontier. Using a large-scale dataset of papers, patents and software products developed over 60 years, the authors show that while smaller teams are associated with more disruptive work and exploration, larger ones are systematically linked to advancing existing ideas and execution.

In terms of team formation, empirical evidence shows that researchers typically source

collaborations through their professional networks (Freeman, Ganguli, and Murciano-Goroff 2014), through serendipitous interactions with colocated individuals (Catalini, 2017) and conferences (Boudreau et al., 2017; Campos, de Leon & Mcquilin, 2018; Chai & Freeman 2018), and by relying on the information disclosed in scientific publications. As Walsh & Nee (2015) highlight, science is organized around increasingly complex teams that resemble the operations of small R&D-intensive firms, with knowledge as their core output.

Substantial differences, however, exist across different scientific disciplines in how collaboration and research is organized: for example, while mathematicians and theoretical physicists rarely work in labs, most research in chemistry, life sciences and experimental physics – also because of different capital, talent and infrastructure requirements – takes place in labs (Stephan 2012).

Our core analysis is focused on collaborations within chemistry. While chemistry largely remains a lab-based science, it has also not embraced the larger scale, big science projects observed in physics. Team size in chemistry – as measured by the number of co-authors – is lower than in biology and physics, though higher than in mathematics (Adams et al. 2005). Chemistry labs are run by a faculty member (principal investigator) who obtains funding for the lab, directs research projects, appears as a co-author on all publications, oversees resource allocation and effectively decides whether to collaborate or not with other labs. In our sample, the median number of co-authors per paper is 4, and many of the authors are graduate students, post-docs or technicians. These perform most of the experiments and day-to-day work on a project.

While many research projects involve a single principal investigator, collaborations between labs and principal investigators are common as well. Consistent with the findings from large scale surveys of scientists (Freeman, Ganguli, and Murciano-Goroff 2014), in our conversations with U.S. principal investigators in chemistry, complementary expertise, skills, materials or new types of experiments are all mentioned as reasons for collaboration across labs. As in other fields of science, collaborations in chemistry are sourced through the principal investigators and junior members' professional networks, serendipitous interac-

tions at conferences, email, etc. In the paper, we focus on collaborations between principal investigators, which are essentially collaborations between different labs.

## 2.1 Data Sources and Key Outcomes of Interest

To examine the effect of the changes in travel costs induced by the entry of Southwest Airlines on scientific collaboration, we combine data on scientists with publication records and air transportation information. Within the chemistry and mathematics samples, biographical information on scientists enables us to effectively disambiguate publication data, while also allowing us to separate faculty members from other types of authors. We now discuss in more detail the data sources we use and key outcomes we focus on throughout the paper.

***Air Transportation Data*** - To recover information on when Southwest operated flights between different routes, as well as information on prices, passengers and miles flown, we use data from the Airline Origin and Destination Survey (DB1B) of the U.S. Bureau of Transportation Statistics. The DB1B is a 10% random sample of airline tickets from reporting carriers in each quarter. For each itinerary, the DB1B records all connecting airports (including origin and destination), the itinerary fare, and other information. This data is publicly available only from 1993, hence we will focus on Southwest entry decisions that occur after 1993.

***Match Between Airports and Universities***. We compute distances between airports and universities using Google Maps. The matching between universities and airports is complicated by the fact that the same metropolitan area could be served by multiple airports (e.g. O'Hare and Midway in Chicago), or that a college town could be half-way between two airports. We chose to match universities to all airports within a 50 miles radius. We code the year of Southwest entry for a pair of universities as the first year in which Southwest operates a flight on any route whose endpoints (airports) are within 50 miles of the respective universities. Results are robust to narrowing this definition further (e.g. 25 miles, 10 miles), see Appendix Table A-1.

***Data on Scientists***. Our focus is on collaborations between faculty members (and



therefore effectively across labs) in the discipline of chemistry<sup>6</sup>, in part due to data availability, and in part because of the short publications cycles in this discipline. For biographical information on scientists, our data source is the directory of graduate research published by the American Chemical Society. Intended as a source of information for prospective graduate students, this directory provides comprehensive listings of faculty affiliated with U.S. departments granting PhDs in chemistry, chemical engineering and biochemistry. Besides faculty names and departmental affiliations, the directory provides information on year of birth, gender and education. The directory is published biannually in print and since 1999 on the web.<sup>7</sup> We combine the directories from 1991 to 2013 to build a longitudinal panel of over 20,000 scientists. We complement this information with department-level R&D expenditures from The National Science Foundation (NSF) Survey of Research and Development Expenditures at Universities and Colleges.

**Publication Data.** We match faculty names to publication data from Scopus covering more than 200 chemistry journals (including all journals from the American Chemical Society), multidisciplinary journals and major journals in neighboring disciplines.<sup>8</sup> Within chemistry, the match between publications and scientists is facilitated by the fact that we know institutional affiliations from the American Chemical Society faculty data. We match publications to faculty based on last name, first and (if non-missing) middle initials, department and university affiliation. From publication data, we construct for each scientist time-varying measures of past productivity (with a moving average over the last three years of publication counts weighted by journal impact factor). We also infer our main outcome, copublications, from bibliometric data combined with faculty data.

A key strength of our data is that we know when individuals enter and exit the profession

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<sup>6</sup>Chemistry, which focuses on the composition, structure, transformations and properties of matter, is a large discipline, with chemistry PhD graduates accounting for around 15% of U.S. PhD life and physical science graduates (NSF 2015).

<sup>7</sup>The American Chemical Society also produced a CD-ROM for the years 1991-1993.

<sup>8</sup>Scopus is one of the two major bibliometric databases (along with ISI Web of Science). Our set of chemistry journals includes all journals from the American Chemical Society, as well as any chemistry journal with an impact factor above 2. Our set of multidisciplinary journals includes *Nature*, *Science*, *Cell* and the *Proceedings of the National Academy of Sciences*. Our set of major journals in neighboring disciplines includes all journals with an impact factor above 6 in physics, biology, material science and nanotechnology.

and therefore are at risk of collaborating with others. If we were inferring copublications from publication data only, we could hardly distinguish between active scholars and individuals that have retired or are not doing research in the field. Papers are counted as a copublication between all pairs of faculty members involved.<sup>9</sup>

***Additional Key Outcomes.*** For part of the analysis, we weight copublications by the citations they receive as a proxy for their impact and quality. Citation counts originate from Scopus, are at the article level and are counted from the year of publication until 2013. We also construct two distinct groups of measures related to novelty using author keywords. These are based on the entire corpus of articles within chemistry journals and related fields. The first group of measures is based on established approaches from the innovation literature (Boudreau, Guinan, Lakhani, and Riedl, 2016; Criscuolo, Dahlander, Grohsjean and Salter, 2017; Azoulay, Gler, Koak, Murciano-Goroff, and Anttila-Hughes, 2012), and relies on calculating the share of keywords in any given paper that have not been observed before. This allows us to capture both novel uses that gain traction, as well as those that do not. To check the robustness of our results, we also experimented with different definitions of what constitutes a novel use (e.g. bottom 5%, 10% or 25% of the keyword-use distribution), as well as with different aggregation methods (mean share of novelty, max share of novelty, total novelty for the focal papers), finding consistent results. We then replicate this approach for subfields to see if a specific use might have been considered novel in aggregate, but not within a smaller community of science.<sup>10</sup>

The second dimension of novelty we explore is what we label as ‘novel trends’. With this measure we are not focused on making sure we capture both failed and successful attempts at developing new concepts (i.e. the variance in outcomes), and instead prioritize identifying

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<sup>9</sup>The majority (75%) of papers matched to a faculty member have exactly one faculty author, 21% percent have two, and less than 4% have more than two authors. Both papers with one faculty author and papers with multiple faculty authors typically have several non-faculty authors. We focus on faculty authors because they are the ones usually making the decision to collaborate. Papers in chemistry journals that are not matched to any of our U.S. faculty authors are likely to be from foreign scientists, scientists working in corporate environments and federal labs.

<sup>10</sup>We are interested in the tension between aggregate and subfield-specific novelty because our theoretical framework predicts that across field collaborations should disproportionately benefit from reductions in travel costs, and want to test if some of these represent arbitrage of ideas between subfields of science.

emerging new trends in science. A drawback from this measure is that it selects on successful cases where scientists work on concepts that end up gaining broader adoption afterwards. The reason why we find this measure interesting is because it proxies for the focal researchers working on topics that were about to become ‘hot’. To do this, for each keyword, we calculate the share of papers in a given year that contains the keyword – a proxy for how popular it is at any point in time. We then calculate the first and second derivative of this measure relative to the previous year. If both the first and second derivative are positive, then the keyword is classified as part of a novel trend since its use is quickly accelerating. Additionally, if the first derivative is zero, and the second is positive, then we are at a local minimum right before a keyword takes off, which we also consider as a novel trend. Aggregating up at the paper level, a publication is considered part of a novel trend if it has an above the median number of novel-trend keywords (results are similar if we impose a higher threshold).

We also constructed proxies for the equipment-intensity of publications by first collecting a large-scale list of keywords associated with chemistry equipment,<sup>11</sup> and then checking this list against the keywords used in each paper. Papers with an above the median number of equipment-related keywords are classified as equipment-intensive (similar results are obtained when using the count of equipment keywords). To test the robustness of our approach to a completely different definition, we also classified areas of chemistry as equipment-intensive versus not using the NSF Survey of Federal Funds for Research and Development. In particular, we first use NSF data to compute the share of departmental R&D expenditures devoted to capital. We then calculate the specialization of each department across fields of chemistry to assess which areas they are specialized in. Last, we use these measures to run a regression at the department level linking department-level capital intensity to the relative prevalence of different subfields of chemistry, and rely on the estimates to classify collaborations based on the type of departments they are originated from.<sup>12</sup>

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<sup>11</sup>This was built by scraping and compiling an inventory of equipment for sale in online catalogues and stores targeted at a wide range of chemistry labs.

<sup>12</sup>The regression yields the following classification: physical chemistry, analytical chemistry, and biochemistry as capital-intensive fields; and organic chemistry, inorganic chemistry, and material science as not capital intensive. Discussions with domain experts as well as anecdotal evidence supports this classification.

## 2.2 Descriptive Statistics for the Main Sample

Our dataset covers over 20,000 scientists and their collaborations. However, we focus on a specific subset of pairs of scientists who experience Southwest entry and for whom we have variation in collaboration over time. Since all regressions include scientist-pair fixed effects, pairs that never collaborate drop out of the sample. In the Appendix, we show that our main result is robust to replacing scientist-pair fixed effects with city-pair fixed effects and including a random sample of non-collaborating pairs. Our results are also robust to replacing pair fixed effects with individual researcher fixed effects.

We have 15,244 pairs of scientists who collaborate at least once.<sup>13</sup> Excluding co-authors that are in the same department, we have 8,311 pairs of scientists in our sample. Only a minority (1,158) of these pairs experience Southwest entry during our analysis period of 1993-2012, either because for the other 7,153 pairs Southwest is already operating a flight, or because Southwest never flies between the relevant endpoints. We drop pairs in locations where Southwest enters but then leaves within two years, as well as pairs where Southwest entry coincides with the move of a scientist.<sup>14</sup> Finally, we also exclude pairs that are within less than 200 miles of each other as air travel is unlikely to be their main travel option.<sup>15</sup> Our final analysis covers 758 pairs of scientists corresponding to 845 individuals.

[Insert Table 1 about here]

Table 1 displays descriptive statistics for our chemistry sample at different levels of analysis: individual, individual-pair and individual-pair-year.<sup>16</sup> Most individuals in the sample are male (90%) with an average age at the time of Southwest entry of 49.6 years. We do not observe individual research budgets but as a proxy we use departmental R&D expenses

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<sup>13</sup>Our dyadic data is not directed, and thus is symmetric: the pair between  $i$  and  $j$  is the mirror image of the pair between  $j$  and  $i$ . The 15,244 figure is after dropping an equal number of symmetric observations.

<sup>14</sup>Scientists in our sample may move from one department to another, in some cases leading to a change in whether they are connected by Southwest or not. We want changes in Southwest status to be driven by Southwest entry decisions rather than by scientist location decisions, and thus exclude pairs who happen to move in the same year as Southwest enters, the year before or the year after.

<sup>15</sup>Results are robust to decreasing this threshold to 100 or 50 miles.

<sup>16</sup>Correlation tables, as well as summary statistics for all samples are in the Appendix.

divided by the number of faculty members in the department. The average in our sample is \$279,880 at the time of Southwest entry. According to the NSF Survey, R&D expenses include compensation for R&D personnel, equipment and indirect costs. In terms of specialization,<sup>17</sup> the largest area is physical chemistry (32%), followed by biochemistry (22%), inorganic chemistry (14%), organic chemistry (13%) and material science (11%).

We observe the 758 pairs for 17 years on average,<sup>18</sup> corresponding to 13,147 observations at the individual-pair-year level. Southwest entry events map to 413 distinct new routes. The median pair experiences Southwest entry in 1999, but we observe Southwest entry from 1994 to 2011. The mean number of copublications over the whole period is 1.9, but the majority of pairs copublishes once. Only 9% of pairs collaborates both before and after Southwest entry.

[Insert Table 2 about here]

It is useful to compare our analysis sample to other distant pairs that do not experience Southwest entry. We have approximately 6,000 such pairs. These include pairs where Southwest is already present in the relevant market prior to 1993 when our sample starts, or has not entered by 2012 when it ends. They also include cases where one of the pair members is a new faculty hired after Southwest has already entered. The comparison is shown in Table 2. The pairs that experience Southwest are not statistically different from the others in terms of publications, but are slightly older (51 versus 49 years) and are observed on average for a slightly longer period of time (17 versus 14 years).<sup>19</sup> Importantly, there is no significant difference in terms of R&D budgets or propensity to be in different subfields of chemistry.

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<sup>17</sup>Specialization is inferred from the journals in which a scientist publishes. For instance, a faculty member who often publishes in the Journal of Biological Chemistry is assumed to be specialized in biochemistry. See Appendix Table A-4.

<sup>18</sup>A pair is in our sample for a maximum of 22 years (from 1991 to 2011). We observe some pairs for less than 22 years due to pair members starting their first faculty appointment after 1991, retiring before 2011, or otherwise no longer being listed in the ACS faculty directory (e.g. because they moved to industry or to a foreign country).

<sup>19</sup>This makes sense since a longer observation period mechanically increases the chances of experiencing Southwest entry.

### 3 Empirical Strategy and Main Results

Our empirical specification is a straightforward difference-in-differences framework at the scientist-pair level where we exploit variation in Southwest entry across different airport pairs over time. It includes scientist-pair fixed effects and is estimated using a Poisson model:

$$Y_{ijt} = \beta \text{AfterSW}_{ijt} + \mu_t + \gamma_{ij} + \epsilon_{ijt}$$

where  $Y_{ijt}$  is the number of copublications between scientist  $i$  and scientist  $j$  in year  $t$ ,  $\text{AfterSW}_{ijt}$  is an indicator variable that takes value 1 after Southwest entry,  $\mu_t$  is a year fixed effect,  $\gamma_{ij}$  is a pair fixed effect to control for unobservable, time-invariant differences between pairs of scientists, and  $\epsilon_{ijt}$  is an idiosyncratic error term.

Our analysis examines the change in the rate of collaboration and in the types of papers that emerge over time for pairs that co-author at least once. Since our unit of analysis is the scientist-pair-year and we include pair fixed effects, our main source of variation is the change in Southwest status for treated pairs, where control pairs are constituted by pairs that never experience entry or will experience it in the future. The pair fixed effects completely capture pairs of scientists for which we never see activity, and thus we remove these from the analysis without empirical consequences. Robust standard errors are clustered at the pair level.

#### 3.1 Southwest Entry and Changes in Passengers, Prices, Miles and Transfers

Before our main analysis, we check how the arrival of Southwest affects some of the key passenger and fare metrics of interest in the air travel industry. In this exercise, we run regressions at the airport-pair level, and compare a number of outcomes before and after Southwest entry. Regressions include airport-pair fixed effects and year fixed effects. The coefficients in Table 3 reflect the types of changes one would expect to take place after the arrival of a low-cost competitor: the increase in the number of passengers is between 54% and 57%<sup>20</sup>, and prices drop by 17% to 19%. We do not find any effect on the average

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<sup>20</sup>The 95% confidence interval for the number of passengers expressed in percentage change is  $[\exp(0.4437 - 1.96 * 0.005) - 1; \exp(0.4437 + 1.96 * 0.005) + 1]$  or  $[0.543; 0.574]$ .

miles flown<sup>21</sup> or on direct flights, and the reduction in the number of transfers is extremely small. Overall, results are consistent with Southwest lowering the cost of air travel without drastically changing the types of routes available or the number of miles passengers have to fly to connect between two endpoints.

[Insert Table 3 about here]

### 3.2 Changes in Collaboration and Evidence for a Causal Interpretation

After a reduction in travel costs, the relative attractiveness of the global pool of potential co-authors should increase, since working with distant collaborators becomes more cost effective. This should lead to an increase in collaboration between the affected locations. As can be seen in Column 1 of Table 4 (which uses our main econometric specification), after Southwest enters we observe a large and significant increase in collaboration between scientists at the connected end points.<sup>22</sup> Relying on a 95% confidence interval, we estimate that scientific collaboration increases between 0.3 and 1.1 times.<sup>23</sup> While the magnitude of the effect is large, it is off a small base (the mean of the dependent variable is approximately 0.1), and comparable with previous studies on the impact of communications, search costs and co-location on scientific collaboration: Agrawal & Goldfarb (2008) find that Bitnet increased the likelihood of collaboration between pairs of universities by 40%; Boudreau et al. (2017) find that a 90-minute structured information sharing session led to a 75% higher probability of co-applying for a grant; Catalini (2017) estimates that exogenous co-location increased the chance of a collaboration between labs on the Jussieu campus of Paris by 3.5 times.

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<sup>21</sup>The data from the Bureau of Public Transportation includes the number of miles flown for each itinerary. Differences in miles flown arise from the number of connections an itinerary involves. We compute average miles flown as the average across all passengers travelling between two airports in a given year.

<sup>22</sup>Collaboration between scientists is increasing over time. In our regressions, this trend is captured by the inclusion of year fixed effects. Therefore, one can interpret our estimates as the relative percentage increase in collaboration due to Southwest entry once the underlying increasing trend in collaboration has been accounted for.

<sup>23</sup>The point estimate is  $\hat{\beta} = 0.505$  and the standard error 0.121. So the lower bound of the 95% confidence interval expressed in percentages is  $(\exp(0.505 - 1.96 * 0.121) - 1)/100=30.6\%$ , and the upper bound is  $(\exp(0.505 + 1.96 * 0.121) - 1)/100=110.2\%$

[Insert Table 4 about here]

One may worry that Southwest entry is systematically correlated with time-varying factors such as growth of the universities (or the regional economies) at both ends of the routes, and therefore that collaboration would have increased even in the absence of a reduction in travel costs. While our main specification already controls for aggregate time trends through year fixed effects, the validity of our results could be threatened by systematic, time-varying factors that affect the target locations around the time of Southwest entry. In Column 2, we mitigate these concerns by controlling for two possible time-varying confounders: the age of the scientist pair, and the (log of) departmental R&D budget per faculty member. The first one accounts for changes in the incentives to collaborate as scientists progress in their careers, the second for changes in the local economies. Whereas the coefficients for the controls are positive and significant, our main result is unaffected. In Column 3, we additionally control for the number of years that have passed since both scientists obtained their PhD, a proxy for their ability to both decide who they want to collaborate with. This estimated coefficient is negative and significant but again does not affect the estimate for *Southwest Entry*. In Column 4 we study the dynamic effects of the reduction in travel costs by replacing the treatment indicator for *Southwest Entry* from Column 1 with a set of four dummy variables capturing the years around the treatment. For example, the indicator *Southwest Entry (-1)* is equal to one if the focal scientist-pair observation is recorded one year prior to the treatment. The other indicator variables are defined analogously with respect to the year of treatment (0), the first year after treatment (1), and two or more years after treatment (2+).<sup>24</sup> The coefficient for *Southwest entry (-1)*, which would capture any ‘effect’ of the new airline routes before their introduction, is insignificant, suggesting that there is no collaboration pre-trend in the data, i.e. it is only once travel costs are reduced that the coefficients turn positive and statistically significant.

[Insert Figure 1 about here]

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<sup>24</sup>We adopt this particular specification because it is the same used by Bernstein, Giroud, and Townsend (2015) in their study of venture capital monitoring and air travel costs, but show robustness to specifications with additional years in Figure 1.



A graphical version of a similar exercise with a full set of coefficient estimates for the 5 years before and 5 years after Southwest entry is displayed in Figure 1. There is again no collaboration pre-trend before Southwest launches a route, and it is only after the new route is available that the estimated coefficients are positive and steadily increasing in magnitude.<sup>25</sup> It is useful to highlight that publication lags in chemistry are substantially shorter than in the social sciences: when studying the 10 major analytical chemistry journals (1985-1999), Diospatonyi et al. (2001) find median lags between submission and publication of 3 to 10 months, with some journals publishing papers within 2 months of first submission.

In Column 5 of Table 4, we conduct a placebo test where we randomly allocate Southwest entry events to scientist pairs. The coefficient for *'Fake Southwest Entry'* is not significant and close to zero, suggesting that it is not the structure of the panel or changes in the data over time that are driving the result. In Column 6 of Table 4, we conduct one more falsification test by looking at entry events (not included in the other regressions) where Southwest withdraws from the market within two years. For these cases, the point estimate of Southwest entry is close to zero and insignificant.<sup>26</sup>

Overall, we believe results in Table 4 and Figure 1 provide robust support for a causal interpretation of our main effect, and reassure us that we are not simply measuring some underlying, unobservable process that takes place with each entry event<sup>27</sup> and drives both Southwest decisions and the increase in scientific collaboration.

While Southwest is the largest U.S. low-cost carrier in terms of number of passengers transported, there are other low cost airlines operating within the same market. In Appendix Table A-12, we explore how our results vary depending on whether a low-cost airline is already operating on a route, as well as whether they differ when other airlines (low-cost

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<sup>25</sup>We repeat the same graph within the large sample at the CBSA-pair level in Figure A-1.

<sup>26</sup>While it may seem counterintuitive that early withdrawals are not associated with an effect, but that we also obtain a positive estimate for *Southwest Entry (0)* in Column 4 of Table 4, it is important to highlight that: a) 98% of early withdrawals occur in the year of entry; b) when we separate entry events by quarter of the year, we only find a positive effect for the year of entry when Southwest starts serving a route in the first or second quarter of the year; c) the confidence interval for *Southwest Entry (0)* in Column 4, and for early withdrawals in Column 6 overlap - so in a statistical sense, we cannot rule out the possibility of a positive effect for these short spells associated with a withdrawal, even if the estimate is noisy.

<sup>27</sup>We observe Southwest arrival across multiple locations and years, which makes it unlikely that some other simultaneous event is always co-occurring with the shocks we use.

or not) start operating a flight in the same year as Southwest. Consistent with the impact of Southwest on travel costs being largest when no low-cost alternatives existed on the same route, estimates are larger when Southwest is the first low-cost to enter (Table A-12, Column 2),<sup>28</sup> and are positive, but non-significant when another low-cost was already operating between the same airports (Column 3). Results are instead essentially unchanged if we exclude cases where other low-cost airlines enter at the same time (Column 4), other major airlines<sup>29</sup> enter at the same time (Column 5), or any other airline enters at the same time (Column 6). We conclude that our results are robust to considering concurrent entry by other airlines.<sup>30</sup>

Results are also not driven by the fact that our sample includes only pairs that ever collaborate: when we include a random sample of non-collaborating pairs and replace individual-pair fixed effects with university-pair fixed effects<sup>31</sup>, we find comparable effects of Southwest entry (see Table A-3). In Appendix Table A-14 we decompose the main effect by pairs of scientists who collaborate both before and after Southwest entry (intensive margin pairs) versus pairs of scientists who collaborate either before or after entry, but not both (extensive margin pairs).<sup>32</sup> We find a stronger effect for intensive margin pairs (Column 3), although the cheaper fares also seem to enable experimentation in the form of new collaborations over distance (Column 2).<sup>33</sup>

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<sup>28</sup>Our list of low-cost airlines includes AirTran Airways Corporation, JetBlue, Frontier Airlines, Spirit Air Lines, ATA Airlines, Allegiant Air, Virgin America, Sun Country Airlines, ValuJet Airlines and Vanguard Airlines.

<sup>29</sup>We classify as major airlines: Delta, American Airlines, United Airlines, US Airways, Northwest Airlines, Continental, America West Airlines, Alaska Airlines, Trans World Airlines and Envoy Air. These correspond to the 10 companies with the largest numbers of passengers carried between 1993 and 2012.

<sup>30</sup>One might also wonder about additional modes of transportation. As shown in Appendix Table A-13, we find no effect of Southwest entry in the Northeast corridor, where train travel has been a consistent alternative to flying.

<sup>31</sup>If we were to run this regression with individual-pair fixed effects, the non-collaborating pairs would be dropped from the estimation.

<sup>32</sup>It is useful to highlight that since our unit of analysis is a pair-year and we observe the extensive margin pairs both before and after entry, the Southwest entry variable is not absorbed by the pair fixed effects and does have variation within these pairs. Some of these extensive margin pairs do not collaborate before the event, and others do not collaborate after, so the estimated effect is a composition of the behavior of both types of pairs. Since the estimate is positive and significant, we infer that on average Southwest entry is associated with more non-previously-collaborating pairs engaging in collaboration than the other way around.

<sup>33</sup>We also investigate whether the stronger effect for the intensive margin pairs is driven by continued

In the Appendix, we perform additional robustness to different econometric approaches, functional forms, clustering of standard errors, treatment of outliers and inclusion in the sample of non-collaborating pairs. Briefly, we obtain qualitatively and quantitatively similar results using ordinary least squares instead of Poisson (see Table A-16, Column 2). We also obtain a positive and significant coefficient for Southwest entry (though of a somewhat smaller magnitude) when we run a linear probability model with an indicator variable for any copublication in the focal year as the dependent variable (see Table A-16, Column 3). Clustering at the city-pair level, rather than at the individual-pair level, hardly impacts the standard errors (see Column 3 of Table A-2). The coefficient on Southwest entry remains significant when we exclude pairs that have more than two copublications over the entire observation period, or winsorize observations with more than two copublications (see Table A-17).

### 3.3 Extensions and External Validity in Different Samples

The analysis and results presented in the previous section describe the effect of Southwest entry on the rate of collaboration between chemistry faculty members. While this approach has the advantage of leveraging rich individual-level data and offers a cleaner identification strategy, one may also be interested in replicating the analysis within a field with slightly different characteristics, as well as testing external validity within a broader set of fields. To do so, we first perform a deep-dive within mathematics (a field for which we have also collected individual-level data), and then explore regressions at the region-pair level for biology, physics and engineering. Results show that the effect we have identified within chemistry is also present across these samples.

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collaboration between former doctoral students or postdocs, and their advisors and sponsoring principal investigators. To code these advisor-advisee pairs, we take advantage of the ordering convention for author names in the field (juniors are typically listed first and principal investigators last). Results are reported in Table A-15, and are suggestive of stronger effects among these special pairs, but the sample size is too small to reach a conclusion.

### 3.3.1 Increases in the Rate of Collaboration within Mathematics

The dataset we use for mathematics includes all U.S. faculty members that have advised at least one PhD student.<sup>34</sup> We observe 431 pairs of individuals that experienced Southwest entry between 1993 and 2012 and have at least one copublication in that period. We adopt the same empirical strategy as in the chemistry sample and regress copublications on an indicator variable for Southwest entry, controlling for pair fixed effects and year fixed effects. Results show that Southwest entry significantly increases copublications in mathematics too.

[Insert Table 5 about here]

### 3.3.2 Increases in the Rate of Collaboration Across Regions

To test if the availability of cheaper flights had an effect on scientific collaboration across a broader set of fields, we also use a large-scale publication dataset covering close to a million papers matched to U.S. regions (defined in terms of CBSAs – Core-Based Statistical Areas).<sup>35</sup> Specifically, we explore how collaboration between any two CBSAs changed after Southwest starts operating a new route between them. The unit of analysis is the CBSA-pair-year (48,274 pairs), and we include CBSA-pair fixed effects and year fixed effects to respectively control for underlying differences across regions that are consistent over time, and overall time trend.<sup>36</sup> The regressions also include linear time trends for the origin and destination

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<sup>34</sup>This database is based on MathSciNet, an abstracting service run by the American Mathematical Project and the Mathematics Genealogy Project, which is targeted at tracking PhD theses in mathematics. We construct a sample of US-based mathematicians who advise at least one PhD student, and deduce their location from the institution their students graduate from.

<sup>35</sup>The starting point for the construction of this sample is the population of scientific articles published in the top 477 scientific journals in biology, chemistry, physics and engineering between 1991 and 2012. We have a total of 2,773,560 papers, of which 1,169,458 have at least one author with a U.S. address. Out of all papers with U.S. addresses, we are able to successfully map 994,672 (85%) to a U.S. CBSA using a combination of three different geocoding services (Google Maps API, Bing Maps API, and the Data Science Toolkit). This allows us to link the vast majority of U.S. papers to the geographic regions involved in their production.

<sup>36</sup>While this approach has the advantage of considering different fields of science, it also has important limitations. We can no longer include scientist-pair fixed effects and account for idiosyncratic, unobservable, and time invariant reasons that may drive collaboration between any two scientists. Core-based statistical areas (CBSAs) may also be too large as a unit of analysis for correctly measuring the effects of interest. Finally, our ability to test the full set of predictions of the model is limited.

CBSA. For the estimation, we use a Poisson model with standard errors clustered at the CBSA-pair level.

[Insert Table 6 about here]

Results are displayed in Table 6: the point estimates for Southwest entry at this more aggregated level of analysis are significant not just in chemistry but also in biology, physics and engineering. While the estimated coefficients for chemistry, physics and engineering are not statistically different from each other, the difference between chemistry and biology is significant.

Overall, we conclude that the results from the chemistry sample are generalizable to other fields, and that the increase in collaboration is larger within biology. We now turn to developing a simple model to place our main finding into the broader context of how geographic frictions shape collaboration, and to guide the empirical exploration of additional predictions.

## 4 Theoretical Framework

The objective of this section is to develop a simple theoretical framework to highlight key trade-offs scientists face when deciding if they should collaborate with a local or a distant co-author, and how much effort they should dedicate to a collaboration based on its intrinsic potential. The model generates novel predictions about how travel costs shape collaboration decisions, which we then test using our data.

We start by assuming that because the global pool of potential co-authors offers more variety than the local one, it is on average possible to find better matches when team formation is not constrained by geographic distance. The quality of a match may depend on complementary ideas, knowledge, skills, equipment, and resources that a co-author brings to a project. Of course, because of agglomeration forces, as the size, specialization and quality of a region’s local pool increases, scientists will rely less on distant co-authors. To account for this, in an extension of the baseline model we allow for the share of ‘first best’ co-authors

available locally to vary.<sup>37</sup>

Our setup is straightforward: ideas are born with intrinsic quality  $q$ , but require effort  $e$  to be developed and achieve their full potential  $v$ . Since scientists observe a noisy signal of  $q$  before starting a project, they will allocate more effort, time and resources to projects that have higher potential (i.e. in our model, effort is endogenous to potential). At the same time, since research constitutes an uncertain endeavour, even when scientists apply effort projects are only successful with probability  $p$ , which depends on the quality of the co-author match. Thus, the realized value of a project can be expressed as  $v = p_i q e$ , where  $p_i$  (with  $i = G, B$ ) is higher when a good match between co-authors is achieved ( $p_G$ ), relative to a bad match ( $p_B$ ). Although a two-sided matching framework would be more realistic, in the model we abstract away from a setup where collaboration decisions are influenced by both sides. Our approach follows a partial equilibrium model in which the pool of potential applicants always accepts a collaboration when invited, and where proposers invite potential co-authors only if they know the project is a fit for them and an interesting one for them to pursue.

Whereas the global pool may offer a better match between co-authors (i.e.  $p_G$ ) and increase the chances of realizing a project's full potential  $v$ , collaborating over distance introduces additional costs, as scientists have to travel for face-to-face interactions, and may be less effective at communicating complex information remotely. As a result, scientists face a trade-off between less choice locally, and increased communication and travel costs over distance.

It is important to highlight that the model is not focused on the decision to collaborate or not (see for example Bikard, Murray & Gans, 2015), nor on the type of project to pursue (this is discovered by the scientist at the start), but is explicitly centered on a situation where a scientist is looking for the best co-author for a particular idea. Since we cannot empirically measure search behavior and search frictions, the model also abstracts away from

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<sup>37</sup>The fraction of first best co-authors in the global pool is assumed to be  $z$ . Since the global pool can be seen as an average over all possible local pools, the fraction of first best co-authors in a given local pool  $w$  can be either higher, lower, or equal to  $z$ . If  $w > z$ , then scientists will never collaborate over distance, as they would incur additional costs but would not be more likely to find an ideal co-author over distance. Therefore, the range of values of  $w$  that provides a meaningful trade-off is  $0 \leq w < z$ . To simplify the exposition, in the paper we will assume  $w = 0$ . More general cases are discussed in the Appendix.

search costs, and assumes that the broader talent pool a scientist is considering is formed by all the individuals a focal researcher is already aware of, has previously met at a conference, has been colocated with, or has read work by. Boudreau et al. (2017) find that search costs constitute a key friction to collaboration even within the same institution, so we make the simplifying assumption of these costs being present both for local and distant collaboration decisions. In the model, we also abstract away from scientists' budgets: although in the regressions we are able to use departmental R&D budget data to explore heterogeneous effects, in the theory we do not account for the fact that scientists at better funded institutions, or more productive scientists in general may have access to larger budgets and may be therefore less sensitive to changes in travel costs. If that were the case, then the reduction in travel costs could disproportionately help lower productivity researchers. We also do not model endogenous time to project completion as a function of travel intensity, or team dynamics beyond two authors. Nevertheless, the stylized framework allows us to obtain several additional predictions that we then take to the data.

In the next sections, we perform comparative statics and explore the main tensions of the model in more detail.

## 4.1 Local versus Distant Collaborations

The scientist's payoff from developing an idea with a local co-author for a given level of effort  $e$  is:

$$\pi_L(e) = p_B q e - c(e) \tag{1}$$

where  $c(e)$  is the cost of effort which we assume for tractability to have the following convex function:  $c(e) = \frac{\alpha}{2} e^2$ . Thus, equation (1) can be re-written as:

$$\pi_L(e) = p_B q e - \frac{\alpha}{2} e^2 \tag{2}$$

The first order condition yields an optimal effort level of  $e_L^* = \frac{p_B q}{\alpha}$ , which is increasing both in project quality  $q$ , and in the quality of the co-author match  $p_B$ . Intuitively, scientists are more willing to apply effort to projects with higher potential, and to projects they are

working on with better matched co-authors. Inserting  $e_L^*$  back into (2), we obtain a scientist's payoff for a local collaboration given the optimal effort level as:

$$\pi_L^* = \frac{(p_B q)^2}{2\alpha}. \quad (3)$$

How does this compare to a distant collaboration? In our setup, over distance, scientists have a higher chance of securing the ideal co-author because the global pool offers more variety. At the same time, this does not happen all the time, and scientists have to incur additional communication and travel costs  $t_i$  to develop a project over distance. We assume that with probability  $z$  scientists find a first best co-author and secure  $p_G$ , and with probability  $(1 - z)$  they land a co-author of the exact same level they would have found in the local pool  $p_B$ . Thus, the payoff for a distant collaboration can be written as:

$$\pi_D(e, t) = (1 - z)[p_B q e_B - \alpha \frac{e_B^2}{1 + t_B} - \beta t_B^2] + z[p_G q e_G - \alpha \frac{e_G^2}{1 + t_G} - \beta t_G^2], \quad (4)$$

where  $e_i$  and  $t_i$  with  $i = G, B$  are the optimally chosen levels of effort and travel for perfectly matched co-authors ( $p_G$ ) versus imperfectly matched ones ( $p_B$ ).

Traveling enters as a convex cost<sup>38</sup> ( $t_i = [0, 1]$ , scaled by a parameter  $\beta$ )<sup>39</sup>, but also increases the chances of success because it improves the ability to communicate complex information, coordinate work and make progress on a project through face-to-face interactions. This trade-off allows for interesting cases to emerge where temporary co-location between distant co-authors is expensive but also helpful, and can therefore lead to both higher and lower payoffs relative to a collaboration on the same project with a local co-author.

The basic dynamic we want to capture is one in which collaboration and communication require less effort when scientists are colocated, but where travel can also be strategically used to recreate the same efficiencies experienced in local collaborations. When communicating over email or phone co-authors may need more time and effort to convey the same concepts

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<sup>38</sup>Longer air travel incurs additional costs, including a longer time in the air, additional transfers, the inability to perform the trip within a day, accommodation costs, time zones and fatigue. In the data, fare prices also increase more than proportionally with distance (i.e. when we estimate  $TicketPrice = b_0 + b_1 * Distance + b_2 * Distance^2$ , we systematically obtain a coefficient  $b_2 > 0$ ).

<sup>39</sup>If  $t_i = 1$ , face-to-face communication is always available (as with a local collaborator), and the cost of effort would be the same under both scenarios. Advancements in communication technology and virtual reality can be therefore thought of as changes in  $t_i$ .



and avoid misunderstanding, and when in person meetings are infrequent it may take more time for a team to get everyone up to speed and make progress. In the model, distant co-authors can either spend more effort and communicate over distance, or invest in travel and rely on more effective face-to-face interactions.<sup>40</sup>

For simplicity, we assume that once a local versus distant co-author has been chosen for a project, it is too costly to switch type without starting a completely new project.<sup>41</sup> We also assume that before a substantial amount of effort and travel is dedicated to a project, the quality of a co-author match has been revealed. The first order conditions with respect to effort and traveling are, respectively:

$$e_D(t_D) = \frac{p_i q (1 + t_D)}{2\alpha}$$

$$\alpha \frac{e_D^2}{(1 + t_D)^2} - 2\beta t_D = 0,$$

where  $i = B, G$  depends on whether the distant co-author has led to a first best or a second best match. Combining both first order conditions, one can show that the optimal levels of travel and effort for a distant collaboration are:

$$t_i^* = \frac{(p_i q)^2}{8\alpha\beta} \tag{5}$$

$$e_i^* = \frac{p_i q}{2\alpha} (1 + t_i^*) = \frac{p_i q}{2\alpha} \left(1 + \frac{(p_i q)^2}{8\alpha\beta}\right) \tag{6}$$

If we plug these back into the payoff function we obtain:

$$\pi_i^* = \frac{(p_i q)^2}{4\alpha} \left(1 + \frac{1}{2} t_i^*\right) = \frac{(p_i q)^2}{4\alpha} \left(1 + \frac{(p_i q)^2}{16\alpha\beta}\right) \tag{7}$$

Thus, the overall payoff over distance is:

$$\pi_D^* = (1 - z)\pi_B^* + z\pi_G^* \tag{8}$$

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<sup>40</sup>Furthermore, since as co-authors invest more in travel and in-person meetings one could imagine the two effort functions should look more similar, in our model as  $t$  converges to its upper bound ( $t = 1$ ),  $e_i^2/(1 + t_i)$  converges to the cost local co-authors face,  $e_i^2/2$ . I.e., one can think of the effort cost under co-location to be a particular case of a distant collaboration facing the minimum possible cost of effort.

<sup>41</sup>One intuitive way to think about changing a co-author within our simple framework is to imagine the original project failing, and a new one being launched with a different team.

Comparing payoff equation (3) for local collaborations with equation (8) for distant ones is informative independent of travel costs (which we discuss in detail in the next section). For example, it allows us to explore how the relative appeal of a local versus a distant collaboration changes as the comparative advantage of the global pool ( $z$ ) over the local one varies:

$$\frac{\partial[\pi_D^* - \pi_L^*]}{\partial z} = \pi_G^* - \pi_B^* > 0 \quad (9)$$

Intuitively, an increase in the likelihood of finding a first best co-author in the global pool will lead to a relative increase in the payoff for distant over local collaborations. Similarly, if scientists enjoy a high quality local environment with good matches (e.g. they are in an agglomerated research cluster), they will find limited benefits from collaborating over distance.

Until now, we have assumed that all scientific projects have the same probability of failure. At the same time, novel, exploratory and cross-disciplinary projects are more likely to fail relative to incremental research or work that does not attempt to recombine knowledge across different disciplines (National Academies, 2004; Wang, J., Veugelers, R., Stephan, P., 2017). To account for this we introduce  $\gamma$ , and link it to the overall probability of success through  $p_G = (1 + \gamma)p_B$ . What we want to capture with  $\gamma$  is a tension between exploitation (low  $\gamma$ ) and exploration (high  $\gamma$ ). Exploratory projects, whether because they are novel or because they bring together disciplines that rarely interact with each other, are more likely to fail, and benefit disproportionately from finding the right co-author. The intuition here is that for cross-disciplinary research a scientist needs to find the exact specialist to pursue the project with,<sup>42</sup> and similarly when novelty is high, the returns from working with a better co-author are also higher. Novel projects are very likely to fail to begin with, and may be particularly sensitive to the weakest member of a team, as discussed in Kremer's (1993) O-Ring theory. Whereas in the baseline model we discuss novelty and across-field specialization together – as they share many similarities and both fit under the broader framework of exploratory versus exploitative research – in the extensions we separate the

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<sup>42</sup>For example one that is able to understand and interpret the contributions and language from another discipline.

two constructs further by incorporating uncertainty and higher variance about the potential states of the world in Appendix D, and by modeling co-author specialization directly in Appendix E. Since the implications are similar, in the paper we focus on the simplified implementation based on  $\gamma$ .

For a given  $p_B$ , a low  $\gamma$  means that the quality of the match between co-authors will have a minor influence on the chances of realizing a project's full potential. Low  $\gamma$ , exploitation projects are therefore relatively more straightforward research where most of the techniques and ideas are established (or everyone has access to similar infrastructure to work on them), and the gap between working with the best possible co-author versus anyone else is small. When  $\gamma$  is low, the relative appeal of the global talent pool is more limited. For high  $\gamma$ , exploration projects, instead, scientists will be more willing to travel to work with the ideal co-author and increase their chances of success:

$$\frac{\partial[\pi_D^* - \pi_L^*]}{\partial\gamma} = z \frac{\partial[\pi_G^*]}{\partial\gamma} > 0 \quad (10)$$

An extreme example of this from a specialization perspective is a project for which there are only a few leading experts or key labs with the right equipment (e.g. CERN, LIGO etc.), and the difference between working with them relative to working with a local alternative is large.

Last, when comparing local versus distant collaborations, it is useful to point out that increases in the underlying, intrinsic project quality ( $q$ ) have an ambiguous effect on the choice of co-author type. As shown in the Appendix, which type of collaborations prevail still depends on the basic trade-off between the quality of the match between scientists and travel costs (since a distant collaboration can still leave a scientist with a match of similar quality to the local alternative).

## 4.2 Reductions in Travel Costs and Changes in the Types of Collaborations

How does a reduction in travel costs affect the types of collaborations scientists engage in? In this section, we perform comparative statics to see how cheaper fares like the ones brought by

a low-cost airline change the relative attractiveness of local versus distant collaborations, and how this effect varies for projects of different types (higher versus lower potential, novelty, or interdisciplinarity, etc.). To simplify the notation and exposition, we define  $\theta = \frac{1}{\beta}$  (which is the inverse of travel costs) as the “ease of travel”. One can think of an improvement in  $\theta$  as better infrastructure that allows scientists to meet with their distant co-authors at a lower cost and with lower frictions. The derivative of relative returns with respect to  $\theta$  is:

$$\frac{\partial[\pi_D^* - \pi_L^*]}{\partial\theta} = \left(\frac{q^2}{8\alpha}\right)^2[(1-z)p_B^4 + zp_G^4] > 0 \quad (11)$$

$\theta$  does not matter for the returns to local collaborations ( $\pi_L^*$ ) as no travel is required, but it makes face-to-face interactions with distant co-authors less expensive. Therefore, it is intuitive that with better travel technology the relative attractiveness of the global talent pool increases,<sup>43</sup> as accessing it is now more cost effective.

But how does this effect vary with the ex-ante relative competitiveness of the local pool? I.e., how does this vary for regions that offer better versus worse alternatives to begin with? Remember that this is captured in our framework by the share of first best co-authors that are in the global pool  $z$ . Taking the first order condition with respect to  $\theta$  and  $z$  we obtain:

$$\frac{\partial[\pi_D^* - \pi_L^*]}{\partial\theta\partial z} = \left(\frac{q^2}{8\alpha}\right)^2(p_G^4 - p_B^4) > 0 \quad (12)$$

Which leads to the following prediction:

**Prediction 1:** *A reduction in travel costs will be especially beneficial for researchers that have access, ex-ante, to a relatively worse pool of local co-authors.*

Empirically, if highly productive researchers embedded in worse local environments start substituting local collaborations with better matched ones over distance, we should see evidence of crowding out behavior.

If we take the derivative of relative returns with respect to quality too:

$$\frac{\partial[\pi_D^* - \pi_L^*]}{\partial\theta\partial q} = q\left(\frac{q^2}{4\alpha}\right)^2[(1-z)p_B^4 + zp_G^4] > 0 \quad (13)$$

we see that after an improvement in the ease of travel, higher quality projects are more likely to be undertaken with better matched co-authors (which are more abundant over distance).

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<sup>43</sup>Notice that this holds for the general case of  $0 \leq w < z$ , and is not limited to cases where  $w = 0$ .

The intuition here is that as travel costs fall, scientists are more likely to travel to match with a better co-author. This is disproportionately valuable when the returns to travel and effort on a project are high to begin with, i.e. for ideas of high potential. This leads us to our second prediction:

**Prediction 2:** *A reduction in travel costs will be especially beneficial for distant collaborations on higher quality projects.*

Last, if we do not assume that all projects have the same probability of failure and separate exploratory from exploitative projects by introducing  $\gamma$ , we obtain:

$$\frac{\partial[\pi_D^* - \pi_L^*]}{\partial\theta\partial\gamma} = z\left(\frac{q^2}{4\alpha}\right)^2(1 + \gamma)^3 p_B^4 > 0 \quad (14)$$

which shows that a reduction in travel costs makes the global pool disproportionately more appealing for exploratory projects (high  $\gamma$ ),<sup>44</sup> or restated:

**Prediction 3:** *A reduction in travel costs will be especially beneficial for distant collaboration on novel or cross-disciplinary projects.*

Empirically, to proxy for  $\gamma$  we will rely on how novel the keywords used by the authors on a focal paper are, as well as explore results for collaborations that span different sub-fields of chemistry versus not. We now return to the data to test these predictions.

## 5 Testing the Theoretical Framework

### 5.1 Types of Scientists Affected and Crowding Out

Having shown evidence in Section 3 that Southwest entry led to a plausibly causal increase in collaboration between the affected scientists, we now take advantage of this source of exogenous variation to test the additional predictions from the theoretical framework.

The first prediction of the model focuses on the impact travel costs have on scientists embedded within better versus worse local research environments. Intuitively, agglomerated regions with a greater number of potential collaborators offer on average better local matches to begin with, which makes the global scientist pool relatively less appealing. Since it is

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<sup>44</sup>Intuitively, in our model this is a result of the complementarity between  $\gamma$  and the quality of a co-author match.

difficult to build accurate proxies for the number of ideal co-authors a specific scientist may have access to without traveling, we rely on past productivity to assess if a scientist from a given department is more versus less likely to find a good match locally.

As can be seen in Panel A of Table 7 (Columns 1 to 3), the increase in collaboration we observe in chemistry after the arrival of Southwest is driven by scientist pairs where at least one member is more productive than her local peers, and is even more pronounced when both scientists are more productive than their colleagues. In the mathematics sample, where we only have a small number of observations in Column 1, the effect is positive and significant only for pairs that are both more productive than their local peers (Column 3), possibly because distant collaborations are more rare and selected in this field to begin with.

Overall, the cheaper fares seem to be particularly helpful for individuals that are talented, but potentially do not have access to co-authors of comparable quality within their local environment. They might be in peripheral institutions because of imperfections in the labor market, or simply because of their geographic preferences. With lower travel costs, these individuals are able to find and sustain better matches over distance.

[Insert Table 7 about here]

As mentioned in the theoretical framework, a natural consequence of highly productive scientists prioritizing distant co-authors in their collaboration portfolio because of the lower fares is a crowding out effect on local collaborations. In Table 8, we explore if the cheaper fares have a negative impact on the local collaboration environment. While local copublications are slightly increasing (Column 1), this result is really a composition of two different, countervailing effects. On the one hand, less productive pairs seem to be working together more with each other (Column 3). On the other hand, we see a sharp decline in collaborations between above average productivity scientists and their local, below average productivity peers (Column 2).<sup>45</sup> This makes sense as higher productivity individuals are also the ones that respond the most to Southwest entry to begin with. Interestingly, we find

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<sup>45</sup> Additionally, when we consider the quality of the local collaborations of these above average productivity scientists and their local, below average productivity peers, we find that it goes down following Southwest entry (Appendix Table A-18 column 1), although the effect on novelty is insignificant (Column 2).

this crowding out pattern both in the chemistry and in the mathematics sample, suggesting that when better options become available over distance, highly productive scientists substitute local collaborations with potentially better matched ones over distance.

[Insert Table 8 about here]

In Appendix Table A-19 we present additional splits of the data beyond those predicted by the theory. The effect of Southwest is stronger for younger scientists (Panel A), and scientists that are more distant from each other (Panel C). We do not find a statistically significant difference in effect size by departmental R&D budget (Panel B), even though the estimate for departments with low budgets is almost twice as large as the other ones. Last, pairs where one or both scientists are female do not respond to lower travel costs, possibly because women may have more constrained travel schedules.

## 5.2 Changes in the Type of Projects

The next set of predictions of the model link the reduction in geographic frictions to an increase in the amount of time and effort allocated to higher quality (Prediction 2) and more novel or cross-disciplinary projects with distant co-authors (Prediction 3). As discussed in Section 2.1, we proxy for the quality of projects using citations, and for high  $\gamma$  projects by looking both at projects that span different sub-fields, as well as research that uses novel keywords or belongs to an emerging novel trend. If exploratory projects (high  $\gamma$ ) are more likely to fail, then our estimate will likely underestimate the full impact of a reduction in travel costs on this set of projects, as many will be abandoned and never turn into a publication to begin with.

In terms of project quality, in Column 1 of Table 9 we condition on collaboration and weight the dependent variable, copublications, by citations received (a proxy for scientific impact and quality). Consistent with Prediction 2 and with the idea that lower travel costs induce scientists to allocate disproportionately more effort to distant collaborations as quality increases, we observe a larger effect of Southwest entry on right tail projects.

In terms of interdisciplinarity, in Panel A, Columns 4 to 5 of Table 7 we see that after Southwest enters, collaborations between scientists specialized in different sub-fields of chemistry increase disproportionately relative to other types of collaborations.<sup>46</sup> These interdisciplinary projects may benefit more from face-to-face interactions because of a greater need to exchange complex information which may be new for at least one of the participants, or because these pairs cannot rely on a shared, discipline-specific vocabulary to streamline communications over distance.

Beyond complementarities in ideas and knowledge, cross-disciplinary work between specialized labs can also be captured through complementarities in equipment and infrastructure. In Columns 7 and 8 of Table 9, the dependent variable is respectively the number of equipment-intensive copublications (Equipment 1), and the count of equipment-related keywords used in the focal papers (Equipment 2).<sup>47</sup> Although these types of collaborations are more rare, the effect of Southwest entry is large and significant, suggesting that at least within chemistry specialization driven by equipment may play a key role in how scientists select into distant collaborations. In Appendix Table A-20, we perform a similar regression without relying on equipment-related keywords, but taking advantage of data on capital intensity by department: results show that the lower fares have the largest effect on collaborations where one of the scientists belongs to a capital-intensive group, and the other one does not.

Last, in Columns 2 to 6 of Table 9 we directly look at different measures of novelty. Results are consistent across the dependent variables and provide further support for Prediction 3. After Southwest enters, we see an increase in collaborations that focus on emerging novel trends and topics that are about to become ‘hot’ (Column 2), as well as an increase in the use of novel keywords (Columns 3 to 5). Interestingly, when we define novelty within the more narrow confines of a sub-domain (Column 6), the result is insignificant, possibly because some of the novel uses from Columns 2 to 5 may represent ideas that are being slowly

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<sup>46</sup>A result that we do not find in mathematics, possibly because some of the cross-disciplinary collaborations within chemistry may be driven by access to specialized equipment.

<sup>47</sup>See Section 2.1 for additional details.



incubated within a sub-domain, but have not diffused more broadly yet.

[Insert Table 9 about here]

Next, we study how Southwest entry changes the types of collaborations that are pursued at the regional level. The analysis of collaborations at the dyadic level between chemistry faculty members already suggested that lower travel costs are particularly beneficial for higher-quality, interdisciplinary, equipment-intensive and more novel projects. However, the estimates are significant but noisy because of the smaller sample size when looking at these rare outcomes. We therefore replicate our analysis within a broader set of papers in chemistry and related fields. The analysis is at the CBSA-pair-year level and includes CBSA-pair fixed effects and year fixed effects.

[Insert Table 10 about here]

The effects (see Table 10) are consistent with our previous findings, and highlight that lower travel costs have a disproportionate effect on the right tail of the quality distribution, and on more novel, cross-disciplinary, and equipment-intensive projects. The impact of these changes is large, with increases in aggregate output between 9% and 40% (Column 1). This corresponds to roughly 300 extra copublications per year.<sup>48</sup> We also find support for the other predictions. Novel ideas increase between 15% to 80% (Column 4), but the estimate is noisy (as in Table 9) when we estimate novelty within sub-fields, and equipment-intensive collaborations increase between 5% and 50%.

Our analyses so far have focused on how reductions in travel costs induced by Southwest entry affect pair-level outcomes. Our model, however, also predicts that higher impact projects should mostly result from distant rather than local collaborations. To test this implication, we focus on individual researchers and change the unit of analysis to a paper, which we flag as local or distant.<sup>49</sup> We then regress the number of citations a paper receives

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<sup>48</sup>The sample mean for copublications at the CBSA-pair-year level is around 2.1, leading to an increase of 0.42 per CBSA-pair-year. We have around 2,100 pairs per year, of which around one third are in treatment status. So a back of envelope estimate is  $0.42 \times 2100 \times 0.33 = 294$ .

<sup>49</sup>Papers not involving a collaboration with another faculty member are excluded

on an indicator variable for whether the paper resulted from a distant collaboration (the omitted category being a local collaboration), controlling for year fixed effects and scientists fixed effects. We find that long-distance collaborations get 6% to 7% more cites than local ones in chemistry, and 21% to 24% more in mathematics (Table 11). These results are consistent with the often reported stylized fact that distant collaborations are more heavily cited (Jones, Uzzi & Wuchty 2008).

[Insert Table 11 about here]

## 6 Conclusions

The paper explores how geographic frictions, and in particular travel costs, shape the rate and direction of scientific research. Whereas previous work has mostly focused on communication costs and their impact on the rate of collaboration, our paper emphasizes other effects distance-related frictions can have on innovative outcomes, including the type of projects that are pursued with local versus distant teams.

While both Gaspar & Glaeser (1998) and Kim, Morse & Zingales (2008) have suggested that the secular decline in air travel costs might have led to an increase in scientific collaborations, they do not take their prediction to the data, making this the first study to do so. We build on a vibrant literature that has looked at how scientists respond to reductions in communication costs, and how changes in the infrastructure of collaboration can counterbalance preexisting geographic frictions (Agrawal & Goldfarb 2008, Ding et al. 2010). In particular, our finding that highly productive scientists that are embedded in worse local environments disproportionately benefit from reductions in travel costs is complementary to Agrawal & Goldfarb's (2008) result that reductions in communication costs allow for better matches between top tier and middle tier institutions from the same region. It is also consistent with Ding et al.'s (2010) finding that lower communication costs help scientists from non-elite institutions. Relative to reductions in communication costs, which Ding et al. (2010) show have a positive effect on women scientists, in our setting lower travel costs only help men –

possibly because women scientists have more constraints on their travel schedules.<sup>50</sup>

The paper also extends work that studied the effect of geographic frictions at a much smaller scale, as it provides insights on how easier access to better distant collaborators influences local collaboration decisions. While studies at the microgeographic level have shown that co-location influences the probability and quality of collaboration (Catalini, 2017), we show that additional, and at times opposing forces may be at work at a larger scale through travel costs. Our findings also call for more research on the exact form search costs take when scientists explore collaborations with local versus distant co-authors: while our model abstracts away from these frictions, Boudreau et al. (2017) show that they are a major obstacle even for colocated individuals.

Our theoretical framework builds on a tension between lower collaboration costs when co-located, and the availability of a broader set of potential collaborators over distance. We start from this basic trade-off and then explore some of the key choices scientists face when deciding if they should collaborate locally versus over distance, how much effort to allocate to projects of different potential, and who they should pursue a more novel or interdisciplinary project with. We test the predictions from this framework by taking advantage of a source of plausibly exogenous variation in travel costs: the differential timing of entry by a low-cost carrier across multiple U.S. airports. Our difference-in-differences empirical strategy, combined with a series of robustness and falsification tests, supports the idea that the availability of lower fares had a causal effect on the probability and intensity of collaboration between scientists.<sup>51</sup> The effect is particularly pronounced for scientists that are less likely to

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<sup>50</sup>Policies targeted at reducing geographic frictions through lower travel costs may therefore need to account for the total cost of travel (including the opportunity cost of time) different types of individuals actually face.

<sup>51</sup>A back of the envelope calculation suggests that Southwest entry induced close to 400 copublications among chemistry faculty pairs. The sample mean of 0.1 copublications per year increases by 50% to 0.15 copublications per year after Southwest entry. We have 750 pairs and 10 post-entry years on average, leading to a back of envelope estimate of  $0.05 * 750 * 10 = 375$  copublications. While this number is sizeable, it is small relative to the total number of copublications among chemistry faculty members in this period. However, Southwest entry corresponds to a 20% price reduction affecting only a fraction of faculty pairs (a large fraction of pairs are served by Southwest or other low-cost carriers before our observation period). Over the last 30 years, the cost per mile for air travel across all routes within the U.S. dropped by 50% (Perry 2014). This suggests that reductions in air transportation overall could have had a substantial aggregate effect on collaboration, above and beyond the particular source of variation in air travel cost we use in this paper.

find co-authors of the same quality within their local environment, is present across multiple fields of science (chemistry, physics, biology, engineering, mathematics), and is robust to controlling for idiosyncratic scientist-pair characteristics, trends in collaboration over time, and department R&D budgets. Moreover, we do not observe a pre-trend in collaboration between scientist pairs that are going to experience lower air travel costs in the future.

Consistent with the theory, the reduction in geographic frictions also transforms the type of projects that emerge, influencing the direction of innovation:<sup>52</sup> our estimates suggest a sizable increase in higher quality papers, in projects that span different sub-disciplines, are more intensive in their use of specialized equipment, and are more novel. Comparisons between our findings in chemistry and mathematics suggest that complementarities in specialized equipment – while important for collaboration decisions between distant labs – are not the only driver behind the observed increase in joint projects over distance. Scientists also launch more experimental projects and projects that seem to take advantage of the complementary skills, ideas and knowledge that a distant lab may contribute to a collaboration.

Beyond the lower fares introduced by the low-cost airline we study in the paper, the cost per mile in the United States has dropped by over 50% in the last 30 years (Perry, 2014),<sup>53</sup> and convenience and routes have greatly improved. Our results should be therefore interpreted within this broader context of improvements in our ability to travel and work with distant collaborators. Whereas we cleanly estimate the impact of only part of these changes, improvements in air travel are likely to affect a much larger population of individuals. This includes inventors and researchers working within firms or public organizations that have multiple sites and face a similar trade-off between the ability to form ideal teams when not constraining their search for participants to one location, and the additional communication, coordination and travel costs geographically dispersed teams entail.<sup>54</sup> As advancements

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<sup>52</sup>For additional work focused on changes in the direction of research see Furman & Teodoridis, 2017; Catalini, 2017.

<sup>53</sup>We would expect similar effects in Europe, where low-cost airlines had even more of an effect on market structure and competition, as well as on uniting different economies.

<sup>54</sup>There is an interesting parallel here with the literature on communications costs and collaboration: while Agrawal & Goldfarb (2008) focused on academic collaborations, Forman & Zeebroeck (2012) subsequently found that the internet fostered R&D collaborations within firms. In principle, one could make progress on this related question using patenting and co-invention data together with our empirical strategy.

in communication technology make online interactions increasingly closer in latency and fidelity to offline ones, more research is needed to understand why face-to-face exchanges still appear to be a complement rather than a substitute to remote ones. In particular, while online exchanges seem to work well for executing on existing ideas, offline ones may still offer greater serendipity (Catalini, 2017). Moreover, trust between participants – often a prerequisite for collaboration when uncertainty makes it difficult to precisely evaluate individual contributions and effort – seems to still depend on individuals having spent enough unstructured time together in the same location.

Overall, relative to location decisions, which are extremely expensive for organizations to shape in the short run, the paper shows that investments targeted at facilitating travel and incentivizing face-to-face interactions may have higher returns than previously expected. This is because they not only affect the intensity of collaboration, but also the quality and impact of the resulting work. By facilitating better matches and more productive teams, they also support novel recombinations of ideas. While it may be tempting for firms and public funding agencies to assume that they can rely on technology to reduce costs and replace travel, our results support the view that – at least in the case of innovative outcomes – this is unlikely to be the case. The findings also show that this constitutes an opportunity for these organizations, as support for travel is a flexible policy lever that can be adjusted over time to shape R&D trajectories.

Whereas geographic distance acts as a sizable disincentive to collaboration and idea recombination, organizations can institutionalize and encourage travel to offset its effect. From a policy perspective, support for travel and better infrastructure can also be used to offer better opportunities to individuals and organizations that are located away from key economic and innovation hubs, influence the strategic location decisions of firms, and ultimately support economic growth. Firms with a geographically dispersed customer base rely on air travel (both for cargo and for employees) to remain competitive, and there have been multiple cases in recent years of headquarters being moved, among other reasons, for better access to flights: for example in 2017 Caterpillar Inc. decided to move from Peoria

(IL) to Chicago to better serve its customers, the vast majority of which are international.<sup>55</sup> Further exploring the trade-offs geographic frictions introduce for individuals and firms in these different contexts is a fruitful area for subsequent research.

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<sup>55</sup>See <https://www.vox.com/the-goods/2018/11/12/18080806/air-service-small-cities-crucial>

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# Tables

Table 1: Summary Statistics (Main Sample)

Variable	Mean	Std. Dev.
<b>Individual Scientist Level (n=845)</b>		
Age	49.6	11.0
Female	.10	.30
Average R&D budget in dept. (1000s USD)	279.88	226.75
Speciality:		
Physical chemistry	.32	.47
Biochemistry	.22	.41
Inorganic chemistry	.14	.34
Organic chemistry	.13	.34
Material science	.11	.31
Other	.08	.27
<b>Individual-Pair Level (n=758)</b>		
Year of Southwest entry	2001	4.5
Distance (in miles)	1232	808.6
Years in sample	17.3	4.6
Total copublications	1.9	3.4
Copub. both before and after	.09	.28
Copub. before Southwest entry	.49	.50
Copub. after Southwest entry	.60	.49
<b>Individual-Pair-Year Level (n=13,147)</b>		
Copublications	.11	.41
Local copubs	1.83	2.69
Local copubs with less productive colleagues	0.62	1.42
Different type of chemistry	0.46	0.50
One above average	0.67	0.47
Both above average	0.26	0.44
<b>Individual-pair-year level conditional on copublication (n=1,177)</b>		
Cites	44.95	71.83
Equipment Intensive 1	0.35	0.58
Equipment Intensive 2	0.48	0.87
Novel trends	0.12	0.36
Mean share novel	0.10	0.24
Max share novel	0.12	0.26
Total share novel	0.13	0.31
Novel in field	0.30	0.40

Table 2: **Comparing Pairs in the Analysis Sample to Pairs Not Experiencing Southwest**

	Distant Pairs Not Experiencing SW	Analysis Sample	P-Value For Equality of Means
Total copublications	1.78	1.90	0.19
Number of years observed	13.85	17.51	<0.01
Age (average in pair)	49.16	51.23	<0.01
Different type of chemistry	0.46	0.45	0.79
Average R&D budget in dept.	288.9	278.1	0.17
Observations	5,954	758	

Table 3: **Effects of Southwest Entry on Price, Passengers and Routes**

	(1) Passengers (log)	(2) Mean Price (log)	(3) Average Miles Flown (log)	(4) Direct Flight	(5) Number of Transfers
Southwest Entry	0.4437*** (0.0050)	-0.1910*** (0.0024)	0.0007 (0.0006)	0.0002 (0.0004)	-0.0174*** (0.0017)
Airport-Pair Fixed Effects	Yes	Yes	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes	Yes	Yes
Mean of Dep. Variable	4.238	5.454	7.066	0.007	1.239
Number of Pairs	55750	55750	55739	55750	55750
Number of Observations	956029	956029	955983	956029	956029

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ , Southwest entry is an indicator variable that takes value 1 if Southwest has started operating a flight between airports. All specifications include airport-pair fixed effects and year fixed effects. Estimation by ordinary least squares.

Table 4: **Effect of Southwest Entry on Copublications at the Individual-Pair Level**

DV=Copublications	(1) Baseline	(2) Controls	(3) Controls	(4) Timing	(5) Placebo 1	(6) Placebo 2
Southwest Entry	0.505*** (0.121)	0.526*** (0.121)	0.526*** (0.121)			-0.029 (0.216)
Mean Age		0.153*** (0.008)	0.268*** (0.015)			
Dept R&D Budget per Faculty (log)		0.364*** (0.127)	0.364*** (0.127)			
Years Since Both Have a PhD			-0.230*** (0.022)			
Southwest Entry (-1)				0.078 (0.152)		
Southwest Entry (0)				0.485*** (0.150)		
Southwest Entry (1)				0.518*** (0.166)		
Southwest Entry (2)				0.582*** (0.181)		
Fake Southwest Entry (Random Timing)					0.095 (0.121)	
Individual Pair FE	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	Yes	Yes	Yes	Yes	Yes	Yes
Number of Pairs	758	758	758	758	758	171
Number of Obs.	13,147	13,147	13,147	13,147	13,147	2,945

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . The dependent variable is the number of copublications between pairs of scientists. Southwest entry is an indicator variable that takes value 1 if Southwest has started operating a flight from airports close to the respective scientists. All specifications include individual-pair fixed effects and year fixed effects. Column 1 is our baseline specification. Column 2 adds controls for the age of the pair members and departmental R&D budget per faculty (both variables are means across the two pairs members). Column 3 additionally controls for the numbers of years that have passed since both pairs members obtained their PhD. Column 4 replaces Southwest entry with a set of indicator variables corresponding to different times from or since entry: Southwest entry (-1) is an indicator variable if the observation is in the year preceding Southwest entry; Southwest entry (0), Southwest entry (1), Southwest entry (2+) are defined analogously for the year of Southwest entry, the year after Southwest entry, and two years or more after Southwest entry. Column 5 is a placebo where we pretend Southwest entry has occurred in a random year for each pair. Column 6 is a placebo where we look at the set of pairs (not included in the baseline specification) who experienced Southwest entry followed by a Southwest exit event shortly thereafter. Estimation by Poisson Quasi-Maximum Likelihood.

Table 5: **Effect of Southwest Entry on Collaboration Among Mathematicians**

DV=Copublications	(1) Mathematics
Southwest Entry	0.247** (0.123)
Pair Fixed Effects	Yes
Year Fixed Effects	Yes
Number of Pairs	431
Number of Observations	5,514

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . These regressions are based on a dataset of U.S. mathematicians constructed using MathSciNet and the Mathematics Genealogy Project.

Table 6: **Southwest Entry and Collaborations Between U.S. Regions (CBSAs)**

DV=Copublications	(1) All	(2) Chemistry	(3) Biology	(4) Physics	(5) Engineering
Southwest Entry	0.503*** (0.020)	0.159*** (0.033)	0.494*** (0.032)	0.141*** (0.031)	0.238*** (0.055)
CBSA Pair Fixed Effects	Yes	Yes	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes	Yes	Yes
City Trends	Yes	Yes	Yes	Yes	Yes
Testing $H_0$			(3)=(2)	(4)=(2)	(5)=(2)
p-value			0.045	0.701	0.216
Number of Pairs	48,274	15,303	22,079	15,872	7,635
Number of Observations	965,480	306,060	441,580	317,440	152,700

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . These regressions are run at the CBSA-pair level. The dependent variable is the number of copublications between pairs of CBSAs. Southwest entry is an indicator variable that takes value 1 if Southwest has started operating a flight from airports close to the respective cities. Column 1 is based on copublications in all journals in our sample. Columns 2, 3, 4, 5, are based on chemistry, biology, physics and engineering journals respectively. All specifications include CBSA-pair fixed effects, year fixed effects, an origin-CBSA time trend and a destination-CBSA time trend. We also report p-values of statistical tests for the equality of Southwest Entry coefficients across samples. Estimation by Poisson Quasi-Maximum Likelihood.

Table 7: **Effect of Southwest Entry on Copublications: Which Pairs Are Most Affected?**

<b>Panel A: Chemistry</b>					
DV=Copublications	(1) Both Less Productive	(2) One More Productive	(3) Both More Productive	(4) Same type of Chemistry	(5) Different Type of Chemistry
Southwest Entry	0.228 (0.292)	0.566*** (0.153)	0.863*** (0.272)	0.340** (0.164)	0.668*** (0.165)
Pair Fixed Effects	Yes	Yes	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes	Yes	Yes
Testing $H_0$		(2)=(1)	(3)=(1)		(4)=(5)
p-value		0.015	0.003		0.002
Number of pairs	154	403	101	417	341
Number of observations	2498	6597	1630	7183	5964
<b>Panel B: Mathematics</b>					
DV=Copublications	(1) Both Less Productive	(2) One more Productive	(3) Both More Productive	(4) Same type Math	(5) Different Type of Math
Southwest Entry	0.196 (0.364)	0.002 (0.192)	0.374** (0.155)	0.260 (0.173)	0.204 (0.158)
Pair Fixed Effects	Yes	Yes	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes	Yes	Yes
Testing $H_0$		(2)=(1)	(3)=(1)		(4)=(5)
p-value		0.638	0.653		0.900
Number of pairs	61	155	263	180	299
Number of observations	859	2001	3318	2223	3955

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Panel A corresponds to our main chemistry sample. Panel B corresponds to the mathematics sample. The dependent variables in all specifications is the number of copublications. Different columns correspond to different subsamples in terms of productivity (columns 1-3) and whether both pair-members are specialized in the same subfield (column 4) or not. Productivity is measured as of the time of Southwest entry. All specifications are estimated by Poisson Quasi-Maximum Likelihood and include year fixed effects and pair fixed effects.



Table 8: **Southwest Entry and Local Copublications**

<b>Panel A: Chemistry</b>	(1)	(2)	(3)
	All Local Pairs	More Productive Pairs with Less Productive Local Colleagues	Less Productive Pairs With Less Productive Local Colleagues
Southwest Entry	0.092** (0.040)	-0.686*** (0.262)	0.169** (0.070)
Pair Fixed Effects	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes
Nr of pairs	741	126	547
Nr of obs.	12,939	2,305	9,533

<b>Panel B: Mathematics</b>	(1)	(2)	(3)
	All Local Pairs	More Productive Pairs with Less Productive Local Colleagues	Less Productive Pairs With Less Productive Local Colleagues
Southwest Entry	0.196*** (0.064)	-0.838* (0.457)	0.471** (0.219)
Pair Fixed Effects	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes
Nr of pairs	896	60	186
Nr of obs.	11,665	848	2,484

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Panel A corresponds to our main chemistry sample. Panel B corresponds to the mathematics sample. In both samples, we construct the set of local copublications of pairs affected by Southwest entry and use it as the dependent variable in column A. We also tag the set of local copublications with local colleagues whose productivity is below departmental average in the years preceding Southwest entry, and use it as the dependent variable in columns 2 and 3. The specification of column 2 is run on the sample of pairs where both members are above departmental average in productivity. The specification of column 3 is run on the sample of pairs where one or both members are below departmental average in productivity. Pairs that have all zero outcomes are dropped from the respective regressions, which results in the number of observations in columns 2 and 3 not summing up to the number of observations in column 1. All specifications are estimated by Poisson Quasi-Maximum Likelihood and include pair fixed effects and year fixed effects.

Table 9: Effect of Southwest Entry on the Type of Collaborations

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Cites	Novel Trends	Novel Trends	Mean Share Novel	Max Share Novel	Total Share Novel	Novel In field	Equipment Intensive 1	Equipment Intensive 2
Southwest Entry	0.420* (0.234)	1.175* (0.687)	0.057* (0.034)	0.069* (0.040)	0.093* (0.048)	-0.008 (0.085)	0.839*** (0.303)	0.647* (0.356)
Pair Fixed Effects	Yes	Yes		Yes	Yes	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Number of pairs	189	37	758	758	758	758	74	74
Number of obs.	606	137	1,177	1,177	1,177	1,177	261	261

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . These regressions are conditional on the pair having collaborated in the focal year. The dependent variables are the number of cites received (column 1), various measures of novelty based on keywords (column 2-6) and the number of equipment-intensive collaborations based on keywords (column 7 and 8). Pairs that never have a non-zero value of the dependent variable are dropped from the regressions. All specifications include individual-pair fixed effects and year fixed effects. Estimation by Poisson Quasi-Maximum Likelihood.

Table 10: Effect of Southwest Entry on the Type of Collaborations (CBSA-Level)

	Quality				Type			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Copubs		Cites-Weighted Copubs	Novel Trends	Novel Overall	Novel in Field	Across-Field Copubs	Equipment Intensive 1	Equipment Intensive 2
Southwest Entry	0.208*** (0.0647)	0.240** (0.0995)	0.235*** (0.0710)	0.365*** (0.112)	0.125* (0.0687)	0.266** (0.116)	0.264*** (0.0871)	0.236** (0.0939)
Year FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
CBSA-Pair FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Obs.	40,227	39,517	34,847	28,616	36,704	27,379	29,767	29,767

Notes: Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . These regressions are run at the CBSA-pair level and are based on a large sample of publications in chemistry and chemistry-related fields. The dependent variable in column 1 is the number of copublications between pairs of CBSAs. Column 2 uses citation-weighted copublications as dependent variable. Columns 3 to 5 use our different specifications of novelty. Columns 6 counts across-field collaborations, and Columns 7 and 8 equipment intensive ones, as inferred from the keywords associated with the papers.

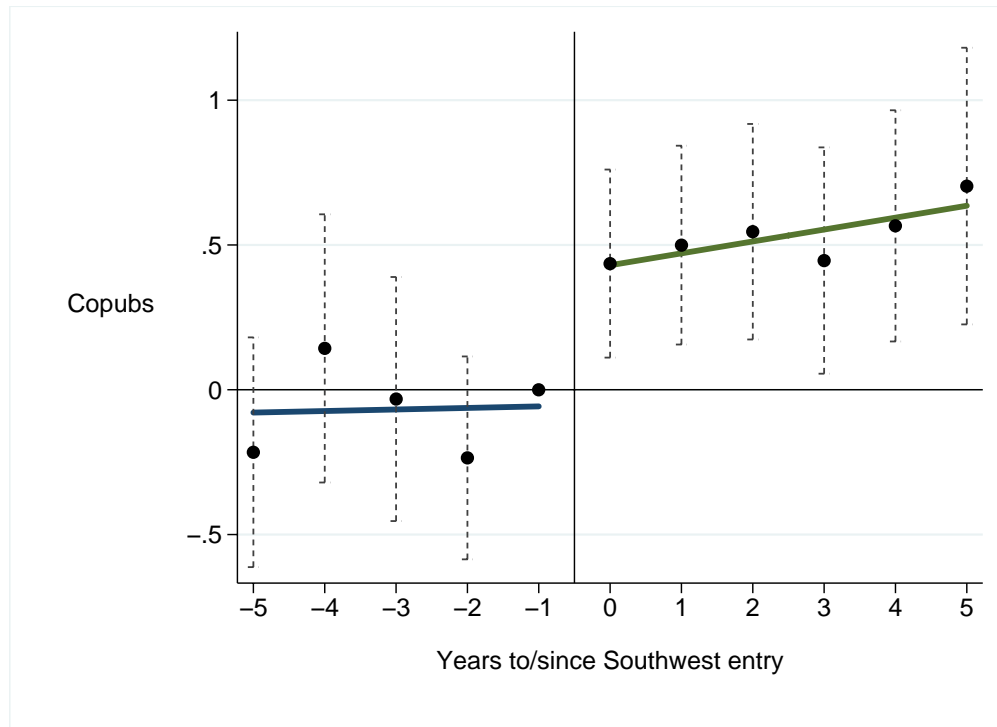
Table 11: **Quality of Distant and Local Collaborations at the Scientist Level**

	(1)	(2)
	Cites	Cites
	Chemistry	Math
Distant collaboration	0.0674***	0.2277***
	(0.0026)	(0.0069)
(Local collaboration omitted)		
Individual Fixed Effects	Yes	Yes
Year Fixed Effects	Yes	Yes
Number of pairs	4,737	2,104
Number of observations	46,060	12,187

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . The data underlying these regressions is the same as in the main analyses but is now structured at the individual scientist level. An observation is a paper which we tag as either a distant or a local collaboration; papers not involving a collaboration with another faculty member are excluded. The dependent variable is the number of cites received, and the variable of interest is whether the paper is a distant collaboration, with local collaborations as the omitted category. All specifications are estimated by Poisson Quasi-Maximum Likelihood and include individual scientist fixed effects and year fixed effects.

# Figures

Figure 1: Dynamics of the Effect of Southwest Entry: Individual-Pair Level



*Notes:* To generate this graph, we regress individual copublications on year fixed effects, pair effects and a set of indicator variables corresponding to 5 years before Southwest entry, 4 years before Southwest entry, ..., 4 years after Southwest entry, 5 years after Southwest entry (1 year before Southwest entry is omitted). We then plot the coefficients associated with these indicator variables against time to/from Southwest entry, superimposing a linear fit line before entry and after entry. The vertical bars represent 95% confidence intervals. The coefficient for the year immediately before entry is set to zero and displayed without a confidence interval since it our baseline year.

# Online Appendix

## Appendix A: Tables

Table A-1: **Changing the Definition of Proximate Airport**

	(1)	(2)	(3)	(4)
	10 miles	25 miles	50 miles	100 miles
Southwest Entry	0.896*** (0.206)	0.583*** (0.116)	0.505*** (0.095)	0.323*** (0.073)
Pair Fixed Effects	Yes	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes	Yes
Number of pairs	150	433	758	1,127
Number of obs.	2,600	7,275	13,147	19,292

*Notes:* Standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

Table A-2: **Inference with City-Pair Clustering**

DV=Copublications	(1)	(2)	(3)
Southwest Entry	0.505*** (0.121)	0.505*** (0.100)	0.492*** (0.134)
Pair Fixed Effects	Individual pair	Individual pair	City Pair
Year Fixed Effects	Yes	Yes	Yes
Clustering	Individual pair	City pair	City pair
Number of obs.	13,147	13,147	13,147

*Notes:* \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . The dependent variable is the number of copublications between pairs of scientists. Southwest entry is an indicator variable that takes value 1 if Southwest has started operating a flight from airports close to the respective scientists. Column 1 is our baseline regressions with individual pair fixed effects and individual pair clustering. In column 2, we keep individual pair fixed effects but cluster at the city pair level, using the POI2HDFE Stata command that implements the algorithm from Guimaraes & Portugal (2010). In column 3, we replace individual pair fixed effects with city pair fixed effects and cluster by city pair fixed effects.

Table A-3: **Robustness to Including Non-Collaborating Pairs**

DV=copublications	(1)	(2)	(3)
Southwest Entry	0.505*** (0.121)	0.500*** (0.129)	0.337*** (0.124)
Pair Fixed Effects	Individual pair	University Pair	University Pair
Year Fixed Effects	Yes	Yes	Yes
Sample includes non-collaborating pairs	No	No	Yes
Number of obs.	13,147	13,147	1,425,523

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Southwest entry is an indicator variable that takes value 1 if Southwest has started operating a flight from airports close to the respective scientists. Estimation by Poisson Quasi-Maximum Likelihood. Column 1 is our baseline regression, which includes only pairs of scientists who collaborate at some point. Column 2 keeps the same sample but replaces individual pair fixed effect by university pair fixed effects. Column 3 adds a 10% random sample of non-collaboration pairs to the sample of collaborating pairs and is run with university pair fixed effects.

Table A-4: **Using Journals to Define Fields of Specialization**

Field	Journal (examples)
Biochemistry	<i>Journal of Biological Chemistry, Biochemistry</i>
Inorganic Chemistry	<i>Inorganic Chemistry</i>
Material Science	<i>Macromolecules, Advanced Materials</i>
Physical Chemistry	<i>Journal of Physical Chemistry</i>
Organic Chemistry	<i>Journal of Organic Chemistry, Organic Letters</i>

The area of specialization for a given faculty member is inferred from the journals s/he publishes in. For instance, a faculty member who publishes often in the *Journal of Biological Chemistry* is assumed to be specialized in biochemistry.

Table A-5: **Summary Statistics (CBSA Sample Across Fields)**

	Mean	Standard deviation
All copubs	1.42	23.51
Chemistry copubs	0.22	3.89
Biology copubs	0.35	6.83
Physics copubs	0.21	3.10
Engineering copubs	0.07	1.33
Southwest entry	0.16	0.36
Year	1999.50	5.77

*Notes:* The unit of observation is a CBSA-pair-year. N=965,480. This data is used in Table 6.

Table A-6: **Summary Statistics (Mathematics Sample)**

	Mean	Standard deviation
Copublications	0.05	0.27
Southwest entry	0.60	0.49
Year	2001.55	5.20
One more productive	0.41	0.49
Both more productive	0.44	0.50
Different speciality	0.68	0.47
Local copublications	0.26	0.67
Local copublications with less productive colleagues	0.03	0.18

*Notes:* The unit of observation is a scientist-pair year. This data is used in Tables 5, 7 (panel B), and 8 (panel C).



Table A-7: **Summary Statistics (Chemistry Collaborations Across CBSAs)**

Copublications	2.11	5.45
Citation-Weighted Copubs	81.64	310.41
Novel Copubs 1	0.59	1.66
Novel Copubs 2	0.36	1.17
Novel in Field 1	0.49	0.91
Across-Field Copubs	0.31	1.05
Equipment Intensive Copubs 1	0.35	1.09
Equipment Intensive Copubs 2	0.51	1.66
Southwest entry	0.36	0.48
Year	2007.22	4.72

*Notes:* The unit of observation is a CBSA-pair-year (N=40,227). This data is used in Table 10.

Table A-8: Cross-Correlation Table (Main sample)

Variables	Copubs	Southwest Entry	Year	Cites	Equip. Intensive 1	Equip. Intensive 2	Novel trends	Mean share novel	Max share novel	Total share novel	Novel in field	Local copubs	Local with less productive	Different type of chemistry	One above average	Both above average
Copublications	1.000															
Southwest entry	0.015	1.000														
Year	-0.006	0.644	1.000													
Cites	0.320	-0.109	-0.220	1.000												
Equipment Intensive 1	0.403	0.073	-0.016	0.103	1.000											
Equipment Intensive 2	0.345	0.086	0.005	0.053	0.884	1.000										
Novel trends	0.055	0.030	0.102	-0.055	0.015	0.006	1.000									
Mean share novel	0.002	-0.147	-0.218	0.030	0.026	0.034	0.034	1.000								
Max share novel	0.178	-0.146	-0.220	0.093	0.090	0.106	0.035	0.953	1.000							
Total share novel	0.275	-0.151	-0.228	0.140	0.126	0.120	0.020	0.870	0.942	1.000						
Novel in field	0.196	-0.045	-0.082	0.113	0.113	0.101	0.131	0.528	0.549	0.518	1.000					
Local copubs	0.018	0.106	0.122	-0.030	-0.009	0.024	0.059	-0.029	-0.044	-0.037	-0.045	1.000				
Local copubs with less productive colleagues	-0.000	0.018	0.029	-0.051	0.003	0.021	0.116	-0.004	-0.013	0.014	-0.012	0.437	1.000			
Different type of chemistry	-0.005	0.010	0.017	0.041	0.133	0.144	0.015	0.057	0.059	0.051	0.119	0.077	0.024	1.000		
One above average	-0.017	0.040	0.009	-0.085	-0.002	0.008	-0.024	-0.017	-0.038	-0.048	-0.031	0.055	0.035	-0.048	1.000	
Both above average	0.035	-0.053	-0.030	0.109	0.017	0.011	-0.072	-0.082	-0.079	-0.078	-0.011	0.103	-0.069	0.032	-0.198	1.000

Table A-9: Cross Correlation Table (CBSA Sample Across Fields)

Variables	All Copubs	Chemistry Copubs	Biology Copubs	Physics Copubs	Engineering Copubs	Southwest	Year
All copubs	1.000						
Chemistry copubs	0.879	1.000					
Biology copubs	0.976	0.807	1.000				
Physics copubs	0.862	0.871	0.775	1.000			
Engineering copubs	0.712	0.829	0.604	0.770	1.000		
Southwest entry	-0.006	-0.018	-0.007	-0.010	-0.016	1.000	
Year	0.019	0.014	0.014	0.011	0.016	0.245	1.000

Notes: The unit of observation in this dataset is a CBSA-pair year. N=965,480. This data underpins Table 6.

Table A-10: Cross-Correlation Table (Mathematics Sample)

Variables	Copublications	Southwest entry	Year	One more productive	Both more productive	Different speciality	Local copubs	Local copubs with less productive colleagues
Copublications	1.000							
Southwest entry	-0.023	1.000						
Year	-0.018	0.545	1.000					
One more productive	-0.049	-0.015	-0.021	1.000				
Both more productive	0.073	0.041	0.009	-0.739	1.000			
Different speciality	-0.063	0.001	0.014	0.035	-0.063	1.000		
Local copubs	0.034	-0.007	0.030	-0.061	0.112	0.022	1.000	
Local copubs with less productive colleagues	-0.013	0.009	0.023	0.026	-0.062	0.008	0.236	1.000

*Notes:* The unit of observation in this dataset is a scientist-pair year. N=965,480. This data underpins table 5, 7 (panel B), and 8 (panel C)

Table A-11: Cross-Correlation Table (Chemistry Collaborations Across CBSAs)

Variables	Copubs	Citation-Weighted Copubs	Novel Copubs 1	Novel Copubs 2	Novel in Field 1	Across-Field Copubs	Equipment Intensive Copubs 1	Equipment Intensive Copubs 2	Southwest entry	Year
Copublications	1.000									
Citation-Weighted Copubs	0.753	1.000								
Novel Copubs 1	0.868	0.638	1.000							
Novel Copubs 2	0.709	0.544	0.718	1.000						
Novel in Field 1	0.084	0.065	0.195	0.311	1.000					
Across-Field Copubs	0.752	0.578	0.638	0.568	0.076	1.000				
Equipment Intensive Copubs 1	0.811	0.621	0.739	0.616	0.097	0.701	1.000			
Equipment Intensive Copubs 2	0.764	0.579	0.700	0.586	0.098	0.665	0.956	1.000		
Southwest entry	0.141	0.126	0.129	0.091	-0.001	0.063	0.098	0.088	1.000	
Year	0.162	0.076	0.183	0.068	0.166	0.039	0.122	0.119	0.066	1.000

Notes: The unit of observation in this dataset is a CBSA-pair year. N=40,227. This data underpins table 10.

Table A-12: **Presence and Concurrent Entry of Other Airlines**

	(1)	(2)	(3)	(4)	(5)	(6)
DV=Copubs	Baseline	No LC before SW	LC before SW	excl. entry by other LC	excl. entry by MC	excl. entry by LC and MC
Southwest Entry	0.505*** (0.121)	0.625*** (0.157)	0.291 (0.199)	0.498*** (0.131)	0.477*** (0.145)	0.441*** (0.157)
Pair Fixed Effects	Yes	Yes	Yes	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes	Yes	Yes	Yes
Number of pairs	758	479	279	649	556	480
Number of obs.	13,147	8,349	4,798	11,261	9,711	8,367

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Column 1 is the baseline regression. Column 2 limits the sample to pairs in locations where no other low-cost airline was operating in the year before Southwest entry. Conversely, column 3 limits the sample to pairs in locations where another low-cost airline was operating in the year before Southwest entry. Column 4 excludes cases where another low cost company entered in the same year as Southwest, column 5 excludes cases where another major company entered in the same year as Southwest and column 6 excludes both of these groups.

Table A-13: **Northeast Corridor Falsification Test**

	(1)	(2)	(3)
DV= Copublications	All	NE corridor only	excluding NE corridor
Southwest Entry	0.505*** (0.121)	-0.745 (0.597)	0.559*** (0.123)
Pair Fixed Effects	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes
Number of pairs	758	31	727
Number of observations	13,147	564	12,583

*Notes:* Standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

Table A-14: **Intensive and Extensive Margin of Southwest Entry**

	(1)	(2)	(3)
DV= Copublications	Baseline	Extensive Margin	Intensive Margin
Southwest Entry	0.505*** (0.121)	0.410*** (0.135)	0.806*** (0.234)
Pair Fixed Effects	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes
Number of pairs	758	692	66
Number of observations	13,147	11,969	1,178

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . The dependent variable is the number of copublications between pairs of scientists. Southwest entry is an indicator variable that takes value 1 if Southwest has started operating a flight from airports close to the respective scientists. Column 1 is the baseline specification. Column 2 restricts the sample to pairs of scientists who collaborate either before or after entry, but not both; Column 3 restricts the sample to pairs of scientists who collaborate both before and after entry. All specifications include individual-pair fixed effects and year fixed effects. Estimation by Poisson Quasi-Maximum Likelihood.

Table A-15: **Intensive Margin Pairs: Advisor-Advisee Pairs versus Others**

	(1)	(2)
DV=Copublications	Intensive Margin Advisor-Advisee Pairs	Intensive Margin Other Pairs
Southwest Entry	1.020* (0.566)	0.844*** (0.241)
Pair Fixed Effects	Yes	Yes
Year Fixed Effects	Yes	Yes
Number of observations	126	1,052

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Advisor-advisee pair takes value one for pairs with at least one copublication before Southwest entry where one of is the first author and the other is the last.

Table A-16: **Alternative Functional Forms**

	(1)	(2)	(3)
	Poisson	OLS	OLS
	Copublications	Copublications	Any copublication
Southwest Entry	0.505*** (0.121)	0.052*** (0.015)	0.030*** (0.010)
Pair Fixed Effects	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes
Number of pairs	758	758	758
Number of obs.	13,147	13,147	13,147

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . Southwest entry is an indicator variable that takes value 1 if Southwest has started operating a flight from airports close to the respective scientists. Column 1 is the baseline regression at the individual pair level (estimated by Poisson Quasi-Maximum Likelihood.) Column 2 estimates the same specification with ordinary least squares. Column 3 is a linear probability model with an indicator variable for any copublication.



Table A-17: **Sensitivity to Outliers**

DV=Copublications	(1)	(2)	(3)
Southwest Entry	0.505*** (0.121)	0.308*** (0.118)	0.424*** (0.112)
Pair Fixed Effects	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes
Comment	Baseline	Excluding outlier pairs	Winsorizing outliers
Number of pairs	758	732	758
Number of obs.	13,147	12,701	13,147

*Notes:* \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ , The dependent variable is the number of copublications between pairs of scientists. Southwest entry is an indicator variable that takes value 1 if Southwest has started operating a flight from airports close to the respective scientists. Column 1 is our baseline regression with individual pair fixed effects and individual pair clustering. In column 2, we exclude pairs that have more than two copublications in any given year. In column 3, we winsorize observations that have more than two copublications to two.

Table A-18: **Quality and Novelty of Copublications of Productive Scientists with Less Productive Local Colleagues**

	(1) Cites-weighted	(2) Novelty-weighted
Southwest entry	-1.054*** (0.322)	-0.008 (0.046)
Pair Fixed Effects	Yes	Yes
Year Fixed Effects	Yes	Yes
Number of pairs	125	124
Number of observations	1,547	419

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table A-19: **Heterogeneous Effects of Southwest Entry**

Panel A: Age at Entry			
	(1)	(2)	
	One or Both Above 50	Both Below 50	
Southwest Entry	0.354*** (0.129)	0.796*** (0.249)	
Testing $H_0$		(1)=(2)	
p-value		0.0471	
Number of pairs	552	206	
Number of observations	9,803	3,344	
Panel B: R&D Budget at Entry			
	(1)	(2)	(3)
	Both Below Average	One Below Average	Both Above Average
Southwest Entry	0.719*** (0.228)	0.426* (0.249)	0.431** (0.200)
Testing $H_0$		(1)=(2)	(1)=(3)
p-value		0.253	0.262
Number of pairs	244	220	294
Number of observations	4,245	3,833	5,069
Panel C: Distance at Entry			
	(1)	(2)	(3)
	Less than 1000	Between 1000 and 2000	Above 2000
Southwest Entry	0.135 (0.219)	0.506*** (0.174)	0.907*** (0.252)
Testing $H_0$		(1)=(2)	(1)=(3)
p-value		0.126	0.007
Number of pairs	188	392	178
Number of observations	3,097	7,001	3,049
Panel D: Gender			
	(1)	(2)	
	One or Both Female	Both Male	
Southwest Entry	0.059 (0.322)	0.585*** (0.130)	
Testing $H_0$		(1)=(2)	
p-value		0.0665	
Number of pairs	118	640	
Number of observations	1,906	11,241	

*Notes:* The dependent variable in all specifications is the number of copublications. All specifications include pair fixed effects and year fixed effects. Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

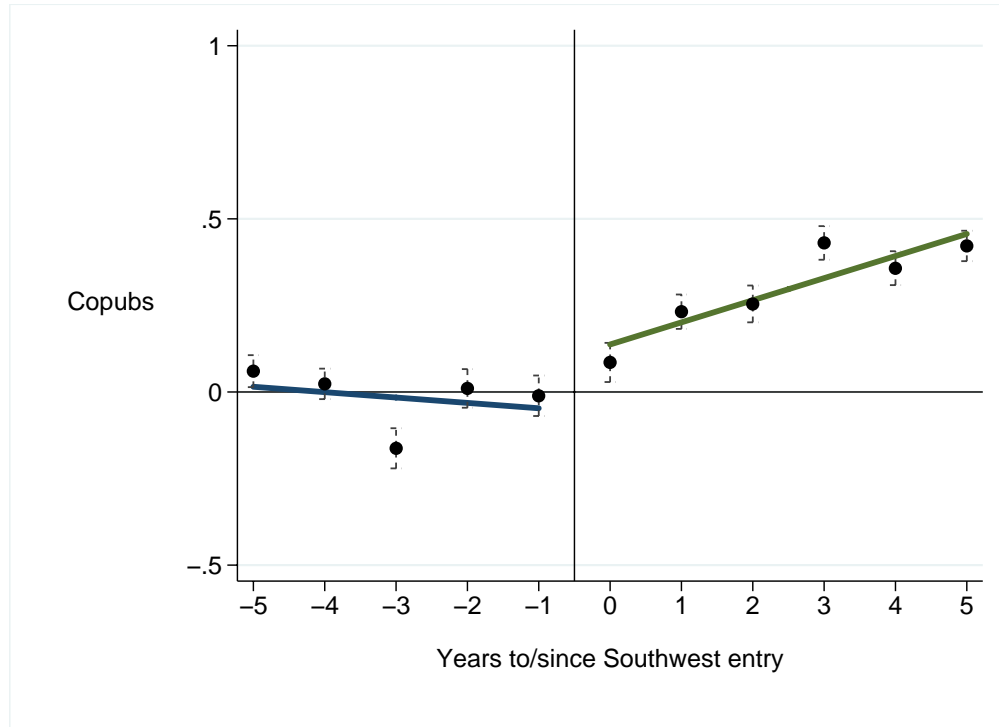
Table A-20: **Southwest Entry in Capital-Intensive Subfields of Chemistry**

DV=Copublications	(1) Neither in K-intensive	(2) One in K-intensive	(3) Both in K-intensive
Southwest entry	0.364 (0.226)	0.679*** (0.208)	0.349** (0.175)
Pair Fixed Effects	Yes	Yes	Yes
Year Fixed Effects	Yes	Yes	Yes
Number of pairs	206	229	323
Number of observations	3,502	3,983	5,662

*Notes:* Robust standard errors in parentheses. \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

# Appendix A: Figures

Figure A-1: Dynamics of The Effect of Southwest Entry: CBSA-Pair Level



*Notes:* To generate this graph, we regress CBSA copublications on year fixed effects, pair effects, origin-CBSA and destination-CBSA time trends, and a set of indicator variables corresponding to 5 years before Southwest entry, 4 years before Southwest entry, ..., 4 years after Southwest entry, 5 years after Southwest entry (5 years before Southwest entry omitted). We then plot the coefficients associated with these indicator variables against time to/from Southwest entry, superimposing a linear fit line before entry and after entry.

## Appendix B: Theoretical Model Under $w=0$

**B1:**  $\pi_D^* > \pi_{G,L}^*$

Intuitively, if scientists go on the global market for co-authors and end up with someone of the same quality as a local co-author (the outside option), then returns will be lower because they have to incur travel costs:

$$\begin{aligned} \frac{(p_L q)^2}{2\alpha} &> \frac{(p_L q)^2}{4\alpha} \left(1 + \frac{1}{2} t_L^*\right) > 0 \\ \iff \frac{1}{2} &> \frac{1}{4} + \frac{1}{8} t_L^* > 0, \end{aligned} \tag{15}$$

since the maximum value of  $t_L^*$  is 1. Q.E.D.

**B2:**  $\pi_{G,H}^* > \pi_D^*$  requires an assumption on parameters

$$\pi_{G,H}^* - \pi_D^* > 0 \tag{16}$$

$$\begin{aligned} \iff \frac{(p_H q)^2}{4\alpha} \left(1 + \frac{1}{2} t_H^*\right) - \frac{(p_L q)^2}{2\alpha} &> 0 \\ \iff (p_H)^2 \left(1 + \frac{1}{2} t_H^*\right) &> 2(p_L)^2 \end{aligned} \tag{17}$$

Intuitively, the benefit of a better co-author has to compensate for higher travel costs.

**B3:**  $\pi_G^* > \pi_D^*$  requires an assumption on parameters

$$(1 - z)\pi_{G,L}^* + z\pi_{G,H}^* - \pi_D^* > 0 \quad (18)$$

$$\begin{aligned} \iff (1 - z)\frac{(p_L q)^2}{4\alpha}\left(1 + \frac{1}{2}t_L^*\right) + z\frac{(p_H q)^2}{4\alpha}\left(1 + \frac{1}{2}t_H^*\right) - \frac{(p_L q)^2}{2\alpha} &> 0 \\ \iff (1 - z)(p_L)^2\left(1 + \frac{1}{2}t_L^*\right) + z(p_H)^2\left(1 + \frac{1}{2}t_H^*\right) - 2(p_L)^2 &> 0 \\ \iff z\left[(p_H)^2\left(1 + \frac{1}{2}t_H^*\right) - (p_L)^2\left(1 + \frac{1}{2}t_L^*\right)\right] - (p_L)^2\left(1 - \frac{1}{2}t_L^*\right) &> 0 \end{aligned} \quad (19)$$

Intuitively, the benefit of a better co-author has to compensate for higher travel costs. Conditional on this being the case, the greater the likelihood that this will happen (i.e.  $z$  high), the more attractive the choice of a distant co-author becomes.

**B4:**  $\frac{\partial[\pi_G^* - \pi_D^*]}{\partial q} > 0$  requires an assumption on parameters

First,  $\frac{\partial[\pi_{G,H}^*]}{\partial q}$ :

$$\begin{aligned} \frac{\partial\pi_{G,H}^*}{\partial q} &= \frac{(p_H)^2 q}{2\alpha}\left(1 + \frac{1}{2}t_H^*\right) + \frac{(p_H q)^2}{4\alpha} \frac{1}{2} \frac{(p_H)^2 q}{4\alpha\beta} \\ &= \frac{(p_H)^2 q}{2\alpha}(1 + t_H^*) \end{aligned} \quad (20)$$

Second,  $\frac{\partial[\pi_{G,L}^*]}{\partial q}$ :

$$\frac{\partial\pi_{G,L}^*}{\partial q} = \frac{(p_L)^2 q}{2\alpha}(1 + t_L^*) \quad (21)$$

Third,  $\frac{\partial[\pi_D^*]}{\partial q}$ :

$$\frac{\partial\pi_D^*}{\partial q} = \frac{(p_L)^2 q}{\alpha} \quad (22)$$

Finally, use the three previous results to show:  $\frac{\partial[\pi_G^* - \pi_D^*]}{\partial q}$

$$(1 - z) \frac{\partial \pi_{G,L}^*}{\partial q} + z \frac{\partial \pi_{G,H}^*}{\partial q} > \frac{\partial \pi_D^*}{\partial q} \quad (23)$$

$$\iff (1 - z) \frac{(p_L)^2 q}{2\alpha} (1 + t_L^*) + z \frac{(p_H)^2 q}{2\alpha} (1 + t_H^*) > \frac{(p_L)^2 q}{\alpha}$$

$$\iff \frac{q}{2\alpha} [(1 - z)(p_L)^2 (t_L^* - 1) + z[(p_H)^2 (1 + t_H^*) - 2(p_L)^2]] > 0 \quad (24)$$

The **first term** is negative since  $t_L^*$  is between 0 and 1. This reflects the idea that if a scientist uses the global market, s/he can end up with a co-author of the same quality as a local co-author, but with the addition of travel costs. The **second term** is positive and reflects the idea that if a scientist uses the global market s/he can end up with a co-author of higher quality with probability  $z$ . Ignoring the first term  $\frac{q}{2\alpha}$ , we simplify and re-write:

$$\iff (1 - z)(p_L)^2 (t_L^* - 1) + z[(p_H)^2 (1 + t_H^*) - 2(p_L)^2] > 0$$

$$\iff z[(p_H)^2 (1 + t_H^*) - (p_L)^2 (1 + t_L^*)] - (p_L)^2 (1 - t_L^*) > 0 \quad (25)$$

Note that the condition to satisfy  $\frac{\partial[\pi_G^* - \pi_D^*]}{\partial q}$  is very similar to the condition to satisfy  $\pi_G^* - \pi_D^* > 0$ . The subsection below will show that if the level condition ( $\pi_G^* - \pi_D^* > 0$ ) is satisfied, then the difference condition ( $\frac{\partial[\pi_G^* - \pi_D^*]}{\partial q}$ ) is automatically satisfied. Therefore, no further assumptions are needed in addition to the level condition.

## B5: Comparing the level condition to the difference condition

Level condition:  $\pi_G^* - \pi_D^* > 0$  means:

$$z[(p_H)^2 (1 + \frac{1}{2}t_H^*) - (p_L)^2 (1 + \frac{1}{2}t_L^*)] - (p_L)^2 (1 - \frac{1}{2}t_L^*) > 0 \quad (26)$$

Difference condition:  $\frac{\partial[\pi_G^* - \pi_D^*]}{\partial q} > 0$  means

$$z[(p_H)^2 (1 + t_H^*) - (p_L)^2 (1 + t_L^*)] - (p_L)^2 (1 - t_L^*) > 0 \quad (27)$$

Therefore, by combining both conditions one can show that:

$$\begin{aligned}
& \frac{\partial[\pi_G^* - \pi_D^*]}{\partial q} > \pi_G^* - \pi_D^* \\
\iff & z[(p_H)^2(1 + t_H^*) - (p_L)^2(1 + t_L^*)] - (p_L)^2(1 - t_L^*) > \\
& z[(p_H)^2(1 + \frac{1}{2}t_H^*) - (p_L)^2(1 + \frac{1}{2}t_L^*)] - (p_L)^2(1 - \frac{1}{2}t_L^*) \\
& \iff z[(p_H)^2(t_H^*) - (p_L)^2(t_L^*)] + (p_L)^2(t_L^*) > \\
& z[(p_H)^2(\frac{1}{2}t_H^*) - (p_L)^2(\frac{1}{2}t_L^*)] + (p_L)^2(\frac{1}{2}t_L^*) \\
& \iff 2z[(p_H)^2(t_H^*) - (p_L)^2(t_L^*)] + 2(p_L)^2(t_L^*) > \\
& z[(p_H)^2(t_H^*) - (p_L)^2(t_L^*)] + (p_L)^2(t_L^*) \\
& \iff z[(p_H)^2(t_H^*) - (p_L)^2(t_L^*)] + (p_L)^2(t_L^*) > 0
\end{aligned}$$

Therefore, if a given project is preferred under a distant co-author, this effect is magnified when project quality rises.



## Appendix C: Theoretical Model Under $z > w > 0$

Here we address the general case where  $z > w > 0$  (in the main text we solved for the specific case where  $w = 0$ ).

Returns under a *local co-author* are:

$$\pi_L^* = (1 - w) \frac{(p_B q)^2}{2\alpha} + w \frac{(p_G q)^2}{2\alpha} = \frac{q^2}{2\alpha} (w p_G^2 + (1 - w) p_B^2). \quad (28)$$

Returns under a *distant co-author* are:

$$\pi_D^* = \frac{q^2}{4\alpha} \left( z p_G^2 \left( 1 + \frac{t_G^*}{2} \right) + (1 - z) p_B^2 \left( 1 + \frac{t_B^*}{2} \right) \right), \quad (29)$$

where  $t_i^* = \frac{(p_i q)^2}{8\alpha\beta}$ .

Therefore, the difference between both profit levels can be shown to be:

$$\pi_D^* - \pi_L^* = \frac{q^2}{4\alpha} \left[ p_G^2 \left( z \left( 1 + \frac{t_G^*}{2} \right) - 2w \right) + p_B^2 \left( (1 - z) \left( 1 + \frac{t_B^*}{2} \right) - 2(1 - w) \right) \right]. \quad (30)$$

By again using the definition of  $\beta = \frac{1}{\theta}$ , one can show that:

$$\frac{\partial[\pi_D^* - \pi_L^*]}{\partial\theta} = \left( \frac{q^2}{8\alpha} \right)^2 (z p_G^4 + (1 - z) p_B^4) > 0, \quad (31)$$

and from here one can see that the three predictions of the main text are satisfied.

## Appendix D: Extension of the Theoretical Model Based on Project Variance

In the baseline model, the parameter  $q$  representing project quality is taken as given, and the probability of success of a project depends on the quality of a match with a given co-author. While the global pool of co-authors offers potentially better matches and therefore a higher probability of success, working with distant co-authors incurs travel costs. In an effort to streamline the exposition, the baseline model also does not incorporate any form of project heterogeneity beyond  $q$ .

In this extension, we introduce two elements: first, there are two states of the world (good and bad) each occurring with 50% probability. Second, there are two types of projects (safe and risky) with the same expected value  $q$ , but different variance. The greater variance is modelled by a bigger gap in the realization of  $q$  between the two states of the world.

The underlying logic of the extension is to represent a world in which novel projects are high variance (i.e. they are more risky), while incremental ones are relatively safe. The goal of the extension is to explore how reductions in travel costs affect high versus low variance projects.

The two types of projects have the following properties:

- If the *safe project* is chosen, there are two possible states of the world, each happening with 50% probability:  $q_2$  in the bad state of the world and  $q_3$  in the good state of the world
- If the *risky project* is chosen, there are two possible states of the world, each happening with 50% probability:  $q_1$  in the bad state of the world and  $q_4$  in the good state of the world.

Assume that the expected value of  $q$  is the same for both safe and risky projects, such that  $q_2 + q_3 = q_1 + q_4$ . We impose this so that the results of our comparative statics are purely driven by the differences in variance, not the mean. As a result, the risky project has more

extreme outcomes both in the good and in the bad states of the world:

$$q_1 < q_2 \leq q_3 < q_4$$

One can think of novel projects as being extremely successful in the good state of the world and, at the same time, complete failures in the bad state of the world. For a scientist, the alternative is to develop a safer, incremental project with profitability levels that depend less on the state of the world. The next two subsections compare how outcomes for a given project type depend on the co-author being local versus distant.

## D1 Local Co-Author

Once the state of the world is realized, scientists choose effort according to the equations of the baseline model, such that:

$$e_i^* = \frac{p_i q_i}{\alpha}$$

$$\pi_{L,i}^* = \frac{(p_i q_i)^2}{2\alpha}$$

One can show that expected returns under the *safe project* are:

$$\pi_{L,S}^* = \frac{p_L^2}{4\alpha} (q_2^2 + q_3^2)$$

Similarly, expected returns under the *risky project* are:

$$\pi_{L,R}^* = \frac{p_L^2}{4\alpha} (q_1^2 + q_4^2)$$

## D2 Distant Co-author

Expected returns under the *safe project*, after optimally choosing effort are:

$$\pi_{D,S}^* = \frac{p_H^2}{8\alpha} \left[ q_2^2 \left( 1 + \frac{(p_H q_2)^2}{16\alpha\beta} \right) + q_3^2 \left( 1 + \frac{(p_H q_3)^2}{16\alpha\beta} \right) \right]$$

Similarly, expected returns under the *risky project* are:

$$\pi_{D,R}^* = \frac{p_H^2}{8\alpha} \left[ q_1^2 \left( 1 + \frac{(p_H q_1)^2}{16\alpha\beta} \right) + q_4^2 \left( 1 + \frac{(p_H q_4)^2}{16\alpha\beta} \right) \right]$$

## D3 Choice of Co-Author After a Reduction in Travel Costs

Recall that  $\beta = \frac{1}{\theta}$ , such that an increase in  $\theta$  is an increase in travel efficiency (or equivalently a reduction in travel costs). In the comparative statics presented here, we want to evaluate how reductions in travel costs affect the optimal choice of co-author. Conditional on choosing the *safe project*, the reduction in travel costs makes the choice of a distant co-author more likely because of the following increase in expected returns:

$$\frac{\partial[\pi_{D,S}^* - \pi_{L,S}^*]}{\partial\theta} = \frac{p_H^4}{2(8\alpha)^2}(q_2^4 + q_3^4) > 0$$

Similarly, conditional on choosing the *risky project*, the reduction in travel costs makes the choice of a distant co-author more likely by the following increase in expected returns:

$$\frac{\partial[\pi_{D,R}^* - \pi_{L,R}^*]}{\partial\theta} = \frac{p_H^4}{2(8\alpha)^2}(q_1^4 + q_4^4) > 0$$

The key question we need to address is which type of project is affected the most. By combining the two previous equations and comparing them in relative terms, we see that the reduction in travel costs will make distant co-authors more appealing for risky projects if  $\frac{\partial[\pi_{D,R}^* - \pi_{L,R}^*]}{\partial\theta} > \frac{\partial[\pi_{D,S}^* - \pi_{L,S}^*]}{\partial\theta}$ . One can easily show that this equation can be simplified to:

$$q_1^4 + q_4^4 > q_2^4 + q_3^4$$

For simplicity, let us assume a scenario in which  $q_2 = q_3 = q$  such that the variance under the *safe project* is zero, and where we still assume that the expected value is the same for both the safe and the risky project:

$$2q = q_1 + q_4$$

One can generalize the values of  $q_1$  and  $q_4$  in the following way:  $q_1 = \alpha q$  and  $q_4 = (2 - \alpha)q$ , with  $0 < \alpha < 1$ . It is possible to show that for any value of  $\alpha$ , the condition  $\frac{\partial[\pi_{D,R}^* - \pi_{L,R}^*]}{\partial\theta} > \frac{\partial[\pi_{D,S}^* - \pi_{L,S}^*]}{\partial\theta}$  is satisfied. This implies that reductions in travel costs make distant co-authors more appealing especially for projects with higher variance in outcomes (e.g. more novel, interdisciplinary).

# Appendix E: Extension of the Theoretical Model Based on Project Specialization

In the baseline version of the theoretical framework, projects are homogeneous and co-authors only differ in how they influence the probability of success. In this extension, instead, we allow for a continuum of projects that have different needs for specialized (versus not) co-authors.

More specialized projects can have higher returns, but only if the scientist manages to attract a suitable co-author with specialized knowledge. If the co-author match is bad, then these specialized projects underperform. Conditional on a project's need for specialization, scientists endogenously choose the optimal type of co-author both in terms of distance as well as specialization. Intuitively, the trade-off is the following: in the global pool of co-authors, one can find all possible types of specialists, but at the cost of incurring both travel as well as search costs that increase with the co-author's specialization. In the local pool, one is more limited in the scope of co-author specialization available, but there are no travel costs.

The goal of this exercise is to explore how a reduction in travel costs affects the optimal type of co-author chosen depending on the degree of specialization of the project. In particular, we want to understand which type of project benefits more.

## E1 Distant Co-Author

The profit function under a distant co-author is:

$$\pi_D(e, t) = \gamma \frac{\theta - B}{\theta - s} e - \alpha \frac{e^2}{t} - \beta t^2 - s,$$

where  $B > 1$ .<sup>56</sup> There is a continuum of projects  $\theta \in (B, \bar{\theta})$  that differ in their degree of specialization. The fact that  $B > 1$  ensures that the higher the value of  $\theta$ , the more beneficial it is to have a more specialized co-author ( $s$ ).

The co-author's degree of specialization  $s \in [0, 1]$  is endogenously chosen in response to the project's need for specialization. Intuitively, the more specialized the project is (higher

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<sup>56</sup>The goal of the constant  $B$  is to ensure that this ratio increases with the project's degree of specialization  $\theta$ .

$\theta$ ), the more important it is to find a specialized co-author (higher  $s$ ), so that  $s$  and  $\theta$  are complements. The search cost is a linear function of the degree of specialization reflecting the fact that the more specialized the co-author one desires, the more one has to incur search costs.

Last, the travel cost is  $t \in [0, 1]$ . At the highest possible value of the travel cost ( $t = 1$ ), the effort cost reduces to  $-\alpha e^2$ , which is exactly the cost faced by the scientist choosing a local co-author (as we will see in the next subsection). The first order conditions with respect to  $e$  and  $t$  are, respectively:

$$\begin{aligned} e &= \frac{\gamma t \theta - B}{2\alpha \theta - s} \\ t^3 &= \frac{\alpha e^2}{2\beta} \end{aligned}$$

By plugging the former in the latter, we obtain the following optimal amount of travel:

$$t(s) = \frac{1}{2\alpha\beta} \left(\frac{\gamma}{2}\right)^2 \left(\frac{\theta - B}{\theta - s}\right)^2 \in [0, 1]$$

Plugging this back into FOC( $e$ ), we get:

$$e(s) = \frac{\gamma^3}{16\alpha^2\beta} \left(\frac{\theta - B}{\theta - s}\right)^3$$

Plugging these two last equations into the profit function, we obtain:

$$\pi_D^*(s) = \frac{1}{4\alpha^2\beta} \left(\frac{\gamma}{2}\right)^4 \left(\frac{\theta - B}{\theta - s}\right)^4 - s \quad (32)$$

## E2 Local Co-Author

The profit function under a local co-author is:

$$\pi_L(e) = \gamma \frac{\theta - B}{\theta - s} e - \alpha e^2 - s,$$

where there are no travel costs. In the local pool of co-authors, one can only find up to a given limit of co-author specialization ( $s_L$ ).

The FOC with respect to effort is:

$$e(s) = \frac{\gamma}{2\alpha} \frac{\theta - B}{\theta - s}$$

Plugging this equation into the profit function, we obtain:

$$\pi_L^*(s) = \frac{1}{\alpha} \left(\frac{\gamma}{2}\right)^2 \left(\frac{\theta - B}{\theta - s}\right)^2 - s$$

### **E3 Case of Non-Binding $s_L$ : Comparison of Returns Under Local vs Distant Co-Authors**

This case occurs for projects with a low level of  $\theta$ , i.e. not too specialized ones. As the optimal co-author in terms of specialization for these projects is rather low, the local market is perfectly capable of providing such a co-author. For this reason, there is intuitively no need to go to the distant market. This argument is illustrated in equations below.

By plugging the equation for the optimal amount of travel,  $t(s) = \frac{1}{2\alpha\beta} \left(\frac{\gamma}{2}\right)^2 \left(\frac{\theta - B}{\theta - s}\right)^2 \in [0, 1]$  into the profit functions of both distant and local co-authors, we can easily compare profit levels for these two co-author types. For distant co-authors,

$$\begin{aligned} \pi_D^*(s) &= \frac{1}{4\alpha^2\beta} \left(\frac{\gamma}{2}\right)^4 \left(\frac{\theta - B}{\theta - s}\right)^4 - s \\ &= \frac{1}{2\alpha} \left(\frac{\gamma}{2}\right)^2 \left(\frac{\theta - B}{\theta - s}\right)^2 t(s) - s \\ &= \beta t(s)^2 \end{aligned}$$

For local co-authors,

$$\begin{aligned} \pi_L^*(s) &= \frac{1}{\alpha} \left(\frac{\gamma}{2}\right)^2 \left(\frac{\theta - B}{\theta - s}\right)^2 - s \\ &= \beta 2t(s) \end{aligned}$$

Under the same level of co-author specialization ( $s$ ), the local co-author will always be

preferred:

$$\begin{aligned}
& \pi_L^*(s) > \pi_D^*(s) \\
\iff & 2\beta t(s) > \beta t(s)^2 \\
\iff & 2 > t(s)
\end{aligned}$$

*Q.E.D.*

Intuitively, this is driven by the fact that collaborating with the local co-author does not add travel costs. Therefore, if the local pool offers a degree of specialization that is enough for the project, there is no need to search for co-authors over distance.

## E4 Case of Binding $s_L$ : Comparison of Returns Under Local vs Distant Co-authors

If the upper bound of local specialization is below the optimal desired level of co-author specialization for a given project, then it is *a priori* not clear anymore if local co-authors will be preferred. By working with a local co-author, the scientist saves travel costs but does not reach the desired specialization level for the projects (i.e.  $s_L < s$  leading to  $t(s_L) < t(s)$ ), which harms the overall returns due to the complementarity between project specialization and co-author fit.

If the specialization gap under a local co-author is too large, then switching to the distant co-author will be the optimal choice:

$$\begin{aligned}
& \pi_L^*(s) < \pi_D^*(s) \\
\iff & 2\beta t(s_L) < \beta t(s)^2 \\
\iff & 2t(s_L) < t(s)^2 \\
\iff & 2 \left( \frac{\theta - B}{\theta - s_L} \right)^2 < \frac{1}{2\alpha\beta} \left( \frac{\gamma}{2} \right)^2 \left( \frac{\theta - B}{\theta - s} \right)^4 \\
\iff & \frac{(\theta - s)^4}{(\theta - s_L)^2} < \frac{1}{4\alpha\beta} \left( \frac{\gamma}{2} \right)^2 (\theta - B)^2
\end{aligned}$$



For example, if  $t(s_L) = 0.3$  and  $t(s) = 0.8$ , then a distant co-author is preferred (i.e. if the project's level of specialization is high enough for it to require a co-author that differs substantially from the ones available locally).

Importantly, one can easily prove  $\frac{\delta t(s)}{\delta \theta} > 0$ : it is more likely for scientists to choose distant co-authors for projects with high  $\theta$ , i.e. more specialized projects. These are therefore the projects that benefit disproportionately from lower travel costs.