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CONTRACEPTION AND FERTILITY:

HOUSEHOLD PRODUCTION UNDER UNCERTAINTY

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## I Introduction

Over the past century fertility behavior in the United States has undergone profound changes. Measured by cohort fertility the average number of children per married woman has declined from about 5.5 children at the time of the Civil War to 2.4 children at the time of the Great Depression. It is seldom emphasized however that an even greater relative change took place in the dispersion of fertility among these women: the percentage of women with, say, seven or more children declined from 36% to under 6%.<sup>1</sup> While students of population have offered reasonably convincing explanations for the decline in fertility over time, they have not succeeded in explaining the fluctuations in the trend and have made surprisingly little effort to explain the large and systematic decline in the dispersion of fertility over time. In this paper we attempt to study contraception behavior and its effects on fertility. One of the effects on which we focus considerable attention is the dispersion or variance in fertility. Our analysis is applied to cross-sectional data but it also provides an explanation for the decline in the variance of fertility over time.

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<sup>1</sup>These figures are taken from the report of the President's Commission on Population Growth and the American Future (see Taeuber 1972). They are indicated below in Table I-1.

The study of fertility behavior has received increasing attention by economists in the past few years. Much of this analysis has been conducted in the context of the new theory of consumer behavior pioneered by Becker (1965) and Lancaster (1966). The work on fertility behavior complements many other studies dealing with aspects of household production. One of the specific topics in the fertility literature has been the relationship between child bearing and several life cycle production decisions such as marriage, schooling, women's career choices, life cycle time and money allocations and so forth. A second and related topic of the economics of fertility behavior is the tradeoff in household production between the family's number of children and the expenditure of resources per child, particularly the expenditure of time devoted to children at the pre-school age. A third focus of this research has been the fertility demand function--the form and stability over time and across groups of the household's demand function for children.<sup>2</sup>

Nearly all of these studies of household fertility behavior assume that the household can produce exactly the number of children it wants, costlessly and with certainty. We have previously pointed out that costly fertility control operates as a subsidy to child bearing,

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<sup>2</sup>For a thorough model of fertility demand and the quantity-quality trade-offs in the context of a static framework see Willis (1973). For an extensive set of papers pertaining to topics in fertility behavior see the two NBER conference volumes New Economic Approaches to Fertility and Marriage, Family Human Capital and Fertility, Schultz (ed.) (1973) and (1974). These volumes indicate, we think, that much of observed fertility behavior is amenable to economic analysis.

lowering the marginal cost of having additional children (see Willis 1971) and we have suggested a framework for analyzing the household's fertility control decisions (see Michael 1973). In this paper we consider the household's fertility control behavior both in terms of the selection of specific fertility control strategies (the costs and benefits of specific contraceptive techniques) and in terms of the effects of different control strategies on household fertility.

One could introduce fertility control costs into a deterministic model of fertility behavior by treating these costs as transaction costs associated with acquiring any given level of fertility. In this framework, the household can select any number of children with certainty provided it pays the requisite costs of fertility control. Assuming that total fertility control costs are larger the smaller the number of children chosen, the positive marginal cost of fertility control raises average fertility by acting as a subsidy to childbearing. In this paper, we have treated the costs of fertility control in a somewhat different framework. We have adopted a model in which the household can select with certainty any particular monthly probability of conception,<sup>3</sup> but in which the household's actual fertility,  $N$ , is a stochastic variable. By selecting and producing a particular monthly

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<sup>3</sup>The probability is bounded by zero and by the probability implied by natural or intrinsic fecundability, say, a probability of about 0.2 per month.

probability of conception, the household selects a distribution of fertility outcomes. The mean of that distribution is its expected fertility; its variance indicates the uncertainty that the household faces.

As Table I-1 indicates, the decline in average fertility

Table I-1

Mean Number of Children and Frequency Distribution  
by Number of Children for Selected Cohorts  
from 1835-1930 for Ever-Married Women

Cohort	Mean No. of Children	Percent Distribution				
		0 Children	1-2 Children	3-4 Children	5-6 Children	≥ 7 Children
1835-39	5.40	7.7	17.3	20.0	18.7	36.3
1845-49	5.27	8.2	18.5	20.3	18.3	34.8
1855-59	4.97	8.9	20.6	21.3	17.9	31.3
1865-69	3.90	12.3	26.6	26.1	16.0	18.9
1875-79	3.46	15.0	30.4	25.2	14.4	15.0
1885-89	3.15	16.6	33.1	25.1	13.1	12.2
1895-99	2.71	18.6	39.0	23.9	10.0	8.4
1905-09	2.36	20.8	43.2	22.4	7.8	5.9
1915-19	2.60	13.9	43.7	28.1	8.9	5.4
1926-30	3.08	8.0	36.5	36.2	12.3	7.0

Source: Taeuber 1972 (with the exception of the most recent cohort the women were at least age 45 at the time of the enumeration; for the 1926-30 cohort the women were age 40-44 at the time of the survey.)

over the past century has been accompanied by a significant reduction in the dispersion of fertility. The stochastic model of fertility control which we discuss in the following section emphasizes the relationship between the mean and variance of fertility and offers an explanation for the observed decline in the dispersion in fertility over time.

Child rearing is an exceptionally costly activity, both in terms of direct dollar outlays and forgone time and human capital.<sup>4</sup> Probably few events in one's lifetime affect subsequent behavior more extensively than having a child. While other important life cycle decisions such as marriage and career choice are subject to considerable uncertainty, the uncertainty generally pertains to the quality or the characteristics of the object of choice. Uncertainty about the characteristics of the prospective child also exist of course, but in addition there exists the uncertainty which we are emphasizing--uncertainty about the acquisition of a child itself.

At the individual household level this uncertainty about the number of children affects at least three aspects of behavior. First, it may affect decisions about the expenditures of resources on existing children--if ordinary substitution between quantity and quality is relevant to children, then not knowing the final number of children may affect the household's expenditure decisions on its first-born children. Second, there exists substitution between expenditures on

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<sup>4</sup>For estimates of the direct costs of children see Cain (1971) or Reed and McIntosh (1972). Lindert (1973) presents a useful discussion of existing evidence on various aspects of the costs of children. Michael and Lazear (1971) emphasize the potential cost of children in terms of forgone human capital and Mincer and Polachek (1974) estimate the depreciation in the mother's human capital related to her nonmarket child rearing activity.

children and on other household goods and services and also between expenditures over time. So uncertainty about the number of children, and about the timing or spacing of children, can be expected to affect the composition and timing of consumer expenditure and savings behavior. Third, because of important interactions with other household production and with the relative value of family member's time, uncertainty about the number of children may have effects on the parents' occupation choices, schooling decisions and general orientation toward market and household activities.

At the aggregate level, positive fertility control costs and the stochastic nature of fertility behavior affect the observed mean and variance of fertility. The size and growth rate of the population affect the age distribution of the population and the rate of growth and the composition of the economy's output.<sup>5</sup> The variance in fertility, on the other hand, influences the distribution of income and of wealth. If uncertainty about fertility outcomes affects household investment and savings decisions, it may have an important influence on the distribution of inherited wealth across generations.

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<sup>5</sup>See Kelley (1972) for a recent discussion of population growth and economic progress. See Kuznets (1960) and other essays in Demographic and Economic Change in Developed Countries for discussions of the effects of population on output employment and demand.



These considerations are not the focus of our paper, but we think the points we emphasize here--the costs of fertility control, fertility as a stochastic process and the relationship between the mean and variance of fertility--have important implications for the level and distribution of the economy's wealth. We do not explore these aggregate relationships, nor do we resolve many of the more esoteric problems which we encounter in our analysis. We do however attempt to integrate into an analysis of contraceptive choice and optimum fertility behavior the constraints imposed by biological limitations and resource (or economic) limitations. We indicate how the choice of contraceptive technique affects the observed mean and variance of fertility. We also analyze the choice of contraceptive technique, in particular the adoption of the new oral contraceptive in the United States in the first half of the 1960's.

## II The Analytical Framework

The theory of the choice of a fertility control strategy treats the fertility goals of the household as given, while the economic theory of fertility demand focuses on the factors which determine these goals. If fertility control is costly, however, these costs as well as the resource costs of bearing and rearing children influence the couple's choice of fertility goals. The link between the theory of the choice of birth control technique and the theory of fertility behavior is provided by assuming that the household maximizes its lifetime utility subject to the constraint of a fertility control cost function as well as the conventional economic resource constraint. The fertility control cost function is simply the combination (the envelope) of least-cost birth control strategies for all possible fertility outcomes.

In this section we describe a stochastic model of reproduction, emphasizing the relationship between the mean and variance of fertility outcomes. We then discuss the economic benefits and costs of fertility control and conclude with an exposition of the optimal fertility control strategy.

## 1 Birth control and the distribution of fertility outcomes

The number of children born to a couple and the pace at which these children are born is ultimately constrained by the fact that reproduction is a biological process. The observed reproductive behavior of an individual woman over her life cycle may be regarded as the outcome of this biological process as it is modified by non-volitional social and cultural factors and by the effects of deliberate attempts to control fertility. In the past two decades, the nature of the biological constraint on fertility choices has been greatly clarified and given rigorous expression in stochastic models of the reproductive process by Henry, Potter, Perrin and Sheps, and others. The basic reasoning underlying these models and their main implications for average fertility were recently summarized by Keyfitz (1971).

These models suggest that the number of children a woman bears during her lifetime is a random variable whose mean and variance depend on her (and her partner's) choice of a fertility control strategy. In this section we draw heavily on this literature in order to present, under simplifying assumptions, analytical expressions for the mean and variance of live births as a function of two sets of parameters, one representing the couple's biological characteristics and the other its fertility control strategy.

The simple observation that it takes a random amount of time to produce a baby provides the point of departure for recent biological models of fertility. Suppose that a woman faces a probability  $p$  of

conceiving in a given month. If that monthly probability of conception is constant over time, the probability that she will conceive in exactly the  $j$ th month ( $j = 1, 2, \dots$ ) is  $p(1-p)^{j-1}$  where  $(1-p)^{j-1}$  is the probability that she fails to conceive in the first  $j-1$  months. Employing the demographer's term "conceptive delay", i.e. the number of months,  $v$ , it takes a fecund woman to conceive, the random variable  $v$  is distributed geometrically with mean  $\mu_v = (1-p)/p$  and variance  $\sigma_v^2 = (1-p)/p^2$ .<sup>6</sup>

Once a woman conceives she becomes sterile during her pregnancy and the anovulatory period following pregnancy. The length of the sterile period,  $s$ , is also a random variable whose value depends on the type of pregnancy termination (i.e., fetal loss or still-birth or live birth) and on the physiological and social factors (e.g., age, parity, breast feeding practices, time to resumption of sexual activity) which determine the length of the anovulatory period following each type of pregnancy termination. For simplicity, we shall assume that all pregnancies terminate in a live birth and that the length of the sterile period,  $s$ , is of fixed, nonrandom length.<sup>7</sup> The length of one reproductive cycle--the number of months

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<sup>6</sup>See Sheps (1964) for a derivation of this result. It should be noted that conceptive delay is defined to be zero months if the woman conceives in the first month.

<sup>7</sup>See Perrin and Sheps (1964) for a model in which pregnancy terminations other than live births are allowed and the sterile period associated with each type of pregnancy is of random length. Compared with the formulas we shall present, the Perrin and Sheps model implies a smaller mean and larger variance in the number of live births a woman has over her reproductive span.

it takes a fecund woman to become pregnant, give birth and revert to a fecund, non-pregnant status--can be expressed as  $t = v + s$ , a random variable with mean  $\mu_t = \mu_v + s$  and variance  $\sigma_t^2 = \sigma_v^2$ .

The number of children the woman bears during a lifetime, say a reproductive span of  $T$  months, depends on the number of reproductive cycles completed during this period. Since each cycle is of random length, the woman's fertility will also be a random variable. The probability distribution of the number of births can be represented in a simple way if the model of reproduction is represented as a Markov renewal process. In order to qualify as a renewal process, the intervals between successive births must behave as independent, identically distributed random variables (Potter, 1970). To meet these qualifications, it is necessary to assume that all of the parameters of the reproductive process (i.e.,  $p$  and  $s$ ) are constant over time and that the reproductive period,  $T$ , is sufficiently long (i.e., infinity). Assuming reproduction to be a renewal process, the distribution of the number of births  $N$  is asymptotically normal with mean<sup>8</sup>

$$\mu_N = T/\mu_t \quad (1)$$

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<sup>8</sup> See Sheps and Perrin (1966) who warn that the asymptotic normal distribution above does not adequately approximate the exact probability of  $N$  for the relevant (finite) range of  $T$ . In another paper Perrin and Sheps (1964) suggest more accurate approximate expressions for the first two moments of  $N$ . Since the qualitative implications of these approximations are quite similar to those of the more exact approximations, it does not seem necessary to encumber the discussion with more complicated expressions for mean and variance. A more serious problem is suggested by Jain (1968) who found that the actual mean of natural fertility tends to fall progressively below the theoretical mean given by the Perrin-Sheps model as  $T$  increases, while actual variance rises progressively above the theoretical variance.

and variance

$$\sigma^2_N = T \frac{\sigma_t^2}{\mu_t^3} \cdot \quad (2)$$

Equations (1) and (2) provide a useful way of summarizing the insights provided by mathematical demography into the determinants of a woman's fertility behavior. Her mean fertility varies in direct proportion with the length of her reproductive span, T, and in inverse proportion with the expected length of her reproductive cycle. The variance of her fertility outcome depends upon these same two factors and also upon the variance in the length of her reproductive cycle. The three variables which determine both the mean fertility  $\mu_N$  and the variance of fertility  $\sigma^2_N$  in this framework are the length of the reproductive span T, the monthly probability of conception p, and the length of the sterile period s following conception. Given T, p and s, the mean and variance are jointly determined. Treating T and s as parameters, the mean and variance of N are related, at all values of p, as

$$\sigma^2_N = \mu_N - k_1 \mu_N^2 + k_2 \mu_N^3, \quad (3)$$

$$\text{with } k_1 = \frac{2s-1}{T} ; k_2 = \frac{s(s-1)}{T^2}$$

where  $k_1$  and  $k_2$  are positive constants.

Consider, next, the determinants of the monthly probability of conception p. A woman's biological capacity to reproduce may be represented by her intrinsic fecundability,  $\beta$ , which is defined as the probability of conception from a single unprotected act of coition at a random time during the menstrual cycle (which is assumed to be one month in length). In the absence of conception the probability

$p$  that she will conceive during a given month is then equal to the product of  $\hat{p}$  and her monthly frequency of coition  $c$ .<sup>9</sup> Demographers frequently discuss "natural fertility" defined, following Henry (1961), as the number of live births a woman expects to have in a reproductive lifespan of  $T$  months in the absence of any deliberate attempt to control fertility. If we suppose there is some "natural" level of coital frequency  $\hat{c}$  for a given couple, then  $\hat{c}\hat{p} = p^*$ , the couple's monthly probability of conception in the absence of any fertility control. We will generally assume  $\hat{c} = 7$  and  $\hat{p} = 0.03$  (see Tietze, 1960), hence we will assume that  $p^* = 0.2$ . Given  $p^*$  and given the reproductive time span  $T$  and the length of the period of infertility  $s$ , the mean and variance of natural fertility,  $\mu_N^{\wedge}$  and  $\sigma_N^2$ , are defined by equations 1 and 2.

Variations across couples in the monthly probability of conception  $p$  may result from variations in fecundity--which affect  $\hat{p}$ -- or from variations in coital frequency. Variations may also result from contraception. If the adoption of a particular contraceptive strategy  $i$  reduces the monthly probability of conception by  $e_i$  percent, then

$$p_i = p^* (1 - e_i) \quad 0 \leq e_i \leq 1.$$

So the couple's actual monthly probability of conception  $p_i$  is determined by its fecundity, coital frequency and contraceptive practice.

As we emphasized above, to qualify as a Markov renewal process of reproduction the monthly probability  $p_i$  is assumed to be constant

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Intrinsic fecundability is discussed in the demographic literature, which contains an empirical justification for expressing  $p$  as approximately proportional to  $c$  over the relevant range of variation in monthly frequency of coition.

for all fertile months in the reproductive timespan of T months. Under these circumstances figure II-1 indicates the relationship between the mean and variance of fertility as summarized in equation 3. Curve A assumes T=240 months ( a reproductive span of 20 years) and s=17 months (a 17 month period of sterility following conception, see Keyfitz 1971). Each point on Curve A corresponds to a different constant monthly probability of conception ranging from p=0.0 ( the origin) to p=0.2 (point f on the curve).<sup>10</sup> Curve B in figure II-1 depicts the same relationship under the assumption that T equals 240 months and s equals only 11 months.

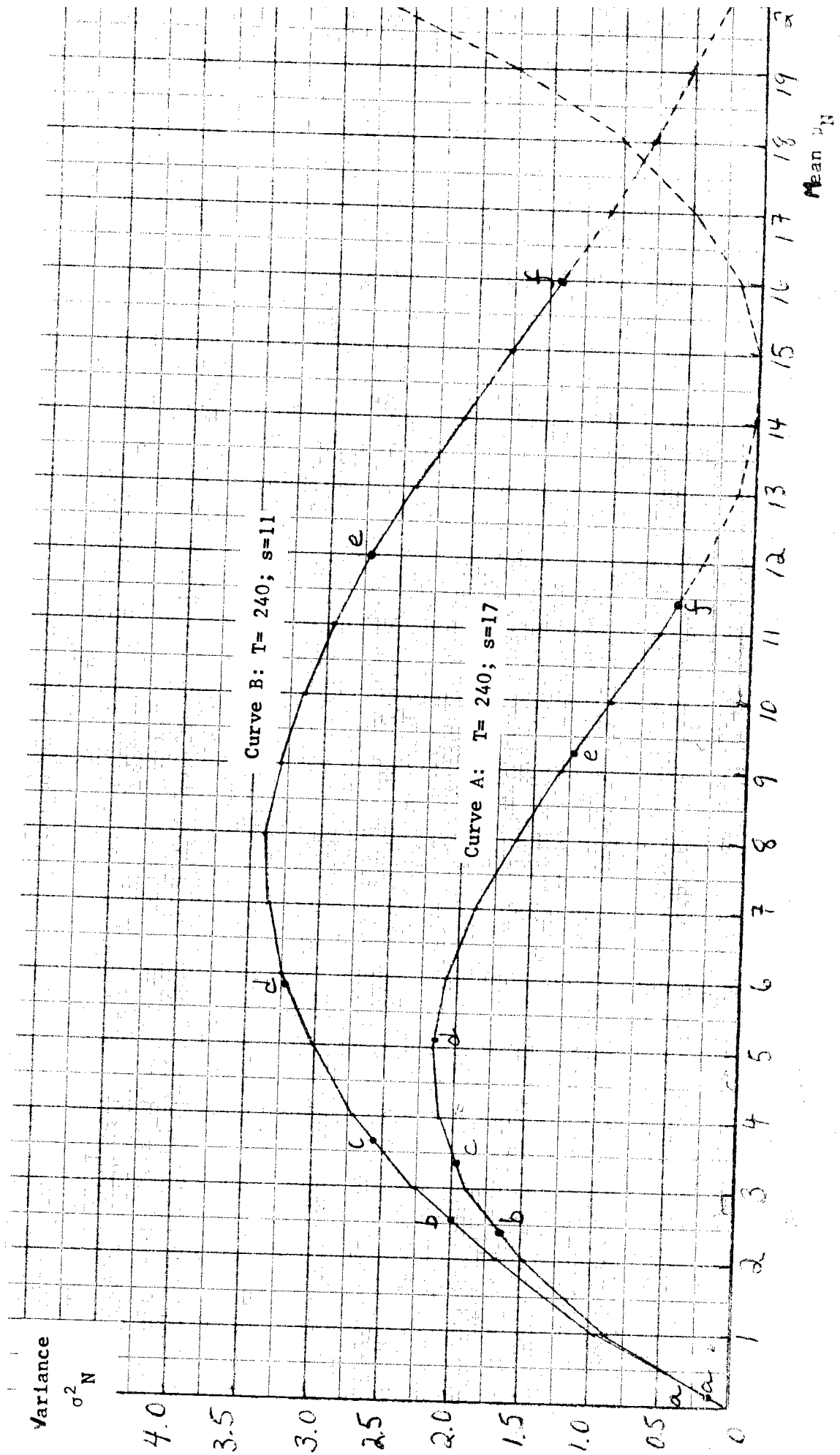
Suppose a couple, at the time of their marriage, were characterized by the parameter values T = 240 months, s = 17 months and  $p^* = 0.2$ . They would then face an ex ante distribution of fertility outcomes with a mean of 11.4 births and a variance of 0.5 (point f on Curve A in figure II-1). The couple could, however, alter this expected outcome by adopting a strategy of fertility control which lowered their constant monthly probability of conception below  $p^*$ . If, for example, the couple selected a contraceptive technique with efficiency  $e_1 = 0.5$ , then their monthly probability would be  $p_1 = p^*(1-0.5) = 0.1$  and their ex ante distribution of fertility outcomes would have a mean of 9.2 births and a variance of 1.2 (point e on

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<sup>10</sup> For example, if  $p=0.0008$  then  $\mu_N = 0.19$  births and  $\sigma_N^2 = 0.18$  which is shown as point a on Curve A. The values of p which correspond to points a,b,c,d,e,f in the figure are 0.0008, 0.0120, 0.0182, 0.0336, 0.1000, 0.2000 respectively. These values refer to specific forms of fertility control and are discussed below.



Figure II-1



Curve A). Thus the couple affects its expected fertility and its uncertainty about the number of its births by its selection of a contraceptive strategy.<sup>11</sup>

Both the contraceptive technique used and the care with which it is used can affect  $e_i$  which in turn determines  $p_i$ . In general, the couple can also affect  $p_i$  by altering its frequency of coition,  $c$ , and can also affect the distribution of fertility outcomes for any given  $p_i$  by altering the length of the reproductive period at risk,  $T$ , through decisions about the age at marriage and the age at which either partner is sterilized.<sup>12</sup> So the full range of fertility control strategies includes considerations other than the choice of contraceptive method, but it is the contraceptive choice on which

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<sup>11</sup>In reality, it is plausible to suppose that the frequency of coition and contraceptive efficiency will tend to vary over time to accommodate a couple's preferences for childspacing as well as for total number of births, or to accommodate any changes in their fertility goals.

It is also plausible to suppose that the choices of  $c$  and  $e_i$  will be conditioned on past pregnancy and birth outcomes. Thus, in general, the values of  $p$  in a given month will tend to be a function of time (i.e., age), past reproductive history (i.e., parity) and random fluctuations in variables that determine fertility goals.

Unfortunately, the analytical simplicity of considering the stochastic model of reproduction as a renewal process is lost under these conditions. While it is possible to write out probability statements in which a couple's contraceptive strategy (i.e., its choices of  $c$  and  $e_i$ ) is defined conditional on all possible fertility outcomes at each period of time, it is not possible to derive the implications of the resulting stochastic process for completed fertility outcomes using analytic methods. Moreover, the dynamic optimization problem involved in selecting a contraceptive strategy that maximizes a couple's expected utility under conditions of uncertainty may itself be analytically intractable. At this stage, it appears wiser to minimize the formal difficulties, thus the contraceptive parameters,  $c$  and  $e_i$ , as well as the biological parameters,  $\hat{p}$  and  $s$ , are assumed to remain constant over time.

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Interruptions of exposure within the time span  $T$  caused by cessations of sexual relations due to divorce or separation are ruled out by the assumption that the parameters of the process remain constant over time.

we will focus. By selecting contraceptive strategy  $i$  which yields a monthly probability of conception  $p_i$ , the couple has in effect selected a particular ex ante distribution of its fertility outcomes. The mean of that distribution is  $\mu_{N_i}$  and its variance is  $\sigma_{N_i}^2$ . We will assume for now that the couple is constrained to a pure contraceptive strategy, defined as the adoption of some form of fertility control which sets  $p$  at some fixed level (during fertile periods) for the entire reproductive span.

From studies of the average contraceptive failure rates of various contraceptive techniques, the efficiency-in-use of each technique can be computed. Table II-1 lists the contraceptive efficiency  $e_i$  and the implied monthly probability of conception  $p_i$  for several contraceptive techniques.<sup>13</sup> The table also computes for each technique the mean and variance of the length of the reproductive cycle and the mean and variance of the fertility outcomes in a 20 year reproductive span. Thus, given the biological constraint on its fertility (e.g. Curve A in figure II-1) the couple can determine the expected distribution of its fertility ( its  $\mu_N$  and  $\sigma_N^2$ ) by selecting a contraceptive strategy which achieves any particular  $p_i$ . The various points labeled on Curve A indicate the mean and variance of births associated with various contraceptive techniques (point a: pill; b: diaphragm; c. suppository; d: rhythm; e: 50% reduction in coital frequency and no other contraception; f: no fertility control).

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<sup>13</sup>The estimates of contraceptive efficiency were compiled by Michael (1973) from the demographic literature ( see especially Tietze (1959) and (1962) ). See Michael (1973) for a discussion of the difficulties in estimating and the hazards in using this comparative list of the efficiency of contraceptive methods. In particular, note that these values represent average observed use-effectiveness and will in general be affected by the intensity and care with which they are used.

Table II-1

Contraceptive Technique	Observed Use-Effectiveness R	Contraceptive Efficiency e	Monthly Birth Probability p*	Mean Length of Reproductive Cycle $\mu_t$	Variance of Reproductive Cycle $\sigma_t^2$	Reproductive lifetime of 240 months	
						Expected Number of Children $\mu_N$	Variance of Number of Children $\sigma_N^2$
Pill	1.0	.9958	.0008	1266	1,561,250	.19	.18
IUD	2.5	.9896	.0021	492	226,100	.48	.46
Condom	13.8	.9425	.0115	103	7,482	2.33	1.64
Diaphragm	14.4	.9400	.0120	100	6,723	2.40	1.61
Withdrawal	16.8	.9300	.0140	87	4,970	2.74	1.79
Jelly	18.8	.9217	.0157	80	4,032	3.00	1.89
Foam Tablets	20.1	.9167	.0168	76	3,540	3.17	1.94
Suppositories	21.9	.9088	.0182	71	2,970	3.36	1.99
Rhythm	38.5	.8396	.0312	47	930	5.11	2.15
Douche	40.3	.8321	.0336	46	870	5.21	2.15
50% Reduction in Coital Frequency		.5000	.1000	26	90	9.24	1.23
No Method		.0000	.2000	21	20	11.42	.51

R = (no. of conceptions/no. of months of exposure) x 1200 = failure rate per 100 years of use (Source: Tietze, 1959, 1962, 1970)

$p_f = R_f \div 1200$  = monthly probability of conception

$e_f = (p^* - p_f) \div p^*$  = percentage reduction in  $p^*$  where  $p^*$  is a constant equal to 0.20

$\mu_t = \left[ \frac{(1 - p_f) \div p_f}{s} \right] + s$  where  $s = 17$

$\sigma_t^2 = (1 - p_f) \div p_f^2$

$\mu_N = T \div \mu_t$  where  $T = 240$

$\sigma_N^2 = T\sigma_t^2 \div \mu_t^3$

In this framework the number of children born to a couple is a random variable which results from a stochastic process. Ex post, the couple has only one number of children  $N$ . But the couple cannot determine its number of children with absolute certainty. Rather it can select any particular value of  $p$ , the monthly probability of conception, which yields a particular distribution of fertility outcomes summarized by the distribution's mean and variance.

So long as we assume that the couple selects one value for  $p$  and retains that particular monthly probability of conception for all the fertile months in the 20 year span--an assumption we will characterize as a "pure" strategy model-- the couple cannot alter its expected fertility  $\mu_N$  without also altering the variance  $\sigma_N^2$ . In short, the pure strategy model restricts the couple to the biological constraint (Curve A in figure II-1 if  $T=240$ ,  $s=17$ ). Before we relax the assumption of the pure strategy we discuss the determinants of the couple's choice of its most preferred position along the biological constraint. We consider in turn the benefits and the costs of fertility control.

#### ii Benefits of fertility control

Recent economic theories of household behavior postulate the existence of several constraints (e.g., a money income constraint, a time constraint, production function limitations) on the household's maximization of utility. The utility is derived from a broad set of desiderata which are produced by the household itself in the nonmarket sector, using purchased market goods and services and the household

members' own time as the inputs in the production.<sup>14</sup> These production functions emphasize the distinction between the household's wants (the output) and the means used to satisfy these wants (the goods and time inputs).

Willis (1973) recently utilized the household production framework to formulate an economic model of human fertility control. We will generalize a simple version of Willis' model to deal with imperfect and costly fertility control. The formal analysis is conducted in a static lifetime framework, although we informally suggest how the implications of the model might be altered in a more dynamic or sequential decision making framework.<sup>15</sup>

In Willis' model it is assumed that the satisfaction parents receive from each of their  $N$  children is represented by  $Q_1, Q_2, \dots, Q_N$  and the satisfaction from other sources of enjoyment is represented by  $S$ . The  $Q_i$ , or "quality" of each child, and the other composite commodity  $S$  are produced within and by the household using the family members' time and purchased market goods as inputs. The household production functions characterize the relationship between inputs of time and goods and the outputs of  $Q_i$  and  $S$ . Assuming (among other things) that

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See Becker (1965) and Lancaster (1966) for early statements of this model. Several monographs and articles in recent years, notably through NBER, have utilized this framework. For a recent brief survey see Michael and Becker (1973).

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For a model of sequential decision making regarding contraceptive behavior in a heterogenous population see Heckman and Willis (1973).

parents treat all their children alike, the total amount of child-services  $C$  may be written as a product of quality per child  $Q$  and the number of children  $N$ :  $C = NQ$ . It is assumed that  $Q$  is positively related to the amount of time and market goods devoted to each child-- $Q$  is perhaps best considered an index of the child's human capital. The household's preferences for number and quality of children and for all other forms of satisfaction are summarized by its lifetime utility function

$$U = U(N, Q, S). \quad (4)$$

The household's capacity to produce  $C$  and  $S$  is limited by its lifetime real income and by the quantity of its nonmarket time. Willis (1973) discusses in detail the relationship between the relative prices of  $N$  and  $Q$ , and considers how various changes in the household's characteristics and circumstances would be expected to affect its demand for  $N$ ,  $Q$  and  $S$ .

One important implication of the economic model of fertility demand should be noted. From an assumption that children are relatively time-intensive in the wife's time (i.e.  $C$ -production requires more of the wife's time per dollar of goods-input than does  $S$ -production), the relative cost of  $C$  rises as the wife's wage rate rises. Hence the cost of both number of children  $N$  and quality of children  $Q$  also rises with the wife's wage rate. If the relative price of  $N$  rises with the wife's wage rate, then abstracting from the change in income, women with higher wages (or higher levels of education) are expected to have lower fertility. This is the basis of the "cost of time" hypothesis (Ben-Porath,

1973) which has, since Mincer's pioneering paper (Mincer, 1963), received much attention as an explanation for the observed negative relationship between the wife's wage and her fertility.

The household's lifetime money income constraint, its time constraint and its production function constraints can be treated as a single constraint on the household's lifetime full real income,  $I$ . Defining the marginal costs of childservices  $\pi_c$  and the composite other commodity  $\pi_g$ , the formal optimization problem characterizing the household's choice is the maximization of the utility function (equation 4) subject to the full real income constraint.

$$\max \{ U(N, Q, S) - \lambda [I - \phi(\pi_c, \pi_g)] \} \quad (5)$$

where  $\lambda$  is the Lagrangean multiplier. This optimization problem assumes that the household can costlessly and with certainty select any number of children ( $N$ ) it wishes and can achieve any given level of the child's human capital ( $Q$ ) it chooses.

To relax this assumption, we consider the benefits to the household of achieving any given number of children. Suppose the household had the utility function and the full real income constraint indicated in equation 5 but that the household was endowed with some arbitrary number of children  $N'$  where  $N' = 0, 1, \dots$ . (To simplify the mathematics  $N'$  will be treated as a continuous variable.) Given its arbitrary  $N'$ , the household's only remaining choices would be the optimal values of child quality,  $Q$ , and other satisfaction,  $S$ , which must be chosen subject to the lifetime full real income constraint. If  $N$  and  $Q$  as well as  $N$  and  $S$  are substitutes in terms of the parents' preferences,



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the levels of both Q and S will tend to fall as N' is increased. This analysis implies that sub-optimal values of Q and S would be chosen if fertility were arbitrarily constrained. Q and S would tend to be larger than (or smaller than) the optimal values, Q\* and S\*, as the arbitrarily constrained level of fertility, N', is smaller (or larger) than the freely chosen desired level of fertility, N\*. <sup>16</sup>

This hypothetical experiment of assigning some arbitrary N to the household is equivalent to maximizing equation 5 while treating N as a parameter, for all possible values of N. Such an exercise yields the household's net utility level as a function of its assigned level of fertility N' and the economic variables. Written as an implicit function the net utility, V, is

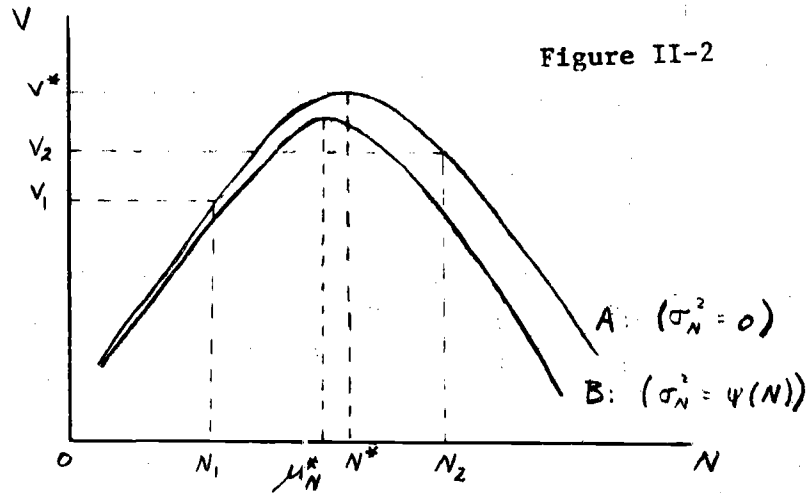
$$V = V(N; I, \pi_c, \pi_s). \quad (6)$$

For each arbitrarily assigned value of N' there is a maximum achievable level of utility, obtained by the appropriate mix of Q and S. By definition, the maximum value of V, indicated as V\*, will be achieved at the desired level of N (N\* = N'), as depicted on Curve A in Figure II-2.

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It should be noted that we are implicitly assuming that the couple knows in advance what number of children it will have and can plan accordingly for its level of child quality and S.



Deviations of fertility from  $N^*$  in either direction, such as  $N_1$  or  $N_2$ , result in reduced utility levels such as  $V_1$  or  $V_2$  in Curve A of figure II-2. Given the emphasis in discussions of family planning on the problem of excess fertility and unwanted births (i.e.  $N > N^*$ ), it is worth stressing that deficit fertility (i.e.  $N < N^*$ ) may reduce welfare by at least as much as excess fertility.<sup>17</sup>

The opportunity cost of deficit or excess fertility ( $V^* - V_1$  or  $V^* - V_2$ ) is a measure of the benefits from improved fertility control. If in the absence of fertility control the household's fertility would have been  $N_1$  (yielding  $V_1$ ) but with a given level of fertility control the household achieved  $N_j$  (yielding  $V_j$  such that  $V_j > V_1$ ) then  $(V^* - V_1) - (V^* - V_j) = V_j - V_1$  is a measure of the benefit from that level of

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As indicated above, in the case of deficit fertility quality per child,  $Q$ , would be higher than it would be in the case of optimal fertility or a fortiori for excess fertility. If from some ethical point of view parents are judged to place too little weight on their children's welfare, and if we measure child-welfare by the level of  $Q$ , it could be argued that the effect on  $Q$  of deficit fertility reduces or outweighs the parents' welfare loss  $V^* - V_1$ .

fertility control. The utility benefits, then, are the gains in utility which accrue from moving nearer the optimal allocation of resources which would exist if fertility control were perfect and costless (i.e.,  $N^*$ ,  $Q^*$  and  $S^*$ ).

We suggest above that the choice of particular values of fertility control parameters such as  $e_i$  yield particular ex ante distributions of fertility summarized by  $\mu_N$  and  $\sigma_N^2$ . We can, therefore, formulate the discussion of the benefits of fertility control in terms of  $\mu_N$  and  $\sigma_N^2$ . For purposes of illustration, suppose the functional form of equation 6 is quadratic in  $N$ :

$$V = a + b N + \frac{c}{2} N^2 \quad ; \quad b > 0, \quad c < 0 \quad (7)$$

where  $I$ ,  $\pi_c$  and  $\pi_s$  are held constant. Since the unconstrained maximum of  $V$  gives the desired value of fertility,  $N^*$ , it follows from (7) that

$$N^* = -\frac{b}{c} . \quad (8)$$

We may now treat  $N$  as a random variable and take the expected value of (7) to obtain:

$$E(V) = a + b\mu_N + \frac{c}{2} \mu_N^2 + \frac{c}{2} \sigma_N^2 . \quad (9)$$

Recall that equation (3) indicates the relationship between  $\sigma_N^2$  and  $\mu_N^2$ , ( $\sigma_{N_i}^2 = \psi(\mu_{N_i})$  where  $\psi$  is a cubic function). If  $E(V_0)$  is the value of equation 9 in the absence of any fertility control, and  $E(V_i)$  is its value when fertility control of  $e_i$  (yielding  $\mu_{N_i}$ ,  $\sigma_{N_i}^2$ ) is employed, then the benefit from fertility control strategy  $i$  is  $E(V_i) - E(V_0)$ .

If we maximize the expected  $V$  (equation 9) with respect to  $\mu_N$ ,

$$\frac{d}{d\mu_N} (E(V)) = 0 = b + c\mu_N + \frac{c}{2} \psi' , \quad (10)$$

where  $\psi'$  is a quadratic equation derived from equation 3. Noting that  $\psi'$  may be roughly approximated by a positive constant  $K(\psi' = K)$  for values of  $\mu_N < 5$  (see figure II-1), the optimal value of  $\mu_N$ , from equation (10) is

$$\mu_N^* = \frac{-b}{c} - \frac{K}{2} = N^* - \frac{K}{2}. \quad (11)$$

The optimal expected fertility  $\mu_N^*$  is somewhat lower than the desired fertility  $N^*$ . Equation 9 implies that the lower the variance  $\sigma_N^2$  the higher the net benefit  $V$ , ceteris paribus. Since the mean and variance are positively related at low values of  $N$  (i.e.,  $K = \psi' > 0$  if  $N < 5$ ), the induced reduction in mean fertility represents the adjustment to the variance associated with  $N^*$ . This is indicated by Curve B in figure 18 II-2. By subtracting an amount proportional to  $N$  from Curve A, Curve B peaks at a level of  $N$  below  $N^*$ .

So far, we have shown that imperfect fertility control implies that a couple's actual fertility  $N$  is a stochastic variable. The couple, by its choice of a contraceptive strategy, acquires some distribution of expected fertility outcomes, and by the nature of the biological process involved, the higher the mean of the distribution the greater its variance (up to at least  $N = 5$ ). Thus the greater the expected fertility, the greater the uncertainty about the actual fertility, or the greater the expected deviation between the mean and the actual fertility. Since deviations from

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<sup>18</sup>As  $N$  increases along Curve A,  $\sigma_N^2$  rises by  $\frac{K}{2}$  per unit of  $N$  (up to  $N = 5$ ), thus an amount of  $V$  proportional to  $N$ ,  $(\frac{cK}{2})N$  is subtracted from Curve A yielding Curve B.

desired fertility reduce net utility,  $V$ , and since higher levels of expected fertility,  $\mu_N$ , are associated with greater uncertainty, the household is induced to reduce its optimal expected fertility  $\mu_N^*$  below its desired fertility  $N^*$  in order to reduce the uncertainty or variance  $\sigma_N^2$ .

### iii Costs of fertility control

The costs of fertility control are the amounts of other desirables forgone in achieving the control. These costs include money costs but also include forgone time, sexual pleasure, religious principles, health, and so forth. It is, at best, difficult to measure these costs empirically. We will instead discuss some of the determinants of these costs and seek to derive testable hypotheses about the relationship between observed fertility and contraceptive behavior.

By definition, couples which avoid all the costs of fertility control have an expected level of fertility  $\mu_N^{\wedge}$ , which is frequently referred to as "natural" fertility.<sup>19</sup> We will assume that costly fertility control strategies are limited to two dimensions: (1) the choice of contraceptive technique (including regulation of coital frequency) and

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<sup>19</sup> See the discussion of natural fertility in an earlier section. Throughout section II of this paper we continue to assume that the couple's fertility control strategy is determined at the outset of the period-at-risk of conception and remains constant throughout the reproductive span. "Natural" fertility results when the age at marriage (which affects  $T$ ) and the rate of coition,  $c$ , are determined without regard to effects on fertility.

(2) the care or intensity with which a given technique is used (i.e. we assume that the contraceptive efficiency,  $e_i$ , of the  $i^{\text{th}}$  technique is a variable which may be increased at increased cost to the couple).<sup>20</sup> In this section we also restrict the choice of contraception strategies to "pure" strategies. Thus each strategy yields a different monthly probability of conception and hence a different point on the mean-variance curve (say, Curve A) in figure II-1.

Associated with each contraceptive strategy is an opportunity cost measured by the utility loss associated with the change in behavior required to implement that strategy. Some strategies cost money, some cost sexual satisfaction, some cost real or imagined decreases in physical health. The assumption of utility maximizing behavior implies that couples will choose the least costly strategy they are aware of in order to achieve any give level of  $p$  which yields  $\mu_N$  and its associated  $\sigma_N^2$ .

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Thus we rule out for now abortion and sterilization and we assume age at marriage to be exogenous. To emphasize this restriction we will use the term "contraception" in place of "fertility control" in discussing costs, strategies, etc. For a study of abortion as a means of fertility control see Potter (1972) or Keyfitz (1971) or for an economic analysis see a study in progress by Kramer (1973).

Suppose the couple's cost schedule for achieving any given  $p_i$  or its fertility outcome  $\mu_N$  is

$$F = F(\mu_N) \quad (12)$$

where  $F$  is the total cost of achieving  $\mu_N$  using the least costly contraceptive strategy. More specifically let the cost of the  $i^{\text{th}}$  contraceptive strategy be the simple linear function

$$F_i = \alpha_i + \beta_i B = \alpha_i + \beta_i (\mu_N^* - \mu_N) \quad (13)$$

where  $B$  is the difference between  $\mu_N^*$ , the couple's natural fertility, and  $\mu_N$ , the mean of the distribution of its expected fertility while using strategy  $i$ . Thus  $B$  is the expected number of births averted.

Equation (13) implies that the total cost of contraception using the  $i^{\text{th}}$  technique may be divided into two components: (1) a fixed cost,  $\alpha_i$ , ( $\alpha_i \geq 0$ ) which must be incurred if the  $i^{\text{th}}$  technique is to be used



at all, and (2) a variable cost,  $\beta_i B$ , which is proportional to the number of births averted by the use of technique  $i$ .<sup>21</sup> The term  $\beta_i$  ( $\beta_i \geq 0$ ) is the marginal cost per birth averted.<sup>22</sup>

It is important to stress that the classification of the contraception costs  $F_i$  of a given technique as fixed ( $\alpha_i$ ) or variable ( $\beta_i B$ ) is distinct from the classification of costs by their source. An economic (e.g. money or time), sociological (e.g. teachings of the Catholic church, deviation from class norms), psychological (e.g. interference with sexual pleasure, fear of adverse effects on health) or physiological (e.g. health) cost may be either fixed or variable.

Some factors, however, are more likely to affect  $\alpha_i$  than  $\beta_i$  or vice versa. Lack of contraceptive knowledge, for instance, is often cited as a reason for imperfect control. To the extent this is true, it is sensible to suppose that the acquisition of information about fertility control methods is costly. A characteristic of the cost of information is that it does not depend on the amount of use to which the information is put. It follows that the costs of information tend

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Each contraceptive strategy involves both the adoption of a contraceptive technique and the care and precision in its use. The adoption of technique  $i$  and its careless use results in less efficient contraception, a lower  $e_i$ , higher  $p_i$  and fewer births averted. Nearly all contraceptive techniques are capable of achieving a low  $e_i$  with careless use or a high  $e_i$  with proficient use.

<sup>22</sup>The linearity of the cost functions in (13) is not a particularly crucial assumption in the sense that the implications to be derived could be obtained under less restrictive assumptions.



to influence the fixed costs of contraception (the  $\alpha_i$ 's), but not  
the marginal costs (the  $\beta_i$ 's).<sup>23</sup> The cost to a Catholic of violating  
the Church's precepts with respect to the use of a contraceptive, for  
example, might be a once-and-for-all cost in which case  $\alpha_i$  is higher  
for Catholics than non-Catholics for all forbidden contraceptive tech-  
niques. Alternatively (or additionally), a Catholic may experience  
greater guilt the more intensively the technique is used, in which  
case  $\beta_i$  is higher to Catholics than to non-Catholics.<sup>24</sup>

The loss of sexual pleasure occasioned by contraception almost  
surely affects only the marginal costs of contraception and not the  
fixed costs. Thus, the number of births averted by condoms depends  
on how frequently and with what care condoms are used. The most ancient  
contraceptive techniques--abstinence or reduced coital frequency, and  
withdrawal--probably have zero fixed cost and rather high (psychological)  
marginal costs. By way of example, consider the choice between re-  
duced coital frequency or withdrawal as alternative contraceptive  
techniques.

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This argument should be qualified to the extent that information  
is acquired by a process of "learning by doing" or that information de-  
teriorates with disuse by a process of forgetting. In this case, the  
marginal cost of the  $i^{\text{th}}$  technique ( $\beta_i$ ) would tend to shift downward  
as the volume of use increases. Analytically, the learning hypothesis  
and the once-and-for-all hypothesis have the same implication, namely,  
that the average contraception cost per birth averted decreases as B  
increases.

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The cost to an individual Catholic of violating the Church's  
precepts may also be a function of the behavior of other Catholics or  
of other members of the society at large. Thus, the dynamics of dif-  
fusion of the pill use among Catholics might be interpreted, in part,  
as the progressive lowering in the cost of contraception to each indi-  
vidual Catholic as he or she sees others using the pill. Of course, the  
equivocation within the Church itself also presumably lowers the costs  
of using forbidden techniques (see Ryder and Westoff, 1971, Chapter 8  
for evidence on the effect of the Papal Encyclical on the contraceptive  
behavior of Catholics.).

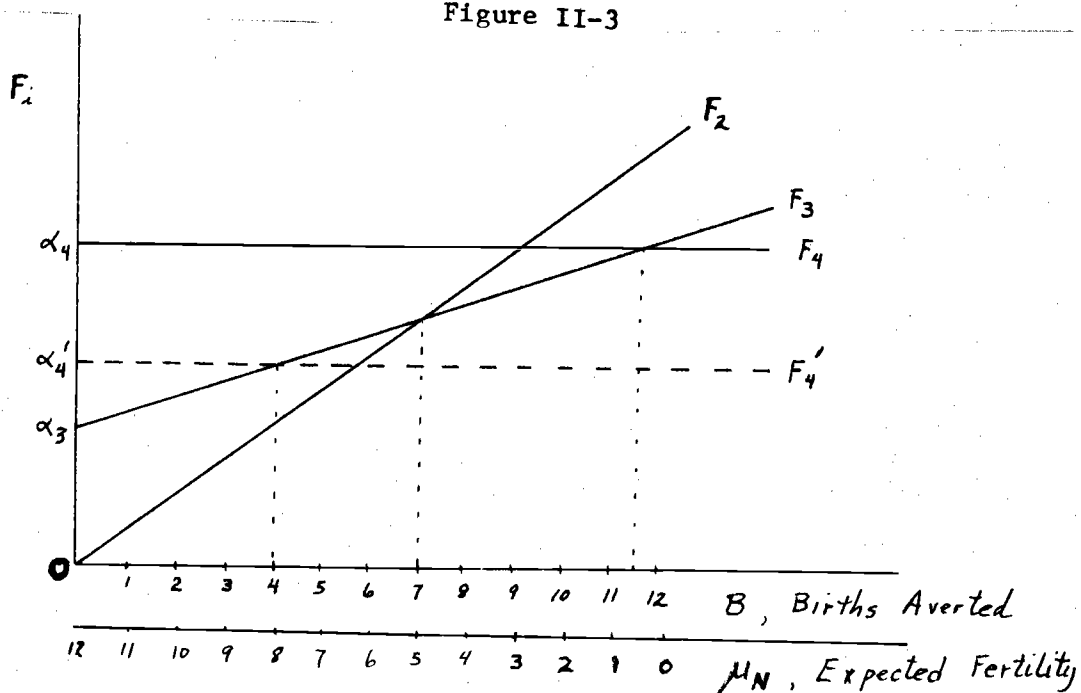
If a husband and wife use neither technique at all, they will expect to have  $\mu_N^0$  births and they will avert no births (i.e.  $B = 0$ ). The more persistently either technique is used, the smaller will be the expected fertility, the larger the expected number of births averted, and the larger the total contraception costs. Which technique is least costly depends solely on which technique has the lower marginal cost. If the marginal cost of reduced coital frequency exceeds the marginal cost of withdrawal, for example, the couple would not use the former technique whatever its desired number of averted births (note that we are limiting the choice at this point to pure strategies).

It is not always the case that one technique dominates all others for all possible fertility goals. Suppose, for example, that a third technique, condoms ( $i = 3$ ), has a lower marginal cost than does withdrawal ( $i = 2$ ) (i.e.  $\beta_3 < \beta_2$ ) but that it has a positive fixed cost (i.e.  $\alpha_3 > \alpha_2 = 0$ ). This situation is depicted in figure II-3 where line  $OF_2$  indicates the total contraception cost incurred if withdrawal is used to achieve each possible value of expected births averted (reading the upper horizontal scale from left to right) or, equivalently, at each possible level of expected fertility (reading the lower horizontal scale from right to left). Similarly, line  $\alpha_3 F_3$  shows the cost of using condoms to achieve each possible outcome.

To avert fewer than seven births (i.e. to have five or more children), the least cost strategy in figure II-3 is withdrawal. How-

ever, to avert more than seven births, condoms are a less costly contraceptive method. The point of equal costs (where lines  $OF_2$  and  $\alpha_3 F_3$  intersect) is called the "switching point": as the number of births to be averted rises, at some point (e.g. seven in the example illustrated in the figure) it becomes cheaper to incur the fixed costs or make the investment in an alternative technique--to switch to the technique with the lower marginal cost.

Figure II-3



Additional contraceptive techniques with still higher fixed costs and lower marginal costs may have lower average cost at higher numbers of averted births. The total cost function  $F$  (equation 12) is defined as the collection of line segments which represent the least-cost method

of achieving each number of averted births. It is the envelope of segments of the  $F_i$  curves in figure II-3. The limiting case would be a contraceptive with zero marginal cost (i.e. method  $i=4$  in the figure). A relatively low marginal cost appears to be a major advantage of modern contraceptive methods such as the pill and IUD. If line  $\alpha F_{44}$  represents such a technique in figure II-3, note that it represents the optimal contraceptive choice only if the couple wishes to have fewer than one child (as the figure happens to be drawn). That is, only couples wishing to avert nearly all potential births would select that high fixed cost, zero marginal cost technique.

#### iv. Optimal fertility control strategy

The preceding sections have discussed the separate elements in the determination of an optimal fertility control strategy. Using the simplifying assumption that a household must follow a pure strategy (i.e. must choose a constant value of  $p$  for the entire reproductive span), we derived a biological constraint on fertility choices illustrated by the mean-variance curve in figure II-1.

Next, we derived the expected utility of the household as a function of the mean and variance of fertility outcomes. This relationship is depicted in figure II-2. The fertility level with the highest net value,  $N^*$ , under conditions of certainty (i.e.  $\sigma_N^2 = 0$ ) and costless contraception is defined as the couple's "desired fertility". If, however, the couple is constrained to choose points on

the mean-variance curve in figure II-1 (but may choose any point without incurring any fertility control cost), the decrease in uncertainty associated with decreases in expected fertility makes it optimal to choose a level of expected fertility  $\mu_N^*$  that is somewhat lower than its desired fertility,  $N^*$ .

Finally, we introduced the (utility) costs of controlling fertility by means of specific contraceptive techniques. We derived a fertility control cost function as the envelope of least-cost segments of the cost curves of the individual techniques which is shown in figure II-3. Holding the couple's expected natural fertility,  $\mu_N^{\wedge}$ , constant, its total cost of fertility control,  $F$ , is larger, the smaller the level of its expected fertility,  $\mu_N$ .

The selection of the couple's optimal fertility control strategy involves two steps. First, for any given choice of  $\mu_N$  and, jointly,  $\sigma_N^2$ , the couple selects the least costly contraception technique-- say the  $i^{\text{th}}$  technique--with total cost  $F_i = \alpha_i + \beta_i (\mu_N^{\wedge} - \mu_N)$ . Second, it selects a level of  $\mu_N$  (and  $\sigma_N^2$ ) such that total expected utility (i.e., the expected net utility from children,  $E(V)$ , minus the total cost  $F_i$ ) is maximized. That is, using the specific functional forms in equations 9 and 13,

$$\max \{ E(V) - F_i \} = \max \left\{ a + b\mu_N + \frac{c}{2}\mu_N^2 + \frac{c}{2}\sigma_N^2 - \alpha_i - \beta_i(\mu_N^{\wedge} - \mu_N) \right\}. \quad (14)$$

The necessary condition for maximization with respect to  $\mu_N$  is

$$0 = b + c\mu_N + \frac{c}{2} \psi' + \beta_i. \quad (15)$$

Solving for  $\mu_N$  and again approximating  $\psi'$  (see equation 10) by the constant  $K > 0$

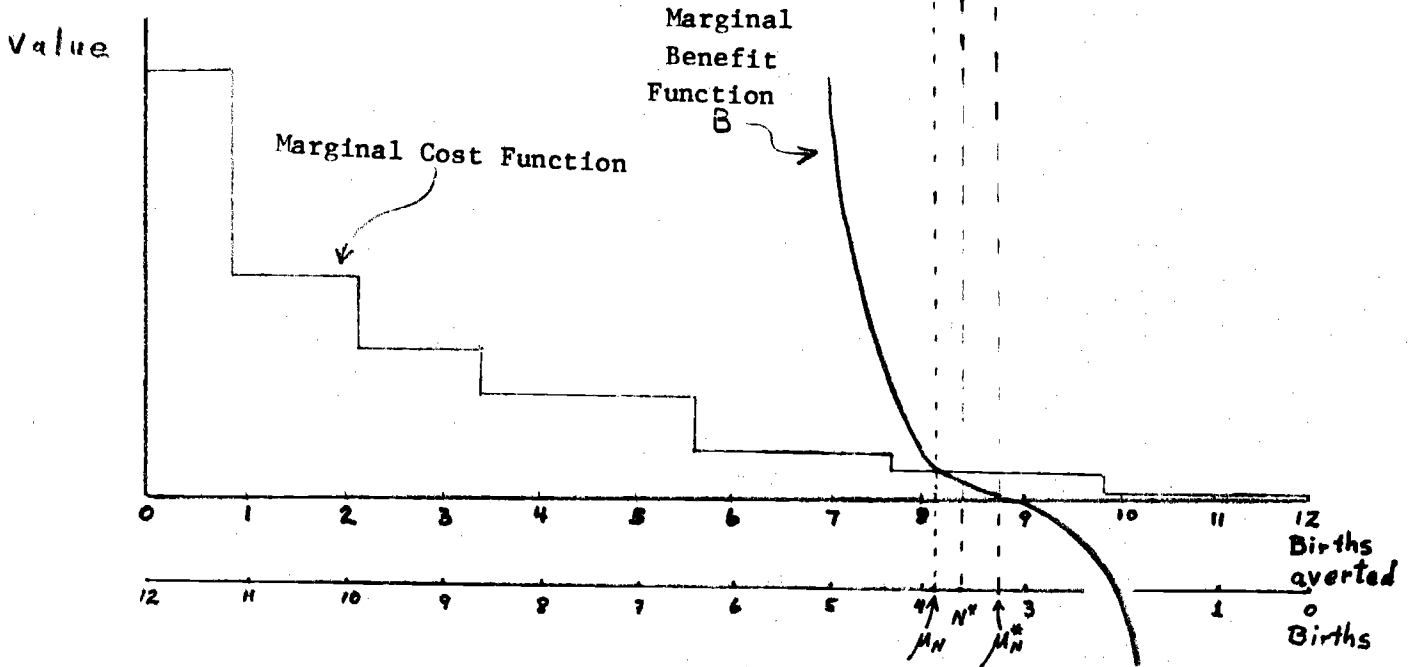
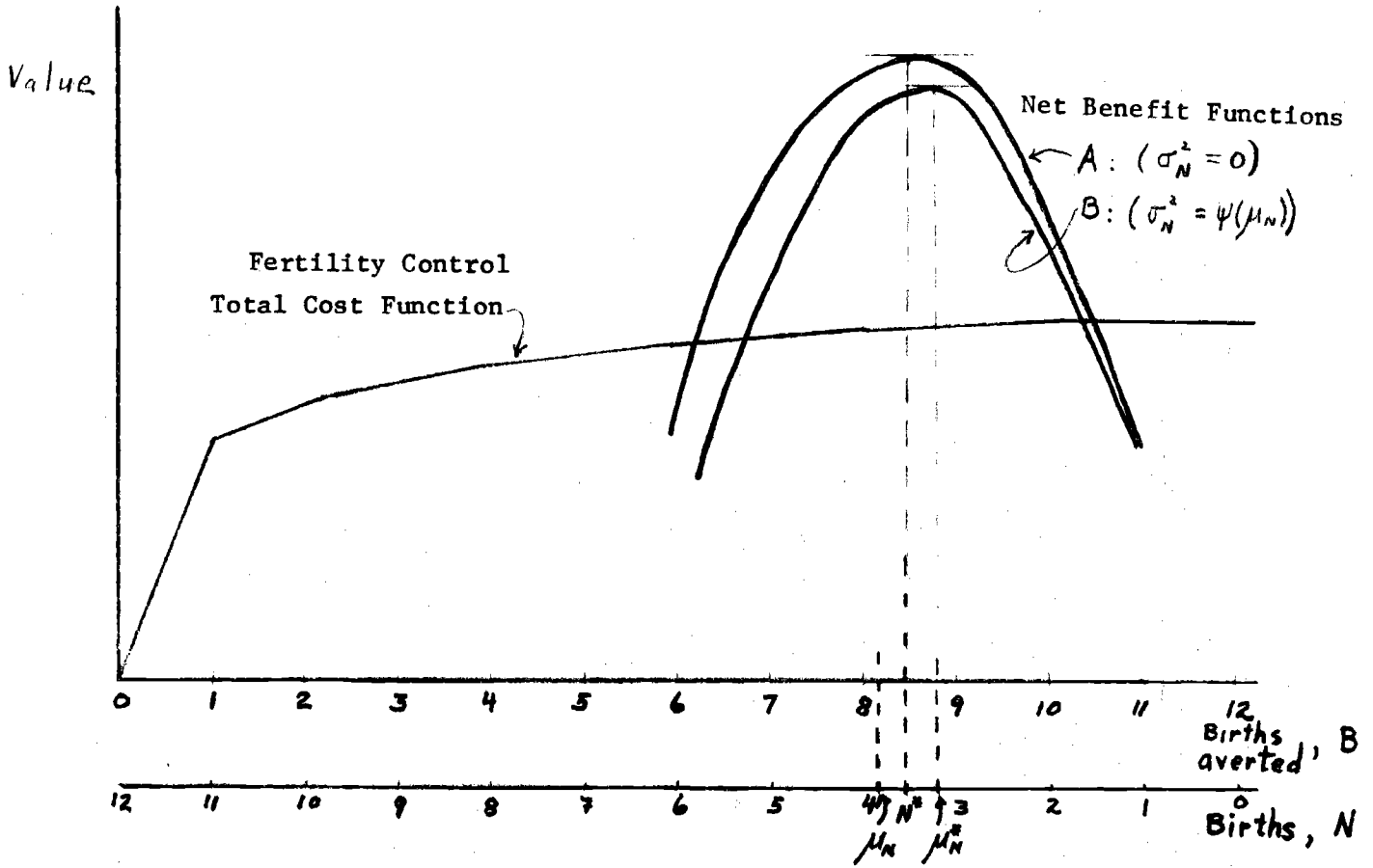
$$\mu_N = -\left( \frac{b}{c} + \frac{K}{2} + \frac{\beta_i}{c} \right) = N^* - \frac{K}{2} + \beta_i k = \mu_N^* + \beta_i k, \quad (16)$$

where  $k = \frac{-1}{c} > 0$ .

Equation 16 summarizes the influence on fertility outcomes of imperfect and costly fertility control. If the couple could select any level of fertility costlessly and with certainty, it would select  $N^*$ , its "desired fertility" determined by its tastes, economic circumstances and so forth. Since fertility outcomes involve an element of uncertainty (i.e. result from a stochastic process), there is a variance associated with each level of fertility, and since this uncertainty lowers the expected benefits or utility from each level of  $N$ , this uncertainty induces the couple to prefer an expected level of fertility  $\mu_N^*$  lower than  $N^*$  ( $\mu_N^* = N^* - \frac{K}{2}$ ). Furthermore since it is costly to avert births, the costs of fertility control further modify the optimal fertility outcome, raising the optimal above the level  $\mu_N^*$  ( $\mu_N = \mu_N^* + \beta k$ ). In effect,  $\beta_1$  is a per unit subsidy to child bearing because a couple may reduce its contraceptive cost by  $\beta_1$  with every additional birth it has.

The relationship between these various levels of fertility can be indicated by combining figures II-2 and II-3 (see figure II-4). Utility would be maximized at a level of  $N = N^*$  if the costs of fertility control were zero and the uncertainty about fertility outcomes were ignored (e.g. the maximum of the net benefit function A is at a level of  $N = N^*$ ). The presence of uncertainty modifies the optimum by raising the peak of the net benefit function (function B) to the level  $N = \mu_N^*$ . The presence of fertility control costs further modifies the optimum by lowering the preferred  $N$  to  $\mu_N$ , the intersection of the marginal cost of fertility control and the marginal benefit from fertility control.

Figure II-4



### III Pure and Mixed Contraceptive Strategies

In the previous section we restricted the discussion of contraceptive strategies to a "pure" strategy defined as one in which the couple selects some specific contraceptive technique and uses it with some specific amount of care throughout the reproductive span of  $T$  months. The pure strategy model, in which the couple uses one technique continuously implies a constant monthly probability of conception  $p$  over the fertile months in the reproductive span. This implication is essential for the Markov renewal process on which are based the equations for the mean and variance of fertility outcomes (equations 1 and 2) and the mean-variance curves depicted in figure II-1. So the pure strategy model lends itself to a simple analytical structure and represents a boundary on the relationship between mean and variance of fertility. As we have noted, this Markov renewal model also underlies much of the important analytical work done in the past decade in mathematical demography.

However, it is evident that in reality a couple is not restricted from altering its contraceptive technique over the reproductive span. Even in the context of a lifetime strategy which could be mapped out initially and carried out over time, a couple might choose to use different contraceptive techniques (including no contraception) in various segments of that span of time. Furthermore, since the discussion of the benefits of fertility control emphasized that under fairly general conditions couples prefer to reduce the variance in their expected fertility,



it is economically sensible as well as technically feasible for couples to select a mixed contraceptive strategy which may result in a lower  $\sigma^2_N$  for any level of expected fertility  $\mu$ <sup>25</sup>.

For example, if a couple used the oral contraceptive at its average observed use-effectiveness throughout its twenty year reproductive span, Table II-1 indicates that the distribution of its expected fertility has a mean of 0.19 births and a variance of 0.18 (point a on the mean-variance curve in figure II-1). If this couple wished to have about three children, it could use a less effective technique or use the pill somewhat carelessly, thereby achieving a monthly probability of conception of 0.0182 which over a twenty year span would yield an expected fertility of 3.4 and a variance of 1.99 (point c on the mean-variance curve in figure II-1). The couple could also achieve a mean fertility of 3.4 however by combining periods of

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The economic rationale, offered in Section II, for preferring a reduced variance in fertility was that any deviation in actual fertility from the desired level of fertility implies a reduction in utility. The greater the variance  $\sigma^2_N$ , the greater the likelihood that the discrepancy between actual and desired fertility will be relatively large.

There is an additional economic reason for generally preferring a lower variance in the distribution of expected outcomes, risk preference aside. The more certain the couple is about the number of children it will eventually have, the more efficiently it can optimize on the allocation of its resources. The couple which is more certain about the timing and number of its children can more efficiently plan its savings pattern, select an optimal size home, automobile, etc., plan the labor force behavior of the wife, and so forth. The same principles apply here as in the case of a firm which can achieve lower average cost of production if its rate of output is constant over the long run than if its rate of output varies significantly from season to season or from year to year.

pill use with periods of time in which no contraception was used. The result would be a mean of 3.4 children and a variance considerably below that indicated by point c on the mean-variance curve. Indeed the variance would be no greater than that indicated by point f, the variance associated with the use of no contraception over the entire twenty year time span.

26  
Furthermore, the pure strategy implies that the births will arrive at random intervals over the twenty year span, while the mixed strategy permits the couple to achieve the same number of children with considerable control over their spacing. 27 By combining the use of a highly effective

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The logical extreme would be a mixed strategy in which natural fertility (no contraception) is pursued in those segments of the reproductive span in which a birth is desired and perfect contraception (i.e.  $e_1 = 0$ ) at all other times. This strategy would enable a normally fecund couple to achieve any given number of children fewer than, say, five with virtual certainty and would also enable them to approximate many plausible desired spacing patterns fairly closely (see Potter and Sakoda, 1967).

Such a strategy is not only a logical possibility, but it is also technically feasible since the monthly probability may be set to zero at any time by reducing coital frequency to zero. The fact that couples do not appear to follow this "perfect contraception" strategy suggests that the problem of fertility control is not a matter of technical feasibility. The biological constraint on fertility choices must be considered simultaneously with other constraints on behavior, with fertility goals viewed as competing with other family goals.

27  
In the pure strategy case the variance of the interval of time between successive births,  $\sigma_t^2$ , is inversely related to  $p$  and hence  $\mu_N$ . Thus, reduction in expected fertility along the mean-variance curve is accompanied by an increase in the variance  $\sigma_t^2$ .

technique with periods of no contraception, a couple can achieve its desired number of children with a relatively low variance  $\sigma_N^2$  and relatively little uncertainty about the spacing of its children.

The mean-variance curve in figure II-1 represents the biological constraint on the distribution of expected fertility when a particular monthly birth probability persists for the entire span of T months. By using a mixed strategy, combining contraception with periods of no contraception, the biological constraint is no longer an effective constraint-- the couple can move off the mean-variance curve toward the horizontal axis representing a distribution of fertility outcomes of mean  $\mu_N$  and zero variance. The more efficient the contraceptive technique chosen during periods of contraception, the smaller is the achievable variance  $\sigma_N^2$  for any given mean  $\mu_N$ . Thus the more efficient the contraceptive technique chosen, the weaker is the relationship between  $\mu_N$  and  $\sigma_N^2$ ,<sup>28</sup> and the smaller is the incentive to lower the mean fertility as a mechanism for reducing uncertainty or variance.

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The mixed strategy (defined in terms of using one specific technique while contracepting and no contraception otherwise) is feasible only when the expected number of births from the continuous use of the technique is less than the number of children the couple desires. So this form of

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<sup>28</sup> Some evidence that the correlation between mean and variance of fertility is positive is found in the 1960 U.S. Census of the Population. Grouping white women married once and husband present into cells defined by husband's occupation (8 categories), husband's education (5 categories) and wife's education (3 categories), the unweighted simple correlation between  $\mu_N$  and  $\sigma_N^2$  across cells is 0.89 for women aged 45-54 and 0.77 for women aged 35-44. If the younger cohort used better contraceptive methods on the average, then the reduction in this correlation across cohorts is consistent with the implication of a weaker correlation among users of better contraception.

mixed strategy is more likely to be used the greater is the efficiency of the contraceptive technique chosen.

In this discussion of mixed strategies of contraception we have focused upon one particular type of mixed strategy-- that of going on and going off one contraceptive technique. Although shifting from contraceptive technique to technique is another possibility, the theoretical discussion of the costs of contraception suggested that this would not be the case. The fixed costs associated with the adoption of modern techniques would inhibit technique switching. Consistent with the model's implication, evidence from the 1965 National Fertility Survey suggests that technique switching has not been a prevalent practice in the U.S. in the past two decades. Ranking contraceptive methods by their mean monthly probability of conception (as indicated in Table II-1) and limiting the subsample to women who had used some contraception in each of their first three birth intervals, Michael (1973) found that the correlation among techniques used across the three pregnancy intervals was quite high (ranging from 0.57 to 0.97) for non-Catholic women partitioned by color and age cohort.<sup>29</sup> Ryder and Westoff (1971) study the relationship

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<sup>29</sup>See Michael (1973) Table 4. One note of caution. The NFS data are oriented by the woman's pregnancy intervals, so Michael had information on only the best technique used by the woman in each interval. He could therefore identify switches in contraceptive techniques from pregnancy interval to interval, but not from technique to technique within a given pregnancy interval.

between use and non-use of contraception across intervals and the relationship between contraceptive failures in successive intervals. They find considerable continuity of contraceptive status across intervals both in terms of whether a woman does or does not use contraception in successive intervals and in terms of the degree of success of use across intervals.<sup>30</sup>

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For example, 90% of women who used some contraceptive technique in the first pregnancy interval (from marriage to first pregnancy) used a contraceptive in the second interval while only 36% of non-users in the first interval used a contraceptive in the second interval. Similar percentages are found for each successive pair of intervals (i.e. from the fourth to the fifth interval the comparable percentages are 95% and 18%). See Ryder and Westoff (1971), Table IX-19 p. 255. Or, 95% of the women who had used contraception in each of the first three pregnancy intervals used contraception in the fourth interval while only 13% of women who had not used contraception in any of the first three intervals used contraception in the fourth (see Ryder and Westoff, (1971) Table IX-23, p. 260).

Evidence of consistency of use across intervals is indicated by the following rather remarkable statistic: of women who used a contraceptive "successfully" in the first three pregnancy intervals, 20% experienced a contraceptive "failure" in the fourth pregnancy interval, while of those who had experienced a contraceptive "failure" in each of the first three intervals, 77% experienced a "failure" again in their fourth interval (see Ryder and Westoff for definitions of success and failure).

This statistic and others support quite strongly, we think, the contention that couples act as if they adopt a lifetime strategy toward contraception and that that strategy involves considerable continuity in the use of a technique throughout a lifetime. (The Princeton Study begun in 1957 also suggested that across-interval changes in fertility control are "clearly not a matter of couples shifting from ineffective to effective methods " of contraception. See Westoff, Potter and Sagi 1963 (pp. 232-235).)

#### IV Contraception and Fertility Outcomes

In the model described in Section II the household's number of children,  $N$ , is a random variable. The household adopts a contraceptive strategy which yields a particular value of  $p$ , the monthly probability of conception. Given  $p$  as a known and fixed parameter, the household has an ex ante distribution of fertility characterized by a mean  $\mu_N$  and variance  $\sigma_N^2$ .

The discussion has focused on the ex ante distribution of fertility outcomes for a single household, but in our empirical analysis we focus on the corresponding distribution for relatively homogenous groups of households. It is assumed that the observed mean and variance in births among households with relatively homogeneous demographic-economic characteristics reflect the mean and variance of the distribution of fertility outcomes faced by each of the households in that group. Recall that the equation for the variance in number of children (equation 2) assumed that the unprotected monthly probability of conception and the length of the period of infertility were constant over the couple's reproductive lifetime. To apply the model across households implies not only constancy of these parameters over time for a given household, but also constancy across households. Heckman and Willis (1973) deal explicitly with the problem of estimating the average monthly probability  $p$  in heterogeneous groups of households. For our purposes, we will not pursue this issue.<sup>31</sup>

<sup>31</sup> Consider two populations of fecund, non-pregnant women with identical mean monthly probabilities of conception,  $\bar{p}$ . One population is homogeneous in the sense that  $p$  is identical for all members of the population and the other is heterogeneous in the sense that  $p$  varies across women according to some distribution with positive variance. It is known that the mean waiting time to conception in the heterogeneous population will be longer and the average birth rate lower than in the homogeneous population, and that this difference is a function of the distribution of  $p$  in the heterogeneous population (see, for example, Sheps, 1964).

The model in Section II was set out in a lifetime context and considered fertility control in terms of a lifetime strategy. Accordingly, in our empirical work we frequently use information about contraceptive behavior at one point in the couple's marriage as an indication or index of the lifetime contraceptive strategy. As we indicated in Section III there is considerable evidence that contraceptive behavior is not characterized by switching from contraceptive technique to technique over the lifetime. Consequently, in this section we will distinguish couples either by the best contraceptive technique used in the time interval from marriage to their first pregnancy or by the best technique used at any time in their marriage. <sup>32</sup>

<sup>32</sup> In addition to the evidence cited above regarding consistency of technique use between pregnancy intervals, the following table indicates the percentage of users of a specific contraceptive technique in the first interval who also used that technique in the second interval. The second column indicates the percentage who used either that same technique or no contraception in the second interval. These figures pertain to white non-Catholic women age 40-44 from the 1965 NFS.

First Interval: Technique Used	Second Interval:	
	% Using Same Technique	% Using Same Technique or No Contraception
diaphragm	77%	89%
condom	72	82
withdrawal	78	85
jelly, foam, suppository	61	67
rhythm	65	85
douche	62	73

The table indicates, for example, that of those couples which used the diaphragm in the first pregnancy interval, 77% also used the diaphragm in the second pregnancy interval. Furthermore, of that same group another 12% used no contraception in the second pregnancy interval thus a total of 89% used either that same contraceptive method or no method in the second interval. (The second interval here is defined as either the period of time from the first to the second pregnancy or from the first pregnancy to the time of the survey if no second pregnancy occurred.)

Since these data were collected by interview at the time these women were 40-44 years of age and pertain to periods of time shortly after marriage, there may be a tendency to give the same response for successive intervals. If so, these percentages overstate the consistency of technique selection across intervals.

In this and the following section we use the 1965 National Fertility Survey which was conducted by the Office of Population Research at Princeton University. <sup>33</sup> This cross-section survey of some 5600 women aged 55 and under, currently married and living with their spouse, contains information on the specific contraceptive technique used in each pregnancy interval, as well as information on the couple's actual fertility outcome. In this section we use this data set to document the relationship between contraception use and fertility outcomes. Since we are interested in studying the variance in fertility we group the data into cells and study between-cell differences in observed behavior.

In this section we explore how contraception behavior is related to the observed distribution of fertility across groups of households; we do not attempt to explain why couples differ in their desired fertility or in the dispersion of their fertility. Although we indicated in Section II that the model is capable of treating contraception choice and fertility control choice in a simultaneous system of equations, we do not attempt to estimate the parameters of those

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We wish to thank Charles F. Westoff and Norman B. Ryder for their help in obtaining these data. Our previously published research from this data set (Michael 1973) used a small data file obtained from Professor Westoff. Our current research uses the publicly available data tape from the 1965 NFS which was acquired through Professor Larry Bumpass. The data set is fully described in Ryder and Westoff, Reproduction in the United States 1965, Princeton University Press, 1971.



structural equations. In the tables below, we partition the data set by household characteristics including color and religion and we use either the wife's education or expected fertility and either age of wife or marriage duration to isolate relatively homogeneous groups of households. In the context of these homogeneous groups we study contraception strategies as the mechanism for affecting fertility outcomes.

Figure IV-1 indicates the frequency distribution of live births for women age 35-55 for groups defined by wife's level of schooling, color and religion. Among the white non-Catholics (figure IV-1a) the distributions appear to be less positively skewed and less dispersed (or more peaked) among women with higher levels of schooling.<sup>34</sup> By contrast, the distributions for nonwhite non-Catholics (figure IV-1b) are considerably less peaked and more skewed. Among this latter group the level of schooling does not distinguish the frequency distributions so clearly, although the percentage of households with, say, seven or more births appears to decline as the level of schooling rises. Among Catholics (figure IV-1c) the distributions are somewhat less dispersed than among the nonwhite non-Catholics, and the level of schooling does not appear to influence

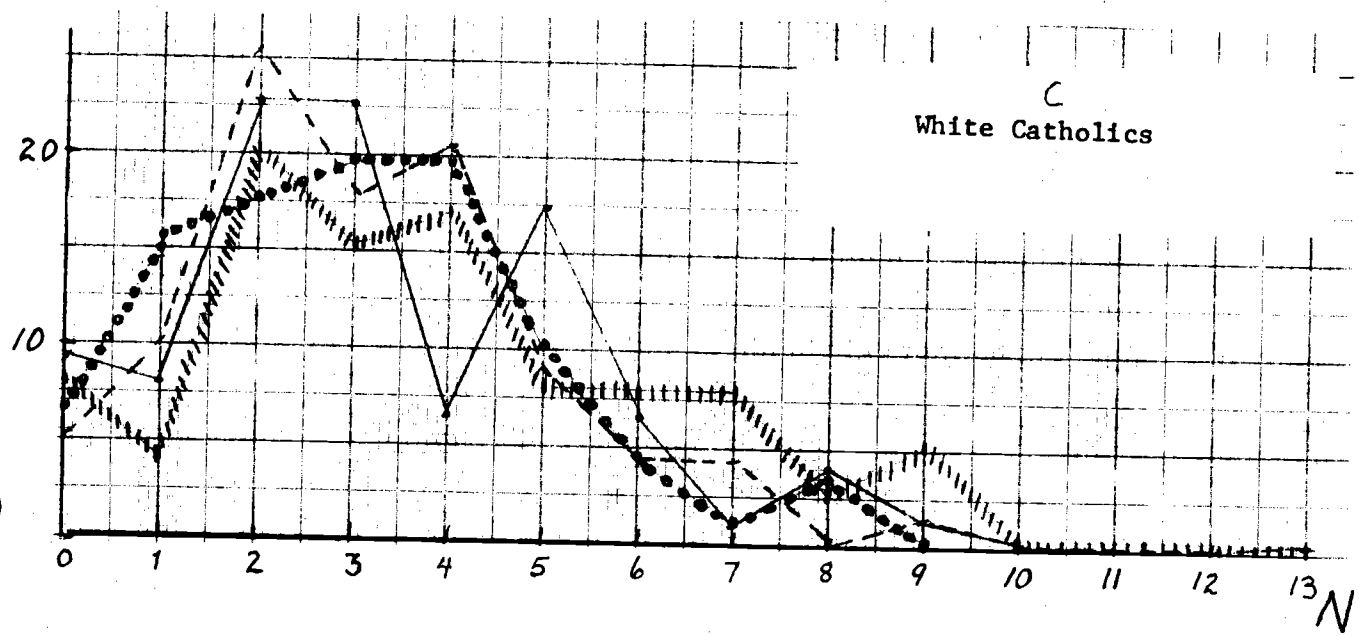
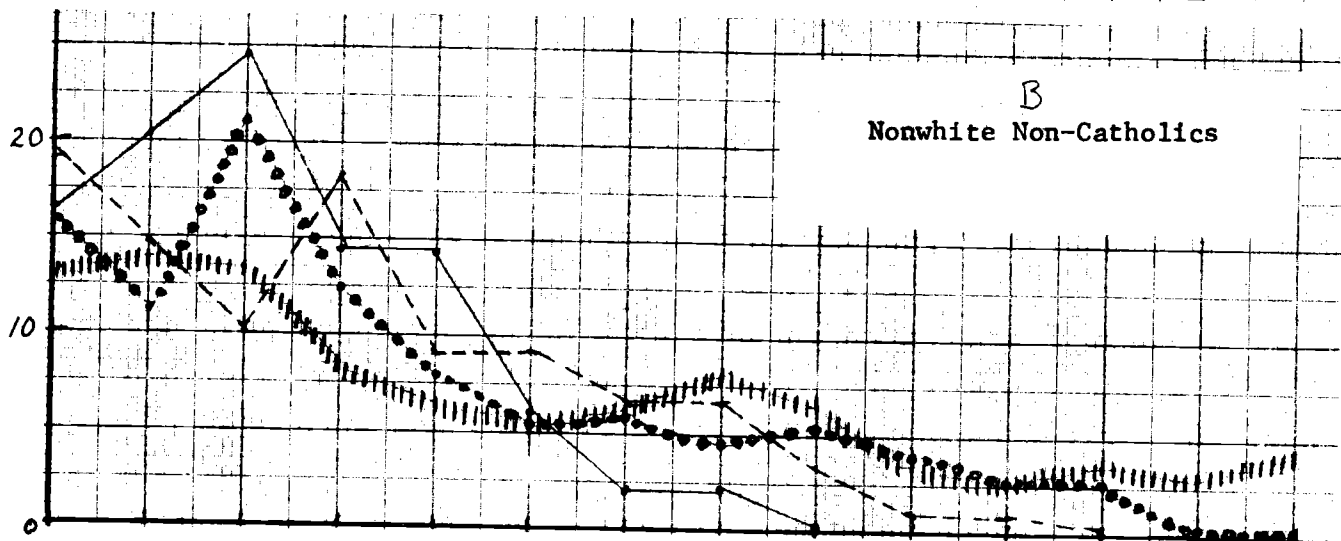
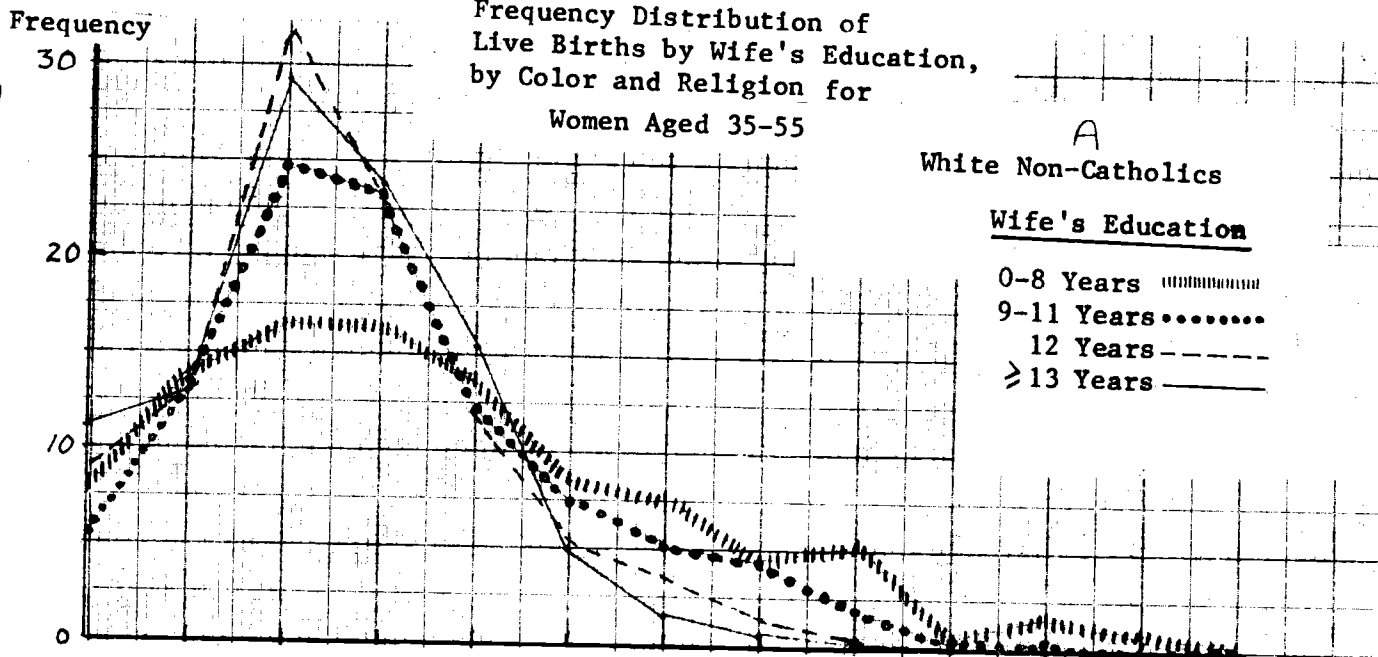
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A few Chi-squared tests have been performed on pairs of distributions of live births for groups of white non-Catholics with different education levels from specific 5-year cohorts. These tests imply rejection (at  $\alpha = .05$ ) of the hypothesis that the grade school women's distribution of live births and the college women's distribution of live births might have been drawn from the same population.

For example,  $\chi^2 = 36.8$  with 12 degrees of freedom for a comparison of  $\leq 8$  years versus  $\geq 13$  years of schooling for women aged 40-44. The critical value for  $\chi^2$  with 12 d.f. at  $\alpha = .01$  is 26.2.

Figure IV-1  
 Frequency Distribution of  
 Live Births by Wife's Education,  
 by Color and Religion for  
 Women Aged 35-55



the distributions systematically.

Among white non-Catholics there appear to be distinctly different distributions of live births by wife's education. These differences are further emphasized by the following table derived from Appendix Table 1.

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Percentage of Women with Six or More Live Births  
(Calculated from Appendix Table 1)  
White Non-Catholic Women

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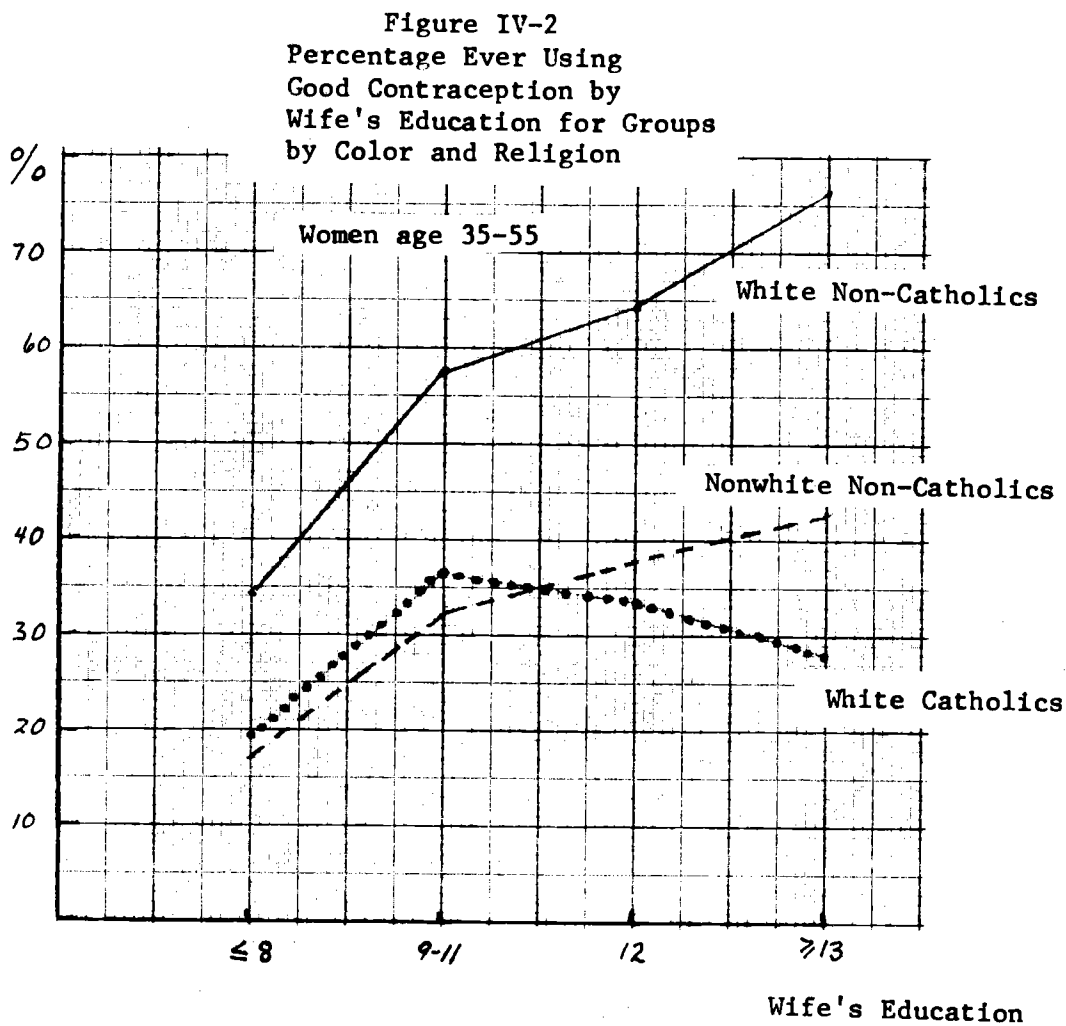
Wife's Education	Wife's Age			
	35-39	40-44	45-49	50-54
≤ 8	22.7	22.4	23.4	19.6
9 - 11	11.5	15.6	5.7	13.1
12	6.8	6.7	4.8	1.3
≥ 13	2.7	1.6	1.6	2.2
Total	8.8	10.1	7.9	8.7

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Also, the groups of nonwhites and Catholics appear from figure IV-1 to have considerably different frequency distributions than the most highly schooled white non-Catholic groups. The discussion above has suggested that different contraceptive strategies yield different distributions of fertility, so we expect to find that groups which differ in the distribution of their actual fertility also differ in their contraception behavior. The NFS data contain information on the particular contraceptive method used by each woman in each pregnancy interval. The data do not indicate

how extensively, regularly or carefully a contraceptive method was used; we know only that the woman indicated that between the time she married and the time she first became pregnant, for example, she used contraceptive method 1 (including no method at all).

Figure IV-2 indicates the percentage of women in each education,



color and religion group which ever used a "good"<sup>35</sup> contraceptive technique. The relationship between these percentages and the distributions of births indicated in figure IV-1 is striking. Among white non-Catholics, the percentage of users of "good" contraceptive techniques rises significantly with the wife's level of schooling, and this is also the group for which schooling most clearly distinguishes the distributions of births. The percentages rise less rapidly by education among the nonwhites, and for these groups the frequency distributions of births are less clearly delineated by education. Among Catholics neither the percentage of users of good contraceptive techniques nor the frequency distribution of births seems to be closely related to wife's education. Thus, in comparisons among and within color and religion groups, there appears to be a quite consistent relationship: groups characterized by a relatively high percentage of users of good contraception are also characterized by relatively low dispersion in fertility.

Table IV -1 summarizes this same relationship. The table indicates for each education, color and religion group the percentage of couples ever using "good" contraception and the actual mean and standard deviation of live births. Groups characterized by a high percentage of users

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<sup>35</sup>Throughout this section, "good" contraception is defined to include pill, IUD, condom and diaphragm. For older cohorts the best available contraceptive methods were the condom and diaphragm while for the younger cohorts the more reliable pill and IUD were also available. Since we are attempting at this time to distinguish use of highly reliable from less reliable techniques, the distinction was arbitrarily made as indicated. All contraceptive methods other than the pill, IUD, condom and diaphragm are categorized as "poor" methods. The use of no contraception is called "none" and is distinguished from the "poor" methods.

Table IV-1  
 Percentage of Couples Using "Good" Contraception,  
 Mean and Standard Deviation of Live Births; by Wife's  
 Education, Age, Color and Religion

		Education of Wife				Total
		≤ 8	9-11	12	≥ 13	
<u>White Non-Catholics</u>						
Age 35-44	% good *	37.4	62.9	67.4	79.9	64.9
	$\mu_{\text{N}}$	3.658	3.284	2.738	2.583	2.946
	$\sigma_{\text{N}}$	2.635	1.942	1.567	1.472	1.849
	(n)**	(155)	(232)	(503)	(230)	(1120)
Age 45-54	% good	29.9	45.4	57.2	67.8	50.9
	$\mu_{\text{N}}$	3.645	2.616	2.136	2.189	2.579
	$\sigma_{\text{N}}$	2.459	2.004	1.593	1.415	1.962
	(n)	(121)	(99)	(206)	(106)	(532)
<u>Nonwhite Non-Catholics</u>						
Age 35-44	% good	21.7	35.3	41.8	46.9	32.7
	$\mu_{\text{N}}$	4.968	3.613	3.493	2.531	3.980
	$\sigma_{\text{N}}$	3.855	3.231	2.636	1.796	3.355
	(n)	(124)	(119)	(67)	(32)	(342)
Age 45-54	% good	11.3	21.2	25.0	35.3	17.4
	$\mu_{\text{N}}$	4.474	3.848	1.800	1.765	3.754
	$\sigma_{\text{N}}$	4.309	3.173	1.765	1.480	3.797
	(n)	(97)	(33)	(20)	(17)	(167)
<u>White Catholics</u>						
Age 35-44	% good	18.0	41.6	31.7	28.3	31.6
	$\mu_{\text{N}}$	4.213	3.217	3.353	3.567	3.467
	$\sigma_{\text{N}}$	2.659	1.831	2.027	2.126	2.111
	(n)	(61)	(106)	(218)	(60)	(445)
Age 45-54	% good	20.9	23.0	39.3	26.7	28.8
	$\mu_{\text{N}}$	3.646	2.564	2.705	2.133	2.896
	$\sigma_{\text{N}}$	2.497	1.832	1.395	1.598	1.949
	(n)	(48)	(39)	(61)	(15)	(163)

\* percentage of couples in the cell which ever used "good" contraception (i.e. pill, IUD, condom or diaphragm).

\*\* cell size

of good contraception are characterized by relatively lower variance and somewhat lower mean fertility. Simple correlations across these groups between the percentage of couples which ever used good contraception (at any time since marriage), "% good," and the mean fertility  $\mu_N$ , and between "% good" and the standard deviation  $\sigma_N$  are consistently negative. As emphasized in the theoretical discussion of the mean-variance curve, the observed correlation between the mean and variance (or standard deviation) in fertility is positive in all cases:

Simple correlation coefficients between % Good (the percent of couples which ever used good contraception),  $\mu_N$  (the group's mean number of live births), and  $\sigma_N$  (the standard deviation in the group's number of live births)

	Simple Correlation			(n)
	% Good, $\mu_N$	% Good, $\sigma_N$	$\mu_N, \sigma_N$	
White non-Catholics	-.647	-.896	.884	(8)
Nonwhite non-Catholics	-.508	-.675	.952	(8)
White Catholics	-.334	-.700	.879	(8)
All Combined	-.475	-.601	.864	(24)

The relationship between contraception behavior and fertility outcomes indicated by Table IV-1 is somewhat circumstantial--groups of households which had relatively high rates of use of good contraception also had relatively low mean and variance in fertility. To relate contraception use to fertility outcomes more directly we partition the age, education groups of white non-Catholics by their contraceptive strategies. Table IV-2, for example, indicates the separate frequency distribution of live births for users of good contraception and users of poor contraception in the first pregnancy interval by wife's age for women with 12 years of schooling. That is, in Table IV-2 two of the cells in Table IV-1

(defined by wife's education equal to 12 years for women age 35-44 and age 45-54) are partitioned by the contraceptive technique used in the first birth interval.

Table IV-2

Frequency Distribution of Live Births for White Non-Catholic Women with 12 Years of Education by Contraceptive Method Used in the First Pregnancy Interval and by Wife's Age

Contraceptive Method	Number of Live Births					cell size
	0	1	2-4	5-6	≥ 7	
<u>Age 35-44</u>						
good *	1.4	7.0	81.2	8.9	1.4	(213)
poor	2.0	13.9	69.3	11.9	3.0	(101)
none	15.3	14.8	55.6	11.1	3.2	(189)
total	6.8	11.3	69.2	10.3	2.4	(503)
<u>Age 45-54</u>						
good	0.0	16.7	80.3	3.0	0.0	(66)
poor	0.0	12.8	71.8	7.7	7.7	(39)
none	31.7	20.8	41.6	5.0	1.0	(101)
total	15.5	18.0	59.7	4.9	1.9	(206)

\* "good" methods are defined to be pill, IUD, condom and diaphragm; "poor" methods include all other contraceptive methods excluding abstinence

For each of the two age groups the distribution of live births is considerably less dispersed among the users of good contraception than among users of poor contraception (e.g. the percentage of households with five or more live births was 10.3 and 3.0 among users of good contraception while the percentages were 14.9 and 15.4 among users of poor contraception). Notice, too, the large percentage of non-users in the first interval which had zero live births. Presumably a relatively large fraction of the users of no contraception knew themselves to be sterile or subfecund. In the



terminology of the model developed in Section II, couples with a relatively low natural fecundity and a low expected fertility need avert fewer births to achieve any given level of desired fertility. So these couples have less incentive to use any contraception in general and less incentive to adopt high fixed cost techniques in particular.

Table IV-3 also indicates the relationship between contraception use and fertility outcomes. While Table IV-2 shows a frequency distribution of live births by contraception use for two of the cells of white non-Catholic women from Table IV-1, Table IV-3 indicates the mean and standard deviation of the live births by contraception use for each of the 5 cells of white non-Catholic women age 35-44 from Table IV-1.

Compare the fertility behavior of the "good" and "poor" contraceptors for some given level of schooling in Table IV-3. In terms of mean fertility, couples which used a "good" contraceptive method (pill, IUD, condom or diaphragm) in the first birth interval had somewhat lower mean fertility than couples which used relatively "poor" contraceptive methods. As Panel B indicates, however, very few of the differences in means are statistically significant. By contrast, the comparison of differences in the standard deviation of the fertility outcomes does exhibit statistical significance--the users of poor contraception have appreciably higher variation in their fertility outcomes than do users of good contraception.<sup>36</sup>

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<sup>36</sup>Note that in the tests of significance of the variances, the few pairwise comparisons which were not significant involved the relatively small cells containing 40 or fewer observations.

Table IV-3

Mean Number of Live Births,  $\mu_N$ , and Standard Deviation of Number of Live Births,  $\sigma_N$ , for White Non-Catholic Women by Wife's Education and by the Contraceptive Method (Good, Poor, None) Used in the First Birth Interval, for Women Aged 35-44

(A)

Contraceptive Method Used in First Birth Interval (Cell Size)		Education of Wife				Total
		$\leq 8$	9-11	12	$\geq 13$	
Good	$\mu_N$	3.769	3.069	2.784	2.887	2.920
	$\sigma_N$	2.065	1.476	1.274	1.076	1.334
	(n)	(26)	(72)	(213)	(124)	(435)
Poor	$\mu_N$	3.840	3.800	2.941	2.706	3.185
	$\sigma_N$	2.444	2.028	1.515	1.488	1.802
	(n)	(25)	(40)	(101)	(34)	(200)
None	$\mu_N$	3.587	3.242	2.577	2.000	2.872
	$\sigma_N$	2.817	2.134	1.860	1.854	2.224
	(n)	(104)	(120)	(189)	(72)	(485)
Total	$\mu_N$	3.658	3.284	2.738	2.583	2.946
	$\sigma_N$	2.635	1.942	1.567	1.472	1.849
	(n)	(155)	(232)	(503)	(230)	(1120)

(B)

Tests of Statistical Significance of Differences in Means and Variances in Number of Live Births, for Specific Pairs of Cells

Difference by Contraceptive Method	Education of Wife			
	$\leq 8$	9-11	12	$\geq 13$
Test of Difference in Means (Student's t-test)				
Good vs. poor	0.11	2.03**	0.91	0.67
Poor vs. none	0.45	1.47	1.82*	2.08**
Good vs. none	0.37	0.64	1.20	3.70***
Tests of Difference in Variance (F-test)				
Good vs. poor	1.40	1.88**	1.42**	1.91***
Poor vs. none	1.33	1.11	1.50**	1.56
Good vs. none	1.86**	2.09***	2.14***	2.97***

Test of Differences in Mean and Variance of Number of Live Births by Wife's Education for Users of Good Contraception Only

Difference by Wife's Education	Difference in		Difference by Wife's Education	Difference in	
	Means (t-test)	Variance (F-test)		Means (t-test)	Variance (F-test)
$\leq 8$ vs. 9-11	1.59	1.95**	9-11 vs. 12	1.46	1.34**
$\leq 8$ vs. 12	2.35**	2.63***	9-11 vs. $\geq 13$	0.92	1.88***
$\leq 8$ vs. $\geq 13$	2.12**	3.67***	12 vs. $\geq 13$	0.78	1.40**

- \* Implies statistical difference at  $\alpha = .10$  (two-tailed t-test).  
 \*\* Implies statistical difference at  $\alpha = .05$  (two-tailed t-test).  
 \*\*\* Implies statistical difference at  $\alpha = .01$  (two-tailed t-test).

The lack of a stronger association between contraception use and mean fertility in Table IV-3 is somewhat surprising. However, recall that equation 16 emphasized two opposing forces influencing optimal mean fertility. The marginal costs of fertility control raised optimal mean fertility ( $\beta_1 k$ ), although the positive relationship between the mean and variance lowered optimal fertility ( $-K/2$ ) as a mechanism for reducing the uncertainty about the number of births. We showed that couples wishing to avert more births would be induced to adopt better (higher fixed cost, lower marginal cost) contraceptive techniques. We also suggested that users of better techniques are more likely to use a mixed strategy of contraception which implies a weaker relationship between the mean and variance of fertility. So users of good contraception are expected to have lower marginal costs of fertility control ( a lower  $\beta_1$  ) and also a weaker relationship between mean and variance ( a lower  $K$  ). Consequently, if wife's education sorts couples by their desired fertility in Table IV-3, the further partitioning by good or poor contraception may not have a systematic effect on mean fertility. The users of good contraception have lower marginal cost of averting births, but less incentive to reduce mean fertility as a mechanism for lowering the uncertainty or variance of fertility. As the relationships in Tables IV-1--Table IV-3 indicate, across relatively homogeneous

groups there is a negative relationship between the use of good contraception and mean fertility, but within the homogenous groups, the use of good contraception systematically affects only the variance of fertility outcomes. Couples wishing to avert relatively more births have greater incentives to use good contraception, and within a group homogenous with respect to their desired fertility, those who use good contraception achieve a lower variance of fertility.

Table IV-3 also indicates that there is a tendency for the more educated women to have lower mean fertility and smaller variance of fertility for each contraception category. The differences in the means do not often exhibit statistical significance (see Panel B), but the differences in the variances among users of good contraception are statistically significant, often at  $\alpha = .01$ . The observed relationship between  $\mu_N$  and  $\sigma_N$  across education groups for good and for poor contraception users separately, mirror the observed relationship between  $\mu_N$  and  $\sigma_N$  across good and poor contraception users holding education constant. This observation is quite consistent with more educated couples being more proficient users of each given contraceptive method--the partitioning of the sample of women aged 35-44 by "good" and "poor" (holding education constant) yields the same qualitative differences as the partitioning by more and less education (holding contraception quality constant).

To obtain another measure of the relationship between contraception choice, the wife's education, and fertility outcomes as indicated in Table IV - 3, a multiple regression was run using the twelve education-

contraception method cells. Let  $\bar{N}_j$  and  $\bar{V}_j$  be the mean and standard deviation of live births in cell  $j$ ,  $G_j$  and  $P_j$  be dummy variables reflecting the use of good contraception (compared to poor) and poor contraception (compared to none), and  $E_j$  be the wife's education level (assigned the values 7, 10, 12, and 14 for the respective columns).

The regressions, weighted by the square root of the cell sizes, yielded:

$$\bar{N}_j = 5.23 - 0.18(E_j) + 0.34(G_j) + 0.42(P_j)$$

(8.51) (-3.63) (1.13) (1.43)<sup>j</sup>

and

$$\bar{V}_j = 3.38 - 0.15(E_j) - 0.34(G_j) - 0.31(P_j) \quad .37$$

(15.45) (-8.27) (-3.25) (-3.03)<sup>j</sup>

The wife's education has a significant negative effect on both the mean and the standard deviation of the number of live births. Contraception use had no significant effect on the mean number of live births, but users of poor contraception had a significantly lower standard deviation in live births than users of no contraception, and users of good contraception had a significantly lower standard deviation in live births than users of poor contraception.

Table IV-4 partitions this set of households, the white non-Catholics, by duration of marriage and the expected number of children, for women married only once and aged 35 and above. Since age at marriage differs systematically by several socio-economic characteristics,

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The standard deviations of  $N_j$  and  $V_j$  were 0.559 and 0.517 and the standard errors of the estimates were 0.359 and 0.127 respectively.

Table IV-4  
 Mean Number of Live Births  $\mu_N$  and Variance in Number of Live Births  $\sigma_N^2$  for White Non-Catholic Women Aged 35 or above and Married Once; by Marriage Duration, Expected Number of Children and Contraceptive Method (Good, Poor, None) Used in the First Pregnancy Interval

	Estimated Number of Births*									
	2		3		4		Total		Total	
	Marriage Duration 15-19 Years	Total	Marriage Duration 20-24 Years	Total	Marriage Duration 25 Years	Total	Marriage Duration 25 Years	Total	Marriage Duration 25 Years	Total
Good $\mu_N$	2.815	2.907	2.724	2.803	3.455	2.809	2.277	2.842	4.000	2.746
$\sigma_N^2$	1.162	1.573	1.349	1.661	4.873	1.745	1.596	1.528	4.286	2.138
(n)	(108)	(204)	(76)	(61)	(11)	(157)	(47)	(57)	(8)	(114)
Poor $\mu_N$	3.300	3.301	2.952	3.000	3.200	2.946	2.864	3.333	4.636	3.383
$\sigma_N^2$	2.537	2.713	1.948	2.455	3.700	2.326	1.933	3.121	6.655	3.464
(n)	(20)	(73)	(21)	(45)	(5)	(74)	(22)	(48)	(11)	(81)
None $\mu_N$	2.286	2.640	2.225	3.109	4.333	3.012	1.836	3.106	4.182	3.184
$\sigma_N^2$	2.790	4.006	3.512	3.087	5.884	4.182	2.991	5.860	7.466	7.001
(n)	(56)	(178)	(40)	(92)	(24)	(166)	(55)	(94)	(33)	(201)
Total $\mu_N$	2.707	2.866	2.613	2.990	3.950	2.919	2.185	3.085	4.250	3.098
$\sigma_N^2$	1.881	2.746	2.107	2.497	5.331	2.867	2.380	3.957	6.623	4.914
(n)	(184)	(455)	(137)	(198)	(40)	(397)	(124)	(199)	(52)	(396)

\* Estimated from a fertility demand equation (see text footnote, p.54). The cells are defined on intervals: 2 is 1.5 to 2.5; 3 is 2.5 to 3.5, 4 is 3.5 to 4.5. Total includes observations with estimated number of births <1.5 and >4.5 as well.

the partitioning by marriage duration should more adequately standardize for the length of the period of time at risk of conception. To standardize further for the incentives to avert births, we have used a definition of expected number of children derived by regressing the actual number of live births on a set of economic and demographic characteristics.<sup>38</sup> We used this fertility demand function to estimate  $\hat{N}$  for each household, then grouped households into cells defined by intervals of  $\hat{N}$ .<sup>39</sup>

One observes in Table IV-4 that when the marriage duration and the household's expected fertility  $\hat{N}$  are held constant, users of good contraception had smaller variation in their actual fertility than did users of poor or no contraception. Also, standardized for marriage duration and contraception choice, households characterized by  $\hat{N}$  equal to three tended to experience a larger variation in actual fertility than households characterized by  $\hat{N}$  equal to two. That is, there appears

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<sup>38</sup>

The regression was estimated from the 1965 NFS for white non-Catholic women aged 35 or above. The estimation yielded:  $N = 10.09616 - 0.35473 (MARI) - 0.00042 (I) - 0.34462 (ED) + 0.00003 (EDXI) + 0.30078 (URB) + 0.09881 (MARD) - 0.12119 (AGE)$  where MARI equals 1 if married more than once or equals zero otherwise; I is an estimate of the husband's income at age 40 based on an estimated earnings function using husband's education and market experience; ED is the wife's education level; EDXI is a multiplicative interaction term using ED and I; URB equals 1 if the household lives in a rural area or equals zero otherwise; MARD is the duration of the current marriage in years; and AGE is the wife's age.

<sup>39</sup>

This procedure partitions the group of households into cells on the basis of the economic and demographic characteristics which, on average, are associated with one, two, three or four children. For our purposes this procedure suffices, but it does not resolve the problem of partitioning the household's actual fertility into the "desired" and the "unwanted" components.

to be a positive association between mean fertility and the standard deviation of fertility. Furthermore, as figure IV-3 indicates, the positive relationship between  $\hat{N}$  and  $\sigma_N$  appears to be strongest among users of no contraception.<sup>40</sup>

Since more educated women tend to marry at later ages, the total length of time at risk of conception probably differs by education for women of a given age. Table IV-3 partitions the sample by wife's age and education; for comparison Table IV-5 partitions the sample by wife's education and marriage duration for women married once and aged 35 and above. Although the cell sizes in Table IV-5 are smaller and the results somewhat more erratic, one again observes a tendency for users of good contraception to have somewhat lower mean fertility and, more systematically, lower variation in their fertility.

The demographic literature has emphasized a distinction between wanted and "unwanted" fertility, and it is tempting to try to partition actual fertility into these two components for separate analysis.<sup>41</sup>

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The figure plots only cells based on 20 or more observations. Unfortunately the cell sizes for the users of poor contraception are quite small, making generalizations difficult.

<sup>41</sup>In addition to Ryder and Westoff(1971), see Bumpass and Westoff(1970), Ryder (1973) and Part IV "Unwanted Fertility" of Volume 1 of the Commission on Population Growth (1972), particularly the essay by Ryder and Westoff (1972).



Figure IV-3  
 Variance in Live Births  
 by Expected Number of Births  
 by Contraception Method  
 and Marriage Duration

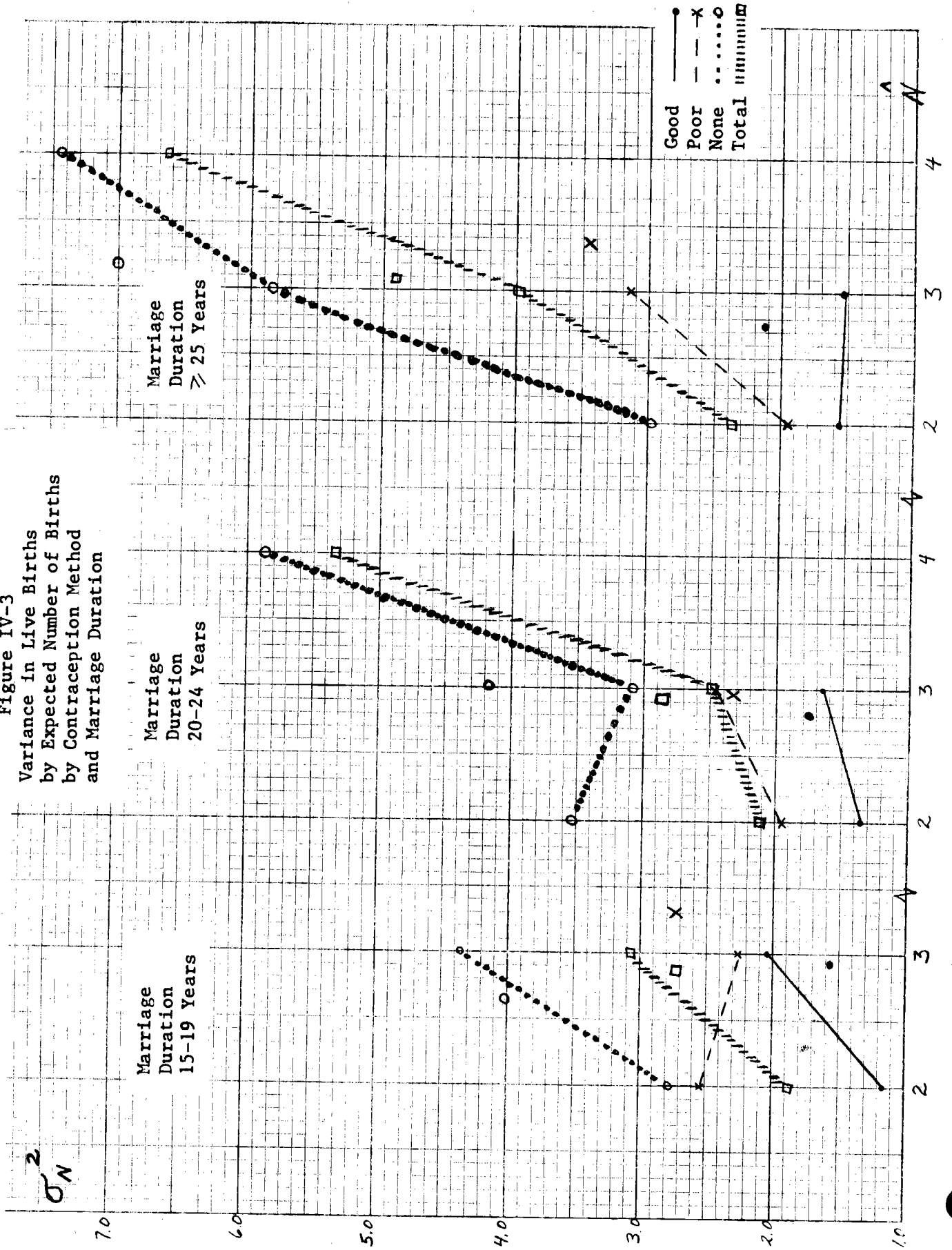


Table IV-5  
 Mean and Standard Deviation of Live Births and Unwanted Births  
 for White Non-Catholic Women Aged 35 or above and Married Once,  
 by Marriage Duration, Wife's Education, and Contraceptive Method  
 (Good, Poor, None) Used in the First Pregnancy Interval

(A)

Marriage Duration 15-19 Years					
Contraceptive Method	Education of Wife				Total
	≤ 8	9-11	12	≥ 13	
<u>Live Births</u>					
Good $\mu_N$	3.167	3.192	2.766	3.000	2.907
	$\sigma_N$ 1.722	1.833	1.202	0.984	1.254
	(n)** (6)	(26)	(107)	(65)	(204)
Poor $\mu_N$	4.182	2.857	3.098	3.429	3.301
	$\sigma_N$ 1.991	2.410	1.428	1.453	1.647
	(n) (11)	(7)	(41)	(14)	(73)
None $\mu_N$	2.708	3.146	2.600	2.061	2.640
	$\sigma_N$ 2.458	2.056	1.946	1.580	2.001
	(n) (24)	(41)	(80)	(33)	(178)
Total $\mu_N$	3.171	3.135	2.768	2.777	2.866
	$\sigma_N$ 2.290	1.988	1.543	1.327	1.657
	(n) (41)	(74)	(228)	(112)	(455)
<u>Unwanted Births*</u>					
Good $\mu_U$	0.167	0.962	0.523	0.431	0.539
	$\sigma_U$ 0.408	1.587	1.049	0.770	1.052
	(n) (6)	(26)	(107)	(65)	(204)
Poor $\mu_U$	2.091	1.714	0.512	0.714	0.904
	$\sigma_U$ 1.921	2.563	0.952	1.139	1.474
	(n) (11)	(7)	(41)	(14)	(73)
None $\mu_U$	0.667	1.268	0.837	0.455	0.843
	$\sigma_U$ 1.341	2.013	1.354	0.905	1.480
	(n) (24)	(41)	(80)	(33)	(178)
Total $\mu_U$	0.976	1.203	0.632	0.473	0.716
	$\sigma_U$ 1.573	1.916	1.155	0.859	1.311
	(n) (41)	(74)	(228)	(112)	(455)

\* For definition see text

\*\* Cell size

Table IV-5  
(continued)  
B

		Marriage Duration 20-24 Years				
Contraceptive Method		Education of Wife				Total
		≤8	9-11	12	≥13	
<u>Live Births</u>						
Good	$\mu_N$	4.000	2.571	2.784	2.610	2.809
	$\sigma_N$	1.961	1.513	1.219	1.115	1.321
	(n)	(14)	(28)	(74)	(41)	(157)
Poor	$\mu_N$	2.000	3.842	2.686	2.917	2.946
	$\sigma_N$	0.926	1.642	1.430	1.379	1.525
	(n)	(8)	(19)	(35)	(12)	(74)
None	$\mu_N$	4.114	3.366	2.701	1.609	3.012
	$\sigma_N$	2.285	2.009	1.715	1.616	2.045
	(n)	(35)	(41)	(67)	(23)	(166)
Total	$\mu_N$	3.789	3.216	2.733	2.355	2.919
	$\sigma_N$	2.169	1.765	1.459	1.402	1.693
	(n)	(57)	(88)	(176)	(76)	(397)
<u>Unwanted Births</u>						
Good	$\mu_U$	1.571	0.571	0.311	0.171	0.433
	$\sigma_U$	1.910	0.920	0.720	0.442	0.942
	(n)	(14)	(28)	(74)	(41)	(157)
Poor	$\mu_U$	0.0	1.211	0.400	0.167	0.527
	$\sigma_U$	0.0	1.273	0.604	0.389	0.879
	(n)	(8)	(19)	(35)	(12)	(74)
None	$\mu_U$	1.971	1.561	0.881	0.391	1.211
	$\sigma_U$	2.419	1.988	1.332	1.076	1.822
	(n)	(35)	(41)	(67)	(23)	(166)
Total	$\mu_U$	1.596	1.170	0.545	0.237	0.776
	$\sigma_U$	2.203	1.613	1.013	0.690	1.419
	(n)	(57)	(88)	(176)	(76)	(397)

Table IV-5  
(concluded)

C

Marriage Duration  $\geq$  25 Years

Contraceptive Method	Education of Wife				Total
	$\leq 8$	9-11	12	$\geq 13$	
<u>Live Births</u>					
Good $\mu_{NN}$	3.688	2.739	2.471	2.708	2.746
$\sigma_{NN}$	1.957	1.789	1.065	1.301	1.462
(n)	(16)	(23)	(51)	(24)	(114)
Poor $\mu_{NN}$	4.105	3.400	3.152	2.667	3.383
$\sigma_{NN}$	2.105	2.113	1.523	1.658	1.861
(n)	(19)	(20)	(33)	(9)	(81)
None $\mu_{NN}$	4.240	3.580	1.932	1.706	3.184
$\sigma_{NN}$	3.004	2.548	1.680	1.359	2.646
(n)	(75)	(50)	(59)	(17)	(201)
Total $\mu_{NN}$	4.136	3.333	2.406	2.360	3.098
$\sigma_{NN}$	2.724	2.295	1.516	1.439	2.217
(n)	(110)	(93)	(143)	(50)	(396)
<u>Unwanted Births</u>					
Good $\mu_{UU}$	1.063	0.435	0.392	0.208	0.456
$\sigma_{UU}$	1.843	0.896	0.666	0.509	0.961
(n)	(16)	(23)	(51)	(24)	(114)
Poor $\mu_{UU}$	1.000	0.700	0.879	0.111	0.778
$\sigma_{UU}$	1.291	0.979	1.453	0.333	1.235
(n)	(19)	(20)	(33)	(9)	(81)
None $\mu_{UU}$	1.733	1.400	0.644	0.353	1.214
$\sigma_{UU}$	2.658	2.356	1.283	0.702	2.182
(n)	(75)	(50)	(59)	(17)	(201)
Total $\mu_{UU}$	1.509	1.011	0.608	0.240	0.907
$\sigma_{UU}$	2.376	1.879	1.157	0.555	1.759
(n)	(110)	(93)	(143)	(50)	(396)

The difficulty, of course, is in obtaining an estimate of wanted births distinct from the household's actual fertility. The problem is not simply one of estimation. Viewed in the context of the stochastic model described above, it is not possible even in principle to designate each pregnancy as "desired" or undesired. Also, we have emphasized that the household's optimal number of children is affected by fertility control costs (i.e.  $\beta_1 k$  in equation 16) and the relation between mean and variance of fertility (i.e.  $-K/2$ ). So the definition of "desired" fertility depends critically upon what is assumed about fertility control costs, the variance in actual fertility, etc. <sup>42</sup>

Recognizing these limitations, we nevertheless attempted to consider "unwanted" fertility since the NFS data set contains retrospective information on the couple's fertility goals prior to each pregnancy. By summing

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<sup>42</sup> That is, in the terminology of equation 16 one might define unwanted fertility as  $\mu_N - N^*$  or  $\mu_N - \mu_N^*$ . However, since  $\sigma_N > 0$ , actual fertility  $N$  differs in general from  $\mu_N$ , so "unwanted" fertility presumably includes not only the discrepancy between some fixed target fertility and optimal mean fertility  $\mu_N$  but also the variation in actual fertility around  $\mu_N$ .

up the number of pregnancies "wanted"<sup>43</sup> and subtracting this number from the total number of live births, we obtained an estimate of the number of births "unwanted" in each household. Panel B of Table IV-5 indicates the mean number of "unwanted" births by marriage duration, wife's education, and contraceptive method used in the first pregnancy interval. There appears to be a relatively strong negative relation

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The NFS data contain retrospective information about the contraceptive behavior and the husband's and wife's attitudes about another pregnancy prior to each of the wife's pregnancies. Following Ryder and Westoff (1971) we considered each live birth as "wanted" or "unwanted" on the basis of the behavioral and attitudinal circumstances prior to that pregnancy. The birth was considered "wanted" if any of the following three conditions was met: (1) The birth was "wanted" if the couple had used no contraception in the interval prior to that pregnancy and responded "yes" to the question "Was the only reason you did not use any method then because you wanted to have a baby as soon as possible?" (2) The birth was "wanted" if the couple had used a contraceptive method in the interval prior to that pregnancy and had "stopped using a method in order to have a child." (3) If the couple was not using contraception but responded "no" to the question quoted above or if the couple had conceived while using a method or while having stopped using a method but "did not want to become pregnant at that time," then the couple was asked two additional questions. If the response was "yes" to either of these additional questions the birth was considered "wanted." The two questions were "Before you became pregnant this time did you want to have a (another) child sometime?" and "Did your husband want to have a (another) child sometime?" Our definition of "wanted" differs slightly from the definition used by Ryder and Westoff.

between the number of "unwanted" births and the wife's education level,<sup>44</sup> and a somewhat systematic relationship between the number of "unwanted" births and the use of good contraception.

We want to stress that the results pertaining to unwanted births are subject to many qualifications and are included here primarily as some evidence that this one measure of the intuitively appealing notion of an unwanted pregnancy seems to be related to contraception use as one would expect. The arbitrariness of the precise definition of an unwanted pregnancy helps convince us that it is more useful in studying the uncertainty related to fertility behavior to focus on the distribution of actual births than to concentrate on partitioning observed fertility into "desired" and "unwanted" fertility. We have shown in this section that across broadly defined groups of households there appears to be a systematic relationship between contraception strategies and the mean and variance of observed fertility. Couples characterized by the use of more effective contraception appear to have somewhat lower mean fertility, lower variance of fertility, a weaker relationship between their mean and variance of fertility and perhaps a lower level of "unwanted" fertility.

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<sup>44</sup> This finding is consistent with Ryder and Westoff's conclusion that "there is a strong negative association of education and unwanted fertility" (See Ryder and Westoff 1972, p.483). Their conclusion is also based on the 1965 NFS data and its sequel, the 1970 NFS, which is not yet available to us. It must be stressed however that Ryder and Westoff's definition of "unwanted" pregnancies and their criteria for selection of the subsample studied differ from ours and one should not make inferences about fertility behavior from comparisons between their tables and ours.

## V Diffusion of the Pill

The 1965 National Fertility Study and its sequel, the 1970 National Fertility Study, provide a unique opportunity to follow, at the household level, the diffusion of a major technical innovation--the oral contraceptive--from its introduction for sale in the U.S. in 1960. In addition to the intrinsic interest of studying the diffusion of new technology, the observed pattern of adoption of the pill provides an important test of hypotheses derived from our theory of contraceptive choice (see Section II).

There are two main reasons why the study of pill adoption provides a more powerful test of our model of contraceptive choice than would be afforded by studying differential choices among techniques existing before 1960. First, the introduction of the pill was an exogenous event from the standpoint of potential adopters. Second, the pill is a truly new kind of contraception in comparison with alternative methods available prior to 1960--it is significantly more effective and less coitus-related than alternative methods. This second consideration suggests that the pill has a significantly lower marginal cost of fertility control than other methods, at least in terms of the psychic costs associated with forgone sexual pleasure. Thus couples were confronted, after 1960, with a significantly different set of potential contraceptive methods to choose from, and we investigate in this section the differential rates of adoption of the pill among women with different initial conditions in 1960 (e.g., marital status, age, parity and prior contraceptive practices).

We first present a few hypotheses about the expected pattern of diffusion assuming contraceptive choice is governed by factors considered in our theoretic-



cal model. Next, we utilize data from the 1965 NFS to test these hypotheses and to estimate the probability of pill use as a logistic function of the household's economic characteristics, parity in 1960, and prior use of contraception.

The main behavioral hypothesis underlying our theory of choice of contraceptive technique is that couples choose the least costly technique to achieve a given fertility goal (e.g. a given value of the mean and variance of fertility outcomes). It was shown in Section II-iii that the same technique will not be least costly for all possible goals unless the technique with the smallest marginal cost is also the technique with the lowest fixed cost. This proposition led to the derivation of the contraception cost curve for a "typical" couple in Figure II-3 as the envelope of least costly segments of the curve associated with particular contraceptive techniques. According to this analysis, the more births a couple expects to avert the more likely it is to choose a technique with relatively low marginal cost and high fixed cost.<sup>45</sup>

To derive hypotheses about the adoption of the pill from this theory, we shall assume that for most couples the pill has a lower marginal cost than other contraceptive techniques available in 1960-65. Since it is easy to provide examples of components of marginal cost which are higher for the pill than for alternative techniques (e.g., money cost, side effects on health,

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<sup>45</sup> For the time being, we shall ignore risk considerations and argue in terms of variations in expected fertility within an essentially deterministic framework.

etc.), the plausibility of this assumption depends on the further assumption that the major component of the marginal cost of contraception for most people stems from the conflict between effective use of a method and sexual pleasure.<sup>46</sup> The fixed cost of the pill includes the cost of acquiring information about its existence, characteristics and method of distribution in addition to the money cost of visiting a doctor to obtain a prescription and various psychic costs (e.g. religious principles) which are not related to information acquisition.

We can use figure II-3 (page 31) to illustrate the hypothetical fertility control costs faced by a typical couple. Recall that the line  $OF_2$  represents a zero fixed-cost, high marginal cost technique such as withdrawal, and line  $\alpha_3 F_3$  represents another technique such as condom. Suppose that the high fixed-cost, low marginal cost technique depicted by  $\alpha_4 F_4$  represents the costs of the pill at the time of its introduction on the market in 1960.<sup>47</sup>

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<sup>46</sup> The IUD, a close rival of the pill in terms of effectiveness and coitus-related costs, was not widely available until after 1965.

<sup>47</sup> While the line  $\alpha F$  in the figure has a slope, or marginal cost  $\beta_4$ , of approximately zero,<sup>4</sup> the cost curve for the pill could have been drawn with a positive although relatively low marginal cost.

Also, for simplicity of exposition, we will assume here that these three techniques, withdrawal, condom, and pill, represent the entire envelope cost curve. In our empirical analysis all available techniques are included.

If figure II-3 depicts the cost conditions faced by a typical couple newly married in 1960, the couple will adopt the pill only if it wishes to avert all of its potential births (i.e. if it wishes to have zero children). If the couple wishes to avert between seven and eleven births (i.e. if it wishes to have one to five children) it will use the condom as its contraceptive method. As the figure is drawn, if the couple wishes to avert fewer than seven births (i.e. if it wishes to have more than five children) withdrawal will be used as its contraceptive method.

The probability  $P$  that a couple will adopt the pill is equal to the probability that the total contraception cost of using the pill, is less than the total cost of using the least costly alternative technique.

This statement may be expressed as:

$$P = \Pr [F_p < \min(F_i)] \quad (17)$$

$$= \Pr \left[ \alpha_p + \beta_p (\mu_N^A - \mu_N) < \min(\alpha_i + \beta_i (\mu_N^A - \mu_N)) \right]$$

where  $F_p = \alpha_p + \beta_p (\mu_N^A - \mu_N)$  is the total cost of the pill and  $F_i = \alpha_i + \beta_i (\mu_N^A - \mu_N)$  is the total cost of the  $i^{\text{th}}$  alternative method ( $i = 1, 2, \dots$ ). If all the  $\alpha$ 's,  $\beta$ 's,  $\mu_N^A$  and  $\mu_N$  were identical constants for all members of the population,  $P$  would be either zero or one for everyone and the pill would be used either universally or not at all.

In fact, of course, these variables will be distributed according to some joint probability distribution across the population so that we should expect to find some fraction,  $0 \leq P \leq 1$ , for whom the pill is least costly. Moreover, our theory suggests that each of these variables is a function of exogenous variables as well as purely random factors. For example,  $\mu_N$  is a function of husband's lifetime income and wife's price of time, two

variables which help determine the couple's demand for children. Likewise,  $u_N^{\wedge}$  is a function of the couple's natural fecundability and the wife's age at marriage, and the  $\alpha$ 's and  $\beta$ 's are functions of variables such as religion and education which determine the fixed and variable costs of each contraceptive technique.

These considerations suggest that we may express the aggregate proportion of households using the pill as

$$P = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} \Pr \left[ (F_p - \min(F_1) | x_1, \dots, x_n) < u \right] h(x_1, \dots, x_n, u) dx_1 \dots dx_n du \quad (18)$$

where  $\Pr(\cdot)$  is the probability of using the pill conditional on the values of the exogenous variables  $x_1, \dots, x_n$  which determine natural fertility, the demand for children and the costs of contraception, and where  $u$  is a random variable and  $h(x_1, \dots, x_n, u)$  is a joint density function of the  $x$ 's and  $u$ .

If we assume that  $u$  is distributed logistically and is independent of the  $x$ 's, the conditional probability of using the pill may be expressed as a logistic function of the form

$$P = \left[ 1 + \exp \left( - (a + b_1 x_1 + \dots + b_n x_n + u) \right) \right]^{-1} \quad (19)$$

or, alternatively, the natural log of the odds of using the pill becomes a linear function of the form

$$\ln\left(\frac{P}{1-P}\right) = a + b_1 x_1 + \dots + b_n x_n + u \quad (20)$$

which can be estimated using standard maximum likelihood procedures.

After discussing several hypotheses about the set of variables  $x_1, x_2, \dots,$

$x_n$ , we estimate the parameters of logistic functions of the form suggested in equation (20).

We shall first consider the pattern of pill use in a static setting in which the distributions of natural fertility,  $\mu_N^A$ , the demand for children,  $\mu_N$ , and the costs of contraception, the  $\alpha$ 's and  $\beta$ 's, are stable in the population. Subsequently we consider the question of diffusion of the pill over time. For convenience, assume for the moment that all couples in the population face identical cost functions for the three techniques depicted in figure II-3, but differ in their natural fertility or in their demand for children. If couples are then distributed by the number of births they wish to avert, figure II-3 implies that the proportion of couples in the population which adopts the pill will be equal to the proportion wishing to avert 11.5 or more births; the proportion which uses the condom will be equal to the proportion wishing to avert between seven and 11.5 births, and the proportion using withdrawal will equal the proportion wishing to avert fewer than seven births. In short, given a distribution of couples by the births they wish to avert, the switching points on the envelope total cost curve determine the proportion of couples using the various available techniques.

We have suggested three sets of factors determining the proportion of couples which might be expected to adopt the pill. We will consider each in turn. Holding constant factors affecting the fertility control total cost curve (the  $\alpha$ 's and  $\beta$ 's) and the demand for children (the  $\mu_N$ ), any factor which increases natural fertility  $\mu_N^A$  will increase expected births averted ( $B = \mu_N^A - \mu_N$ ). Thus for a group of households, an increase in  $\mu_N^A$

will, *ceteris paribus*, increase the proportion of households using the pill (and decrease the proportion using withdrawal).<sup>48</sup> The major observable variable related to  $\mu_N^A$  is the wife's age when the pill became available. So our hypothesis is that, *ceteris paribus*, the older the wife was in 1960 the less likely she is to adopt the pill.<sup>49</sup>

Alternatively, if we hold constant natural fertility,  $\mu_N^A$  and the fertility control total cost curve, factors which increase the demand for children  $\mu_N$  reduce the expected averted births  $B$  and, therefore reduce the probability of adopting the pill. In studies of completed fertility,

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<sup>48</sup> In the simple three-technique case the effect on condom usage is uncertain since some who previously used the condom are expected to switch to the pill while some who previously used withdrawal are expected to switch to the condom. In the more general case of several techniques an increase in  $B$  is expected to increase the proportion using the high fixed cost, low marginal cost technique, *ceteris paribus*.

<sup>49</sup> That is, the older she is the shorter the remaining reproductive time span, so the smaller her remaining natural fertility and thus the smaller her remaining expected averted births. Since she wishes to avert fewer births, *ceteris paribus*, her incentive to adopt the high-fixed-cost pill is relatively slight.

This hypothesis should be qualified in two respects. First, there may be an advantage to postponing the investment in the fixed cost associated with pill adoption until the benefits of the reduced variable costs associated with pill use are close at hand. If these benefits are greatest when the couple wishes to contracept with high efficiency in order to prevent any additional pregnancy, the adoption of the pill might be postponed until desired fertility is reached. Second, age at marriage is likely to be inversely correlated with the couple's demand for children. So the younger the woman is at a given parity the more likely she wants to have a large family, which would tend to offset the effect of age on expected births averted, unless desired fertility is explicitly held constant. Thus our inability to hold desired fertility precisely fixed may introduce biases on the other variables.

the husband's lifetime income, H, and the wife's education, W, (as a proxy for her potential lifetime wage or price of time), have been found to be the most important economic variables. Sanderson and Willis (1971) and Willis (1973) argued on theoretical grounds for a positive interaction effect between H and W.<sup>50</sup> Estimating an equation of the form

$$\mu_N = \gamma_0 + \gamma_1 H + \gamma_2 W + \gamma_3 HW \quad (21)$$

in a number of samples of U. S. women, they found that  $\gamma_1$  and  $\gamma_2$  are negative and  $\gamma_3$  is positive, as expected. If we include these three variables in the equation for P, since we hypothesize a negative relationship between  $\mu_N$  and P, we expect to observe positive coefficients on H and W and a negative coefficient for HW.<sup>51</sup>

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Briefly, the argument is this. When the wife is not supplying labor to the market, the shadow price of her time is higher than her potential market wage and is an increasing function of her husband's income. When she supplies labor to the market, her price of time is equal to her market wage and is independent of her husband's income. Since children are assumed to be time-intensive, the positive income effect of H on the demand for children is offset by a substitution effect against children in families with non-working wives, while there is no offset in families with working wives. Since the wife's labor force participation is negatively related to H and positively related to W, Willis (1973) shows that the effects of H and W on number of children will be nonlinear with a positive coefficient on the interaction variable, HW. See, however, Ben Porath (1973) for alternative interpretations of nonlinearity in the demand function for children.

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The husband's lifetime income and the wife's education are not, of course, the only variables relevant for the demand curve for children. The demand curve used in the previous section, for example, included several additional variables (see footnote 38) and also pertained to white non-Catholics only. For the analysis of pill adoption in this section we again restrict ourselves to white non-Catholics.

Estimating the effects on pill use of variation in the fixed and marginal costs of the pill and alternative methods of contraception presents a number of difficult problems. These costs are likely to be dominated by psychic or nonmarket components which may vary widely across households but cannot be measured directly. Certain variables such as Catholicism are known to influence the costs of certain forms of contraception and it is frequently argued that education reduces the cost of acquiring birth control information (see Michael 1973). Unfortunately, both of these variables also help determine the demand for children and it is not easy to see how this influence can be disentangled from their influence on the cost of contraception.





One advantage of studying the adoption of the pill is the fact that we may study the response to the introduction of the pill by women whose initial conditions differed in 1960. This procedure enables us to study the effect of prior use of other forms of contraception before 1960 on the probability of adopting the pill after 1960. The theory suggests that couples which have incurred the fixed costs associated with some other technique will, *ceteris paribus*, be less likely to adopt the pill. Hence, we expect prior use of the diaphragm and condom, for example, to be negatively associated with pill adoption.<sup>52</sup> A second aspect of prior use which may influence the probability of adopting the pill is the success the couple has with the previous method. If the couple's previous method has had high marginal costs, our theory suggests that it would be used relatively inefficiently. The couple would have, therefore, a higher risk of an "accidental" pregnancy while using that method. Since the pill has a relatively low marginal cost, the higher the marginal cost of the alternative method, the more likely the pill will be adopted.

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<sup>52</sup> That is, after the fixed costs are borne, the fertility control costs fall from  $\alpha_i + \beta_i B$  to simply  $\beta_i B$ . There are difficulties in this test of the sunk-cost hypothesis, however. A woman's prior contraception history is not independent of residual variance caused by variation in the couple's demand for children, fecundability, or costs of contraception which we are unable to hold constant with the other variables in the model. Several potential biases tend to work against the "sunk cost" hypothesis. For example, subfecundity or sterility, which may be one of the reasons that a woman has not contracepted in the past, would also tend to reduce her probability of adopting the pill. If this bias dominated, we might find that prior users of contraception have a higher rather than a lower probability of using the pill compared with prior nonusers.

Consequently, we expect prior "contraceptive failure" to be positively related to the probability of adopting the pill. This expectation is strengthened by the likelihood that couples which have experienced contraceptive failure confront the prospect of larger losses of expected utility from additional "unwanted" births than do couples who have successfully contracepted in the past.

While we have discussed three separate sets of factors influencing the probability of pill adoption in a static framework, we have not as yet considered the diffusion of the pill over time. A major driving force in any process of diffusion of new technology is the reduction over time of the cost of acquiring information about the new innovation. By the simple act of adopting and using the new technique, early adopters convey information to later adopters about the existence of the technique, how it is distributed, etc. Since the pill is a prescription drug, its adoption is affected by the diffusion of information among doctors as well as among potential adopters. The dynamics surrounding the cost of information about the pill and the speed and pattern of its diffusion will also be related to socio-economic differentials in rates of adoption since this information is spread by word of mouth.

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Although they are not information costs, a similar mechanism may operate to reduce the costs associated with deviation from group norms as individuals in the group witness increased nonconformity with these norms. In the case of the pill, the interaction between the teachings of the Catholic Church and the behavior of individual Catholics might be interpreted along these lines (see Ryder and Westoff, 1971, Chapter 8).

To consider the effect of decreasing information costs of pill adoption, we again make use of figure II-3 (page 31). With the passage of time the fixed cost (which includes the information cost) of the pill for the average household may fall from  $0\alpha_4$  to  $0\alpha'_4$ . Thus the switching point--the number of averted births at which pill adoption is warranted--falls from about 11.5 births averted to about 6 births averted, as the figure is drawn. Obviously, in the aggregate, the reduction in the fixed cost increases the proportion of couples using the pill. Notice that the model implies that the users of the next-best technique will be those who most readily adopt the new technique as its costs of information fall over time. The new low-marginal-cost technique first displaces the existing technique with the lowest marginal cost.<sup>54</sup>

To test the hypotheses advanced above, we have selected three samples of white non-Catholic women from the 1965 National Fertility

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Hence, should a new contraceptive technique be introduced which further lowers the marginal cost of contraception, our theory implies that the pill could be the first technique to be displaced. Ryder (1972) shows that by 1970 pill use differentials by education, race and religion had converged. Much of the convergence, however, was caused by an absolute decline in the use of the pill after 1967 by highly educated white non-Catholics, the group which has the highest rate of pill use. It is interesting to speculate whether new techniques such as the IUD and the increased popularity of the vasectomy and tubal ligation had begun to displace the pill in this group. Ryder emphasizes the effects of fears about long-term adverse health effects of the pill, but these alleged effects were not widely publicized until the U. S. Senate Hearings in 1969, well after the 1967 peak in pill usage in the high-use group.

Study who began their first, second or third pregnancy interval in the period 1960-1964 (see Table V-1 for a description of these samples).

In each sample we estimate the probability that a woman uses the pill in the specified interval as a function of three sets of variables that determine, respectively, (1) the woman's potential (i.e. natural) fertility from the beginning of the interval until menopause; (2) the couple's demand for children; and (3) the couple's costs of contraception.

These variables, which are listed in Table V-1, may be grouped as follows:

<u>Variable</u>	<u>(Expected Effect on P)</u>
I. Potential Fertility	
Wife's age in 1965	(-)
II. Demand for Children	
Wife's education	(+)
Husband's income at age 40	(+)
Income-education interaction	(-)
III. Cost of Contraception	
(A) Date interval began	(+)
(B) Used diaphragm in previous interval	(-)
Used condom in previous interval	(-)
Used pill in previous interval	(+)
Used other method in previous interval	(-)
(C) Contraceptive failure in previous interval	(+)
IV. Age at Marriage	(+)

The sign accompanying each variable indicates the hypothesized direction of effect of that variable on the probability of using the pill in a particular pregnancy interval. These hypotheses stem from the discussion on the preceding few pages, and most seem to require no further discussion. Note that the "date the interval began" operates as a time trend in

Table V-1  
Means and Variances of Variables Used in Logistic Estimates of Probability of Pill Use\*

Variable	First Interval (N = 495)		Second Interval (N = 502)		Third Interval (N = 472)	
	Mean	Variance	Mean	Variance	Mean	Variance
% Contracepted in Interval	.471	.249	.775	.174	.828	.142
% Used Pill in Interval	.083	.076	.183	.148	.242	.183
% Used Diaphragm in Previous Interval	----	----	.080	.073	.153	.129
% Used Condom in Previous Interval	----	----	.267	.196	.324	.219
% Used Pill in Previous Interval	----	----	.046	.044	.028	.027
% Used Other Method in Previous Interval	----	----	.173	.143	.269	.197
% Had Contraceptive Failure in Previous Interval	----	----	.129	.112	.229	.176
Date of Beginning of Interval						
(Century Months)	748.2	297.2	750.1	272.5	750.8	320.8
Wife's Age at Marriage	20.4	15.5	20.1	9.8	19.9	9.6
Wife's Age in 1965	23.6	17.3	25.1	19.4	27.0	23.6
Education of Wife	12.1	4.1	12.1	4.3	11.9	4.2
Husband's Predicted Income at Age 40 (\$1000)	8.58	2.2	8.56	2.6	8.44	2.7
Interaction (Husband's Predicted Income x Wife's Education(\$1000))	105.9	1041	106.2	1196	102.7	1186

\* The samples consist of white Non-Catholic women married once whose pregnancy interval began in 1960-1964. The beginning of the interval is defined as the date of marriage for the first interval, the date of birth of the first child for the second interval and the date of birth of the second child for the third interval. These dates are stated in "century months." To find the year in which the interval began divide by 12 and add 1900 (e.g. 750 century months equals 1962.5). <sup>the</sup>

this analysis, so it is assumed to be negatively related to the information cost of pill adoption. The prior use of other specific contraceptive techniques is compared with prior non-use of contraception, hence the fixed costs associated with each technique are expected to deter adoption of the pill.<sup>55</sup>

We have estimated the probability of adopting the pill,  $P$ , as a logistic function of the form in equation (20) by a maximum likelihood method.<sup>56</sup> The results are indicated in Tables V-2 and V-3. Two different versions of the pill adoption model are investigated. First, we considered the choice of pill versus all other techniques including no contraception. These results are labeled as pertaining to the "total sample". Second, we considered the conditional choice of pill versus all other techniques, given that some contraceptive technique was used. These results pertain to the sample of "contraceptors". This latter dichotomy is the appropriate one if pill adoption is characterized by the two-stage decision: (1) contracept or not

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<sup>55</sup> One might in fact offer hypotheses about the relative magnitudes of these negative effects based on the assumed ranking of the fixed cost components (the  $\alpha_i$ ) of each. But there may be persistent or serially correlated error terms across intervals, so we have refrained from emphasizing this hypothesis. For example, couples which chose the condom in the previous interval presumably did so for reasons only some of which we have accounted for. Also, each technique has its own set of characteristics which may be related to pill adoption (eg., the diaphragm is a prescription method, so some of its fixed costs which are related to a medical examination may in fact lower the fixed cost of pill adoption).

<sup>56</sup> The computer program was written by Kenneth Maurer of the Rand Corporation. The advantage of the maximum likelihood estimation procedure is that the data need not be grouped. Thus the effects of a relatively large number of independent variables may be estimated from relatively small samples.

Table V-2

Estimates of Probability of Pill Use by White Non-Catholic Women in Pregnancy  
 Intervals Beginning in 1960-1964\*  
 (asymptotic t-ratio in parentheses)

(total sample)

	First Interval		Second Interval		Third Interval	
	(1)	(2)	(1)	(2)	(1)	(2)
	$(\bar{p} = .08)$		$(\bar{p} = .15)$		$(\bar{p} = .24)$	
Age at Marriage	-0.003 (-0.77)	-0.005 (-0.99)	-0.004 (-0.29)	-0.004 (-0.29)	-0.028 (-2.30)	-0.028 (-2.29)
Age of Wife	- -	- -	-0.023 (-2.05)	-0.023 (-2.04)	0.001 (0.13)	0.001 (0.21)
Education of Wife	0.008 (1.01)	0.090 (1.64)	0.019 (1.36)	0.022 (0.36)	0.015 (0.98)	0.084 (2.17)
Husband's Predicted Income (\$1000)	0.03 (2.30)	0.14 (1.85)	0.06 (3.05)	0.06 (0.76)	0.03 (1.71)	0.19 (2.32)
Income-Education Interaction (\$1000)	- -	-0.009 (-1.52)	- -	-0.000 (-0.05)	- -	-0.013 (-2.01)
Date Interval Began	0.002 (3.22)	0.002 (2.86)	0.009 (5.63)	0.009 (5.62)	0.007 (4.72)	0.007 (4.77)

\* The coefficients reported above the t-ratios are analogous to coefficients that would be estimated in a linear probability model of the form  $p = a + b_1x_1 + \dots + b_nx_n$ . These are obtained from the  $\beta$ 's estimated by maximum likelihood in the logistic function  $p = 1/[1 + e^{\alpha + \beta_1x_1 + \dots + \beta_nx_n}]$  by the relation  $b_j = \frac{\partial p}{\partial x_j} = \bar{p}(1-\bar{p})\beta_j$  where  $\bar{p}$  is the mean proportion using the pill. The asymptotic t-ratios, of course, pertain to the  $\beta$ 's.

Table V-3

Estimates of Probability of Pill Use by White Non-Catholic Women in Pregnancy Intervals Beginning in 1960-1964 for Total Sample and for Subsample of Women who Contracepted During the Interval\* (asymptotic t-ratios in parentheses)

	Second Interval		Third Interval	
	(1) **	(2)	(3)	(4)
Age at Marriage	-0.014 (-0.95)	-0.017 (-0.86)	-0.024 (-1.92)	-0.029 (-2.00)
Age of Wife	-0.015 (-1.27)	-0.014 (-0.89)	.004 (0.55)	.004 (0.49)
Education of Wife	.017 (1.15)	.027 (1.29)	.137 (2.07)	.154 (2.07)
Predicted Income of Husband (\$1000)	.050 (2.6)	.041 (1.66)	.190 (2.23)	.206 (2.14)
Income-Education Interaction (\$1000)	----	----	-.013 (-1.94)	-.015 (-1.90)
Date Interval Began	.009 (5.33)	.011 (5.19)	.006 (4.51)	.007 (4.25)
Used Diaphragm in Previous Interval	-.024 (-0.29)	-.077 (-0.75)	.076 (0.94)	-.117 (-1.17)
Used Condom in Previous Interval	-.089 (-1.56)	-.195 (-2.71)	.118 (1.74)	-.082 (-0.96)
Used Pill in Previous Interval	.286 (3.24)	.419 (2.80)	.370 (2.84)	.228 (1.41)
Used Other Method in Previous Interval	-.119 (1.85)	-.223 (-2.70)	.102 (1.39)	-.095 (-1.04)
Contraceptive Failure in Previous Interval	.142 (2.27)	.198 (2.54)	.138 (2.69)	.153 (2.62)
p	.15	.23	.24	.29

\* See note at bottom of Table V-2

\*\* Minimization technique did not converge after 11 iterations



contracept and (2) select a contraceptive technique.

Table V-2 indicates the estimates on the total sample for each of the first three pregnancy intervals, excluding the variables which indicate prior contraception use.<sup>57</sup> The time trend (the date interval began) is positive and statistically significant. This conforms with our hypothesis regarding the effects of a decline in information costs over time. The age of the wife has the expected negative sign in only the second interval, while the age at marriage has an unexpected negative effect on the probability of pill adoption.<sup>58</sup>

The effects of the variables related to the demand for children were computed both with and without the income-education interaction term. In all cases the variables exhibited the expected signs. The relatively stronger effect of husband's income than wife's education, however, is quite surprising. In several studies estimating the effects of these variables on fertility demand, the wife's education has the

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<sup>57</sup>The estimates were computed on both the total sample (for the unconditional P) and the sample of contraceptors (for the conditional P). The results were quite similar in the two cases so only the former are reported here.

<sup>58</sup>We have no explanation for the consistently negative effect of age at marriage. Holding wife's age and current parity constant, a higher age at marriage implies a shorter duration of time from marriage to current parity. Thus we think age at marriage in these estimates may be positively related to relatively high fecundity, relatively high rates of coital frequency, or relatively low demand for children, but each of these factors implies a positive effect of age at marriage on the probability of pill adoption, *ceteris paribus*.

stronger effect. The introduction of the interaction effect does strengthen the effect of the wife's education in the first and third intervals.<sup>59</sup> In general, the signs and magnitudes of the effects of these three variables tend to support the hypothesis that these variables affect fertility demand and have an effect of opposite sign on the probability of using the pill, given the wife's age, parity and the time sequence of the pregnancy interval.

Turning to Table V-3, the effect of prior use of contraception is added to the estimating equations. The table includes the results for the total sample and for the subsample of contraceptors. The effects of the fertility demand variables, the information cost variable, and the age of the wife and age at marriage variables are only slightly affected by the introduction of the set of prior-use variables, so they will not be discussed again here. The effect of failure in the preceding interval is positive as hypothesized. The effects of prior use of the diaphragm, condom or other contraception are negative as hypothesized in the subsample of contraceptors, but are seldom statistically significant; the signs are not as hypothesized in the third interval for the total sample. The expected positive effect of prior pill use is quite strong in most cases.

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<sup>59</sup>The effect of the wife's education,  $W$ , on the probability of pill adoption in, say, the third interval, is  $\delta P / \delta W = .084 - .013H$  which is positive at lower values of husband's income and negative at high values of husband's income.

While it is tempting to discuss in detail several of these estimated coefficients, we will not do so here. We think the qualitative results of our study of pill use offer rather strong support for the hypotheses we developed earlier in this section. In addition, the model can help us interpret the observed trend and differential use of the pill since 1960. In the preceding section we showed that the implications about the relationship between distributions of fertility outcomes and contraception behavior are also supported by the observed behavior from the 1965 National Fertility Survey.

Appendix Table 1  
 Frequency Distribution of Number of Live Births  
 by Wife's Age, Wife's Education, for White Non-Catholic Women

		%											
Wife's Age and Education	Cell Size	Number of Live Births											
		0	1	2	3	4	5	6	7	8	9	10	11
<u>Age 35-39</u>													
≤ 8	66	10.6	9.1	16.7	18.2	16.7	6.1	7.6	4.5	6.1	1.5	0	3.0
9-11	104	2.9	8.7	25.0	21.2	17.3	13.5	6.7	2.9	1.9	0	0	0
12	249	8.4	11.2	28.9	25.7	14.1	4.8	4.0	2.4	0.4	0	0	0
≥ 13	112	8.0	10.7	28.6	22.3	18.8	8.9	0.9	0.9	0.9	0	0	0
Total	531	7.5	10.4	26.6	23.2	16.0	7.5	4.3	2.4	1.5	0.2	0	0.4
<u>Age 40-44</u>													
≤ 8	89	2.2	16.9	25.8	18.0	5.6	9.0	6.7	3.4	6.7	1.1	3.4	0
9-11	128	5.5	13.3	21.1	27.3	12.5	4.7	3.1	7.8	3.1	1.6	0	0
12	254	5.1	11.8	31.9	25.6	11.8	7.1	5.1	1.6	0	0	0	0
≥ 13	118	11.9	8.5	31.4	28.8	13.6	4.2	0.8	0.8	0	0	0	0
Total	589	6.1	12.2	28.5	25.5	11.4	6.3	4.1	3.1	1.7	0.5	0.5	0
<u>Age 45-49</u>													
≤ 8	60	8.3	13.3	10.0	15.0	21.7	8.3	11.7	3.3	5.0	0	1.7	1.7
9-11	53	9.4	22.6	28.3	20.8	7.5	5.7	3.8	1.9	0	0	0	0
12	126	7.9	19.8	34.9	16.7	13.5	2.4	1.6	1.6	0.8	0.8	0	0
≥ 13	61	8.2	11.5	29.5	31.1	18.0	0	1.6	0	0	0	0	0
Total	300	8.3	17.3	27.7	20.0	15.0	3.7	4.0	1.7	1.3	0.3	0.3	0.3



Appendix Table 2 a

Frequency Distribution of Best Contraceptive Method Ever Used  
by Education for White Non-Catholic Women

Method	Education of Wife				Total
	< 8	9-11	12	> 13	
<u>Wife Aged 35-44</u>					
Pill	3.2	12.1	10.9	13.0	10.5
IUD	0.6	0.4	0.8	1.7	0.9
Condom	31.0	35.8	37.4	37.4	36.2
Diaphragm	2.6	14.6	18.3	27.8	17.3
Withdrawal	8.4	4.7	4.4	1.7	4.5
Jelly	3.2	0.9	1.8	2.2	1.9
Foam	0.0	0.9	0.4	0.9	0.5
Suppository	1.3	0.4	2.0	0.4	1.2
Rhythm	3.9	3.9	5.6	2.6	4.4
Douche	3.9	7.3	4.0	1.3	4.1
Other	0.0	0.0	0.0	0.0	0.0
None	41.9	19.0	14.5	10.9	18.5
(n)	(155)	(232)	(503)	(230)	(1120)
<u>Wife Aged 45-54</u>					
Pill	0.0	0.0	2.4	6.6	2.2
IUD	0.0	0.0	0.5	0.9	0.4
Condom	24.0	32.3	32.0	35.8	31.0
Diaphragm	5.9	13.1	22.3	24.5	17.3
Withdrawal	9.9	9.1	5.3	5.7	7.1
Jelly	0.8	2.0	4.8	1.9	2.8
Foam	0.0	0.0	0.0	0.0	0.0
Suppository	2.5	2.0	0.5	0.0	1.1
Rhythm	3.3	3.0	3.4	2.8	3.2
Douche	8.3	11.1	1.9	4.7	5.6
Other	0.8	0.0	0.0	0.0	0.2
None	44.6	27.3	26.7	17.0	28.9
(n)	(121)	(99)	(206)	(106)	(532)

Appendix Table 2 b

Frequency Distribution of Best Contraceptive Method Ever Used  
by Education for Non-White, Non-Catholic Women

Method	Education of Wife				Total
	≤ 8	9-11	12	≥ 13	
<u>Wife Aged 35-44</u>					
Pill	2.4	10.9	3.0	9.4	6.1
IUD	0.0	0.0	0.0	0.0	0.0
Condom	18.5	18.5	25.4	21.9	20.2
Diaphragm	0.8	5.9	13.4	15.6	6.4
Withdrawal	4.0	3.4	6.0	3.1	4.1
Jelly	4.0	4.2	7.5	6.3	5.0
Foam	0.0	1.7	1.5	3.1	1.2
Suppository	1.6	3.4	4.5	0.0	2.6
Rhythm	0.8	0.8	1.5	6.3	1.5
Douche	12.1	8.4	9.0	12.5	10.2
Other	0.0	0.0	0.0	0.0	0.0
None	39.9	42.9	28.4	21.9	42.7
(n)	(124)	(119)	(67)	(32)	(342)
<u>Wife Aged 45-54</u>					
Pill	0.0	3.0	0.0	0.0	0.6
IUD	1.0	0.0	0.0	0.0	0.6
Condom	8.2	15.2	10.0	35.3	12.6
Diaphragm	2.1	3.0	15.0	0.0	3.6
Withdrawal	1.0	0.0	0.0	0.0	0.6
Jelly	3.1	3.0	10.0	0.0	3.6
Foam	0.0	3.0	0.0	0.0	0.6
Suppository	0.0	0.0	5.0	0.0	0.6
Rhythm	0.0	3.0	0.0	0.0	0.6
Douche	19.6	21.2	5.0	11.8	17.4
Other	0.0	0.0	0.0	0.0	0.0
None	64.9	48.5	55.0	52.9	59.3
(n)	(97)	(33)	(20)	(17)	(167)

Appendix Table 2c

Frequency Distribution of Best Contraceptive Method Ever Used  
by Education for White Catholic Women

Method	Education of Wife				Total
	< 8	9-11	12	> 13	
<u>Wife Aged 35-44</u>					
Pill	6.6	10.4	5.5	5.0	6.7
IUD	0.0	0.0	0.5	1.7	0.4
Condom	9.8	25.5	17.4	18.3	18.4
Diaphragm	1.6	5.7	8.3	3.3	6.1
Withdrawal	21.3	11.3	6.9	5.0	9.7
Jelly	3.3	1.9	0.0	3.3	1.3
Foam	0.0	0.9	0.0	1.7	0.4
Suppository	0.0	1.9	1.4	0.0	1.1
Rhythm	13.1	16.0	35.8	38.3	28.3
Douche	6.6	0.9	2.3	1.7	2.5
Other	0.0	0.0	0.0	0.0	0.0
None	37.7	25.5	22.0	21.7	24.9
(n)	(61)	(106)	(218)	(60)	(445)
<u>Wife Aged 45-54</u>					
Pill	0.0	0.0	1.6	0.0	0.6
IUD	0.0	0.0	0.0	0.0	0.0
Condom	14.6	17.9	27.9	26.7	21.5
Diaphragm	6.3	5.1	9.8	0.0	6.7
Withdrawal	6.3	15.4	6.6	13.3	9.2
Jelly	2.1	0.0	0.0	6.7	1.2
Foam	0.0	0.0	0.0	0.0	0.0
Suppository	0.0	0.0	0.0	0.0	0.0
Rhythm	10.4	15.4	19.7	13.3	15.3
Douche	10.4	5.1	3.3	0.0	5.5
Other	0.0	0.0	1.6	0.0	0.6
None	50.0	41.0	29.5	40.0	39.3
(n)	(48)	(39)	(61)	(15)	(163)



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