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# Evaluating Effects of Tax Preferences on Health Care Spending and Federal Revenues

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## Executive Summary

In this paper, we calculate the consequences for health spending and the federal budget of an above-the-line deduction for out-of-pocket health spending. We show how the response of spending to this expansion in the tax preference can be specified as a function of a small number of behavioral parameters that have been estimated in the existing literature. We compare our estimates to those from other researchers. And, we use our analysis to derive some implications for tax policy toward HSAs.

## 1. Introduction

As Pauly's (1986) classic review shows, virtually all observers of health policy since Feldstein's (1973) seminal article have agreed that the tax preference for employer-provided health insurance—under which employer contributions to employee health insurance are deductible to the employer and non-taxable to the employee—encourages overconsumption of health services in the United States. By making health spending in general, and insured health spending in particular, appear less costly than they are, the tax preference gives employees the incentive to take compensation as health insurance rather than cash, even if they would otherwise prefer not to.

Both the budget cost of the tax preference, and its potential implications for efficiency in markets for health services, are large. Table 3.1 provides three estimates of the federal revenue loss from the tax preference in 2004. Shiels and Haught (2004) estimate the revenue loss to be \$188.5 billion. According to them, the loss from the exclusion from the personal income tax base of employer contributions to employee

**Table 3.1**

Cost to the Federal Budget of Existing Tax Preferences for Health Spending 2004  
(in billions of dollars)

	Study		
	Shiels and Haught (2003)	JCT (2003)	US Treasury in OMB (2003)
Total	\$188.5		
Exclusion of employer contribution to HI premiums from...			
Social security payroll tax base	\$52.2		
Medicare payroll tax base	\$14.2		
Personal income tax base	\$114.7	\$101.0	\$123.9
Deduction for out-of-pocket expenses in excess of 7.5% of adjusted gross income	\$7.4	\$5.9	\$6.3

Note: Three studies' estimation methods differ slightly; see Shiels and Haught (2003) for discussion.

health insurance alone was \$114.7 billion. The Joint Committee on Taxation (2003) estimates this loss to be slightly less (\$101.0 billion); the Department of the Treasury in the Office of Management and Budget (OMB 2003) estimates it to be slightly more (\$123.9 billion). Everyone agrees, however, that it dwarfs the revenue loss from all other health tax preferences (such as the deduction for health spending in excess of 7.5 percent of adjusted gross income), and indeed is the single largest tax preference in the federal budget. According to the Department of the Treasury in OMB (2003), the loss from the employer exclusion surpasses the loss from the deductibility of mortgage interest, state and local property taxes, and all capital gains tax preferences. Indeed, the only personal tax preferences that come close are the various exclusions for retirement savings contributions.

Table 3.2 presents trends in health spending by payor and form of spending for 1993–2003. The table documents the well-known growth in the magnitude of real spending over this period: Particularly noteworthy is the change in the *form* of health spending, from (largely taxable) out-of-pocket to tax-preferred insured spending. According to the table, real employer and employee payments for insured health spending rose about 50 percent over the period, while out-of-pocket spending rose less than half as much. If the tax preference contributed to this, and

**Table 3.2**

Health Spending, by Payor and Form of Spending 1993–2003 (in billions of 2003 dollars)

Type of Payor and Form of Spending	1993	Year 1998	2003	% change 1993–2003
<b>TOTAL</b>	\$1,087.5	\$1,257.2	\$1,614.2	48.4%
<b>Private</b>	\$694.4	\$806.8	\$992.2	42.9%
<i>Private Business</i>	\$280.9	\$325.1	\$423.0	50.6%
Employer contribution to HI premiums	\$205.1	\$237.2	\$320.6	56.3%
Employer payments of Medicare payroll taxes	\$45.5	\$60.8	\$64.3	41.4%
Workers' compensation and other	\$30.4	\$27.1	\$38.1	25.5%
<i>Household</i>	\$367.5	\$423.5	\$512.6	39.5%
Employee contribution to HI premiums	\$110.4	\$129.5	\$174.1	57.8%
Employee payments of Medicare payroll taxes	\$55.5	\$78.2	\$86.0	55.0%
Individual payments of Medicare SMI premiums	\$15.1	\$17.5	\$22.0	45.6%
Out-of-pocket spending	\$186.6	\$198.4	\$230.5	23.6%
<i>Other Private</i>	\$46.0	\$58.2	\$56.6	23.1%
<b>Public</b>	\$393.1	\$450.5	\$622.0	58.2%
<i>Federal Government</i>	\$223.0	\$243.0	\$344.0	54.3%
Employer contributions to HI premiums	\$14.6	\$12.9	\$19.7	34.9%
Medicaid	\$99.2	\$115.0	\$160.9	62.2%
Medicare (net of payroll tax receipts) and other	\$106.2	\$112.3	\$160.2	50.9%
<i>State and Local Government</i>	\$170.1	\$207.5	\$278.1	63.5%
Employer contributions to HI premiums	\$48.5	\$55.1	\$86.2	77.7%
Medicaid	\$58.2	\$83.1	\$111.8	92.2%
Other	\$57.2	\$61.9	\$71.5	25.1%

Source: Health United States (2005), table 127, deflated with the CPI.

insured spending is subject to a greater degree of moral hazard, then the impact of the preference on efficiency could be substantial.

Taken together, these factors have led academic researchers to focus on the consequences of revoking the tax preference. Yet, policymakers over the past 30 years have taken an alternative approach; they have sought to level the tax playing field by expanding the tax preference rather than eliminating it.<sup>1</sup> In 1978, changes to section 125 of the Internal Revenue Code allowed health expenditures made through an employer-provided Flexible Spending Account (FSA) to be deductible to the employer but nontaxable to the employee.<sup>2</sup> In 1996, the Health

Insurance Portability and Accountability Act allowed employees of small businesses who were covered by certain high-deductible health plans (HDHPs) to make tax-free contributions to a Medical Savings Account (MSA). Funds from an MSA can be withdrawn, tax free, to pay for medical expenses in the present or the future; if used for other purposes after age 65, MSA distributions are taxed as ordinary income. Under Treasury regulations issued in 2002, sections 105 and 106 of the Internal Revenue Code allow health reimbursement accounts (HRAs) to reimburse employees for medical expenses with before-tax dollars, without the use-it-or-lose-it provision of section 125 cafeteria plans.<sup>3</sup> In 2003, the Medicare Prescription Drug, Improvement, and Modernization Act allowed employers and individuals with any HDHP to make tax-free contributions to a health savings account (HSA). President Bush has proposed expanding the use of HSAs by liberalizing their contribution limits.

Conditional on the tax preference for insurance remaining in place, the consequences of these expansions for health spending, and economic efficiency, are theoretically indeterminate. Expanding the tax preference has two opposing effects. First, expansion *lowers* the overall price of health care relative to other goods and services, which increases distortionary spending. Second, expansion *raises* the price of purchasing health care through insurance relative to out-of-pocket. The second effect induces people to shift to health plans with higher deductibles and coinsurance rates, which, in turn, lowers distortionary spending.

Thus, assessing the effects of expanding the tax preference to out-of-pocket spending is important for evaluating existing and proposed tax policies toward health care. Yet, very little work has sought to estimate these effects and to understand their sensitivity to assumptions about the demand for health services and insurance. In this paper, we present a simplified version of the approach in John Cogan, R. Glenn Hubbard, and Daniel Kessler (CHK 2005) to calculating the effects of an above-the-line deduction for out-of-pocket health spending, which we term "full deductibility." In that paper, we show how the response of total health spending to an expansion in the tax preference can be specified as a function of a small number of behavioral parameters that have been estimated in the existing literature. This paper expands on that work in three ways. First, it calculates the effects of full deductibility on out-of-pocket spending, total spending, and the government budget under a range of parameter values. Second, it compares our estimates

to those from other researchers. Third, it uses our analysis to derive some implications for tax policy toward HSAs.

## 2. Assessing the Effects of Tax Deductibility on Health Spending

As reviews by Pauly (1986) and, more recently, Selden and Moeller (2000) show, a substantial body of research has sought to assess the effects of revoking the tax preference for employer-provided health insurance. Considerably less work has focused on the effects of extending the preference to out-of-pocket spending. Jack and Sheiner (1997) simulate the effects on insurance policy choice, health spending, and efficiency of both revoking and extending the tax preference. Those authors show that extending deductibility might actually reduce health spending and improve efficiency, by leading to such a large increase in the effective coinsurance rate that the gain from the reduction in moral hazard swamps the loss from the reduction in the overall price of health care. A recent working paper by Jack, Levinson, and Rahardja (2005) provides empirical support for this hypothesis. They show that, correcting for selection effects, FSAs are associated with effective coinsurance rates that are about 7 percentage points higher, relative to a sample average coinsurance rate of 17 percent. This finding suggests that making out-of-pocket health spending deductible, which an FSA effectively does, would significantly change the form of the average health insurance contract.

In CHK (2005), we derive the relationship between the impact on health spending of making out-of-pocket expenses tax deductible and two parameters from economic studies: the price elasticity of health spending, and the elasticity of the coinsurance rate with respect to the tax preference for insured spending. We specify health spending  $E$  as a function of the after-tax price of health services relative to all other goods  $p$  and the tax preference for out-of-pocket spending relative to insured spending  $t_o/t_i$ ,  $E(p, t_o/t_i)$ . In a world without taxes,  $p$  is the price of health services  $p^*$ . In a world with tax preferences,  $p$  is  $p^*$  multiplied by the weighted average of the tax preferences for out-of-pocket spending  $t_o$  and insured spending  $t_i$ ,  $p = p^* \times [ct_o + (1 - c)t_i]$ , where  $t_o$  and  $t_i$  are weighted by the quantity shares of out-of-pocket and insured spending  $c$  and  $(1 - c)$ , respectively. The share  $c$  can also be thought of as the coinsurance rate—that is, the share of spending out-of-pocket in the absence of tax preferences.

So

$$\frac{dE}{dt_o} = \frac{\partial E}{\partial p} \times \frac{\partial p}{\partial t_o} + \frac{\partial E}{\partial(t_o/t_i)} \times \frac{\partial(t_o/t_i)}{\partial t_o} = \frac{\partial E}{\partial p} \times \frac{\partial p}{\partial t_o} + \frac{\partial E}{\partial(t_o/t_i)} \times \frac{1}{t_i}$$

and

$$\frac{dE}{dt_i} = \frac{\partial E}{\partial p} \times \frac{\partial p}{\partial t_i} + \frac{\partial E}{\partial(t_o/t_i)} \times \frac{\partial(t_o/t_i)}{\partial t_i} = \frac{\partial E}{\partial p} \times \frac{\partial p}{\partial t_i} - \frac{\partial E}{\partial(t_o/t_i)} \times \frac{t_o}{t_i^2}$$

Then the sum of these two equations, in elasticity terms, is:

$$\begin{aligned} \frac{dE/E}{dt_o/t_o} + \frac{dE/E}{dt_i/t_i} &= \frac{\partial E/E}{\partial p/p} \times \frac{\partial p/p}{\partial t_o/t_o} + \frac{\partial E/E}{\partial(t_o/t_i)} \times \frac{t_o}{t_i} + \frac{\partial E/E}{\partial p/p} \\ &\times \frac{\partial p/p}{\partial t_i/t_i} - \frac{\partial E/E}{\partial(t_o/t_i)} \times \frac{t_o}{t_i} \end{aligned}$$

or

$$\begin{aligned} \frac{dE/E}{dt_o/t_o} + \frac{dE/E}{dt_i/t_i} &= \frac{\partial E/E}{\partial p/p} \times \left[ \frac{\partial p/p}{\partial t_o/t_o} + \frac{\partial p/p}{\partial t_i/t_i} \right] \\ &= \frac{\partial E/E}{\partial p/p} \times \left[ \frac{p^* \times [ct_o + (1-c)t_i]}{p} + \theta(t_o, t_i, p^*) \right] \end{aligned}$$

or

$$e(t_o) + e(t_i) = e(p) \times [1 + \theta(t_o, t_i, p^*)],$$

where  $e(t_o)$  is the elasticity of spending with respect to the tax preference for out-of-pocket spending;  $e(t_i)$  is the elasticity of spending with respect to the preference for insured spending; and  $e(p)$  is the price elasticity of spending.

In CHK (2005), we show that under reasonable assumptions and current market conditions and tax preferences,<sup>4</sup>  $\theta(\cdot)$  is small, so

$$e(t_o) + e(t_i) \approx e(p).$$

Finally, we translate the results from previous studies into these terms. For example, the equation above can be rewritten as

$$e(t_o) + e(c) \times e(c, t_i) \approx e(p)$$

where  $e(c)$  is the elasticity of spending with respect to the coinsurance rate and  $e(c, t_i)$  is the elasticity of the coinsurance rate with respect to

the tax preference for insured spending. If demand curves are locally linear, then  $e(p) = e(ap)$  for any positive constant  $a$ , so  $e(p) = e(c)$ , which implies:

$$e(t_i) \approx e(p) \times (1 - e(c, t_i)).$$

Assessing the effects of extending deductibility thus requires estimates of  $e(p)$  and  $e(c, t_i)$ . There is a range of estimates of  $e(p)$ . Based on the RAND Health Insurance Experiment, Manning and colleagues (1987) estimate  $e(p) = -0.2$  in arc elasticity terms.<sup>5</sup> In more recent work, Eichner (1998, table 1) estimates  $e(p) = -0.7$  (average for all employees 1990–92, also in arc elasticity terms).<sup>6</sup> In addition, even the high end of this range may understate the impact of a market-wide change in incentives such as extending deductibility. All of the estimates of  $e(p)$  are based on responses to individual-level changes in copayments, which may be smaller than the responses to more widespread changes in insurance contracts that fundamentally alter how doctors practice medicine. Finkelstein (2005), for example, shows that the change in hospital spending associated with the introduction of Medicare was far greater than the elasticities from the RAND Experiment would have predicted.

Less disagreement surrounds the magnitude of  $e(c, t_i)$ . Several studies have assessed the effect of the tax preference on coinsurance rates. These can be used to compute  $e(c, t_i)$ . Early simulations by Feldstein and Friedman (1977) suggest that revoking the tax preference for employer-provided insurance would lead to a doubling in the coinsurance rate (from approximately 25 to 50 percent). This finding is consistent with an unpublished estimate by Phelps (1986). More recent work leads to virtually the same conclusions. At conservative levels of consumer risk aversion and  $e(p)$ , simulations by Jack and Sheiner (1997, table 2) find that the tax preference for insurance has led optimal coinsurance rates to shrink from 33–67 percent to 20–30 percent. Assuming an average marginal (payroll plus income) tax rate of 30 percent,<sup>7</sup> revocation leading to doubling of coinsurance rates from  $c$  to  $2c$  implies an  $e(c, t_i)$  in arc elasticity terms of 1.9, because:

$$\frac{\frac{2c - c}{[2c + c] \frac{1}{2}}}{\frac{1 - (1 - .3)}{[1 + (1 - .3)] \frac{1}{2}}} = 1.9.$$



Table 3.3 presents calculations of the effects of revoking and extending the income and payroll tax preference for  $e(p) = -0.2, -0.45, \text{ and } -0.7$  at  $e(c, t_i) = 1.9$ . The first two rows of table 3.3 present the effects of revoking the income and income-and-payroll tax preference which is simply

$$e(p) \times e(c, t_i) \times \frac{dt_i}{t_i}.$$

The third and fourth rows of table 3.3 present the effects of extending the income and income-and-payroll tax preference which is simply

$$e(p) \times (1 - e(c, t_i)) \times \frac{dt_o}{t_o}.$$

Our estimates of the effects of tax policy on spending are within the range of those from other research. For example, according to Gruber (2002, table 5), removing the income tax subsidy for health insurance would result in a 32.8 percent decline in health spending, expressed as a percentage point change from its initial value. Expressed as a percentage-point change at the average (in order to make his estimate comparable to those in table 3.1), this amounts to a 39.2 percent decline—larger than the 31.8 percent decline in spending that we would predict even assuming an elasticity of spending with respect to the coinsurance rate of 0.7. Gruber's estimate, when combined with the consensus estimate

**Table 3.3**  
Effect on Health Spending of Changing the Tax Preference

	Elasticity of Spending with Respect to After-Tax Price of Health Care		
	-0.2	-0.45	-0.7
Effect of revoking:			
Income tax preference	-9.1%	-20.4%	-31.8%
Income plus payroll tax preference	-13.4%	-30.2%	-46.9%
Effect of extending to out-of-pocket:			
Income tax preference	-2.7%	-6.1%	-9.5%
Income plus payroll tax preference	-5.6%	-12.6%	-19.7%

Note: Assumes a health-spending-weighted average marginal income tax rate of .19, out-of-pocket-spending-weighted average marginal income tax rate of .14, and an average payroll tax rate of .13. Average marginal tax rates were calculated using MEPS, and include both households with and without income tax liabilities. See CHK (2005) for details.

of  $e(c, t_i) = 1.9$ , implies that extending the income tax preference to out-of-pocket spending would result in a decline in overall spending of 11.8 percent.<sup>8</sup> According to Jack and Sheiner (1997, table 1), extending the income and payroll tax preference to out-of-pocket spending would lead to a decline in spending of 4.9 percent ( $=0.3/6.15$ ), slightly lower than the decline in spending that we would predict assuming an elasticity of spending of 0.2.

### 3. Assessing the Budget Implications of Extending Deductibility to Out-of-Pocket Spending

In addition to reducing inefficient health spending, full deductibility reduces federal tax revenues. Deductibility has two effects on revenues—a loss from making previously taxable spending deductible, and a gain from the shift away from previously deductible health spending. We do not account for any possible spillover effects from privately-purchased health care to the Medicare or Medicaid programs.

The revenue *loss* consists of two components—the loss from allowing the above-the-line deduction of out-of-pocket spending, and the loss from purchases on health insurance being deducted above-the-line that are currently not deducted or deductible. The revenue *gain* also consists of two components. Tax revenues rise because higher policy deductibles will translate into a shift in employees' compensation away from excludable health spending to taxable wages.<sup>9</sup> The government picks up both payroll and income taxes on the portion of the wage increase directed to non-health spending (first component), and payroll taxes on the portion directed to out-of-pocket health spending (second component).

Table 3.4 presents our calculations of these losses and gains on an annual basis, in 2004 dollars. The top panel of the table shows the two components of the gross losses from full deductibility; the middle panel shows the two components of the gross gains; and the bottom panel shows the intermediate steps underlying the calculation of each of the components of the gross gains.

As the top panel shows, the gross losses are a mechanical consequence of the policy; they do not depend on behavior. We calculate that full deductibility would have a gross revenue cost of \$26.8 ( $= \$16.4 + \$10.4$ ) billion per year.

**Table 3.4**  
Effect on Tax Revenues of Full Deductibility

	Elasticity of Spending with Respect to After-Tax Price of Health Care		
	-0.2	-0.45	-0.7
<i>Revenue Losses</i>			
Loss from deduction of taxable out-of-pocket spending	-\$16.4	-\$16.4	-\$16.4
Loss from deduction of taxed or taxable insurance payments	-\$10.4	-\$10.4	-\$10.4
<i>Revenue Gains</i>			
Shift away from health spending			
Pickup of payroll taxes	\$2.4	\$5.3	\$8.1
Pickup of income taxes	\$3.5	\$7.7	\$11.8
Shift away from employer-insured spending toward out-of-pocket			
Pickup of payroll taxes	\$6.3	\$5.3	\$4.3
<i>Total</i>	-\$14.5	-\$8.4	-\$2.5
<i>Intermediate calculations underlying estimates of revenue gains</i>			
Reduction in health spending	-\$18.4	-\$40.7	-\$62.3
New coinsurance rate	33.3%	33.3%	33.3%
Reduction in employer-insured spending	-\$67.2	-\$81.6	-\$95.6
Increase in out-of-pocket spending (difference)	\$48.8	\$40.9	\$33.3

Note: Assumes \$117b of currently taxable out-of-pocket spending, \$74b of current taxed or taxable insurance payments, \$500b of current spending on employer health insurance, \$688b of total private health spending, an average coinsurance rate of 25%, a health-spending weighted average marginal income tax rate of .19, out-of-pocket spending-weighted average marginal income tax rate of .14, and an average payroll tax rate of .13. See note to table 3.1 and CHK (2005) for details on calculation of average marginal tax rates.

The first two rows of the middle panel present estimates of the first component of gross gains, the increase in tax revenues from the shift from insured spending to wages. Given the percentage effects on spending from table 3.3, full deductibility leads to a decline in spending of \$18.4 to \$62.3 billion (bottom panel, first row),<sup>10</sup> and in turn to an increase in payroll (income) tax revenues of \$2.4 to \$8.1 billion (\$3.5 to \$11.8 billion).

The third row of the middle panel presents estimates of the second component of gross gains, the increase in tax revenues from the shift

from insured to out-of-pocket spending. To calculate this shift, we make the (conservative) assumption that spending for individuals with employer-provided health insurance responds to full deductibility the same as does all private health spending. Under this assumption, the coinsurance rate that obtains under full deductibility  $c'$  (bottom panel, second row) is equal to

$$c \times \left[ 1 + \frac{2(\eta/e(p))}{2 - (\eta/e(p))} \right] / (1 - \tau)$$

where  $\tau$  is the average marginal tax rate, and  $2(\eta/e(p))/[2 - (\eta/e(p))]$  is the implied rise in the after-tax coinsurance rate necessary to induce the spending decline from table 3.3. The coinsurance rate under full deductibility would be 33.3 percent, regardless of the elasticity of demand for health services, up from a pre-deductibility average of 25 percent. (That the coinsurance rate does not vary with the elasticity of demand is, of course, a product of our model's assumption of a constant  $e(c, t)$ .)

This translates into a decline in spending on employer-provided health insurance, in percentage terms, of

$$\left( \frac{1 - c'}{1 - c} \right) \times (1 + \eta^*),$$

or in dollar terms, of \$67.2 to \$95.6 billion (bottom panel, third row). The increase in out-of-pocket spending subject to the payroll tax (bottom panel, fourth row) is thus the difference between this decline and the decline in total spending. According to this simple model, then full deductibility would lead to a significant but plausible increase in out-of-pocket spending, from \$149 billion in 2004 dollars (=\$117 billion in currently non-deductible out-of-pocket spending plus \$32 billion in deductible spending, see CHK (2005, Appendix E)) to \$190 billion (=\$149 billion + \$41 billion, see column 2), or 27.5 percent (=( \$190 billion - \$149 billion)/\$149 billion).

The calculations presented in table 3.4 make the important point that much of the gross revenue losses from full deductibility will be made up by revenue gains from the reduction in the inefficiency due to the highly-distortionary existing tax preference. Indeed, even assuming an elasticity of demand for health services of  $-0.2$ , \$12.2 billion of the \$26.8 billion (or 46 percent) of the losses will be undone; if the elasticity of

demand is  $-0.45$ , fully \$18.3 billion of the \$26.8 billion (or 68 percent) of losses are undone.

#### 4. Implications for Policy toward HSAs

Like full deductibility, allowing HSA contributions to be tax-deductible gives a tax preference to out-of-pocket spending. Under current law, a holder of an HDHP (i.e., a health plan with a deductible of at least \$1,050/\$2,100 in 2006 (individual/family)) can contribute to an HSA the amount of the deductible, but not more than \$2,700 (individual) or \$5,450 (family).<sup>11</sup> The contribution is deductible from federal income taxes and from income taxes in 44 states.<sup>12</sup> If the contribution is made by a person's employer, it is also excludable from the Social Security tax base. The contribution accumulates interest tax-free and is non-taxable on distribution, if spent on health services; it is taxable as ordinary income if distributed for any other purpose after age 65.

HSAs differ from full deductibility in three key ways. First, an individual can only have an HSA if they are enrolled in an HDHP. Second, an individual can deduct HSA contributions from his or her taxable income up to the amount of their HDHP's deductible, whether or not they incur any health expenses, but can not deduct more, even if they have coinsurance payments in excess of the deductible. Third, an HSA allows an individual to save tax-free for future health expenses or retirement, whereas full deductibility only allows deduction of current health expenses.<sup>13</sup>

For consumers who use HSAs only as a vehicle to deduct current health expenses, the most important difference between HSAs and full deductibility is the minimum deductible requirement of HDHPs. If, for these individuals, all of the expenditure-reducing incentive effects of full deductibility were channeled through insurance policy deductibles (rather than coinsurance rates), then deductibles would have to rise from a typical value of \$221 (2004 dollars)<sup>14</sup> to approximately \$290,<sup>15</sup> far less than the 2006 mandated HDHP minimum of \$1,050. Because HSAs are indistinguishable from full deductibility for these consumers, this implies that they would prefer a lower deductible than the mandated minimum.

Because most taxpayers do not exhaust their existing retirement savings incentives (CBO 2003), and therefore are likely to treat HSAs

primarily as a vehicle to deduct current expenses, HSAs as currently formulated will likely be taken up by fewer people than would full deductibility. However, among HSA enrollees, HSAs bring health spending much closer to the efficient level (the level that would be preferred in the absence of any tax preference) than full deductibility. If all of the expenditure-reducing incentive effects of revoking the income and payroll tax preferences were channeled through insurance policy deductibles, then deductibles would rise to approximately \$1,680 in 2004 dollars.<sup>16</sup> These back-of-the-envelope comparisons between full deductibility and HSAs are consistent with empirical studies of HSAs (see, for example, Melinda Buntin et al. 2005; and Roger Feldman et al. 2005) and MSAs (see, for example, Larry Ozanne 1996; and Emmett Keeler et al. 1996).

The most efficient way to expand HSAs would be to allow deductibility of all out-of-pocket payments for people with insurance (not just those toward the policy deductible), but limit the budget consequences of HSAs by capping deductible contributions at a fixed dollar amount (such as \$1,000/\$2,000 for an individual/family, indexed to inflation) in excess of current health expenses. Two considerations support this shift.

First, recall that the back-of-the-envelope calculations above suggest that lowering the minimum deductible requirement is an important policy for increasing the take-up rate. For most consumers, the minimum deductible requirement is simply too high, given the magnitude of the existing tax preference for employer-provided insurance. Allowing people to choose their policy deductible will solve this problem. Just as the take-up of managed care had beneficial spillovers to fee-for-service insurance (for example, Laurence Baker 1997), the take-up of HSAs and insurance plans with more cost sharing will as well.

Second, evidence from the RAND Experiment suggests that most of the expenditure-reducing effects of policy deductibles occur at low levels of deductibles (for example, Emmett Keeler et al. 1988). Extending deductibility to out-of-pocket expenses above the policy deductible will provide an important incentive for individuals to increase coinsurance rates as well. The results from the RAND experiment suggest that a mix of higher deductibles and coinsurance rates would achieve greater efficiency in health spending than mandating that all of the savings be channeled through the deductible.

## 5. Conclusion

The U.S. health care system, the envy of the world in innovation, faces criticisms from policymakers about the cost of care. From an economic perspective, an alternative approach is to ask whether private consumers of health care—and taxpayers who fund public programs—are obtaining the highest “value” for the resources devoted to health care. Healthy, competitive markets generally offer the greatest opportunity to maximize value.

As academic researchers have long observed, limiting or revoking altogether the tax preference for health insurance would improve the performance of markets for health services on this dimension. Current policy generally allows individuals to receive employer-provided health insurance expenditures tax-free, but requires direct out-of-pocket medical spending to be financed from after-tax income. This tax preference has given consumers the incentive to purchase health care through low-deductible, low-copayment insurance instead of out-of-pocket. However, likely because the vast majority of voters benefit from this preference, policymakers over the past 30 years have instead sought to level the tax playing field by expanding the tax preference rather than eliminating it.

In this paper, we show that extending deductibility to out-of-pocket spending, while a second-best policy change, is nonetheless likely to lead to significant improvements in efficiency under a range of assumptions about demand for health care and health insurance. Although we are not the first to recognize this fact, we quantify the actual health spending and revenue effects of such a policy using a transparent accounting model and a small number of behavioral parameters from existing studies.

Providing additional evidence on the sensitivity of health insurance contracts to tax changes is a subject for future research. Also, while not emphasized here, expanding deductibility may also significantly reduce rates of uninsurance by lowering the cost of health insurance. Finally, we view as an important topic for future work more analysis of the relationship between tax deductibility and Health Savings Accounts.

## Notes

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1. Such a pattern likely reflects politics as much as economics: the vast majority of voters benefit from the excludability of health insurance. When President Reagan expressed interest in eliminating or even limiting the exclusion, his proposal was soundly rejected in Congress. Indeed, the Clinton health reform plan explicitly rejected any such limitations (Cutler 1994).
2. An FSA allows employees to allocate a portion of their compensation to nontaxable fringe benefits instead of taxable wages. Currently, once the amount of the FSA contribution has been designated, the employee is not allowed to change it or drop the plan during the year unless he or she experiences a change of family status. By law, the employee forfeits any unspent funds in the account at the end of the year.
3. HRAs, however, are owned by the employer and contributions to them are subject to nondiscrimination rules; that is, they can not be at the employee's discretion. See U.S. Department of Labor (2003).
4. For example,  $t_1 = 1$  and  $t_2 = 0.7$ .
5. Because the tax preference leads to large changes in effective prices for health services, point elasticities expressed in (current) after-tax terms will be very different from those expressed in (counterfactual) pre-tax terms. For example, the effect of a 1 percent increase in the effective coinsurance rate from its current (lower) base is much smaller than the effect of a 1 percent decrease in the effective coinsurance rate from its (higher) base in the absence of the tax preference. Some of the studies we review provide estimates in the former terms; some provide estimates in the latter. We follow the convention used in the RAND study and convert all elasticities into arc terms, expressed at the average between pre- and post-tax prices.
6. The published estimates in Eichner (1997, table 1) are based on models that assume that consumers make marginal health spending decisions throughout the year based on the coinsurance rate that they face at the end of the year. This assumption is important because many plans' coinsurance rates vary with a consumer's level of cumulative spending over a calendar year. For example, a plan may have a \$500 deductible (i.e., a coinsurance rate of 100 percent on the first \$500 of spending), a coinsurance rate of 25 percent, and a \$2,000 out-of-pocket maximum (i.e., a coinsurance rate of 0 percent after \$1,500 in coinsurance payments or \$6,500 in total spending). If consumers have rational expectations, then this assumption is correct. Regardless of when in the year the choice to make a (marginal) health expenditure arises, the effective coinsurance rate for any marginal expenditure would be the rate in effect after all of the year's expenditures had occurred. In the text of the article, Eichner points out that estimates of  $e(p)$  from models that do not assume rational expectations are generally lower. However, in our view, the rational expectations assumption is more justifiable than the alternatives, so we use the estimates from the table.
7. We discuss how we assess the magnitude of the average marginal tax rate in more detail below.
8. Gruber's estimate implies  $e(p) = -0.87 = 0.392 / (1.9 * 0.19 * 2 / (0.7 + 0.89))$ , which implies an effect of  $-0.118 = -0.87 * 0.9 * 2 * 0.14 / (1 + 0.86)$ .



9. As Gruber (2000) points out, empirical evidence supports the hypothesis that the costs of health insurance premiums are fully shifted out of wages.

10. Table 3.3 presents percentage effects measured at the average  $\eta$ , so the percentage effect measured at current levels

$$\eta^* = \frac{2\eta}{2-\eta},$$

because

$$\eta = \frac{\eta^*}{[1+(1+\eta^*)]^{1/2}}$$

11. This summary is taken from the detailed explanation of the tax treatment of HSAs in Internal Revenue Service (2004).

12. See National Conference of State Legislatures (2006).

13. Of course, adoption of full deductibility does not preclude HSAs. Indeed, full deductibility enhances the incentive to finance current health spending out of pocket, while HSAs (when used as a savings vehicle) enhance the incentive to accumulate assets to finance future health expenses out of pocket.

14. See Gabel and Rice (2003).

15. We reach this conclusion by using parameters from the RAND Health Insurance experiment, the approach suggested by Phelps (2003), and the increase in health spending reported by CMS (2006). In 1984 dollars, the current deductible of \$221 would be equivalent to \$44. According to CMS (2006), spending per private health insurance enrollee rose from \$675 in 1984 to \$3,379 in 2004, a factor of five. According to Phelps (2003, table 5.6), to achieve the spending reduction of 2.7 percent from full deductibility predicted by the RAND Experiment, deductibles would have to have risen to \$58 in 1984, or \$290 (=58\*5) in 2004.

16. To achieve the spending reduction of 13.4 percent from revoking the income and payroll tax preferences predicted by the RAND Experiment, deductibles would have to have risen to \$336 in 1984, or \$1,680 (=336\*5) in 2004.

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