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LOCKED IN BY LEVERAGE:
JOB SEARCH DURING THE HOUSING CRISIS

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Locked in by Leverage: Job Search during the Housing Crisis
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ABSTRACT

This paper examines how housing market distress affects job search. Using data from a leading online job search platform during the Great Recession, we find that job seekers in areas with depressed housing markets apply for fewer jobs that require relocation. With their search constrained geographically, job seekers broaden their search to lower level positions nearby. These effects are stronger for job seekers with recourse mortgages, which we confirm using spatial regression discontinuity analysis. Our findings suggest that housing market distress distorts labor market outcomes by impeding household mobility.

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Over the five years following their peak in 2006, U.S. home values fell by more than 22%. The decline in home values, combined with record-high household leverage, evaporated households' housing wealth and set off a chain of events that pushed the U.S. economy into a severe recession that reduced labor demand. During the recession, which started in December 2007, employers laid off more than 3.6 million workers and cut job openings by 44% (Bureau of Labor Statistics 2009, 2012). Much less is known, however, about how the reduction in housing wealth affected labor supply, including the location and types of jobs that individuals were willing to take.

A popular idea is that the housing bust affected labor supply by reducing household mobility. Liquidity-constrained households might not have the cash required to sell a home with negative equity. Even when a homeowner is not underwater, diminished home equity reduces the amount of capital available to finance a down payment for a new home, which also "locks in" workers to their current location (Stein 1995; Genesove and Mayer 1997). Home price declines might further reduce mobility if households are sensitive to nominal loss aversion (Genesove and Mayer 2001; Engelhardt 2003). Although households could avoid selling by becoming landlords and renting elsewhere, preferences for homeownership and rental market frictions make this an imperfect substitute. As a result, workers facing home price declines might be reluctant to apply to positions that require relocation.¹

The housing bust could also affect labor supply through a wealth effect. When unemployed workers have lower home equity and net worth, liquidity constraints can make them more willing to accept lower level positions (Chetty 2008; Herkenhoff and Ohanian 2015;

¹ At the same time, house value declines could increase mobility if foreclosures force defaulting homeowners out of their homes (Molloy and Shan 2013). We estimate the net effect. Even without price changes, transactions costs of transferring homeownership could interfere with the labor market by reducing workers' mobility (McCormick 1983; Head and Lloyd-Ellis 2012; Blanchflower and Oswald 2013).

Herkenhoff, Phillips, and Cohen-Cole 2016). Consequently, households facing price declines might increase applications to lower level jobs and those requiring only limited experience.²

Little is known about the housing bust's effect on labor supply, in part, because it is challenging to separately identify the effect on labor supply from the effect on labor demand. Through an aggregate demand channel, falling home values and tightening credit markets suppress consumer spending, leading to a drop in consumer demand, firm production, and labor demand (Eggertsson and Krugman 2012; Mian and Sufi 2011, 2014; Guerrieri and Lorenzoni 2015). With data only on employment or wages, it is impossible to separate the supply and demand channels empirically. This paper exploits data from a large online job search platform to overcome this identification challenge.³ With microdata on nearly four million job applications across the U.S. between May 2008 and December 2009, we hold demand fixed and examine how labor supply to specific jobs is affected by job seekers' housing market conditions.

During the Great Recession, house values and mortgage leverage varied substantially both across locations and over time. Our first identification strategy exploits the fact that local labor markets encompass many hyperlocal housing markets. Job seekers within a labor market have access to the same job postings, but experience different local housing market developments. Because the job application data contain no information on individual job seeker's homeownership or housing wealth, we examine job applications at the ZIP code-month level. For example, within a given local labor market, we compare the changes in job search strategies of applicants in ZIP codes that experienced relatively stable home prices with applicants in nearby ZIP codes that experienced larger price declines.

² Debt overhang after house price declines could also lead job seekers to be choosier (Donaldson, Piacentino, and Thakor 2015; Bernstein 2016). Again, we estimate the net effect.

³ We were provided the data under a nondisclosure agreement that restricts us from identifying the online platform. This agreement places no constraints on the conclusions of the analysis.

We find that home value declines and limited home equity lead job seekers to focus their searches on jobs within commuting distance of their home. A 30% decline in home value is associated with a 15% decrease in the fraction of applications submitted to jobs out of the applicants' commuting zone. Highly indebted households appear to be particularly tied to their homes: underwater homeowners apply to only half as many distant positions as homeowners with abundant home equity.

With their search constrained geographically, job seekers broaden their search in other ways. They apply for more positions inside their commuting zone by expanding the types of jobs to which they apply. Using job codes assigned by the online platform, we measure the concentration of applicants' job search using a Herfindahl-Hirschman index (HHI) and find that home price declines are associated with less focused local job search. These job codes and job requirements also provide insight into the changing nature of the search. After home value declines, job seekers tilt their search for nearby jobs toward lower level positions. They increase applications to jobs that require little experience, decreasing the fraction of applications to management and executive positions. Consistent with these changes being motivated by the cost of relocation, the effects do not extend to search outside of the commuting zone or hold for ZIP codes dominated by renters.

These effects on job search are particularly pronounced in states with recourse mortgages. Some U.S. states prohibit lenders from pursuing a homeowner's other assets if he or she defaults and the foreclosure sale does not cover the outstanding debt. Because recourse is costly for defaulting homeowners, it makes sense for job seekers with recourse loans to be more locked-in to their homes after a housing market crash than job seekers with nonrecourse loans. Consequently, housing considerations affect job search the most in recourse states.

Although our first empirical approach is designed to control for the greatest identification challenge (recession-induced variation in labor demand), that approach is susceptible to a subtle selection bias among job seekers. If less mobile individuals within a ZIP code, such as those with less education or those with school-age children, are more likely to search for jobs during housing downturns, then the changes in job search that we estimate could reflect differences in the pool of applicants as opposed to differences in housing-related incentives. To address this potential concern, we exploit the discrete change in recourse laws at state borders.

Using a spatial regression discontinuity framework, we compare ZIP codes near a common state border where mortgage recourse changes. These applicants face the same job opportunities and local economic conditions, yet their job search strategies change discontinuously at the border. Job seekers on the side of the border with recourse mortgages pursue fewer jobs that would require relocation and apply to a broader range and lower level of positions, relative to their counterparts immediately across the border in a nonrecourse state. These effects are strongest in high-income ZIP codes experiencing significant home value declines, where we would expect homeowners to have other valuable assets and be most sensitive to recourse.

Together, the results of the panel and regression discontinuity analyses suggest that housing market fluctuations distort job search. These results complement Bernstein (2016), who finds that some homeowners strategically reduce labor supply when presented with the opportunity for an income contingent mortgage modification. Our analysis shows that nonstrategic motivations also affect distressed homeowners' job search by shifting the location and type of jobs sought. We find that distressed homeowners increase labor supply to some jobs while decreasing labor supply to others.

By impeding efficient matching in the labor market, housing market frictions likely reduce aggregate productivity and contribute to the Yagan (2016) finding that individuals' location at the onset of the Great Recession affected their labor market participation years later. When workers do not apply to (and, therefore, are not hired for) jobs where their skills are most valued, they forego opportunities to build their on-the-job experience and general human capital. Constrained job search that leads to lower short-term and long-term earnings could exacerbate households' financial distress and reinforce workers' reduced mobility. Firms may also suffer if they are unable to attract workers from the national labor market, potentially preventing them from hiring workers with appropriate skills. More generally, if migration facilitates knowledge transfer between regions, then job seekers who are reluctant to relocate impose an externality on productivity (Serafinelli 2012).

We find that job seekers act as if declines in home value and negative home equity reduce their ability to relocate. Although Ferreira, Gyourko, and Tracy (2010, 2012) find that negative equity reduces property sales, it might not reduce households' actual mobility if they rent out the property or suffer a foreclosure-related eviction (Schulhofer-Wohl 2011). Using administrative data from the Netherlands, Struyven (2014) finds that negative equity reduced Dutch owner-occupied mobility during the Great Recession by 20%–25%.⁴ U.S. data are less complete, and the results are mixed. Kothari, Saporta-Eksten, and Yu (2013) and Modestino and Dennett (2013) find that the housing bust affected mobility in the U.S., but other studies find no or weak evidence of housing-related mobility declines during the same period (Aaronson and Davis 2011; Donovan and Schnure 2011; Molloy, Smith, and Wozniak 2011; Schmitt and Warner 2011; Farber 2012; Coulson and Grieco 2013; Mumford and Schultz 2013; Bucks and Bricker 2016;

⁴ Henley (1998) and Chan (2001) find qualitatively similar results in the UK and U.S. following home value declines in the 1990s.

Demyanyk, Hryshko, Luengo-Prado, and Sorensen 2016). Valletta (2013) finds no effect on unemployment durations. Our results highlight the importance of recourse, which could explain why housing distress would affect mobility more in the Netherlands than in the U.S. We are not aware of any other empirical work linking the housing market and mobility to workers' actual labor supply decisions.

It is possible that realized mobility (i.e., actual moves) provides only a coarse measure of housing "lock in" during a deep recession, when job finding rates are low irrespective of search intensity. Because the recession was geographically widespread, mobility (i.e., the ability to move) might not have been particularly beneficial for the unemployed. Indeed, the realized mobility of unemployed renters declined substantially during the Great Recession (Farber 2012), even though we find that renters did not significantly reduce applications to jobs outside of their commuting zone. When unable to find a job, even many mobile job seekers will not move. Through this lens, our results are consistent with researchers' conclusions that the persistent unemployment following the Great Recession was more likely explained by reduced aggregate demand than a structural (geographic) mismatch (e.g., Farber 2012; Kothari, Saporta-Eksten, and Yu 2013).

The remainder of the paper is organized as follows. We describe the data in Section 1 and results from the panel of ZIP codes in Section 2. Section 3 reports results from the spatial regression discontinuity analysis of mortgage recourse at state borders, and Section 4 concludes.

1. Data

1.1. Job applications

To explore the impact of housing market distress on labor supply, we examine the

relationship between job seekers' applications and their local housing market conditions. The job application data come from a large online job search platform. The platform allows firms to post job listings and job seekers to apply to these positions. The platform earns revenues from companies posting positions and from advertising; job seekers use the platform for free. All job listings include the job's location and some list experience requirements. Unfortunately, few listings include wages. Job seekers can browse job categories—filtering by location and job characteristics—or search the platform using key words. For most applicants, we only observe his or her ZIP code.

As reported in Table 1, Panel A, the sample includes data on approximately 4 million job applications to almost 60,000 jobs posted in the financial services industry between May 2008 and December 2009.⁵ The job search platform has declined to provide us with data from other industries or time periods. The postings include a wide range of positions, including jobs in retail branches (e.g., tellers, account executives, and financial advisors) and back-office jobs (e.g., telephone bankers, financial analysts, software engineers, and administrative assistants) and are spread across all fifty U.S. states roughly in proportion to population. Applications are also widely dispersed, coming from job seekers in 12,157 ZIP codes. To match the housing market data, we aggregate applications to the ZIP code-month level for our analysis.

Panel B of Table 1 reports characteristics of the applications. To classify the geographic breadth of applicants' search, we match each job applicant and posting to one of 709 commuting zones, defined in 2000 by the U.S. Economic Research Service. Commuting zones delineate local economies, not political boundaries, and approximate local labor markets (Autor and Dorn 2009, 2013). We scale the volume of applications using the ZIP code's number of finance

⁵ The data are a subsample of the more than 5.5 million applications used by Brown and Matsa (2016) to study labor supply to financially distressed firms. The current analysis examines all applications that include the job seekers' ZIP code for which we observe housing market information.

workers (from the American Community Survey, 2008–2012). Job seekers apply to nearby positions more frequently than they apply to positions located farther away. About three-quarters of applications are to jobs inside the applicant’s commuting zone. In an average month, we observe 18.1 applications per 1,000 finance workers inside the commuting zone and 7.2 applications outside.

We also characterize the breadth of job search using a classification provided by the online platform. Posted jobs are assigned up to four of nineteen job codes, such as administrative and clerical, sales, customer service, management, executive, and entry level. We account for heterogeneity in the positions by considering each combination of job codes as a unique job type. By our job type classification, for example, a customer service management position is distinct from a customer service entry-level position. We measure the concentration of applicants’ search using a Herfindahl-Hirschman index (HHI), which is the sum of the squared market shares of the job types in a ZIP code-month. Job search concentration, which can range from zero to 10,000, averages about 2,500 and 3,000 inside and outside of the commuting zone, respectively.

Postings’ job codes and experience requirements also indicate a job’s level of seniority. Six to eight percent of applications are to management positions, less than one percent are to executive positions, and two to three percent are to entry-level positions. About 11% of job postings (8.5% of the total applications) specify the years of experience required for the position. When applying to jobs in the applicant’s commuting zone that specify experience requirements, 15.2% of applications are to jobs requiring less than one year of experience. This percentage is lower, 10.9%, when applying outside the commuting zone.

1.2. Housing and labor market conditions

Housing and labor market conditions are summarized in Panel C of Table 1. To measure housing market conditions, we use monthly estimates of median home values at the ZIP code-level from Zillow to calculate how much home value declined in each ZIP code since January 2006, when home values peaked nationally. Large, positive values reflect large price drops. Between May 2008 and December 2009 (our sample period), home values had fallen an average of 9.4% from their peak.

We estimate the impact of these price movements on homeowners' home equity using loan servicing information from CoreLogic's Loan-Level Market Analytics database. For purchase and refinance loans, we estimate the property value monthly by inflating (or deflating) the appraised value at origination using the Zillow ZIP code home value index. To filter likely data entry errors, we drop loans in the 1% tails of loan-to-value at origination. We also exclude loans on commercial property, mixed-use property, residential property with five or more units, manufactured housing, and property for which type is missing. On average in the sample period, 32.3% of mortgagors had loan-to-value ratios greater than 80%, and 10.3% owed more than their home was worth.

We also measure local labor market conditions in each month. We measure labor supply in the commuting zone using the unemployment rate (constructed as the population-weighted average of county-level data from the U.S. Bureau of Labor Statistics) and measure labor demand in the commuting zone using the number of finance jobs posted (in our application data) divided by the total number of finance workers in March 2008 (U.S. Census Bureau's County Business Patterns). On average, the unemployment rate is 7.5%, and there are 5.6 job postings per 1,000 workers.

2. Panel analysis of job applications and home values

2.1. Geographic scope

By reducing the funds available to finance the down payment for a new home (Stein 1995) or requiring a deficiency payment from the owner in order to sell the home, home value declines could affect homeowners' job search behavior by reducing their mobility. Figure 1 presents a binned scatterplot of the percent of out-of-commuting zone applications against home value declines, along with a smoothed kernel-weighted local polynomial and its 95% confidence interval. The pattern in the figure is striking: although nearly 30% of applications are to jobs out of the local labor market when homes have appreciated, these applications fall sharply when home values declines. Relative to places where prices remained flat, ZIP codes in which median home values declined by 20 percentage points are associated with 10 percentage points fewer applications submitted to jobs outside of the local labor market—about a one-third decrease.

These results are robust to controlling for detailed fixed effects and labor market conditions. Table 2 presents results from ZIP code-month regressions, where standard errors are adjusted for clustering at the ZIP code level to account for correlation in unobserved determinants of job seekers' behavior in a ZIP code over time. In a specification without controls, the estimated coefficient on home value decline is negative and statistically significant—a 30% decrease in home value is associated with a 6.4 percentage point decline in applications to jobs out of the commuting zone, which is a 27% drop relative to the mean (column 1; $p < 0.01$). Month fixed effects, which account for changing aggregate economic conditions, have little effect on the estimates (column 2). The specifications in columns 3 and 4 account for geography- and labor market-specific differences using 305 MSA or 353 commuting

zone fixed effects. The estimated coefficient on the home value decline, albeit smaller in magnitude, remains negative and highly statistically significant.

Commuting zone fixed effects control only for fixed differences between local labor markets. To control also for local developments in these markets over time, we include time-varying proxies for labor supply and labor demand in the commuting zone: the local unemployment rate and the local job posting-to-worker ratio. As reported in column 5, the estimated coefficients on these controls have the expected signs. More local unemployment is associated with more applications out of the commuting zone, suggesting that applicants expand their search geographically when the local market is competitive. A scarcity of open positions relative to the local labor force also encourages workers to consider moving to a different market where they might be more likely to find a job. Including these controls for local labor market conditions, however, has little effect on the coefficient on home value decline. A specification with 691 three-digit ZIP code fixed effects accounts for even finer geographic heterogeneity and also yields negative and statistically significant results (column 6; $p < 0.01$).

The most demanding specification, presented in column 7, includes ZIP code fixed effects in addition to the local labor market controls and month fixed effects. In this specification, the relationship between home value and job seekers' application behavior is identified from changes in home value within the 12,157 individual ZIP codes over time and after controlling for both aggregate temporal patterns and developments in the local labor market. The estimated coefficient on the home value decline suggests that a 30% drop in value is associated with a 4.7 percentage point decrease in out-of-commuting zone applications, which is a nearly 20% decrease relative to the mean ($p < 0.01$).

This decline in the percent of applications to jobs outside of the commuting zone reflects

both a reduction in applications to distant jobs and an increase in applications to nearby ones. In Table 3, we report the effects on the application volume separately for jobs inside and outside the commuting zone. For each type of application, we examine an indicator for whether any applications were submitted in a ZIP code-month (extensive margin) and the log of the monthly number of applications submitted per worker residing in the ZIP code (intensive margin). Similar to column 7 of Table 2, these regressions control for ZIP code and month fixed effects and the local labor market controls.

A depressed home value is associated with decreased applications to jobs outside of the commuting zone on both the extensive and intensive margins. A 30% drop in value is associated with a 3.2 percentage point decrease in any distant applications (column 3; $p < 0.01$) and, when there are distant applications, a 12.0 log point drop in the number (column 4; $p < 0.01$). These results reinforce those in Table 2 and suggest that applicants facing housing distress shift their job search away from positions that would require relocation.

2.2. Concentration and job type

Housing market distress and lock-in also affect the *types* of jobs to which job seekers apply. With less ability to pursue positions farther away, job seekers with diminished home equity expand their job search inside their commuting zone and apply to positions they would not have considered otherwise. As reported in Table 3, a 30% drop in home value is associated with a 1.8 percentage point increase in any local applications (column 1; $p < 0.01$) and, when there are local applications, a 6.0 log point increase in their number (column 2; $p < 0.01$).

As these job seekers expand their search locally, they also broaden the types of positions to which they apply. To study the breadth of applicants' search, we examine the concentration

measure described in Section 1.1. A larger value reflects a more focused job search. Table 4 presents regressions of the concentration measure on home value decline and controls for month fixed effects and labor market conditions. Columns 1–4, which include increasingly demanding geographic controls, examine the concentration of applications to jobs inside the commuting zone.

After home values collapse, searches for nearby jobs become less focused. After controlling for detailed geographic fixed effects at the three- or five-digit ZIP code levels, home value declines are associated with less concentrated search for a position within the commuting zone.⁶ A 30% decline in home value is associated with a 100–150 unit decrease in the concentration measure, which is a 4%–6% change relative to its mean (columns 3 and 4; $p < 0.01$). In contrast, as we would expect, home value declines do not decrease search concentration among positions that require relocation. When examining applications to positions outside the commuting zone, the coefficient on home value decline is positive, relatively small, and not statistically significant (column 5).

To explore the change in job search in greater detail, we examine job postings' type codes and experience requirements. In analysis reported in Table 5, we relate home value decline to the percent of applications submitted to management positions, executive positions, entry-level positions, and positions requiring less than one year of experience. The analysis in Panel A examines applications to nearby jobs, whereas the analysis in Panel B examines applications to distant ones. Similar to column 7 of Table 2, all specifications control for ZIP code and month fixed effects and the local labor market controls.

⁶ Estimates from specifications with less detailed geographic controls, reported in columns 1 and 2, are small and not statistically significant. This pattern, which is consistent with depressed ZIP codes having more focused searches before prices fall, underlines the importance of accounting for geographic differences across ZIP codes in this analysis.

As job seekers in distressed housing markets broaden their search, they become less ambitious. When home values decrease, job seekers shift their local search away from management and executive positions. A 30% decline in home values is associated with 13% decrease in the fraction of local applications submitted to management positions (column 1; $p < 0.01$) and a 65% decrease in the fraction of applications to local executive job postings (column 2; $p < 0.01$), relative to their respective means. The opposite may be true for entry-level positions, to which applications increase by 8%, although this estimate is not statistically significant (column 3). Consistent with this interpretation, job seekers increase applications to nearby positions requiring little work experience. A 30% decrease in home values is associated with a 14 percentage point increase in applications to jobs requiring less than a year of experience, almost doubling applications to those positions (column 4; $p < 0.01$).

In contrast, distressed job seekers do not downgrade their search when applying to positions that require relocation. When examining applications to positions outside of the commuting zone, a 30% decrease in home values is associated with a 46% decrease in applications to entry-level jobs, relative to the mean (column 7; $p < 0.01$). Although this effect is small in absolute terms, it is consistent with distressed job seekers being unwilling to bear the elevated costs of relocation for a low-level position. Although home value decreases are not significantly related to the share of applications submitted to management positions (column 5), executive positions (column 6), or positions requiring less than a year of work experience (column 8), the point estimates for management positions and experience are consistent with the entry-level results.

2.3. Heterogeneity in effect

2.3.1. Renters

Housing market distress directly affects job seekers who own their home. We expect renters to be much less sensitive to changes in local home prices, at least in partial equilibrium, and therefore examine their job search as a falsification exercise. Although we cannot identify individual applicants' home ownership status, the homeownership rate in their ZIP code provides a proxy for the likelihood the applicant is constrained by housing debt. To limit the sample to mostly renters, we examine ZIP codes with homeownership rates below 25% in the 2000 U.S. Census. We repeat the previous analysis on this sample and report the results in Table 6.

As we would expect, renters' job search patterns are less affected by the housing market's collapse. In ZIP codes dominated by renters, a 30% decline in home values is associated with only 0.8 percentage points fewer applications out of the commuting zone (column 1). The estimate for the breadth of job search is also smaller (column 2). Neither estimate is statistically significant. These estimates are consistent with homeowners' housing market distress explaining the significant effects estimated in our main analytic sample.

2.3.2. Home equity

Although loss averse homeowners would hesitate to sell even an unmortgaged home after its value declines, we expect home price declines to have the greatest impact on highly indebted mortgagors. Many households considering relocation rely on their current home equity to fund a down payment on a new home. Thus, the less home equity available, the harder it is to move. In addition to lacking equity to put toward a down payment, households holding underwater mortgages must find cash to cover the shortfall to sell their home.

We examine the role of mortgage debt in shifting application behavior in analysis reported in Table 7. As expected, low or negative home equity is associated with fewer applications out of the commuting zone (column 1) and less concentrated local job search (column 2). Assuming that only the applications of high loan-to-value mortgagors are affected, the estimates imply that being underwater leads a homeowner to decrease applications out of his or her local labor market by 12.6 percentage points, a more than 50% decline relative to the mean. The effect is half as big for a household whose equity position is positive but less than 20% of the property value: these households decrease applications out of the commuting zone by 7.1 percentage points, an almost 30% decline. The pattern is reversed for job search concentration, which decreases by over 50% for households with low positive home equity and by 13% for underwater homeowners. Home equity thus appears to be an important mechanism through which housing market distress affects homeowners' job search strategies.

2.3.3. Recourse law

Throughout the U.S., residential mortgage loans are secured by the property. If the borrower defaults, the lender can seize and sell the property to collect on the debt. Homeowners are also personally liable for their mortgage debt in most U.S. states. In these "recourse" states, lenders can pursue other assets from the borrower to collect on any debt not covered by the foreclosure sale. In states without recourse, however, borrowers can default on underwater mortgages without any additional financial liability, because lenders cannot pursue borrowers' other assets to recover the remaining debt. Although not all residents may be aware of a state's debtor protection laws, potential homebuyers and sellers are likely to be informed about recourse by their realtors and lenders. Indeed, personal liability for mortgage debt has real effects.

Borrowers are less likely to default on recourse mortgages, even when they have substantial negative home equity (Jones 1993; Bhutta, Dokko and Shan 2010; Ghent and Kudlyak 2011).

If housing market conditions affect job seekers' labor supply decisions, then we would expect to find different patterns in recourse and nonrecourse states. Figure 2 displays the nine states with nonrecourse mortgages, based on the classification by Ghent and Kudlyak (2011).⁷ Because recourse is costly for defaulting homeowners, we expect job seekers with recourse loans to be more locked-in to their homes after home values crash than job seekers with nonrecourse loans. As a result, job seekers facing diminished home values in recourse states are likely to engage in more geographically constrained search than other job seekers.

To explore this possibility, we interact home value declines with a recourse indicator in our regression framework and control for local labor market conditions, ZIP code fixed effects, and separate month fixed effects in recourse and nonrecourse states. Because recourse varies only at the state level, its main effect is absorbed by the ZIP code fixed effects. The results are reported in Table 8.

The geographic breadth of job search is most sensitive to home value declines in recourse states, where defaulting is particularly costly. Housing market conditions have little effect on the geographic breadth of job search in states with nonrecourse mortgages: the coefficient estimate on the uninteracted home value measure is small and not statistically significant. However, the coefficient estimate on the interaction is sizeable and statistically significant, suggesting that job search in recourse states is much more sensitive to housing market conditions. Relative to nonrecourse states, a 30% decline in home values in a recourse state is associated with 5.1 percentage points fewer applications out of the commuting zone, which is more than a fifth of

⁷ California and North Carolina allow a mix of recourse and nonrecourse mortgages. Because we do not observe individual loans, we classify these states as recourse, but the results are similar if we classify them as nonrecourse or drop these states altogether.

the sample mean (column 1; $p < 0.01$).

In-commuting zone job search concentration also responds to housing market distress differently in recourse and nonrecourse states. Whereas home price declines are associated with more focused job search in nonrecourse states, the declines are associated with less focused search in recourse areas, where job seekers are more locked-in to their homes. Relative to nonrecourse states, a 30% decline in home values in a recourse state is associated with a 310 unit increase in search concentration, which is about 12% of the sample mean (column 2; $p < 0.01$).

3. Regression discontinuity analysis

Although our panel analysis finds a robust relationship between home value and the breadth of job search in both geography and type, one might wonder whether the results reflect differences in the type of workers seeking jobs during economic downturns instead of changes in a consistent type of workers' job search. For example, less educated workers are both less mobile (Machin, Salvanes, and Pelkonen 2012) and more exposed to job loss during recessions (Hoynes, Miller, and Schaller 2012). Thus one might ask if the shift in applications toward jobs inside the commuting zone could be because less educated workers are searching, not because of the changing housing market conditions. To address this possibility, we employ a second identification strategy that compares the applications of job seekers, in a narrow region near a state border, who face the same local economic conditions but different housing market-related incentives because of different laws regarding mortgage recourse on either side of the border.

To exploit the discontinuity in recourse regimes at state borders, we restrict the sample to ZIP codes within 50 miles of borders at which recourse laws change. Figure 2 marks the centroids of these ZIP codes on a map of the U.S., using red circles in recourse states and blue

Xs in nonrecourse ones. This cross-state variation in debtor protection laws results from historical factors, such as foreclosure rates on farms during the Great Depression, and is unrelated to recent aggregate shocks or these narrow regions of the states that we analyze (Ghent 2014; Dobbie and Goldsmith-Pinkham 2015). By comparing job seekers across these neighboring jurisdictions with different recourse regimes, we can confirm that housing market incentives influence job seekers' applications.

The ZIP codes on either side of these borders are quite similar. Table 9 summarizes demographic and economic characteristics of the two sets of ZIP codes, including age, race, education, household income, marital status, household size, population density, owner occupancy, historical mobility, and prevalence of low home equity. The ZIP codes appear very similar across these dimensions. Table 9 also reports the p -value of the difference in means, adjusting for clustering at the state level. There is no statistically significant difference in almost all of the ZIP code characteristics analyzed. Of the 38 variables analyzed, only two age measures (under 5 years and over 65 years) and one race measure (black) are statistically different at the 10% level. As discussed below, the results of the regression discontinuity analysis are robust to including controls for age and race.

To further ensure that we only compare ZIP codes that are in close proximity, we follow Mian, Sufi, and Trebbi (2015) and divide the borders into 10-mile strips. The resulting 10-mile by 50-mile strips of ZIP codes are perpendicular to the state border, and each strip is specific to a border pair (e.g., Minnesota-South Dakota). Throughout our analysis, we control for very local conditions by including fixed effects for these strips. As such, we compare only ZIP codes that lie in the same 10-mile strip running across the state border in question. We also interact these border strip fixed effects with commuting zone fixed effects to ensure that all comparisons are

within the same commuting zone. For conciseness, we refer to these interacted fixed effects as simply “location” fixed effects.

3.1 Geographic scope

Table 10 reports results from regressions of the percent of applications out of the commuting zone on an indicator for whether the state allows recourse mortgages, as well as month and location fixed effects. Job seekers submit 37 percentage points fewer applications to jobs outside of the commuting zone on the recourse side of the border (columns 1; $p < 0.01$). After including additional controls for the distance to the border, distance squared, and distance cubed, the estimate’s magnitude reduces to 21 percentage points but is still highly statistically significant (columns 2; $p < 0.01$).

Because recourse allows lenders to seek deficiency payments from defaulting homeowners, recourse is most salient for homeowners with other valuable assets. In contrast, borrowers without substantial wealth beyond their homes face similar repercussions regardless of recourse. As a proxy for homeowners’ wealth, we use ZIP codes’ average adjusted gross income (AGI) from the Internal Revenue Service’s Statistics of Income for September 2008. We repeat the analysis from column 2 separately on above-median and below-median AGI ZIP codes and report the results in columns 3 and 4, respectively. Recourse is associated with geographically narrower job search in high-income areas but not in low-income areas. In high-income ZIP codes, job seekers apply to 22.5 percentage point more jobs outside of the commuting zone on the recourse side of the border than on the nonrecourse side (column 3; $p < 0.01$). But there is essentially no difference in lower income areas (column 4; $p = 0.83$).

The discontinuity at the border is readily apparent in a nonparametric analysis. Figure 3

depicts the percent of applications submitted to jobs outside of the commuting zone in high-income ZIP codes near state borders where recourse changes. We regress this variable on 50 indicators for each one-mile interval on either side of the border, with negative values in states with recourse mortgages, as well as month and location fixed effects. The figure plots the coefficient estimates on the distance indicators in those regressions, along with separate smoothed kernel-weighted local polynomials on either side of the border and their 95% confidence intervals. The vertical line represents the state border. ZIP codes in states with recourse mortgages are on the left and ZIP codes without recourse are on the right.

Figure 3 reveals a sizeable jump in applications to distant positions. Job seekers in states with recourse mortgages submit substantially fewer applications to jobs outside of their commuting zone relative to their counterparts, just across the state border, who hold nonrecourse mortgages. Although the average fraction of applications to distant positions is somewhat flat as you approach the state border from either side, it jumps discontinuously at the border. Like in the panel analysis reported in Section 2, these reductions seem to reflect both a decrease in applications to distant positions and an increase in applications to positions nearby (see Appendix Figure A1 and Appendix Table A1).

Border ZIP codes in recourse and nonrecourse states are statistically indistinguishable in almost any dimension (see Table 9); however, there are small but statistically significant differences in the prevalence of child, elderly, and black residents. We suspect that these differences result from random variation. Nevertheless, to confirm that these differences cannot explain the regression discontinuity results, we augment the specification in column 3 of Table 10 with the full set of age and race variables listed in Table 9 and report the results in column 5. The jump at the border remains similar in magnitude and is statistically significant ($p < 0.05$).

We would expect recourse to be most limiting after home values decline substantially. Absent a price drop, homeowners are unlikely to be underwater or face a shortfall in default. In a final specification, we use an indicator to control for ZIP codes with above median declines in home value and interact this variable with the recourse indicator. As expected, we find that applications to distant positions are particularly sensitive to recourse eligibility in places that experience large price drops. The coefficient estimate on the interaction of recourse and large price declines is negative and statistically significant. The jump at the border in applications to jobs outside of the commuting zone is 6.1 percentage points larger in areas suffering more substantial price declines (column 6; $p < 0.05$). This interaction provides further evidence that the differences in job search behavior are tied to housing related lock-in.

3.2 Concentration and job types

Results from the regression discontinuity analysis also confirm our findings about the concentration of applications and types of jobs sought. Table 12 reports results for the concentration of local job search. Except for the dependent variable, the specifications are the same as those presented in Table 11. With their search constrained geographically, job seekers in states with recourse mortgages search more broadly for positions within commuting distance from their home. Job seekers on the recourse side of the border search with 1,000–1,900 units less concentration than job seekers on the opposite side, a 40%–75% reduction relative to the overall sample mean (columns 1 and 2; $p < 0.05$).

The effect of recourse on job search concentration comes from high-income ZIP codes, as we would expect. In high-income areas, job search concentration is 1,200 units lower on the recourse side of the border than on the nonrecourse side, a 50% reduction relative to the mean of

2,400 in high-income areas (column 3; $p < 0.01$). This pattern is not present in low-income ZIP codes, whether the difference is positive, smaller, and not statistically significant (column 4; $p = 0.37$).

Figure 4, Panel A, displays these results nonparametrically and reveals a discrete jump in local search concentration at the state border. When searching for jobs in their commuting zone, job seekers in states with recourse mortgages submit substantially less concentrated applications than their counterparts, just across the state border, who hold nonrecourse mortgages. Furthermore, the applications' average concentration is mostly flat as you approach the state border from either side, and it jumps discontinuously at the border.

We also examine the concentration of search for distant positions as a falsification test. If job seekers in states with recourse mortgages search more broadly for nearby positions because reduced mobility constrains their search geographically, then we would not expect the same effects on their search for distant positions. Figure 4, Panel B, plots the concentration of applications to positions outside of the job seeker's commuting zone. The concentration of these applications is flat over the whole region, including at the border. The regression analysis also detects no significant difference (Table 11, column 7; $p = 0.24$). This result provides further evidence of the similarity of job seekers on either side of the border in actions not affected by recourse.

Finally, we confirm that recourse reduces local job search concentration the most in areas suffering large price declines. As reported in column 6 of Table 11, the coefficient estimate on the interaction of recourse and large price declines is negative and statistically significant. The drop in the concentration of in-commuting zone applications at the border is 330 units larger in areas suffering substantial price declines (column 6; $p < 0.01$). This interaction further reinforces

the conclusion that housing distress creates lock-in that affects labor supply.

The regression discontinuity analysis also confirms that, as job seekers in distressed housing markets broaden their search, they become less ambitious. Table 12 reports effects on the percent of applications submitted to each of the job types examined in Table 5, presented separately for jobs in (Panel A) and out (Panel B) of the commuting zone. The regressions, which control for month and location fixed effects and a flexible function of distance to the border, are restricted to ZIP codes with above median income.

Job seekers on the side of state borders with recourse mortgages tilt their search toward lower level jobs nearby. When searching for positions in their commuting zone, job seekers submit 59% more applications to entry-level positions (column 3; $p < 0.10$) and 40% more applications to jobs requiring less than one year of experience (column 4; $p < 0.01$), relative to the overall sample means. As a result, the share of applications to management positions drops by 28% (column 1; $p < 0.05$). Although it is imprecisely estimated, the point estimate suggests that recourse also reduces applications to executive positions by 54% (column 2; $p = 0.43$). Appendix Figure A2 plots the discontinuities in these variables at the state borders.

As we would expect, recourse does not cause a similar tilt in applicants' search for jobs that require relocation. We find no evidence that jobs seekers with recourse mortgages apply for more entry-level or limited-experience positions when searching for positions outside of their commuting zone (columns 7 and 8). If anything, these job seekers appear to shift their search for distant jobs in the other direction, increasing applications to management positions by 35% (column 5; $p < 0.01$). Although it is again imprecisely estimated, the point estimate suggests that recourse also increases applications to executive positions by 84% (column 6; $p = 0.27$). Appendix Figure A3 plots the discontinuities in these variables at the state borders. These effects

are consistent with locked-in job seekers being willing to bear the elevated costs of relocation only for a high-level position.

4. Conclusion

We match a unique data set from a large online job search platform to housing market data to explore the relationship between home values and workers' labor supply decisions. We find that home value declines and negative equity lead job seekers to focus their searches on jobs within their commuting zone. Constrained geographically, job seekers broaden their search in other ways by considering lower level positions and those requiring less experience. These effects are more pronounced in recourse states where lenders can go after a defaulting homeowner's other assets. As a result, we find discontinuous jumps in workers' job search strategies at borders between recourse and nonrecourse states. After house values decline, it appears that job seekers lack the liquidity to fund the down payment for a new home or a protracted job search.

Housing-related constraints on job search have mixed implications for firms. When workers focus their searches on jobs nearby, firms lose access to the national labor market. At the same time, in areas where job seekers are locked in to their homes, firms face less competition from the broader labor market. These firms gain access to workers who are more qualified for a given position and, potentially, hire workers at lower cost.

From the workers' perspective, these outcomes are costly and potentially long-lived. Workers who accept positions below their skill or experience levels forego opportunities to build their human capital. Moreover, constrained workers, who apply to jobs they would otherwise have avoided, crowd out other job seekers. For example, an experienced manager who applies to

a lower level position can displace a younger worker from that job opportunity. Even if housing market constraints are short-lived, they can impose long run costs on both workers. Indeed, employment status during a recession affects labor market outcomes years later (Yagan 2016). The broader labor market impact and distributional consequences of constrained mobility are important areas for future research.

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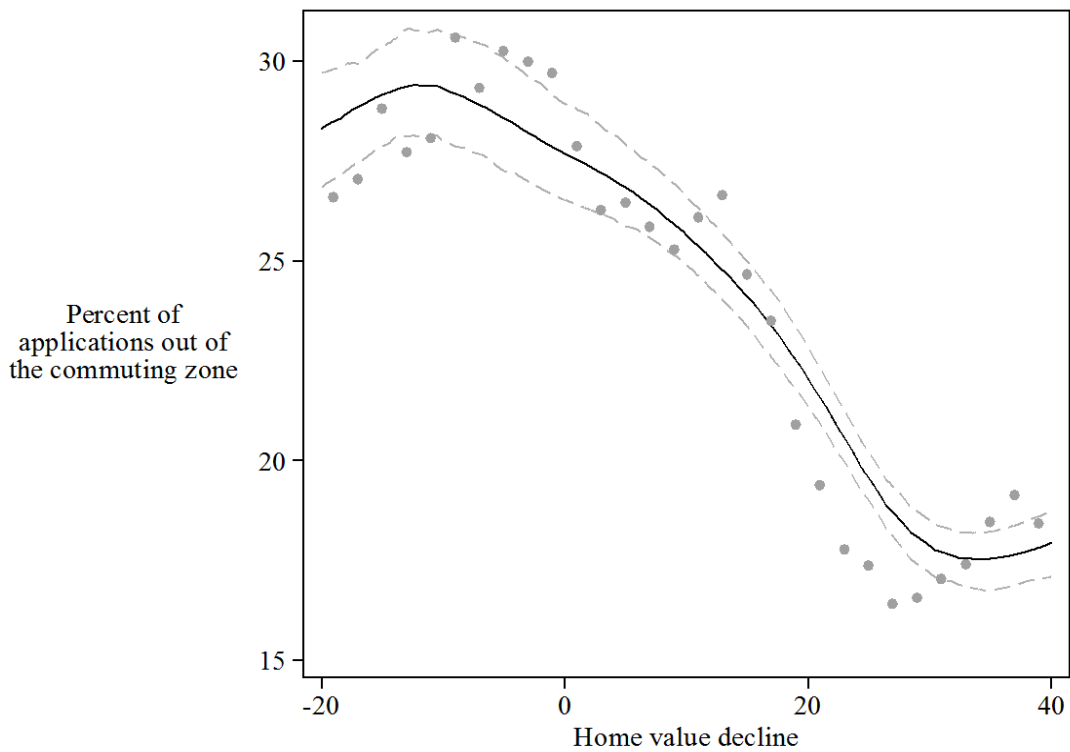
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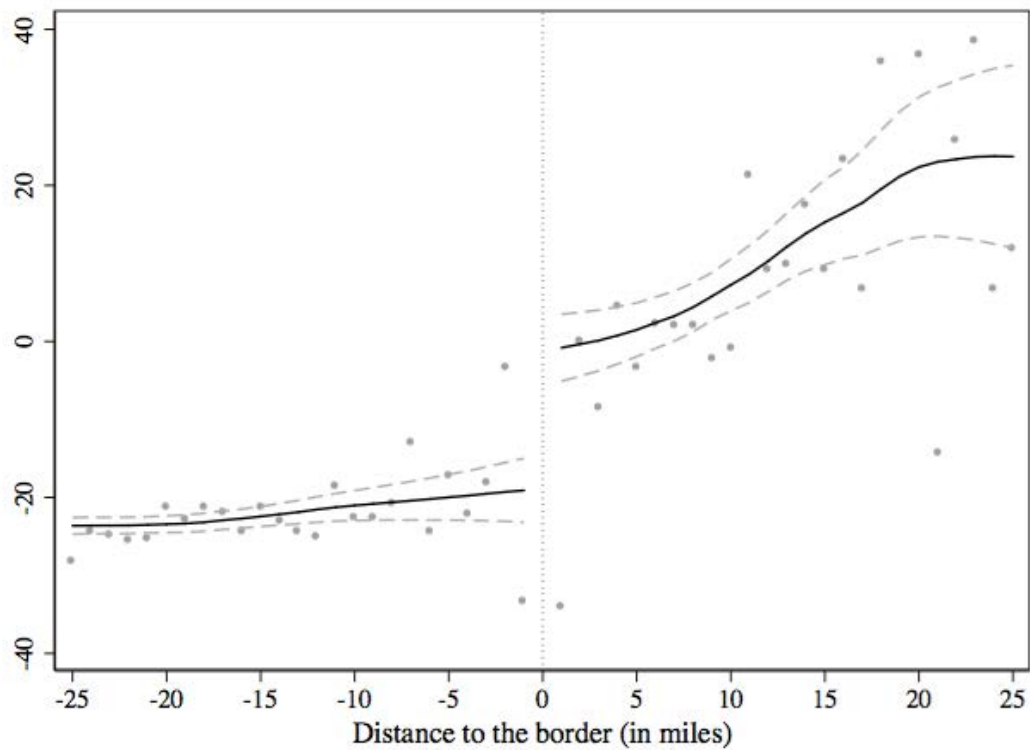
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Figure 1
Applications out of the commuting zone and home values



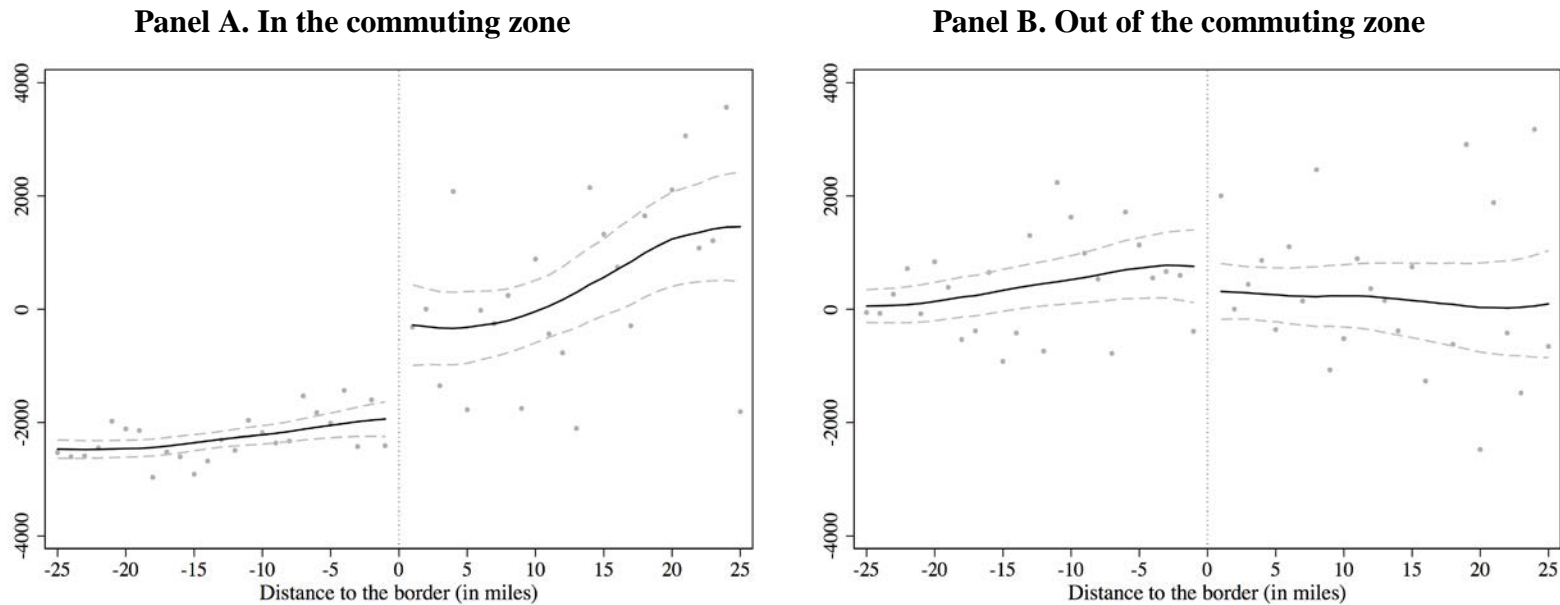
This figure presents a binned scatterplot of the percent of applications to jobs out of the commuting zone against home value decline (from January 2006, %). The sample consist of ZIP code-months between May 2008 and December 2009, and the bins are two units wide. Percent of applications out of the commuting zone is weighted by the total number of applications submitted from the ZIP code that month. The lines represent a smoothed kernel-weighted local polynomial and its 95% confidence interval, calculated by bootstrapping clustered at the ZIP code level.

Figure 3
Percent of applications to jobs out of the commuting zone



This figure plots the percent of applications to jobs out of the commuting zone for the months between May 2008 and December 2009 for U.S. ZIP codes that are near a state border where recourse law changes and had above median adjusted gross income in 2006. We regress the percent of applications on one-mile distance-to-the-border indicators, and month and location (10-mile border strip \times commuting zone) fixed effects. The distances are labeled as negative for states with recourse mortgages. The figure plots the coefficients on these dummies, along with separate smoothed kernel-weighted local polynomials on either side of the border and their 95% confidence intervals.

Figure 4
Concentration of job applications



These figures plot the concentration of applications to jobs in (Panel A) and out (Panel B) of the commuting zone for the months between May 2008 and December 2009 for U.S. ZIP codes that are near a state border where recourse law changes and had above median adjusted gross income in 2006. Variables are defined as in Table 1, and the methodology used is the same as in Figure 3. The distances are labeled as negative for states with recourse mortgages.

Table 1
Summary statistics

Panel A. Job postings				
Number of job postings	59,469			
Number of applications	3,997,972			
Number of ZIP codes	12,157			
Panel B. Application characteristics (ZIP code-month level, $N = 180,232$)				
	In commuting zone		Out of commuting zone	
	<i>Mean</i>	<i>SD</i>	<i>Mean</i>	<i>SD</i>
Quantity				
Percent of total applications	76.1	25.7	23.9	25.5
Any applications (%)	88.0	32.4	74.9	43.4
Applications per 1,000 workers	18.1	26.6	7.2	17.1
Type				
Concentration of applications	2508	2017	2939	2439
Management position (%)	5.6	10.2	8.2	13.7
Executive position (%)	0.3	2.0	0.4	3.2
Entry-level position (%)	2.6	8.2	2.3	8.3
Requires <1 year of experience	15.2	29.0	10.7	24.7
Panel C. Local economic conditions				
	<i>Mean</i>	<i>SD</i>		
Home value decline (from January 2006, %)	9.4	16.4		
Percent with low positive home equity [$0 \leq HE < 20\%$]	22.0	8.6		
Percent with negative home equity [$HE < 0$]	10.3	14.2		
Unemployment rate in commuting zone (%)	7.5	1.8		
Jobs posted per 1,000 workers in the commuting zone	5.6	3.6		

This table reports summary statistics for the application sample, which includes applications to 59,469 jobs at large financial services firms between May 2008 and December 2009. Panel A describes the job postings. Panel B summarizes application characteristics separately for jobs in and out of the applicant's commuting zone, as defined by the U.S. Economic Research Service in 2000. The percent of total applications in (and out) of the commuting zone is weighted by the number of applications submitted from the ZIP code that month. The volume of applications is scaled by the number of finance workers (in thousands) in the ZIP code, based on the American Community Survey, 2008–2012. Posted jobs are assigned up to four of nineteen job codes by the online platform, and each combination of job codes defines a unique job type. The concentration of applications is measured by a Herfindahl-Hirschman index based on the shares of these job types in a ZIP code-month. Management, executive, and entry-level positions are classified using the job codes. These job type variables are weighted by the number of applications submitted to jobs in (or out) of the commuting zone from the ZIP code that month. Required experience is weighted by the number of applications for which experience information is available (in the commuting zone, mean = 5.2, $SD = 7.5$, $N = 40,984$; out of the commuting zone, mean = 3.5 $SD = 3.1$, $N = 29,628$). Panel C describes local economic conditions. ZIP code-level home value decline since January 2006 is based on monthly estimates of median home value from Zillow. Home equity measures are estimated from CoreLogic's Loan-Level Market Analytics database. The commuting zone unemployment rate is the weighted average of county-level data from the U.S. Bureau of Labor Statistics. The number of jobs posted in the commuting zone is divided by the total number of finance workers (in thousands) in March 2008, based on the U.S. Census Bureau's County Business Patterns.

Table 2
Applications out of the commuting zone and home values

<i>Dependent variable: Percent of applications out of the commuting zone</i>							
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Home value decline	-0.212*** (0.020)	-0.232*** (0.021)	-0.138*** (0.042)	-0.107*** (0.023)	-0.115*** (0.023)	-0.249*** (0.026)	-0.155*** (0.016)
Local labor market conditions % unemployment					0.730*** (0.117)	0.623** (0.310)	0.853*** (0.111)
Posted jobs per 1,000 workers					-18.578*** (0.482)	-31.496*** (2.726)	-17.403*** (0.474)
R^2	0.02	0.02	0.27	0.57	0.59	0.54	0.72
N	180,232	180,232	180,232	180,232	180,232	180,232	180,232
Fixed effects							
Month		X	X	X	X	X	X
MSA			X				
Commuting zone				X	X		
Three-digit ZIP code						X	
ZIP code							X

This table summarizes results from regressions of the percent of applications out of the commuting zone on the ZIP code-level home price decline. Where indicated, regressions include monthly measures of local labor market conditions and fixed effects for month, metropolitan statistical area (MSA), commuting zone, three-digit ZIP code, and five-digit ZIP code. Variables are defined as in Table 1. Regressions are weighted by the number of applications submitted from the ZIP code that month. Standard errors, adjusted for clustering at the ZIP code level, are reported in parentheses. ** and *** indicate statistical significance at the 5% and 1% levels, respectively.

Table 3
Number of applications

<i>Dependent variable:</i>	In commuting zone		Out of commuting zone	
	Any applications?	Log applications per worker	Any applications?	Log applications per worker
	(1)	(2)	(3)	(4)
Home value decline	0.060*** (0.021)	0.002*** (0.001)	-0.108*** (0.029)	-0.004*** (0.001)
R^2	0.39	0.65	0.37	0.67
N	180,232	158,452	180,232	134,746
Local labor market conditions	X	X	X	X
Fixed effects				
Month	X	X	X	X
ZIP code	X	X	X	X

This table summarizes results from regressions of measures of the volume of applications on the ZIP code-level home value decline. Regressions include monthly measures of local labor market conditions and fixed effects for month and ZIP code. Variables are defined as in Table 1. Standard errors, adjusted for clustering at the ZIP code level, are reported in parentheses. *** indicates statistical significance at the 1% level.

Table 4
Breadth of job search

Dependent variable: Concentration of applications

	In commuting zone				Out of commuting zone
	(1)	(2)	(3)	(4)	(5)
Home value decline	1.03 (1.36)	1.35 (1.19)	-4.77*** (1.36)	-2.88* (1.54)	0.71 (1.83)
R^2	0.42	0.51	0.52	0.68	0.59
N	158,598	158,598	158,598	158,598	134,872
Local labor market conditions	X	X	X	X	X
Fixed effects					
Month	X	X	X	X	X
MSA	X				
Commuting zone		X			
Three-digit ZIP code			X		
ZIP code				X	X

This table summarizes results from regressions of the concentration of applications on the ZIP code-level home value decline. All regressions include monthly measures of local labor market conditions and month fixed effects. Where indicated, regressions also include fixed effects for metropolitan statistical area (MSA), commuting zone, three-digit ZIP code, and five-digit ZIP code. Variables are defined as in Table 1. Regressions are weighted by the number of applications submitted to jobs in (columns 1–4) or out (column 5) of the commuting zone from the ZIP code that month. Standard errors, adjusted for clustering at the ZIP code level, are reported in parentheses. * and *** indicate statistical significance at the 10% and 1% level, respectively.

Table 5
Types of jobs

<i>Dependent variable:</i>	Management	Executive	Entry level	Requires <1 year of experience
Panel A. In commuting zone				
	(1)	(2)	(3)	(4)
Home value decline	-0.024** (0.010)	-0.006*** (0.001)	0.007 (0.009)	0.470*** (0.107)
R^2	0.32	0.12	0.16	0.61
N	158,598	158,598	158,598	40,984
Panel B. Out of commuting zone				
	(5)	(6)	(7)	(8)
Home value decline	0.012 (0.014)	0.000 (0.003)	-0.035*** (0.013)	-0.107 (0.140)
R^2	0.19	0.12	0.15	0.47
N	134,872	134,872	134,872	29,628
Local labor market conditions	X	X	X	X
Fixed effects				
Month	X	X	X	X
ZIP code	X	X	X	X

This table summarizes results from regressions of measures of job type on the ZIP code-level home value decline. All regressions include monthly measures of local labor market conditions and fixed effects for month and ZIP code. Regressions are weighted by the number of applications submitted to jobs in (columns 1–3) or out (columns 5–7) of the commuting zone from the ZIP code that month or the number of number of applications submitted to jobs in (column 4) or out (column 8) of the commuting zone for which experience information is available. Variables are as defined in Table 1. Standard errors, adjusted for clustering at the ZIP code level, are reported in parentheses. ** and *** indicate statistical significance at the 5% and 1% level, respectively.

Table 6
Falsification test: ZIP codes with predominantly renters

<i>Dependent variable:</i>	Percent of applications out of the commuting zone	Concentration of applications in the commuting zone
	(1)	(2)
Home value decline	-0.027 (0.070)	-1.78 (6.27)
R^2	0.87	0.55
N	4,166	3,997
Local labor market conditions	X	X
Fixed effects		
Month	X	X
ZIP code	X	X

This table summarizes results from regressions of the percent of applications out of the commuting zone or the concentration of applications in the commuting zone on the ZIP code-level home value decline, after restricting the sample to ZIP codes with homeownership rates below 25%. Regressions include the monthly measures of local labor market conditions and fixed effects for month and ZIP code. The regression in column 1 is weighted by the total number of applications from the ZIP code that month, whereas the regression in column 2 is weighted by the number of applications submitted to jobs in the commuting zone. Variables are defined as in Table 1. Standard errors, adjusted for clustering at the ZIP code level, are reported in parentheses.

Table 7
Importance of home equity

<i>Dependent variable:</i>	Percent of applications out of the commuting zone	Concentration of applications in the commuting zone
	(1)	(2)
Percent with low positive home equity [0 ≤ HE < 20%]	-0.071*** (0.022)	-13.23*** (2.18)
Percent with negative home equity [HE < 0]	-0.126*** (0.017)	-3.17** (1.59)
R^2	0.72	0.68
N	180,161	158,540
Local labor market conditions	X	X
Fixed effects		
Month	X	X
ZIP code	X	X

This table summarizes results from regressions of the percent of applications out of the commuting zone or the concentration of applications in the commuting zone on the percent of mortgagors in the ZIP code with different amounts of home equity (HE). Regressions include the monthly measures of local labor market conditions and fixed effects for month and ZIP code. The regression in column 1 is weighted by the total number of applications from the ZIP code that month, whereas the regression in column 2 is weighted by the number of applications submitted to jobs in the commuting zone. Variables are defined as in Table 1. Standard errors, adjusted for clustering at the ZIP code level, are reported in parentheses. ** and *** indicate statistical significance at the 5% and 1% levels, respectively.

Table 8
Importance of recourse laws

<i>Dependent variable:</i>	Percent of applications out of the commuting zone	Concentration of applications in the commuting zone
	(1)	(2)
Home value decline	0.008 (0.045)	6.84* (3.74)
Home value decline × Recourse	−0.169*** (0.047)	−10.40*** (3.99)
R^2	0.72	0.68
N	180,232	158,598
Local labor market conditions	X	X
Fixed effects		
Month × Recourse	X	X
ZIP code	X	X

This table summarizes results from regressions of the percent of applications out of the commuting zone or the concentration of applications in the commuting zone on the ZIP code-level home value decline and its interaction with an indicator for whether mortgage lenders in the state typically have recourse against borrowers' other assets. All regressions include monthly measures of local labor market conditions and fixed effects for month and ZIP code. The main effect of recourse is absorbed by the ZIP code fixed effects. The regression in column 1 is weighted by the total number of applications from the ZIP code that month, whereas the regression in column 2 is weighted by the number of applications submitted to jobs in the commuting zone. Variables are defined as in Table 1. Standard errors, adjusted for clustering at the ZIP code level, are reported in parentheses. * and *** indicate statistical significance at the 10% and 1% level, respectively.

Table 9
Characteristics of ZIP codes near state borders

	Nonrecourse (<i>N</i> = 997)		Recourse (<i>N</i> = 1,070)		<i>p</i> - value of difference
	<i>Mean</i>	<i>SD</i>	<i>Mean</i>	<i>SD</i>	
Population density (people per square mile)	199	725	605	2,133	0.16
Owner occupied (%)	76.8	11.5	76.6	13.1	0.88
Mobility (different county from 5 years ago, %)	0.2	0.1	0.2	0.1	0.58
Mean commute time (minutes)	22.8	6.7	24.0	6.9	0.18
Distance to commuting zone border (miles)	13.3	10.0	17.5	13.8	0.12
Home equity (%)					
Low positive [$0 \leq HE < 20\%$]	22.9	8.4	23.5	7.7	0.83
Negative [$HE < 0$]	6.9	7.7	16.5	20.1	0.31
Demographics					
Male (%)	50.3	2.1	50.4	2.9	0.81
Married (%)	0.3	0.0	0.3	0.0	0.65
Mean household size	2.5	0.3	2.6	0.3	0.12
Age (%)					
Less than 5 years old	5.9	1.7	6.4	2.7	0.01
5 to 9 years old	7.0	1.8	7.2	1.8	0.40
10 to 14 years old	7.8	1.9	7.9	2.0	0.81
15 to 19 years old	7.6	2.8	7.5	2.5	0.83
20 to 24 years old	4.7	2.8	4.9	3.1	0.30
25 to 34 years old	10.7	2.7	11.3	3.5	0.20
35 to 44 years old	15.5	2.5	15.5	2.6	0.99
45 to 54 years old	14.3	2.6	14.2	3.2	0.82
55 to 59 years old	5.5	1.6	5.6	1.8	0.81
60 to 64 years old	4.9	1.7	4.8	1.8	0.73
Over 65 years old	16.1	5.9	14.7	5.6	0.08
Race (%)					
White	93.2	15.2	90.8	15.7	0.43
Black	0.6	2.0	1.5	5.2	0.09
Asian	0.5	0.9	1.2	2.9	0.15
Hispanic	3.2	7.8	4.6	8.6	0.40
Education (%)					
High school	38.1	8.8	35.9	10.4	0.37
Some college	7.1	3.2	6.4	3.0	0.12
College	11.1	6.5	12.4	7.5	0.24
Graduate school	4.3	3.7	5.5	5.3	0.14

Table 9 (continued)

Household income (%)					
Less than \$10,000	10.1	7.0	9.4	6.8	0.71
\$10,000 to \$15,000	7.7	4.0	7.2	5.4	0.59
\$15,000 to \$25,000	15.6	5.5	14.7	6.9	0.44
\$25,000 to \$35,000	15.2	4.7	14.5	5.8	0.37
\$35,000 to \$50,000	19.1	5.6	18.6	6.7	0.49
\$50,000 to \$75,000	19.5	6.6	19.5	7.3	0.98
\$75,000 to \$100,000	7.4	4.5	8.0	5.3	0.63
\$100,000 to \$150,000	3.8	3.1	5.1	5.2	0.24
\$150,000 to \$200,000	0.8	1.0	1.3	2.2	0.11
Over \$200,000	0.9	1.1	1.6	3.9	0.13

This table reports summary statistics for ZIP codes within 50 miles of state borders where the recourse law changes. Means, standard deviations, and *p*-values of the difference in means, adjusted for clustering at the state level, are reported. Home equity measures are ZIP code averages for May 2008 to December 2009, estimated from CoreLogic's Loan-Level Market Analytics database. Distance to commuting zone border is the geodesic distance. All other ZIP code characteristics are from the 2000 U.S. Census.

Table 10
Applications out of the commuting zone and recourse laws, regression discontinuity analysis

	All ZIP codes		High income ZIP codes	Low income ZIP codes	High income ZIP codes	
	(1)	(2)	(3)	(4)	(5)	(6)
Recourse	-36.7*** (9.1)	-21.1*** (6.7)	-22.5*** (5.9)	1.6 (7.4)	-23.2** (7.8)	-24.4*** (6.7)
Recourse × Above median price drop						-6.1** (2.5)
R^2	0.60	0.60	0.60	0.70	0.54	0.53
N	10,874	10,874	8,774	2,097	5,838	6,404
Fixed effects						
Month	X	X	X	X	X	X
Border strip × commuting zone	X	X	X	X	X	X
Distance to border		X	X	X	X	X
Distance to border squared		X	X	X	X	X
Distance to border cubed		X	X	X	X	X
ZIP code age and race controls					X	

This table summarizes results from regressions of the percent of applications out of the commuting zone on an indicator for whether mortgage lenders in the state typically have recourse against borrowers' other assets, after restricting the sample to ZIP codes within 50 miles of state borders where recourse law changes. All regressions include month and location (10-mile border strip × commuting zone) fixed effects. Where indicated, regressions include controls for distance to the state border, and the age and race variables listed in Table 9. High (low) wealth ZIP codes are ones with above (below) median adjusted gross income in 2006. The specification in column 6 includes an uninteracted indicator for ZIP codes that experienced an above median price decline. Variables are defined as in Table 1. Regressions are weighted by the number of applications submitted from the ZIP code that month. Standard errors, adjusted for clustering at the state level, are reported in parentheses. ** and *** indicate statistical significance at the 5% and 1% level, respectively.

Table 11
Breadth of job search and recourse laws, regression discontinuity analysis

	In commuting zone						Out of zone
	All ZIP codes		High income ZIP codes	Low income ZIP codes	High income ZIP codes		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Recourse	-1879*** (478)	-1006** (361)	-1232*** (191)	739 (808)	-1080*** (196)	-1099*** (168)	832 (683)
Recourse × Above median price drop						-328*** (67)	
R^2	0.65	0.66	0.65	0.64	0.67	0.65	0.32
N	7,898	7,898	6,808	1,087	5,141	5,484	6,195
Fixed effects							
Month	X	X	X	X	X	X	X
Border strip × commuting zone	X	X	X	X	X	X	X
Distance to border		X	X	X	X	X	X
Distance to border squared		X	X	X	X	X	X
Distance to border cubed		X	X	X	X	X	X
ZIP code age and race controls					X		

This table summarizes results from regressions of the concentration of applications on an indicator for whether mortgage lenders in the state typically have recourse against borrowers' other assets, after restricting the sample to ZIP codes within 50 miles of state borders where recourse law changes. Specifications are the same as those described in Table 10. Variables are defined as in Table 1. Regressions are weighted by the number of applications submitted to jobs in (columns 1–6) or out (column 7) of the commuting zone from the ZIP code that month. Standard errors, adjusted for clustering at the state level, are reported in parentheses. ** and *** indicate statistical significance at the 5% and 1% level, respectively.

Table 12
Number of applications and recourse laws, regression discontinuity analysis

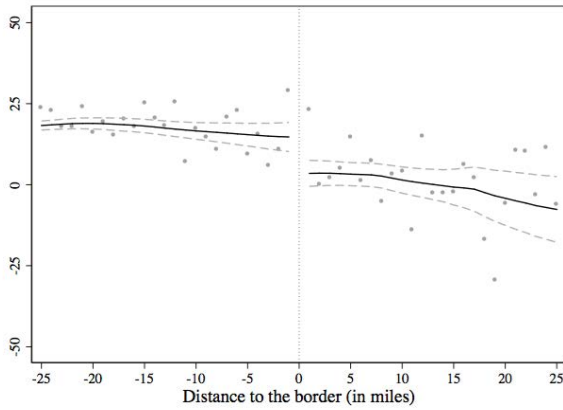
	Management	Executive	Entry level	Requires <1 year of experience
Panel A. In commuting zone				
	(1)	(2)	(3)	(4)
Recourse	-1.59** (0.58)	-0.15 (0.18)	1.51* (0.76)	6.07*** (0.90)
R^2	0.36	0.11	0.30	0.16
N	6,808	6,808	6,808	1,487
Panel B. Out of commuting zone				
	(5)	(6)	(7)	(8)
Recourse	2.92*** (0.93)	0.31 (0.27)	-0.29 (0.88)	1.68 (2.70)
R^2	0.07	0.08	0.06	0.17
N	6,195	6,195	6,195	1,247
Fixed effects				
Month	X	X	X	X
Border strip \times commuting zone	X	X	X	X
Distance to border	X	X	X	X
Distance to border squared	X	X	X	X
Distance to border cubed	X	X	X	X

This table summarizes results from regressions of measures of job type on an indicator for whether mortgage lenders in the state typically have recourse against borrowers' other assets, after restricting the sample to ZIP codes within 50 miles of state borders where recourse law changes. The specification is the same as in column 3 of Table 10. Regressions are weighted by the number of applications submitted to jobs in (columns 1–3) or out (columns 5–7) of the commuting zone from the ZIP code that month, respectively, or the number of number of applications submitted to jobs in (column 4) or out (column 8) of the commuting zone for which experience information is available. Variables are as defined in Table 1. Standard errors, adjusted for clustering at the state level, are reported in parentheses. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

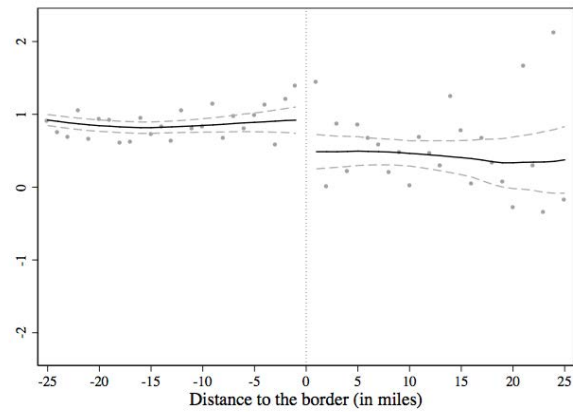
Appendix Figure A1 Volume of applications in and out of the commuting zone

Panel A. In commuting zone

Any applications?

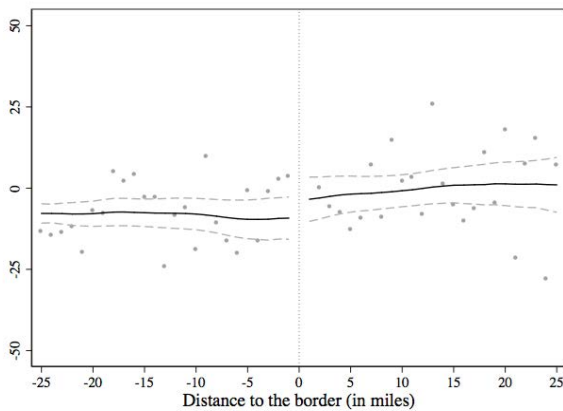


Log applications per worker

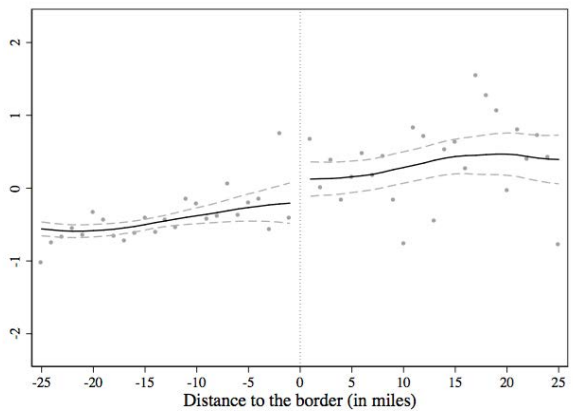


Panel B. Out of commuting zone

Any applications?

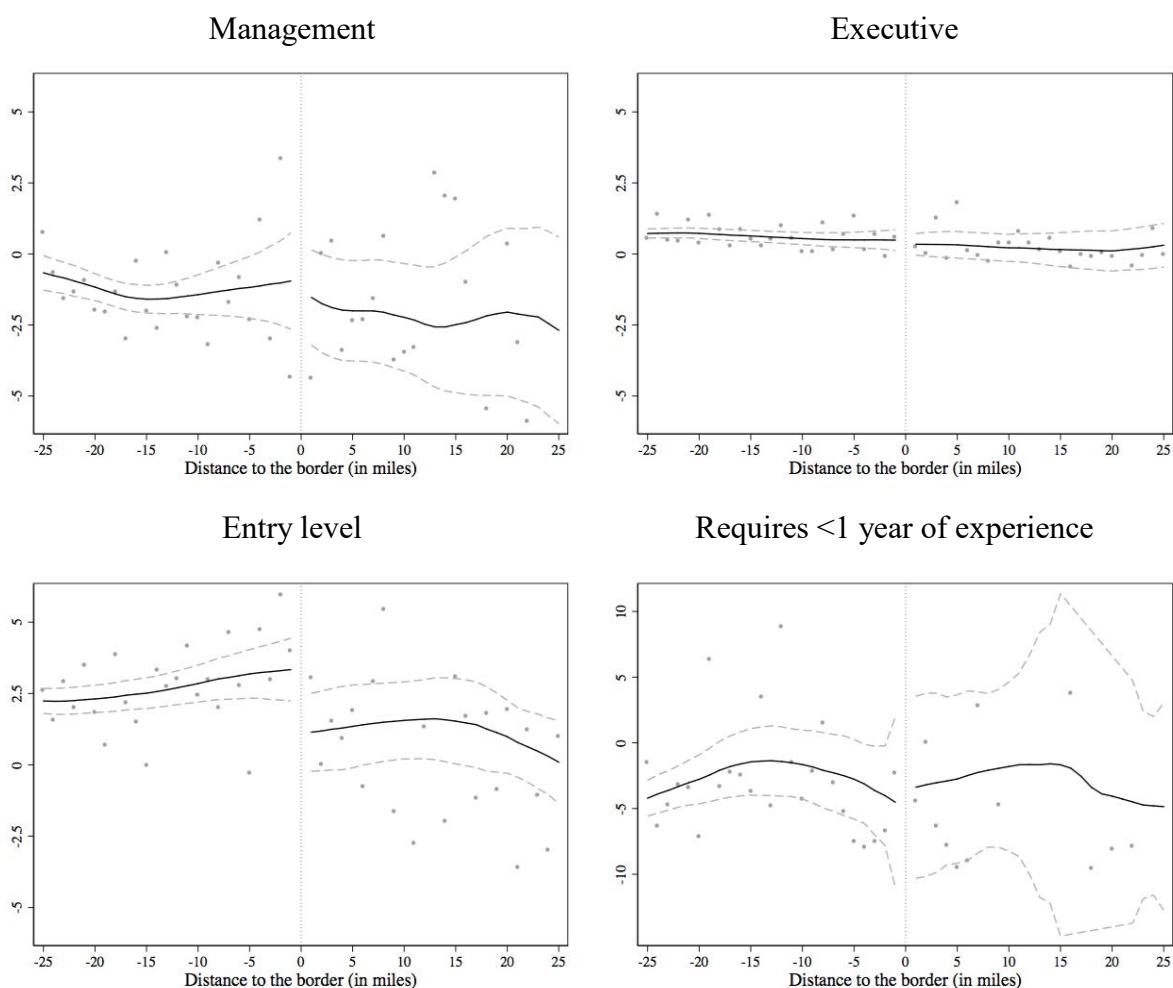


Log applications per worker



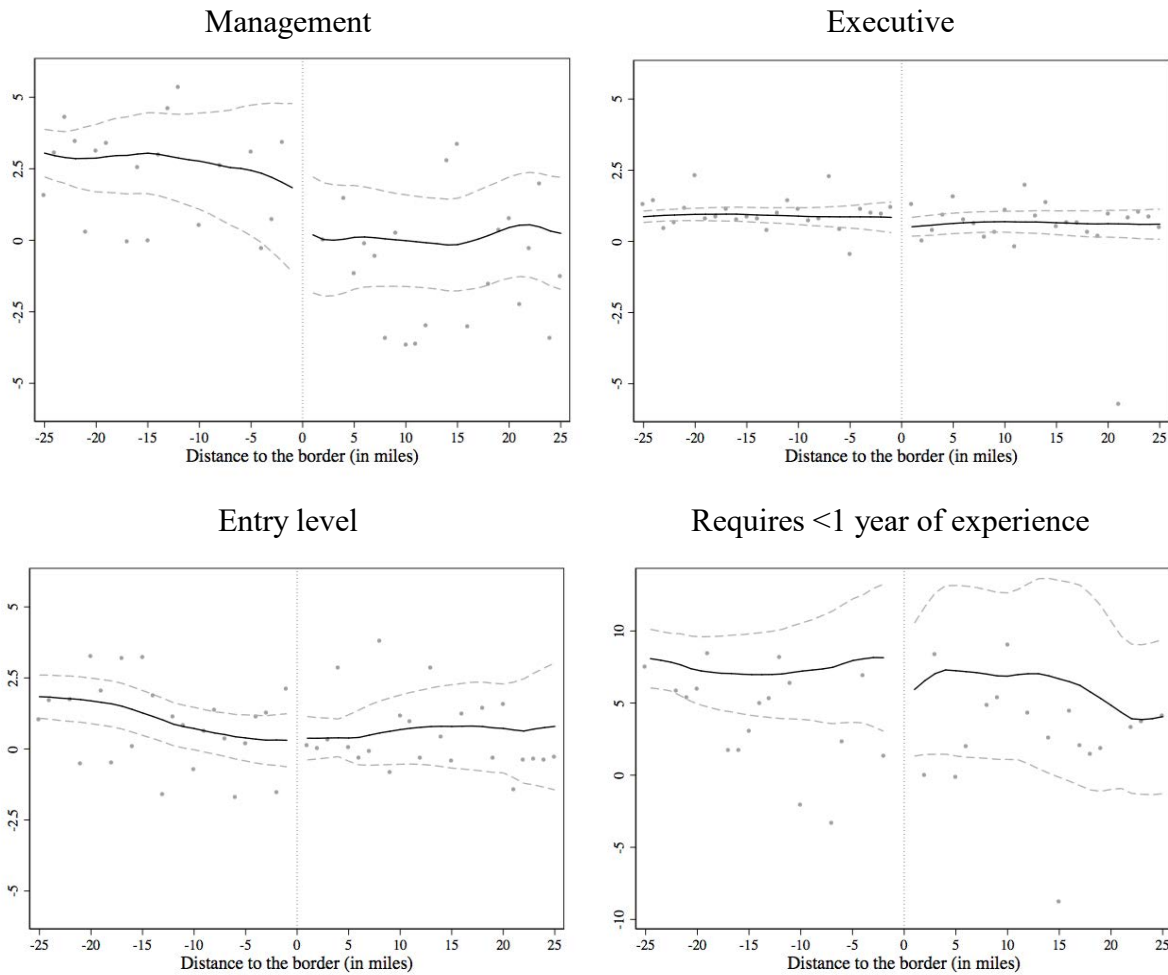
These figures plot indicators for any applications or the number of applications to jobs in (Panel A) or out (Panel B) of the commuting zone for the months between May 2008 and December 2009 for U.S. ZIP codes that are near a state border where recourse law changes and had above median adjusted gross income in 2006. Variables are defined as in Table 1, and the methodology used is the same as in Figure 3. The distances are labeled as negative for states with recourse mortgages.

Appendix Figure A2 Types of jobs applied for in the commuting zone



These figures plot the percent of applications in the commuting zone submitted to jobs of various types for the months between May 2008 and December 2009 for U.S. ZIP codes that are near a state border where recourse law changes and had above median adjusted gross income in 2006. Variables are defined as in Table 1, and the methodology used is the same as in Figure 3. The distances are labeled as negative for states with recourse mortgages.

Appendix Figure A3
Types of jobs applied for out of the commuting zone



These figures plot the percent of applications out of the commuting zone submitted to jobs of various types for the months between May 2008 and December 2009 for U.S. ZIP codes that are near a state border where recourse law changes and had above median adjusted gross income in 2006. Variables are defined as in Table 1, and the methodology used is the same as in Figure 3. The distances are labeled as negative for states with recourse mortgages.

Appendix Table A1
Number of applications and recourse laws, regression discontinuity analysis

	In commuting zone		Out of commuting zone	
	Any applications?	Log applications per worker	Any applications?	Log applications per worker
<i>Dependent variable:</i>	(1)	(2)	(3)	(4)
Recourse	8.66* (4.83)	0.30 (0.23)	-5.77 (6.29)	-0.35 (0.34)
R^2	0.51	0.51	0.21	0.45
N	8,743	6,795	8,743	6,177
Fixed effects				
Month	X	X	X	X
Border strip × commuting zone	X	X	X	X
Distance to border	X	X	X	X
Distance to border squared	X	X	X	X
Distance to border cubed	X	X	X	X

This table summarizes results from regressions of measures of the volume of applications on an indicator for whether mortgage lenders in the state typically have recourse against borrowers' other assets, after restricting the sample to ZIP codes within 50 miles of state borders where recourse law changes. The specification is the same as in column 3 of Table 10. Variables are defined as in Table 1. Standard errors, adjusted for clustering at the state level, are reported in parentheses.