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ESTIMATING THE VALUE ADDED OF ATTENDING PHYSICIANS ON PATIENT OUTCOMES

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ABSTRACT

Despite increasing calls for value-based payments, existing methodologies for determining physicians' "value added" to patient health outcomes have important limitations. We incorporate methods from the value added literature in education research into a health care setting to present the first value added estimates of health care providers in the literature. Like teacher value added measures that calculate student test score gains, we estimate physician value added based on changes in health status during the course of a hospitalization. We then tie our measures of physician value added to patient outcomes, including length of hospital stay, total charges, health status at discharge, and readmission. The estimated value added varied substantially across physicians and was highly stable for individual physicians. Patients of physicians in the 75th versus 25th percentile of value added had, on average, shorter length of stay (4.76 vs 5.08 days), lower total costs (\$17,811 vs \$19,822) and higher discharge health status (8% of a standard deviation). Our findings provide evidence to support a new method of determining physician value added in the context of inpatient care that could have wide applicability across health care setting and in estimating value added of other health care providers (nurses, staff, etc).

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Introduction

Estimating the value provided by health care providers and systems—to encourage areas of value and reduce ineffective areas—remains one of the most challenging and important areas of health policy and practice in the US. Assessing value in health care settings has also become an important practical issue, with large financial implications. For example, as of 2015, the Centers for Medicare & Medicaid Services (CMS) will adjust payments for physicians belonging to large groups to reflect the value of care provided (VanLare, Blum et al. 2012, VanLare 2012). The change at CMS is an effort to motivate physicians and the larger health care system to pursue strategies that improve patient outcomes and reduce unnecessary, costly services.

Current efforts to determine physicians' contributions to patient outcomes have substantial shortcomings. First, risk adjustment and risk standardization methodologies using administrative data are imperfect. Clinicians and researchers have raised concerns that the use of inadequate risk adjustment could incentivize physicians to avoid the most challenging patients (Norton 1992, Dranove D 2002, Shen 2003, Werner and Asch 2005, Hayward and Kent 2008). Second, most existing measures apply to specific diseases, resulting in small sample size problems for physicians with insufficient numbers of patients per disease. Furthermore, disease-specific measures are difficult to use to infer overall value. For example, the Physician Quality Reporting System used by CMS includes 194 individual quality metrics (Centers for Medicare & Medicaid Services 2012), making it cumbersome to interpret. Last, most current methods of measuring physician value focus on outpatient care, while hospital-based physicians now comprise one fifth of generalist physicians (Kuo, Sharma et al. 2009).

What is needed is an estimate of physician value added that is pertinent for a comprehensive range of diseases and treatments, that appropriately risk adjusts for baseline clinical severity to mitigate adverse selection pressure, and that measures outcomes that are important to patients, such as health status and function.

In the field of education research, investigators have developed a value added methodology to isolate teacher effects on student outcomes based on change in test scores during the course of an academic year (Harris 2011) (McCaffrey 2009, Chetty 2011). We apply a similar approach to estimate the value added of physicians based on change in health status during the course of a hospitalization, adjusted for patient severity of illness and characteristics. We then estimate associations between our measure of physician value added and patient outcomes (from a hold-out sample) to further demonstrate the utility of our approach. We find modest but important impacts on patient outcomes—approximately an 8% SD increase in health at discharge associated with assignment to a physician with a value added score in the 75th percentile versus 25th percentile of the distribution, as well as reductions in total costs and length of stay.

Although our measures of value added of attending physicians, like all similar measures used in the literature, are difficult to separate from the value provided by other personnel caring for the patients, the measure should be interpreted broadly to reflect all actions and decisions related to attending physicians. Our findings lay the groundwork for a new approach to support value-based payment reform that is applicable to a broad set of patient diseases and physicians.

Conceptual Framework and Empirical Approach

We extend the education literature to the hospital setting by linking changes in patient health with their specific physician during their hospitalization using a Value

Added Model (VAM). The basic idea is that health (like educational achievement) is a cumulative process that is a function of an individual's health endowment (e.g. genetics, birth weight) and the subsequent investments and decrements to their health as they age:

$$H_{it} = f(H_{io}, I_{i0}...I_{it}, E_{i0}...E_{it}, \mu_{it})$$
(1)

where health at time t is a function of an individual's initial health (H_{io}) , her history of investments $(I_{i0}...I_{it})$ (e.g. physician interactions) and decrements in health due to illness $(E_{i0}...E_{it})$ as well as any positive or negative shocks in the current period (μ_{it}) . Simplifying the relationship to be linear and focusing on investment in health gives:

$$H_{it} = \beta_0 I_{i0} ... + \beta_t I_{it} + \delta H_{io} + \mu_{it}$$
 (2)

Because the data do not contain the complete history of investments and decrements to health between birth and adulthood, we need to assume a geometric distributed lag, which imposes geometric decay of the parameters, for some λ , we have $\beta_s = \lambda^s \beta_0$ so that the effects of all past time-varying inputs decay at the same rate over time and their influence on current health status decrease in the specified manner as their distance from the present increases. This assumption is standard in the education literature (see Guarino et al. 2012). With this assumption, we can create a simpler estimating equation:

$$H_{it} = \alpha_t I_{it} + \lambda H_{it-1} + \varepsilon_{it} \tag{3}$$

This estimating equation now relates patient health outcomes at discharge (H_{it}) with the physician's investments (I_{it}) , adjusting for the patient's health at admission (H_{it-1}) . Note that health measurements at admission include the Rothman Index (defined below) and also diagnosis codes² and demographics.

After conditioning for their diagnosis at admission and other characteristics, α_t will indicate the association between the physician's investment and patient gains in health—their "value added". In practice, and following the education literature, we will

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² The diagnosis codes are from the clinical classifications software (CCS) developed by the Healthcare Cost and Utilization Project (HCUP) and is used to cluster patient diagnosis into a manageable number of clinically meaningful categories—collapsing the over 14,000 ICD diagnosis code into 285 mutually exclusive categories (i.e. Tuberculosis or Burns).

estimate equation (3) using a summary measure (Rothman Index) of patient health and estimate physician fixed effects (indicator variables for each physician in the data).³

Linking Value Added with Patient Outcomes

Taking our results from Equations 1-3 as inputs, we then ask what the effects are of being assigned a "high quality" versus a "low quality" attending physician on patients' health (e.g. Rothman Index) and hospitalization outcomes (e.g. readmission).

$$H_{it} = \rho \hat{I}_i + \pi H_{it-1} + \nu_{it} \tag{4}$$

Where the outcome of patient i at hospital discharge (or post-discharge) is a function of the estimated value added status of her physician (\hat{I}_i) from equation (3), her baseline health status (H_{it-1}) and health shocks that are unrelated to the physician (υ_{it}). The outcomes of interest will include the Rothman Index at discharge, the total costs of the hospitalization, the length of stay, the total costs per day in the hospital, and 30-day readmission for the patient. We used ordinary least squares (OLS) regression analysis for the discharge Rothman Index, total charges, and readmission outcomes and zero-truncated Poisson regression analysis for the length of stay outcome (Long and Freese 2006). The standard errors in this multivariable regression analysis were estimated using Huber/White sandwich estimator to account for clustering of patients by the attending physician (Williams 2000).

An implicit assumption underlying a causal interpretation of equation (4) is that the patients and doctors are not paired due to the patient's health at admission and the provider's value added. To test for any sorting, we will also estimate the relationship between the assigned physician's value added and the Rothman Index at hospital admission (baseline).

$$H_{it-1} = \psi \hat{I}_i + \nu_{it} \tag{5}$$

³ Like in the education literature, we are concerned that estimates from equation (3) could be confounded by hospital practices of sorting patients to physicians. After our controls are added for patient baseline health and diagnosis, our comparison will be over sets of patients who are identical with respect to their diagnosis and admission health status but are assigned to different physicians, and we will estimate the average "adjusted value added" of each physician in the data.

We detail evidence below that ψ =0 in our sample, conditional on diagnosis controls, which suggests a limited role of regular and purposeful sorting between physicians with high value added and the health of the patients to whom they are assigned

Data

We use patient data from all unique medical and surgical inpatients admitted to Yale-New Haven Hospital (YNHH) between July 1, 2011 and June 30, 2012. We obtained data from the hospital electronic medical record (Sunrise Clinical Manager, Allscripts) and from the hospital billing database. From the hospital billing database, we obtained dates of admission and discharge, patient age, sex, payer status (Medicare, Medicaid, other), admitting service (medicine, surgery), an indicator for admission into the intensive care unit (ICU), discharge diagnosis⁴, total hospital charges, and presence of any 30-day readmission. Patient outcomes included length of hospital stay, total charges, the Rothman Index on the day of discharge, and readmission to YNHH for any cause within 30 days of discharge.

From the electronic medical record, we obtained patients' health status as defined by the *Rothman Index*. The Rothman Index is a validated measure of health status (Rothman, Solinger et al. 2012, Bradley, Yakusheva et al. 2013, Rothman, Rothman et al. 2013, Rothman, Rothman et al. 2013, Tepas III, Rimar et al. 2013), which is calculated from vital signs, physical exam findings, laboratory results, the Braden scale for risk of skin breakdown (Bergstrom 1987), and nursing assessments of whether a minimum

⁴ Diagnostic categories were defined using the ICD-9 code, aggregated into one of 285 mutually exclusive categories using the clinical classifications software (CCS) developed by the Healthcare Cost and Utilization Project (HCUP 2012).

standard of function is met in 11 areas (cardiac, respiratory, gastrointestinal, genitourinary, neurological, skin and tissue, safety and fall risk, peripheral vascular, food and nutrition, psychosocial, musculoskeletal). Nursing assessments were recorded for each patient at least once per nurse shift between the time of admission and time of discharge. The Rothman Index is calculated and updated using novel proprietary software whenever new information is entered in the electronic medical record, such as when nurse assessments are entered, blood pressure is recorded, or laboratory results are reported. The Rothman Index has a range of (-91,100), with higher values indicating better health.

A total of 29,857 admissions occurred in the medicine and surgery services during both study periods. Of these, 9,982 were excluded because the admission was not the first admission in our data window⁶, and an additional 3,688 were excluded because they were admitted by physicians who cared for fewer than 10 patients in either the first or second time period. Thus, our final analysis sample included 16,187 patients, comprised of 8,979 patients treated by 163 attending physicians in the first six-month time period and 7,208 patients treated by 163 attending physicians in the second six-month time period.

Patient characteristics and outcomes are shown in **Table 1**. The mean value of the Rothman Index at admission was 76.1; the mean value at discharge was 79.6 and the mean change was 3.5 (standard deviation, 17.7). The mean number of patients seen by

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⁵ For example, the neurological standard was whether the patient was alert; oriented to person, place, time and situation; and speech was coherent.

⁶ To avoid bias due to dynamic assignment of patients to physicians (for example, if patients that had a bad experience on the first hospitalization elected different clinicians on a repeat hospitalization), we included only the first visit within the study period for each patient. In order to generate and then validate the value added measure, we split the data into two six-month periods. We excluded patients whose attending physician at discharge attended for fewer than 10 patients in each six-month period⁶.

the same attending physician was 46 in the first six months and 32 in the second sixmonth period.

Results

Our primary independent variable of interest was physician's value added, which we estimated from a multivariable regression in which the outcome was patient health status at the time of discharge (measured by the Rothman Index at discharge), adjusted for health status at the time of admission (measured by the Rothman Index at admission), a set of patient characteristics, and an indicator variable for each attending physician in our sample. For this analysis, we used the first six-month period of data. The multivariable regression coefficients on the physician indicator variables defined the estimated value added for each physician. We adjusted the standard errors using the Huber/White sandwich estimator to account for the clustering of patients by physician (Williams 2000).

Using equation (3), estimated value added varied from -0.6 to 0.9 on the Rothman Index (standard deviation 0.26) (**Table 2**). Together, physician indicator dummy variables (i.e., the value added estimates for unique physicians) were highly significant (p-value < 0.001). Additionally, value added estimates for the same physician in the first time period and the second time period were strongly correlated (spearman correlation coefficient r = 0.59, p-value < 0.001)⁷.

Physician ranking by value added scores was highly stable over time (**Table 3**).

Of the 41 physicians in the bottom quartile in time period one, nearly 50% were ranked in

⁷ For example, the teaching literature finds correlations of estimated value added across adjacent school years for the same teacher of between 0.3 and 0.7 (e.g. McCaffrey et al. 2009).

the bottom quartile in time period two, 83% were in the bottom two quartiles, and none was ranked in the top quartile. Similarly, of the 40 physicians in the top quartile in time period one, 65% were ranked in the top quartile in time period two, and 87.5% were in the top two quartiles (chi-square test p-value <0.001).

importantly, physician value added was not significantly associated with the Rothman Index upon admission in the first time period (p-value = 0.75) or in the second time period for patients assigned to the same physician (p-value = 0.82), indicating no evidence of more severely ill patients being directed to physicians with higher or lower value added estimates (selection bias based on observables).

We then estimated adjusted associations between physicians' value added determined in the first time period and their patients' outcomes in the second time period (hold-out sample), including length of stay, total charges, discharge Rothman Index, and same-hospital readmission for any cause in the 30 days following discharge. We adjusted for age, age-squared, sex, service, payer, admitting diagnosis, ICU stay, and the month of admission. For ease of interpretation, we discuss average outcomes in period two for physicians in the top 75th percentile of value added (based on period one) compared to those in the bottom 25th percentile of value added.

Patients of physicians in the 75th versus 25th percentile of the value added had, on average, shorter length of stay (4.76 vs 5.08 days, p<0.001), lower total costs (\$17,811 vs \$19,822, p<0.001) and higher discharge Rothman Index scores (80.28 vs 78.92, p<0.001) for patients in the second six-month period (**Table 4**). The estimated value added was not significantly associated with likelihood of readmission within 30 days (p-value=0.78).

Discussion

Our findings provide evidence to support a new method of determining the value added by individual physicians to outcomes of hospitalized patients. Though patient outcomes necessarily also reflect care provided by other team members (e.g., physicians in training, nurses, social workers, physical therapists), we were still able to demonstrate that the value added varies substantially across attending physicians and are highly stable for individual physicians across six-month time periods.

The measure of physician value added was strongly associated with key patient outcomes including length of stay, costs, and health status at discharge, providing external validity to the concept of value added as a novel measure of physician quality. Although the effect sizes were modest, they are meaningful. Specifically, a standard deviation increase in attending physician value added was associated with a reduction in total hospital charges of approximately 6% of a standard deviation, an increase in the discharge Rothman Index of approximately 8% of a standard deviation, a reduction in length of stay of approximately 4% of a standard deviation, but was not significantly associated with the likelihood of 30-day readmission to the same hospital. Understanding the specific mechanisms underlying these associations could lead to improved methods of evaluating the differing influences of physicians on key hospital outcomes.

The value added approach to determining attending physician contributions to patient outcomes has numerous advantages. The method allows physicians' added value to be estimated on the basis of every patient's health status rather than on the basis of relatively rare events such as readmission or mortality, and is therefore less sensitive to sampling variation from a small numbers of patients per physician. Value added can be

calculated based on any measure of health status available at defined care intervals such as admission and discharge: for example, the Modified Early Warning Score (MEWS)(Subbe, Kruger et al. 2001). Physicians' value added could also be calculated based on changes in patient-reported outcomes over time, such as quality of life, symptom burden or global health. Given that hospitals are rapidly adopting electronic health records (Desroches, Charles et al. 2013), basing quality measurement on measures of health status automatically extracted from the medical record, such as the Rothman Index (Rothman, Solinger et al. 2012, Bradley, Yakusheva et al. 2013, Rothman, Rothman et al. 2013, Rothman, Rothman et al. 2013, Tepas III, Rimar et al. 2013, Rothman Forthcoming), is increasingly feasible (Amarasingham, Moore et al. 2010, Escobar, LaGuardia et al. 2012, Johnson, Kramer et al. 2013, Liu, Turk et al. 2013, Wu, Kharrazi et al. 2013).

Our findings should be interpreted in light of several limitations. First, the value added approach attributes all systematic changes in patients' health status between admission and discharge to the assigned attending physician at discharge even though patients have a team of health care providers, potentially including multiple attending physicians, who may contribute to value added. We note that this limitation is common to all measures of physician care quality, and is likely to reduce the signal of the individual physician because other team members are often shared among attending physicians. Hence, like other measures in the literature, our measure of value added should be interpreted broadly, recognizing that such value may be also attributed to the coordination and care efforts of nurses and other staff members who regularly work with the physicians. Future research should attempt to further separate these effects into the

value added of physicians, nurses and other staff members, much as the education literature has attempted to estimate value added of principals, teachers, and other staff members (Eberts and Stone 1988, Brewer 1993, Coelli and Green 2012, Grissom, Kalogrides et al. 2012). Second, the findings may be confounded by selection effects in which more patients with more severe illness are assigned to certain physicians in ways that are not adequately accounted for by our health status measure. To assess this potential source of bias, we examined the association between the estimated value added from the first six-month study period and the health status at admission for the same physician in the second six-month study period, which we found to be non-significant, suggesting no evidence of this bias. Nevertheless, it is possible that heterogeneity across patients could confound the analyses.

Several opportunities exist to enhance and extend our approach. Our methods can, in principle, be scaled up by merging data from multiple hospitals. This step would allow novel cross-organizational comparisons that can be used to begin to evaluate differences in hospital practices that may be linked with physician and nurse performance in increasing patient health during hospitalization. Furthermore, hospitals (or consortia of hospitals) could use these value added methods to evaluate the impacts of specific hospital policies and professional development programs in enhancing physician and nurse value added. For example, a hospital could examine whether estimates of physicians' value added increase following the completion of a professional development program. Last, hospital and medical groups could begin to consider tying reimbursement and retention decisions more explicitly to physician and nurse value added, particularly if future studies establish links with longer-term patient outcomes.

Tables

Characteristic	Study Cohort	Time Period 1	Time Period 2	
	(N=16,187)	(N=8,979)	(N=7,208)	
Age in years				
Mean (SD)	59 (19)	59 (19)	59 (18)	
Range	18, 104	18, 104	18,102	
Female—N (%)	8,076 (50)	4,486 (50)	3,590 (50)	
Primary insurance—N (%)				
Medicare	7,270 (45)	4,116 (46)	3,154 (44)	
Medicaid	3,054 (19)	1,737 (19)	1,317 (18)	
Other	5,862 (36)	3,126 (35)	2,737 (38)	
Medicine service—N (%)	10,448 (65)	5,939 (66)	4,509 (63)	
Any intensive care unit stay—N (%)	2,814 (17)	1,530 (17)	1,284 (18)	
Admission Rothman Index				
Mean (SD)	76.1 (18.9)	75.9 (18.9)	76.4 (18.8)	
Range	-36, 100	-30,100	-36,100	
Discharge Rothman Index				
Mean (SD)	79.6 (17.0)	79 (17.3)	80.4 (16.5)	
Range	-47, 100	-38,100	-47,100	
Change in Rothman Index				
Mean (SD)	3.5 (17.7)	3.2 (14.6)	4.0 (14.8)	
Range	-102, 94	-102, 94	-102, 90	
Same hospital 30-day readmission—N (%)	1619 (10)	899 (11)	657 (8)	
Length of stay—days				
Mean (SD)	4.9 (7.7)	5.1 (8.3)	4.7 (7.0)	
Range	1, 266	1,266	1,166	
Total charges (\$) per admission				
Mean (SD)	\$18,816 (31,984)	18,566(32,753)	19,128(31,000)	
Range	\$593-\$947,113	\$941-\$947,113	\$593-\$818,353	

Notes: RI: Rothman Index. Distribution of Patient Diagnosis Codes available from the authors.

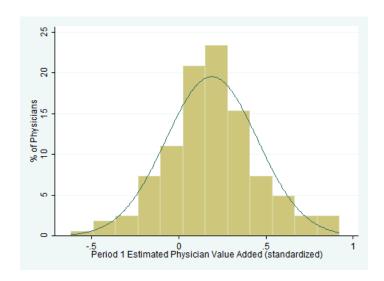


Table 2. Distribution of Estimated Physician Value Added, First Six-Months of Data.

Estimates based on 8,979 patients assigned to 163 physicians.

Table 3: Consistency of Estimated Physician Value Added in Time Period 1 and 2						
		Time Period 2				
		Quartile 1	Quartile 2	Quartile 3	Quartile 4	Total
Time Period 1	Quartile 1	19	15	7	0	41
	Quartile 2	10	16	14	1	41
	Quartile 3	8	9	11	13	41
	Quartile 4	4	1	9	26	40
		41	41	41	40	163

Notes: This figure is a transition matrix of the quartiles of estimated value added for the 163 attending physicians in the sample between the first 6 months of data and the second six months of data.

Table 4. Estimated Associations between Physician Value Added Determined in Time Period One						
and Patient Outcomes in Time Period Two (N=7,208)						
	Discharge Rothman	Length of	Total	30 Day		
Outcome:	Index (SD)	Stay	charges	Readmission		
Physician Value Added (SD)	0.080**	-0.110**	-2,011**	-0.002		
	(0.000)	(0.004)	(0.009)	(0.787)		
	, ,	[-0.315]	, ,	` ,		
R-squared	0.607		0.612			

Notes: Robust p-values are shown in parentheses and are clustered on the discharging physician identification number. Coefficients are presented in Row 1. Marginal effects are listed in brackets for the length of stay outcome. All results use OLS regression analysis except length of stay, which uses zero-truncated Poisson regression analysis. See Table A1 for list of control variables and full results. **1% statistical significance. SD: standard deviation

Appendix Table

Outcome	Admission RI	Discharge RI	Length of Stay	Total Charges	Readmission
Physician Value Added (Std)	-0.004	0.080***	-0.104***	-2,010.774***	-0.002
	(0.860)	(0.000)	(0.000)	(0.009)	(0.783)
Admission RI		0.483***	-0.110***	-1,765.040**	-0.011*
		(0.000)	(0.004)	(0.018)	(0.090)
Age (years)	0.014***	0.013***	0.003	173.010**	0.001
	(0.000)	(0.000)	(0.585)	(0.040)	(0.290)
Age Squared	-0.000***	-0.000***	-0.000	-1.524**	-0.000
	(0.000)	(0.000)	(0.868)	(0.043)	(0.403)
Female	-0.038*	-0.077***	0.031	243.910	-0.009
	(0.055)	(0.000)	(0.292)	(0.656)	(0.339)
Medicare	-0.063**	-0.114***	0.108**	1,014.615	0.007
	(0.015)	(0.000)	(0.015)	(0.190)	(0.492)
Medicaid	-0.073***	-0.041*	0.219***	1,778.195*	-0.001
	(0.004)	(0.084)	(0.000)	(0.050)	(0.924)
Medical Division	-0.125**	0.088*	0.078	-231.860	-0.034
	(0.025)	(0.078)	(0.411)	(0.898)	(0.168)
ICU Indicator	-0.827***	0.009	0.431***	12,205.333***	0.007
	(0.000)	(0.851)	(0.000)	(0.000)	(0.676)
Month = January	-0.024	-0.011	0.034	883.559	0.022
•	(0.436)	(0.618)	(0.451)	(0.270)	(0.110)
Month = February	-0.029	-0.056**	0.039	1,517.023*	-0.001
j	(0.319)	(0.023)	(0.487)	(0.064)	(0.946)
Month = March	0.013	-0.022	0.023	1,417.343*	0.015
	(0.704)	(0.387)	(0.624)	(0.055)	(0.273)
Month = April	-0.017	0.010	0.033	1,199.042	
1	(0.605)	(0.710)	(0.521)	(0.137)	
Month = May	-0.036	-0.056**	0.053	1,464.790	0.005
-	(0.278)	(0.026)	(0.310)	(0.112)	(0.702)
Month = June (Omitted)					
Constant	0.291***	0.014	0.290**	10,882.341***	0.065*
	(0.003)	(0.856)	(0.018)	(2,555.586)	(0.088)
Observations	7,208	7,208	7,208	7,208	5,994
R-squared	0.508	0.607		0.612	0.131

Notes: Robust p-values are shown in parentheses and are clustered on the discharging physician identification number. Coefficients are presented. All results use OLS regression analysis except length of stay, which uses zero-truncated Poisson regression analysis. The sample size for the readmission outcome is smaller because patients who were admitted in the final month of the time period were not used in the analysis due to the right censoring of the outcome. ***1%, **5%, *10% statistical significance. Additional Controls: Principal diagnosis indicator variables.

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