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MICROECONOMIC EVIDENCE ON PRICE-SETTING

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### **ABSTRACT**

The last decade has seen a burst of micro price studies. Many studies analyze data underlying national CPIs and PPIs. Others focus on more granular sub-national grocery store data. We review these studies with an eye toward the role of price setting in business cycles. We summarize with ten stylized facts: Prices change at least once a year, with temporary price discounts and product turnover often playing an important role. After excluding many short-lived prices, prices change closer to once a year. The frequency of price changes differs widely across goods, however, with more cyclical goods exhibiting greater price flexibility. The timing of price changes is little synchronized across sellers. The hazard (and size) of price changes does not increase with the age of the price. The cross-sectional distribution of price changes is thick-tailed, but contains many small price changes too. Finally, strong linkages exist between price changes and wage changes.

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## 1. Introduction

Recent years have seen a wealth of rich micro price data become available. Many studies have examined data underlying nationally-representative consumer and producer price indices from national statistical agencies. A smaller set of studies have focused on finer scanner data for a subset of stores or products. The U.S. and Western European countries have received the most attention, but evidence on emerging markets has grown rapidly.

Such micro data offer many insights on the importance of price stickiness for business cycles. We review the literature by stating a series of ten facts modelers may want to know about price setting.

First, individual prices change at least once a year. The frequency is more like twice a year in the U.S. versus once a year in the Euro Area. Thus we need a “contract multiplier” to explain why real effects of nominal shocks appear to last several years.

Second, temporary price discounts (“sales”) and product turnover are important to micro price flexibility. This is particularly true in the U.S., which plays a role in its greater price flexibility than in the Euro Area. We provide evidence that such sale prices partially cancel out with cross-sectional and time aggregation, but appear to contain macro content.

Third, if one drops a broad set of short-lived prices (i.e., more than just temporary price discounts), a stickier “reference” price emerges that changes about once a year in the U.S. This filtering conceals considerable novelty in non-reference prices, and these deviations could be responding to aggregate shocks as they do not seem to wash out with aggregation. Still, reference price inflation is considerably more persistent than overall inflation, perhaps suggesting some sort of sticky plan and/or sticky information.

Fourth, goods differ greatly in how frequently their prices change. At one extreme are goods that change prices at least once a month (fresh food, energy, airfares), and at the other extreme are services that change prices much less often than once a year. Such heterogeneity makes mean price durations much longer than median durations, and could help explain a big contract multiplier if combined with strategic complementarities.

Fifth, goods with more cyclical quantities (e.g., cars and apparel) exhibit greater micro price flexibility than goods with little business cycle (e.g., medical care). Durables, as a whole, change prices more frequently than nondurables and services. Including temporary price changes, nondurables change price more frequently than services. Such non-random heterogeneity in price stickiness may hold down the contract multiplier.

Sixth, micro price changes are, on average, much bigger than needed to keep up with aggregate inflation, suggesting the dominance of idiosyncratic forces (intertemporal price discrimination, inventory clearance, etc.). In state-dependent pricing models, price changers can be selected on their idiosyncratic shocks, thereby speeding price adjustment and depressing the contract multiplier. Micro evidence exists for such selection, but not as strong as predicted by models with a single menu cost. For example, many price changes are small, as with time-dependent or information-constrained pricing.

Seventh, relative price changes are transitory. Idiosyncratic shocks evidently do not persist as long as aggregate shocks do. Sellers may be implementing price changes for temporary, idiosyncratic reasons while failing to incorporate macro shocks (e.g., as in rational inattention models).

Eighth, the timing of price changes is little synchronized across products. Most movements in inflation (from month to month or quarter to quarter) are due to changes in the

size rather than the frequency of price changes. This may be a byproduct of the stable inflation rates in the past few decades in the U.S. and Euro Area. In countries with more volatile inflation, such as Mexico, the frequency of price changes has shown more meaningful variation. This lack of synchronization is consistent with the importance of idiosyncratic pricing considerations over macro ones. When combined with strategic complementarities, price staggering paves the way for coordination failure and a high contract multiplier. It is also consistent with rational inattention toward macro shocks. Perhaps related, consumer price changes (both increases and decreases) have increased noticeably in the recent U.S. recession.

Ninth, the hazard rate of price changes falls with the age of a price for the first few months (mostly due to sales and returning to regular prices), and is largely flat thereafter (other than a spike at one year for services). This finding holds in the U.S. and Euro Area, and for both consumer and producer prices. Such a pattern is consistent with a mix of Calvo and Taylor time-dependent pricing, but can also be generated under state-dependent pricing. Meanwhile, the size of price changes is largely unrelated to the time between price changes. This fact seems more discriminating, and favors state-dependent over time-dependent pricing. If price spell length is exogenous, more shocks should accumulate and make for bigger price changes after longer price spells. Under state-dependent pricing, longer price spells reflect stable desired prices rather than pent-up demand for price changes.

Tenth and finally, price changes are linked to wage changes. Firms in labor-intensive sectors adjust prices less frequently, potentially because wages adjust less frequently than other input prices. Furthermore, survey evidence suggests synchronization between wage and price adjustments over time. Thus, in addition to contributing directly to a higher

contract multiplier, wage stickiness may be contributing indirectly by lowering the frequency of price changes.

We organize the rest of this chapter as follows. Section 2 briefly outlines the micro data sources commonly used in the recent literature. Section 3 discusses evidence on the frequency of price changes. Section 4 describes what we know about the size of price changes. Section 5 delves into price setting dynamics – for example, synchronization and what types of price changes cancel out with aggregation across products and time. Section 6 reviews, at greater length, the ten stylized facts we just discussed. Section 7 offers conclusions.

## **2. Data Sources**

The recent literature studies data underlying consumer (CPI) and producer (PPI) price indices, scanner and online data collected from retailers, and information gleaned from surveys of price setters. In this section we briefly describe these datasets.

Until recently, empirical evidence on price-setting at the microeconomic level was somewhat limited, consisting mostly of studies that focused on relatively narrow sets of products (e.g., Carlton 1986, Cecchetti 1986, and Kashyap 1995).<sup>1</sup> This changed as datasets underlying official CPIs and PPIs became available to researchers. These datasets, compiled by national statistical agencies, contain a large number of monthly price quotes tracking individual items over several years or more. The samples aim to be broadly representative – in terms of products, outlets, and cities covered – of national consumer expenditure (or industrial production). For example, the CPI Research Database (CPI-RDB), maintained by

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<sup>1</sup> Wolman (2007) provides a comprehensive survey of the older literature, while Mackowiak and Smets (2008) also survey the more recent literature.

the U.S. Bureau of Labor Statistics (BLS), contains prices for all categories of goods and services other than shelter, or about 70% of consumer expenditure. It begins in January 1988 and includes about 85,000 prices per month (Klenow and Kryvtsov, 2008 and Nakamura and Steinsson, 2008a).

Although the CPI and PPI datasets are alike in many ways, Nakamura and Steinsson (2008a) point out that interpreting the PPI data is somewhat more complicated than interpreting evidence on consumer prices.<sup>2</sup> First, the BLS collects PPI data through a survey of firms rather than a sample of “on-the-shelf” prices. Second, the definition of a PPI good is meant to capture all “price-determining variables”, which often include the buyer of the good. Intermediate prices may be part of (explicit or implicit) long-term contracts, and thus observed prices might not reflect the true shadow prices faced by the buyer (Barro, 1977). Related, in wholesale markets the seller may choose to vary quality margins, such as delivery lags, rather than varying the price (Carlton, 1986). Mackowiak and Smets (2008) point out that repeated interactions (say, for legal services) and varying quality margins (say, waiting in order to purchase a good at the published price) are also present in some retail markets.

A critical open question for macroeconomists in interpreting prices is whether they conform to the Keynesian sticky price paradigm of “call options with unlimited quantities.”<sup>3</sup> On-the-shelf consumer prices may have this feature if they are available in inventory (see Bils, 2004, on stockouts in the CPI). Gopinath and Rigobon (2008) say import prices usually appear to be call options for buyers. Still, unlike for consumer prices, it is not clear whether new buyers of producer goods have the option to buy at prices prevailing for existing buyers.

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<sup>2</sup> The challenges described are for United States PPI data, but Euro Area PPI data display similar features (Vermeulen et al., 2007).

<sup>3</sup> We are grateful to Robert Hall for this phrase.



Tables 1 and 2 list several studies that have made use of CPI and PPI data, respectively. These include studies for the United States, for countries in the Euro Area (Austria, Belgium, Finland, France, Germany, Italy, Luxembourg, the Netherlands, Portugal, and Spain), and for a handful of other developed (Denmark, Israel, Japan, Norway, South Africa) and developing economies (Brazil, Colombia, Chile, Hungary, Mexico, Sierra Leone, Slovakia). Although differences in methodology and coverage make cross-country comparisons challenging, the Inflation Persistence Network (IPN) has coordinated efforts of the many researchers in the Euro Area to allow for such comparisons (Dhyne *et al.* 2006 and Vermeulen *et al.* 2007).

A related set of studies has made use of micro data the BLS collects to construct import and export price indices for the United States. These include Gopinath and Rigobon (2008), Gopinath, Itskhoki, and Rigobon (forthcoming), Gopinath and Itskhoki (forthcoming), and Nakamura and Steinsson (2009). The prices are collected from surveys of importing firms and thus represent wholesale markets. One benefit to using international data is the ability to analyze price-setting behavior in response to large, identified shocks (i.e., nominal exchange rate shocks).

Another source of microeconomic evidence on pricing comes from scanner (i.e., barcode) data collected from supermarkets, drugstores, and mass merchandisers. These data cover a narrower set of goods than the data underlying price indices, but they provide deeper information. Scanner data usually cover many more items per outlet, and often contain information on quantities sold (and sometimes wholesale cost). Data are usually collected at a weekly frequency and may come from one particular retailer (e.g., Eichenbaum, Jaimovich

and Rebelo, 2009, or studies using Dominick’s data) or from multiple retailers (e.g., through AC Nielsen). A number of these studies are listed in Table 3.

Other researchers have begun collecting price information from retailers by “scraping” prices from websites. The ongoing “Billion Prices Project” of Cavallo and Rigobon (e.g., Cavallo, 2009) collects daily prices from numerous retailers in over 50 countries. Useful aspects of this dataset include the daily frequency, comparability across many countries, and detailed information on each product including sales and price control indicators. Lünnemann and Wintr (2006) is another example.

A final source of microeconomic information comes from surveying firms about their price-setting practices, as opposed to collecting longitudinal information about individual prices. These surveys allow researchers to ask about aspects of pricing that cannot be captured from datasets of observed prices, such as the frequency with which price setters review prices and the importance of particular theories of price stickiness for explaining their pricing decisions. Blinder et al. (1998) was a pioneering study for the United States, and subsequent surveys have been conducted in many countries, as shown in Table 4. The surveys typically ask firms to focus on their main product (or most important products).

### **3. Frequency of Price Changes**

We begin our review of the substantive findings of the literature by looking at the frequency with which prices change. A theme that will arise throughout the paper is the presence of a great deal of heterogeneity in price-setting behavior, and we therefore report results along several dimensions. These include measures of the “average” time between price changes, how these measures vary across different samples and types of goods, the

importance of temporary sales and product turnover, and some discussion of the determinants of the frequency of price change.

### *3.1 Average Frequency*

Table 1, drawn primarily from a survey by Álvarez (2008), presents estimates of the *mean* frequency of price changes obtained from the datasets underlying national CPIs.<sup>4</sup> Prices clearly exhibit nominal stickiness, as the (unweighted) median across these studies for the estimated mean frequency of price change is 19% (per month). The degree of stickiness varies considerably across countries, with prices in the Euro area appearing to change less frequently than those in the U.S., which in turn change less frequently than those in high-inflation developing countries (Brazil, Chile, Mexico, Sierra Leone, Slovakia). We will return to the question of what explains these cross-country differences after we take a closer look at individual country studies.

We will give particular attention to the multiple U.S. CPI studies (Bils and Klenow 2004, Klenow and Kryvtsov 2008, and Nakamura and Steinsson 2008a) in order to shed light on a number of features of the data and provide understanding on how different methodologies impact results. Moreover, since we have access to the micro data from the BLS, we will be able to construct some new results as we proceed. We begin by describing the structure of the micro data. The BLS divides goods and services into 300 or so categories of consumption known as Entry Level Items (ELIs). Within these categories are prices for particular products sold at particular outlets (which we will refer to as “quote-lines”). The

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<sup>4</sup> For studies that contain information on price changes due to temporary sales, we report the frequency for both all prices (in parentheses) and non-sale prices. In many countries, the prices that are reported during sales periods are prices without rebates (i.e., posted prices *are* essentially non-sale prices), and we thus use the non-sale prices when we describe results across countries.

BLS collects prices monthly for all products in the three largest metropolitan areas (New York, Los Angeles, and Chicago) and for food and fuel products in all areas, and bimonthly for all other prices. The statistics we report from Klenow and Kryvtsov (2008) and Nakamura and Steinsson (2008a) are for prices collected monthly from the top three cities.<sup>5</sup>

To construct their average monthly frequency that we report in Table 1, Klenow and Kryvtsov (2008) first estimate frequencies for each ELI category and then take the weighted *mean* across categories to arrive at a figure of 36.2% (for posted prices) between 1988 and early 2005. Of course, this is not the only possible measure of the “average” frequency of price change. The weighted *median* frequency of price changes is 27.3%. The mean is higher than the median because the distribution of the frequency of price changes across ELIs, shown in Figure 1 for 1998-2009 data, is convex (Jensen’s inequality). Related, the mean implied duration (from the mean of the inverse frequencies across ELIs) of 6.8 months is higher than the median (the inverse of the median frequency) of 3.7 months.

Turning to producer prices, the median country in Table 2 has a mean frequency of price change of 23%. Nakamura and Stiensson (2008a) compare price flexibility with consumer goods by matching 153 ELI categories from the CPI with product codes from the PPI. In general, they find the frequency of price change for producer prices to be similar to that of consumer prices excluding sales. Goldberg and Hellerstein (2009), however, report a higher frequency – closer to consumer prices including sales – and attribute the difference to weighting products by their use of BLS firm and industry weights. Using these weights makes a large difference because larger firms change prices more frequently than small firms,

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<sup>5</sup> Both studies weight ELIs by BLS estimates of their importance in consumer expenditures. Klenow and Kryvtsov also use some BLS weighting information for items within ELIs; Nakamura and Steinsson do not, but it does not seem to affect statistics such as the median duration of prices across ELIs.

they find. Vermeulen et al. (2007) document producer price setting in six European countries. Not controlling for the composition of the CPI and PPI baskets, they find producer prices change somewhat more frequently than what Dhyne et al. (2006) reported for consumer prices. This pattern persists when they focus on similar products in the ‘processed food’ and ‘non-food non-energy consumer goods’ sectors.

Gopinath and Rigobon (2008) find that U.S. export and import (wholesale) prices are sticky in the currency in which they are reported. The median implied duration in the currency of pricing is 10.6 (12.8) months for imports (exports) during their 1994-2005 sample.<sup>6</sup> They go on to compare the duration of their cross-border transactions to domestic transactions by using product category descriptions to match international price (IPP) categories with PPI categories. Restricting their sample to these matched categories (69 of them), they find a mean duration of 10.3 months for the IPP and 10.6 months for the PPI.

Table 3 provides frequencies from a number of scanner data studies. Although the underlying data is weekly, the numbers in the table are monthly frequencies of price change (with the exception of Eichenbaum et al., 2009, who report a weekly frequency). Frequencies vary, even across studies using the same dataset, because of different sample choices and reported measures. For example, for Dominick’s, Midrigan (2009) reports a mean frequency for one store, while Burstein and Hellwig (2007) consider many stores and report the frequency of the median product category. Despite these differences, the studies have very similar results: average posted (non-sale) prices change at least every three (five) months.

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<sup>6</sup> These numbers correspond to their benchmark specification, in which price changes across non-adjacent prices and product substitutions are included and the frequencies of goods whose price never changes are adjusted by the probability of discontinuation.

Table 4 presents information on price flexibility that comes from asking firms how frequently they changed their prices in the past year (or on average in recent years). The firms surveyed tend to sell their main product to other firms, and thus, the survey data pertains primarily to producer prices. The median frequency of price change, about once a year in most countries, exhibits more stickiness than the PPI micro data, although the results are not directly comparable due to different time periods, samples of firms, etc.

To recap, prices do not change continuously but do change “on average” *at least* once a year. We use “on average” in a loose sense, as we have already seen that the complexity of the micro data make it difficult to summarize with one statistic (such as the mean or median). We now explore this complexity in more detail, first investigating heterogeneity in the frequency with which prices change across different types of goods and then discussing the treatment of sales and product substitution.

### *3.2 Heterogeneity*

Figure 1 shows a tremendous amount of heterogeneity across ELI categories, as the price change frequency ranges from 2.7% for “Intracity Mass Transit” to 91% for “Regular Unleaded Gasoline”. Indeed, while half of all prices have an implied duration less than 3.4 months, almost a fifth last longer than a year. This heterogeneity also helps to explain the finding that “average” consumer prices adjust more frequently than in the narrower investigations predating the latest generation of micro studies. For example, Cecchetti (1986) found the length of time between changes in the newsstand prices of US magazines ranged from 1.8 to 14 years, but Nakamura and Steinsson (2008a) find that “Single-Copy Newspapers and Magazines”, with a duration of 17.2 months in their sample, change prices less frequently than 84% of non-housing consumption.

Table 5 illustrates the heterogeneity in the frequency of price changes in additional ways. It reports the weighted median and mean implied price durations in the U.S. CPI from January 1988 through October 2009 separately for posted and regular (i.e., non-sale) prices covering: (a) All Items; (b) Durables, Nondurables, and Services; (c) Raw and Processed goods; and (d) eight Major Groups. For conciseness, consider the mean durations of posted prices. For Durables the mean price duration is 3.0 months, whereas for Nondurables it is 5.8 months and for Services it is 9.4 months. For raw goods (energy and food commodities) prices last about 1.1 months, whereas for processed goods and services it is 6.9 months. Among Major Groups, price durations range from 2.9 months in Apparel to 14.7 months in Other Goods and Services.

We further explore the connection between durability and price change frequency at a more disaggregated level in U.S. consumer prices. For each of 65 Expenditure Classes of the CPI (more aggregated than the 300+ ELIs, less aggregated than the Major Groups), we were able to estimate durability from the data in Bils and Klenow (1998). For interpretability, in Figure 2 we aggregate back up to the Major Groups. The figure plots the average frequency of posted price changes against average durability in years, with each dot proportional to the group's average expenditure weight. Transportation stands out as durable, flexible, and important. For example, Motor Vehicles have a high weight (16% of the non-shelter sample weight), high durability (9 years), and high frequency (38% per month). Food, on the other hand, is *nondurable*, flexible and important. As a result, Table 6 reports no significant relationship between frequency and durability across the 65 Expenditure Classes: the weighted least squares estimate is 0.60 percentage points (standard error 0.69). Excluding the six Expenditure Classes with raw goods (fresh food and energy) – which are nondurable,

flexibly-priced and often dropped from the data for business cycle analysis – the relationship becomes significantly positive. For processed goods, each year of durability goes along with 1.47 percentage points higher frequency (standard error 0.41), so that a good lasting 10 years tends to have 14 percentage points higher price change frequency than a nondurable. The connection is similar for regular price changes among processed goods.

The positive correlation between durability and price flexibility (at least for processed goods and services) could have important implications for business cycles. The more durable the good, the more cyclical expenditures and production tend to be. This is true in both theory and practice (see Bils and Klenow 1998 for one example). Barsky, House and Kimball (2007) present a model in which monetary non-neutrality is closely connected to price stickiness for durables, with the stickiness of nondurables of no importance. That said, we hasten to reiterate that the relationship in the data is not significant if raw good categories are included.

Durability and cyclicalities are not synonymous, of course. We therefore gauged cyclicalities directly for each of 64 BLS Expenditure Classes based on the coefficient from regressing its NIPA real consumption expenditure growth on NIPA aggregate real consumption growth. For each Expenditure Class we ran a single OLS regression (with a constant) for the available NIPA sample from February 1990 through March 2009, which is close to our CPI-RDB 1988-2009 monthly sample. Figure 3 shows a clear positive association between price change frequencies and cyclicalities across the Major Groups. Inside Transportation, for example, Motor Vehicles stands out in terms of its combination of price flexibility (38% frequency), cyclicalities (5.7% higher expenditure growth for every 1% higher aggregate consumption growth), and sampling weight (16% of the non-shelter



sample). Apparel is also fairly flexible (30% frequency if one includes sale prices) and fairly cyclical (cyclicality coefficient 1.75). In Table 6, the WLS regression coefficient of price frequency on cyclicality across the 64 Expenditure Classes is 3.23 percentage points (standard error 1.05). If we look only at the 58 Expenditure Classes for processed goods, the WLS coefficient is similar at 3.29 (but more precisely estimated, with a standard error of 0.59). The relationship is stronger still for regular price changes across processed goods at 3.79 (standard error 0.48).<sup>7</sup>

The frequency-cyclicality nexus could arise because cyclicality induces price changes, the other way around, or because they share driving forces. If price flexibility is responding to cyclical shocks, then the pattern may suggest more macro price flexibility than in models where the variation in price change frequency across goods is exogenous, as in Carvalho (2006), or reflects variation in non-cyclical factors, as in Nakamura and Steinsson (2008b). Causality running from frequency to cyclicality would presumably work for “supply” shocks (for which price flexibility should amplify the response of real expenditure growth) but not for “demand” shocks (for which price flexibility would dampen the response of real expenditure growth) – see Bils, Klenow and Kryvtsov (2003).<sup>8</sup>

Heterogeneity is also evident for producer prices. Nakamura and Steinsson (2008a) report a median implied duration of 8.7 months for finished producer goods, 7.0 months for intermediate goods, and 0.2 months for crude materials from 1998 to 2005. Within finished

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<sup>7</sup> Frequency is positively correlated with cyclicality across categories even when controlling for durability. When we regress the frequency of regular price changes for processed goods on durability and cyclicality across 64 ECs, the coefficient on durability is 0.34 (s.e. 0.05) and the coefficient on cyclicality is 9.24 (s.e. 1.39).

<sup>8</sup> We do not know how the analysis would change with shelter. On the one hand, rents and owner equivalent rents may be sticky. On the other hand, shelter *quantities* may be just as sticky – contrary to the Keynesian paradigm of flexible quantities relative to prices. And housing services are presumably less cyclical than housing construction, for which prices may be more flexible.

producer goods, the median frequency of price change ranges from 1.3% for “Lumber and Wood Products” to 87.5% for “Food Products”. Vermeulen et al. (2007) also document significant heterogeneity across sectors and investigate the causes. Firms with a higher labor share in total costs tend to change price less frequently, whereas firms with higher shares of energy and non-energy intermediate goods change price more frequently. Moreover, they find that the higher the degree of competition, the higher is the frequency of price changes, particularly price decreases.

Another dimension of heterogeneity that apparently affects the frequency of price changes is the type of establishment at which goods are sold. In Europe, consumer prices are more flexible in large outlets, such as supermarkets and department stores, than in smaller retail outlets (e.g., Jonker *et al.* 2004, Dias *et al.* 2004, and Fabiani *et al.* 2006). For U.S. producer prices, Goldberg and Hellerstein (2009) find that large firms change prices two to three times more frequently than small firms. Survey studies have also reported similar patterns (e.g., Amirault *et al.* 2006, Buckle and Carlson 2000, and Fabiani *et al.* 2005).

### *3.3 Sales, Product Turnover, and Reference Prices*

One lesson from the theoretical price-setting literature is that different types of price adjustments (e.g., transitory or permanent, selected or random) can have substantially different macroeconomic implications. Researchers have thus investigated how measures of the frequency of price change are altered when the data is filtered in different ways, such as excluding temporary sales and product turnover. It turns out that the answer can vary considerably depending on how exactly this is done.

In the U.S. CPI data, a “sale” price is (a) temporarily lower than the “regular” price, (b) available to all consumers, and (c) usually identified by a sign or statement on the price tag. Klenow and Kryvtsov (2008) report that roughly 11% of the prices in their sample are identified as sale prices by BLS price collectors. Another approach is to use a “sales filter” to identify “V-shaped” patterns in the data as sales. Nakamura and Steinsson (2008a) report results for a variety of sales filters, allowing for asymmetric and multi-period V’s. Concerning product turnover, “forced item substitutions” occur when an item in the sample has been discontinued from its outlet and the price collector identifies a similar replacement item in the outlet to price going forward, often taking the form of a product upgrade or model changeover. The monthly rate of force item substitutions is about 3% in the BLS sample.<sup>9</sup>

Table 7 demonstrates the impact on the implied duration of prices of applying various filters to the data. We follow Klenow and Kryvtsov (2008) but with a U.S. CPI sample that extends through October 2009 (rather than January 2005). Depending on the treatment of sales, the median duration of prices ranges from 3.4 months (all prices included) to 6.9 months (excluding BLS-flagged sales). The one-period V-shape filter – in which every time a middle price is lower than its identical neighbors it is replaced by its neighbors – produces an intermediate duration of 5.0 months, reflecting that many sales, such as clearance sales, are not V-shaped.<sup>10</sup> “Like” prices compare a regular price only to the previous regular price and a sale price only to the previous sale price, thus allowing for the possibility that sale prices are sticky even if they do not return to the previous regular price. This raises the implied median duration to 5.9 months. The fact that “like” prices change more frequently

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<sup>9</sup> See Broda and Weinstein (2007) for the importance of product turnover in AC Nielsen Homescan data.

<sup>10</sup> Nakamura and Steinsson (2008) found that removing BLS-flagged sales generated a higher estimate of price duration than did any of their V-shaped sales filters.

than regular prices (every 5.9 months vs. every 6.9 months) indicates a sale price is more likely to differ from a previous sale price than a regular price is to differ from a previous regular price.

Table 7 also shows that removing all forced item substitutions from the data increases the median implied duration to 8.3 months from 6.9 months. Klenow and Kryvtsov (2008) report that item substitutions display price changes about 80% of the time, much more frequently than the average over a product's life cycle. Purging substitution-related price changes can imply prices have a longer duration than the products themselves: for apparel, regular prices excluding substitutions change about every 27 months, whereas the average item lasts only 10 months. Finally, comparing only *consecutive* regular monthly prices between substitutions pushes the duration up from 8.3 months to 9.0 months. Price changes are more frequent after items return to stock, come back into season, or return from sales.

Nakamura and Steinsson (2008a) underscore that sales and forced item substitutions are much more important in some categories than others in the U.S. CPI. For example, 87% of price changes in Apparel and 67% in Home Furnishings are sale-related price changes, while Utilities, Vehicle Fuel, and Services have virtually no sale-related changes. The monthly rate of forced item substitutions is about 10% in Apparel and in Transportation, compared to 3% for all goods. This uneven distribution of sales and substitutions is important for explaining why excluding them has a sizable impact on the implied median duration of U.S. consumer prices: the sectors in which sales and substitutions are concentrated are those with a frequency of price change that is relatively close to the median.

Sales have become more important over time for U.S. consumer prices but continue to play a small role in other countries. Nakamura and Steinsson (2008a) document a strong

increase in the frequency of U.S. sales from 1988 to 2005, especially in processed food and apparel where the frequency of sales doubled. Available evidence from European countries suggests that sales are a less important source of price flexibility. Wulfsberg (2009) reports that sale prices account for only 3% of price observations in Norwegian CPI data, and removing these observations increases the mean duration by only 0.3 months. Dhyne et al. (2006) similarly report that sales have small effects on the estimated frequency of price change in France and Austria.

As emphasized by Mankiw and Reis (2002), Burstein (2006), Woodford (2009) and others, price changes may be part of a sticky plan and hence fail to incorporate current macro information. This can be true of regular price changes, not just movements between regular and sale prices. One sign of such a plan might be the existence of only a few prices over the life of an item. In the top three cities of the U.S. CPI, the weighted median (mean) length of a quote-line is 50 (53) monthly prices. In Table 8 we report the cumulative share of price quotes represented by the “top” (i.e., most common) 1, 2, 3 and 4 prices over the quote-line. The median (mean) share of the most common price is 31% (38%). The top four prices over the typical 4.3 year quote-line together represent a median (mean) of 70% (66%) of all prices. Table 8 reports the figures for Major Groups as well. There is a natural tendency for higher shares where there are shorter quote-lines (apparel) or less frequent nominal price changes (medical care, recreation). Relative to its moderate frequency of price changes, Food stands out in having a 42% median for the top price and 86% median for the top 4 prices. Even Transportation – which is highly cyclical and exhibits frequent price changes – has a 16% median share of the top price and 46% median share of the top 4 prices.

Given that nominal prices change every four months or so, there are on average around 13 prices per quote-line. A Taylor model (and perhaps Calvo model as well) would therefore imply a *median* top 2 price share of less than 20%, whereas the actual share in the data is over 50%.<sup>11</sup> The disproportionate importance of a few prices appears supportive of downward-sloping hazards and/or sticky nominal plans.<sup>12</sup> In favor of the former, two-thirds of top price spells are uninterrupted by other prices. Before discarding less common prices, however, research could explore how aggregate quantities produced and sold relate to changes in common vs. rare prices. It is conceivable that cyclical quantities are sensitive to the rare prices (e.g., clearance sales in apparel).

Eichenbaum, Jaimovich and Rebelo (2009) usefully propose a way of measuring sticky “reference” prices amidst shorter-lived new prices. Using weekly price data covering 2004-2006 from a large U.S. supermarket chain, they define the reference price for each UPC as the modal price in each quarter. They find such reference prices are responsible for 62% of all weekly prices and 50% of quantities sold. Importantly, they report that reference prices only change every 11.1 months. This is considerably stickier than regular (non-sale) prices in the same supermarket, which change about once per quarter. They go on to present a simple model in which the frequency of reference price changes is the key to monetary non-neutrality, as deviations from reference prices largely reflect idiosyncratic considerations.

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<sup>11</sup> The prediction of a menu cost model for the top price shares would be more sensitive to the distribution and timing of large idiosyncratic changes in the desired price. In Golosov and Lucas (2007) high variance shocks are realized every period, whereas in Gertler and Leahy (2008) they follow a Poisson process.

<sup>12</sup> In addition to costs of collecting and processing information and formulating and implementing new plans, the use of a few prices may reflect “price points.” Levy et al. (2007) find prices ending in 9 are most common (whether in cents, dollars, or tens of dollars), less likely to change, and change by bigger amounts.

Do sticky reference prices exist in the U.S. CPI more broadly? The CPI data is monthly, so it is not possible to implement the exact Eichenbaum et al. methodology on the CPI. We instead defined the reference price in each month (for an item) as the most common price in the 13-month window centered on the current month. We broke ties in favor of the current price. An advantage of a rolling window is that it allows the reference price to change every month, whereas the calendar definition imposes at most one reference price change per quarter. Using this 13-month window, we find that the posted price equals the reference price 78.5% of the time on average (84.2% median) when looking across weighted quote-lines. Table 9 provides reference price statistics for all items and Major Groups. The share of reference prices is modestly higher in Food (88.1% median). The only Major Group with a reference price share below 70% is Transportation (62.5% median); albeit an important exception given its cyclicity. We find that the weighted median duration of reference prices is 11.0 months across ELI's in the CPI. Note that reference prices change less frequently than regular prices (median duration 6.9 months), so that some regular price changes must be temporary deviations from reference prices. The median duration is higher for Food at 13.5 months, and lower for Transportation at 6.1 months. We conclude that the Eichenbaum et al. reference price phenomenon extends not only to most food items, but to most items in the non-shelter CPI more generally.

We add several caveats to our reference price results for the U.S. CPI. First, our definition of reference prices is not strictly comparable to that of Eichenbaum et al. (2009). With our definition, a combination of a high share and high duration of references – which we do observe – is more suggestive of stickiness than either of these without the other. Second, there is considerable variation in the share of reference prices across quote-lines (weighted

standard deviation 41%), even within product categories (see Table 9). Third, it is possible that cyclical quantities are sensitive to non-reference prices along with reference prices.

A final caveat has to do with the modeling implications that can be drawn from the reference-price findings. Although reference prices constitute a large share of total prices and have long durations, our statistics need not imply that sellers choose prices from a small set or that prices display “memory”. Indeed, in the U.S. CPI, we found that only 30% of deviations from reference prices *ever* return to the previous reference price. We next consider two statistics that may be more directly revealing about these issues.

The first statistic is the fraction of prices which are “novel”, which we define as prices that do not appear in any of the previous 12 months for the same item (quote-line). For the top three cities of the U.S. CPI, we find the weighted median (mean) share of prices that are novel to be 16.1% (25.0%). These fractions are consistent with genuinely new prices every four to seven months. We take this to mean that deviations from the most common prices exhibit considerably novelty. Table 10 also breaks the statistic down by Major Group. Novelty is naturally correlated with the frequency of price changes. Food exhibits less than average novelty (10.3% median vs. the overall item median of 16.1%, 13.8% average vs. the overall average of 25.0%) despite having average frequency of price changes. Thus caution may be warranted in drawing lessons from grocery store scanner datasets. Prices appear more novel in more cyclical categories (Transportation, Apparel). If cyclical quantities are linked to these novel prices, they could well contribute to macro price flexibility.

We also compute the fraction of prices that are “comeback” prices. We define the current price as a comeback price if the same price appeared any time in the previous 12



months with a *different* price occurring at least once in between. As a hypothetical example, we would label a current price of \$10 as a comeback price if the price was stuck at \$10 for the previous six months, was at \$11 seven months ago, but was also at \$10 eight months ago. For the top three cities of the U.S. CPI, Table 11 reports the weighted median (mean) share of comeback prices to be 0.8% (14.0%). The Major Groups with the highest share of comeback prices are Apparel and Food, which have means of around 25%. The share of comeback prices for the median quote-line is zero, however, in five of the eight Major Groups.

The upshot is that, outside of Apparel and Food, there appears to be little memory in monthly U.S. CPI prices. A corollary is that most reference prices are not comeback prices. The typical non-reference price must be a short-lived (transition) price in between successive reference prices, at least as far as we can tell. It is possible that monthly observations are obscuring many temporary price changes we would see if we had weekly or even daily data. This is most plausible for Apparel and Food, where a nontrivial share of comeback prices occur after or during sales (roughly  $\frac{3}{4}$  of comeback prices in apparel, and roughly  $\frac{1}{2}$  of comeback prices in food). It would appear less likely for services, such as medical care, where monthly prices largely go from one reference price to another.

### *3.4 Determinants of Frequency*

Researchers have also investigated factors *affecting* the frequency of price change. Some studies have made use of the substantial variation in frequencies along different dimensions to identify important determinants, while others have directly asked price-setters to assess the importance of various theories of price stickiness. A (non-exhaustive) list of determinants include (a) the level and variability of inflation, (b) the frequency and magnitude

of cost and demand shocks (broadly construed to include price discrimination), (c) the structure and degree of market competition (including regulation of discounts), and d) the price collecting methods of statistical agencies (e.g., do they report temporary sales?).

We begin by looking at the cross-country evidence reported in Table 1. Following Golosov and Lucas (2007) and Mackowiak and Smets (2008), Figure 4 simply plots the mean frequency of price change against the average inflation rate for these studies. The OLS regression coefficient of price frequency on inflation is 17.0 (standard error 6.8). Of course, this exercise comes with a few well-known caveats. First, the studies differ in many details, such as sample composition and different price-collecting methodologies. Second, periods of high inflation are often periods of volatile inflation, so it is unclear whether the relationship in Figure 4 reflects the importance of the level or the variability of inflation (probably both). Still, the relationship is provocative.

Dhyne et al. (2006) investigate the impact of inflation (and other factors) on the frequency of price change by running regressions on European data. They regress the frequency of price change across 50 product categories in 9 countries on dummy variables for product type (unprocessed and processed food, energy, non-energy industrial goods, and services), country dummies, the mean and standard deviation of inflation at the product category level, an indicator for whether sale prices occur and are reported, the share of prices set at attractive levels (“price points”), and an indicator for whether the price is typically regulated. They find that mean inflation is *not* significantly correlated with the overall frequency, but is correlated with the frequency of increases and decreases separately. The overall frequency (and increases and decreases, respectively) are significantly higher in sectors in which the variability of inflation is higher. A higher frequency of price change is

also found when sales and temporary price cuts are included, when the share of attractive prices is lower, and when prices are not subject to regulatory control.

For U.S. consumer prices, Bils and Klenow (2004) consider regressions relating the frequency of price change in different product categories to measures of market structure: the concentration ratio, wholesale markup, and rate of non-comparable substitutions in those categories. After controlling for whether a good is raw or processed, they find that the first two measures do not have statistically significant effects, while the rate of product turnover remains a robust predictor of the frequency of price changes. They interpret the role of product turnover and raw materials in explaining the frequency of price changes as reflecting the importance of the volatility of shocks to the supply of and demand for goods.

Boivin, Giannoni, and Mihov (2009) find a relationship between the volatility of sectoral shocks and the frequency of price change. Using disaggregated PCE inflation series, they disentangle inflation fluctuations due to sector-specific conditions from those due to macroeconomic factors. They find that sectors with more volatile sector-specific shocks have a higher frequency of price change.

Other studies provide evidence consistent with the importance of cost shocks, as goods with more sticky input prices tend to change price less frequently. For example, Nakamura and Steinsson (2008a) report a high correlation between the frequency of price changes upstream (PPI) and downstream (CPI), and Eichenbaum et al. (2009) find a similar pattern using detailed cost and price data for one large U.S. retailer. Looking at Euro Area producer prices and noting that wages tend to be stickier than goods prices, Vermeulen et al. (2007) find that goods with a higher labor share in total costs tend to change price less frequently, whereas firms with a higher share of intermediate goods change prices more

frequently. Peneva (2009) matches categories of U.S. consumer goods to the manufacturing industries in which they are produced and also finds that higher labor intensity is associated with less frequent price changes.

Survey data also provides useful insights on the determinants of price change. Fabiani et al. (2005) report the top four reasons Euro area firms refrain from changing prices include: (1-2) implicit and explicit contracts with customers; (3) cost-based pricing (i.e., input costs are slow to change); and (4) coordination failure (not wanting to raise one's price out of fear of losing market share to competitors who do not follow suit). These reasons were also ranked in the top five by firms in the United States (Blinder et al. 1998), United Kingdom (Hall, Walsh and Yates, 2000), Sweden (Apel, Friberg and Hallsten, 2005) and Canada (Amirault et al. 2006). On the other hand, physical costs (menu costs) and costly information are among the reasons least favored by firms. Finally, the main impediments to more frequent price adjustment do appear to be associated with price *changes* rather than price *reviews* – that is, surveyed firms report reviewing prices much more often than changing prices.

Drilling down further into the survey evidence for cost-based pricing and, in particular, the role of labor costs, Álvarez and Hernando (2007a) report that Spanish sectors with relatively high labor costs tend to contain a small number of firms that change prices often. Druant et al. (2009) provide insight into the relationship between wage and price rigidity based on a survey on wage and pricing policies of Euro Area firms. They find that 40 percent of firms indicate a relationship (formal or informal) between the timing of their wage and price adjustment decisions. Moreover, firms with a higher labor cost share report a tighter link between wage and price changes and a lower frequency of price adjustment (as

wages change less frequently than prices). Finally, even accounting for the likely simultaneity between price and wage changes, a statistically significant relationship is found, running from the frequency of wage changes to that of prices.

Dhyne et al. (2006) try to account for the higher frequency of price changes in the U.S. than in the Euro Area. The U.S. had somewhat higher level and volatility of inflation, but to arguably little effect. Differences in consumption patterns do not help at all, as the expenditure share of more flexible components of the CPI is actually larger in the Euro Area. Heterogeneity of outlets may play a role: smaller shops, which change prices less frequently, have a higher market share in Euro area. Differences in occurrence and treatment of temporary sales are important. For example, 1 in 5 price changes is related to sales in U.S., compared to less than 1 in 8 in France. Many other Euro Area countries do not record sale prices. Finally, a higher variability of wages (and other input prices) and less anti-competitive regulation may help explain the higher frequency of price changes in the U.S.

#### **4. Size of Price Changes**

We now move from the extensive margin – how often prices change – to the intensive margin – how large are the price changes. Again, there is substantial heterogeneity in the micro data, and it is thus useful to characterize the distribution of the size of price changes.

##### *4.1 Average magnitude*

A common finding across studies is that price changes are large on average. For example, in the U.S. CPI Klenow and Kryvtsov (2008) report a mean (median) absolute change in posted prices of 14% (11.5%), while regular price changes are smaller but still

large with a mean (median) of 11% (10%). The average consumer price decrease (increase) is 10% (8%) in the Euro area (Dhyne et al. 2006), and emerging markets also display large changes (e.g., Barros et al. 2009, Konieczny and Skrzypacz 2005). For U.S. finished goods producer prices, Nakamura and Steinsson (2008a) report a median magnitude of 7.7%.

Given the low level and volatility of aggregate inflation in the U.S. and Europe, most price changes are not simply keeping up with overall inflation (i.e., indexed). But perhaps many micro price changes incorporate idiosyncratic or sectoral considerations, but not aggregate shocks. See Mackowiak and Smets (2008) for a further discussion of this rational inattention hypothesis in the context of large micro price changes.

#### *4.2 Increases vs. Decreases*

A second feature of the size distribution is that price declines are very common. Nakamura and Steinsson (2008a) report that around 40% of both CPI and PPI monthly price changes in the U.S. are decreases, and Dhyne et al. (2006) report similar numbers for the Euro area. These facts help reconcile the finding of large average *absolute* price changes with small average price changes (14% vs. 0.8% according to Klenow and Kryvtsov 2008) and suggest an important role for idiosyncratic shocks (or price discrimination) in driving price changes. The prevalence of price declines also varies across sectors; in particular, they are relatively uncommon in the services sector (Dhyne et al. 2006 and Nakamura and Steinsson 2008a).

#### *4.3 Higher Moments of the Size Distribution*

Using scanner data, Midrigan (2009) emphasizes that the distribution of non-zero price changes has more weight in the vicinity of zero than predicted by a normal distribution,

while the tails are somewhat fatter. Formally, the distribution of price changes is leptokurtic (i.e., has positive excess kurtosis). We have confirmed this pattern in the U.S. CPI data, where the kurtosis of the price change distribution is 10.0 for posted prices and 17.4 for regular prices (vs. 3 for a normal distribution). As maintained by Midrigan, fat tails suggest weaker selection, with more price changes large and therefore infra-marginal. For this reason, the fraction of price increases vs. decreases can be less sensitive to macro shocks (see also Gertler and Leahy, 2008).

Other studies have documented the prevalence of small price changes more directly. In the U.S. CPI, around 44% of regular price changes are smaller than 5% in absolute value, 25% are smaller than 2.5%, and 12% are smaller than 1% (Klenow and Kryvtsov 2008). Note that this is not just due to frequent shopper cards (as in scanner data that report average weekly prices inclusive of coupons and frequent shopper discounts). Wulfsberg (2009) reports that 45% of price changes are smaller than 5% in Norway, while in Brazil, where the mean absolute size is 13%, over a third of price changes are smaller than 5% (Barros et al. 2009). For producer prices, Vermeulen et al. (2007) find that a quarter of both increases and decreases are smaller than 1%, compared to a mean price change of 4% in the Euro area.

## **5. Dynamic Features of Price Changes**

In addition to the unconditional statistics that we have highlighted so far, researchers have documented a number of features concerning how prices change over time. These include the synchronization of price changes, how the frequency and size of price changes correlate with the duration of the existing price, and the response of prices to shocks that would be expected to alter a firm's desired price. These features are of interest because in the

presence of nominal stickiness (like we see in the data), price setters have dynamic decision problems, and thus, dynamic features of the data are particularly helpful in distinguishing between the various theories of price setting.

### 5.1 Synchronization

Since at least Taylor (1980), staggered price-setting has played an important role in modeling persistent real effects of monetary shocks. The staggering of price adjustments is readily apparent from the observation that not all prices change in any given period. Moreover, when price setters do change their price, they often take the prevailing price of their competition into consideration (Levy et al. 1998), which only makes sense if (at least some) competitors' prices are expected to remain active for a period of time. Recent studies have looked at time variation in the frequency of price changes as a measure of the extent of synchronization (or, conversely, uniform staggering) in price setting. Time variation also provides relevant evidence for distinguishing between some time-dependent vs. state-dependent pricing models.

Klenow and Kryvtsov (2008) decompose monthly inflation ( $\pi_t$ ) into the fraction ( $fr_t$ ) of items with price changes and the average size ( $sz_t$ ) of those price changes:  $\pi_t = fr_t \times sz_t$ . In their sample (U.S., 1988-2004), they found the fraction to be relatively stable and not so correlated with inflation (correlation 0.25), while the average size was more volatile and had commoved almost perfectly with inflation (correlation 0.99). In addition, they decompose the variance of inflation over time into an “intensive margin” (IM) and “extensive margin” (EM) as follows:



$$\text{var}(\pi_t) = \underbrace{\text{var}(sz_t) \cdot \overline{fr}^2}_{IM} + \underbrace{\text{var}(fr_t) \cdot \overline{sz}^2 + 2 \cdot \overline{fr} \cdot \overline{sz} \cdot \text{cov}(fr_t, sz_t)}_{EM} + O_t.$$

This decomposition is interesting because different models of price-setting have distinct implications for this decomposition. For example, in staggered TDP models, the intensive margin will account for all of inflation's variance, whereas the fraction of items changing price plays a substantial role in some SDP models, such as Dotsey, King and Wolman (1999). Klenow and Kryvtsov (2008) find that the IM term accounts for between 86% and 113% of the variance of inflation, implying that the fraction of items changing price are a relatively unimportant source of fluctuations in inflation, at least for their sample period. Note that other SDP models (such as Golosov and Lucas, 2007), can fit this decomposition; the key is to include (realistically) large idiosyncratic price changes so that aggregate shocks have offsetting effects on the frequency of increases vs. decreases.

Gagnon (2009) first emphasized the usefulness of further decomposing inflation into terms due to price increases and decreases. Note that  $fr = fr^+ + fr^-$  and  $sz = fr^+sz^+ - fr^-sz^-$ , where  $fr^+$  and  $fr^-$  ( $sz^+$  and  $sz^-$ ) denote the frequency (absolute size) of price increases and decreases, respectively. So, even if the average size of price increases and decreases remain constant over time, significant variation in the average size of price changes could result due to offsetting movements in the frequency of price increases and decreases, which would also imply little variation in the overall frequency of price changes. This is the type of pattern we see in the U.S.: Klenow and Kryvtsov (2008) report that a 1 percentage point increase in inflation is associated with a 5.5 (-3.1) percentage point change in the fraction of price increases (decreases), and a 0.6 (-1) percentage point change in the size of price increases (decreases). When they additively decompose the variance of inflation

(splitting the covariance term), they find fluctuations in increases and decreases equally important to inflation volatility.<sup>13</sup>

While inflation was relatively low and stable in the United States during the sample period considered by Klenow and Kryvtsov (2008), Gagnon (2009) notes that Mexico experienced episodes of both high and low inflation from 1994 to 2002. He finds that when the annual rate of inflation was below 10-15%, the average frequency (size) of price changes co-moves weakly (strongly) with inflation due to offsetting movements in the frequency of price increases and decreases. When inflation rises beyond 10-15%, however, few price decreases are observed and both the frequency and average size are important determinants of inflation. Thus, over the entire sample period, the extensive margin accounts for more than half of the variance of inflation.

Wulfsberg (2009) studies Norwegian consumer price data over a 30-year period in which inflation was at first high and volatile (1975-1989) and then low and stable (1990-2004). He uses a different metric to assess the importance of the extensive and intensive margins. Specifically, he constructs the average monthly inflation in year  $t$  as the weighed product-sum of item-specific average frequencies and magnitudes of price changes,

$$\hat{\pi}_t = \sum_i \omega_{i,t} [fr_{i,t}^+ \cdot sz_{i,t}^+ + fr_{i,t}^- \cdot sz_{i,t}^-],$$

where  $fr_{i,t}^+$  is the average frequency of price increases,

and so on. To assess the importance of the extensive (intensive) margin, he computes the conditional inflation rate where the size (frequency) of price changes are kept constant at their means while allowing the frequency (size) of price changes to vary. He finds that the

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<sup>13</sup> Nakamura and Steinsson (2008a), on the other hand, find that the frequency of price increases, and not decreases, are important for driving inflation movements. They look at the median frequency of price changes across sectors rather than cross-sector means.

extensive margin is strongly correlated with CPI inflation (0.91), while the intensive margin is negatively correlated with CPI inflation (-0.12) and interprets this as evidence of strong state dependence in price setting. Restricting his sample to the low and stable inflation period of 1990-2004, he finds the intensive margin has a higher correlation with CPI inflation: 0.51 vs. 0.36.

We next consider evidence of price-change synchronization in the U.S. during the recent recession. Figure 5 plots the average frequency of price changes in the top three cities, based on regular prices for processed items. We first calculated the monthly frequency of price changes (the weighted mean across ELIs) from April 1988 through September 2009. We then took out seasonal (monthly) dummies, and calculated deviations from them. Finally, we averaged the monthly deviations in each quarter, and added back the mean across all quarters to produce quarterly data from 1988:2 through 2009:3. The *monthly* frequency of price changes increased a couple of percentage points from the end of 2007 onward, from about 18.5% to about 21%. We do not have a good metric for whether this is an economically large shift, but the issue deserves deeper investigation. For example, did the frequency increase more rapidly for cyclical goods? Did the increase reflect endogenously more attention given the magnitude of the recession? We do know that the increase holds for non-raw posted prices as well as all posted prices. Also, the size of increases and decreases for regular processed items were little changed (if we include raw items, there were some unusually large energy price declines near the end of the sample).

Another form of synchronization is seasonality. Nakamura and Steinsson (2008a) find that the weighted median frequency of (regular) consumer price changes declines monotonically over the four quarters of the year, with local spikes in the first month of each

quarter. The quarterly seasonal pattern in producer prices mirrors the seasonal pattern in consumer prices qualitatively, but is substantially larger, with the frequency of price change in January more than twice the average for the rest of the year. For the Euro area, Dhyne et al. (2006) emphasize that various goods are more likely to exhibit seasonal patterns: unprocessed foods due to seasonality in agricultural producer prices, certain industrial goods due to “end-of-season” sales, and services because they show an inclination to change prices at the beginning of the year and refrain from changing price at the end of the year.

The Euro area consumer price studies (Dhyne et al. 2006) also investigated synchronization at the product level using the Fisher–Konieczny (2000) measure, which takes a value of 1 in case of perfect synchronization and a value of 0 in the case of uniform staggering of price changes. They calculated the measure for each of the 50 product categories in their common sample. The degree of synchronization was, in general, rather low except for energy prices, with the median synchronization ratio across products ranging between 0.13 in Germany and 0.48 in Luxembourg. The higher ratio observed in Luxembourg compared to Germany likely reflects the difference in the size of the market upon which the ratio is computed and the relatively small number of outlets in Luxembourg.

Other studies have focused on the synchronization of prices at the outlet level. Using scanner data for a number of retailers in the United States, Nakamura (2008) decomposes price variation for individual products into variation that is common across all stores (16%), variation that is common only to stores within the same retail chain (65%), and variation that is completely idiosyncratic to particular stores (17%). These findings suggest that retail-level shocks, rather than manufacturer-level shocks, may be quite important for understanding fluctuations in retail prices. Midrigan (2009) finds the probability a particular product

experiences a price change depends on the fraction of other prices within its store that change, especially those in its own product category. He also finds some evidence, albeit weaker, of synchronization across stores in a particular city. Related, Lach and Tsiddon (1992, 1996) analyze prices of different meat products and wines at retail stores in Israel. They find that, when stores change price, they seem to change price for most of their products at the same time, but these adjustments occur at different times for different stores.

### *5.2 Sales, Reference Prices, and Aggregate Inflation*

A key question for macroeconomists when thinking about sale-related price changes is whether they respond to macro shocks or instead reflect entirely idiosyncratic forces. Midrigan (2009) and Nakamura and Steinsson (2008a) implicitly take the latter view when they replace sale prices with previous regular prices in their empirical analysis. Guimaraes and Sheedy (2008) rationalize the latter in a specific model, generating oscillation between sticky regular and sale prices as an optimal form of price discrimination. Similarly, Eichenbaum, Jaimovich and Rebelo (2009) describe a model in which sellers choose a “sticky pair” of prices that they can freely bounce between, but a menu cost applies whenever the pair is changed. Kehoe and Midrigan (2008) take an intermediate position, modeling sale-related price changes as subject to (smaller) menu costs and responding to macro as well as idiosyncratic shocks. Because sale prices tend to revert to previous regular prices (both in their theory and in the data), their theoretical sale prices contribute notably less to macro price flexibility than do regular price changes (which can be arbitrarily persistent).

The answer to whether sale-related price changes contribute to macro price flexibility is ultimately empirical, of course. According to Klenow and Kryvtsov (2008), more than 40% of sale price episodes do *not* return to the previous regular price, opening the door to

more macro price flexibility. Klenow and Willis (2007), using bi-monthly data for all cities in the CPI, find that the magnitude of price discounts indeed correlates with cumulative inflation since the item last changed price. Another possibility, which has not yet been investigated, is that clearance sales (more common for apparel and appliances, less common for food) react to unwanted inventory build-up at the macro level.

To provide some evidence on the macro content of sale prices, we calculated inflation for posted prices and regular prices separately in the U.S. CPI from February 1988 through October 2009 for the top three cities. We took out separate monthly dummies for each to remove seasonal effects. Table 12 summarizes some of the resulting moments for posted vs. regular price inflation in the U.S. CPI. For the residuals, the variance of regular price inflation equaled 87.5% of the variance of posted price inflation, so that sale prices "accounted for" 12.5% of the aggregate variance. By comparison, sale prices represent about 19% of all price changes (see Klenow and Kryvtsov, 2008). When we aggregated up to the quarterly level, however, sale prices contributed only 7.5% of the variance of quarterly posted price inflation. Thus sale-related price changes do not fully wash out with cross-sectional aggregation, but do significantly cancel out with time aggregation.

The serial correlation of posted price inflation in the U.S. CPI (0.394, standard error 0.053) is similar to that of regular price inflation (0.420, s.e. 0.051) – see also Bils, Klenow and Malin (2009). The same is true of quarterly inflation rates (0.174 serial correlation for posted price inflation, 0.142 for regular price inflation). To get at whether price changes build or fade, we regressed cumulative inflation from month  $t$  to  $t+12$  on inflation from month  $t$  to  $t+1$ . As given in Table 12, following a 1% price increase in the first month, regular prices are 1.48% (s.e. 0.27) higher after 12 months. The aggregate component of

sale-related price changes is surprisingly persistent: the difference between posted and regular prices is 1.09% (s.e. 0.17) higher after 12 months (following a 1% increase in the first month in the difference between posted and regular price inflation).

Table 13 provides statistics on aggregate *reference* price inflation, again obtained as residuals from separate monthly dummies for the top three cities in the U.S CPI. Recall that, in the spirit of Eichenbaum et al. (2009), we defined the reference price as the most common price in the 13-month window centered on the current month for each item. As discussed, reference prices represent about 80% of all prices in the CPI by this definition, and change every 11 months. By comparison, regular prices represent about 90% of all prices and change roughly every 7 months.<sup>14</sup> Thus, in practice, reference prices more aggressively filter out short-lived prices (though this need not be the case in principle, as our definition allows the reference price to conceivably change every month). The variance of aggregate *monthly* reference price inflation is 46% of the variance of posted price inflation in the U.S. CPI. The monthly posted price inflation rate is similarly correlated with reference price inflation (0.69) and with the deviations from reference price inflation (0.73). Monthly reference price inflation is more serially correlated (0.46, standard error 0.05), however, than are the deviations from reference price inflation (0.12, s.e. 0.06). This is not by construction, as the deviations from reference prices could have exhibited more persistent changes. Reference price changes build to 3.07% (s.e. 0.40) after one year, whereas deviations from reference prices neither build nor fade (coefficient 0.91%, s.e. 0.16).

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<sup>14</sup> The correlation between *monthly* regular and reference price inflation rates is 0.69 (standard error 0.03), whereas that between *quarterly* regular and reference price inflation rates is 0.79 (0.04).

As Table 13 shows, reference price inflation becomes more volatile and persistent relative to posted price inflation when we aggregate up to the quarterly level. The variance of reference price inflation rises to about 59% of overall quarterly inflation (from about 46% at the monthly frequency). At the quarterly level, posted price inflation is more correlated with reference price inflation (0.82) than with deviations from it (0.64). The serial correlation of quarterly reference price inflation is 0.44 (standard error 0.09), vs. -0.21 (0.11) for deviations from quarterly reference price inflation. Figure 6 plots these quarterly series. [Both rates plunge at the end of 2008 due to energy price declines, but these two quarters do not drive any of the statistics.] We conclude that reference price inflation picks up responses to more persistent shocks to inflation, whereas deviations from reference price inflation reflect more fleeting disturbances.

### *5.3 Hazard Rates*

Another “dynamic” feature that has been documented in many studies is the shape of the hazard function of price change, i.e., how the probability of changing price varies depending upon the age of the price. One reason for this is that price-setting models often have fairly stark predictions for the shape of the hazard function. The original Calvo (1983) model assumes a flat hazard function, while the deterministic timing of price adjustment in the Taylor (1980) model predicts a zero hazard except at a single age, where the hazard is one. Menu-cost models can generate a variety of shapes, depending on, among other things, the relative importance of transitory and permanent shocks to marginal costs. Permanent shocks, which accumulate over time, tend to yield an upward sloping hazard function, while transitory shocks tend to flatten or even produce a downward-sloping hazard (e.g., sellers



may be more attentive to getting prices right when revenue is temporarily high for a product due to idiosyncratic supply or demand considerations).

The general finding in the literature is that hazard rates for individual products are *not* upward-sloping.<sup>15</sup> Klenow and Kryvtsov (2008) find the frequency of (regular) price changes conditional on reaching a given age is downward sloping if all goods are considered together, but note that this could simply reflect a mix of heterogeneous flat hazards and survivor bias. Once they take out decile fixed effects (i.e., average frequencies of price change in each decile of price change frequency), they characterize the hazard rate as flat (other than a spike at 12 months). Álvarez (2008) reports a similar pattern in Euro area consumer price data. Nakamura and Steinsson (2008a) estimate separate hazard functions for each Major Group, however, and find hazards are downward sloping for the first few months and generally flat after that. They find similar patterns for producer prices. Finally, Cavallo (2009) finds initially downward-sloping hazards for daily online prices from a large supermarket in each of several Latin American countries.

#### 5.4 *Size vs. Age*

Klenow and Kryvtsov (2008) undertake an analogous exercise for the (absolute) size of price changes. They find that the size of price changes rises with age, but once they control for heterogeneity across goods, the size of price changes is unrelated to the age of the price for a given item. Likewise, Eden (2001), using data from Israel and several different approaches, finds no strong correlation between the size and age, and Álvarez and Hernando

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<sup>15</sup> One exception is Ikeda and Nishioka (2007), who use a finite mixture model (to allow for heterogeneity across price setters) and assume price changes occur according to a Weibull distribution (to allow for increasing, flat, or decreasing hazards). They estimate increasing hazard functions for some Japanese products, and flat or Taylor-type spiked hazards for others.

(2004), using Spanish data and Heckman's sample selection correction, report that age does not significantly contribute to the size of price changes.

In contrast, many models predict a positive relationship between size and age. For example, models in which shocks accumulate over time and the timing of price changes are exogenous (Calvo 1983) or driven primarily by having a low menu-cost draw (Dotsey et al. 1999) fall into this category.

### *5.5 Transitory Relative Price Changes*

The persistence of relative price changes can shed light on the relative importance of idiosyncratic versus aggregate shocks and also help to distinguish between different price-setting models. Indirect evidence on this aspect of the data comes from investigating the behavior of sales. Nakamura and Steinsson (2008a) report that (a) sales spells are short – the average length is just 1.8–2.3 months; (b) sales price changes are twice as large as other price changes on average; and (c) many prices return to their original price following a sale. Taking these facts together suggests highly transient *relative* price changes for both sales prices and all posted prices, as sales are quite common in the U.S. CPI data (representing about 20% of price changes according to Klenow and Kryvtsov, 2008). Scanner data studies uncover similar patterns. For example, Kehoe and Midrigan (2008) find that price changes are large and dispersed and most are of a “temporary” nature, again suggesting transitory relative price changes.

The transitory nature of relative price changes is not just a sales-related phenomenon, however, as it has been found after filtering out sales as well. Campbell and Eden (2005) document that grocers choose, and then quickly abandon, extreme relative prices (rather than

arriving at those prices through the gradual erosion of their fixed nominal price by other sellers' price adjustments). Midrigan (2009) reports that the probability a firm's next price change will have the same sign as the current one is fairly low (between 32% and 41%). Klenow and Willis (2006) construct relative prices as the ratio of individual prices to the weighted average of prices in a sector, and report an across-sector weighted mean serial correlation of *new* relative prices (i.e., across months with price changes) of 0.32.

To provide a metric for how transitory these relative price changes are, it is useful to consider what they imply for the persistence of shocks in structural models. The general finding is that transitory, idiosyncratic shocks are an important input that enables structural models to replicate pricing patterns in the data. Klenow and Willis (2006) estimate a monthly persistence of 0.7 for the idiosyncratic technology shock in their model in order to match the serial correlation of *new* relative prices. The shocks that trigger price changes in Midrigan (2009) are similarly transitory and tend to be reversed. Nakamura and Steinsson (2008a) calibrate the size of the menu cost and the idiosyncratic-shock persistence and volatility to match the frequency of (regular) price changes, the fraction of changes that are increases, and the size of changes. The idiosyncratic shock is large and has a serial correlation of 0.66. Other models that produce transitory relative price changes include Burstein and Hellwig (2007) who incorporate transitory demand shocks in addition to marginal cost shocks, Kehoe and Midrigan (2008) who give firms the option of a "one-period markdown" at a lower menu cost than a permanent price change, and Eichenbaum et al. (2009) whose sticky-plan firms can switch between a small number of prices at no cost.

## 5.6 Response to Shocks

Evidence on the response of prices to identified shocks is particularly useful for shedding light on the nature of firms' price-setting decisions because most (if not all) theories of price-setting posit a tight link between (desired) prices and the underlying cost and demand conditions facing the firm. Of course, precisely identifying shocks in the data is not an easy task, but recent studies have made use of novel datasets to do just this. These include studies of the response of import/export prices or of prices on either side of a border to exchange rate shocks, evidence of price-setting around the time of a significant event like the Euro changeover or an increase in VAT rates, and studies of scanner data which contain both wholesale costs and retail prices.

Researchers have made use of well-identified and sizable movements in exchange rates to evaluate how prices respond to shocks. Gopinath and Rigobon (2008) measure *medium-run* exchange rate pass-through as the change in a price in response to the cumulative change in the exchange rate since the price was last changed and find that the trade weighted exchange rate pass-through into U.S. import prices is low, at 22%. Gopinath, Itskhoki and Rigobon (forthcoming) document that non-dollar priced goods display much higher medium-run pass-through than do dollar priced goods. They present a sticky-price model in which firms choose their currency based on their average desired pass-through over the period of price non-adjustment. Incomplete desired pass-through – whether it is driven by variable mark-ups, imported inputs or decreasing returns to scale in production – is thus seen as the reason that the majority of U.S. import goods are priced in dollars. Gopinath and Itskhoki (forthcoming) find further evidence consistent with variable mark-ups or variable marginal costs in the form of a positive correlation between the frequency of price

adjustment and *long-run* pass-through, the life-long change in the price of a good (relative to U.S. inflation) on the (real) exchange rate over the same period. Fitzgerald and Haller (2009) use micro data on domestic and export prices set by Irish producers. Because they have matched price quotes for the same product produced in the same plant but sold in multiple markets, they are able to control for changes in marginal costs in addition to exchange rate movements. They find strong evidence of pricing-to-market for products whose export prices are invoiced in the destination currency: conditioning on price changes in both markets, relative prices move one-for-one with exchange rate changes (i.e., zero pass-through of exchange rate changes).

Other research has found evidence that borders effectively segment markets. Burstein and Jaimovich (2009) and Gopinath et al. (2009) evaluate scanner data from a major retailer that has multiple locations in Canada and the US. Gopinath et al. (2009) find large discontinuities of retail prices and wholesale costs between adjacent stores on either side of the border, and note that these border gaps move almost one-for-one with changes in the nominal exchange rate. Burstein and Jaimovich (2009) document substantial pricing-to-market for traded goods. Using a dataset of online book prices (and some information on quantities) for a number of U.S. and Canadian retailers, Boivin, Clark, and Vincent (2009) conclude that market segmentation is probably behind the lack of exchange rate pass-through in their sample.

Other identifiable aggregate shocks include changes in tax rates or in the nation's currency. A number of European studies, summarized in Dhyne et al. (2005), have consistently found that changes in the value-added tax rate lead to temporary increases in the frequency of CPI price changes. The frequency of price change also increased noticeably

after countries converted to the euro in January 2002 (Dhyne et al. 2005). Although the response of aggregate inflation to the introduction of the euro was not particularly large, Hobijn et al. (2006) emphasize the dramatic increase in euro area restaurant prices and explain this increase with a menu-cost model in which firms decide when to adopt the new currency.

Eichenbaum et al. (2009) are able to provide information on the price response to (potentially disaggregated) shocks. They investigate the relationship of prices to costs using scanner data for one large US retailer. Prices typically do not change in the absence of a change in cost, but a cost change is not sufficient to induce a change in price. This allows markups (both actual and reference) to display substantial variation. When the retailer does decide to change its reference price, it almost always re-establishes the unconditional markup over the reference cost. That is, the retailer passes through 100% of the cumulative change in reference cost that occurred since the last reference price change.

Nakamura and Zerom (forthcoming) document delayed and incomplete pass-through of (observable) commodity cost shocks to wholesale/retail prices in the coffee industry. They estimate that, relative to a CES benchmark, local costs reduce long-run pass-through (after 6 quarters) by 59% and markup adjustment reduces it by an additional 33%. Barriers to price adjustment are important for the delayed response of prices to cost but have a negligible role in incomplete long-run pass-through.

Bils, Klenow, and Malin (2009) do not condition upon identified shocks but instead document how “reset prices” evolve over time in response to reduced-form impulses. A reset price for an individual firm is the price it would choose if it implemented a price change in the current period. Models with sizeable strategic complementarities predict that the impact of a

shock to reset prices will build over time, as price setters wait for the average price to respond. In the U.S. CPI data, however, the reset price actually overshoots – it changes more on impact than it does in the long-run. At the same time, Klenow and Willis (2007) find some evidence that micro price changes respond to “old” inflation (i.e., since before the item’s previous price change), as in sticky information models.

### *5.7 Higher Moments of Price Changes and Aggregate Inflation*

Various authors have investigated the relationship between higher moments of the distribution of price changes and inflation (i.e., the first moment). Relative price variability is generally found to be positively correlated with inflation, as predicted by menu cost or incomplete information models. Lach and Tsiddon (1992) point out that menu cost models imply that the variability is affected by *expected* inflation, while incomplete information models imply a relationship with *unexpected* inflation. They find that expected inflation had a stronger effect on variability than unexpected inflation for a sample of foodstuffs in Israel. Koneiczny and Skrzypacz (2005) reach similar conclusions using data on 52 goods in Poland from 1990 to 1996.

Ball and Mankiw (1995) focus on the third moment and find that innovations in inflation are positively associated with the skewness in relative-price changes for four-digit U.S. PPI data from 1949 to 1989. Indeed, the inflation-skewness relationship is stronger than the inflation-variance relationship. They show these findings are consistent with a menu-cost model: when price adjustment is costly, firms adjust to large shocks but not to small shocks, and therefore large shocks have disproportionate effects on the price level. They also explore the idea that asymmetric relative-price changes represent aggregate supply shocks.

Domberger and Fiebig (1993) look at individual price changes in 80 disaggregated industry groups in the UK. They also find that an increasing (decreasing) average price is associated with a rightward (leftward) skew, but note that when the absolute magnitude of inflation (deflation) increases, the distribution of individual price changes becomes less skewed. They suggest that the higher the skewness of the price-change distribution, the more staggered (less synchronized) are changes.

We complement these studies by calculating monthly measures of higher moments of the distribution of price changes for the top three cities in the U.S. CPI from 1988 to 2009. We then construct the correlation of the variance, skewness, and kurtosis of (non-zero) price changes with inflation. We find few robust significant relationships. The variance is *negatively* correlated with inflation, both for posted prices (-0.16, s.e. 0.06) and for regular prices (-0.29, s.e. 0.06). These results are fragile, however, being highly influenced by deflation during the last few months of 2008. Omitting the last 15 months of the sample (August 2008 – October 2009), the inflation-variance correlation is still statistically significant for regular prices (-0.15, s.e. 0.06) but not for posted prices (-0.07, s.e. 0.06). Moreover, the correlation between skewness and inflation becomes significant: correlation of 0.15 (s.e. 0.06) for regular prices and 0.17 (s.e. 0.06) for posted prices. That is, the late-2008 plunge in inflation was accompanied by a spike in the variance of price changes, but no fall in skewness. We find that skewness is more robustly (if modestly) related to *absolute* inflation (0.15, s.e. 0.06 for both posted prices and regular prices).

More study of the relationship between inflation and higher moments thus seems warranted. For example, Ball and Mankiw's interpretation of skewness as representing an aggregate supply shock suggests a more nuanced inflation-skewness relationship than might



be picked up by a simple correlation. Indeed, they argue that the absence of asymmetry when inflation fell in the 1975 and 1982 recessions can be viewed as evidence in favor of their hypothesis, as it suggests that causation does not run from inflation to skewness.

## **6. Ten Facts and Implications for Macro Models**

We now summarize the findings from the empirical literature on individual price data with ten stylized facts and their potential lessons for models. These touch on two broad themes. The first is the frequency and nature (e.g., state-dependent vs. time-dependent pricing) of micro price changes. The second is the potential magnitude and source of the “contract multiplier”, i.e., how long it takes for the macro effects of price stickiness to fade after a permanent shock relative to the duration of individual prices.

*Fact 1: Prices change at least once a year.*

In the U.S. CPI, prices change every 4 months or so. Prices appear stickier in the U.S. PPI (around 6-8 months for the median), and stickier still in the Euro Area (around once a year for the median). But median price durations in emerging markets (tilted toward food, as actual household budgets are) are typically closer to the U.S. than the Euro Area.

In contrast, a large literature has estimated that permanent monetary shocks have real effects lasting several years.<sup>16</sup> Real effects of nominal shocks therefore last three to five times longer than individual prices. Nominal stickiness appears insufficient to explain why aggregate prices respond so sluggishly to monetary policy shocks. For this reason, nominal price stickiness is usually combined with a “contract multiplier” (in Taylor’s 1980 phrase).

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<sup>16</sup> Christiano, Eichenbaum, and Evans (1999), Romer and Romer (2004), and Bernanke, Boivin, and Elias (2005) are a few examples based on U.S. data. Peersman and Smets (2003) report similar Euro Area results.

Ball and Romer (1990) and Kimball (1995) are early examples, and Christiano, Eichenbaum and Evans (2005) a more recent one. Possible sources of big contract multipliers include strategic complementarities, countercyclical markups, sticky plans, and sticky information (the latter including rational inattention, for example Mackowiak and Wiederholt, 2008).

Of course, prices do not change continuously either. Nominal stickiness is a pervasive feature of the data in all country-years with moderate inflation rates. Models that generate persistence in aggregate prices without any nominal stickiness, for example via indexation (Christiano et al., 2005) or convex costs of price adjustment (Rotemberg 1982), are at odds with this fact.

*Fact 2: Sales and product turnover are often important for micro price flexibility.*

Large price discounts and product turnover account for about one-third of consumer price changes in the United States. Sales are much more frequent for some types of goods (food and apparel), and usually revert to the previous regular price afterward (more so for food, less so for apparel). Sales are much less frequent in the Euro Area than in the U.S. Within the U.S., sales are less prevalent for producer prices than consumer prices. Product turnover ushers in nominal price changes, and is much more common for durables (and apparel) than food or services. Both sales and product turnover exhibit clear seasonal patterns, suggesting a time-dependent component.

Given their temporary and seasonal nature, sale-related price changes may contribute more to micro than macro price flexibility. Put differently, the contract multiplier may be larger for price changes related to sales and product turnover. That said, sale- and turnover-

related price changes do not wash out with aggregation in the U.S., consistent with macro content. It remains an open question whether such prices respond to aggregate shocks.

In addition to interesting empirical questions concerning the macroeconomic importance of sales and substitutions, more theoretical research is necessary for us to better understand how these types of price changes affect the monetary transmission mechanism. It is awkward to simply apply menu cost models to data in which some price changes have been filtered out. Alternative approaches have been pursued in the literature. Kehoe and Midrigan (2008) model one-period price discounts as incurring a lower menu cost. They find sales-related price changes can contribute to macro price flexibility, but not as much as regular price changes. Guimaraes and Sheedy (2008) build a model in which firms face consumers with different price sensitivities and therefore choose a pricing strategy with a high price at some moments and a low price at other moments. Even though “sales” decisions are completely flexible at the micro level, this does not translate into flexibility at the macro level because sales are strategic substitutes: a firm gains little from a sale if its competitors are having sales. Thus, different models of sales have different implications for aggregate flexibility. Empirical work could help distinguish between and refine this theorizing.

*Fact 3: Reference prices are stickier and more persistent than regular prices.*

Eichenbaum, Jaimovich and Rebelo (2009) document many temporary regular price changes in a major U.S. grocery store chain. We find a similar pattern of sticky “reference” prices in most U.S. CPI items. Whereas posted prices change every 4 months and regular prices every 7 months, reference prices change every 10-11 months.

Deviations from reference prices do not cancel out across items, thereby affecting aggregate inflation. Perhaps not surprising, reference price inflation is considerably more persistent than either posted or regular price inflation. Changes in reference prices build over time, whereas other price changes tend to be much more transitory. Reference pricing behavior may suggest some form of sticky plan and/or sticky information, rather than menu costs per price change, and hence could explain the sizable contract multiplier. That said, a novel price (relative to prices in the previous 12 months) tends to appear every six months or so for a given item. And the vast majority of prices are not “comeback” prices in the sense of having existed in an earlier price spell in the prior 12 months.

A related fact is that U.S. CPI quote-lines, which typically span over four years, are dominated by a few prices. The top price alone captures one-third of prices, and the top two prices more than one-half of prices. This fact may reflect downward-sloping hazards rather than sticky information or plans, though, given that two-thirds of top prices occur without interruption by other prices.

Future research could explore whether the reference price phenomenon extends to U.S. producer prices and to prices in other countries. There might be more frequent reference price changes in countries and timeframes with more persistent inflation than in recent U.S. decades. Moreover, a vital question remains: how do reference vs. non-reference prices respond to aggregate shocks?

*Fact 4: There is substantial heterogeneity in the frequency of price change across goods.*

Every micro dataset to date has displayed large, persistent heterogeneity in the frequency of price changes across types of goods. Universally, service prices are stickier

than those of goods. Among goods, “raw” goods (energy, fresh produce) are more flexible than “processed” goods (goods with low intermediate shares and high labor shares).

As stressed by Carvalho (2006) and Nakamura and Steinsson (2008b), such heterogeneity can combine powerfully with strategic complementarities to boost the contract multiplier. But evidence for strategic complementarities in micro price data is mixed. In the supportive column, Gopinath and collaborators have consistently found that exchange rate changes gradually pass-through into import prices. In the skeptical column, Klenow and Willis (2006) stress that sellers routinely change prices relative to close substitutes (e.g., temporary sales are not synchronized among competing brands of a narrow product) and argue that large relative price movements are inconsistent with strong “micro” real rigidities. Using evidence on product-level prices and market shares, Burstein and Hellwig (2007) also conclude that “micro” pricing complementarities are too weak to generate large aggregate real effects. Kryvtsov and Midrigan (2009) note that strong “macro” complementarities imply stable markups, and hence are hard to reconcile with countercyclical inventory-sales ratios. And Bils, Klenow and Malin (2009) find that consumer price changers tend to overshoot, rather than undershoot as predicted by slow pass-through.

*Fact 5: More cyclical goods change prices more frequently.*

In the U.S. CPI, prices change more frequently in categories with more procyclical real consumption growth – in particular transportation (cars, airfares) and apparel. This fact probably extends to U.S. producer prices, given the high correlation between upstream and downstream price change frequency (Nakamura and Steinsson, 2008a). In other countries, we also note the strong tendency for durables to change prices more frequently than services.

This form of price change heterogeneity could have the opposite effect of idiosyncratic sources of heterogeneity (such as different-sized menu costs or idiosyncratic shocks); that is, higher frequency price changes among cyclical goods could reduce the contract multiplier. See, in particular, Barsky, House and Kimball (2007). Natural topics for further investigation include the source of this connection (does cyclicality cause flexibility or the other way around?) and its quantitative importance for the contract multiplier in various models.

*Fact 6: Price changes are big on average, but many small changes occur.*

Micro price changes are an order of magnitude larger than needed to keep up with aggregate inflation in the U.S. and Euro Area. Thus idiosyncratic forces dominate macro shocks. Golosov and Lucas (2007) show that the former can limit the contract multiplier. In their SDP model, a positive aggregate monetary shock results in more idiosyncratic price increases (and fewer idiosyncratic price decreases), thereby speeding the response of the aggregate price level. Bilal, Klenow and Malin (2009) find that “reset price inflation” behaves as if it strongly affected by such selection.

Many price changes *are* small, too, so that there is a “missing middle” in the distribution of price changes emerging from SDP models with a single, large menu cost. This need not be true if menu costs are variable (as in Dotsey, King and Wolman, 1998) or are small but shocks arrive infrequently (as in Gertler and Leahy, 2008).

Meanwhile, Midrigan (2009) makes the point that TDP models can easily fit the whole size distribution. He argues that the missing middle in SDP models is a byproduct of overstated selection. Alternatively, the small price changes generated by his SDP model with multiproduct firms reflect weakened selection and hence greater departures from monetary

neutrality, akin to TDP models. Woodford (2009) lays out a model in which information constraints can lead to small price changes as well (costly information updating results in little change *ex post* sometimes, though not on average). As in TDP, the result of information constraints can be a bigger contract multiplier than under SDP.

*Fact 7: Relative price changes are transitory.*

A corollary to large idiosyncratic price changes is big relative price movements. Even within narrower categories of the CPI (Expenditure Classes, ELIs) and U.S. grocery scanner data, these relative price changes are large. Such relative price movements tend to fade over time – they are far less persistent than a random walk. This is true across stores of different chains, but also across competing brands within chains. It is also true even for regular prices, i.e., excluding the (often temporary) sale price discounts from regular prices. This fact is a corollary to many regular price changes being temporary in nominal terms.

The persistence of relative prices can matter for macroeconomic questions. First, in menu cost models the Ss band for non-adjustment is narrower when firms are selling more and wider when firms are selling less. Thus the selection effect of macro shocks can change with the persistence of idiosyncratic shocks. Second, big transitory movements in relative prices require either that micro complementarities be weak, or that idiosyncratic shocks be large. Weaker complementarities imply a smaller contract multiplier in response to monetary shocks. Third, transitory movements may reflect micro flexibility amidst a sticky macro plan/information (i.e., firms frequently respond to big, temporary idiosyncratic shocks within the plan, but less frequently change the macro plan/information given that persistent macro shocks are smaller).

*Fact 8: Price changes are typically not synchronized over the business cycle.*

The contract multiplier can be lower if sellers accelerate price increases in response to positive monetary shocks and postpone them in response to negative monetary shocks. For moderate inflation episodes such as the last two decades in the U.S., sellers do not seem to synchronize their timing in this way. This is consistent with time-dependent pricing, but need not nullify the selection effect of state-dependent pricing completely. As predicted by state-dependent models and time-dependent models alike, the composition of price increases vs. decreases correlates positively with inflation movements. But lack of synchronization may mean weaker selection, or even a preoccupation with idiosyncratic over aggregate shocks (e.g., rational inattention).

*Staggered* nominal price stickiness is also critical for how strategic complementarities can help produce a big contract multiplier. Synchronized price changes are not subject to coordination failure in magnitude. The less price changes are synchronized, the more sellers may slow their adjustment as they wait for other prices (of competitors and input suppliers) to adjust. That said, the evidence for strategic complementarities appears mixed. The lack of synchronization even among closely competing brands, itself, is hard to understand if micro real rigidities are strong.

Periods of greater macro volatility may exhibit more synchronization. In Mexico in recent decades, the frequency of price changes moved up and down with aggregate inflation movements. In the recent U.S. recession, the frequency of consumer price changes appears to have surged. Intriguingly, both price increases and decreases became more common, perhaps because of more frequent updating of sticky plans or sticky information.



*Fact 9: Neither frequency nor size is increasing in the age of a price.*

In both the U.S. and Euro Area, and for both consumer and producer prices, the hazard rate of price changes is falling over the first few months and largely flat afterward. The exception is a spike in price change frequency for services, suggestive of annual updating for the stickiest category of products. The downward slope is much less pronounced if one looks at regular prices (i.e., excludes sale-related price changes), and may be flat if one fully controls for survivor bias.

For a state-dependent pricing model, standard intuition would have the hazard of price changes increasing in the time since a price change as shocks accumulate and the desired price drifts farther away from the current price. This force can be weakened by thicker-tailed or Poisson shocks, and by periods with wider vs. narrower  $S_s$  bands (which create an intertemporal form of survivor bias). Scanner studies find clear evidence of state-dependence in rising hazards in distance from the average markup. Again, state-dependence can imply a smaller contract multiplier than time-dependent pricing.

For a time-dependent model, analogous intuition would imply that the *size* of price changes is *increasing* in the duration of stickiness, as more shocks accumulate the longer the spell between price changes. Evidence on this question is somewhat more limited, but there seems to be little connection between the size of price changes and the duration of price spells. This pattern is consistent with state-dependent pricing, under which spell length is endogenous to the shocks accumulated. A long spell without a price change may signal that the desired price did not move much, so that it need not change unduly when the spell ends. This question could use more study in the U.S. and elsewhere.

*Fact 10: Price changes are linked to wage changes.*

Recent research has revealed a noticeable link between price and wage rigidity. In the cross-section, firms (or categories of goods) with a higher share of labor costs in total costs make less frequent price adjustments, potentially resulting from the fact that wages adjust less frequently than other input prices. Survey evidence also suggests synchronization between wage and price adjustments over time, as well as a cross-sectional correlation between wage flexibility and price flexibility.

Wage stickiness can contribute directly to a high contract multiplier, of course. But the aforementioned evidence suggests it may be contributing indirectly, as well, by lowering the frequency of price changes. This relationship is all the stronger where production is labor-intensive. Moreover, wage adjustment exhibits a substantial degree of time-dependence. Firms tend to concentrate wage changes in a specific month, mostly January in a majority of European countries (Druant et al. 2009). This could lend a degree of time-dependence into price-setting, further contributing to a higher contract multiplier.

*Summary: Model features and the facts.*

Table 14 provides a quick summary of how some common features of macro models of price-setting stack up against the stylized facts about micro price setting. Most of the entries should be self-explanatory given the preceding discussion, but a few warrant elaboration. Take the fact that many consumer price changes are reversed (i.e., prices sometimes exhibit “memory”, particularly after sales in food and apparel in the U.S.). This pattern is consistent with sellers periodically choosing a “sticky set” of a few prices which

they bounce between, a form of sticky plan advocated by Eichenbaum et al. (2009). A fuller explanation might be price discrimination *a la* Guimaraes and Sheedy (2008).

Probably the most robust fact across all micro pricing studies is that goods persistently differ in their frequency of price changes. The variation is far from random, as “raw” goods (fresh produce, energy) change price more frequently than do services in country after country. Thus price flexibility appears to respond to economic fundamentals (e.g. the average size of sectoral shocks and the trend rate of inflation). And, in the U.S. at least, more cyclical categories exhibit more consumer price flexibility.

Another consistent finding is that most micro price changes are much larger than needed to keep up with average inflation. This fact would seemingly point to a combination of big idiosyncratic shocks and big menu costs. But rational inattention could co-exist with these ingredients. At the same time, the number of small price changes in the data is far from trivial. This may be because smaller menu costs are combined with sticky information (sellers sometimes find only a small price change is needed after a periodic information update) or with sticky plans (sellers periodically update a path of prices, as in Burstein, 2006).

## **7. Conclusion**

We review the recent empirical literature on individual price data and distill ten salient facts for macro models. Prices change quite frequently, although much of this flexibility is associated with price movements that are temporary in nature. Even if all short-lived prices are excluded, however, the resulting nominal stickiness, by itself, appears insufficient to account for the sluggish movement of aggregate prices. These findings point to the need for a large contract multiplier to bridge the gap between micro flexibility and macro inertia.

Other micro price facts provide evidence on the plausibility of various mechanisms for generating a large contract multiplier. The lack of synchronization in the timing of price changes provides scope for strategic complementarities to amplify the real effects of nominal stickiness. The presence of substantial heterogeneity across goods in the average frequency of price change can bolster this channel, although the fact that more cyclical goods exhibit greater price flexibility may work in the opposite direction. The presence of many large, transitory price changes raises concerns about the relevance of real rigidities at the individual-item level, but is consistent with macro rigidities or with rationally inattentive sellers who respond to sizeable idiosyncratic shocks but not to smaller aggregate impulses. Finally, the fact that the size of price changes does not increase with the age of a price provides evidence for state-dependent pricing and thus selection. The stronger the selection effect, the smaller the contract multiplier, although the presence of many small price changes suggests the selection effect may be muted.

A number of open empirical (and related theoretical) questions remain. How do temporary price changes (related to sales, product turnover, or a movement to a non-reference price) respond to aggregate shocks? How should these temporary movements be modeled, and what impact do they have on the contract multiplier? What drives the relationship between the cyclicalities of goods and the frequency with which prices change? Related, what aspects of micro heterogeneity in price-setting are important to consider in macro models? Finally, what evidence can be used to distinguish between different sources of the contract multiplier, such as rational inattention, sticky plans, or strategic complementarities?

**Table 1****Monthly Mean Frequency of CPI Price Changes**

Country	Paper	Sample Period	Frequency ( in %)
Austria	Baumgartner <i>et al.</i> (2005)	1996:01 – 2003:12	15.1
Belgium	Aucremanne and Dhyne (2004)	1989:01 – 2001:01	16.9
Brazil	Barros <i>et al.</i> (2009)	1996:03 – 2008:12	37.2
	Gouvea (2007)	1996:01 – 2006:12	37.0
Chile	Medina <i>et al.</i> (2007)	1999:01 – 2005:07	46.1
Denmark	Hansen and Hansen (2006)	1997:01 – 2005:12	17.3
Euro Area	Dhyne <i>et al.</i> (2006)	1996:01 – 2001:01	15.1
Finland	Vilmunen and Laakkonen (2005)	1997:01 – 2003:12	16.5
France	Baudry <i>et al.</i> (2007)	1994:07 – 2003:02	18.9
Germany	Hoffmann and Kurz-Kim (2006)	1998:02 – 2004:01	11.3
Hungary	Gabriel and Reiff (2008)	2001:12 – 2007:06	15.1
Israel	Baharad and Eden (2004)	1991:01 – 1992:12	24.5
Italy	Fabiani <i>et al.</i> (2006)	1996:01 – 2003:12	10.0
Japan	Saita <i>et al.</i> (2006)	1999:01 – 2003:12	23.1
Luxembourg	Lunnemann and Matha (2005)	1999:01 – 2004:12	17.0
Mexico	Gagnon (2009)	1994:01 – 2004:12	29.4
Netherlands	Jonker <i>et al.</i> (2004)	1998:11 – 2003:04	16.5
Norway	Wulfsberg (2009)	1975:01 – 2004:12	21.3 (21.9)
Portugal	Dias <i>et al.</i> (2004)	1992:01 – 2001:01	22.2
Sierra Leone	Kovanen (2006)	1999:01 – 2003:04	51.5
Slovakia	Coricelli and Horvath (2006)	1997:01 – 2001:12	34.0
South Africa	Creamer and Rankin (2008)	2001:12 – 2006:02	16.0
Spain	Álvarez and Hernando (2006)	1993:01 – 2001:12	15.0
United Kingdom	Bunn and Ellis (2009)	1996:01 – 2006:01	15 (19)
United States	Bils and Klenow (2004)	1995:01 – 1997:12	26.1
	Klenow and Kryvtsov (2008)	1988:02 – 2005:01	29.9 (36.2)
	Nakamura and Steinsson (2008a)	1988:01 – 2005:12	21.1 (26.5)

Notes: Source: Álvarez (2008) with three additional studies, Barros *et al.* (2009), Bunn and Ellis (2009) and Wulfsberg (2009), and updated versions of Gagnon (2009), Creamer and Rankin (2008), and Klenow and Kryvtsov (2008). For studies that report frequencies of both regular (i.e., non-sale) and posted prices, the figures in parentheses correspond to posted prices. Frequencies for Nakamura and Steinsson (2008a) correspond to the 1998-2005 sample period (for contiguous observations, excluding substitutions). For Germany, frequencies refer to the sample considering item replacements and non-quality adjusted data. The Spanish sample excludes energy products, which lowers the aggregate frequency.

**Table 2****Monthly Mean Frequency of PPI Price Changes**

<b>Country</b>	<b>Paper</b>	<b>Sample Period</b>	<b>Frequency (in %)</b>
Belgium	Cornille and Dossche (2008)	2001:01 – 2005:01	24
Colombia	Julio and Zarate (2008)	1999:06 – 2006:10	20.2
Euro Area	Vermuelen <i>et al.</i> (2007)	Various	21
France	Gautier (2008)	1994:01 – 2005:06	25
Germany	Stahl (2006)	1997:01 – 2003:09	22
Italy	Sabbatini <i>et al.</i> (2006)	1997:01 – 2002:12	15
Portugal	Dias <i>et al.</i> (2004)	1995:01 – 2001:01	23
South Africa	Creamer (2008)	2001:12 – 2006:02	19.5
Spain	Álvarez <i>et al.</i> (2008)	1991:01 – 1999:02	21
United Kingdom	Bunn and Ellis (2009)	2003:01 – 2007:12	26
United States	Nakamura-Steinsson (2008a)	1988:01 – 2005:12	24.7
	Goldberg-Hellerstein (2009)	1987:01 – 2008:10	33.3

Note: Source: Álvarez (2008), Bunn and Ellis (2009), Goldberg and Hellerstein (2009) and the published versions of Cornille and Dossche (2008) and Gautier (2008). Frequencies for Nakamura and Steinsson (2008a) correspond to finished goods. The Italian sample excludes energy products, while the French sample does not include business services.

**Table 3****Frequency of Price Change in Scanner Data Sets**

<b>Data Source</b>	<b>Paper</b>	<b>Sample Period</b>	<b>Frequency (in %)</b>
Dominick's	Midrigan (2009)	1989 – 1997	45 (23)
	Burstein and Hellwig (2007)		41 (26)
Large U.S. Retailer	Eichenbaum et al. (2009)	2004 – 2006	43
AC Nielsen ERIM Database	Campbell and Eden (2005)	1985 – 1987	47
	Midrigan (2009)		36 (25)
AC Nielsen ScanTrak	Nakamura (2008)	2004	44 (19)
AC Nielsen Homescan	Broda and Weinstein (2007)	1994 – 2003	NA

Note: For most studies, weekly data is collected, but a monthly frequency of price change is reported. Eichenbaum et al. (2009) report a weekly frequency of price change. Frequencies are for *posted* prices, and numbers in parentheses are for regular (i.e., non-sale) prices. Frequencies vary across studies using the same data set because of different sample choices and reported measures (e.g., for Dominick's, Midrigan (2009) reports mean frequencies for one store, while Burstein and Hellwig (2007) consider many stores and report the frequency of the median product category).

**Table 4****Number of Price Changes per year (%) in Survey Data**

Country	Paper	<1	1	2-3	≥4	Median	Mean (in months)
Austria	Kwapil <i>et al.</i> (2005)	24	51	15	11	1	12.7
Belgium	Aucremanne and Druant (2005)	18	55	18	8	1	11.9
Canada	Amirault <i>et al.</i> (2006)	8	27	23	44	2-3	6.8
Estonia	Dabusinskas and Randveer (2006)	14	43	25	18	1	10.0
Euro Area	Fabiani <i>et al.</i> (2005)	27	39	20	14	1	12.3
France	Loupias and Ricart (2004)	21	46	24	9	1	11.8
Germany	Stahl (2005)	44	14	21	21	1	13.5
Italy	Fabiani <i>et al.</i> (2007)	20	50	19	11	1	11.9
Japan	Nakagawa <i>et al.</i> (2000)	23	52	11	14	1	12.5
Luxembourg	Lunnemann and Matha (2006)	15	31	27	27	2-3	9.0
Mexico	Castanon <i>et al.</i> (2008)	-	-	-	-	-	5.7
Netherlands	Hoeberichts and Stokman (2006)	10	60	19	11	1	10.7
Portugal	Martins (2005)	24	51	14	12	1	12.7
Romania	Copaciu <i>et al.</i> (2007)	-	-	-	-	-	4.1
Spain	Álvarez and Hernando (2007a)	14	57	15	14	1	11.1
Sweden	Apel, Friberg and Hallsten (2005)	29	43	6	20	1	12.7
Turkey	Sahinoz and Saracoglu (2008)	-	-	-	-	-	3.0
United Kingdom	Hall, Walsh and Yates (2000)	6	37	44	14	2-3	8.2
United States	Blinder <i>et al.</i> (1998)	10	39	29	22	1	8.8

Note: Source: Álvarez (2008), Table 3. Mean implicit durations obtained from the interval-grouped data using the following assumptions: for firms declaring “at least four price changes per year”, 8 price changes are considered (i.e. mean duration of 1.33 months); for those declaring “two or three price changes per year”, 2.5 price changes are considered (i.e., 4.8 months); for those declaring “one change per year” a duration of 12 months, and for those declaring “less than one price change per year”, a change every two years is considered (i.e., 24 months).



**Table 5****Price Durations by Category in the U.S. CPI**

Durations in Months	<i>Posted</i>		<i>Regular</i>		% of CPI
	Median	Mean	Median	Mean	
<i>All Items</i>	3.4	6.2	6.9	8.0	100.0%
Durable Goods	1.8	3.0	1.8	5.0	21.7
Nondurable Goods	3.4	5.8	7.3	8.3	48.6
Services	7.6	9.4	7.6	9.6	29.7
Raw Goods	1.0	1.1	1.0	1.2	12.0
Processed Goods	4.4	6.9	7.7	8.9	88.0
Apparel	2.8	2.9	9.2	10.1	7.0
Education and Communication	5.4	6.2	6.7	6.3	7.3
Food	3.4	6.9	8.5	9.3	22.4
Home Furnishings	1.9	3.5	2.0	5.4	17.0
Medical Care	10.0	14.2	12.6	14.7	7.8
Recreation	6.3	7.5	9.4	9.8	8.5
Transportation	1.8	3.7	1.8	3.8	24.5
Other Goods and Services	8.6	14.7	12.1	16.7	5.5

*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) from February 1988 through October 2009. Durations are weighted medians or means of implied durations from weighted average frequencies within ELIs. Durables, Nondurables, and Services coincide with U.S. National Income and Product Account classifications. Raw goods include gasoline, motor oil and coolants, fuel oil and other fuels, electricity, natural gas, meats, fish, eggs, fresh fruits, fresh vegetables, and fresh milk and cream. Apparel, etc. are Major Groups in the CPI (1998-onward definition).

**Table 6**

**Frequency vs. Durability and Cyclicity**  
**Across U.S. CPI Categories**

WLS of Frequency on →	<i>Durability</i>	<i>Cyclicity</i>
<i>All Items</i>		
Posted Prices	0.60 (0.69)	3.23 (1.05)
Regular Prices	0.49 (0.71)	3.77 (1.07)
<i>Processed Items</i>		
Posted Prices	1.47 (0.41)	3.29 (0.58)
Regular Prices	1.42 (0.39)	3.79 (0.48)

*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) from January 1988 through October 2009. Entries are Weighted Least Square (WLS) Coefficients across Expenditure Classes, where the weights are based on average shares of consumer expenditure. Regressions include a constant. Coefficient standard errors are in parentheses. We were able to match estimates of durability (cyclicity) for 65 (64) BLS Expenditure Classes (1998-onward definition) for All Items, and 59 (58) for Processed Items. Frequency is calculated as the weighted mean across ELIs (with each ELI mean itself a weighted average across quotes within the ELI). Durability is based on estimates reported in Bils and Klenow (1998). Cyclicity is based on an OLS regression coefficient for each Expenditure Class of *real* monthly consumption growth on aggregate real monthly consumption growth, using seasonally adjusted Detailed Expenditure data from the U.S. National Income and Product Accounts, January 1990 through May 2009.

**Table 7****U.S. CPI Price Durations under Various Exclusions**

Case	Implied Median Duration	Implied Mean Duration
Posted Prices	3.4 months	6.2 months
No V Shapes	5.0	7.1
Like Prices	5.9	7.2
Regular Prices	6.9	8.0
No Substitutions	8.3	10.1
Adjacent Prices	9.0	11.2
1988-1997 Posted	3.6	6.3
1998-2009 Posted	3.4	6.6

*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) from January 1988 through October 2009. The implied durations are inverses of the monthly frequencies. Means and medians use ELI weights based on the BLS consumer expenditure surveys and unpublished weights for each quote based on BLS point-of-purchase surveys.

No V shapes: lower middle prices are replaced with identical neighbors.

Like prices: regular (sale) price is compared only to the previous regular (sale) price.

Regular prices: posted prices excluding sale prices.

No substitutions: only regular prices in between item substitutions are compared.

Adjacent prices: only consecutive monthly regular prices in between substitutions.

1988–1997 only: sample is confined to 1988–1997 (posted prices).

1998–2009 only: sample is confined to 1998–2009 (posted prices).

**Table 8****Share of “Top” Prices in Each U.S. CPI Quote-Line**

<b>Median % (Mean %)</b>	<b>Top Price</b>	<b>Top 2 Prices</b>	<b>Top 3 Prices</b>	<b>Top 4 Prices</b>	<b># of Quotes</b>
<i>All Items</i>	31.4 (37.6)	50.9 (53.2)	62.7 (61.3)	70.1 (66.2)	50 (53.1)
<i>By Major Group</i>					
Apparel	36.7 (41.2)	58.3 (59.9)	71.4 (70.5)	79.0 (77.0)	38 (41.2)
Education and Communication	29.2 (33.1)	47.4 (50.1)	61.3 (61.7)	73.0 (69.3)	42 (52.3)
Food	42.4 (47.0)	66.7 (66.7)	79.2 (76.3)	85.5 (81.3)	51 (51.7)
Home Furnishings	24.0 (30.5)	40.0 (43.5)	50.0 (50.4)	56.9 (54.7)	56 (63.4)
Medical Care	48.9 (53.3)	72.1 (71.1)	82.8 (79.0)	88.0 (84.1)	50 (49.5)
Recreation	40.4 (46.2)	64.9 (65.4)	77.6 (75.4)	85.1 (81.2)	49 (48.9)
Transportation	16.3 (22.6)	28.6 (34.8)	38.0 (43.1)	45.9 (49.8)	49 (53.1)
Other Goods and Services	50.0 (54.1)	70.6 (68.5)	79.3 (74.6)	84.0 (78.2)	51 (52.6)

*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) from January 1988 through October 2009. Both regular and sale prices are included. Entries are weighted medians (means) across quotelines of the top (most common) *n* prices in each quote-line as a share of all prices in the quote-line.

**Table 9****Share of “Reference” Prices in Each U.S. CPI Quote-Line**

	Median %	Mean %	S.D. %	Median Duration (Months)
<i>All Items</i>	84.2	78.5	41.0	11.0
<i>By Major Group</i>				
Apparel	77.8	77.1	35.1	11.0
Education and Communication	93.1	83.0	57.4	7.3
Food	88.1	82.2	27.2	13.5
Home Furnishings	82.1	77.9	46.8	6.9
Medical Care	100.0	93.7	27.1	18.8
Recreation	92.3	87.4	36.3	14.6
Transportation	62.5	64.2	70.4	6.1
Other Goods and Services	96.3	89.4	36.9	19.3

*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) from January 1988 through October 2009. Both regular and sale prices are included. “Reference” prices are the most common prices in the 13 month window centered on the current month within each quote-line. The share of reference prices in all prices was calculated for each quote-line, and then the weighted median (mean) and standard deviation of this % was calculated across quote-lines. We calculated the weighted median duration of reference prices across Major Groups using MLE as in Klenow and Kryvtsov (2008).

**Table 10**

**Share of Prices that are “Novel”  
in Each U.S. CPI Quote-Line**

	Median	Mean
<i>All Items</i>	16.1	25.0
<i>By Major Group</i>		
Apparel	21.2	26.3
Education and Communication	23.3	28.3
Food	10.3	13.8
Home Furnishings	23.6	35.9
Medical Care	6.7	8.2
Recreation	9.9	12.9
Transportation	35.0	38.3
Other Goods and Services	7.1	10.7

*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) from January 1988 through October 2009. Both regular and sale prices are included. “Novel” prices are prices that did not appear in any of the previous 12 months for the same quote-line.

**Table 11**

**Share of “Comeback” Prices  
in Each U.S. CPI Quote-Line**

	<b>Median</b>	<b>Mean</b>
<i>All Items</i>	0.8	14.0
<i>By Major Group</i>		
Apparel	19.5	25.2
Education and Communication	0	4.8
Food	14.9	24.7
Home Furnishings	6.3	16.4
Medical Care	0	3.7
Recreation	0	11.0
Transportation	0	6.2
Other Goods and Services	0	8.8

*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) from January 1989 through October 2009. Both regular and sale prices are included. “Comeback” prices are current prices that both occurred and were interrupted in the previous 12 months for the same quote-line.

**Table 12**

**Posted vs. Regular Price Inflation in the U.S. CPI**

	Posted Price $\pi$	Regular Price $\pi$	Sale-Related $\pi$
<i>Monthly</i>			
Variance relative to that of Posted Price $\pi$	1	0.875	0.125
Correlation with Posted Price $\pi$	1	0.963 (0.005)	0.365 (0.054)
Serial Correlation	0.394 (0.053)	0.420 (0.051)	-0.076 (0.062)
Cumulative Inflation After 1 Year	1.507 (0.270)	1.480 (0.267)	1.089 (0.168)
<i>Quarterly</i>			
Variance relative to that of Posted Price $\pi$	1	0.925	0.075
Correlation with Posted Price $\pi$	1	0.979 (0.005)	0.285 (0.100)
Serial Correlation	0.174 (0.106)	0.142 (0.108)	-0.017 (0.110)

*Source:* CPI-RDB. Inflation rates are for the top three cities (New York, Los Angeles, and Chicago). Monthly series go from February 1988 through October 2009, and quarterly series from 1988:2 through 2009:3. Posted Prices include sale prices, whereas Regular Prices replace sale prices with the previous regular price. Monthly (seasonal) dummies are removed separately from Posted Price Inflation and Regular Price Inflation before moments are calculated. “Sale-related inflation” refers to the series obtained by subtracting Regular Price Inflation from Posted Price Inflation. The “Cumulative Inflation After 1 Year” is obtained from regressing  $\ln P_{t+12} - \ln P_t$  on  $\ln P_{t+1} - \ln P_t$ .



**Table 13****Posted vs. Reference Price Inflation in the U.S. CPI**

	Posted Price $\pi$	Reference Price $\pi$	Non- Reference $\pi$
<i>Monthly</i>			
Variance relative to that of Posted Price $\pi$	1	0.462	0.538
Correlation with Posted Price $\pi$	1	0.689 (0.033)	0.734 (0.029)
Serial Correlation	0.398 (0.054)	0.461 (0.050)	0.119 (0.063)
Cumulative Inflation After 1 Year	1.569 (0.281)	3.067 (0.395)	0.906 (0.162)
<i>Quarterly</i>			
Variance relative to that of Posted Price $\pi$	1	0.592	0.408
Correlation with Posted Price $\pi$	1	0.822 (0.037)	0.642 (0.066)
Serial Correlation	0.175 (0.109)	0.441 (0.091)	-0.217 (0.107)

*Source:* CPI-RDB. Inflation rates are for the top three cities (New York, Los Angeles, and Chicago). Monthly series go from August 1988 through April 2009, and quarterly series from 1988:3 through 2009:1. Reference Prices represent the most common price in the 13 month window centered on the current month for each item. Monthly (seasonal) dummies are removed separately from Posted Price Inflation and Reference Price Inflation before moments are calculated. “Non-reference Inflation” refers to the series obtained by subtracting Reference Price Inflation from Posted Price Inflation. The “Cumulative Inflation After 1 Year” is obtained from regressing  $\ln P_{t+12} - \ln P_t$  on  $\ln P_{t+1} - \ln P_t$ .

## Table 14

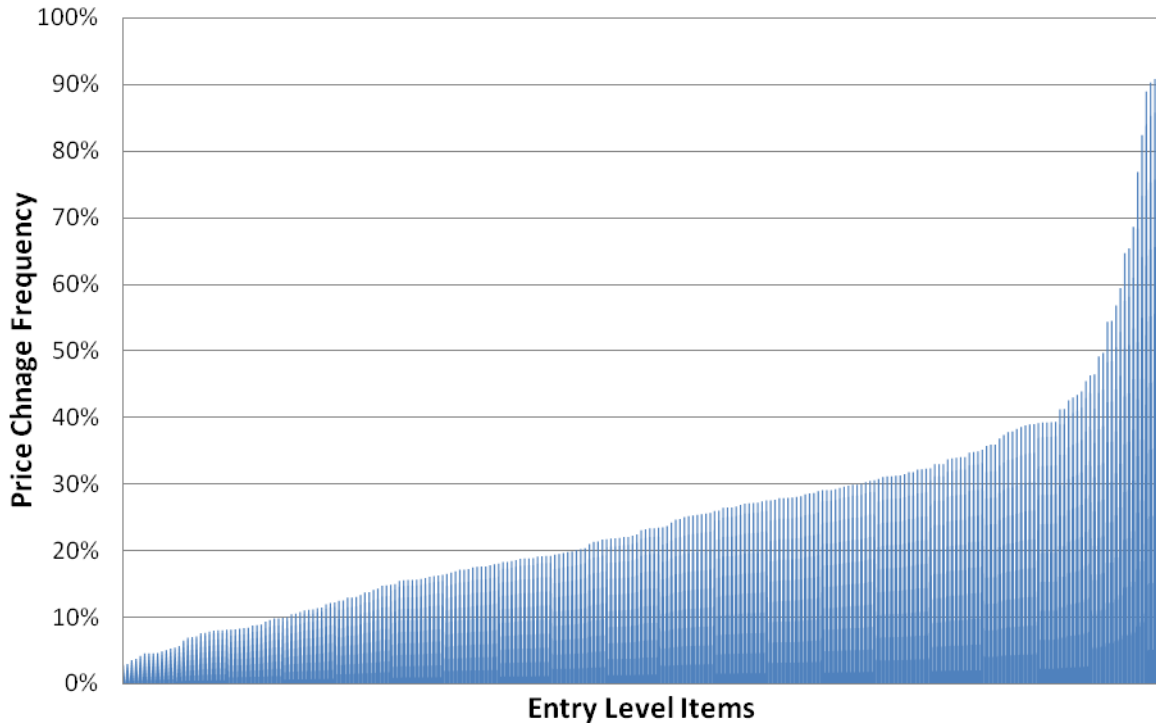
### Model Features and the Facts

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	Consistent Features	Inconsistent Features
<b>Half of prices change only a few times a year</b>	Menu costs; Sticky input prices	Price indexation; Convex adjustment costs
<b>Temporary price changes are common</b>	Sticky set; Price discrimination	Menu costs with flexible marginal cost
<b>Frequency differs persistently across goods</b>	Menu costs; Sticky information	Exogenous frequency of price changes
<b>Price changes are large on average</b>	Big menu costs; Rational inattention	Small idiosyncratic shocks with strong complementarities
<b>Many price changes are small</b>	Sticky information; Sticky path	Large menu costs for changing each individual price
<b>Relative price changes are transitory</b>	Transitory Idiosyncratic shocks	Strong micro complementarities
<b>Price changes are not well synchronized</b>	Big idiosyncratic shocks	Small idiosyncratic shocks with strong complementarities
<b>Old prices do not change by bigger amounts</b>	Menu costs; Lumpy shocks	Time-dependent pricing with persistent shocks

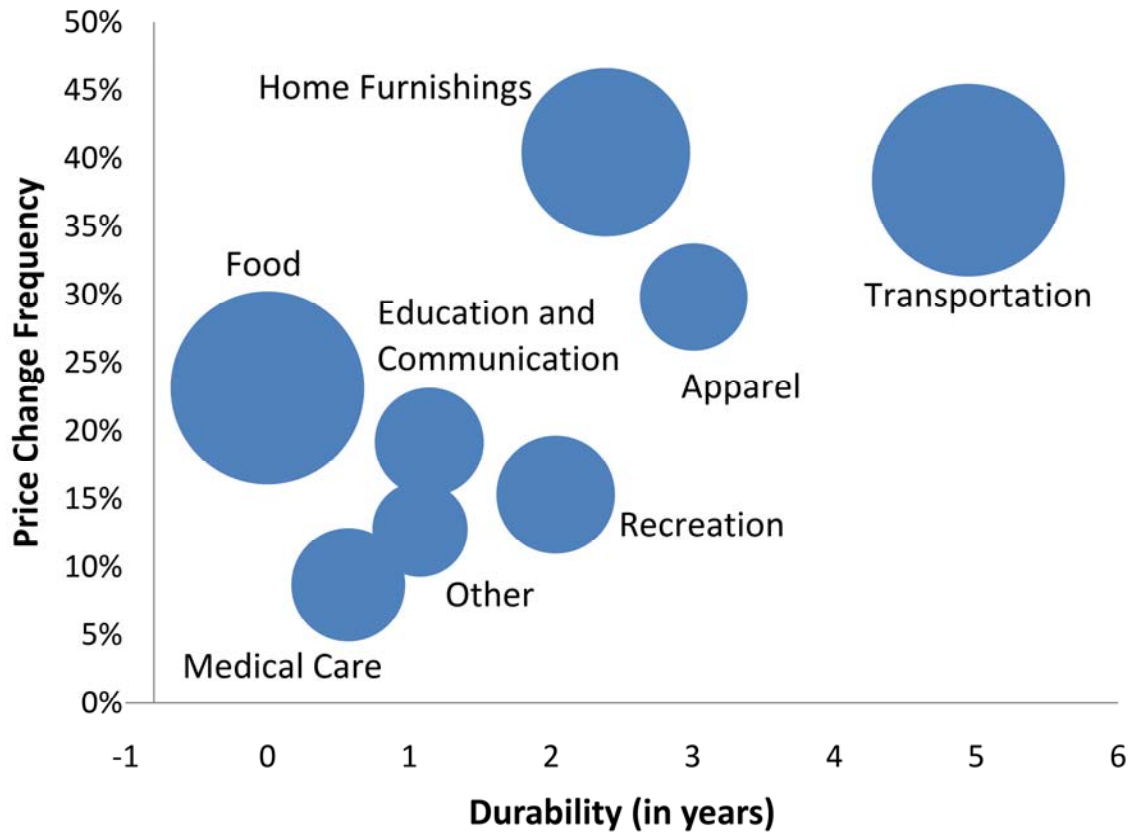
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**Figure 1: Price Change Frequency by Product Category**



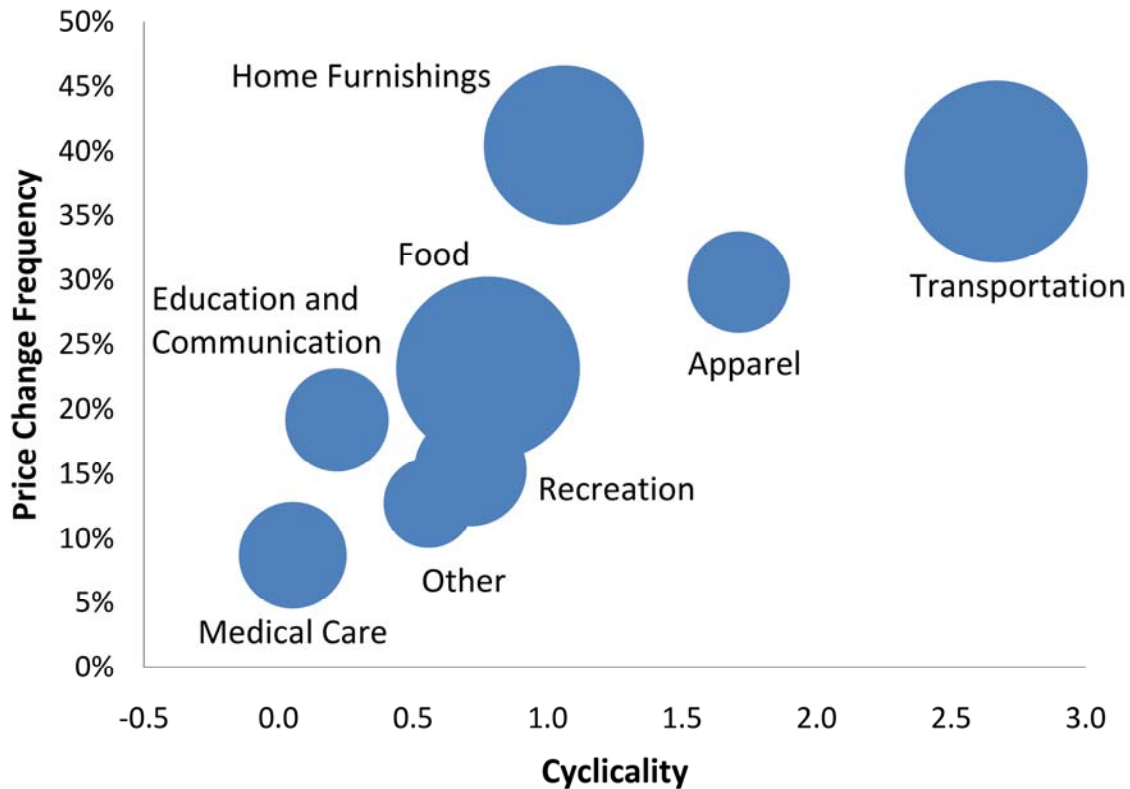
*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) January 1988 through October 2009. Each bar corresponds to an ELI product category (weighted by expenditure), and price change frequency is calculated as the weighted average across quotes within the ELI. Prices include sales and regular prices.

**Figure 2: Frequency vs. Durability in the U.S. CPI**



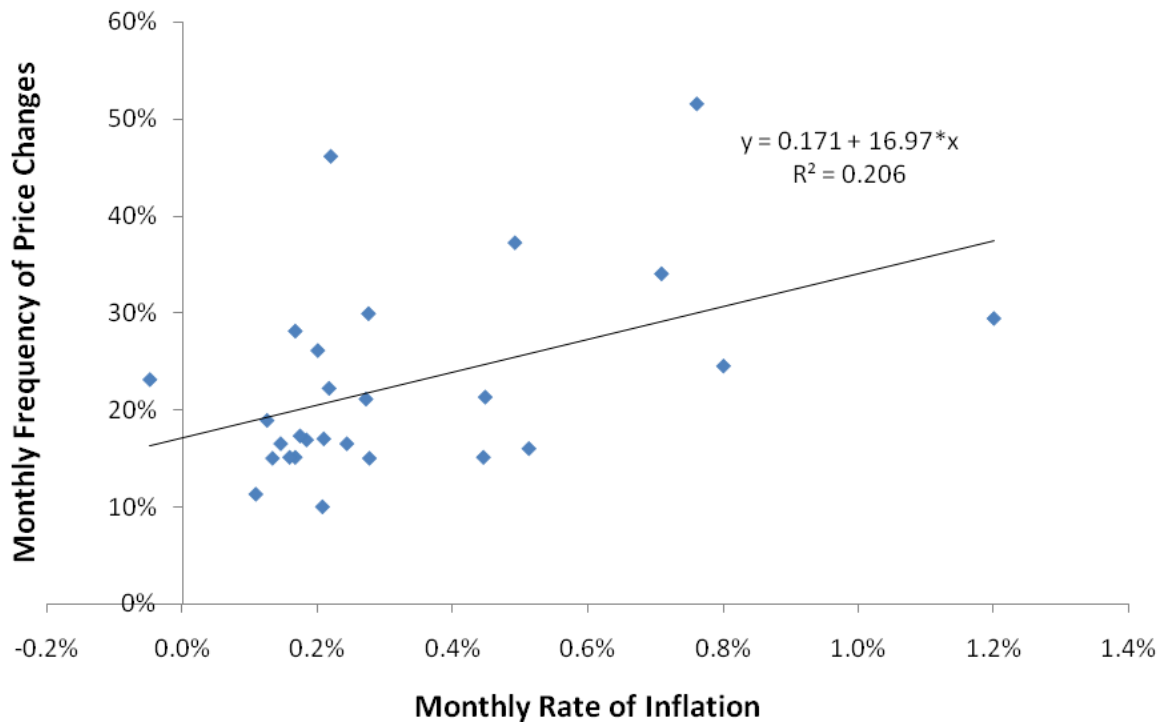
*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) January 1988 through October 2009. Each circle is one of eight Major Groups in the CPI (1998-onward definition), with the area proportional to the average expenditure weight over the sample. Prices include sales and regular prices. Frequency is calculated as the weighted mean across ELIs (with each ELI mean itself a weighted average across quotes within the ELI). Durability is based on estimates reported in Bils and Klenow (1998).

**Figure 3: Frequency vs. Cyclicity in the U.S. CPI**



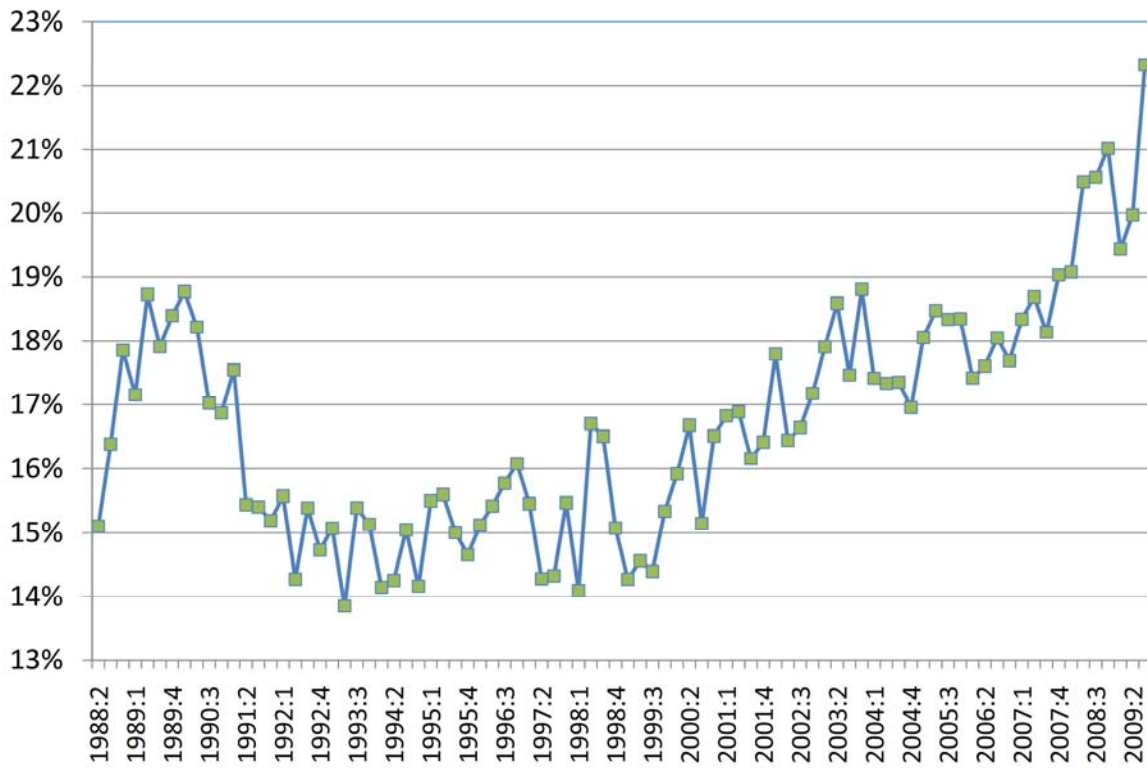
*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) January 1988 through October 2009. Each circle is one of eight Major Groups in the CPI (1998-onward definition), with the area proportional to the average expenditure weight over the sample. Prices include sales and regular prices. Frequency is calculated as the weighted mean across ELIs (with each ELI mean itself a weighted average across quotes within the ELI). Cyclicity is based on an OLS regression coefficient for each Major Group of *real* monthly consumption growth on aggregate real monthly consumption growth, using seasonally adjusted Detailed Expenditure data from the U.S. National Income and Product Accounts, January 1990 through May 2009.

**Figure 4: Cross-country Frequency vs. Inflation**



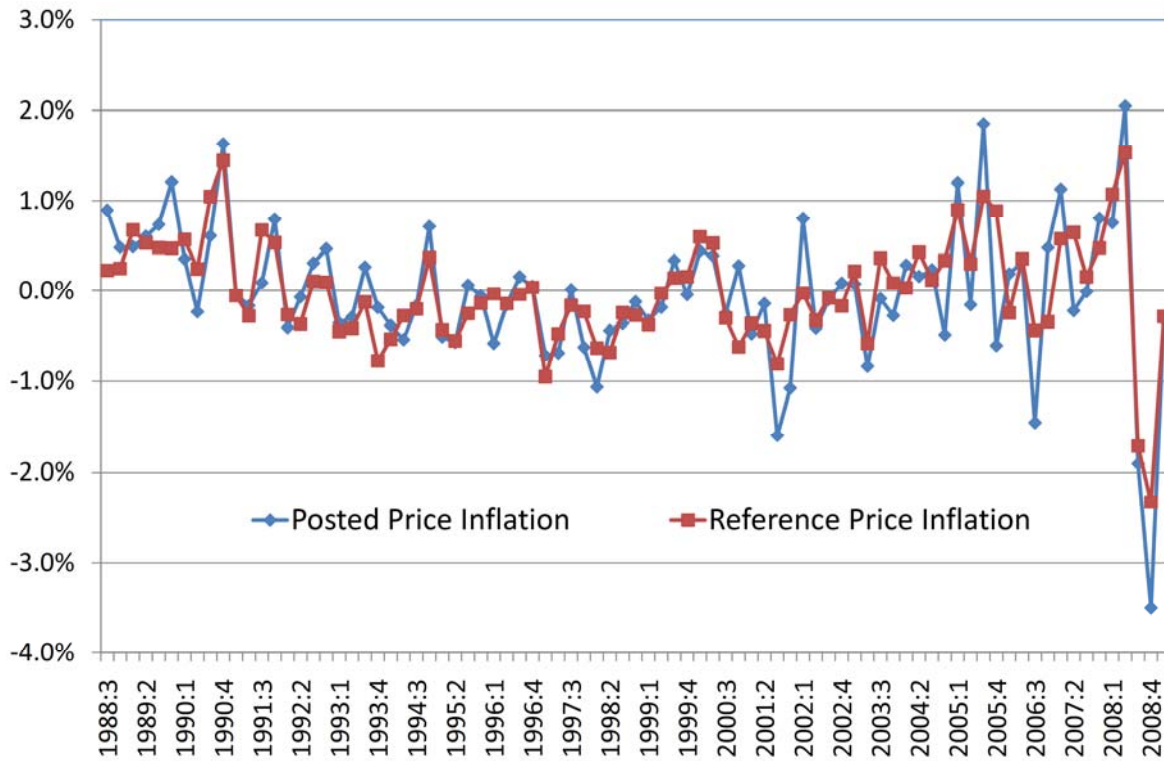
*Notes:* Each data point represents one study from Table 1, with the Gouvea (2007) study excluded because of its overlap with the Barros et al. (2009) study. The monthly frequency of price change is as reported in Table 1, while the monthly rate of inflation is based on authors' calculations.

**Figure 5: U.S. Monthly CPI Frequency Over Time**



*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago), based on regular prices for processed items. We first calculated the monthly frequency of price changes (weighted mean across ELIs) from April 1988 through September 2009. We then took out seasonal (monthly) dummies, and calculated deviations from them. Finally, we averaged the monthly deviations in each quarter, and added back the mean across all quarters, to produce quarterly data from 1988:2 through 2009:3.

**Figure 6: Posted vs. Reference Price Inflation, U.S. CPI**



*Source:* CPI-RDB. Data are for the top three cities (New York, Los Angeles, and Chicago) for the quarters 1988:3 through 2009:1. Posted price inflation includes sale and regular prices. Reference price inflation is based on replacing the current posted price with the modal price in the 13-month window centered on the current month (for each micro datapoint). Inflation rates were calculated at the monthly level, and monthly (seasonal) dummies were removed for each series separately. The monthly residuals were summed to create quarterly data.



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