

## **Where Do Teachers Go?**

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## **Where Do Teachers Go?**

### **Abstract**

Using new and unique administrative data from Georgia, this paper analyzes transitions from full-time elementary and high school teaching. Contrary to public perception, we find very strong evidence that new female teachers are not leaving the teaching profession for high paying jobs in alternative occupations. For our sample of female teachers, only 3.8 percent of elementary teachers and 5.4 percent of high school teachers who left full-time teaching take a non-education sector job in Georgia that paid more than the state minimum teaching wage. This implies that less than one percent of new female teachers leave full-time teaching for a relatively high paying non-education job in Georgia after the first year of teaching. Other groups of teachers, including males, also have low rates of exits to relatively high paying occupations. Given that our results are in direct contrast to public discussion on the issue, we consult the 1994-95 Teacher Followup Survey in an effort to provide some independent validation of our conclusions. While this national survey of teachers does not provide direct evidence on what individuals actually do when they leave teaching, its circumstantial evidence in the form of motives and anticipated activities is strongly consistent with our results.

“New college graduates as well as seasoned teachers are being lured to other professions with handsome salary offers while the teaching profession often isn’t even in the horse race. (July 5, 2000)”<sup>1</sup>

Sandra Feldman, President American Federation of Teachers (AFT)

## I. Introduction

Despite a large recent literature that examines the labor force decisions of elementary and secondary school teachers, some of the most fundamental issues related to teacher attrition are not well-understood. As one prominent example, there is almost no evidence regarding the very common perception that the majority of teacher attrition is caused by the attractiveness of high-paying alternative occupations. Understanding “where teachers go” when they leave the teaching profession is crucial for numerous current policy debates, especially issues of teacher compensation.<sup>2</sup> In this paper we examine this issue using new and unique data. We have assembled this database by merging information obtained from the public school system in Georgia with administrative payroll records from its state unemployment insurance system. In these payroll records, we observe the actual wages paid to all teachers and former teachers in the state of Georgia.

Projections of increases in the demand for teachers and increases in the number of teachers who will retire in the near future has heightened interest in designing effective policies to address teacher attrition (Darling-Hammond, 2000; U.S. Department of Education, 1997).<sup>3</sup> A necessary condition in this endeavor is to understand what individuals do after they leave teaching. As the quote above and the quotes in Figure 1 help to illustrate, teachers’ unions, the popular press, academics who are active in education policy debates, and politicians regularly portray a scenario in which teachers are lured away from teaching by high-paying jobs in other occupations. This view of teacher attrition is also adopted in

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<sup>1</sup> This quote can be found in an AFT press release available at: <http://www.aft.org/press/2000/070500.html>

<sup>2</sup> Using information in the 2001 Digest of Education Statistics, we estimate that nationally about \$124 billion was spent on wages for the almost three million public elementary and secondary school teachers in 1998-99. For context, this sum is more than the U.S. federal government spends on TANF, food stamps, SSI, and housing assistance combined.

<sup>3</sup> The recruitment of new teachers is also important in ensuring that there are enough good teachers. Although we do not focus on this issue in this paper, we stress that the quality of the teachers who are recruited and retained is clearly important from the standpoint of school quality. For example, Rivkin et al. (1998) finds that variation in teacher quality—largely unmeasured teacher quality—has a large impact on student achievement.

the work of economists who often describe the relevant decision margin as a choice between staying in teaching and moving to a new career.<sup>4</sup> Nonetheless, despite the strength of this perception, there exists virtually no research which examines whether it is true. The reason for this is largely a practical matter. Virtually all previous research in this area utilized “teacher-specific” data that were constructed from the administrative records of a particular state or school district, and, as a result, contained no information on the labor force status of individuals after they left teaching.

An exception is Stinebrickner (2002) who found evidence that the common perception may not be true in the U.S. Using data from the National Longitudinal Study of the High School Class of 1972 (NLS72), he found that 67 percent of all exiting female teachers left the workforce altogether. Of those who remained in the workforce, only half earned a higher wage in their new occupation. Among general longitudinal data sources, the NLS72 was a logical choice because the survey design involved oversampling teachers and because the data includes information from a teacher supplement. However, even with oversampling, the number of teachers in the NLS72 is very small relative to the number of teachers in teacher-specific data used by other researchers. Perhaps more importantly from the standpoint of current policy debates, the majority of teachers in his sample graduated from college in 1976.<sup>5</sup>

The fundamental policy importance of the finding that the large majority of teaching exits are not related to higher paying alternative occupations and the possibility that this result may not exist today given the dramatic changes in fertility and female labor supply that have occurred in the last twenty-five years, motivate us to examine this issue using very recent data from the state of Georgia. We find very

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<sup>4</sup> For example, the theoretical models used to motivate the empirical work in Murnane and Olsen (1989, 1990) assumed that teachers who leave teaching do so to begin work at their best non-teaching job alternative. Similarly, recent work by Gritz and Theobald (1996) suggested that a teacher's "decision to remain in his or her current teaching assignment depends upon the perceived benefits of this choice relative to alternative career opportunities."

<sup>5</sup> Dolton and van der Klaauw (1999), using a sample of 923 UK teachers who graduated college in 1980, also had information on reasons for exits from teaching. They estimated two competing risks models of teacher transitions that differentiated between exits to a non-teaching occupation and exits out of the workforce. Like Stinebrickner (2002), their sample size is small. Neither paper considers as a separate transition former teachers remaining in the public education sector in some other capacity.

strong evidence that new female teachers are not leaving the teaching profession for higher paying jobs in alternative occupations. Only 3.8 percent of elementary teachers and 5.4 percent of high school teachers who left full-time teaching during our sample period took a non-education sector job in Georgia that paid more than the state minimum teaching wage in Georgia. Further, simulations using a competing risks model suggest that only a very weak relationship exists between teaching wages and the teaching exits of females.

To focus the discussion as much as possible, we concentrate primarily on women because this group represents approximately 83 percent of the Georgia teaching force.<sup>6</sup> Given possible differences in the way that males and females respond to family changes, it is possible that males have different exit patterns and/or are more responsive to wages. We discuss this possibility in the conclusion, at which point we refer to results that we have included, for the sake of paper brevity, in Appendix A. Although we do find certain differences between male and female teachers, the general spirit of the results for males is very similar to that for females. For example, we find that males also exhibit low rates of exit to relatively high paying non-education sector jobs.

Given that our results are in direct contrast to public discussion on the issue, we consulted the 1994-95 Teacher Followup Survey in an effort to provide some independent validation of our conclusions. While this national survey of teachers does not provide direct evidence on what individuals actually do when they leave teaching, as discussed below, its circumstantial evidence in the form of motives and anticipated activities is strongly consistent with our results.

Section II describes the data, while section III contains an analysis of the transitions made by full-time teachers. Given new and unique observations of “where teachers go,” we estimate competing risks models of teacher transitions that differentiate between types of exits from full-time teaching. The specification and results of these models are presented in section IV. Section V provides a discussion of

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<sup>6</sup> This percentage has remained remarkably stable over the 1990s.

the policy implications of the results for teacher compensation and teacher retention policies and concluding remarks.

## II. Data

To analyze the transitions of teachers, we merged three sources of data on all public schools teachers and all public schools in Georgia.<sup>7</sup> Data on the characteristics of individual teachers from the 1991-92 school year to the 2000-01 school year were obtained from the administrative records kept by the Georgia Professional Standards Commission (GAPSC). Characteristics of individual schools from 1994-95 to 1999-2000, including racial composition, average student achievement on standardized exams, and percent of students eligible for free or reduced lunch were provided by the Georgia Department of Education (GADOE). What makes the assembled database most unique comes from the third source of data. Actual quarterly wages paid to the teachers *and former teachers* comes from administrative payroll records from the state unemployment insurance (UI) system maintained by the Georgia Department of Labor (GADOL). These data, referred to as ES202 data, are described in detail by White, et al. (1990). All employers covered in the unemployment insurance system report each employee's wages to the GADOL on a quarterly basis. The ES202 data identify the industry (4-digit SIC code), but not the occupation of each individual.

Using ES202, the actual wages paid to teachers and former teachers were matched with the teacher records in the GAPSC files by social security number. For individuals listed in the GAPSC files as

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<sup>7</sup> Georgia is in many ways roughly an “average” state. Georgia’s median household income in 1999 was about \$39,500, ranking 26<sup>th</sup> in the U.S. Women working in Georgia earned \$934 more than the national median for women. Georgia students in the 3<sup>rd</sup>, 5<sup>th</sup>, and 8<sup>th</sup> grades scored very close to or at the national average on the battery of Stanford 9 exams administered in 2001 (Georgia Department of Education, 2002). Almost 83 percent of Georgia residents have at least a high school diploma compared to 84 percent nationwide. According to the National Education Association, teacher salaries in Georgia are slightly below the national average, the highest in the southeast, and 17<sup>th</sup> highest in the nation (Salzer, 2001). The rate of adults over age 25 with at least a college degree is lower in Georgia relative to the nation—23.1 percent in Georgia and 25.6 percent overall. Georgia differs from the rest of the nation with respect to homeownership and racial composition as well. In 1999, 71.3 percent of Georgians owned their homes compared to a rate of 66.8 percent nationally. Nationally, 30.9 percent of individuals are nonwhite or of mixed race, compared to 37.4 percent of Georgians. Unless otherwise noted, all information in this footnote comes from the U.S. Department of the Census web site, [www.census.gov](http://www.census.gov).

teachers for a given year, the match of wages to teachers was almost perfect.<sup>8</sup> Wages paid to former teachers are also observed in the ES202 data *if they are employed in Georgia*. When an individual teacher disappeared from the GAPSC data after teaching the previous year, we searched the ES202 data to see if the individual earned wages in another occupation in that and subsequent years.

Virtually all employees are subject to the UI tax, and thus virtually all wages in non-teaching occupations are observed.<sup>9</sup> As a result, if a former teacher does not have a wage in the ES202 file then she is either living in Georgia but not working, living in another state and not working, living in another state and working as a teacher, or living in another state and working in a non-teaching occupation. The data do not allow us to differentiate between these possibilities so we group them together in a “leaving the Georgia workforce” category. In subsequent sections we pay careful attention to the implications of this classification. Appendix B discusses how we determine if someone is a teacher, if a teacher has left teaching, construct teaching wages, and define new teachers.

### **III. Transitions out of Full-time Teaching**

For the primary analysis in this paper, we examine exits out of full-time public school teaching for the 10,145 new female elementary teachers and the 4,750 new female high school teachers who began their teaching careers between the academic years 1994-95 and 1999-00 and were less than twenty-seven years old at the time that their teaching careers began. Although our teacher identifiers and unemployment insurance records begin in 1991-1992, we use 1994-1995 as the initial period for our primary analysis because, as discussed in Section II, we do not observe school characteristics before this

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<sup>8</sup> As evidence of the quality of the ES202 data, we found that of the roughly 820,000 teacher records over the 10-year period, only 7 records could not be matched to wage information in the ES202 data files.

<sup>9</sup> The list of excluded employment is quite small. The three main categories perhaps most likely to apply to former teachers are the self-employed, elected officials, and occupations where the individual is paid solely by commission (However, many salespeople are paid a small wage along with any commission earnings.) Thus, individuals who were working solely in an excluded occupation would be mislabeled as not working in Georgia. Also, we would understate wages for individuals with more than one job during the year, where at least one job was not covered by UI. Other occupations not covered by UI include caddies, some maids, and newspaper carriers (Section 108.02(15) of the UI Handbook for Employers). Of course, ES202 will miss wages paid in the underground economy.

year. This is important because school characteristics may play an important role in determining why teachers leave teaching and because the school characteristic information is needed to make a potentially important distinction between elementary and high school teachers.

We classify an individual as a full-time teacher if she receives at least the state minimum pay in any of the quarters during the teaching year.<sup>10</sup> The primary rationale for this definition, which is discussed fully in Appendix B, is that it allows us to avoid classifying very short spells out of the workforce as exits. Note that the use of quarterly wage data from the unemployment insurance system is beneficial in this regard because it allows us to differentiate between a full-time teacher who may have taken a short break from the labor force (perhaps, for example, for health or maternity reasons) and a teacher who is part-time for the entire year.

Figures 2a and 2b show Kaplan-Meier survivor functions associated with the first teaching spell of female elementary and high school teachers respectively. The Kaplan-Meier Survivor function takes into account that the first teaching spells are censored for 6,498 elementary teachers and 2,870 high school teachers. That is, these teachers remain in full-time teaching continuously through the end of the sample period, which occurs in 2000-2001. A given survivor function evaluated at time  $t$  shows the probability that a teacher in our sample will teach  $t$  or more years before leaving full-time teaching. Thus, only approximately half of all elementary and high school teachers have first teaching spells that last more than five (six or more) years. This is consistent with the findings of previous studies of teacher attrition.

The rate at which teachers leave teaching, as shown in Figure 2a and Figure 2b, is well-known. What is not well-known is what teachers do after leaving teaching. For each of the 3,647 elementary and 1,880 high school teachers who exited teaching during the sample period we examine the activity state in the first year that she is not a full-time teacher. Let  $T-1$  be the last year of full-time teaching for

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<sup>10</sup> Mandated by state law and updated annually, Georgia public schools systems face minimum salary schedules for all teachers they employ. The salary schedule lists the minimum teacher salary that can be paid based on teacher certification status, experience, and education. Local school districts may pay teachers a local supplement to the salary schedule and a large majority do so. Local supplements to the salary schedule vary widely.



individuals who exit the teaching force. Denoting the exit year as  $T$ , we determine whether each exiting teacher continued to work for the Georgia public educational system in some position other than as a full-time teacher in year  $T$  (e.g., as an administrator, counselor, part-time teacher, or substitute teacher) and whether the person held a job outside the public educational system in Georgia in year  $T$ .<sup>11</sup> We call the former the “education sector.” We inflate all earnings to 2001 dollars (using the CPI-U) and discretize the earnings in each of the two “sectors” (education and non-education) into the following earnings categories:  $\$0$ ,  $(\$0, \$10,000)$ ,  $[\$10,000, \text{Minwage}(T))$ , and  $[\text{Minwage}(T), \infty)$ , where  $\text{Minwage}(T)$  is the mandated minimum teaching wage in Georgia in year  $T$ .

Table 1 shows bivariate sample probabilities for elementary teachers in year  $T$ , where a particular element in the table shows the proportion of exiting teachers in the sample who, in their first year out of full-time teaching, have a particular combination of earnings in the education and non-education sectors in Georgia. For example, the number 0.0359 indicates that, in the year following their exit from teaching, about 3.6 percent of all exiting teachers earned a wage between  $(\$0, \$10,000]$  in both the education and the non-education sector. The bivariate sample probabilities for high school teachers are shown in Table 2.

The marginal sample probabilities for earnings in the Georgia education sector are shown in the last column of the table. The marginal probabilities indicate that many teachers who leave full-time teaching remain in the public education sector in some capacity.<sup>12</sup> We refer to (non-teaching) jobs in the education sector with earnings above the teaching minimum as administrative jobs. Six percent of exiting elementary teachers have earnings in the public education sector that are greater than the minimum teaching salary in the state. Twenty-three percent of exiting elementary teachers have earnings in the public education sector that are greater than  $\$10,000$ . The jobs below the minimum are a combination

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<sup>11</sup> Clearly the exit year,  $T$ , varies by person. We suppress person specific notation on this variable.

<sup>12</sup> This is consistent with Brewer (1997) who reports that it is very common for administrators in New York State to be former teachers.

of part-time teaching jobs and part-time administration jobs.<sup>13</sup> This evidence from Georgia is consistent with circumstantial evidence from the 1994-95 Teacher Followup Survey in which 22.2 percent of exiting teachers reported that they expected to be “working in a non-teaching occupation in education” in the year after leaving teaching (NCES, 1997).

The marginal sample probabilities for earnings in the non-education sector are shown in the last row of each of the two tables. The marginal probabilities provide strong evidence that the primary reason for teaching exits is not the attractiveness of non-teaching jobs. Only 3.8 percent of all exiting elementary teachers accept a non-education job in Georgia that pays more than the minimum teaching wage in the state. Only 9.9 percent of exiting elementary teachers work in a non-education job in Georgia that pays more than \$10,000 in 2001 dollars. As shown in Table II, these numbers are only slightly higher, 5.4 percent and 12.5 percent, for high school teachers.<sup>14</sup>

Our methodology of examining earnings in the first year after an exit from teaching will tend to understate the importance of high-paying alternatives if some teachers leave for occupations that initially pay below current teaching salaries but have higher future wage growth. To examine this possibility, we computed marginal sample probabilities based on the highest annual non-teaching wages earned after exiting the public education sector.<sup>15</sup> We found that only 6.4 percent of elementary and 9.5 percent of high school teachers had non-education earnings of more than the Georgia minimum teaching wage at any point before the end of the sample period. Further, this finding that teachers are not leaving for higher

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<sup>13</sup> Stinebrickner (2002) finds that 0.67 of exiting teachers leave the full-time workforce entirely after leaving teaching. Included in the remaining 0.33 are administrators, counselors, and others who remain in the education system. The importance of exits to other jobs in the education system suggests that the 0.33 number may substantially overstate the number of people who are working in the non-education sector.

<sup>14</sup> One reason that a former teacher may have left the Georgia workforce is to pursue additional education. However, Stinebrickner (2002) finds that “The data indicate that returning to school full-time is not a common occurrence for teachers; only 8 of the 172 female teachers who are observed exiting the work force enter school full-time.” Further, his data suggest that former female teachers are far less likely to go back to school than females who exit other occupations. Specifically, in his data, 66 out of 239 (28 percent) of females who left other occupations went back to school.

<sup>15</sup> This exercise may also allow us to capture the future wages on former teachers who initially went back to school after exiting the public education sector.

paying non-education jobs in Georgia was found to be robust to changes in the ages of the teachers that are examined and robust to a change in the sample period that allows us to follow teachers for more years after exit.<sup>16</sup>

Our evidence that very few teachers leave the profession for higher paying jobs in alternative occupations is consistent with circumstantial evidence from the 1994-95 Teacher Followup Survey in which only 12.1 percent of teachers who left teaching reported that the “main reason for leaving the teaching profession” was to pursue another career and only 6.5 percent reported that they left teaching for better salary or benefits (NCES, 1997). When asked about their expected main activity for the following year, 9.4 percent of exiting teachers responded that it was “working outside the field of education” (NCES, 1997).

Recall that our data contain information regarding only individuals who are working in Georgia. Thus, we do not observe the number of exiting teachers who work in a non-education job in another state, and, as a result, we cannot calculate the total percentage of exiting teachers who work in a non-education job. Nonetheless, based on both supply and demand considerations, it seems natural to think that observing a non-trivial number of individuals leaving teaching for non-education jobs in Georgia would be a necessary condition for believing that many individuals are leaving teaching to work in non-education jobs outside of Georgia. From a supply perspective, it would seem that if teachers have a strong desire to work in non-education jobs many would find searching for such jobs in Georgia to be desirable. In general, finding potential job matches and participating in the relevant interviewing/hiring process is likely

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<sup>16</sup> For example, one group considered was female teachers who were aged 27 to 30 when they began their Georgia teaching career (although they may have been teachers in other states). This group is about 55 percent as large as the group of teachers considered in this paper (new female teachers aged 26 or less when they began). For the group aged 27 to 30, these teachers were much more likely to remain in full-time teaching than the younger female teachers considered in tables 1 and 2. More importantly, teachers in this slightly older age group had virtually the same rate of leaving for non-education sector occupations and earning more than Minwage. In another example, of all females (in all age groups) in the teaching force in 1991-92 who exited teaching in one of the next two years, less than 9 percent earned non-education sector wages in Georgia that exceeded the Minwage at any point between 1992 and 2001.

to be more costly outside of one's current geographic location. In addition, all else equal, it seems likely that the individuals in our sample (who have already revealed a preference for living in Georgia) would tend to value jobs close to their current locations to minimize or avoid incurring costs of moving. From a demand perspective, the economy in Georgia was very strong in absolute and relative terms during the time period covered by our data. According to the Census Bureau, the population of Georgia increased from about 6.5 million residents in 1990 to 8.2 million in 2000, the 4<sup>th</sup> largest increase in total population among states. Throughout the 1990s the unemployment rate in Georgia was below, and typically well below, the national unemployment rate.

Combining these supply and demand considerations, it seems highly unlikely that so few teachers would be observed in non-education jobs in Georgia if it is true that both a large number of teachers would like to leave teaching for non-education jobs and employers of high-paying jobs have a high demand for former teachers.<sup>17</sup> While we believe that this intuitive argument (that the total number of teachers who leave teaching for high-paying non-education jobs is small) is compelling, it is worth exploring whether informative upper bounds can be established on the total proportion of exiting teachers that could be working in high paying jobs in the non-education sector (inside and outside of GA). These upper bounds are generated by making the very strong assumption that *all former Georgia teachers who leave the state are employed in a non-education sector occupation that exceeds Georgia's minimum teaching wage*. In reality, many teachers who leave Georgia are likely to remain in teaching or other education jobs or leave the full-time workforce altogether.

For our calculations, we require an estimate of the proportion of former female teachers who leave Georgia. We use two strategies. The first strategy assumes that all of the teachers who leave our teaching sample and never work in Georgia again during our sample period are individuals who left the state. This

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<sup>17</sup> Explanations for a low demand for former teachers include teachers having lower average ability or that in order to become a teacher, one must invest in human capital that does not have a high market value in non-teaching occupations. Although this paper does not address this issue, if the latter were true, then this required investment in teaching-specific human capital could substantially hamper the recruiting of teachers.

is likely to be an extremely conservative assumption given that Stinebrickner (2002) finds that women who leave teaching (often to care for young children) frequently remain out of the workforce entirely for a large number of years. The second strategy uses information on male teachers to construct an estimate of the proportion of former female teachers who leave the state. The appeal of this approach is that male teachers who leave teaching and do not return to the workforce in Georgia within several years have almost certainly left the state.

Based on these two strategies, which are described in detail in Appendix C, our upper bound estimates of the proportion of exiting elementary school teachers who could be working in a non-education job and earning more than the minimum teaching wage are 0.355 and 0.266, respectively. Similar calculations for high school teachers lead to upper bounds of 0.358 and 0.272, respectively. We stress again that these upper bounds are almost certainly much too high since it is unlikely that a high percentage of teachers who leave Georgia work in high paying non-education jobs given that virtually no exiting teachers enter high paying non-education jobs in Georgia.<sup>18</sup> Nonetheless, the bounds are of interest because, despite being extreme, they still do not indicate the type of important role for high paying non-education jobs that is often portrayed in public discussion.

The above analysis considers the proportions of *exiting* teachers who take high paying alternative jobs. For policy purposes, it seems equally important to compute the proportion of *all* new teachers who leave teaching for high paying jobs in the non-education sector. For our sample of new teachers, less than one percent leave teaching for a high paying non-education job in Georgia after the first year of teaching. This is consistent with circumstantial evidence from the 1994-95 Teacher Followup Survey in which 0.6 percent of current teachers reported that they expected to be “working outside the field of education” in the following year. Further, using the most extreme estimate of the proportion of exiting

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<sup>18</sup> For example, our findings within Georgia suggest that even assuming that 50 percent of all exiting teachers who leave Georgia do so to work in a non-education job that pays more than the minimum teaching wage would be very cautious. In this case the upper bounds for elementary teachers become 0.196 and 0.152 and the bounds for high school teachers become 0.206 and 0.163.

teachers who accept high paying non-education sector jobs from above, only about five percent of new teachers exit to a high paying non-education jobs (inside or outside of Georgia) after the first year of teaching.<sup>19</sup> These percentages decline with tenure because the hazard rate out of teaching declines with tenure. It is worth stressing again that this number is computed under the extremely conservative assumption that all former teachers who are not observed earning wages in Georgia after exit are living in another state, working in a non-education sector job, and earning more than the Georgia state minimum teaching wage. Thus, for policy purposes, it seems safe to conclude that exits to high paying jobs do not represent an important phenomena for young female teachers.

#### **IV. How Sensitive are Teaching Decisions to Pay? A Competing Risks Model**

As suggested in Dolton and Van der Klaauw (1999), it is intuitively appealing to think that teachers who are considering leaving full-time teaching for higher paying jobs in alternative occupations would be more sensitive to changes in pay than teachers who are considering a reduction in hours worked or a complete departure from the workforce (perhaps e.g., for reasons related to children). As a result, our new findings about the rarity of teacher exits to high-paying alternative jobs suggest that it is worthwhile to examine the relationship between teaching exits and wages using our new data.

The large literature on teacher attrition typically finds a negative relationship between teaching wages and teaching exits. The joint presence of three valuable features of our data allows us to make a new contribution to this literature. First, we are able to differentiate between different reasons for exits. From the standpoint of understanding why higher wages matter, this will be important if higher wages have stronger effects on certain types of exits than on other types. Second, we are able to include information about non-wage characteristics of jobs. As discussed in Loeb and Page (2000) and Hanushek et al. (2001), this will be important if wages are correlated with school characteristics. Finally, as

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<sup>19</sup> These calculations are  $(0.136*0.355=0.048)$  for female elementary teachers and  $(0.153*0.358=0.055)$  for female high school teachers. The first number in each parentheses is the probability of leaving after the first year of teaching as displayed in figures 2a and 2b. The second number in each parentheses is the most conservative estimate of the total proportion of exiting teachers who earn more than the minwage in a non-education sector job.

discussed in Appendix B, our access to quarterly wage data allows us to pay careful attention to the construction of appropriate salary measures. In other studies that use actual yearly earnings records from educational data sources, a spurious relationship between earnings and exits may be generated by the fact that a teacher who leaves teaching before the end of a year will have lower earnings than a person with the same salary who does not leave teaching during the year.<sup>20</sup>

To examine the effects of wages (and other characteristics) on teacher exits, we estimate a discrete-time competing risks model. A hazard function for discrete-time data in a single risk duration model represents the probability that a person leaves full-time teaching (F) for any reason in a given year  $t$ , conditional on not having left before year  $t$ . The competing risks hazard model used here makes a further distinction between exits to other jobs within the Georgia public education system (E), exits to other jobs not in the Georgia public education system (N), and exits out of the Georgia workforce (O).<sup>21</sup> We classify individuals who earn less than \$10,000 in annual wages as out of the Georgia workforce.

The model is estimated by maximum likelihood. Define  $P_{it}^j$  to be the probability that at the end of her  $t^{\text{th}}$  year in teaching, teacher  $i$  chooses activity  $j$ ,  $j=F,E,N,O$ , for time  $t+1$ . There are two cases to consider. First, suppose a person's spell in teaching is censored after  $S$  years in teaching. In this case, the likelihood contribution for teacher  $i$  is the probability that at the end of years  $1,2,\dots,S-1$ , the person decides to return to teaching for the next year

$$(1) \quad L_i = P_{i1}^F \cdot P_{i2}^F \cdot \dots \cdot P_{iS-1}^F.$$

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<sup>20</sup> By using the highest quarterly wage to construct our measure of annual teaching wages, we more closely obtain a measure of wages on which individual teachers base their transition decisions. Administrative or survey data that have information only on total wages paid during the year would be more likely to find larger effects of teaching wages on exits on full-time teaching—because teachers with artificially low observed “annual” wages (artificial, because teachers left in the middle of a year) may be more likely to not be teaching in the following year.

Given the rigidity of the salary structure in most public schools, a reasonable alternative approach is to construct salary information for individual teachers from district or school specific salary schedules as was done in Hanushek, et al. (2001). For most studies of teacher attrition the source of information on teaching wages is not reported.

<sup>21</sup> Stinebrickner (2002) uses a continuous-time hazard model. There is no obvious benefit to the continuous-time model here given that our activity data are at a yearly level. Results did not change in important ways when we estimated the continuous-time analog to the specification described below.

The likelihood contribution is similar in the alternative case where the person is not censored. Suppose a person teaches for  $S$  years and then leaves teaching for option  $k \in \{E, N, O\}$ . In this case, the likelihood contribution for the person is the joint probability that at the end of years 1, 2, ...,  $S-1$ , the person decides to return to teaching for the next year and decides at the end of year  $S$  to have activity state  $k$  in time  $S+1$

$$(2) \quad L_i = P_{i1}^F \cdot P_{i2}^F \cdot \dots \cdot P_{iS-1}^F \cdot P_{iS}^k.$$

We define  $P_{it}^j$  to have a multinomial logit form

$$(3) \quad P_{it}^j = \frac{e^{X_{it}\beta^j + B^j(t)}}{e^{X_{it}\beta^F + B^F(t)} + e^{X_{it}\beta^E + B^E(t)} + e^{X_{it}\beta^N + B^N(t)} + e^{X_{it}\beta^O + B^O(t)}}$$

where  $X_{it}$  is a vector of observable characteristics of person  $i$  and the school at which person  $i$  works at time  $t$ .  $B^j(t)$  is a function which is used to determine how the probability of choosing a particular option  $j$  changes with the number of years,  $t$ , that a person has been in her teaching spell. We assume a non-parametric form

$$(4) \quad B^j(t) = \delta_1^j I(t=1) + \delta_2^j I(t=2) + \delta_3^j I(t=3) + \dots + \delta_6^j I(t=6)$$

where  $I$  is an indicator function that is equal to one if its argument is true. The number of terms in equation (4) comes from the fact that a maximum number of six decision years can be observed for person  $i$ .<sup>22</sup>

The likelihood function for the sample is given by  $\prod_i L_i$ .<sup>23</sup> The coefficient vector  $\beta^F$  and the coefficients in  $B^F(t)$  are normalized to zero and the remaining coefficient vectors  $\beta^j$ ,  $j=E, N, O$  and the parameters of  $B^j(t)$ ,  $j=E, N, O$  are estimated by maximum likelihood.

The observable characteristics which are included in  $X_{it}$  are described in Table 3 for the 10,145 elementary teachers and 4,750 new high school teachers who fit the sample criteria--female and under age

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<sup>22</sup> The first teaching year is 1994-95. The last year of our sample is 2000-01.

<sup>23</sup> In this specification, conditional on the observable characteristics, the year specific likelihood contributions are independent. We have also estimated the model allowing for unobserved heterogeneity. In this case, the likelihood contributions for person  $i$  are found by integrating equations (1) and (2) over the distributions of the heterogeneity. The results from this estimation were very similar to those from the homogenous model that is described below.



27 when they began teaching in Georgia. Along with log teaching wages, variables included in  $X_{it}$  are whether the teacher is nonwhite and three school-level characteristics: an average test score of the students, proportion of students eligible for free or reduced lunch (poverty), and percent of students who are black. To capture local labor market conditions, we included proximity to Atlanta and other cities,<sup>24</sup> a dummy variable to indicate metropolitan statistical area, and county unemployment rate.<sup>25</sup>

The estimation results from the competing risks models are shown in Tables 4 and 5 for elementary and high school teachers, respectively. As discussed above, for both models the reference activity is full-time teaching (F). Here we focus the discussion primarily on the effects of wages.<sup>26</sup> Higher teaching wages are found to decrease the benefit of non-education jobs relative to remaining in full-time teaching for high school teachers (t-statistic of -2.24) but not for elementary teachers (t-statistic of -0.88). For both elementary and high school teachers, higher teaching wages decrease the benefit of other public education sector jobs (administration or part-time teaching) relative to staying in full-time teaching (t-statistics of -1.8 and -2.0 respectively). However, our results indicate no relationship between teaching wages and the benefit of leaving the Georgia workforce relative to remaining in full-time teaching.<sup>27</sup>

In order to quantify the importance of the wages, we use our estimates in Table 4 and Table 5 to compute the first-year exit probability associated with each of the exit reasons for a “baseline” elementary

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<sup>24</sup> Distances to Atlanta and to medium-sized cities were measured from the geographic centroid of each school district to the centroid of Atlanta and the centroid of the nearest city with over 75,000 population (excluding Atlanta).

<sup>25</sup> Murnane and Olsen (1989, 1990) and Dolton and Van der Klaauw (1999) constructed estimates of “opportunity wages” for individual teachers based on wages paid to non-teachers. Both studies found that opportunity wages were positively and significantly related to exits from teaching. Given that identification of opportunity wages is tenuous, we do not attempt to construct estimates of them. There are many other variables such as fertility, number of children, total household income, and ability that could also explain exits out of full-time teaching but are not available in our data. Many of these variables have been used in other studies and have been found to be related to teacher exits.

<sup>26</sup> See Scafidi, Sjoquist, and Stinebrickner (2002) for an examination of the relationship between school characteristics and teacher decisions that uses these data.

<sup>27</sup> Dolton and van der Klaauw (1999) estimated two competing risks models of teacher transitions and found a negative and statistically significant relationship between teaching wages and exits out of the workforce.

teacher and a “baseline” high school teacher at “baseline schools” and then compare these probabilities to those obtained after increasing the teaching wage by \$5,000.<sup>28</sup> The sum of the first-year exit probabilities associated with the three transition risks is 0.162 and 0.147 for the elementary and high school baseline teachers respectively. Increasing wages by \$5,000 decreases the predicted probability of leaving the first job in the first year by very little—to 0.160 and 0.141 respectively.<sup>29 30</sup> This evidence that the relationship between teaching wages and exits may be weak is consistent with circumstantial evidence from the 1994-95 Teacher Followup Survey. For teachers who left the teaching profession at the end of the 1993-94 academic year and expressed “dissatisfaction with teaching as a career” as one of three main reasons for leaving the profession, only 10.7 percent listed “poor salary” as the main area of dissatisfaction with teaching (NCES, 1997).<sup>31</sup>

We estimated a variety of specifications of the competing risks models that are not reported here. Specifically, we included different variables measuring local labor market conditions (region unemployment rates, median county earnings, median region earnings, and dummy variables for 12

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<sup>28</sup> The teaching wage used in all simulations in the paper is \$31,571. This is the wage of a person who has the average log wage of elementary teachers in our sample. The baseline person is a white teacher with all other explanatory variables set to the sample means.

<sup>29</sup> For the elementary teacher, the baseline probabilities associated with the non-education job, new education job, and out of the Georgia workforce alternatives are 0.015, 0.033, and 0.114 respectively. The probabilities associated with the wage increase are 0.014, 0.031, and 0.115 respectively. For the high school teacher, the baseline probabilities associated with the non-education job, new education job, and out of the Georgia workforce alternatives are 0.015, 0.040, and 0.091 respectively. The probabilities associated with the wage increase are .012, .036, and .091 respectively.

<sup>30</sup> We find a statistically significant and positive impact of percent black students on exits from the education sector, and, curiously, a negative effect of percent students in poverty on these transitions. Both results are consistent with the findings in Hanushek et al (2001) and Clotfelter et al (2002). Hanushek et al (2001) and Clotfelter et al (2002) use administrative data on public school teachers from Texas and North Carolina, respectively, to estimate models of exits from teaching. Although both studies find that percent of students in poverty has a negative effect on exits from teaching, neither study can differentiate reasons for exits from teaching.

<sup>31</sup> In descending order, teachers who were dissatisfied with teaching were most likely to report student discipline problems, poor student motivation to learn, inadequate support from administration, and lack of recognition and support from administration as the main reason for leaving the teaching profession.

regions of the state). We also estimated models that allowed for the presence of unobserved heterogeneity. The results reported here are robust to these alternative specifications.

## **V. Conclusions and Policy Implications.**

In this paper we examine “where teachers go” when they exit the teaching profession. We find that leaving teaching for higher paying alternative occupations is the exception not the rule. In fact, only 3.8 percent of exiting female elementary teachers and 5.4 percent of exiting female high school teachers in our sample took a non-education sector job in Georgia that paid more than the Georgia state minimum teaching wage. Larger percentages of exiting teachers remained employed in the public education sector in an administrative or part-time capacity. These transitions into other education sector jobs have been largely ignored in previous work on teacher attrition.<sup>32</sup> For our sample of new teachers, less than one percent leave teaching for a high paying non-education job in Georgia after the first year of teaching. Extremely conservative upper bounds suggest that, at most, approximately 5 percent of new teachers leave teaching for a “high-paying” non-education sector job (inside or outside of Georgia) after their first year of teaching. As discussed throughout the paper, this evidence from Georgia is consistent with circumstantial evidence about the motives and anticipated activities of teachers from the 1994-95 Teacher Followup Survey (NCES, 1997).

Nationally, a very large amount of taxpayer money (about \$124B) is spent on teacher salaries, and the quality of schools are directly at stake when teacher compensation policies are discussed. Nonetheless, despite the importance of this issue, our findings are at direct odds with the public dialogue on this issue. It seems extremely difficult to design effective policies involving teachers if, at the very least, the most fundamental issue of what teachers do when they

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<sup>32</sup> An exception is Brewer (1997).

leave teaching is not understood. Thus, in the short-run, more public discussion about these issues would seem to be warranted.

Policy debates involving teachers often seem to be premised on the assumption that teacher attrition rates are unacceptably high. Whether or not this is true is a complicated issue that depends on many factors such as the current supply and demand of teachers and the relationship between teaching experience and teacher quality. While we do not address this issue directly, our findings about the scarcity of exits to non-education jobs raise the possibility that many exiting teachers may eventually return to teaching (e.g., after children reach school-age). A thorough examination of the extent to which this is true would provide valuable information for policymakers, but would require a longer panel than that used in this paper.

Our paper does not attempt to say anything about whether teacher attrition rates are too high. However, if this is deemed to be the case, our results about what teachers do when they leave teaching suggest that it is worth considering policy initiatives that are unrelated to teacher pay. For example, it seems very possible that schools, which already have the infrastructure in place to take care of young children, might find it cost-effective to provide inexpensive on-site child-care.<sup>33</sup> In effect, this would give parents the new, potentially appealing option of being able to both work and be close to their young children. Other initiatives such as job-sharing may also be promising. We stress that, although our results suggest that the exit decisions of female teachers, when viewed as a whole, are not strongly related to teaching wages, raising wages may be important from the standpoint of retaining particular types of current teachers, recruiting gifted teachers in the future (Loeb and Page, 2000), and luring former teachers back into the teaching

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<sup>33</sup> Using NLS72, Stinebrickner (2001a, 2001b, 2002) finds a very strong relationship between the birth of children and teacher attrition. The fact that the types of exits in our data are very similar to those in the NLS72 data suggests that family reasons are likely to still be important determinants of teacher exits. However, we cannot examine this directly because our administrative data do not contain family information.

profession. However, our findings suggest that careful thought should be given to whether or not all future wage changes should be of the standard across-the-board variety.

In an effort to focus the discussion the paper as much as possible, we have concentrated on female teachers who represent over 80 percent of the new teachers in Georgia. In Appendix A we present analogous results for men. We do find some differences between male and female teachers. For example, men, who are presumably less likely to leave the workforce for family reasons, are less likely to disappear from our data entirely after leaving teaching, and males who leave teaching are roughly twice as likely as females who leave teaching to appear in non-teaching jobs that pay more than the Georgia minimum teaching wage. In addition, simulations using the competing risks model indicate that the relationship between wages and exits for males is somewhat stronger than for females. Nonetheless, the general spirit of the results for males is very similar to that for females. Specifically, many males who leave full-time teaching remain in the education sector in some capacity, and extremely conservative estimates indicate that, at most, one in twenty male teachers leave teaching after the first year for a non-education job that pays more than the minimum Georgia teaching wage.

We also note that our data do not allow us to examine the extent to which our results are representative of certain subgroups of teachers such as math/science teachers and teachers with high academic ability. Nonetheless, while these groups should be of interest to policymakers, it is worth noting that, in part due to the strength of teacher unions, the majority of current policy decisions are made on the basis of beliefs about the set of teachers viewed as a single entity. Thus, our results seem to provide timely new information that is directly relevant to current policy discussions.

Many previous studies have examined whether teachers leave the profession. Nevertheless, our results suggest the importance of using a new and large database that analyzes not only whether teachers leave, but “where teachers go” when they exit the teaching profession.

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## Appendix A

In this appendix we replicate the analyses for new male teachers. Given the small number of new teachers who are male (and the small proportion of males in the Georgia teaching force as a whole), we treat them as one group in this analysis. Thus, in all analyses that follow, male elementary and male high school teachers are analyzed together.

### *Attrition of Male Teachers*

Figure A.1 shows the Kaplan-Meier survivor function associated with the first teaching spell of all male teachers in the sample. Although males exhibit similar exit rates as females in the first couple of years, they are slightly more likely to teach more than five years. This is consistent with the findings of previous studies of teacher attrition.

### *Where Do Male Teachers Go?*

Table A.1 shows bivariate sample probabilities for male teachers in year T, where a particular element in the table shows the proportion of exiting teachers in the sample who, in their first year out of full-time teaching, have a particular combination of earnings in the education and non-education sectors in Georgia. Relative to females, higher proportions of exiting males have non-education sector wages in the year after leaving the Georgia teaching force. However, the proportions remain low: 9.5 percent of exiting male teachers accept a non-education sector job that pays more than the minimum teaching wage and 21.5 percent earn more than \$10,000 in the non-education sector.<sup>34</sup> As expected given likely differences in the way that males and females respond to family changes such as the birth of children, exiting males teachers are less likely to be observed earning without wages in Georgia after leaving teaching. Therefore, even though males are more likely to accept higher paying non-education sector jobs within Georgia, the extremely conservative upper bound estimates of the total proportion of exiting male teachers who earn more than the minimum Georgia teaching wage are only

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<sup>34</sup> Only 17.7 of males have non-education sector earnings greater than the Minwage in any year after leaving teaching.



0.315 and 0.264. These numbers represent extremely conservative upper bounds of the proportions *exiting* teachers who take high paying alternative jobs. For policy purposes, it seems equally important to compute the proportion of *all* new teachers who leave teaching for high paying jobs in the non-education sector. Using the most conservative estimates of the proportion of exiting teachers who accept high paying non-education sector jobs from above, about 4.5 percent of new teachers will exit to a high paying alternative job after the first year of teaching.<sup>35</sup>

As was the case for females, a non-trivial proportion of male teachers have earnings in the public education sector in the year after exiting teaching. However, males are much more likely to receive relatively high earnings in the public education sector. Exiting male teachers are more than twice as likely as exiting female teachers to accept non-teaching jobs in the education sector that pay more than the teaching minimum wage (marginal probability for exiting males equals 0.136).

#### *Competing Risks Results*

Summary statistics are reported in Table A.2, and the competing risks results for all male teachers can be found in Table A.3. Statistically significant and negative relationships exist between wages and exits to non-teaching jobs and exits out of the Georgia workforce. This latter finding is in contrast to females.<sup>36</sup> Simulations indicate that the relationship between wages and exits is somewhat larger than for females. The sum of the first-year exit probabilities associated

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<sup>35</sup> This calculation is  $(0.143 \times 0.315 = 0.045)$  for all male teachers. The first number is the probability of leaving after the first year of teaching as displayed in Figure A.1. The second number is the most conservative estimate of the total proportion of exiting male teachers who earn more than the Minwage in a non-education sector job.

<sup>36</sup> Males also seem to be more sensitive to local labor market conditions as there is a negative and statistically significant relationship between unemployment rates and exits to non-teaching jobs. Finally, there is no statistically significant relationship between student poverty and exits for males.

with the three transition risks is 0.137 for the baseline male teacher. Increasing wages by \$5,000 decreases the predicted probability of leaving the first job in the first year to 0.125.<sup>37</sup>

## Appendix B

In this data appendix we discuss three issues particular to our database that had to be addressed in order to conduct the analysis of teacher attrition: who is a teacher, how should teaching wages be constructed, and how should “new teachers” be defined.

### *Who Is a Teacher?*

Each record in the GAPSC data contains a job code, which is used to determine which individuals are teaching. Nevertheless, a large number of these teachers have low actual wages as reported by ES202. Perhaps these teachers were working for only part of the academic year. Since we do not know why these individuals are not earning a full-time annual teaching wage, we did not want to characterize their wage necessarily as their observed annual wage (see discussion about wages below).

In addition, some individuals who were a teacher in a given year and not listed in the GAPSC teacher files in a later year have a wage from the school district reported in ES202. The employer numbers in ES202 correspond to school districts, not individual schools. Therefore, we cannot identify which individual school or whether the district office employs these individuals. Furthermore, what they do for the school district is not observed. Although the work performed (teaching or non-teaching) by these individuals is unobserved, many of them have low annual wages relative to the minimum teaching wages mandated by Georgia’s statewide salary schedule.<sup>38</sup>

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<sup>37</sup> For the male teacher, the baseline probabilities associated with the non-education job, new education job, and out of the Georgia workforce alternatives are 0.019, 0.028, and 0.090 respectively. The probabilities associated with the wage increase are 0.015, 0.028, and 0.082 respectively.

<sup>38</sup> Mandated by state law and updated annually, Georgia public schools systems face minimum salary schedules for all teachers they employ. Local school districts may pay teachers a local supplement to the salary

Regarding the work performed by these individuals, who are not reported to be teachers but have reported wages from a school district, there are several possibilities. Some could be administrators in the district's central office. Others may be temporary, substitute, or part-time teachers or teachers who left teaching for part of the academic year. It is possible that some teachers were not reported to the state—either through a reporting error by the district or because these individuals were not teaching at the times teacher data are sent to the state. Districts are mandated to send information on all employed teachers to the state twice per academic year. Districts typically have a three week window in which to report teacher information. The first time window begins in mid-October and the second begins in mid-March. Perhaps some teachers who were employed only before, only after, or only in-between these windows are not reported to the state. Districts may report information on teachers outside this time window, but they are not required and have no incentive to do so.

To analyze teacher transitions, we had to classify these individuals (who are paid by the district, but are not recorded as teachers in the GAPSC teacher files) in some manner. Given the method of data reporting and typical low wages paid to these individuals, we deemed the most appropriate classification as “part-time teachers/administrators.” That is, these individuals may typically be former full-time teachers who have transitioned into part-time teaching, teaching only part of an academic year, or administration. In the empirical work discussed below, whether we include these individuals in this manner or treat them as full-time teachers does not impact the results.

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schedule and many do so. The salary schedule lists the minimum teacher salary that can be paid based on teacher certification status, experience, and education.

### *Constructing Teaching Wages*

Teaching wages are observed quarterly in the ES202 data: January-March, April-June, July-September, and October-December. Georgia teachers are paid on 12-month contracts.<sup>39</sup> Since the quarterly data do not match the school year, care had to be taken in constructing annual teaching wages. In the 3<sup>rd</sup> quarter of the calendar year, the ES202 data will contain wages for teachers from two different academic years. To avoid this issue and the issue of teachers leaving in the middle of an academic year, we took the highest quarterly teaching wage and annualized that figure. Teachers making decisions on whether to leave the profession surely consider the wage they would be paid for the entire academic year as the wage offered in teaching.

### *Who is a New Teacher?*

The time period that covers the sample of new teachers is the 1994-95 academic year through the 1999-00 academic year, giving us six years of data. In this paper, we focus on new teachers who were under the age of 27 when they began their teaching career. Georgia does not collect consistent information on teacher experience. Therefore, teachers are defined as “new teachers” if they had not been a teacher in Georgia since the 1991-92 academic year, the first year of our teacher files. Thus, teachers deemed new in 1994-95 will have not taught in a Georgia public school in any of the previous three years. This method of defining new teachers would include teachers who are new to Georgia, but have taught previously in another state.

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<sup>39</sup> In the early 1990s, some school districts gave teachers the option to be paid over 10 months. For districts this was difficult because money from the state earmarked for teacher salaries was paid in 12-month installments. Although ended in the mid 1990s, the 10-month option was grandfathered in some districts. Thus, it is possible that some veteran teachers are paid over ten months. However, the GAPSC has assured us that “99 percent” of teachers are paid on a 12-month calendar, and have been during our sample period. Since we are considering only new teachers in this paper, this issue is not a concern.

## Appendix C

In this appendix, we describe the methods used to construct an upper bound of the total number of former teachers who earn a wage in a non-teaching occupation that exceeds the minimum Georgia teacher salary (Minwage).

For sake of argument, consider the set of elementary teachers in year T-1 who exit teaching and live either inside or outside of Georgia in year T. The proportion of this group working in year T in non-education jobs that pay more than the Georgia minimum teaching wage is

$$\begin{aligned} \text{(C1) } & \text{PR}(\text{exiting teacher work in non-education sector, earnings} > \text{Minwage}(T)) = \\ & \text{PR}(\text{exiting teacher remains GA in } T \cap \text{work in non-education, earnings} > \text{Minwage}(T)) \\ & + \text{PR}(\text{exiting teacher leaves GA in } T \cap \text{work in non-education, earnings} > \text{Minwage}(T)) \\ & = .038 + \text{PR}(\text{exiting teacher leaves GA in } T) \\ & \quad * \text{PR}(\text{work in non-education, earnings} > \text{Minwage}(T) | \text{leave GA in } T), \end{aligned}$$

where PR represents probability.<sup>40</sup> We have no information about the term on the last line of equation (C1). Therefore, set this number to unity and compute an upper bound for equation (C1) as

$$\text{(C2) } \quad .038 + \text{PR}(\text{exiting teacher leaves GA in } T).$$

This implies that our bound assumes that **all** individuals who leave the state work in a high-paying non-education job in year T. This is not realistic since many of the teachers who leave the state will remain in teaching and many are likely not to work. Thus, we stress that the bound in (C2) will almost certainly be extremely conservative.

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<sup>40</sup> Recall that Minwage(T) equals the minimum teaching wage in Georgia at time T.

We use two approaches for obtaining the necessary information about PR(exiting teacher leaves GA in T). The first stems from the likelihood that virtually all teachers who leave the workforce at time T will eventually return to work in the future.<sup>41</sup> This implies that

$$(C3) \quad PR(\text{exiting teacher leaves GA at T})$$

$$\approx PR(\text{exiting teacher doesn't work in GA at T or after})$$

$$= PR(\text{exiting teacher doesn't work in GA in T})$$

$$*PR(\text{exiting teacher doesn't work in GA after T} | \text{exiting teacher doesn't work in GA at T})$$

$$=.458 * PR(\text{exiting teacher doesn't work in GA after T} | \text{exiting teacher doesn't work in GA at T})$$

where the .458 is the upper left entry in Table 1. If we had many years of UI records after T we could compute the probability term on the last line of equation (C3). Unfortunately, this is not the case. Instead, we compute the approximation of (C3) given by

$$(C4) \quad 0.458 * PR(\text{exiting teacher doesn't work in GA after T and before 2001/2002} | \text{exiting teacher doesn't work in GA at T}).$$

This will be a conservative approximation of  $0.458 * PR(\text{exiting teacher leaves GA at T})$ , if, as reported by Stinebrickner (2002), some women remain out of the workforce entirely for a non-trivial number of years before returning. Thus, using the approximation in equation (C4) to when computing equation (C2) will serve to make our upper bound even more conservative.

In an effort to tighten the bound as much as possible we estimate the probability in equation (C4) using only the exit years  $T=1993, 1994, 1995, 1996, 1997$ . Note that here we use several years that are earlier than our standard sample period in an effort to obtain a reasonable number of exits

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<sup>41</sup> Here we ignore the possibility that people leave Georgia and return at a future time.

where we observe a minimum of four years after  $T$ .<sup>42</sup> We observe a total of 970 exits during this time for which the exiting teacher received no earnings at time  $T$ .<sup>43</sup> Of these 970, 673 (69.4%) do not return to teaching after time  $T$  and before 2002. Thus, our estimate of equation (C4) is  $0.458 * 0.694 = 0.317$ , and thus our upper bound in equation (C2) of the proportion of exiting teachers who could be working in a non-education job earning more than the minimum teaching wage is  $0.038 + 0.317 = 0.355$ .

Our second approach is similar in spirit to the first but attempts to obtain information about departures from Georgia using information about male teachers. This approach is justified under an assumption that the proportion of young female teachers who leave Georgia is similar to the proportion of young male teachers who leave Georgia. The benefit of examining men is that we can potentially form a more accurate approximation of the proportion of exiting teachers who leave Georgia because few men are expected to leave the workforce for extended periods of time. This implies that if a young male does not have an income in Georgia for several years it is almost certainly the case that the person has left the state while, as discussed earlier, this may be less true for young females. We return to using teachers from our original sample period because observing a large number of years after  $T$  does not seem as important as it was for the females. Males spend a total of 5,450 person years in their first teaching spells between 1995 and 1999.<sup>44</sup> Of these person years, 160 (2.93%) are followed by an exit from teaching in which no earnings are observed in the

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<sup>42</sup> Our data contain information for the years 1991-92 to 2000-01. In other parts of the paper we have concentrated on new teachers who began teaching in 1994-95 or after. One reason for this is that information on school characteristics that will be used in subsequent analysis is only available during this period.

<sup>43</sup> For this exercise we have pooled all elementary, middle school, and high school teachers. This is necessitated by fact that we do not observe what type of school a teacher works in before 1995.

<sup>44</sup> Given the relatively small number of male teachers, for this exercise we pool all elementary, middle school, and high school male teachers in our data. We make 1999 the last year to ensure that each individual will have at least two years before the end of the sample period after his last year of full-time teaching.

exit year or any year after. Thus, this 2.93% is our upper bound estimate of the percentage of total teachers that leave the state in a particular year. This state exit rate estimate combined with information about the number of women who are teaching in each of the sample years and how many exiting teachers have zero earnings in time T yields an estimate that 49.9% of all the exiting female teachers who have no earnings in year T leave the state.<sup>45</sup> Thus, our estimate of equation (C4) is  $0.458 * 0.499 = 0.228$  and our upper bound in equation (C2) of the proportion of exiting teachers who could be working in a non-education job and earning more than the minimum teaching wage is  $0.038 + 0.228 = 0.266$ .

Similar calculations for high school teachers lead to upper bounds of 0.358 and 0.272 respectively under the two strategies.<sup>46</sup>

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<sup>45</sup> Assuming male and female former teachers have the same rate of interstate moves, the approximate proportion of female teachers who leave the state in year T after teaching in in T-1 is 0.029. The proportion of exiting females with zero earnings in year T who leave the state is found by dividing the total number of women who leave the state by the total number of exiting teachers with zero earnings in their exit year. We also computed numbers that took into account that the proportion of people who exit the state may vary with how many years a person has spent in teaching. This approach led to results very similar to the ones obtained by pooling all years.

<sup>46</sup> These numbers are calculated as  $0.054 + 0.438 * 0.694$  and  $0.054 + 0.438 * 0.499$  respectively.



## Figure 1

“We must begin to think about making teacher salaries competitive with other professions in order to convince our good teachers to remain and to attract college students who are pursuing other professions because of better salaries, benefits, and working conditions.”

Ralph Noble, President Georgia Association of Educators

October 24, 2001

<http://www.gae.org/about/communications/press/01/01salaryupdate.html>

“There is a little bit of data about where people go when they leave teaching ... In general, people who leave teaching go into fields that don't pay a significantly higher salary, but I believe working in Silicon Valley these days, that's probably not the case for math and science teachers.”

Linda Darling-Hammond, Charles E. Ducommun Professor of Teaching and Teacher Education

Stanford University

quoted in “Before It's Too Late: A Report to the

Nation from the National Commission on

Mathematics and Science Teaching for the 21<sup>st</sup> Century.”

(September 23, 1999)

<http://www.ed.gov/inits/Math/glenn/LDHtran2.html>

“Pay teachers respectable salaries. I have proposed a 10 percent pay raise for our teachers. Some say that is “unreasonable,” but we are in the middle of a major teacher shortage and Georgia needs to be able to attract and retain quality teachers for our children. How can we expect to keep our teachers in the classroom when the private sector can give them a higher salary? For the past two years, the governor has only recommended 3 percent pay raises for our teachers. That is unacceptable.”

Linda Shrenko, Georgia State School Superintendent (elected) and candidate in 2002 Governor's race

July 2001 Guest Editorial available at:

<http://www.doe.k12.ga.us/communications/releases/00/103000.html>

“The demand for new teachers is primarily due to teachers moving from or leaving their jobs and while it is true that teacher retirements are increasing, teacher turnover appears to have little to do with a graying workforce. In contrast, the high rates of teacher turnover that plague schools, teachers report, are far more often a result of two related causes: teachers seeking to better their careers and/or teachers dissatisfied with teaching as a career.”

Statement of Richard M. Ingersoll

Graduate School of Education

University of Pennsylvania

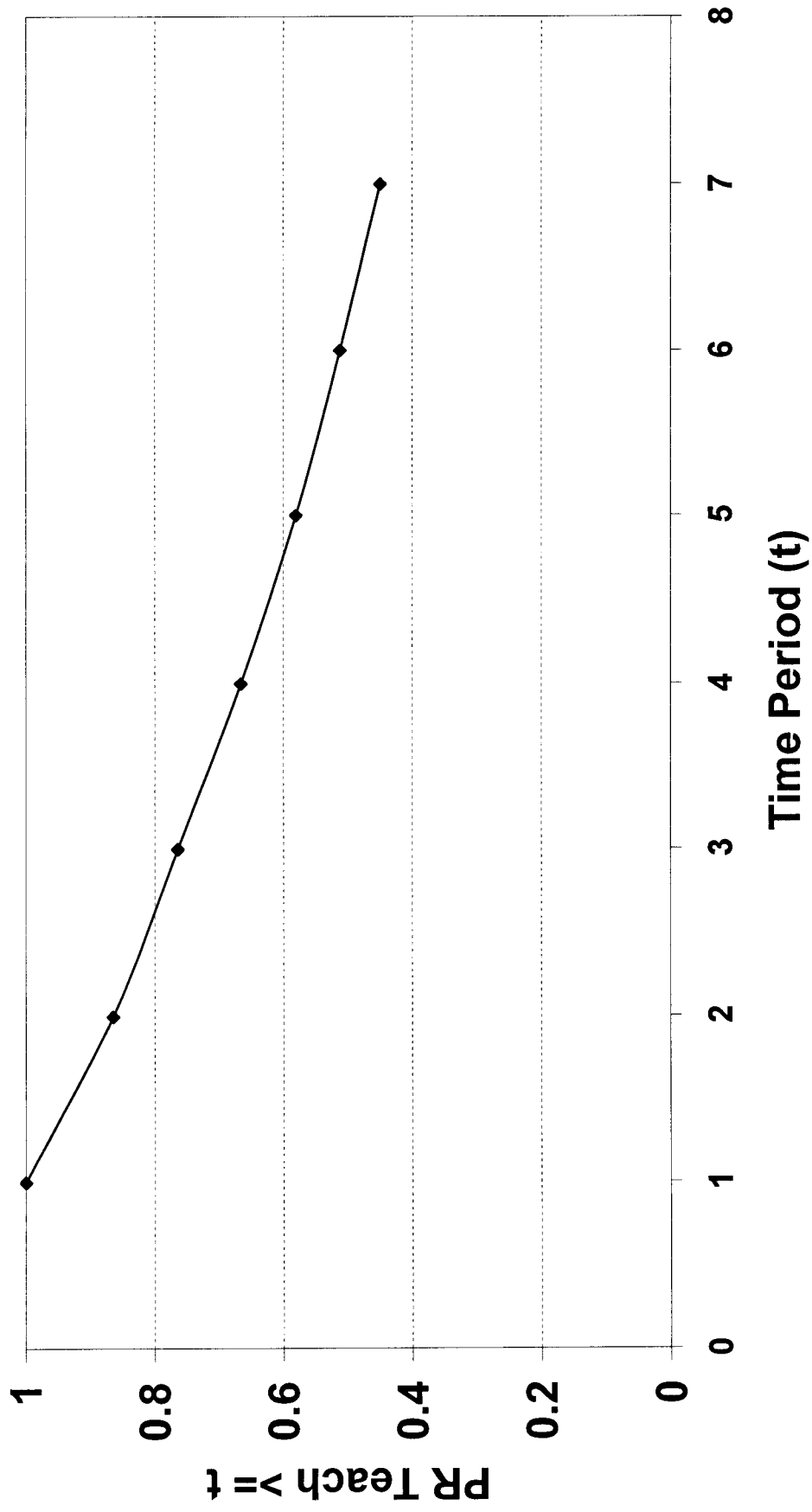
Testimony to the Subcommittee on Early Childhood, Youth and Families

Committee on Education and the Workforce

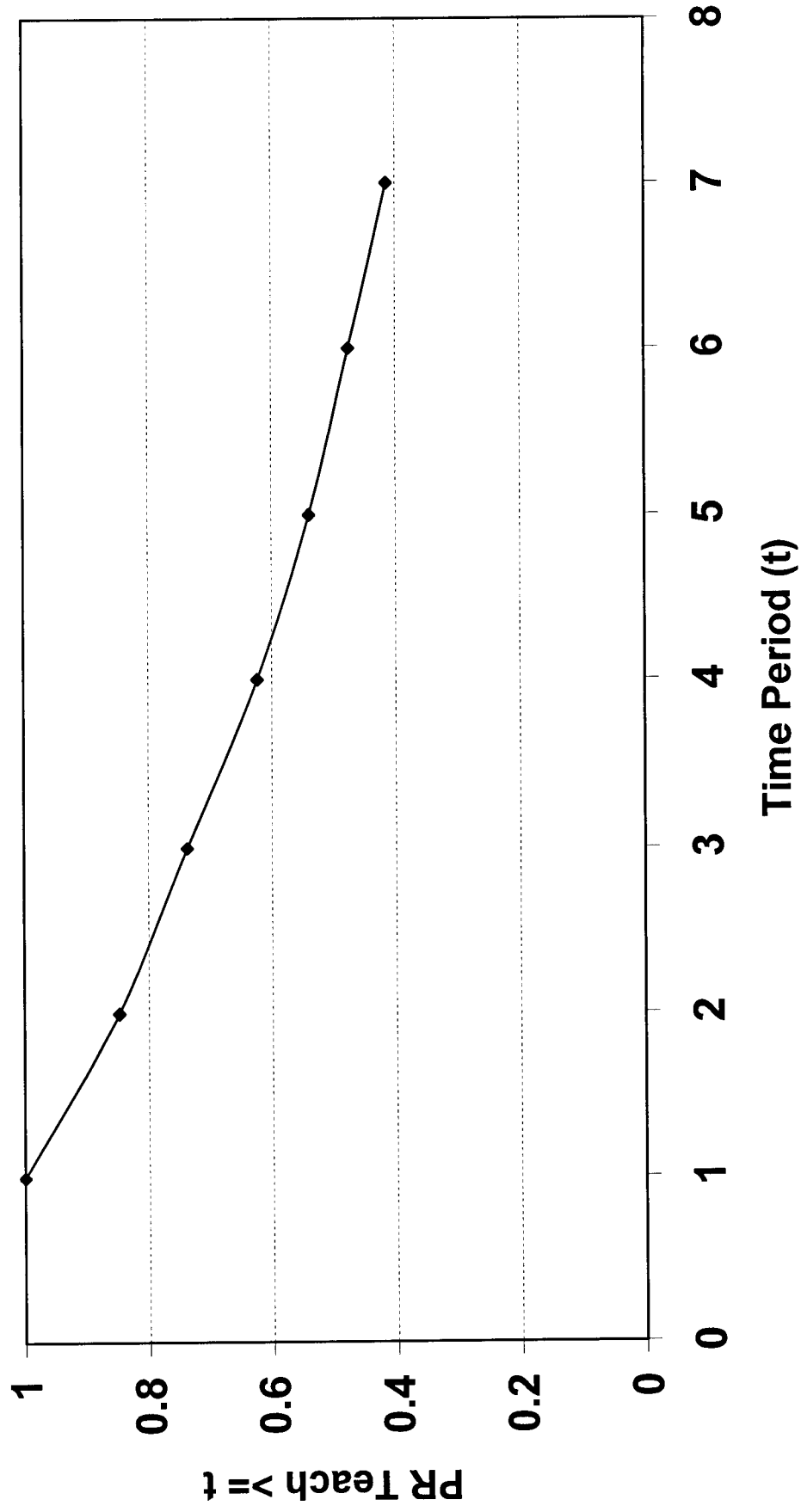
United States House of Representatives

(February 24, 1998)

**Figure 2a. Kaplan-Meier Survivor Functions for  
Elementary Teachers**



**Figure 2b. Kaplan-Meier Survivor Function for  
High School Teachers**



**Table 1**

**Education Sector Wages and Non-Teaching Wages of Former Full-Time Elementary School Teachers**

**Non-Education Sector Wages**

	0	(0-10,000]	(10,000-Minwage(T))	[Minwage(T), infinity)	Marginal Probability
<b>Education Sector Wages</b>	0.4579	0.0751	0.0455	0.0302	0.6087
<b>Non-Teaching Wages</b>	0.1012	0.0359	0.0148	0.0071	0.1590
<b>Teaching Wages</b>	0.1382	0.0332	0.0008	0.0003	0.1725
<b>Non-Teaching Wages</b>	0.0458	0.0137	0.0003	0.0000	0.0598
<b>Marginal Probability</b>	0.7431	0.1579	0.0614	0.0376	1.0000

**Table 2**

**Education Sector Wages and Non-Teaching Wages of Former Full-Time High School School Teachers**

**Non-Education Sector Wages**

	0	(0-10,000]	(10,000-Minwage(T))	[Minwage(T), infinity)	Marginal Probability
<b>Education Sector Wages</b>	0.4378	0.1011	0.0559	0.0447	0.6394
<b>Non-Teaching Wages</b>	0.0883	0.0356	0.0133	0.0096	0.1468
<b>Teaching Wages</b>	0.1165	0.0314	0.0016	0.0000	0.1495
<b>Non-Teaching Wages</b>	0.0489	0.0154	0.0000	0.0000	0.0644
<b>Marginal Probability</b>	0.6915	0.1835	0.0707	0.0543	1.0000

**Table 3****Summary Statistics**

<b>Variable</b>	<b>Elementary Teachers</b>		<b>High School Teachers</b>	
	<b>Mean</b>	<b>Std. Dev.</b>	<b>Mean</b>	<b>Std. Dev.</b>
<b>Log Teaching Wages</b>	10.36	0.154	10.36	0.159
<b>Nonwhite*</b>	0.172	0.377	0.185	0.388
<b>Test Score**</b>	54.83	14.56	68.90	16.57
<b>Poverty***</b>	0.462	0.273	0.366	0.249
<b>Percent Black Students</b>	0.369	0.314	0.378	0.309
<b>Urban</b>	0.771	0.420	0.756	0.430
<b>Close to Atlanta</b>	0.600	0.490	0.562	0.496
<b>Close to City with &gt;75K</b>	0.152	0.368	0.182	0.386
<b>County Unemployment Rate</b>	4.28	1.75	4.40	1.89
<b>N</b>	10,145		4,750	

\* Mean and SD reported for first year of teaching spell only.

\*\* For teachers in elementary schools test score equals the average 3rd grade percentile rank on the Iowa Test of Basic Skills Exam (ITBS Math + ITBS Reading)/2. For high school teachers, the test score is the average overall score on the Georgia High School Graduation Test. For a given school, both test scores could range from 1 to 100.

\*\*\* Poverty equals the proportion of children eligible for free or reduced price lunch.

\*\*\*\* Close to Atlanta equals "1" if school district is within 50 miles of Atlanta. Close to a city with 75,000 residents equals "1" if the school district is within 50 miles of a city that contains at least 75,000 residents, not including Atlanta.

Table 4

## Competing Risks Results for Elementary School Teachers

<u>Risk</u>	<u>Variable</u>	<u>Coefficient</u>	<u>Std. Error</u>	<u>T-Stat</u>
<b>Non-Teaching Job</b>				
	Log Teaching Wage	-0.443	0.502	-0.88
	$\delta N1$	-2.437	1.161	-2.10
	$\delta N2$	-2.668	1.194	-2.23
	$\delta N3$	-2.564	1.213	-2.11
	$\delta N4$	-2.587	1.245	-2.08
	$\delta N5$	-3.171	1.262	-2.51
	$\delta N6$	-2.553	1.322	-1.93
	Test Score	-0.0096	0.0047	-2.03
	Poverty	-1.960	0.363	-5.39
	Percent Black Students	1.732	0.316	5.47
	Nonwhite	-0.232	0.150	-1.55
	Urban	0.439	0.226	1.94
	Close to Atlanta	-0.061	0.215	-0.28
	Close to City with >75K	0.108	0.210	0.52
	County Unemployment Rate	-0.041	0.048	-0.85
<b>Education Sector Job</b>				
	Log Teaching Wage	-0.460	0.256	-1.80
	$\delta E1$	-1.902	0.637	-2.99
	$\delta E2$	-2.530	0.656	-3.86
	$\delta E3$	-2.658	0.674	-3.95
	$\delta E4$	-2.575	0.681	-3.78
	$\delta E5$	-2.434	0.715	-3.41
	$\delta E6$	-2.188	0.754	-2.90
	Test Score	-0.0011	0.0036	-0.31
	Poverty	-0.006	0.238	-0.03
	Percent Black Students	0.086	0.200	0.43
	Nonwhite	-0.225	0.113	-2.00
	Urban	-0.559	0.121	-4.64
	Close to Atlanta	0.148	0.140	1.06
	Close to City with >75K	0.234	0.113	2.07
	County Unemployment Rate	0.082	0.023	3.61
<b>Leave Working in GA</b>				
	Log Teaching Wage	0.042	0.196	0.21
	$\delta L1$	-2.351	0.471	-4.99
	$\delta L2$	-2.357	0.480	-4.91
	$\delta L3$	-2.227	0.489	-4.56
	$\delta L4$	-2.163	0.499	-4.33
	$\delta L5$	-2.353	0.513	-4.59
	$\delta L6$	-2.386	0.531	-4.50
	Test Score	0.0012	0.0020	0.61
	Poverty	-0.849	0.144	-5.89
	Percent Black Students	1.078	0.130	8.32
	Nonwhite	-0.624	0.069	-9.03
	Urban	0.335	0.078	4.27
	Close to Atlanta	-0.245	0.080	-3.06
	Close to City with >75K	-0.060	0.078	-0.77
	County Unemployment Rate	-0.034	0.017	-1.94

Table 5

## Competing Risks Results for High School Teachers

<b>Risk</b>	<b>Variable</b>	<b>Coefficient</b>	<b>Std. Error</b>	<b>T-Stat</b>
<b>Non-Teaching Job</b>	<b>Log Teaching Wage</b>	-1.202	0.538	-2.24
	$\delta N1$	-1.109	1.348	-0.82
	$\delta N2$	-1.317	1.373	-0.96
	$\delta N3$	-0.703	1.396	-0.50
	$\delta N4$	-1.536	1.445	-1.06
	$\delta N5$	-1.184	1.461	-0.81
	$\delta N6$	-2.578	1.702	-1.51
	<b>Test Score</b>	0.0037	0.0068	0.55
	<b>Poverty</b>	-1.176	0.496	-2.37
	<b>Percent Black Students</b>	1.264	0.414	3.05
	<b>Nonwhite</b>	-0.288	0.201	-1.43
	<b>Urban</b>	-0.059	0.254	-0.23
	<b>Close to Atlanta</b>	0.174	0.256	0.68
	<b>Close to City with &gt;75K</b>	-0.140	0.250	-0.56
<b>County Unemployment Rate</b>	-0.016	0.049	-0.33	
<b>Education Sector Job</b>	<b>Log Teaching Wage</b>	-0.680	0.341	-2.00
	$\delta E1$	-1.133	0.860	-1.32
	$\delta E2$	-1.762	0.872	-2.02
	$\delta E3$	-1.765	0.891	-1.98
	$\delta E4$	-1.958	0.914	-2.14
	$\delta E5$	-1.632	0.952	-1.72
	$\delta E6$	-1.514	1.016	-1.49
	<b>Test Score</b>	-0.0051	0.0045	-1.13
	<b>Poverty</b>	0.177	0.331	0.54
	<b>Percent Black Students</b>	-0.056	0.295	-0.19
	<b>Nonwhite</b>	-0.097	0.156	-0.62
	<b>Urban</b>	-0.618	0.170	-3.64
	<b>Close to Atlanta</b>	0.517	0.201	2.57
	<b>Close to City with &gt;75K</b>	0.411	0.158	2.60
<b>County Unemployment Rate</b>	0.061	0.032	1.88	
<b>Leave Working in GA</b>	<b>Log Teaching Wage</b>	-0.046	0.251	-0.18
	$\delta L1$	-2.690	0.613	-4.39
	$\delta L2$	-2.719	0.624	-4.35
	$\delta L3$	-2.611	0.636	-4.10
	$\delta L4$	-2.629	0.647	-4.06
	$\delta L5$	-2.944	0.667	-4.42
	$\delta L6$	-2.916	0.698	-4.18
	<b>Test Score</b>	0.0061	0.0027	2.25
	<b>Poverty</b>	-0.513	0.205	-2.50
	<b>Percent Black Students</b>	0.711	0.197	3.60
	<b>Nonwhite</b>	-0.480	0.096	-5.01
	<b>Urban</b>	0.496	0.107	4.62
	<b>Close to Atlanta</b>	-0.199	0.111	-1.80
	<b>Close to City with &gt;75K</b>	0.100	0.106	0.94
<b>County Unemployment Rate</b>	-0.001	0.023	-0.05	

Table A.1

Education Sector Wages and Non-Teaching Wages  
of Male Former Full-Time Teachers (Elementary and High School)

Non-Education Sector Wages

	0	(0-10,000]	(10,000-Minwage(T))	[Minwage(T), infinity)	Marginal Probability
0	0.3387	0.0995	0.0961	0.0824	0.6167
Education Sector Wages	0.0435	0.0412	0.0183	0.0114	0.1144
(0-10,000]	0.0881	0.0389	0.0057	0.0000	0.1327
(10,000-Minwage(T))	0.1178	0.0172	0.0000	0.0011	0.1361
[Minwage(T), infinity)	0.5881	0.1968	0.1201	0.0950	1.0000
Marginal Probability					



**Table A.2****Summary Statistics for Male Teachers (Elementary and High School)**

<b>Male Teachers</b>		
<u>Variable</u>	<u>Mean</u>	<u>Std. Dev.</u>
Log Teaching Wages	10.39	0.169
Nonwhite*	0.205	0.404
Poverty***	0.396	0.244
Percent Black Students	0.199	0.399
Urban	0.671	0.470
Close to Atlanta	0.506	0.500
Close to City with >75K	0.200	0.400
County Unemployment Rate	4.57	2.01
N	2,652	

\* Mean and SD reported for first year of teaching spell only.

\*\* For teachers in elementary schools test score equals the average 3rd grade percentile rank on the Iowa Test of Basic Skills Exam (ITBS Math + ITBS Reading)/2. For high school teachers, the test score is the average overall score on the Georgia High School Graduation Test. For a given school, both test scores could range from 1 to 100.

\*\*\* Poverty equals the proportion of children eligible for free or reduced price lunch.

\*\*\*\* Close to Atlanta equals "1" if school district is within 50 miles of Atlanta. Close to a city with 75,000 residents equals "1" if the school district is within 50 miles of a city that contains at least 75,000 residents, not including Atlanta.

**Table A.3**

**Competing Risks Results for Male Teachers (Elementary and High School)\***

<b>Risk</b>	<b>Variable</b>	<b>Coefficient</b>	<b>Std. Error</b>	<b>T-Stat</b>
<b>Non-Teaching Job</b>	<b>Log Teaching Wage</b>	-1.428	0.581	-2.46
	<b>δN1</b>	0.336	1.388	0.24
	<b>δN2</b>	0.343	1.425	0.24
	<b>δN3</b>	0.648	1.463	0.44
	<b>δN4</b>	0.064	1.510	0.04
	<b>δN5</b>	-0.251	1.574	-0.16
	<b>Poverty</b>	-0.603	0.562	-1.07
	<b>Percent Black Students</b>	0.929	0.518	1.79
	<b>Nonwhite</b>	-0.419	0.248	-1.69
	<b>Urban</b>	-0.220	0.302	-0.73
	<b>Close to Atlanta</b>	0.200	0.316	0.63
	<b>Close to City with &gt;75K</b>	-0.154	0.282	-0.55
	<b>County Unemployment Rate</b>	-0.106	0.054	-1.97
<b>Education Sector Job</b>	<b>Log Teaching Wage</b>	-0.128	0.430	-0.30
	<b>δE1</b>	-3.293	1.041	-3.16
	<b>δE2</b>	-3.839	1.060	-3.62
	<b>δE3</b>	-3.541	1.108	-3.20
	<b>δE4</b>	-3.111	1.150	-2.70
	<b>δE5</b>	-3.053	1.163	-2.62
	<b>Poverty</b>	-0.675	0.460	-1.47
	<b>Percent Black Students</b>	0.731	0.363	2.01
	<b>Nonwhite</b>	0.319	0.196	1.63
	<b>Urban</b>	0.089	0.233	0.38
	<b>Close to Atlanta</b>	-0.088	0.262	-0.34
	<b>Close to City with &gt;75K</b>	-0.070	0.220	-0.32
	<b>County Unemployment Rate</b>	0.049	0.051	0.97
<b>Leave Working in GA</b>	<b>Log Teaching Wage</b>	-0.730	0.393	-1.86
	<b>δL1</b>	-0.728	0.953	-0.76
	<b>δL2</b>	-1.014	0.970	-1.05
	<b>δL3</b>	-1.069	0.990	-1.08
	<b>δL4</b>	-1.648	1.016	-1.62
	<b>δL5</b>	-2.193	1.075	-2.04
	<b>Poverty</b>	-0.635	0.336	-1.89
	<b>Percent Black Students</b>	1.203	0.282	4.27
	<b>Nonwhite</b>	-0.147	0.140	-1.05
	<b>Urban</b>	0.274	0.163	1.68
	<b>Close to Atlanta</b>	-0.265	0.175	-1.52
	<b>Close to City with &gt;75K</b>	-0.272	0.163	-1.67
	<b>County Unemployment Rate</b>	-0.014	0.034	-0.42

\* Baseline hazards were combined for years 5 and 6 because of small cell sizes. Test scores were not included as an explanatory variable as they are not necessarily comparable between elementary and high schools.

**Figure A.1 Kaplan-Meier Survivor Function for Male Teachers (Elementary and High School)**

