HAVE NEWER CARDIOVASCULAR DRUGS REDUCED HOSPITALIZATION?
EVIDENCE FROM LONGITUDINAL COUNTRY-LEVEL DATA ON 20 OECD
COUNTRIES, 1995-2003

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<u>ABSTRACT</u>

This study examines the effect of changes in the vintage distribution of cardiovascular system drugs on hospitalization and mortality due to cardiovascular disease using longitudinal country-level data. The vintage of a drug is the first year in which it was marketed anywhere in the world. We use annual data on the utilization of over 1100 cardiovascular drugs (active ingredients) in 20 OECD countries during the period 1995-2003.

Countries with larger increases in the share of cardiovascular drug doses that contained post-1990 or post-1995 ingredients had smaller increases in the cardiovascular disease hospital discharge rate, controlling for the quantity of cardiovascular medications consumed per person, the use of other medical innovations (CT scanners & MRI units), potential risk factors (average consumption of calories, tobacco, and alcohol), and demographic variables (population size & age structure, income, and educational attainment). The estimates also indicate that use of newer cardiovascular drugs has reduced average length of stay and the age-adjusted cardiovascular mortality rate, but not the number of potential years of life lost due to cardiovascular disease before age 70 per 100,000 population.

The estimates indicate that if drug vintage had not increased during 1995-2004, hospitalization and mortality would have been higher in 2004. We estimate that per capita expenditure on cardiovascular hospital stays would have been 70% ($89) higher in 2004 had drug vintage not increased during 1995-2004. Per capita expenditure on cardiovascular drugs would have been lower in 2004 had drug vintage not increased during 1995-2004. But our estimate of the increase in expenditure on cardiovascular hospital stays is about 3.7 times as large as our estimate of the reduction in per capita expenditure for cardiovascular drugs that would have occurred ($24).

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New drugs tend to be more expensive than old drugs. Many people might therefore expect greater use of new drugs to increase total medical expenditure. This is not necessarily the case, however. If new drugs are more effective than older drugs, use of newer drugs might reduce other medical (e.g. hospital) expenditure more than it increases drug expenditure.

A number of previous studies have investigated the hypothesis that use of newer drugs has reduced other medical expenditure. Some of these studies have investigated whether new drugs *in general*—i.e., drugs in all therapeutic classes and for all diseases—have reduced other medical expenditure. These studies have used a variety of methodologies and kinds of data. Lichtenberg (2001) examined condition-by-patient-level data from the 1996 Medical Expenditure Panel Survey. Lichtenberg (2006a) examined longitudinal state-level data from the period 1997-2003, and Lichtenberg (2007a) examined longitudinal disease-level data from the period 1990-2003. All three of these studies, which were based on U.S. data, provided support for the hypothesis that, overall, using newer drugs has reduced other types of medical expenditure, particularly hospital expenditure.

Other studies have examined specific classes of drugs. Miller et al (2005) investigated whether use of newer cardiovascular drugs reduced other medical expenditure. They concluded that there was no relationship between use of newer drugs and non-drug expenditure. However Lichtenberg (2006b) argued that their analysis had some serious flaws, and that their conclusions should be viewed with considerable skepticism. In particular, they controlled for factors that should not have been controlled for, and they controlled poorly for factors that should have been controlled for.

Weisfeldt and Zieman (2007) surveyed evidence from clinical trials about drug and device efficacy in the prevention and treatment of coronary artery disease and heart failure during the period 1970-2010. They concluded that pharmaceutical agents played a major role in prevention of atheroscierosis and its consequences (heart attack, stroke, and heart failure), and that the marked reduction in cardiovascular disease and its consequences was largely driven by the development and implementation of drugs for long-term use and by complicated and costly procedures and operations for acute disease management. Two other lessons that emerged from their survey of the clinical literature

are worth noting here. First, the introduction of new drugs within a class may improve the health of the population, even if newer agents do not strictly dominate older agents (i.e., they are not superior for every patient), because not all members of the population benefit similarly from the use of specific agents in a particular drug class.[1] Second, evidence from clinical trials showing the efficacy of a drug may first be published many years after the class was launched. For example, according to Weisfeldt and Zieman, evidence from clinical trials showing the efficacy of aldosterone antagonists for preventing and treating coronary heart disease was first published in 1999, 39 years after the first drug in the class (spironolactone) was approved by the FDA.[2]

This study will examine the effect of changes in the vintage distribution of cardiovascular system drugs on hospitalization and mortality due to cardiovascular disease using longitudinal country-level data. The vintage of a drug is defined as the year in which the drug was first sold anywhere in the world. The concept of vintage has hardly been used in health economics, but it is well established in both theoretical and empirical analysis in other areas of economics.[3]

We will use annual data on the utilization of over 1100 cardiovascular drugs (active ingredients) in 20 OECD countries during the period 1995-2004. Cardiovascular drugs comprise the largest therapeutic class of drugs, in terms of expenditure, in 13 key global markets IMS Health (2006). Heart disease was the most frequent cause of hospitalization in the U.S. in 1999 (Popovic and Hall (2001)).

Section I of this paper describes the econometric models that we will estimate. Data sources and descriptive statistics are discussed in Section II. Empirical results are

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[1] Thus, the introduction of new statins might have benefited patients even though Zhou et al (2006) concluded from published statin randomized placebo-controlled trials that pravastatin, simvastatin, and atorvastatin, when used at their standard dosages, showed no statistically significant difference in their effect on long-term cardiovascular prevention. Moreover Cannon et al (2004) found that patients taking 80 milligrams of atorvastatin (launched in January 1997) were 16% less likely to die, have a heart attack, undergo bypass or angioplasty surgery, or develop worsening chest pain than those taking 40 milligrams of pravastatin (launched in October 1989).

[2] The class includes just one other drug: eplerenone, which was approved by the FDA in 2002.

[3] See Boucekkine et al (2006) for a brief survey of the theoretical literature on vintage capital. Empirical studies by Hulten (1992), Bahk and Gort (1993), and Sakellaris and Wilson (2004) have demonstrated that the vintage of capital equipment has an important effect on productivity in manufacturing.

presented in Section III. Implications of the estimates are considered in Section IV. The final section presents a summary.

I. Methodology

We will estimate models of the following form, using longitudinal country-level data:

$$\ln Y_{it} = \beta \, rx\_vint_{it} + \gamma \, Z_{it} + \alpha_i + \delta_t + \varepsilon_{it} \qquad (1)$$

where

Y_{it} = a measure of hospital use or mortality due to cardiovascular disease in country i (i = 1,…,21) in year t (t = 1995,…,2003)

rx_vint_{it} = a measure of the vintage distribution of cardiovascular system (ATC class C) drugs used in country i in year t

Z_{it} = other potential determinants of hospital use or mortality due to cardiovascular disease in country i in year t

α_i = a fixed effect for country i

δ_t = a fixed effect for year t

Since the model includes country and year fixed effects, it is a difference-in-differences model. Negative and significant estimates of β would indicate that, *ceteris paribus*, countries with above-average increases in cardiovascular drug vintage had above-average declines (or below-average increases) in cardiovascular hospital use and mortality.

This model is static: it is predicated on the assumption that hospital use and mortality depend on *contemporaneous* values of drug vintage and other variables. In reality, health outcomes might depend on *lagged* as well as contemporaneous values of the explanatory variables.[4] But there are two practical reasons to estimate a static model. First, the time-series is relatively short, since 1995 is the first year for which the drug data are available. Including lagged vintage would therefore significantly reduce the number of sample observations. Second, many of the explanatory variables exhibit high serial correlation. This makes precise identification of the lag structure challenging, and would cause multicollinearity if lagged as well as current values of regressors were included.

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[4] Weisfeldt and Zieman (2007) observe that, unlike therapies used in infectious diseases, cardiovascular agents offer no rapid "cure." Rather, they prevent or reduce the progression of disease when ingested continuously.

Dependent variables

We will estimate models of the following measures of hospital use or mortality due to cardiovascular disease:

- The number of hospital discharges[5] in which the principal diagnosis was cardiovascular disease per 100,000 population (n_discharges)
- Average length of stay[6] in hospital discharges in which the principal diagnosis was cardiovascular disease (alos)
- The number of hospital days in which the principal diagnosis was cardiovascular disease per 100,000 population (n_days = n_discharges * alos)
- The (age-standardized) number of deaths caused by diseases of the circulatory system[7] per 100,000 population (n_deaths)
- Potential years of life lost due to diseases of the circulatory system before age 70 per 100,000 population[8] (pyll)

Measures of the vintage distribution of cardiovascular system drugs

We will use two different measures of the vintage distribution of cardiovascular system drugs. Both are based on the following general definition of drug vintage:

$$rx\_vint_{it} = \Sigma_d \, n\_units_{dit} \, f(launch\_year_d) \, / \, \Sigma_d \, n\_units_{dit}$$

where

n_units_{dit} = the number of standard units[9] of cardiovascular drug d sold in country i in year t

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[5] Discharge is the formal release of an in-patient from an acute care institution after a period of "hospitalization". It includes deaths in hospitals, but excludes same-day separations and transfers to other care units within the same institution. However, the following countries include at least some same-day separations: Austria, Czech Republic, France, Hungary, New Zealand and the United States.

[6] Average length of stay (ALOS) is calculated by dividing the number of days stayed (from the date of admission in an in-patient institution) by the number of discharges (including deaths).

[7] The age-standardized number of deaths is calculated as the population times the age-standardized death rate. The latter is calculated by the OECD Secretariat, using the total OECD population for 1980 as the reference population. The age-standardized death rates are necessary for comparing the level of mortality across countries and over time since they take into account the differences in age structure of the populations.

[8] Potential years of life lost (PYLL) is a summary measure of premature mortality which provides an explicit way of weighting deaths occurring at younger ages, which are, a priori, preventable. The calculation of PYLL involves summing up deaths occurring at each age and multiplying this with the number of remaining years to live up to a selected age limit. The limit of 70 years has been chosen for the calculations in *OECD Health Data*. In order to assure cross- country and trend comparison, the PYLL are standardized for each country *i* and each year *t*.

[9] A standard unit is equivalent to a standard dose of medication (Aldonas (2005)). Standard Units (SU) are determined by taking the number of counting units (e.g. actual number of ml in a bottle) and dividing it by

launch_year$_d$ = the year of initial world launch of cardiovascular drug d

The two measures are based on two different forms of the function f(launch_year$_d$):

$$f_1(\text{launch\_year}_d) \quad = 1 \text{ if launch\_year}_d > 1995$$
$$= 0 \text{ if launch\_year}_d \leq 1995$$

$$f_2(\text{launch\_year}_d) \quad = 1 \text{ if launch\_year}_d > 1990$$
$$= 0 \text{ if launch\_year}_d \leq 1990$$

The two corresponding measures of cardiovascular drug vintage are:

rx_post1995%$_{it}$ = Σ_d n_units$_{dit}$ f$_1$(launch_year$_d$) / Σ_d n_units$_{dit}$
= the fraction of standard units whose active ingredients
were launched after 1995

rx_post1990%$_{it}$ = Σ_d n_units$_{dit}$ f$_2$(launch_year$_d$) / Σ_d n_units$_{dit}$
= the fraction of standard units whose active ingredients
were launched after 1990

Other potential determinants of hospital use or mortality due to cardiovascular disease

In addition to cardiovascular drug vintage, the models of cardiovascular hospitalization and mortality we estimate will include several other types of explanatory variables: the *quantity* of cardiovascular medications consumed per person; indicators of the use of other medical innovations (CT scanners & MRI units); indicators of potentially important cardiovascular risk factor (calories consumed per person per day and per capita expenditure on tobacco and alcohol); and demographic variables (population size & age structure, income, and educational attainment). In particular, the models will include the following covariates:

ln(n_cv_su$_{it}$) = the log of the number of standard units of cardiovascular drugs per person
ln(ct_scan$_{it}$) = the log of the number of computed tomography scanners per million population
ln(mri$_{it}$) = the log of the number of magnetic resonance imaging units per million population

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the smallest common dose of a product as defined by IMS Health. For example, oral solids have an SU of one tablet or capsule, whereas for syrup forms it is 5ml and for injectable forms it is one ampoule or vial.

$\ln(\text{calories}_{it})$ = the log of the average number of calories consumed per person per day[10]

$\ln(\text{tobacco}_{it})$ = the log of per capita expenditure on tobacco (in PPP US dollars)

$\ln(\text{alcohol}_{it})$ = the log of per capita expenditure on alcohol (in PPP US dollars)

$\ln(\text{pop}_{it})$ = the log of the population

$\text{age\_ge\_65\%}_{it}$ = the fraction of the population age 65 and over

$\ln(\text{gdp}_{it})$ = the log of per capita GDP

$\text{edu\_low\%}_{it}$ = the fraction of the population with education at or below the lower secondary level (ISCED[11] 0/1/2)

$\text{edu\_mid\%}_{it}$ = the fraction of the population with education at the upper secondary or post-secondary, non-tertiary level (ISCED 3/4)

$\text{edu\_max\%}_{it}$ = the fraction of the population with education at the tertiary level (ISCED 5A/6)[12]

Cardiovascular disease prevalence. Ideally, we could control directly for the *prevalence* of cardiovascular disease, i.e. the fraction of the population at risk for cardiovascular hospitalization or death. Unfortunately, reliable data on the prevalence of cardiovascular disease, by country and year, are not available.[13] The cardiovascular drug utilization rate (n_cv_su) may control, to some extent, for the cardiovascular disease prevalence rate, but it also reflects intensity of use among people with the disease:

$$\text{n\_cv\_su} = \frac{\text{su}}{\text{prev}} \times \frac{\text{prev}}{\text{pop}}$$

An increase in the cardiovascular disease prevalence rate (prev / pop) would increase cardiovascular hospitalization and mortality, but an increase in treatment intensity (su / prev) might be expected to reduce cardiovascular hospitalization and mortality. Hence, the net effect of an increase in per capita consumption of cardiovascular drugs on cardiovascular hospitalization and mortality is theoretically ambiguous. As indicated

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[10] The main data source is the Food and Agriculture Organization of the United Nations Nutrition database (FAOSTAT). See http://faostat.fao.org/.

[11] ISCED denotes the International Standard Classification of Education from 1997. ISCED 5B (programs at the tertiary level that focus on practical, technical or occupational skills for direct entry into the labor market) is the omitted category.

[12] The omitted education category is the fraction of the population with education at the ISCED 5B level. These programs are typically shorter than those of ISCED 5A and focus on practical, technical or occupational skills for direct entry into the labor market, although some theoretical foundations may be covered in the respective programs. They have a minimum duration of two years full-time equivalent at the tertiary level.

[13] HIV/AIDS and cancer are the only chronic diseases for which there are reasonably reliable longitudinal estimates of disease prevalence, due to the existence of disease registries. Lichtenberg (2006c, 2007b) uses these data to investigate the effects of pharmaceutical use and innovation on the survival rates of HIV/AIDS and cancer patients in the U.S.

above, we also control for per capita income and educational attainment, which may influence both disease prevalence and the quality of disease management, hence the rate of cardiovascular hospitalization and mortality.

Cardiovascular risk factors. Instead of, or in addition to, controlling for cardiovascular disease prevalence, it is desirable to control for cardiovascular disease risk factors. Advanced age, obesity and smoking are known to be cardiovascular disease risk factors. Our models of hospital utilization will control for the fraction of the population age 65 and over.[14] The *OECD Health Data 2006* database contains some information about the prevalence of obesity and tobacco use. Unfortunately, the data on these indicators are quite incomplete. For example, data on obesity are available for only 37% (59 out of 159) of the country-year observations in our sample; data on tobacco use were available for 57% (91 out of 159) of the observations. Including a measure of obesity in the model would reduce the number of residual degrees of freedom by 83% (from 111 to 19). Therefore the basic models we estimate will not include measures of obesity. However we will control for an important (and frequently reported) determinant of obesity: calories consumed per person per day. We will also control for per capita expenditure (in PPP US dollars) on tobacco and on alcohol; data on these variables were obtained from the *OECD Annual National Accounts* database.

Other medical innovation. Weisfeldt and Zieman (2007) found that the marked reduction in cardiovascular disease and its consequences was largely driven by complicated and costly procedures and operations for acute disease management as well as by the development and implementation of drugs for long-term use. It would therefore be ideal to control for the rate of adoption of non-pharmaceutical as well as pharmaceutical medical innovations relevant to cardiovascular disease.

The *OECD Health Data 2006* database contains five indicators of non-pharmaceutical medical technology. Data on all of these indicators are quite incomplete. We will control for the use of two medical technologies: CT scanners and MRI units. These are the two technologies with the least incomplete data, and that are most relevant

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[14] Our models of cardiovascular mortality will not control for the fraction of the population age 65 and over, because the mortality measures are age-adjusted.

to the diagnosis and treatment of cardiovascular disease.[15] Over 73,000 articles in MEDLINE refer to both "Cardiovascular Diseases" and either "Tomography, X-Ray Computed" or "Magnetic Resonance Imaging".[16] However CT scanners and MRI units are also used to diagnose many other diseases. These measures may not adequately control for non-pharmaceutical cardiovascular disease innovation.

If pharmaceutical and non-pharmaceutical cardiovascular disease innovation are "complements" (i.e., they are positively correlated across countries), estimates of β could be biased away from zero. On the other hand, if pharmaceutical and non-pharmaceutical innovation are "substitutes" (i.e., they are negatively correlated across countries), estimates of β could be biased towards zero. It is not possible to determine the sign of the correlation between pharmaceutical and non-pharmaceutical cardiovascular disease innovation across *countries*, but it is possible to determine the sign of this correlation across *U.S. states*. Using data on the two largest public health insurance programs, we constructed the following indicators of the adoption of pharmaceutical and non-pharmaceutical (i.e. surgical) cardiovascular disease innovations by state in 2004:

$rx\_post1995\%_{s,2004}$ = the fraction of Medicaid[17] cardiovascular drug prescriptions in state s in 2004 whose active ingredients were approved by the FDA after 1995

$rx\_post1990\%_{s,2004}$ = the fraction of Medicaid cardiovascular drug prescriptions in state s in 2004 whose active ingredients were approved by the FDA after 1990

$surg\_post1995\%_{s,2004}$ = the fraction of Medicare major cardiovascular surgical procedures[18] in state s in 2004 with procedure codes established by the American Medical Association after 1995

$surg\_post1990\%_{s,2004}$ = the fraction of Medicare major cardiovascular surgical procedures in state s in 2004 with

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[15] The other three technologies are radiation therapy equipment, lithotriptors, and mammographs.

[16] See, for example, Harrigan et al (2008) and Schmid et al (2008).

[17] Medicaid is a health insurance program available only to certain low-income individuals and families who fit into an eligibility group that is recognized by federal and state law. There were about 82 million Medicaid cardiovascular drug prescriptions in 2004.

[18] Medicare is a health insurance program for people age 65 or older, people under age 65 with certain disabilities, and people of all ages with End-Stage Renal Disease (permanent kidney failure requiring dialysis or a kidney transplant). There were about 4 million Medicare major cardiovascular procedures performed in 2004. Major cardiovascular surgical procedures are procedures with Berenson-Eggers Type of Service (BETOS) codes P2A through P2F, i.e CABG, aneurysm repair, thromboendarterectomy, coronary angioplasty (PTCA), pacemaker insertion, and other major cardiovascular procedures. See http://www.cms.hhs.gov/hcpcsreleasecodesets/20_betos.asp.

procedure codes established by the American
Medical Association after 1990

The data used to construct these measures were obtained from a number of sources.[19] As discussed in Lichtenberg (2007b), the surgery innovation measures are probably less reliable than the drug innovation measures, because FDA approval of a drug is a more meaningful indicator than AMA establishment of a new surgical procedure code. In particular, some procedures with "new" procedure codes may be just relabeled or reclassified old procedures, rather than true innovations.

The following table shows the correlations across states (N = 50) between the pharmaceutical and surgical innovation measures, both unweighted and weighted by the number of prescriptions (n_rx) or the number of major surgical procedures (n_proc):

	unweighted	weight=n_rx	weight=n_proc
corr(rx_post1990, surg_post1990)	-0.129	-0.213	-0.180
p-value	0.378	0.142	0.217
corr(rx_post1995, surg_post1995)	-0.229	-0.473	-0.257
p-value	0.113	0.001	0.075

All six correlation coefficients are negative, and one is highly significant. This suggests that, within the U.S., pharmaceutical and non-pharmaceutical cardiovascular disease innovation may be substitutes rather than complements. Therefore, failure to control adequately for non-pharmaceutical cardiovascular disease innovation is more likely to bias estimates of β towards zero than away from zero.

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[19] State-level data on Medicaid cardiovascular drug prescriptions were obtained from Medicaid State Drug Utilization Files (http://www.cms.hhs.gov/MedicaidDrugRebateProgram/SDUD/List.asp). Data on FDA approval dates were obtained from the Drugs@FDA database (http://www.fda.gov/cder/drugsatfda/datafiles/default.htm). State-level data on Medicare major cardiovascular surgical procedures were obtained from the Physician/Supplier Procedure Summary Master File (http://www.cms.hhs.gov/NonIdentifiableDataFiles/06_PhysicianSupplierProcedureSummaryMasterFile.asp). Data on the dates at which procedure codes were established by the American Medical Association were obtained from the AMA's *Current Procedural Terminology Assistant Archives* (https://catalog.ama-assn.org/Catalog/product/product_detail.jsp?childName=&parentCategoryName=&parentCategory=&productId=prod780015&categoryName=&prodId=&start=&parentId=).

II. Data sources and descriptive statistics

Data on cardiovascular drug use (number of standard units), by pharmaceutical subclass, active ingredient, country, and year, were provided by IMS Health. There are over 1100 ingredients in 29 subclasses in ATC C. About 1.5 trillion standard units were sold in all classes combined in the 21 countries during 1995-2005. Table 1 shows illustrative data, by ingredient, for one subclass (C10A cholesterol & triglyceride regulators) in one country (the U.S.) in one year (2005).

The 'launch date' is the date when a product was first made available to the market (although there are circumstances when the launch date is considered to be the date when the first pack was included in the IMS data, or first accrued sales). The world launch date of some ingredients is unknown. Many of these ingredients are naturally occurring. The fraction of standard units with missing launch dates declined from 12% in 1995 to 6% in 2005.

Two features of the data are worth noting. First, a given ingredient may appear in several subclasses. Second, a given product may contain several ingredients. To avoid giving too much weight to combination products when calculating mean vintage, we used a two-step procedure to calculate mean vintage. First, we used the ingredient-level data to calculate mean vintage in each subclass, country, and year. Then, we used the subclass-level data to calculate mean vintage in each country and year, weighting by the number of sub-class SUs (*not* the sum of the number of ingredient SUs within the sub-class).

The mean vintage of cardiovascular drugs varies considerably across countries. Figure 1 shows post-1995 SUs as a percentage of total SUs in 2004, by country. In the top 5 countries, at least 13% of SUs were post-1995 SUs. In the bottom 4, only 1% of SUs were post-1995 SUs. It is not surprising that these are the emerging economies of Eastern Europe. However there are also sharp differences in cardiovascular drug vintage between countries that are otherwise similar. Switzerland's post-1995 share (11%) is more than twice as high as Austria's (5%). Australia's post-1995 share (20%) is four times as large as New Zealand's (5%). Table 2 further documents these differences by

showing the top 10 (ranked by number of SUs) cardiovascular drugs in these four countries in 2005. Descriptive statistics for all variables are shown in Table 3.

III. Empirical results

Estimates of eq. (1) based on the first measure of cardiovascular drug vintage (rx_post1995%, the fraction of standard units whose active ingredients were launched after 1995) are shown in Table 4. Estimates based on the second measure of cardiovascular drug vintage (rx_post1990%) are shown in Table 5.

First we will discuss the estimates shown in Table 4. In model 1, the dependent variable is the log of the number of cardiovascular disease hospital discharges per 100,000 population. The rx_post1995% coefficient is negative and highly significant (p-value < .0001). This signifies that countries with larger increases in the share of cardiovascular drug SUs that were post-1995 SUs had smaller increases in the cardiovascular disease hospital discharge rate, *ceteris paribus*.

The coefficient on per capita cardiovascular drug use is not significantly different from zero. As noted above, the net effect of an increase in per capita consumption of cardiovascular drugs on cardiovascular hospitalization is theoretically ambiguous.

Coefficients on four of the covariates are statistically significant (p-value < .05). The coefficient on age_ge_65% is positive and significant, as one would expect: cardiovascular hospital utilization increases when the fraction of the population that is elderly rises. The coefficient on the log of the number of CT scanners per million population is also positive and significant. This is consistent with the view that some diagnostic innovations may increase utilization of hospital care.[20] Increased consumption of tobacco is also positively associated with cardiovascular hospital utilization, but calorie consumption is inversely related to cardiovascular hospital utilization.

The dependent variable in model 2 is the log of average length of cardiovascular disease hospital stays. Once again, the rx_post1995% coefficient is negative and significant (p-value = .03). This implies that using newer drugs has reduced average

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[20] The Kaiser Family Foundation (2007), citing Rettig (1994), says that "advances in medical technology have contributed to rising overall U.S. health care spending."

length of stay. The coefficient on GDP is also negative and significant: average length of stay has declined faster in countries with higher growth in per capita GDP. Length of stay is positively correlated with calorie and alcohol consumption, but negatively correlated with tobacco consumption. The fact that the coefficient on edu_max% is positive and significant suggests that ALOS has declined more slowly in countries where the fraction of the population that is highly educated has increased the most.[21] However the implied effect of education on ALOS is quite small: if the entire population shifted from the low education to the maximum education category, ALOS would increase by less than 4%.

The dependent variable in model 3 is the log of the number of hospital days in which the principal diagnosis was cardiovascular disease per 100,000 population. The coefficients in model 3 are equal to the sums of the coefficients in models 1 and 2. Once again, the rx_post1995% coefficient is negative and highly significant. The only other variables whose coefficients are significant are ln(ct_scan), age_ge_65%, and edu_max%.

The dependent variable in model 4 is the (age-standardized) number of deaths caused by diseases of the circulatory system per 100,000 population. The drug vintage coefficient is also negative and highly significant (p-value = .0059) in this model. This implies that using newer cardiovascular drugs has reduced the age-adjusted cardiovascular disease mortality rate.   The coefficient on per capita drug consumption is positive and significant. Presumably this is because increased per capita cardiovascular drug consumption is indicative of increased cardiovascular disease prevalence. The coefficient on ln(calories) is positive and significant at the 10% level. This may reflect the effect of obesity on cardiovascular mortality. The education coefficients are significant, but the implied effect of changes in the educational attainment distribution on cardiovascular mortality is quite small.

The dependent variable in model 5 is the log of the number of potential years of life lost due to cardiovascular disease before age 70 per 100,000 population. This measure gives a great deal of weight to deaths that occur well before the age of 70, and

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[21] Lichtenberg (2007c) finds a positive association *across U.S. states* between per capita hospital expenditure growth and growth in educational attainment during the period 1991-2004, controlling for a number of other factors.

no weight to deaths that occur after the age of 70. This is the only model in which the drug vintage coefficient is not statistically significant.

To summarize, estimates based on one measure of cardiovascular drug vintage—the fraction of standard units whose active ingredients were launched after 1995—are consistent with the hypothesis that use of newer cardiovascular drugs has reduced the number of cardiovascular hospital discharges, average length of stay, and hospital days. The estimates also indicate that use of newer cardiovascular drugs has reduced the age-adjusted cardiovascular mortality rate, but not the number of potential years of life lost due to cardiovascular disease before age 70 per 100,000 population. Table 5 shows that similar results are obtained when we use an alternative measure of cardiovascular drug vintage—the fraction of standard units whose active ingredients were launched after 1990.

IV. Implications of the estimates

During the period 1995-2004, the mean vintage of cardiovascular drugs increased: the mean fraction of post-1995 drugs increased from 0.0% to 9.6%, and the mean fraction of post-1990 drugs increased from 1.4% to 14.7%. The vintage coefficient estimates presented in Tables 4 and 5 imply that this increase in vintage reduced the cardiovascular hospitalization rate, length of stay, and age-adjusted mortality. In other words, if drug vintage had not increased during 1995-2004, the number of hospital discharges, hospital days, and deaths would have been higher in 2004. The magnitudes of the implied increases are shown in Table 6. The two vintage measures yield similar estimates of the increase in the number of discharges in 2004 had there been no increase in vintage after 1995. The mean of these estimates is 51%. The mean estimate of the increase in average length of stay is 12%, and of the increase in hospital days is 70%.

We estimate that per capita expenditure for cardiovascular hospital stays would have been $89 higher in 2004 had drug vintage not increased during 1995-2004. This is how we derived this estimate. In 2004, mean per capita expenditure on hospital care for *all* diagnostic categories in OECD countries was $926. We estimate that cardiovascular

hospitalizations account for about 14% of total hospital expenditure.[22] Hence mean per capita expenditure on cardiovascular hospital care in OECD countries was \$128 (= 14% * \$926). The vintage coefficient estimates imply that, absent the 1995-2004 increase in drug vintage, there would have been 70% more cardiovascular hospital days in 2004. Hence per capita expenditure for cardiovascular hospital stays would have been \$89 (= 70% * \$128) higher had drug vintage not increased.

It is reasonable to expect that per capita expenditure on cardiovascular *drugs* would have been *lower* in 2004 had drug vintage not increased during 1995-2004. The *OECD Health Data 2006* database contains some information on sales of pharmaceutical products by selected Anatomic Therapeutic Chemical (ATC) groups, based on retail prices. In particular, it contains information about per capita sales of cardiovascular system drugs, in U.S. PPP dollars. To assess the effect of cardiovascular drug vintage on per capita expenditure on cardiovascular drugs, we estimated the following model:

$$\ln(\text{cardio\_rx\_expend}_{it}) = \pi\,\text{rx\_vint}_{it} + \gamma\,\ln(\text{n\_cv\_su}_{it}) + \alpha_i + \delta_t + \varepsilon_{it}$$

Estimates of π based on the two different drug vintage measures are shown in Table 7. Both estimates are positive and significant, as one would expect: countries with larger increases in the share of SUs that were for post-1995 or post-1990 drugs had larger increases in per capita drug expenditure, conditional on growth in the number of SUs per capita.

To be conservative, we will use the larger (more negative) estimate of the reduction in 2004 per capita drug expenditure that would have occurred had drug vintage not increased after 1995: 25.1%. Mean per capita expenditure on all prescription drugs in OECD countries was \$477 in 2004. Cardiovascular drug expenditure accounted for about 20% of total drug expenditure, so per capita expenditure on cardiovascular drugs in OECD countries was \$95 (= 20% * \$477) in 2004. The more conservative estimate implies that per capita cardiovascular drug expenditure would have been \$24 (= 25.1% * \$95) lower in 2004 had drug vintage not increased after 1995. Our estimate of the increase in per capita expenditure for cardiovascular hospital stays that would occurred

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[22] Cardiovascular hospitalizations account for about 14% of total hospital bed days.

had drug vintage not increased ($89) is about 3.7 times as large as our estimate of the reduction in per capita expenditure for cardiovascular drugs that would have occurred ($24).

V. Summary

This study has examined the effect of changes in the vintage distribution of cardiovascular system drugs on hospitalization and mortality due to cardiovascular disease using longitudinal country-level data on 20 OECD countries during the period 1995-2003.

Countries with larger increases in the share of cardiovascular drug SUs that were post-1995 SUs had smaller increases in the cardiovascular disease hospital discharge rate, controlling for the *quantity* of cardiovascular medications consumed per person, the use of other medical innovations (CT scanners & MRI units), consumption of calories, tobacco, and alcohol, and demographic variables (population size & age structure, income, and educational attainment). The estimates also indicate that use of newer cardiovascular drugs has reduced average length of stay and the age-adjusted cardiovascular mortality rate, but not the number of potential years of life lost due to cardiovascular disease before age 70 per 100,000 population.

The estimates indicate that if drug vintage had not increased during 1995-2004, hospitalization and mortality would have been higher in 2004. We estimate that per capita expenditure on cardiovascular hospital stays would have been 70% ($89) higher in 2004 had drug vintage not increased during 1995-2004. Per capita expenditure on cardiovascular drugs would have been lower in 2004 had drug vintage not increased during 1995-2004. But our estimate of the increase in expenditure on cardiovascular hospital stays is about 3.7 times as large as our estimate of the reduction in per capita expenditure for cardiovascular drugs that would have occurred ($24).

Although our data on hospital use and especially on drug utilization were quite complete and reliable, data on cardiovascular risk factors were less complete. We can think of little reason to expect these risk factors to be correlated with drug vintage, conditional on income, education, and average rate of drug utilization (which we control

for). Nevertheless, obtaining better information on these risk factors would certainly be desirable.

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Table 1

Number of standard units of cholesterol & triglyceride regulator agents sold in the U.S. in 2005, by active ingredient

subclass/ingredient	world launch month/year	thousands of standard units
C10A cholesterol & triglyceride regulator		8,464,846
ATORVASTATIN	1/1997	3,144,330
SIMVASTATIN	6/1988	1,555,881
GEMFIBROZIL	2/1982	548,762
EZETIMIBE	11/2002	539,047
LOVASTATIN	9/1987	525,054
PRAVASTATIN	10/1989	493,161
FENOFIBRATE	4/1975	428,381
NICOTINIC ACID		366,023
ROSUVASTATIN	2/2003	343,580
COLESEVELAM	9/2000	193,975
FLUVASTATIN	2/1994	182,766
COLESTYRAMINE	4/1967	96,573
COLESTIPOL	3/1977	47,313
ACETYLSALICYLIC ACID	1/1910	213
ALLIUM SATIVUM	1/1919	0
ASCORBIC ACID		0
CERIVASTATIN	4/1997	0
CLOFIBRATE		0
DEXTROTHYROXINE		0
GLYCINE MAX		0
LECITHIN		0
NICOTINAMIDE		0
PROBUCOL	2/1977	0
PYRIDOXINE	1/1908	0
RIBOFLAVIN		0
THIAMINE		0
VITAMIN E		0

Figure 1
Post-1995 cardiovascular Standard Units (Sus) as a percentage of total SUs in 2004, by country

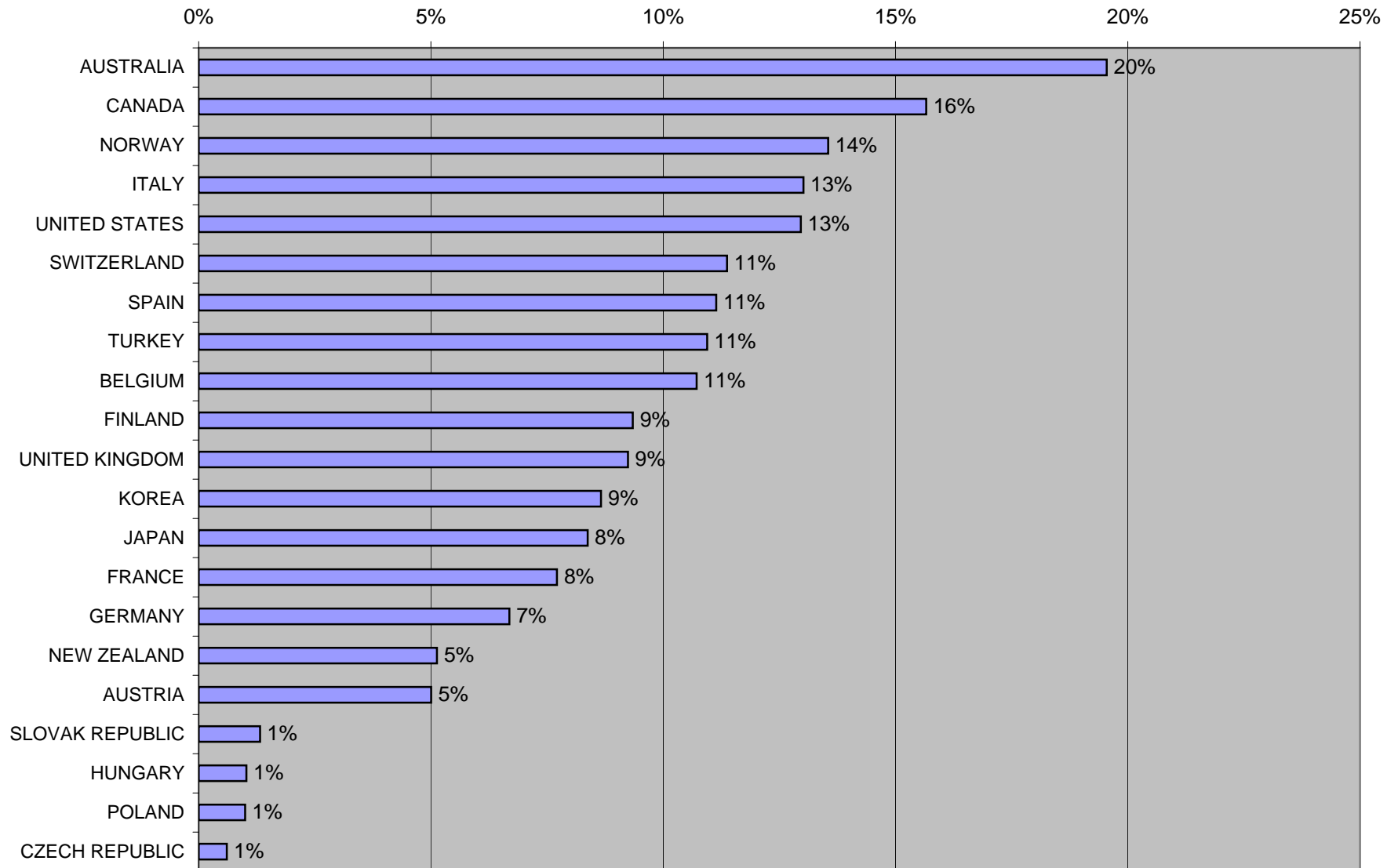| Country	Percentage
AUSTRALIA	20%
CANADA	16%
NORWAY	14%
ITALY	13%
UNITED STATES	13%
SWITZERLAND	11%
SPAIN	11%
TURKEY	11%
BELGIUM	11%
FINLAND	9%
UNITED KINGDOM	9%
KOREA	9%
JAPAN	8%
FRANCE	8%
GERMANY	7%
NEW ZEALAND	5%
AUSTRIA	5%
SLOVAK REPUBLIC	1%
HUNGARY	1%
POLAND	1%
CZECH REPUBLIC	1%

Table 2
Top 10 (ranked by no. of standard units) cardiovascular drugs in 4 countries in 2005

molecule	launch year	% of std units	class		molecule	launch year	% of std units	class
AUSTRALIA (N=2,997,258)					**NEW ZEALAND (N=620,001)**			
ATORVASTATIN	1997	9%	C10A CHOLEST&TRIGLY.REGULATOR		SIMVASTATIN	1988	14%	C10A CHOLEST&TRIGLY.REGULATOR
SIMVASTATIN	1988	5%	C10A CHOLEST&TRIGLY.REGULATOR		METOPROLOL	1975	8%	C7A B-BLOCKING AGENTS,PLAIN
HYDROCHLOROTHIAZIDE	1959	5%	C9D ANGIOTEN-II ANTAG, COMB		QUINAPRIL	1989	7%	C9A ACE INHIBITORS PLAIN
FUROSEMIDE	1964	4%	C3A DIURETICS		NITROGLYCERIN	1915	6%	C1E NITRITES AND NITRATES
IRBESARTAN	1997	4%	C9C ANGIOTEN-II ANTAG, PLAIN		FUROSEMIDE	1964	6%	C3A DIURETICS
PERINDOPRIL	1988	4%	C9A ACE INHIBITORS PLAIN		BENDROFLUMETHIAZIDE	1957	6%	C3A DIURETICS
ATENOLOL	1976	4%	C7A B-BLOCKING AGENTS,PLAIN		CILAZAPRIL	1986	6%	C9A ACE INHIBITORS PLAIN
RAMIPRIL	1989	4%	C9A ACE INHIBITORS PLAIN		FELODIPINE	1988	4%	C8A CALCIUM ANTAGONIST PLAIN
METOPROLOL	1975	3%	C7A B-BLOCKING AGENTS,PLAIN		DILTIAZEM	1974	4%	C8A CALCIUM ANTAGONIST PLAIN
NITROGLYCERIN	1915	3%	C1E NITRITES AND NITRATES		HYDROCHLOROTHIAZIDE	1959	3%	C9B ACE INHIBITORS COMBINAT
SWITZERLAND (N=1,903,135)					**AUSTRIA (N=1,890,382)**			
HYDROCHLOROTHIAZIDE	1959	3%	C9D ANGIOTEN-II ANTAG, COMB		HYDROCHLOROTHIAZIDE	1959	6%	C9B ACE INHIBITORS COMBINAT
HEPARIN	1910	3%	C5B VARICOSE THERAPY,TOPICAL		METOPROLOL	1975	3%	C7A B-BLOCKING AGENTS,PLAIN
ATORVASTATIN	1997	2%	C10A CHOLEST&TRIGLY.REGULATOR		AMLODIPINE	1990	3%	C8A CALCIUM ANTAGONIST PLAIN
METOPROLOL	1975	2%	C7A B-BLOCKING AGENTS,PLAIN		CRATAEGUS OXYACANTHA	.	3%	C1X ALL OTHER CARDIAC PREPS
HYDROCHLOROTHIAZIDE	1959	2%	C9B ACE INHIBITORS COMBINAT		FUROSEMIDE	1964	3%	C3A DIURETICS
TORASEMIDE	1992	2%	C3A DIURETICS		BISOPROLOL	1986	3%	C7A B-BLOCKING AGENTS,PLAIN
AMLODIPINE	1990	2%	C8A CALCIUM ANTAGONIST PLAIN		GINKGO BILOBA	1965	3%	C4A CEREBR/PERIPH VASOTHERAP
ATENOLOL	1976	2%	C7A B-BLOCKING AGENTS,PLAIN		SIMVASTATIN	1988	3%	C10A CHOLEST&TRIGLY.REGULATOR
SIMVASTATIN	1988	2%	C10A CHOLEST&TRIGLY.REGULATOR		HEPARIN	1910	3%	C5B VARICOSE THERAPY,TOPICAL

Table 3
Descriptive statistics

Variable	MEAN	STD. DEV.	MIN	MAX
number of hospital discharges in which the principal diagnosis was cardiovascular disease per 100,000 population	2330	963	478	4710
number of hospital days in which the principal diagnosis was cardiovascular disease per 100,000 population	23701	14935	4646	67314
average length of stay in hospital discharges in which the principal diagnosis was cardiovascular disease	9.708	2.949	4.7	17.3
age-standardized number of deaths caused by diseases of the circulatory system per 100,000 population	280	115	123	562
potential years of life lost due to diseases of the circulatory system before age 70 per 100,000 population	743.84	361.8	333	1988
the number of standard units of cardiovascular drugs per person	0.143	0.062	0.005	0.279
fraction of standard units whose active ingredients were launched after 1995	4.2%	4.7%	0.0%	21.8%
fraction of standard units whose active ingredients were launched after 1990	7.4%	6.2%	0.0%	24.9%
number of computed tomography scanners per million population	16.281	12.588	3.5	92.6
number of magnetic resonance imaging units per million population	6.752	6.058	0.6	40.1
population (in thousands)	37,658	55,552	267	296,972
fraction of the population age 65 and over	13.9%	3.2%	5.0%	20.0%
per capita GDP (in PPP US dollars)	$22,866	$8,360	$5,558	$48,162
average number of calories consumed per person per day	3334	253	2768	3817
per capita expenditure on tobacco (in PPP US dollars)	$327	$346	$26	$2,662
per capita expenditure on alcohol (in PPP US dollars)	$279	$138	$43	$788
fraction of the population with education at or below the lower secondary level (ISCED 0/1/2)	29.8%	17.2%	10.7%	79.0%
fraction of the population with education at the upper secondary or post-secondary, non-tertiary level (ISCED 3/4)	47.4%	16.3%	11.9%	76.7%
fraction of the population with education at the tertiary level (ISCED 5A/6)	14.8%	5.4%	6.0%	29.8%

Table 4
Estimates of eq. (1): $\ln Y_{it} = \beta\, \text{rx\_post1995\%}_{it} + \gamma\, Z_{it} + \alpha_i + \delta_t + \varepsilon_{it}$

Model	1			2			3			4			5		
Dep. Var.	ln(n_discharges)			ln(alos)			ln(n_days)			ln(n_deaths)			ln(pyll)		
Parameter	estimate	t-value	prob>t	estimate	t-value	prob>t	estimate	t-value	prob>t	estimate	t-value	prob>t	estimate	t-value	prob>t
rx_post1995%	-4.013	-5.51	<.0001	-0.858	-2.30	0.0271	-4.870	-5.83	<.0001	-0.846	-2.88	0.0059	0.333	0.61	0.5447
ln(n_cv_su)	0.006	0.47	0.6418	0.012	1.78	0.0831	0.018	1.20	0.2363	0.014	2.12	0.0391	0.021	1.74	0.0886
ln(ct_scan)	0.289	2.63	0.0123	0.058	1.03	0.3084	0.347	2.75	0.009	-0.029	-0.69	0.4911	-0.041	-0.53	0.5994
ln(mri)	-0.043	-0.76	0.4545	0.046	1.58	0.1224	0.003	0.05	0.9629	0.007	0.22	0.8235	-0.068	-1.24	0.2192
ln(calories)	-1.324	-2.48	0.0178	0.754	2.75	0.009	-0.571	-0.93	0.3583	0.464	1.81	0.0764	1.239	2.60	0.0122
ln(tobacco)	0.463	2.94	0.0055	-0.168	-2.09	0.0434	0.295	1.63	0.1111	0.052	0.70	0.4888	-0.133	-0.97	0.3378
ln(alcohol)	-0.233	-1.03	0.3094	0.539	4.65	<.0001	0.306	1.18	0.2448	-0.061	-0.60	0.5516	-0.195	-1.04	0.3027
ln(population)	1.936	1.37	0.1781	0.476	0.66	0.5141	2.412	1.49	0.1444	-0.030	-0.05	0.9632	0.326	0.27	0.7877
age_ge_65%	0.080	2.58	0.0139	-0.007	-0.41	0.6845	0.073	2.06	0.0458						
ln(gdp)	0.160	0.42	0.6792	-0.518	-2.63	0.0123	-0.357	-0.81	0.4232	-0.072	-0.40	0.6898	0.267	0.80	0.4259
edu_low%	-0.023	-1.93	0.0603	0.004	0.63	0.5323	-0.019	-1.40	0.168	0.015	2.69	0.0098	0.036	3.59	0.0008
edu_mid%	-0.016	-1.01	0.321	0.011	1.35	0.1857	-0.005	-0.27	0.7851	0.015	2.14	0.0371	0.035	2.70	0.0096
edu_max%	0.020	0.82	0.4176	0.039	3.12	0.0034	0.058	2.11	0.0416	0.022	2.54	0.0142	0.034	2.14	0.0371
year 1995	0.182	1.27	0.211	0.290	3.94	0.0003	0.472	2.87	0.0066	0.207	3.56	0.0008	0.301	2.79	0.0074
year 1996	0.100	0.74	0.4663	0.209	3.02	0.0045	0.309	1.99	0.0539	0.162	3.18	0.0025	0.268	2.83	0.0067
year 1997	0.082	0.73	0.4688	0.167	2.93	0.0056	0.249	1.95	0.0587	0.129	2.79	0.0074	0.245	2.86	0.0062
year 1998	0.101	1.09	0.2843	0.126	2.65	0.0115	0.227	2.13	0.0395	0.117	2.94	0.0049	0.217	2.94	0.005
year 1999	0.102	1.39	0.1737	0.106	2.82	0.0076	0.208	2.47	0.0182	0.098	3.01	0.0041	0.185	3.08	0.0034
year 2000	0.095	1.64	0.1091	0.088	2.98	0.005	0.183	2.76	0.0088	0.057	2.20	0.0323	0.136	2.86	0.0063
year 2001	0.065	1.71	0.0959	0.068	3.46	0.0013	0.133	3.03	0.0043	0.028	1.52	0.1352	0.090	2.62	0.0116
year 2002	0.037	1.40	0.1702	0.050	3.71	0.0006	0.087	2.88	0.0065	0.019	1.46	0.1509	0.058	2.34	0.0231
year 2003	0.000	.	.	0.000	.	.	0.000	.	.	0.000	.	.	0.000	.	.
R^2	0.993			0.997			0.996			0.997			0.990		
Degrees of freedom															
Model	36			36			36			33			33		
Error	39			39			39			49			49		
Corrected Total	75			75			75			82			82		

Note: all models include country fixed effects.

Table 5

Estimates of eq. (1): $\ln Y_{it} = \beta\, rx\_post1990\%_{it} + \gamma\, Z_{it} + \alpha_i + \delta_t + \varepsilon_{it}$

Model	1			2			3			4			5		
Dep. Var.	ln(n_discharges)			ln(alos)			ln(n_days)			ln(n_deaths)			ln(pyll)		
Parameter	estimate	t-value	prob>t	estimate	t-value	prob>t	estimate	t-value	prob>t	estimate	t-value	prob>t	estimate	t-value	prob>t
rx_post1990%	-3.270	-3.01	0.0046	-1.087	-2.35	0.0239	-4.357	-3.52	0.0011	-0.915	-2.67	0.0103	-0.296	-0.47	0.6406
ln(n_cv_su)	0.011	0.74	0.465	0.013	1.98	0.055	0.024	1.39	0.1729	0.015	2.32	0.0245	0.019	1.6	0.1169
ln(ct_scan)	0.205	1.58	0.1232	0.051	0.92	0.3629	0.256	1.73	0.0917	-0.028	-0.67	0.5087	-0.058	-0.74	0.4627
ln(mri)	-0.023	-0.34	0.7383	0.052	1.78	0.0828	0.029	0.37	0.7134	0.023	0.77	0.4426	-0.066	-1.18	0.2439
ln(calories)	-1.729	-2.68	0.0106	0.644	2.35	0.0239	-1.085	-1.48	0.1469	0.374	1.43	0.1584	1.218	2.54	0.0144
ln(tobacco)	0.536	2.77	0.0084	-0.139	-1.69	0.0989	0.397	1.81	0.0785	0.082	1.11	0.2727	-0.156	-1.15	0.2541
ln(alcohol)	-0.415	-1.54	0.1318	0.497	4.33	0.0001	0.082	0.27	0.7915	-0.093	-0.91	0.3663	-0.187	-1	0.3232
ln(population)	1.085	0.64	0.5289	0.521	0.72	0.4774	1.606	0.83	0.4134	-0.187	-0.29	0.7697	1.045	0.89	0.376
age_ge_65%	0.097	2.59	0.0133	-0.002	-0.11	0.9099	0.095	2.24	0.0311						
ln(gdp)	0.594	1.3	0.2018	-0.413	-2.12	0.0405	0.182	0.35	0.7291	-0.021	-0.12	0.9081	0.290	0.86	0.3922
edu_low%	-0.031	-2.16	0.0369	0.002	0.31	0.7583	-0.029	-1.78	0.0823	0.012	2.14	0.0378	0.037	3.71	0.0005
edu_mid%	-0.020	-1.06	0.2971	0.010	1.2	0.2363	-0.010	-0.48	0.6345	0.012	1.71	0.0933	0.033	2.51	0.0153
edu_max%	0.007	0.24	0.8082	0.039	3.16	0.0031	0.046	1.4	0.1707	0.023	2.68	0.0099	0.036	2.23	0.0306
year 1995	0.208	1.2	0.2375	0.284	3.85	0.0004	0.491	2.49	0.017	0.214	3.67	0.0006	0.262	2.44	0.0182
year 1996	0.125	0.77	0.4487	0.202	2.89	0.0062	0.327	1.75	0.0873	0.168	3.28	0.0019	0.227	2.42	0.0195
year 1997	0.128	0.95	0.347	0.163	2.85	0.0069	0.292	1.9	0.0645	0.134	2.88	0.0059	0.204	2.38	0.0214
year 1998	0.149	1.34	0.1893	0.125	2.63	0.012	0.275	2.16	0.0371	0.123	3.12	0.003	0.180	2.47	0.017
year 1999	0.133	1.5	0.141	0.105	2.8	0.008	0.238	2.37	0.023	0.101	3.08	0.0033	0.158	2.62	0.0116
year 2000	0.110	1.59	0.1196	0.090	3.04	0.0042	0.200	2.54	0.0153	0.060	2.35	0.0229	0.119	2.51	0.0153
year 2001	0.070	1.52	0.1361	0.069	3.52	0.0011	0.138	2.65	0.0115	0.030	1.58	0.1195	0.080	2.31	0.025
year 2002	0.045	1.43	0.1618	0.052	3.86	0.0004	0.097	2.7	0.0103	0.021	1.56	0.1243	0.055	2.21	0.0314
year 2003	0.000	.	.	0.000	.	.	0.000	.	.	0.000	.	.	0.000	.	.
R^2	0.990			0.997			0.994			0.997			0.990		
Degrees of freedom															
Model	36			36			36			33			33		
Error	39			39			39			49			49		
Corrected Total	75			75			75			82			82		

Note: all models include country fixed effects.

Table 6
% increase in dependent variable in 2004 if no 1995-2004 change in drug vintage

		% increase in dependent variable in 2004 if no 1995-2004 change in drug vintage			
		$[\exp(-\beta \ (rx\_vint_{2004} - rx\_vint_{1995}))] - 1$			
vintage measure (rx_vint)	1995-2004 change in vintage measure $(rx\_vint_{2004} - rx\_vint_{1995})$	n_discharges	alos	n_days	n_deaths
rx_post1995%	10%	47%	9%	60%	9%
rx_post1990%	13%	55%	16%	79%	13%
average		51%	12%	70%	11%

Table 7

Estimates of the effect of cardiovascular drug vintage on per capita expenditure on cardiovascular drugs

vintage measure (rx_vint)	π	t-stat	p-value	$[\exp(-\pi \, (rx\_vint_{2004} - rx\_vint_{1995}))]-1$
rx_post1995%	2.70	2.99	0.004	-22.90%
rx_post1990%	2.16	2.34	0.022	-25.10%

Estimates of π based on the two different drug vintage measures

$\ln(cardio\_rx\_expend_{it}) = \pi \, rx\_vint_{it} + \gamma \, \ln(n\_cv\_su_{it}) + \alpha_i + \delta_t + \varepsilon_{it}$