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THE EFFECT OF EQUITY BARRIERS  
ON FOREIGN INVESTMENT IN  
DEVELOPING COUNTRIES

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ABSTRACT

This paper investigates stock performance in emerging markets in relation to their accessibility by foreign investors (as measured by the investability index of the IFC). Using the Stehle (1977) model, we reject for most markets integration and fail to reject for all segmentation. We find that there is a positive relationship between a stock's P/E-ratio and its investability index for most emerging markets, suggesting that barriers to access by foreigners have a negative impact. For four markets, this result is robust to the inclusion of the world beta and the degree of international spanning of the domestic market. A significant negative relationship between the investability index and stock return is only found for Jordan. This is likely because the effects of changes in the degree of access over time confound the cross-sectional relationship between return and investability indexes.

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## Introduction

Equity portfolio flows to developing countries have increased sharply in magnitude in recent years, especially to the so called emerging countries. Total equity flows to developing countries are estimated to have been \$13.0 billion in 1992, quadruple that of three years earlier (Table 1). Equity flows are quite concentrated among a small group of emerging countries (e.g., Latin America received about 60% of all equity flows to developing countries in 1992). Even though relatively still small for developing countries in aggregate (about 10% of the aggregate net resource flows they received in 1992), these flows are an important source of finance for some developing countries.

Equity flows have taken place in a number of forms: direct equity purchases by investors in the host stock markets; investments through country funds; issue of rights on equities held by depository institutions (American and Global Depository Receipts (ADRs and GDRs)<sup>1</sup>; and direct foreign equity offerings. In the last three years equity flows has largely taken place through DRs. The volume of ADRs/GDRs issued for equity claims of developing countries is estimated to have been about \$13.9 billion over 1989-1993 (first six months of 1993).<sup>2</sup> Until recently, next in importance were (closed-end) country funds: during 1989-1993, new country funds were created for developing countries with an aggregated size of \$8.0 billion. The sharpest relative increase in the last few years has been direct purchases of equities: these are estimated to have been about \$5.8 billion in 1992, up from \$0.8 billion in 1990, and were

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<sup>1</sup>ADRs and GDRs are receipts issued by financial intermediaries in industrial countries against shares held in custody by these intermediaries in the developing countries.

<sup>2</sup>This includes direct offerings on foreign capital markets by corporations in developing countries outside the ADR/GDR structure (under Rule 144A in the US). These have been minimal, however.

second in importance in 1991 and 1992.

The increased importance of direct equity purchases by foreigners on emerging stock markets may in part be attributed to the progressive removal of barriers by developing countries on foreign participation in their stock markets. Many developing countries have removed restrictions on foreign ownership, liberalized capital account transactions, improved their accounting and information standards, and in general have made it easier for foreigners to access their markets (see further, for example, Mathieson and Rojas-Suarez (1993) and Reisen and Fischer (1993)). Particularly in Europe and Latin America, many countries have now no or very few restrictions on access by foreigners to their markets and treat foreign investors in most ways identical to domestic investors.

At the same time, returns on stock markets in emerging countries have been high; for example, the IFC composite index for Latin America was up 294.2 percent over 1988-1992, compared to 108.4 percent for the S&P500. This may also have been a factor motivating the larger inflows of foreign equity. At the same time, however, the volatility of rates of return has also been high reaching, for example, more than 100% for Argentina.

The increase of these equity flows to a number of developing countries and the opening up of their stock markets raise a number of issues. An important one is what the effect has been on the risk-return tradeoff in these markets of the removal of barriers, i.e., how much has the risk-return tradeoff changed. The purpose of this paper is to quantify the effect of barriers to access by foreign investors on stock prices and rates of return.

To answer this question, we use the newly created indexes by the IFC Emerging Markets Data Base (EMDB) on the degree of foreign access or "investability." The IFC investability

indexes capture for each stock the barriers to free access by foreigners (general inflow or outflows restrictions, general or sector-specific ownership restrictions, remittance restrictions, other exchange restrictions, restrictions on capital structure, etc.). These indexes should thus be a good indicator of the relative importance of barriers across securities at a given point in time in one market or across a number of markets, or of changes in barriers over time.

Summarizing our results, we find a positive relationship between P/E-ratios and the degree of access for almost all the countries. For four out of the seven markets we study in detail, this result is robust to the inclusion of the world beta and the degree of international spanning of the domestic market. Only for Jordan and Mexico, however, is this result robust to the inclusion of the additional factor of the supply of stocks. For the relationship between rates of return and the investability index, we find evidence of a negative sign for only Jordan, which is also less robust. For other countries, we do not find that abnormal stock returns are related in a systematic fashion to a stock's investability index.

The outline of this paper is as follows. Section one presents an overview of possible analytical frameworks. Section two describes the data we use and provides some statistics on the rates of return. We then perform the Stehle (1978) test for market segmentation or integration for each market to investigate whether these markets indeed show signs of being segmented. Section three provides the empirical results of these tests of market integration and market segmentation. We then describe in the next section the concept of the investability index as developed by the IFC and provide some statistics on the investability indexes. Section five perform the tests on the (cross-sectional and time series) relationship between the P/E-ratio and the abnormal rate of return on an individual stock on the one hand and the level of its

investability index on the other hand and performs some robustness tests. Section six concludes.

## I. Overview of Possible Analytical Models

*Tests Assuming No Barriers* Without barriers, international integration tests can be performed using the various international asset pricing models that have been developed. Past empirical tests along these lines specifically concerned with developing countries--and which assume no barriers--are, for example, Lessard (1973, 1974), Divecha, Drach and Stefek (1992), Bekaert (1993), Buckberg (1993), Diwan, Errunza and Senbet (1993b), De Santis (1993), Harvey (1993), and Tesar and Werner (1993). All papers find that there are significant diversification benefits available from investing in developing countries. Most of these tests, however, use a specific asset pricing model which assumes full integration. As a result, one doesn't know whether these diversification benefits can be achieved.

*Tests Assuming Barriers* Without explicitly incorporating the type and severity of barriers in an asset-pricing model, several papers have investigated market integration (or segmentation) using the test developed by Stehle (1977). The advantage of the Stehle methodology is that it allows for tests of both full integration and full segmentation. Jorion and Schwartz (1986), focussing on interlisted stocks, reject full market integration between Canada and the US using this test, something which they attribute to legal barriers. Mittoo (1992) investigates the same issue and finds segmentation pre-1981, but integration afterwards, especially for interlisted stocks.

With barriers, assets in different markets may have different expected rates of return even when their risk characteristics are the same. One way of testing integration in the presence of

barriers is to model the barriers explicitly, derive the resulting theoretical equilibrium asset prices, and to verify the model using actual asset prices. Following Jorion and Schwartz (1986), barriers can be classified into indirect barriers, arising from differences in available information, transaction costs, accounting standards, etc.; and legal barriers, arising from the different judicial status of foreign and domestic investors, e.g., ownership restrictions and taxes. Typically only legal barriers are incorporated in asset pricing models as these can easily be modelled explicitly.<sup>3</sup>

Theoretical models here are Black (1974 and 1978), Stulz (1981), and Errunza and Losq (1985 and 1989). For imperfectly accessible stocks (i.e., access up to a share  $\delta$  less than 1), Eun and Janakiraman (1986) and Stulz and Wasserfallen (1992) develop models. These papers find theoretical "mispricing" resulting from the barriers given the specific asset-pricing model used. As expected, the analytical predictions on asset pricing with barriers crucially depend on the type of market segmentation.

There are some empirical investigations building on these models for industrial countries. Hietala (1989) investigates the pricing of individual Finnish stocks which can be owned by foreign as well as domestic investors (unrestricted) versus stocks which can only be owned by domestic investors (restricted). Some other papers have applied these tests to developing countries. Errunza and Losq (1985) find tentative empirical support for a hypothesis of mild<sup>4</sup>

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<sup>3</sup>For these reasons, Bekaert (1993) employs a non-parametric approach for testing the relationship between barriers and measures of market integration.

<sup>4</sup>Defined as a situation where the industrial countries' security markets are well integrated and developing country investors can invest in all these (foreign) security markets but foreign investors can not vice-versa invest in developing countries.

market segmentation. Errunza, Losq and Padmanabhan (1992) find that many emerging markets are neither completely integrated with nor completely segmented from industrial countries.

For imperfectly accessible stocks, Stulz and Wasserfallen (1992) test their model for Swiss stocks and find that a relaxation of investment barriers lowers substantially the value of the shares available to foreigners only relative to the value of the shares available to all investors. Bailey and Jagtiani (1992) use this model to investigate differential pricing of restricted and unrestricted stocks for Thailand. They find that cross-sectional difference in the severity of foreign ownership explains some of the variation in the premiums of unrestricted shares over restricted shares, leading to a mildly segmented capital market.

## II. The Data

The raw data we have cover 20 emerging markets. The price and rate of return data are generally available from 1975 on. Table 2 and 3 provide some basic statistics for the rates of return on the IFC indexes and other market data in these emerging markets over the period 1989-1992. Annex 1 describes the criteria used for creating the indexes.

As can be observed from Table 2, the IFC indexes have in general increased, for some countries by multiple factors (e.g., Argentina). There is also a great variation in the market capitalization across countries.<sup>5</sup> The rates of return in emerging markets are in general high, but so are the standard deviations (Table 3). The highest rate of return is for Argentina, more

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<sup>5</sup>It is important to note that the IFC indexes cover only a subset of all stocks listed on the various exchanges, varying between 39% (Turkey) and 90% (Colombia) in terms of market capitalization. Typically, because of its selection criteria, the IFC index will be weighted towards the larger market capitalization and more liquidly traded stocks.



than 100% on an annual basis. However, Argentina also has the highest standard deviation, almost 130%, and the highest range. In general, the rates of return and standard deviations for the emerging markets are much higher than those for the industrial countries. Table 3 also provides the skewness and kurtosis measures, which indicate that the rates of return are not likely drawn from normal distributions. Jarque-Bera tests for normality bear this out: for most markets it rejects normality (see further Claessens, Dasgupta and Glen, 1993).

Table 4 provides some cross-sectional information on the monthly rates of return of the individual stocks for each market (the methodology used for creating the individual stocks' rates of return is described in Claessens, Dasgupta and Glen, 1993). There is a great cross-sectional variation in the monthly rates of return behavior. Autocorrelation coefficients likewise vary over a wide range.

### III. Test of Market Segmentation

We first use the model of Stehle (1977), as also applied by Jorion and Schwartz (1986), Errunza, Losq and Padmanabhan (1992) and Mittoo (1992), to investigate the hypothesis of market integration or segmentation separately for each emerging market. The Stehle model assumes that the CAPM holds and that exchange risk is not priced. The test requires running the following regressions. First, we project the rate of return of the domestic IFC market indexes,  $j$ ,  $j = 1, \dots, K$ , on the rate of return on a world portfolio index, here approximated by the Morgan Stanley Capital International (MSCI) World Index (the net dividends reinvested series), to get the orthogonal component in the domestic index (note that all time subscripts are omitted):

$$(1) \quad R_j = \alpha_{0j} + \alpha_{1j}R_w + V_{j-w}$$

where  $R_j$  is the rate of return on the index in market  $j$ ,  $R_w$  is the rate of return of the world index, and  $V_{j-w}$  is the component orthogonal to the projection of  $R_j$  on  $R_w$ .

We then regress the world rate of return on the various IFC indexes' rates of return to get the orthogonal components here.

$$(2) \quad R_w = \delta_{0j} + \delta_{1j}R_j + V_{w-j}$$

where  $V_{w-j}$  is the component orthogonal to the projection of  $R_w$  on  $R_j$ .

The Stehle test then involves two cross-section tests, using the orthogonal components of the regressions (1) and (2). Under full integration, the parameter on the slope coefficients (denoted here by  $\beta_{i,j-w}$ ) of the individual stocks' rates of return on the orthogonal component of the regression of the local index on the world portfolio should not be significantly different from zero. Under complete segmentation, the parameter on the slopes coefficients (denoted here by  $\beta_{i,w-j}$ ) of the individual stocks' rates of return on the orthogonal component of the regression of the world portfolio on the local index should be not be significantly different from zero. In other words, assuming complete integration or complete segmentation, equation (3) or (4) should

hold for the rate of return on stock  $i$  in market  $j$ :

$$(3) \quad E(R_{ij}) = \gamma_{0j} + \gamma_{1j}\beta_{ij,w} + \gamma_{2j}\beta_{ij,w-j}$$

$$(4) \quad E(R_{ij}) = \eta_{0j} + \eta_{1j}\beta_{ij} + \eta_{2j}\beta_{i,w-j}$$

where  $R_{ij}$  is the return on stock  $i$  in market  $j$ ,  $i = 1, \dots, N_j$ , where  $N_j$  is the number of stocks in market  $j$ . Under complete integration,  $\gamma_2 = 0$  and  $\eta_2 \neq 0$  and under complete segmentation  $\eta_2 = 0$  and  $\gamma_2 \neq 0$ .

Since we have actual rates of return, we need to decompose the rates of return in an expected component and an unexpected component. Under complete integration:

$$(5) \quad R_{ij} = E(R_{ij}) + \beta_{ij,w} [R_w - E(R_w)] + \beta_{ij,w-j} V_{j-w} + \epsilon_{ij}$$

Under complete segmentation:

$$(6) \quad R_{ij} = E(R_{ij}) + \beta_{ij}[R_j - E(R_j)] + \beta_{i,w-j} V_{w-j} + \theta_{ij}$$

Substituting equation (3) into (5) and (4) into (6) we get two equations which give us the empirical model under market integration (7) and under segmentation (8), respectively:

$$(7) \quad R_{i,j} = \gamma_{0,j}(1-\beta_{i,j,w}) + \gamma_{2,j}\beta_{i,j-w} + \beta_{i,j,w}R_w + \beta_{i,j-w}V_{j-w} + \epsilon_{i,j}$$

$$(8) \quad R_{i,j} = \eta_{0,j}(1-\beta_{i,j}) + \eta_{2,j}\beta_{i,w-j} + \beta_{i,j}R_j + \beta_{i,w-j}V_{w-j} + \theta_{i,j}$$

Estimating these equations using the two-pass approach often used in empirical studies of the traditional asset-pricing models (see Shanken, 1992) is not straightforward here since the  $\beta$ 's are measured with error--there is thus an errors-in-the-variables problem--and the cross-section equations (7) and (8) are biased. To overcome this problem, cross-section tests have traditionally been done using portfolios of stocks, in the expectation that the formation of portfolios will reduce the measurement error (the Fama-Macbeth (1973) method). Because of the limited data we have here, this is difficult (there are few stocks for each country with complete data on rates of return, on average less than 20). We therefore use the rates of return on the individual stocks directly.

We use the non-linear, seemingly unrelated regression (SUR) technique, which is asymptotically efficient and equivalent to MLE (see further Gibbons, 1980 and 1982).<sup>6</sup> This method is consistent, but may not have good small sample properties. We therefore use all securities which are consistently available in a given market over the 1989-1992 period. We estimate for each market  $N$  equations ( $N$  being the number of securities in the market) as a

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<sup>6</sup>We use the SAS routine SYSNLIN (version 5.0) for the NLSUR. Other approaches are the MLE method of Litzenberger-Ramaswamy (1979); the procedure outlined in Gibbons (1980); and the odd/even instrumental variable approach of Mankiw and Shapiro (1986). We did use the odd/even method but this method had a lower power as it could neither reject market segmentation nor integration for any of the countries.

system of equations with cross-equation restrictions on the  $\gamma$  and  $\eta$  coefficients in each market and no restrictions on the  $\beta$ -s (except that they are constant over time). The estimation technique allows for correction of heteroskedasticity across stocks and exploits the contemporaneously correlated errors. The parameter estimates and other statistics are in Table 5.

The  $R^2$ s for the segmentation and integration tests (last column) vary between 0.18 (Jordan) and 0.74 (Nigeria) and are of similar magnitude (by country) for the two tests (reflecting the fact that the two systems are basically run with the same set of fundamental variables). The integration hypothesis is rejected at the 5% level for 10 out of the 16 countries for which we have consistent data.<sup>7</sup> The segmentation hypothesis is not rejected at the 5% level for all countries and at the 10% level for only three countries (India, Korea and the Philippines). For two countries, the segmentation test did not converge (Colombia and Malaysia). Combining the two tests, market integration can and market segmentation cannot be rejected (at the 5% level) for eight countries (Brazil, Greece, Korea, Mexico, Pakistan, the Philippines, Taiwan and Thailand). For six countries neither market segmentation nor market integration can be rejected (Chile, India, Jordan, Nigeria, Venezuela and Zimbabwe), possibly indicating a low power of our test.

The results of these estimation techniques can be compared with the results for Canada-US: Jorion and Schwartz (1986) find strong evidence of market segmentation; and Mittoo (1992) finds evidence of market segmentation for the pre-1981 period, but integration for the post-1981

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<sup>7</sup>Data for individual stock rates of return are missing for Argentina for all years; for Turkey and Indonesia, data were only available since 1987 and 1990 respectively; for Portugal no stock has consistently data available for the 1989-1992 period.

period. For developing countries, Errunza, Losq and Padmanabhan (1992), using IFC EMDB data over the 1976-1987 period, reject complete market integration for all eight developing countries they study and reject complete market segmentation for five (Brazil, Chile, Greece, Korea and Mexico) of these eight countries. They conclude that "mild" segmentation describes the market structure for these five countries best. Compared to their results, we find that relatively fewer countries are not integrated (10 out of 16 compared to 8 out of 8), but more are segmented (14 out of the 14 markets which converged compared to theirs 5 out of 8).

It is worth noting that the overall fit of both cross-section equation (5) as well as (6) improves over time.<sup>8</sup> The fact that both the complete segmentation model as well as the complete integration model describe the cross-sectional behavior of returns better as time progresses is somewhat puzzling. A priori, we expected that the integration model would have performed better over time--as countries opened up--and the segmentation model worse. One explanation is that both equations essentially use the same set of explanatory variables, world and local rates of return, and consequently that the behavior over time of the overall fit has to be similar.

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<sup>8</sup>We first estimate the cross-section equations (5) and (6) for every month during the period December 1988-December 1992, where we use estimates of the various betas obtained from using the previous three years of data (instead of running it as systems with constant betas). We then measure the degree of improvement over time in overall fit for each country through the correlations of the  $R^2$ s of the cross-section equations with an index which runs from 1 (first cross-section equation) to 49 (last cross-section equation). For both equation (5) and (6), 14 out of the 16 correlations are positive (of which 6 significantly so at the 5% level).

#### IV. Barriers and the Investability Indexes

This section provides some statistics on the investability indexes. Barriers to access by foreigners are more severe for developing countries than for industrial countries. While many developing countries have liberalized in recent years, many of these countries have had in the past--and some still have--capital controls affecting the general ability to invest in and repatriate capital out of the host country, restrictions on foreign investment (e.g., restrictions on the general permissible share of foreign ownership), and other sector or company-specific ownership restrictions.

In addition to these legal barriers, other barriers likely limit foreigners' access to these markets.<sup>9</sup> The IFC investability indexes are, however, only concerned with legal barriers. In particular, the investability indexes are compiled on the basis of information on type (and/or changes) of identifiable barriers (in or out, ownership restrictions, remittance restrictions, other foreign exchange restrictions, restrictions on capital structure, etc.). Typically, however, the index reflects the share of stocks which can be held by foreigners, i.e., the  $\delta$ -constraint. Indirect barriers are not incorporated in the index (even though the IFC categorizes the severity of these indirect barriers by market, see the IFC Emerging Markets Factbook, 1993). Annex 3 describes the method used for creating the investability indexes and the restrictions in place as of end-1992 for some selected emerging markets.

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<sup>9</sup>For example, there can be restrictions imposed on investors by the home country (e.g., restrictions on the share of foreign assets held by pension funds) and other regulatory and accounting standards in the home country. Also, indirect barriers may exist, such as: the efficiency of the domestic stock (and other financial) markets; the regulatory, accounting, enforcement, etc. standards in the host country; the different forms of sovereign (or transfer) risk; taxes (see Demirguc-Kunt and Huizinga (1993)) and other transaction costs. We do not analyze these restrictions.

The investability indexes are available since December, 1989, initially for 10 of the 20 markets in the EMDB and later for 18. The investability indexes take on values between 0.0 (complete lack of access by foreigners) and 1.0 (complete access). Table 6 provides information (the number of stocks, mean level, the standard deviation, the range, and the skewness of the indexes) on the cross-sectional distributions of the investability index within a given country, at different points in time. Figure 1 provides the time-series plots for the mean and cross-sectional standard deviation for the seven countries which have consistent data for the investability index and the stock rates of return since 1989 (Chile had missing data for 1991 and could thus not be plotted). As can be observed from the figure and also by comparing the three panels of Table 6, there are sharp movements over time in the degree to which foreigners can access these markets. For Mexico, for example, the index goes up from an average of 0.10 in January 1989 to 0.61 at the end of 1992 and further to 0.80 in March 1993. Similarly, the average for Brazil goes up from 0.18 to 0.53. Except for Malaysia, the cross-sectional mean is lower in Asia, an indication that few markets in this region have opened up.

The cross-section standard deviation of the index at the end of 1992 varies greatly, from 0.00 for Taiwan to 0.51 for Colombia, Greece and Venezuela. In general, the cross-sectional standard deviation is lower in Asia (even though less so for Pakistan, the Philippines and Malaysia), an indication that these countries have mostly market-wide, not sector- or stock-specific restrictions.

The time-series plots of the cross-section variation and Table 7 show that the European and Latin American countries have seen the greatest variation over time in the mean index (STDMN in Table 7 is higher for European, except Jordan, and Latin American, except



Venezuela, countries). Asian countries have the least variation over time. Taking into account also the low cross-sectional variation in Asian countries, this reflects that those Asian countries which opened up during this period did so in a market-wide fashion. There are altogether four markets which have little time-series variation in access (i.e., for which in Table 7  $STD_{MN} \leq 0.04$ ), but a reasonable cross-sectional variation (i.e., for which, according to Table 6,  $STD \geq 0.16$  at any point in time and for which we have complete data on rates of return and investability indexes): Jordan, Malaysia, the Philippines, and Thailand. Of these four, Jordan has the lowest mean index; 0.09 at the end of 1992.

#### **V. Tests of the Relationship between the Investability Indexes, P/E-ratios, and Rates of Return**

So far, we have found evidence of market segmentation for about 10 markets. We now proceed to incorporate barriers more formally in our empirical tests, using the models of Eun and Janakiraman (1986) and Stulz and Wasserfallen (1992), and the application of these models by Bailey and Jagtiani (1992). We start with the assumption that the world and the emerging country have the same numeraire (dollars) so that exchange risk is not priced. Consequently, we focus on the dollar rates of return. We further assume that the residents of the emerging countries have full access to foreign financial markets and foreign stocks. The high levels of flight capital observed for many developing countries indicate that this is a reasonable assumption. Foreigners are, however, restricted from full access to the emerging markets and can only invest up to a fraction  $\delta$  measured by the investability index.

When the  $\delta$ -constraint is binding, two prices for the same security will occur: a

"domestic" price for that share of the stock which can only be held by domestic residents and a "foreign" price for the share of the stock which can be held by both foreigners and domestic residents. Compared to a situation with no restrictions a discount can arise for the domestic price and a premium for the foreign price. The ratio of foreign to domestic prices will, among others, depend on the supply of both classes of shares, relative to domestic and foreign investors' wealth.

Apart from the fact whether the constraint on ownership is binding on the foreigners--which we assume it is, the degree to which the domestic market offers unique risk-return characteristics from a world capital market point of view plays an important role in determining the existence and size of the discount and premium. If the domestic market can be mimicked perfectly using world assets, then foreigners will not be willing to pay a premium for emerging markets' stocks. And if domestic residents can lay off the risk of their emerging market stocks through positions in stocks available in the world capital markets, then they do not require a discount on emerging markets stocks, even if they are forced to hold them because of the  $\delta$ -constraint.<sup>10</sup>

We tests these relationships using individual stocks' P/E-ratios and rates of return for the seven countries for which we have consistent data on returns and investability indexes.<sup>11</sup> Since

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<sup>10</sup>Notice that this approach resembles segmentation/integration tests where the residual of a projection of the local return on the world return (and vice-versa) was used. Here the local index is mimicked more generally using (in principle) all worldwide traded assets. Since barriers and associated "mispricing" of individual securities can affect the overall domestic stock market, (announcements of) barriers on individual securities can lead to a market wide effect through "spill-over" effects (see further Eun, Claessens and Jun, 1993). We do not attempt to control for these effects.

<sup>11</sup>The results for 7 other countries are available upon request.

we are not studying unrestricted and restricted shares of the same firm, but rather shares of individual firms which vary in degree of restrictiveness, we cannot calculate the ratio of foreign to domestic prices here.<sup>12</sup> However, we can study the price-earning ratio of a stock. Similarly, we can use domestic rates of return (instead of the difference between the returns to foreign and domestic shareholders).

We estimate mimicking portfolios on the basis of the IFC-index for the emerging market and the MSCI-indexes for 13 industrial countries. Specifically, the mimicking portfolios are created through OLS-regressions of (the rates of return on) the indexes on (the rates of return on) the 13 MSCI-indexes. We then use the predicted values from this regression as the rates of return on the mimicking portfolio.<sup>13</sup> As in Stulz and Wasserfallen (1992) and Bailey and Jagtiani (1992), we also include in the estimations a size (or supply) variable, here taken as the log of the market capitalization of each stock,  $MV_{i,t}$ . Finally, we use actual instead of expected P/E-ratios or returns.

We thus model the P/E-ratio, or alternatively, the (excess) rate of return of domestic stock  $i$  in market  $j$ , as:

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<sup>12</sup>Even though we have some stocks of the same firm (e.g., Telmex shares A, B, C and L) which differ in degree of investability, the sample of such stocks is small. Other foreign prices are available in the form of country-fund and ADRs prices. For an analysis of country-fund prices, see Hardouvelis, La Porta and Wizman (in this conference) and Diwan, Errunza and Senbet (1993a and 1993b).

<sup>13</sup>The mimicking is, as expected, generally poor as these markets have a low correlation with markets of industrial countries. The residual domestic risks are consequently quite large.

$$(9) \quad (P/E)_{i,t} = \alpha_{0,t} + \alpha_{1,t}\delta_t + \alpha_{2,t}\hat{\beta}_{w,j,t} + \alpha_{3,t}[\hat{\beta}_{j,j,t} - \hat{\beta}_{a,j,t}] + \alpha_{4,t}MV_{i,t} + \epsilon_{i,t}$$

where  $\delta_{i,t}$  indicates the share foreigners can buy of a particular stock  $i$  at time  $t$  (the investability index),  $\beta_{w,i}$  is the slope coefficient of stock  $i$  on the world portfolio (here the MSCI world index),  $\beta_{a,i}$  is the slope coefficient of stock  $i$  on the mimicking portfolio  $A$ ,  $\beta_{j,i}$  is the slope coefficient of stock  $i$  on the local market index  $j$  (note that the  $\alpha$ s are not stock specific), and where the subscript  $t$  for the  $\beta$ -coefficients indicates that these are estimates updated every month using the previous three years of data. The difference between  $\beta_{j,i}$  and  $\beta_{a,i}$  represents the domestic risk that cannot be hedged through positions in foreign assets. When there is no residual risk to bear,  $\beta_{j,i} - \beta_{a,i} = 0$  and the world CAPM prevails. To estimate the betas, we regress the rates of return during the three year period preceding the date on the respective indexes.

The coefficients  $\alpha_{1,t}$  depend on the relative risk aversion, the wealth of both foreign and domestic investors and the total supply of restricted and unrestricted shares. The prior is, when the access constraint is binding, that the coefficients  $\alpha_{1,t}$  are positive (an increase in  $\delta$  relaxes the foreign constraint, decreases the required rate of return and increases the P/E-ratio). The coefficients  $\alpha_{2,t}$  represent the world market price of risk and are expected to be negative. The coefficients  $\alpha_{3,t}$  are expected to be negative as a decrease in the ability to mimic local risk increases the required rate of return and lowers the P/E-ratio. Finally,  $\alpha_{4,t}$  are expected to be positive as an increase in the supply of assets, keeping liquidity constant, raises the required rate of return and lowers the P/E-ratio. When using the rate of return as the dependent variable,

the signs of the  $\alpha$ -coefficients are expected to take the opposite value.

We do not impose time-series restrictions on the coefficients  $\alpha_{0,t}, \alpha_{4,t}$  for each market, i.e., we do not use the SUR-technique we used for the integration/segmentation tests. Rather we employ the Fama-MacBeth (1973) methodology where we estimate a separate cross-section equation for each month in the 1989-1992 period for each market and then calculate averages, standard deviation, etc. of the time series of the slope coefficients.

The results for the P/E-ratio alone are in the first panel of Table 8 which reports the means of the slope coefficients, the t-value for the time-series means, and the means of the individual t-values. The other panels cover the results when we include different combinations of the other right-hand side variables. The results for the rates of return are in Table 9.

The results for regressions of the P/E ratios on the index alone confirm the notion that the P/E ratio is positively related to the degree of access by foreigners, suggesting that barriers to access have a negative impact on prices. As measured by the t-value for the mean slope coefficient, for all seven countries, the mean slope is significantly positive (with Brazil marginally). When including other explanatory variables, we find that the positive sign for  $\delta$  is robust to the inclusion of the world beta and the degree of international spanning of the domestic market for four markets. Only for Jordan and Mexico, is the positive sign for  $\delta$  maintained across all regression specifications. For the other countries the sign for  $\delta$  turns at times negative, e.g., when including the lagged (log) market value, the third panel. For Jordan, the signs for the other explanatory variables are not always as expected, e.g., several of the  $\beta$ s have positive signs. In case of Mexico, the signs for the  $\beta$ s are as hypothesized all negative (e.g., see the very last line of Table 8).

The t-values for the time-series means show that there are quite a number of significant coefficients. For example, for the regression which includes all explanatory variables (the last panel of Table 8), 14 out of 28 coefficients are significant on the basis of the t-values for the mean (the mean of the individual t-values shows, however, that many of the individual regression coefficients were insignificant). But, the signs are often not as expected.

For the rates of return, Jordan is the only country which has the expected negative sign for  $\delta$  (first panel, Table 9). This negative sign is robust, but loses significance, when including  $\beta_w$  and  $\beta_j - \beta_a$ , which themselves also have the expected positive and often significant signs. The negative sign for  $\delta$  disappears when the lagged (log) market value is included. For none of the other markets do we find that returns are negatively related to the investability indexes in a consistent fashion.

The degree to which the model explains the cross-section variation in the P/E-ratios and rates of returns varies greatly across equations and countries. While in general we have low explanatory power, with the time-series mean of the adjusted  $R^2$ s reaching mostly less than 10 percent, at times the mean adjusted  $R^2$  reaches 70 to 80 percent (figures are not reported).

Our findings may be better understood by referring back to Figure 1. This figure showed that there is much erratic behavior in the investability indexes, with large swings from month to month for some countries, e.g., for Greece in late 1990 the mean index falls in one month from 0.77 to 0.57 and then goes back up to 0.75, casting some doubt on the manner in which these data were constructed. More importantly, as was noted before, only for Jordan is the mean index stable and low, while at the same time displaying a relatively large cross-sectional variation.

The behavior of the indexes may explain why we only find consistent results for both P/E-ratios and rates of return for Jordan. For three of the four countries where the investability index was stable while still displaying relatively large cross-sectional variation, the access constraint likely did not bind (Malaysia, Philippines and Thailand). This implied one wouldn't expect a cross-sectional relationship between returns and the indexes. For the other countries, the investability indexes were not stable (the mean increased for Brazil and Mexico and behaved erratically for Greece). This could imply that time-series effects (of opening up or closing of the markets, or of data problems) confounded the cross-sectional relationship between returns and the investability indexes. Or put differently, the ex-post rates of return are likely a poor proxy of the ex-ante rates of return. A positive relationship between the P/E-ratio and the investability indexes could still be maintained if countries opened up (e.g., Brazil and Mexico), however, since then both the cross-sectional and time-series effects would go the same way.

To control for the times-series effects of market-wide opening up, we standardize the P/E-ratio of each stock in a given market by dividing it by the market-average P/E-ratio. In this way, we control for changes in the P/E-ratio of each stock for market-wide developments which may be related to the opening up of the market. Admittedly, this is a crude way of proxying for events which affect a particular stock's P/E ratio over time, but it should provide some indication of how robust our results are to the time-series behavior of the P/E-ratios.

The results are reported in Table 10. Comparing the coefficients in Table 10 with those of Table 8, we find that the cross-sectional relationship between a stock's P/E-ratio and its  $\delta$  is robust to this standardization. While, as expected, the slope-coefficients drop significantly, the t-values are not affected. If anything, the cross-sectional effect of  $\delta$  on the P/E-ratio is

significant at higher levels than in Table 8. By multiplying the slope coefficients with the mean level of  $\delta$  (from Table 7), the relative sensitivity of a stock's P/E-ratio with respect to  $\delta$  can be compared across countries. Excluding Thailand (which has a very low slope coefficient), the mean sensitivity is 0.91, with a standard deviation of only 0.44, indicating some evidence of a common pattern.

We also perform a second robustness test. This involves controlling for the stocks' industry (sector).<sup>14</sup> Our previous results may be capturing differences in P/E-ratios by industry to the extent that foreign ownership restrictions differ systematically by industry. Since we have often a limited number of stocks for each country, we cannot control for each industry without running out of degrees of freedom. We therefore classify stocks in two groups: Non-Banks and Banks. Ownership restrictions appear to differ most systematically between these two groups. We perform this second robustness test for two countries, Malaysia and Brazil. Malaysia has 23 stocks in the Non-banking group and 6 in Banks. The Non-banks group has a much higher  $\delta$  (100%) than the Banks group (30%): the P/E ratio for Non-banks is on average about twice as high as the P/E-ratio of the Bank stocks. Average rates of return do not differ between the two groups. Brazil has 15 stocks in Non-banks and only 3 in Banks. The P/E ratio for Non-banks is on average slightly higher as that of Bank, but average rates of return do not differ between the two groups.

The results of similar regressions as in Tables 8-10 but now with a dummy added for sector (Non-banks = 0, Banks = 1), are reported in Table 11. As can be observed, for both countries and for the P/E-ratio as well as for the rate of return equations, the sector dummies

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<sup>14</sup>We are grateful to Donald Lessard for suggesting this extension.



have the right (negative) sign, indicating that the industry classification affects a stock's P/E-ratio. For Malaysia, however, the dummies are never significant, while for Brazil only three out of eight are significant in case of the P/E-regressions and none for the rates of return regressions. (Note, however, that there are only three stocks in the Brazil Banks group.) Introducing the sector dummy does affect the other slope coefficients, however. In particular, for Malaysia the t-statistics for  $\delta$  become insignificant for almost all specifications. For Brazil, on the other hand, t-statistics often improve. For Malaysia, this raises the possibility that the regressions on the P/E-ratio on  $\delta$  without sector dummy are misspecified because of multicollinearity between  $\delta$  and sector (i.e., sector-specific factors other than  $\delta$  determine a stock's P/E-ratio in such a way that high  $\delta$  sectors end up with high P/E-ratios and vice-versa). As we use no factors other than  $\delta$  and sector-dummy to control for a stock's P/E-ratio, we cannot determine conclusively either way whether it is the sector or the level of  $\delta$  which is driving the relationship between  $\delta$  and P/E-ratio. In case of Brazil, there is no evidence of a coincidence between the industry sector of a stock and its  $\delta$ .

The negative results for the rates of return are consistent with Bekzaert (1993). He finds that there is not a significant relationship between ownership restrictions and the integration of an emerging market with world markets. He conjectures that ownership restrictions are not binding or are being circumvented.

## VI. Conclusions

Tests of market integration using the Stehle (1977) model, employing nonlinear, seemingly unrelated regressions (equivalent to MLE), reject market integration hypothesis for most and fail to reject segmentation for all. In particular, we find that over this period Brazil, Greece, Korea Mexico, Pakistan, the Philippines, Taiwan and Thailand are segmented from international markets.

We have evidence that the degree of investability affects P/E-ratios for seven countries in the expected way. When including other explanatory variables, we find this result to be robust for four markets to the inclusion of the two additional explanatory variables, and for two, Jordan and Mexico, when including three additional variables. It is also robust to the standardization of the P/E-ratios. When using rates of return, we only find for Jordan the expected results.

Our weak results for the rates of return are likely because we cover time-series as well as cross-section effects. Without any change in access, i.e., on a cross-sectional basis, one would expect stocks which are more accessible to have lower return. However, many markets have opened up and as a result stock prices have increased, implying that ex-post returns have been high (even though expected returns may have declined). This implies that on cross-sectional basis one may not find a negative relationship between a stock's return and its investability index. The other possibility, of course, is that the CAPM is not the right model to use.

Our results indicate two possible avenues for improvements: one, we should attempt to keep the degree of access over time constant; and two, to expand the model to test for the

importance of the investability indexes in explaining rate of return behavior. This is left for future research.

**Annex 1: General criteria for inclusion in the IFC indexes<sup>1</sup>**

IFC selects stocks for inclusion in the indexes on the basis of three criteria: size, liquidity, and industry. The indexes include the largest and most actively traded stocks in each market, with a target index total representing the top 60% of total market capitalization at the end of each year and, as a second step, the top 60% of total trading value during each year. Size is measured by market capitalization; liquidity is the total value of shares traded during the year.

Only stocks that are listed on one of the major exchanges in the emerging markets are included in the index. The index will not include stocks whose issuing company is headquartered in an emerging market but listed only on foreign markets.

If several stocks meet the liquidity and size criteria, but only one or two are needed, IFC selects the stocks that represent industries that are not yet well represented in the IFC index.

In a few instances, particularly where multiple classes of stocks are common (e.g., Brazil and Mexico), IFC may include in the IFC index more than one class of stock for the same company even though they are not necessarily actively traded. The purpose is to give a balanced view of the capitalization of companies that have other classes of stock that are actively traded.

Stock market "float" (i.e., the amount of issued stock held by the general public and generally assumed to be available for trading) is not a consideration in weighing the indexes, due to the difficulty of obtaining accurate information in a timely manner.

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<sup>1</sup>This annex and annex 2 is copied from the IFC methodology notes.

## **Annex 2: Criteria used by the IFC for the investable indexes and restrictions on foreign investors in selected countries**

### **Criteria used by the IFC for the investable indexes**

As a first screen, stocks are included in the investable indexes if they are in the global index (see below) and are available for purchase by non-resident investment institutions to some degree; the degree is determined by national laws and by company statutes. Governments and companies impose a variety of restrictions on foreign ownership, which may also differ by sector. In addition, individual firms may restrict foreign ownership of (certain classes of) shares. Several examples are shown here; Annex 2 summarizes the restrictions in effect at the end of 1992 for some selected markets.

General national limits, such as "foreigners as a group may not own more than 10% of any company."

Special class of shares, such as A and B class shares in the Philippines. The two are equivalent except that foreigners may not own A class shares;

Sector restrictions, most commonly used to limit foreign ownership of financial institutions, energy producers, utilities, and the media;

Single foreign holder limitations on general classes of shares, such as Brazil's "no more than 5% of the voting classes, nor more than 20% of aggregate capital" or Colombia's 10% limit per investor. The IFC rule in this regard is to use the aggregate that foreign investors as a whole may acquire.

Example: In Colombia, foreigners may own 100% of most companies, although no single foreigner may own more than 10%. The investable market capitalization would be considered as 100%.

"Foreign Board" adjuncts to the main stock exchange, where foreign investors may trade listed stocks among themselves, assuring that trades conducted there will not cause the foreign ownership content to exceed maximum permitted levels.

Prohibitions on individual foreign investors while permitting multiple foreign mutual funds, if they meet certain criteria, such as minimum fund size and experience. The IFC rule in this regard is to consider the market as open as it is to authorized investors, using the "aggregate investor" rule noted above for individual stock investability factors.

Company statutes that impose limits that differ from national law in some markets. In those cases, IFC uses the most restrictive limit.

Example: The national limit is 49% but a company's articles of incorporation set a limit

of 25%. IFC would use a weight of 25%.

National limits on the aggregate permitted foreign investment. For example, Taiwan, China set a ceiling of US\$2.5 billion on foreign inflows when it opened its stock market in January 1991. At the end of 1992, this represented about 2.5% of total TSE capitalization. In this case, IFC would apply the relative shares of the available stocks within the market against the aggregate limit.

Example: The national limit in a market is \$1 billion, and the investable index in the market consists of two stocks, XYZ Inc. and ABC Corp., which have available market capitalizations of \$2 billion and \$500 million, respectively. In the absence of the aggregate limit, the investable index would use \$2 billion and \$500 million as the available market capitalization.

However, these amounts would exceed the limit, and IFC would apply the relative share approach: two stocks represent 80% and 20% of the available market capitalization, so the investable index would use \$800 million and \$200 million as the two stocks' "available" market capitalization.

For the calculation of the various IFC price and rate of return indexes, the investable market capitalization of each stock is used for its weight in the index instead of the stock's total market capitalization.

Example: XYZ, Ltd. has a total market capitalization of \$100 million but national law prohibits foreign ownership of more than 49% of a company. The IFC Global Index would use the full \$100 million as the stock's market capitalization while the Investable Index would use only \$49 million.

To take concerns regarding illiquidity or relatively small market capitalizations into account, the IFC excludes stocks from the investable indexes if:

- 1) trading value for the year totals less than \$10 million, using total trading value unweighted for foreign access;
- 2) the investable market capitalization is less than \$25 million. An exception occurs when the investable capitalization is small but the trading is large. IFC will not exclude a stock if the value traded exceeds \$100 million for the year, regardless of the stock's investable capitalization.

Example. A stock in Korea has a total capitalization of \$240 million and trading totaling \$1,300 million for the year. With the 10% limit currently in effect in Korea, the investable capitalization is only \$24 million. It is clearly an accessible, large and liquid stock, and foreigners are unlikely to have difficulty in trading it.

In rare cases, the selection screens could produce fewer than five stocks in an investable

market, which is insufficient for an index. If that happens, IFC will select as many stocks as needed to reach the minimum of five stocks, using investable capitalization ranked in decreasing order by size.

#### Restrictions on foreign investors at the end of 1992 in selected countries

- Argentina** The market is considered generally 100% investable; some corporate statute limitations apply.
- Brazil** The market is considered generally investable; since May 1991 foreign institutions may own up to 49% of voting common stock and 100% of non-voting participating preferred stock. Some corporate statute limitations (e.g., Petrobras common are off-limits) apply.
- Chile** Foreign portfolio investment is considered to enter Chile through Law 18657 of 1987 regarding Foreign Capital Investment Funds, which limits aggregate foreign ownership to 25% of a listed company's shares.
- Colombia** The market is considered 100% investable from February 1, 1991.
- Greece** The market is generally 100% investable.
- India** A press note issued by the Ministry of Finance of the Government of India on September 14, 1992, announced that foreign institutional investors (FIIs) could henceforth invest in all listed securities in both primary and secondary markets. FIIs are required to register with the Securities and Exchange Board of India before making any investment. The market is considered effectively open from November 1, 1992.
- Investments are subject to a ceiling of 24% of issued share capital for the total holdings of all registered FIIs and 5% for the holding of a single FII in any one company. The ceiling includes the conversion of fully and partly convertible debentures issued by the company.
- Indonesia** Until December 1987, the market was closed to foreign investment. In December 1987, the government introduced deregulation measures that allowed foreigners to purchase shares in eight non-joint venture companies. On September 16, 1989, the Minister of Finance of the Republic of Indonesia issued Decree Number 1055/KMK.013/1989, which allowed foreigners to purchase up to 49% of all companies' listed shares, including foreign joint ventures but excluding banks. The Bank Act, 1992, enacted on October 30, 1992, allowed foreigners to invest in up to 49% of the listed shares in three categories of banks - private national, state and joint foreign. Currently only private national banks are listed.

In a few markets, such as Indonesia, companies do not list all the shares outstanding. For its indexes, IFC counts only the shares listed at the stock exchange.

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| <b>Jordan</b>      | The market is considered generally 49% investable.   |
| <b>Korea</b>       | Since January 1, 1992, authorized foreign investors have been allowed to acquire up to 10% of the capital of listed companies; some corporate statute limitations apply (e.g., POSCO & KEPCO 8%, and some are permitted up to 25%). The 10% limit applies separately to common and preferred stock. Under the revised regulations of June 22, 1992, effective in July 1992, companies whose foreign holdings already exceeded 10% could apply to Korea's Securities and Exchange Commission to increase their limit to 25%. As of March 1993, four companies had received permission: Korea Electronic Parts, Korea Long-Term Credit Bank, Trigem Computer and Young Chang Akki. The ceiling automatically declines when foreign-held shares are sold to domestic investors.   |
| <b>Malaysia</b>    | The limit on foreign ownership of Malaysian stocks is subject to some debate. Bank Negara, the central bank, restricts the ownership of banks and financial institutions by foreigners to 30%. However, these limits do not appear to be strictly enforced. Under the Banking and Financial Institutions Act, 1989, the approval of the Minister of Finance is required before foreign investors can buy or sell shares of a licensed bank or finance company amounting to 5% or more. Certain non-bank stocks have different foreign share holding limits for tax and other reasons. These are MISC, Proton, Telekom, Tenaga Nasional, Tai Wah Garments and Yantzekiang. All other stocks are open to foreign portfolio investment without any limits. However, the approval of the Foreign Investment Committee is required for acquiring 15% or more of the voting power of a company by any one foreign interest and for acquiring the assets or interests of a company when they exceed M\$5 million, whether by Malaysian or foreign interests. Except for a few specific cases, IFC uses 100% for most stocks and 30% for banks and financial institutions. |
| <b>Mexico</b>      | Foreign portfolio investment is permitted in designated classes of shares, and since May 1989 in most other shares through the use of the Nafinsa Trust arrangement. It is now considered generally 100% investable, except for banks, where foreign ownership is restricted to 30%.   |
| <b>Nigeria</b>     | Closed to foreign investment.  |
| <b>Pakistan</b>    | The market is considered 100% investable from February 22, 1991.   |
| <b>Philippines</b> | National law requires that a minimum of 60% of the issued shares of domestic corporations should be owned by Philippine nationals. To ensure compliance,   |



Philippine companies typically issue two classes of stock: "A" shares, which may be traded only among Philippine nationals, and "B" shares, which may be traded to either Philippine nationals or foreign investors and which usually amount to 40% of the total. Mass media, retail trade and rural banking companies are closed to foreign investors.

<b>Portugal</b>	The market is considered generally 100% investable; some corporate statute limitations apply, particularly regarding shares issued in privatizations.
<b>Taiwan</b>	The market was opened to foreigners on January 1, 1991, though foreign investors must meet high registration requirements and total cash inflows from abroad cannot currently exceed an official ceiling of \$2.5 billion. There is a 10% limit on aggregate foreign ownership of issued capital. The domestic transportation industry is closed to foreign investors.
<b>Thailand</b>	Various Thai laws restrict foreign shareholdings in Thai companies engaged in certain areas of business. The Banking Law restricts foreign ownership in banks to 49%. The Alien Business Law, administered by the Ministry of Commerce, restricts foreign ownership of stocks in specified sectors to 49%. In addition, other laws provide similar restrictions on foreign ownership. Restrictions are also faced by foreign investors through limits imposed by company by-laws which range from 15% to 65%. The Foreign Board was established in 1988 to facilitate trading in shares registered in foreign names.
<b>Turkey</b>	The market is considered 100% investable from August 1989.
<b>Venezuela</b>	Non-financial stocks are considered generally 100% investable from January 1, 1990, but some restricted classes do exist. Bank stocks are currently not available.
<b>Zimbabwe</b>	Effectively closed to foreign investment by virtue of severe exchange controls.

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**Table 1: Equity Flows to Developing Countries**  
 (Millions of dollars, estimates)

Type of flow	1989	1990	1991	1992	1993 <sup>p</sup>	Total 1989-93 <sup>p</sup>
Country Funds	\$2.2	\$2.9	\$1.2	\$1.3	\$0.4	\$8.0
ADRs/GDRs	-	\$0.1	\$4.9	\$5.9	\$3.0	\$13.9
Direct Equity	\$1.3	\$0.8	\$1.5	\$5.8	\$1.8	\$11.2
Total	\$3.5	\$3.8	\$7.6	\$13.0	\$5.2	

Source: World Debt Tables (1993) and Gooptu (1993)

Notes: for the first six-months of 1993.

Table 2: IFC indexes and other data for each market: January 1989 and December 1992  
(millions of US-dollars, unless otherwise noted)

CTRY	STOCKS	IFC INDEX	P/E RATIO	IFC P/BV RATIO	IFC MARKET CAPIT.	IFC VAL TRADED	IFC MARKET CAPIT.	EXCH RATE
IND	63	59.03	12.19	1.40	8661.31	259.65	12037.54	2063.50
IND	60	233.25	415.96	33.74	11624.16	1068.88	2518.98	28.66
KOR	61	130.26	18.46	2.75	54928.72	69661.92	94233.33	15.16
MEX	62	176.52	453.19	21.84	20176.50	188.03	25175.37	788.40
PAK	58	176.52	453.19	21.84	20176.50	188.03	25175.37	2.62
PER	18	1526.25	205.78	12.34	14.33	2590.99	2427.11	2.73
PHI	30	1526.25	205.78	12.34	14.33	2590.99	2427.11	18.95
THA	62	176.52	453.19	21.84	20176.50	188.03	25175.37	25.30
TUR	29	376.89	900.42	12.83	13.93	2.15	39174.36	27.65
USA	29	376.89	900.42	12.83	13.93	2.15	39174.36	25.49
GBR	26	226.14	537.42	10.12	6.89	1.43	3922.75	155.00
JPN	25	132.93	181.76	15.76	1.09	1.48	3365.03	215.30
PER	23	437.84	503.06	15.05	9.05	2.77	9215.36	0.67
TUR	18	134.41	227.01	2.26	6.95	1.48	9930.80	146.92
ARG	24	188.10	1253.14	0.55	37.99	0.08	1876.49	8540.00
BRA	54	95.00	158.92	4.57	26.43	0.44	18632.57	0.00
COL	26	754.93	3315.58	4.10	12.99	0.78	24280.00	0.99
CUB	21	359.32	2171.44	5.39	27.95	0.97	29643.89	12243.00
MEX	52	642.16	2408.24	3.47	12.28	0.58	1144.98	382.33
VEN	13	147.85	523.61	8.80	15.63	1.89	139060.77	81.77
NGO	15	24	33.82	64.43	5.61	8.98	1878.43	2.30
ZIM	17	384.76	2.03	0.31	397.69	0.10	752.72	36.30
					267.97	0.44	627.63	6.90
								21.50
								5.46

Source: EMDB

Note: The first column under each heading refers to January, 1989 and the second to December 1992. The (double) columns are: Level of the IFC index (1984 = 100), IFC P/E-ratio, IFC P/BV-ratio, IFC Market Capitalization, IFC Value Traded, Total Market Capitalization, and exchange rates (LC/\$). The P/E ratios can be misleading in high inflation countries (such as Argentina and Brazil in the late 1980s, as the earnings are measured as the average flow over the last 12 months, and prices are taken at the end of the periods. Similarly, P/BV ratios can be misleading in a highly inflationary environment).

Table 3: Statistics on the Index Rates of Return (1989-1992), by Country.  
(annual percentage changes)

CTRY	N	MEANCHG	STDCHG	MINCHG	MAXCHG	SKEWCHG	KURTCHG	AUTOCORR
IDN 35	35	-16.6	32.1	-250.6	224.9	0.18	0.12	0.25
IND 48	48	22.2	39.6	-292.6	423.2	0.60	0.92	0.17
KOR 48	48	-4.4	31.3	-230.9	319.0	0.94	1.23	-0.18
MAL 48	48	17.6	21.7	-186.8	155.9	-0.56	0.36	-0.10
PAK 48	48	28.7	31.0	-189.9	423.2	1.64	5.19	0.28
PHI 48	48	14.2	34.7	-351.6	325.0	-0.17	1.59	0.34
TAI 48	48	3.7	49.2	-409.6	359.5	0.14	0.18	0.18
THA 48	48	29.1	29.6	-270.0	201.7	-0.59	0.28	0.25
GRE 48	48	31.4	53.4	-206.8	702.9	1.83	4.01	0.13
JOR 48	48	11.8	20.2	-154.1	193.9	-0.18	0.93	-0.16
POR 48	48	-2.4	25.4	-170.1	348.3	1.34	4.31	0.06
TUR 48	48	33.6	70.5	-377.4	829.5	1.12	1.46	0.22
ARG 48	48	109.0	129.5	-779.4	2137.3	2.34	8.98	-0.12
BRA 48	48	41.8	83.4	-682.7	573.8	0.06	-0.14	-0.09
CHI 48	48	43.5	26.0	-109.7	255.4	0.24	-0.68	0.41
COL 48	48	53.4	39.6	-209.5	448.1	1.61	2.83	0.52
MEX 48	48	47.5	27.0	-170.1	235.9	-0.02	-0.37	0.16
VEN 48	48	48.1	54.1	-313.8	582.6	0.62	1.15	0.33
NIG 48	48	16.1	30.2	-507.1	226.0	-2.98	13.54	0.15
ZIM 35	35	-29.0	31.4	-276.6	180.4	-0.34	-0.42	0.29

Note: The monthly rates are multiplied by 12 to obtain the yearly rates. The standard deviation is obtained by multiplying the monthly standard deviation with the square root of 12. N is number of months, MEANCHG refers to the mean change in the rate of return, STDCHG to the standard deviation of the rate of return, MINCHG and MAXCHG to the minimum and maximum change in the rate of return, SKEWCHG to the skewness coefficient, KURTCHG to the kurtosis coefficient, and AUTOCORR to the first order autocorrelation. First observation for Indonesia and Zimbabwe is January 1990.

Table 4: Min and max values (ranges) of cross-sectional values of monthly time-series of rates of return for all stocks.

CTRY	YRS	AVL	LMEAN-HMEAN	LSTD-HSTD	LMIN-HMIN	LMAX-HMAX	LAUTOLAG-HAUTOLAG					
10N	90-92	90	-0.085	0.388	2.182	-0.776	-0.010	0.007	12.484	-0.838	0.569	
1WD	76-92	69	-0.005	0.078	0.074	0.386	-0.565	-0.131	0.265	2.910	-0.304	0.248
KOR	76-92	105	-0.027	0.057	0.054	0.223	-0.648	-0.071	0.084	1.362	-0.696	0.244
MAL	86-92	75	-0.028	0.076	0.045	0.263	-0.499	-0.069	0.109	2.037	-0.538	0.463
PAK	85-92	77	-0.046	0.096	0.040	0.332	-0.443	-0.062	0.069	1.093	-0.418	0.340
PHI	85-92	34	-0.092	0.079	0.037	0.352	-0.593	-0.085	0.000	2.848	-0.350	0.478
TAI	85-92	77	-0.036	0.055	0.108	0.305	-0.725	-0.172	0.205	1.685	-0.260	0.271
THA	76-92	58	-0.030	0.104	0.063	0.389	-0.517	-0.137	0.138	1.774	-0.448	0.366
GRE	76-92	34	-0.054	0.043	0.050	0.235	-0.497	-0.049	0.042	1.408	-0.310	0.416
JOR	78-92	30	-0.023	0.106	0.057	0.170	-0.468	-0.110	0.139	0.753	-0.374	0.250
POR	86-92	30	-0.051	0.066	0.074	0.324	-0.758	-0.124	0.118	1.885	-0.423	0.275
TUR	87-92	25	-0.084	0.094	0.140	0.389	-0.466	-0.251	0.241	2.274	-0.228	0.415
BRA	76-87	80	-0.091	0.208	0.117	0.753	-4.538	-0.180	0.000	3.628	-0.411	0.646
CHI	76-92	44	-0.017	0.075	0.097	0.348	-0.798	-0.109	0.213	3.011	-0.254	0.387
COL	85-92	22	-0.050	0.057	0.082	0.367	-0.475	-0.103	0.000	3.209	-0.220	0.402
MEX	76-92	83	-0.046	0.099	0.019	0.501	-1.000	-0.012	0.064	3.695	-0.333	0.449
VEN	85-92	17	-0.016	0.075	0.144	0.316	-0.572	-0.301	0.273	1.955	-0.361	0.236
NIG	85-92	25	-0.051	0.057	0.093	0.174	-0.621	-0.392	0.145	0.784	-0.077	0.327
ZIM	76-92	21	-0.097	0.048	0.107	0.250	-0.643	-0.040	0.130	1.239	-0.408	0.267

Notes: Avail is the number of stocks for which data are available during the period. Lmean is the lowest mean rate of return for any stock in a market, and h mean the highest rate of return. Lstd is the lowest standard deviation of the rates of return across all stocks in a given market, hstd the highest. Lmin is the lowest minimum rate of return across all stocks in a given market, hmin is the highest minimum rate of return is a given market. Lmax and hmax, the highest. Autolag is the first autocorrelation, with autolag the lowest and hautolag the highest in a given market. No data on individual stock rate of return were available for Argentina.



Table 5: Slope Coefficients for the Integration and Segmentation Tests

N	Integration		R <sup>2</sup>	I	Segmentation		S	R <sup>2</sup>
	$\gamma_{a,j}$	$\gamma_{i,j}$			$\eta_{a,j}$	$\eta_{i,j}$		
IND 18	0.0333 (0.002)	-0.0265 (0.067)	0.46	not reject	0.0121 (0.786)	-0.1956 (0.076)	not reject	0.46
KOR 22	-0.0106 (0.144)	0.0279 (0.0001)	0.40	reject	0.1841 (0.163)	0.2604 (0.053)	not reject	0.40
MAL 29	0.0089 (0.221)	0.0154 (0.0001)	0.37	reject	NC			
PAK 31	0.0132 (0.012)	0.0108 (0.042)	0.29	reject	-2.623 (0.896)	5.954 (0.895)	not reject	0.28
PHI 16	-0.0243 (0.015)	0.0478 (0.0001)	0.30	reject	0.0523 (0.335)	0.194 (0.079)	not reject	0.29
TAI 20	-0.0148 (0.107)	0.045 (0.0001)	0.68	reject	-0.0674 (0.079)	0.490 (0.389)	not reject	0.68
THA 9	0.0560 (0.036)	0.0215 (0.003)	0.46	reject	0.211 (0.525)	0.546 (0.301)	not reject	0.47
GRE 8	-0.035 (0.067)	0.0545 (0.0001)	0.55	reject	-0.152 (0.605)	0.522 (0.418)	not reject	0.56
JOR 9	-0.0056 (0.460)	0.0092 (0.203)	0.18	not reject	0.059 (0.586)	-0.362 (0.575)	not reject	0.19
BRA 18	-0.002 (0.799)	0.0375 (0.0001)	0.45	reject	-12.54 (0.977)	-8.798 (0.977)	not reject	0.45
CHI 22	0.0627 (0.005)	-0.0024 (0.877)	0.37	not reject	0.188 (0.744)	-2.020 (0.489)	not reject	0.40
COL 20	0.0077 (0.234)	0.0435 (0.0001)	0.33	reject	NC			
MEX 21	0.0227 (0.042)	0.0466 (0.0001)	0.32	reject	0.158 (0.941)	-6.748 (0.927)	not reject	0.32
VEN 12	0.0100 (0.479)	0.0199 (0.390)	0.36	not reject	0.0073 (0.941)	0.940 (0.199)	not reject	0.36
NIG 14	1.196 (0.808)	-0.914 (0.807)	0.74	not reject	6.622 (0.962)	-5.914 (0.961)	not reject	0.74
ZIM 10	0.0006 (0.969)	0.0059 (0.805)	0.22	not reject	0.006 (0.673)	-0.030 (0.176)	not reject	0.22

Notes: Approximate p-values (for the t-statistics) are in parentheses. In spite of using many different starting values for the parameters, and even after the maximum iterations was increased up to 2000 and the convergence criteria was raised to 0.0001, no convergence (NC) was obtained for Colombia and Malaysia for the segmentation test. R<sup>2</sup>s are obtained as one minus the ratio of sum of squared residual (totalled for all equations) over sum of squared totals (totalled for all equations).

Table 6: Cross-sectional analysis of the investability index for each country  
in January 1989, June 1990, and December 1992

CTRY	DATE	NOSTOCKS	MEAN	STD	MAX	SKWENESS
MAL	8901	62	0.84	0.34	1.00	-1.74136
PHI	8901	18	0.28	0.46	1.00	1.08486
THA	8901	29	0.30	0.20	1.00	1.14879
GRE	8901	26	0.31	0.47	1.00	0.88525
JOR	8901	25	0.10	0.20	0.49	1.59749
POR	8901	23	0.74	0.45	1.00	-1.16667
ARG	8901	24	0.58	0.50	1.00	-0.36103
BRA	8901	56	0.18	0.19	0.56	0.31331
CHI	8901	26	0.09	0.12	0.25	0.68705
MEX	8901	52	0.10	0.30	1.00	2.82184
MAL	9006	62	0.86	0.32	1.00	-2.03384
PHI	9006	29	0.22	0.41	1.00	1.43347
THA	9006	34	0.29	0.21	1.00	0.91982
GRE	9006	26	0.77	0.43	1.00	-1.35763
JOR	9006	25	0.10	0.20	0.49	1.59749
POR	9006	27	0.67	0.48	1.00	-0.75423
TUR	9006	18	0.89	0.32	1.00	-2.70579
ARG	9006	24	0.42	0.50	1.00	0.36103
BRA	9006	56	0.10	0.17	0.50	1.22881
CHI	9006	28	0.08	0.12	0.25	0.80870
MEX	9006	54	0.56	0.50	1.00	-0.23005
VEN	9006	13	0.38	0.51	1.00	0.53859
IDN	9212	63	0.26	0.25	0.49	-0.09769
IND	9212	62	0.15	0.12	0.24	-0.62193
KOR	9212	91	0.10	0.02	0.24	0.25280
MAL	9212	62	0.85	0.33	1.00	-1.87221
PAK	9212	58	0.09	0.28	1.00	3.02748
PHI	9212	30	0.25	0.43	1.00	1.24847
TAI	9212	70	0.03	0.00	0.05	-5.67578
THA	9212	51	0.27	0.16	0.50	-0.31587
GRE	9212	32	0.47	0.51	1.00	0.13149
JOR	9212	27	0.09	0.19	0.49	1.71783
POR	9212	30	0.38	0.48	1.00	0.56336
TUR	9212	25	0.80	0.41	1.00	-1.59749
ARG	9212	29	0.79	0.41	1.00	-1.52730
BRA	9212	69	0.53	0.47	1.00	-0.12553
CHI	9212	35	0.14	0.13	0.25	-0.17986
COL	9212	20	0.50	0.51	1.00	0.00000
MEX	9212	66	0.61	0.49	1.00	-0.44428
VEN	9212	17	0.41	0.51	1.00	0.39424

Notes: Statistics provide the cross-sectional distribution of the investability index at a given point in time. The cross-sectional minimum is 0.0 in all markets. No data were available for Nigeria and Zimbabwe.

**Table 7: Time-series analysis of the cross-sectional mean of  
the investability indexes  
(December 1988 - March 1993)**

CTRY	N	MEAN	STDMN	MINMN	MAXMN
IDN	31	0.20	0.07	0.10	0.37
IND	5	0.14	0.01	0.14	0.15
KOR	15	0.10	0.00	0.09	0.10
MAL	52	0.87	0.04	0.76	0.93
PAK	25	0.09	0.01	0.09	0.13
PHI	52	0.23	0.04	0.18	0.28
TAI	27	0.03	0.00	0.03	0.03
THA	52	0.28	0.02	0.23	0.30
GRE	52	0.56	0.19	0.31	0.77
JOR	52	0.10	0.01	0.08	0.13
POR	52	0.57	0.14	0.38	0.83
TUR	44	0.79	0.10	0.56	0.89
ARG	52	0.61	0.15	0.33	0.79
BRA	52	0.30	0.18	0.09	0.54
CHI	40	0.10	0.03	0.06	0.14
COL	26	0.38	0.12	0.25	0.50
MEX	52	0.47	0.21	0.10	0.80
VEN	39	0.41	0.03	0.31	0.44

Notes: The statistics apply to the time series of the mean value of the investability index for a given market. Chile has missing data for 1991.

Table 8: Times Series Summary of Cross-Sectional Regressions of P/E-Ratio against Investability Index (1989-1992)

	MN $\delta$	T $\delta$	MN T $\delta$	$\beta_w$	T $\beta_w$	MN T $\beta_w$	MN $\beta_j$	T $\beta_j$	MN T $\beta_j$	MN MV	T MV	MN T MV
Malaysia	40.45	2.65	-0.04									
Philippines	57.62	3.45	1.08									
Thailand	5.1	3.83	0.64									
Greece	28.31	3.06	1									
Jordan	125.87	3.26	0.23									
Brazil	9.18	1.94	0.12									
Mexico	6.19	4.13	0.98									
Malaysia	53.17	1.71	-0.06	37.44	0.78	0.23						
Philippines	-52.22	-3.18	-0.28	54.88	2.87	-0.23						
Thailand	-25.73	-8.52	-1.01	10.99	9.02	1.44						
Greece	140.18	1.85	0.76	-109.3	-1.63	-0.71						
Jordan	4.2	1.02	-0.01	11.74	3.41	0.21						
Brazil	7.33	2.17	0.3	-0.33	-0.48	-0.39						
Mexico	11.69	3.6	0.38	-1.95	-0.76	0						
Malaysia	71.11	1.62	-0.1				-17.9	-0.3	0.37			
Philippines	10.66	1.89	-0.24				-82.77	-3.05	-1.39			
Thailand	-9.53	-5.89	-0.31				11.9	7.71	1.94			
Greece	288.19	1.64	0.21				236.21	1.56	-1.02			
Jordan	26.31	2.24	-0.08				3.28	0.4	-0.45			
Brazil	10.58	2.33	0.55				-3.25	-4.92	-1.02			
Mexico	9.75	4.77	0.48				-5.49	-2.08	-0.25			
Malaysia	-2.3	-0.15	-0.17				-16.31	-2.73	0.31			
Philippines	-38.41	-2.4	0.04				0.77	0.31	-0.77			
Thailand	-13.47	-7.35	-0.47				1.29	4.6	0.55			
Greece	-65.12	-1.55	0.28				60.61	1.81	0.17			
Jordan	43.6	4.53	0.45				-6.58	-5.34	-0.51			
Brazil	-13.99	-2.27	-0.93				3.15	4.37	1.81			
Mexico	10.46	2.67	0				0.07	0.05	0.69			

Table 8 Continued

	MN $\delta$	T $\delta$	MN T $\delta$	$\beta_u$	T $\beta_u$	MN T $\beta_u$	MN $\beta_i - \beta_u$	T $\beta_i - \beta_u$	MN T $\beta_i - \beta_u$	MN MV	T MV	MN T MV
Malaysia	66.62	1.52	-0.08	16.33	0.34	0.37	-10.28	-0.16	0.51			
Philippines	4.2	1.13	0.04	13.16	2.1	-0.75	-77.76	-3.04	-1.22			
Thailand	-18.4	-6.17	-0.58	4.51	4.36	0.41	9.38	6.2	1.11			
Greece	370	1.71	0.35	-108.1	-1.68	-1.21	240.34	1.6	-1.06			
Jordan	16.97	1.82	-0.02	31.29	3.03	0.21	14.94	1.29	-0.25			
Brazil	10.43	2.22	0.67	-2.18	-3.36	-0.79	-3.92	-6.26	-1.39			
Mexico	12.31	4.06	0.57	-4.34	-1.07	-0.11	-5.56	-1.7	-0.45			
Malaysia	22.94	0.97	-0.13	35.19	0.75	0.18				-7.15	-1.38	0.4
Philippines	17.06	2.33	0.28	85.15	2.97	0.07				-25.4	-3.32	-0.87
Thailand	-27.16	-8.59	-1.17	11.95	10.11	1.61				1.68	7.89	0.96
Greece	-35.86	-0.6	0.35	-32.24	-0.78	-0.45				51.08	1.71	0.16
Jordan	51.39	4.54	0.54	20.06	4.63	0.34				-9.28	-5.09	-0.57
Brazil	-0.23	-0.07	-0.64	-0.01	-0.01	-0.29				2.29	7.1	2.01
Mexico	11.6	2.67	0.07	-1.96	-0.8	-0.18				0.14	0.11	0.64
Malaysia	25.87	0.71	-0.19				-19.77	-0.33	0.34	-16.92	-2.69	0.28
Philippines	21.82	3.12	0.32				-83.72	-3.08	-1.54	-4.14	-3.04	-0.95
Thailand	-10.64	-5.92	-0.4				11.85	7.19	1.58	-0.13	-0.72	0.06
Greece	174.61	1.59	0.02				218.29	1.57	-0.86	47.36	1.79	-0.08
Jordan	32.21	4.07	0.49				3.94	0.47	-0.53	-1.55	-0.98	-0.55
Brazil	5.47	0.84	-0.34				-2.9	-3.09	-0.61	1.5	3.36	1.68
Mexico	10.57	2.91	0.09				-5.46	-1.91	-0.44	-0.62	-0.39	0.65
Malaysia	35.48	0.97	-0.16	14.46	0.31	0.34	-10.87	-0.16	0.51	-7.93	-1.45	0.34
Philippines	39	3.06	0.47	32.96	2.62	-0.41	-74.07	-2.99	-1.22	-14.25	-3.19	-0.8
Thailand	-21.33	-6.18	-0.71	6.46	4.37	0.57	8.23	4.56	0.57	0.5	1.99	0.34
Greece	235.3	1.68	0.06	-53.95	-1.24	-0.75	227.95	1.63	-0.84	35.29	1.71	0.11
Jordan	32.15	3.51	0.53	37.82	3.68	0.34	16.36	1.43	-0.35	-3.63	-1.89	-0.55
Brazil	5.88	0.91	-0.19	-1.37	-2.13	-0.7	-3.02	-3.12	-0.89	1.51	3.29	1.45
Mexico	12.47	2.93	0.23	-4.31	-1.13	-0.22	-5.73	-1.68	-0.62	-0.17	-0.11	0.67

Notes: "MN" stands for the time-series mean of the cross-sectional regression coefficients. "Std" represents the standard deviation of the cross-sectional regression coefficients. "T" stands for the t-value of the mean of the time-series of coefficients, i.e.,  $\text{mean} \cdot \sqrt{n}/\text{std}$ . "MN T" stands for the mean of the t-values of the individual cross-sectional regressions. We use stocks with complete observations from 1/86 through 12/92. "MV" are one month lagged log market values. "b" stands for the investability index. " $\beta_i$ " are world betas. " $\beta_u$ " are the betas against the local index and " $\beta_d$ " are the betas for the mimicking portfolios. For the number of stocks, see Table 6 on the investability index. Results are similar when using local betas and betas from mimicking portfolios separately rather than taking the difference between them.

Table 9: Times Series Summary of Cross-Sectional Regressions of Rates of Return against Investability Index (1989-1992)

	MN $\delta$	T $\delta$	MN T $\delta$	$\beta_w$	T $\beta_w$	MN T $\beta_w$	$\beta_j$	T $\beta_j$	MN T $\beta_j$	M MV	T MV	MN T MV
Malaysia	0.02	1.77	0.13									
Philippines	0.02	1.7	0.35									
Thailand	0.13	4.28	0.48									
Greece	0.03	2.3	0.55									
Jordan	-0.03	-2.05	-0.2									
Brazil	0.02	0.68	0.14									
Mexico	0.01	1.08	0.1									
Malaysia	0.02	1.48	0.09	-0.01	-0.7	-0.16						
Philippines	0.02	1.94	0.42	0	-0.1	-0.08						
Thailand	0.1	2.82	0.44	-0.02	-0.69	-0.12						
Greece	0.03	2.2	0.59	0	0.35	0.18						
Jordan	-0.04	-1.98	-0.22	0.03	1.86	0.25						
Brazil	0	-0.08	-0.01	-0.03	-2.08	-0.28						
Mexico	0.01	0.91	0.06	0.01	1.01	0.27						
Malaysia	0.02	1.65	0.12				0.01	0.91	0.02			
Philippines	0.02	2	0.41				-0.01	-0.98	-0.34			
Thailand	0.12	3.19	0.44				0	-0.14	0			
Greece	0.02	1.82	0.6				-0.01	-0.86	-0.11			
Jordan	-0.02	-1.09	-0.16				0.03	1.56	0.02			
Brazil	0.04	0.89	0.08				-0.01	-0.6	-0.05			
Mexico	0.02	1.41	0.16				-0.01	-1.65	-0.3			
Malaysia	0.01	1.49	0.12				0	-1	0.11	0	-1	0.11
Philippines	0.03	2.69	0.29				0	-0.71	-0.01	0	-0.71	-0.01
Thailand	0.13	3.83	0.46				0	0.69	0.18	0	0.69	0.18
Greece	0.03	2.02	0.43				0	-0.54	-0.14	0	-0.54	-0.14
Jordan	0	0.06	0.01				-0.01	-0.93	-0.09	-0.01	-0.93	-0.09
Brazil	0.02	0.57	0.06				0	0.12	0.1	0	0.12	0.1
Mexico	0	-0.36	-0.11				0.01	2.04	0.47	0.01	2.04	0.47

Table 9 Continued

	MN $\delta$	T $\delta$	MN T $\delta$	$\beta_w$	T $\beta_w$	MN T $\beta_w$	$\beta_l \beta_w$	T $\beta_l \beta_w$	MN T $\beta_l \beta_w$	M MV	T MV	MN T MV
Malaysia	0.02	1.46	0.09	0	-0.11	-0.08	0.01	0.69	-0.02			
Philippines	0.03	2.32	0.47	-0.01	-0.69	-0.15	0	-0.15	-0.21			
Thailand	0.1	2.28	0.32	-0.03	-0.95	-0.08	0.01	0.32	-0.08			
Greece	0.03	1.75	0.55	0	-0.2	0.15	-0.01	-0.58	-0.01			
Jordan	-0.02	-1.38	-0.13	0.05	3.25	0.29	0.05	2.61	0.2			
Brazil	0.02	0.45	0	-0.03	-2.26	-0.18	-0.02	-1.51	-0.2			
Mexico	0.01	0.96	0.15	0.02	1.26	0.34	-0.01	-1.42	-0.38			
Malaysia	0.01	1.27	0.1	0	-0.01	-0.05				0	-1.07	0.09
Philippines	0.04	2.97	0.39	0.01	0.67	0				-0.01	-1.15	-0.05
Thailand	0.09	2.22	0.46	-0.01	-0.33	-0.13				0	0.67	0.05
Greece	0.05	2.74	0.59	0.02	1.14	0.33				-0.01	-1.75	-0.31
Jordan	0.02	0.67	0.09	0.04	2.05	0.34				-0.01	-1.59	-0.19
Brazil	-0.01	-0.23	-0.06	-0.03	-1.97	-0.28				0	0.27	0.07
Mexico	-0.01	-0.55	-0.15	0.02	1.08	0.31				0.01	2.22	0.49
Malaysia	0.01	1.34	0.12				0.02	1.14	0.05	0	-1.01	0.11
Philippines	0.03	3.03	0.37				-0.01	-1.11	-0.35	0	-0.76	-0.05
Thailand	0.11	2.79	0.4				-0.01	-0.26	-0.13	0	-0.07	0.03
Greece	0.03	2.02	0.64				-0.02	-1.15	-0.32	-0.01	-1.01	-0.33
Jordan	-0.01	-0.19	0.06				0.02	1.09	-0.06	0	-0.21	-0.11
Brazil	-0.03	-0.77	-0.07				0	0.36	0.05	0.01	0.48	0.08
Mexico	0	-0.14	-0.06				-0.02	-2.47	-0.39	0.01	2.39	0.52
Malaysia	0.01	1.24	0.1	0	0.37	0.01	0.01	0.87	0.02	0	-1.11	0.09
Philippines	0.04	3.36	0.45	0	0.14	-0.06	0	-0.11	-0.18	-0.01	-1.14	-0.05
Thailand	0.04	0.72	0.2	0.01	0.31	0	-0.01	-0.2	-0.3	0	0.46	0.04
Greece	0.04	2.35	0.67	0.01	0.35	0.19	-0.02	-0.72	-0.15	-0.01	-1.91	-0.46
Jordan	0.01	0.4	0.15	0.06	3.17	0.37	0.05	2.35	0.14	-0.01	-1.01	-0.21
Brazil	-0.01	-0.36	-0.06	-0.02	-1.85	-0.12	-0.01	-0.69	-0.09	0	0.08	0
Mexico	-0.01	-0.49	-0.11	0.02	1.31	0.39	-0.02	-1.88	-0.43	0.01	2.44	0.54

Notes: "MN" stands for the time-series mean of the cross-sectional regression coefficients. "T" stands for the t-value of the mean of the time-series of coefficients, i.e.,  $\text{mean} \cdot \text{sqrt}(n) / \text{std}$ . "MN T" stands for the mean of the individual cross-sectional regressions. We use stocks with complete observations from 1/86 through 12/92. "MV" are one month lagged log market values. "δ" stands for the investability index. "β<sub>w</sub>" are the betas against the local index and "β<sub>l</sub>" are the betas from the mimicking portfolios. For the number of stocks, see Table 6 on the investability index. Results are similar when using local betas and betas from mimicking portfolios separately rather than taking the difference between them.

Table 10: Times Series Summary of Cross-Sectional Regressions of Standardized P/E-Ratio against Investability Index (1989-1992)

	MN $\delta$	T $\delta$	MN T $\delta$	MN $\beta_u$	T $\beta_u$	MN T $\beta_u$	MN $\beta_{\beta_u}$	T $\beta_{\beta_u}$	MN T $\beta_{\beta_u}$	MN MV	T MV	MN T MV
Malaysia	1.95	2.94	-0.04									
Philippines	3.92	3.5	1.08									
Thailand	0.35	3.2	0.64									
Greece	1.89	3.7	1									
Jordan	10.13	3.38	0.23									
Brazil	1.72	3.03	0.21									
Mexico	0.63	4.31	0.98									
Malaysia	2.48	1.8	-0.06	2.68	1.37	0.23						
Philippines	-3.54	-3.23	-0.28	3.64	2.85	-0.23						
Thailand	-1.94	-8.34	-1.01	0.81	9.76	1.44						
Greece	8.17	1.95	0.76	-6.43	-1.68	-0.71						
Jordan	0.39	1.07	-0.01	1.16	3.61	0.21						
Brazil	0.58	2.07	0.34	0.01	0.25	-0.24						
Mexico	1.2	2.93	0.38	-0.36	-1.24	0						
Malaysia	3.18	1.59	-0.1				-0.86	-0.32	0.37			
Philippines	0.67	1.74	-0.24				-5.62	-3.07	-1.39			
Thailand		-0.7	-6.04				0.86	8.43	1.94			
Greece	15.95	1.71	0.21				13	1.61	-1.02			
Jordan	2.54	2.37	-0.08				0.55	0.75	-0.45			
Brazil	1.03	2.69	0.59				-0.3	-5.6	-1.02			
Mexico	0.87	3.65	0.48				-0.29	-1.14	-0.25			
Malaysia	-0.04	-0.05	-0.17				-0.7	-2.79	0.31			
Philippines	-2.59	-2.38	0.04				0.03	0.15	-0.77			
Thailand	-0.99	-7.69	-0.47				0.09	4.79	0.55			
Greece	-3.36	-1.47	0.28				3.36	1.82	0.17			
Jordan	3.69	3.89	0.45				-0.53	-4.6	-0.51			
Brazil	-2.49	-2.38	-0.68				0.4	3.33	1.36			
Mexico	1.07	2.25	0				-0.02	-0.15	0.69			



Table 10 Continued

	MN $\delta$	T $\delta$	MN T $\delta$	MN $\beta_u$	T $\beta_u$	MN T $\beta_u$	MN $\beta_1 - \beta_u$	T $\beta_1 - \beta_u$	MN T $\beta_1 - \beta_u$	MN MV	T MV	MN T MV
Malaysia	3.05	1.54	-0.08	1.66	0.83	0.37	-0.65	-0.22	0.51			
Philippines	0.28	1.11	0.04	0.81	1.92	-0.75	-5.29	-3.06	-1.22			
Thailand	-1.41	-5.93	0.34	4.28	0.41	0.67	6.63	1.11				
Greece	20.83	1.78	0.35	-6.32	-1.67	-1.21	13.39	1.67	-1.06			
Jordan	1.61	1.91	-0.02	3.12	3.07	0.21	1.69	1.56	-0.25			
Brazil	1.04	2.61	0.71	-0.2	-3.5	-0.69	-0.4	-7.26	-1.34			
Mexico	1.22	3.16	0.57	-0.74	-1.54	-0.11	-0.22	-0.67	-0.45			
Malaysia	1.17	1.12	-0.13	2.61	1.35	0.18				-0.32	-1.43	0.4
Philippines	1.13	2.34	0.28	5.67	2.97	0.07				-1.72	-3.4	-0.87
Thailand	-2.05	-8.44	-1.17	0.9	10.34	1.61				0.13	8.45	0.96
Greece	-1.41	-0.45	0.35	-2.16	-0.9	-0.45				2.81	1.76	0.16
Jordan	4.41	3.94	0.54	1.84	4.31	0.34				-0.79	-4.44	-0.57
Brazil	-0.28	-0.99	-0.39	0.06	1.19	-0.12				0.24	8.24	1.67
Mexico	1.23	2.31	0.07	-0.36	-1.37	-0.18				0	0.02	0.64
Malaysia	1.22	0.74	-0.19				-0.95	-0.35	0.34	-0.74	-2.78	0.28
Philippines	1.5	3.16	0.32				-5.69	-3.11	-1.54	-0.31	-3.44	-0.95
Thailand	-0.78	-6.11	-0.4				0.84	7.94	1.58	-0.01	-0.49	0.06
Greece	9.69	1.68	0.02				11.98	1.61	-0.86	2.64	1.79	-0.08
Jordan	2.57	3.43	0.49				0.66	0.88	-0.53	-0.04	-0.29	-0.55
Brazil	0.26	0.47	-0.12				-0.22	-2.99	-0.65	0.19	5.51	1.38
Mexico	1.04	2.39	0.09				-0.28	-1	-0.44	-0.12	-0.64	0.65
Malaysia	1.7	1.04	-0.16	1.59	0.82	0.34	-0.69	-0.23	0.51	-0.35	-1.49	0.34
Philippines	2.66	3.09	0.47	2.16	2.56	-0.41	-5.05	-3.01	-1.22	-0.98	-3.31	-0.8
Thailand	-1.66	-5.86	-0.71	0.52	4.29	0.57	0.57	4.7	0.57	0.05	2.48	0.34
Greece	13.53	1.76	0.06	-3.37	-1.25	-0.75	12.7	1.69	-0.84	1.93	1.77	0.11
Jordan	2.59	2.92	0.53	3.59	3.48	0.34	1.8	1.67	-0.35	-0.25	-1.46	-0.55
Brazil	0.44	0.79	0.07	-0.07	-1.37	-0.51	-0.29	-3.9	-0.96	0.16	4.4	1.1
Mexico	1.29	2.47	0.23	-0.73	-1.63	-0.22	-0.23	-0.68	-0.62	-0.05	-0.29	0.67

Notes: "MN" stands for the time-series mean of the cross-sectional regression coefficients. "T" stands for the t-value of the mean of the time-series of coefficients, i.e.,  $\text{mean}(\sqrt{n})/\text{std}$ . "MN T" stands for the mean of the t-values of the individual cross-sectional regressions. We use stocks with complete observations from 1/86 through 12/92. "MV" are one month lagged log market values. " $\delta_u$ " stands for the investability index. " $\beta_u$ " are world betas. " $\beta_j$ " are the betas against the local index and " $\beta_i$ " are the betas from the mimicking portfolios. For the number of stocks, see Table 6 on the investability index.

Table 11: Times Series Summary of Cross-Sectional Regressions of P/E-Ratio and Rates of Return against Investability Index and Sector Dummy (1989-1992)

Panel A: P/E Ratio										
	MN $\delta$	T $\delta$	MN $\beta_{\Delta}$	T $\beta_{\Delta}$	$\beta_j \beta_{\Delta}$	T $\beta_j \beta_{\Delta}$	MN MV	T MN MV	MN SEC	T SEC
Malaysia	11.40	1.21							-19.39	-0.94
Brazil	2.14	0.59							-2.2	-1.53
Malaysia	27.12	1.01	23.03	0.46					-12.43	-0.53
Brazil	10.54	2.34	-1.6	-1.83					-7.58	-2.98
Malaysia	35.20	1.17			-2.22	-0.04			-21.19	-0.89
Brazil	10.69	2.32			-4.25	-8.5			3.01	1.13
Malaysia	-53.37	-2.66					-19.43	-2.77	-31.08	-1.33
Brazil	-11.68	-1.78					3.19	4.2	-2.98	-1.71
Malaysia	33.16	0.93	1.52	0.03	2.27	0.04			-15.5	-0.55
Brazil	10.62	2.25	-3.45	-2.87	-5.10	-8.84			-1.47	-0.42
Malaysia	-17.74	-0.65	19.03	0.39			-9.17	-1.49	-19.5	-0.77
Brazil	2.86	0.60	-1.22	-1.33			2.24	7.05	-9.74	-2.45
Malaysia	-34.04	-1.13			-2.56	-0.04	-20.54	-2.74	-33.94	-1.26
Brazil	6.62	0.96			-3.19	-2.71	1.49	2.93	-3.86	-0.73
Malaysia	-14.42	-0.41	-1.47	-0.03	3.02	0.04	-10.45	-1.55	-23.48	-0.76
Brazil	6.92	1.02	-2.87	-2.39	-3.74	-3.04	-1.59	3.06	-14.2	-2.24

Table 11 Continued

Panel B: Rates of Return										
	MN $\delta$	T $\delta$	MN $\beta_w$	T $\beta_w$	$\beta_j - \beta_w$	T $\beta_j - \beta_w$	MN MV	T MV	MN SEC	T SEC
Malaysia	0.03	1.27							0.00	0.00
Brazil	0.02	0.69							0.00	0.00
Malaysia	0.03	1.04	0.00	-0.30					0.00	0.00
Brazil	0.02	0.37	-0.03	-1.29					0.00	0.00
Malaysia	0.02	1.17			0.00	0.44			0.00	0.00
Brazil	0.05	1.16			-0.03	-1.62			0.00	0.00
Malaysia	0.02	1.01					0.00	0.25	0.00	0.00
Brazil	0.00	0.04					0.00	-1.73	0.00	0.00
Malaysia	0.03	1.04	0.00	0.02	0.00	0.00	0.00	0.25	0.00	0.00
Brazil	0.04	0.81	-0.03	-1.39	-0.04	-0.04	0.00	-1.73	0.00	0.00
Malaysia	0.02	0.81	0.00	0.23			0.00	0.25	0.00	0.00
Brazil	-0.01	-0.31	-0.02	-0.94			0.00	-1.73	0.00	0.00
Malaysia	0.02	0.78	0.01	0.40	0.01	0.01	0.00	0.49	0.00	0.00
Brazil	-0.02	-0.54			0.01	0.01	0.01	0.41	0.01	0.01
Malaysia	0.02	0.78	0.01	0.40	0.01	0.01	0.01	0.49	-0.00	-0.00
Brazil	-0.00	-0.10	-0.02	-1.22	-0.01	-0.01	0.00	-0.38	0.00	0.00

Notes: "MN" stands for the time-series mean of the cross-sectional regression coefficients. "T" stands for the t-value of the mean of the time-series of coefficients, i.e.,  $\text{mean} \sqrt{n}/\text{std}$ . We use stocks with complete observations from 1/86 through 12/92. "MV" are one month lagged log market values. " $\beta$ " stands for the investability index. " $\beta_w$ " are world betas. " $\beta_j$ " are the betas against the local index and " $\beta_w$ " are the betas from the mimicking portfolios. "SEC" stands for the sectoral dummy. For the number of stocks, see Table 6 on the investability index. Results are similar when using local betas and betas from mimicking portfolios separately rather than taking the difference between them.

Figure 1: Mean and Standard Deviation of the Investability Indexes  
(by country)

