AGING AND HOUSING EQUITY: ANOTHER LOOK

by

Steven F. Venti and David A. Wise

Revised July 2001

Prepared for the conference on the Economics of Aging May 17-20, 2001

We thank the National Institute on Aging and the Hoover Institution for Financial Support.

ABSTRACT

Aside from Social Security and, for some, employer-provided pensions, housing equity is the principle asset of a large fraction of older Americans. Many retired persons have essentially no financial assets to support retirement consumption. We use data from the Health and Retirement Study (HRS), the Asset and Health Dynamics Among the Oldest Old (AHEAD), and the Survey of Income and Program Participation (SIPP) to understand the extent to which families use housing equity to support general consumption in retirement. The analysis is based in part on self-assessed home values reported by survey respondents. Because the self-assessments exaggerate actual home equity, much of the analysis is based on the selling price of recently sold homes, together with the reported equity in recently purchased homes. Homeowners can change home equity by either discontinuing ownership or by purchasing another home of lesser or greater value. We find that in the absence of a precipitating shock--death of a spouse or entry of a family member into a nursing home--families are unlikely to discontinue home ownership. And even when there is a precipitating shock, discontinuing ownership is the exception rather than the rule. On average, families that move and purchase a new home tend to <u>increase</u> home equity. We find, however, that income-poor and house-rich families are more likely to reduce equity when they move, while house-poor and income-rich households are more likely to increase housing equity. Overall, accounting for discontinuing ownership and moving to another home, housing equity increases with age until about age 75 and then declines slightly as households grow older. The overall decline among older households (surveyed in the AHEAD) is about 1.76 percent per year, and this decline is largely accounted for by a 7.84 percent decline among households who experience a precipitating shock. Families that remain intact reduce housing equity very little, about 0.11 percent per year for twoperson households and 1.15 percent per year for one-person households. We conclude that, on average, home equity is typically not liquidated to support general non-housing consumption needs as households age.

Except for Social Security and, for some, employer-provided pension assets, housing equity is the most important asset of a large fraction of older Americans. In principle, these assets might be used to support consumption after retirement. In this paper we take another look at the change in the home equity of older families as they age, beginning at ages just before retirement. We use data from the Health and Retirement Study (HRS), the Asset and Health Dynamics Among the Oldest Old (AHEAD) survey, as well as the Survey of Income and Program Participation (SIPP). We distinguish changes in housing equity that may might be thought of as part of a financial plan to use housing equity as a means of general support in retirement from changes in housing equity that are precipitated by family shocks--death or severe illness.

This paper continues the analysis we began in Venti and Wise [2001], based on the SIPP and on the first three waves of AHEAD. In that analysis, we found that in the absence of changes in household structure, most elderly families are unlikely to move. We found that even among movers, those families that continue to own typically do not reduce home equity. However, precipitating shocks, like the death of a spouse or entry to a nursing home, sometimes lead to liquidation of home equity. Home equity, we concluded, is typically not liquidated to support *general* non-housing consumption needs. The present analysis is an extension of this previous work. Here we use both the HRS and AHEAD data, as well as data from eight panels of the SIPP. Again, the

¹The AHEAD initially surveyed persons age 70 and over in 1993 and resurveyed them in 1995 as part of the second wave of AHEAD and resurveyed them again in 1998 as part of the fourth wave of the HRS. For convenience we refer to these surveys as the first three waves of AHEAD.

key question is whether housing wealth is typically used to support the general consumption of older persons as they age, although the analysis is based on more extensive data. The present analysis also presents a more formal accounting for the change in home equity when ownership is discontinued and the change in home equity from moving to another owned unit ("up-sizing" or "down-sizing"). In addition we give brief consideration to parallel changes in non-housing assets as persons age.

In earlier papers--Venti and Wise [1989, 1990]—we concluded that households "don't want to reduce housing equity" as they age. We found that large reductions in home equity were typically associated with the death of a spouse, retirement, or with other precipitating shocks. These analyses were based on the Retirement History Survey (RHS) and covered persons in the 58 to 73 age range. Merrill [1984], based on the Retirement History Survey (RHS), found that unless there was a change in family status there was little if any reduction in housing equity as families aged Feinstein and McFadden [1989], based on the Panel Survey of Income Dynamics (PSID), including households with heads over age 75, also concluded that in the absence of change in family status housing equity was typically not reduced. Megbolugbe, Sa-Aadu, and Shilling [1997] also used the PSID and found that the change in housing equity varied by age. The oldest households (age 75+) were as likely to trade up as to trade down when they moved. Sheiner and Weil (1993) found some decline in home equity at older ages, but these declines were primarily associated with shocks to family status and health. Hurd [1999], in a general analysis of wealth change based on the first two waves of the AHEAD, concluded that there was a modest decline in housing wealth and rates of home ownership for two-person households that survived the two year period intact, but larger declines for two-person households that lost a member between the waves. He also found that total wealth increased between the waves for all types of households and at all ages.

Whether the elderly perceive home equity as a source of funds for general consumption as they grow older is an important issue for at least two reasons. A concern of some is that older households have substantial wealth locked in illiquid housing and would like to release it. A proposed solution to this perceived "problem" is a reverse annuity mortgage that allows the household to draw down home equity while remaining in the home. To date, there has been little apparent interest in reverse mortgages. It is not clear whether the failure is due to unfavorable financial terms of reverse mortgages or simply to a lack of demand for a product that is intended to exhaust housing equity over the life of the occupant. Several studies, including Venti and Wise [1991], Mayer and Simons [1994], and Merrill, Finkel, and Kutty [1994], have shown that a significant segment of the population appears to be "income-poor and house-rich," and might benefit from a reverse mortgage. We concluded in our earlier analyses, however, that the equity choices of older persons were inconsistent with substantial interest in such products. Nonetheless, knowing whether older households wish to withdraw assets from housing equity helps to evaluate the extent of the potential market for reverse mortgages, and we judge it important to revisit the issue.

A second reason to consider whether the elderly plan to, or will, use home equity to support general consumption is to understand the adequacy of saving for retirement. If housing equity is used just like financial assets to support consumption after retirement, then it might also be considered as a substitute for financial wealth and perhaps treated interchangeably with financial wealth in considering the well-being of the elderly. On the other hand, if households do not plan to be drawn down home equity as they age, it may be more realistic to assume that general consumption expenditures will come largely from accumulated financial wealth, including Social Security and other annuities. Analysts considering how well households are prepared for retirement have treated housing equity in various ways. Moore and Mitchell [2000]

include housing wealth in the set of assets that can be used to finance retirement. The Congressional Budget Office [1993] also includes housing wealth with other wealth. On the other hand, Bernheim [1992] in considering "Is the Baby Boom Generation Preparing Adequately for Retirement" excluded housing wealth in making a determination. Engen and Gale [1999] include zero, 50 percent, and 100 percent of housing equity. Gustman and Steinmeier [1999] conduct analyses using zero and 100 percent of home equity.

In this paper we first consider the relationship between age and housing equity over the life cycle, based on data from the SIPP. This analysis is drawn largely from Venti and Wise [2001]. The results are based on cohort analysis and are presented graphically. Next, we present more detailed cohort analysis for older households, based on the HRS and the AHEAD data. In much of this cohort analysis, the data from these two surveys are combined. Put in many instances the cohort data for older households are shown separately.

We then focus on within household changes in housing equity, giving particular attention to the effect of precipitating shocks. We find that on average there is no reduction in housing equity among persons who continue to own homes, even as they age through their eighties and even into their nineties. Indeed, persons who sell one house and buy another tend to increase housing equity, on average. Large reductions in housing equity are, on average, typically associated only with selling and discontinuing home ownership. Giving up ownership is most often associated with the death of a spouse or entry into a nursing home. In these cases, home equity may be used to pay medical expenses or indeed to support more general consumption of a surviving spouse, although we have not attempted here to document such expenditures. In general, however, we find that home equity is not systematically converted to liquid assets to support non-housing consumption.

Finally, our analysis draws attention to two limiting features of the HRS and AHEAD data. The first feature concerns the use of imputations in analysis of panel data. Our earlier analysis of the AHEAD data was based on preliminary releases of AHEAD wave 2 and HRS wave 4 (the third wave of AHEAD). In the current paper we use more recent releases of the second wave of AHEAD and the fourth wave of the HRS that include asset imputations—including home equity--provided by the HRS staff.² Tabulations from the new data sources are similar, to tabulations presented in Venti and Wise [2001] that did not use these imputations. We find, however, that in many instances the imputations appear to increase the "randomness" in the data. This is perhaps not surprising, given that imputed values are "hot-decked," based on contemporaneous cross-section data. In panel applications, the imputed values should be based on both family-specific longitudinal data, as well as cross-section data. In this paper, all analyses using the "selling price" data (section C.5 forward) drop imputed observations.

A second, related, concern is the large number of inconsistent responses in the *reported* data, particularly when comparing "move" and "stay" transitions to "own" and "rent" housing tenures. For example, many households are reported to own in one wave then rent in the next, and then return to ownership in the third wave, without reporting a move between either the first and second waves, or between the second and third waves. Many of these households begin and end with the same (or similar) home equity. Most of these anomalies are apparently reporting errors. Each such error results in two changes in housing equity that are of equal magnitude but opposite sign and thus may have a large effect on calculated changes in home equity. In some of our analyses we have dropped observations that reported a change in tenure but did not

²The newer data also use additional information on death and nursing home entry that has recently become available.

report a move._We also find many unrealistically large wave-to-wave swings in home equity among households that <u>stay</u> in the same home. These apparent errors are comparable in magnitude to the changes in home equity reported by movers.³

Much of the analysis in this paper is based on recent selling prices and on the reported equity in newly purchased homes. We believe these data are likely to be the most reliable data on home equity. We also have given considerable attention to evaluating the extent of bias in self-assessed home values. Thus on balance, while we believe that more attention can be given to improving the data, we are comfortable with our principle conclusions.

A. COHORT DESCRIPTION

1. SIPP Data on Home Ownership and Equity over the Life Course

The SIPP provides housing equity (obtained from home value and mortgage debt) data for seven years - 1984, 1985, 1987, 1988, 1991, 1993 and 1995.⁴ From the random sample of cross-section data in each of these years we have created cohort data. For example, to trace the home equity of persons who were age 26 in 1984, we

Panel Wave Dates in Field

1984 Sept-Dec 1984 1984 Sept-Dec 1985 1985 Sept-Dec 1985 1985 7 Jan-Apr 1987 1986 Jan-Apr 1987 7 Jan-Apr 1988 1986 Feb-May 1988 1987 4 Feb-May 1991 1990 1991 7 Feb-May 1993 Feb-May 1993 1992 4 Feb-May 1995 1993 7

³The HRS is currently using "call-back" procedures to resolve these issues.

⁴The survey panels and wave that provide the data are as follows:

begin with the average home equity of persons age 26, based on the random sample of persons age 26 in 1984 survey. Next we obtain the average equity of persons age 27 from the 1985 survey, age 29 in the 1987 survey, and so forth. We identify cohorts by their age in the 1984 survey. We do this for 17 cohorts defined by the age of the cohort in the first year of the data. In fact, to obtain more precise estimates of housing equity, the data for a cohort, like age 26, is the average of data for a three-year age interval –25, 26, and 27. We do this for cohorts, age 26, 29, ...to age 71,74. All cohorts are followed until age 80 in the SIPP.⁵

Figure 1 shows the percent of two-person households who own a home, by cohort. These data can be affected by differential mortality. For example, suppose that home owners were less likely to die at any age than renters. In this case, the ownership rate would be increased with age simply because the owners lived and the renters died. To account for this possibility, we made a mortality correction to the data, which is explained in the appendix. The mortality-corrected data for two person households is shown in Figure 1. To make the figure easier to read, only selected cohorts are shown. The key message of the figure is that home ownership does not decline with age, through age 79. In addition, there appear to be no important cohort effects until about age 70. That is, there are no large jumps when the data for one cohort ends and the data for another cohort begins. At older ages, however, there do appear to be noticeable cohort effects. Home ownership is lower for the last two cohorts. But like the trends for the other cohorts, there is no evident decline in ownership as these cohorts age.

Home ownership data for one-person households are shown in Figure 2. Again there is no apparent decline in ownership with age, though age 79. Indeed, the data seem to show some increase in ownership at the oldest ages.

⁵ Data for households over age 80 are not used because age is top coded at 80.

Cohort home equity data for two-person families are shown in Figure 3. These data in 1995 dollars and are corrected for mortality. The within-cohort data show no decline in home equity as the cohort ages. The data may even show some increase in equity within cohorts for ages 65 to 79. There do appear to be some cohort effects in equity, as evidenced by the jumps when the data for one cohort ends and the data for another cohort begins.

In estimates reported in Venti and Wise [2001] we show rather systematic cohort effects. The estimates show that both older cohorts—those over age 70 in 1984--and younger cohorts—those younger than 36 in 1984--have lower home equity than the average, while the middle-aged cohorts have higher equity than the average. The cohort effects are likely determined in large part by differences in housing price changes over time.⁶

Figure 4 shows the cohort equity data for one-person households, corrected for mortality and inflation. As with the two-person households, there seems to be no decline in equity through age 79.

2. AT OLDER AGES: HRS and AHEAD

To understand trends in home equity at older ages, we use the AHEAD as well as the HRS. Both are panel studies. The HRS follows persons in households with heads age 51 to 61 in 1992. Members of these households were interviewed in 1992 and again in 1994, 1996, and 1998. In 1998, the heads were age 57 to 67. Thus this age range is included within the SIPP ages. The AHEAD study follows persons in

⁶For example, assume that homes are bought at age 35 on average, and consider the cohort that was age 50 in 1984 compared to the cohort that was age 38 in 1984. The older cohort bought homes in 1969 on average and would have gained from large home price increases in the 1970s. On the other hand, the younger cohort would have bought homes in 1981 on average and would have seen much lower increases in home equity during the 1980s and 1990s.

households with heads age 70 and older in 1993. These households were interviewed in 1993 and again in 1995 and in 1998 (as part of the fourth wave of the HRS.⁷ The AHEAD age range overlaps the older SIPP ages. Thus both HRS and AHEAD allow comparison with components of the longer life cycle SIPP data. Details of the survey design are presented in Juster and Suzman [1995].

In this analysis, we follow households in both the AHEAD and HRS files. One complication is tracking households over time. A household may split through divorce or separation, members may die, or a family member may enter a nursing home. For the purposes of this analysis, we have adopted these conventions: In the first wave of each survey households are identified as either one-person or two-person households (institutionalized persons are excluded from the original sample). In subsequent survey waves we classify each household--according to the change since the prior wave--into one of the following six "states":

- "1" Continuing one-person household
- "2" Continuing two-person household
- "D" One of the original members has died
- "T" Both of the original members have died
- "N" One or more members has entered a nursing home
- "S" Household composition has changed for some other reason (most often a split through divorce or separation or the addition of a new adult member.)
- "0" Household refused the interview or is missing for other reasons

The sequences observed in the HRS and AHEAD are presented in Tables 1.

These sequences are used to distinguish households included in analyses below. In cohort analysis in the next section we restrict attention to continuing two-person or one-person households identified as "2222" or "1111" for the HRS and "222" or "111" for the

⁷Juster and Suzman [1995] provide details of the survey design.

AHEAD. In the following section we consider changes in housing equity and other assets between waves. For this analysis we use each two-period sequence (creating an "interval"), and we focus in particular on the within household relationship between home ownership and home equity on the one hand and change in household composition on the other hand. We consider cohort data on home ownership first. Then we consider cohort data on home equity, as well as non-housing net assets.

a. Home Ownership

To obtain cohort data comparable to the SIPP cohort data, we construct cohorts from the HRS and AHEAD data by grouping households in two-year age intervals.

These constructed cohorts are the basis for the cohort data shown below.

The home ownership cohort data for two-person families are shown in Figure 5a, which covers ages from 50 to 93. To make the individual cohort data easier to view, only selected–largely non-overlapping–cohorts are shown. Overall, the within-cohort data show an increase in home ownership through age 70. Thereafter the cohort data suggest a small decline in ownership. Cohort data for the older AHEAD households are shown separately in Figure 5b. A more detailed analysis of these data, presented below, shows that for the AHEAD sample the within-cohort decline in ownership for continuing two-person households is about 0.66 percent per year for cohorts age 70 to 78 in the initial year and 0.34 percent for cohorts age 80 or more in the initial year. A comparison of these data with the SIPP data in Figure 1 shows that for persons age 50 to 79 the SIPP and the HRS-AHEAD data are very similar. Both data sources show ownership rates of about 90 percent for families over age 60. The within-cohort SIPP data, however, show no decline in ownership through age 79.

The pattern of home ownership for continuing one-person households, shown in Figure 6a, is quite different. The <u>within-cohort</u> data for one person households show a

distinct rise in ownership between ages 50 and 75 and a decline in ownership at older ages. The cohort data for the AHEAD households is shown separately in Figure 6b. The <u>within-cohort</u> decline for the continuing one-person AHEAD households is about 0.66 percent per year for cohorts age 70 to 78 in the initial year and 1.82 percent for cohorts age 80 or more in the initial year. (Some of the decline in these cohort profiles is apparently the result of reporting errors. The within-cohort profiles for the AHEAD households in Figures 6a and 6b become flat when households who report a change in ownership status, but do not report having moved, are deleted from the same, as discussed in section C.2 below.) The HRS-AHEAD data for the 50 to 75 age group are comparable to the SIPP data for this group.

b. Home Equity

Mean home equity cohort data for two-person households are shown in Figure 7a.8 These within cohort data show an increase in home equity through about age 70 or 75. The data for the AHEAD households are shown separately in Figure 7b. At older ages, the randomness in the within cohort trends make it hard to see clear trends, although there appears to be a within cohort decline in equity. In fact, data presented below show that the average mean decline is about \$2,100 per year, which is largely accounted for by the reported decline the same-home equity of continuing owners.

The home equity cohort data for one-person households are shown in Figure 8a, and separately for the AHEAD households in Figure 8b. As with the two-person households, there is a clear within-cohort increase in home equity through age 70 or 75. At older ages a consistent within-cohort trend is not apparent. Data presented below show that the average decline is about \$3,000 per year, again, largely accounted for by

⁸All dollar amounts for the SIPP and AHEAD have been converted to 1998 dollars using the CPI.

the reported decline the same-home equity of continuing owners. There appear to be substantial differences in home equity by cohort, although the randomness in the data make it hard to distinguish cohort effects from within-cohort changes in home equity.

Median cohort data for two- and one-person households are shown in Figures 9 and 10 respectively. There is less randomness in the median data than in the mean data and thus within cohort trends are easier to discern in there figures. For example, for older two-person households the medians suggest modest within cohort decline in home equity beginning at about age 75, but cohort effects are not apparent. On the other hand, the median cohort data for older one-person households show little within-cohort decline in home equity but rather substantial cohort effects. Older cohorts seem to have successively less home equity. Below, we present quantitative estimates of the within-cohort changes in home equity.

c. Non-Home Equity

In considering the equity value of housing as these cohorts aged, it is informative to compare the value of housing with other assets. Cohort data on non-housing assets are shown in Figures 11 through 14. Like the home equity data, mean and median cohort data are shown for two- and one-person households. And separate figures are shown for the older AHEAD households. As with the home equity data, the trends in the non-home equity data for the HRS households is quite clear. But the enormous randomness in the data make the cohort data for the AHEAD households much harder to interpret. Nonetheless, some trends are clear form the cohort data. (Below we show quantitative within-cohort changes in non-home assets, as well as home equity.)

First, it is clear for the HRS households that both home equity and housing increased with age, but the non-housing assets increased much more. For example, from Figure 7a it can be seen that the mean home equity of continuing two-person

households increased from about \$80,000 at age 50 to about \$120,000 for households in their early 70s. There seem to be no apparent cohort effects. In Figure 11a, it can be seen that non-housing assets of the HRS households increased from about \$200,000 at age 50 to close to \$400,000 at age 74, about five times as much as the increase in home equity. Again, cohort effects are not apparent in this age range. In future analysis we will try to determine which components of non-equity assets account for the large increase.

Second, for the older HRS households there are also large within-cohort increases in non-equity assets. For the older households, however, there are also large cohort effects, with successively older cohorts having lower non-housing assets. And, for the older cohorts there is some within-cohort decline in home equity.

It may be that there are in fact very large wave to wave changes in both home equity and non-housing assets. We believe, however, that the data is likely to reflect substantial errors. Thus further "cleaning" of the data might result in more consistent cohort patterns. The further cleaning would have to be based on joint evaluation of all assets over all waves of the HRS and AHEAD surveys—looking perhaps at a X x Y matrix of data for each household. In principle, the cohort data could be parameterized—with assets related to a flexible function of age, with cohort effects, for example. But at this point, we have not found a parameterization f the non-equity data that adequately "fits" the cohort patterns.

C. FAMILY STATUS AND HOME EQUITY: HRS and AHEAD

We now turn to the relationship between changes in home equity and changes in family structure. Again we consider two- and one-person households separately and provide separate estimates for the HRS and the AHEAD families. Before considering within-cohort household transitions, cross-section summary data on household tenure

(own, or rent or other combined) are shown by age and household structure (one-person or two-person) in Table 2. Home ownership of two-person families exceeds 90 percent between ages 54 and 74 and then declines to around 80 percent at ages 85 and older. For one-person families, home ownership increases to about 68 percent for age 70 to 74 households and then declines to about 50 percent for households age 85 and older. The home ownership rate for one-person households peaks in the 70-74 age range, declines modestly over the next decade, then falls sharply after age.

1. Within-Household Transitions

We focus on the events that precipitate changes in home ownership and the changes in home equity that are associated with the ownership changes. Table 3 shows ownership transitions between consecutive survey waves (an "interval"). The first two panels of the table pertain to households that owned a home at the beginning of the interval. The third and fourth panels pertain to households that did not own a home at the beginning of the interval. The table entries show the percent of households who make a transition between adjacent waves of each survey. For example, the transition labeled "22" identifies two-person household at the beginning and at the at the beginning of the interval (the first of the two waves) and at the end of the period (in the subsequent wave). The HRS yields as many as three transitions (wave1 to wave 2, wave 2 to wave 3, and wave 3 to wave 4) and each represents a two year interval. The AHEAD yields two transitions. The first interval is two years and the second three years. All intervals in the HRS are combined to obtain the HRS results, and all interval in the AHEAD are combined to obtain the AHEAD results.

Consider first the top panel of the table which pertains to the HRS households who were homeowners at the beginning of an interval. The first column shows the percent of households that own and the percent that rent (or have some other living

arrangement) at the end of the interval. Of continuing two persons households, 98.3 percent still owned at the end of the interval; 1.7 percent no longer owned. The ownership of initial owners declined about 0.85 percent per year. Now consider continuing two-person HRS households who were non-owners at the beginning of the period shown in the third panel of the. Of these households 22.3 percent became owners during the interval, about 11.1 percent per year. On balance the number of homeowners increased: some initial owners became non-owners, but a larger number of initial non-owners became owners. This net addition to the homeowner group is shown graphically for the younger--HRS-cohorts in Figure 5a. The figure, however, pertains to households who continued as two-person families through all four waves of the HRS. The data for continuing two-person households in the table, however, is based on all households that continued as two person families during any two adjacent survey waves.

Other rows of the first panel of Table 3 show that if a spouse dies (2D), the ownership rate remains high, at 95.6 percent. If a spouse enters a nursing home (2N) the ownership rate declines more, to 88.6 percent, although the sample of nursing home entrants is quite small for the younger HRS households.. For continuing one-person HRS households the ownership rate also remains high, at 95.2 percent. (There are only three single-person households in which the person entered a nursing home during the interval.)

The percent moving between adjacent waves is shown in the next column of Table 3. Of two-person households that own in both waves, 7.1 percent moved over the two-year interval. For two-person households that change from own to rent-or-other, the move rate is an unexpectedly low 65.7 percent. It is possible that ownership is transferred from parents to children, so the parents do not move, but also no longer own. However, this low move rate is more likely a reflection of reporting error.

Inspection of some of these cases shows households owning a house of roughly constant value for three of the four waves. This evidence, combined with the absence of a move (which is verified by survey-takers), suggests errors in reporting or coding for one of the waves. Because there are a relatively small number of these households, a few errors can have a substantial effect on the move rate.

Similar results for the AHEAD sample are presented in the second and fourth panels. Initial homeowners in AHEAD were also likely to remain owners unless there was a change in family status. For example, 96.9 percent of continuing two-person households continued to own. But if one of the members died the ownership rate dropped to 88.8 percent. If one of the members entered a nursing home the rate dropped to 75 percent. For continuing one-person households, 91.3 percent remain owners. But if the single person enters a nursing home, the ownership rate drops to 39.9 percent. Thus, as with the younger HRS households, in the absence on precipitating shock most AHEAD homeowners continue to own. But in the event of a shock, the decline in ownership is greater for older than for younger households. In addition, the decline is greater for one-person than for two-person households.

The move rate for the older AHEAD households that own in both waves is quite low, about 3.9 percent for two-person households and 4.5 percent of one-person households. Since the interval between waves is about 2 ½ years for the AHEAD, the annual move rates are 1.6 percent and 1.8 percent respectively.

Overall, Table 3 suggests that homeowner households in the HRS age group are very likely to remain owners. And even if one of the household members dies or enters a nursing home, the rate of ownership remains high. Homeowners in the AHEAD age group are also likely to continue to own unless there is a change in family status, especially continuing two-person households. When a member of this older household dies or enters a nursing home, the decline in ownership is greater than for younger

households. The greatest decline in ownership is for single-person AHEAD households who enter a nursing home. Even among this group almost 40 percent continue to own.

2. Change in Home Equity

We consider changes in home equity that parallel the transitions shown in Table 3. Home equity changes are presented in two formats. The first format shows changes for all households— initial owners and initial renters-others. It shows changes for households who switch form owning to renting, as well as those switching from renting to owning. And it shows the net change in home equity for both groups combined. The second format is directed to the primary focus of our analysis, the change in home equity for initial homeowners. In this format we give particular attention to the change in the equity of movers who continue to own, compared to stayers, those who remain in the same house. Although we discuss changes based on changes in self-assessed home values here, we show below that the exaggeration of self-assessed home value impart large bias to the implied changes in home equity. Then we consider changes based on home selling prices compared to reported equity in newly purchased homes. We believe these latter data are the most reliable, as discussed below.

In addition, the mover-stayer comparison is complicated by the data inconsistencies discussed in the previous section. Some households report a change in tenure without moving. While such changes are possible, we believe most such cases reflect reporting or coding errors. The information on whether a household moved since the previous wave is likely to be accurate because the prior address is incorporated in the survey question on moving.⁹ In all calculations reported below, we deleting all

⁹For example, in wave 4 of the HRS (also wave 3 of the AHEAD) noninstitutionalized respondents were asks "Are you still living, all of the year or part of the year, in the same apartment/house in previous wave address and city>?"
Respondents in nursing homes were asked "Do you still have the same apartment/house in previous wave address and city>?" If respondents in nursing

observations with apparent transitions involving a change in tenure without a reported move. Following this procedure, 1.1 percent of the HRS households and 3.4 percent of the AHEAD households are deleted.¹⁰

Change in home equity using the first format is presented in Table 4. The family status designations are the same as those used in Table 3. There are four tenure designations: OO, OR, RO, and RR where "O" indicates own and "R" indicates rent or other living arrangement. Large reductions in home equity are typically associated only with a home sale and subsequent rental. Those who move from renting to owning, of course, increase home equity. No matter what the change in family status, there is an increase in the average equity of HRS households (with the exception of the few 1N families). On the other hand, there is a decrease in the mean home equity of AHEAD families, no matter what the change in family status. The greatest decrease occurred when a family member entered a nursing home. For all continuing two-person households, the mean increase in housing equity was \$5,937 in the HRS and -\$4,817 in the AHEAD. The median increase was zero for households in each of the surveys. In general, the median changes are smaller in absolute value than the mean changes, but the relative patterns by family status and change in tenure are similar.

Change in home equity using the second format is shown in Table 5. The key question here is whether continuing homeowners who move and buy another house reduce home equity more than stayers, who can serve as the "control group" in this comparison. If movers typically wanted to use some of the wealth accumulated in home equity to support other non-housing consumption, the home equity of movers would be

homes answered affirmatively, they are may still be homeowners and they are not classified as movers.

¹⁰Deleting all respondents who change tenure without moving reduces the frequency of own to rent transitions. This affects the HRS and AHEAD cohort figures presented above. In particular, the cohort profiles for one-person AHEAD households (Figures 6a and 6b) become flat.

reduced relative to the change in the equity of stayers. The first two panels of Table 5 show the mean change in housing equity for the HRS and AHEAD; the next two panels show medians. Consider the upper panel of the table, which pertains to two-person households in the HRS. The change in family status is shown on the left margin. The ownership status (tenure) at the end of the interval is shown along the top margin. A household can continue to own or become a renter (or have some other living arrangement) at the end of the interval. The change in home equity is shown for continuing owners, for renters-others, and for both groups combined (all). The initial home value for each group is shown in the right column of the table. The mean home equity of continuing two-person households increased by \$5,855, on average. For those who remained home owners, equity increased by \$6,569. Initial homeowners whose transition was to the rent-other group reduced home equity by \$54,155 on average. The average initial home value of continuing two-person households was \$102,310. Thus home equity of the home sellers was only about half of the average equity of all continuing two-person households.

Some of those who continued to own stayed in the same house, others moved and bought a new house. The equity of those who stayed increased by \$6,686. The equity of those who moved and bought a new house also increased, by \$5,074. In somewhat more formal estimation below we use the change in the equity of the stayers as a measure of the increase the movers would have experienced had they not moved. In this case the decrease for movers was \$1,612, about 1.7 percent of the initial home equity of this group. Thus these movers who bought a new home are not typically taking substantial home equity out of housing to support other consumption. By this measure, the greatest decline in home equity occurred in mover households in which a member died, although the sample sizes are small and the means are not precisely measured. For example, the home equity of the small number of two-person households who move

but continue to own when one member dies declines by \$21,935.

The average equity of continuing one-person HRS households declined by \$697, a very small fraction of the \$95,555 average initial home equity. Continuing one-person households who moved but continued to own reduced home equity by \$3,739, and the stayers increased equity by \$935. Using the stayers as a control, the movers reduced equity by 4.8 percent of the initial home equity of this group.

In summary: the average home equity of two-person HRS households increased over this period. This was true for continuing two-person households as well as those in which a member died or in which a member entered a nursing home. The equity of one-person households declined only slightly. Continuing owners who moved typically reduced home equity only marginally, when compared to stayers.. The only substantial reduction in the home equity of continuing owners was for households in which one member died.

For the older AHEAD households, changes in home equity also are typically associated with precipitating shocks. But for the older households the shocks are more frequent. Consider continuing two-person households first. The equity of continuing stayer owners (who do not move) declined by \$4,103 and can serve as a base of comparison for other groups. This reduction, if taken at face value, apparently reflects a fall in the value of the homes of the older households as they continue to live in the homes, but not direct withdrawal of housing equity to support other consumption. This decline is only slightly less than the average reduction for all continuing two-person households, \$5,367. Thus on average we conclude that little housing equity is taken from housing to support other consumption.

Continuing homeowners who move reduce home equity by \$15,877, which is \$11,322 more than the reduction in home equity of the stayers. We take this to represent funds taken from housing and that might be used to support other non-

housing consumption. It represents, however, only about 10.5 percent of initial home equity for these households, and less than 4 percent of their initial non-housing wealth. Remember that the typical older household will only move once from one home to another. So if the reduction in housing equity can only be a one-time addition to funds available for other consumption. Below we show that even this small reduction is probably exaggerated and that in fact the average change is likely positive (an increase in housing equity).

For continuing owners in two-person households in which a member dies or enters a nursing home, the reduction in the home equity of the movers is \$14,655 greater than the reduction for the stayers. The reduction in the home equity of continuing one-person households is also small. In particular movers who continue to own reduce home equity by a small fraction of initial home equity.

In summary: even among the older AHEAD households the reduction in home equity of continuing owners is small relative to initial home equity, even among those who move to a different house. Large reductions in home equity are typically observed only for home owners who move and discontinue home ownership. The probability of such a move is larger in cases of precipitating shocks. But as seen in Tables 3 and 4, even in the event of shocks to family status, most households continue to own and thus do not withdraw equity from housing to support other needs. For all HRS groups, the initial home equity of the seller (rent-other) group was much lower than the equity of the continuing owners. For the older AHEAD households the initial home equity of sellers is also less than the initial home equity of continuing owners, although the difference is much smaller than for the HRS households.

Median changes in home equity are shown in bottom half of Table 5. The pattern of change is essentially the same as the pattern for mean changes. The changes, however, are typically smaller than the mean changes, in particular for the older AHEAD

households. For example, for continuing two-person households in the HRS the median increase in home equity is \$1,474. The increase for continuing owner-movers is only \$2,105 greater than for stayers. For continuing one-person families the median increase is \$222. And the reduction for continuing owner-movers is only \$1,028 greater than for stayers. Again, the conclusion is that for the most part housing equity is substantially reduced only after a precipitating shock, and even in these instances, withdrawal from home ownership is the exception and not the rule.

3. Respondent Estimates of Home Values versus Sales Prices

Before turning to some simple estimation, we emphasize that respondent assessment of home equity likely overestimates home value by a substantial margin. Thus reliance on reported home values yields exaggerated reductions in housing equity when homeowners move. Substantial evidence shows that homeowners overestimate the value of their homes. Kiel and Zabel [1999]) surveyed the literature and concluded that self-reported home values exceed actual sale prices or appraisal values by -2 to 16 percent. Their own analysis showed that homeowners on average overvalue their home by 8 percent, and that owners with long tenure overvalue their houses even more. In other words, when a family moves the realized sale price is typically less than the family's prior estimate of the home value. This creates a bias in our estimate of the change in housing equity among movers. The pre-move estimate is inflated. The post-move price is presumably "accurate" because the purchase transaction was recently completed.

The estimates in Tables 4 and 5 on the change in housing equity between waves are based on HRS and AHEAD respondent self-assessment of home values and are affected by such overvaluation. The tendency to overvalue homes confounds mover-stayer comparisons. Recent movers are likely to know the market value of their homes. Stayers, on the other hand, are likely to overvalue their houses. As a result, the change

in home equity is more likely show a larger price decrease for movers than for stayers. Thus in the previous tables movers, relative to stayers, appear to be taking more equity out of their homes than is actually the case.

Information obtained in both the HRS and the AHEAD allows us to gauge the extent of this bias. For households that have recently moved, the surveys inquired about the sale price of the house. The sale price can be compared to the reported value of the house in the previous wave. The survey also asks for the month and year of the sale; the month and year of the self-assessed value is the interview date. We index the pre-move assessed value of movers and the post-move price of movers to obtain measures in 1998 dollars.¹¹ From these values we obtain estimates of the overvaluation bias.

Mean and median differences between assessed values and sales prices are shown in the Table 6. The results suggest that both the HRS and the AHEAD respondents overestimated their home values by 15 to 20 percent, based on a comparison of mean values. Based on medians, home values are overestimated by 6 to 7 percent. The mean dollar differences are \$20,000 to \$30,000, and median dollar differences are \$6,000 to \$8,000. This suggests that our calculated reductions in the home equity of continuing owner-movers may be due entirely to valuation bias. For example the mean reduction of \$15,887 (or \$11,322 using the stayers as a control) in the home equity of two-person AHEAD families who move and continue to own would be more than accounted for by such bias.

4. More Formal Estimates of Change in Home Equity

¹¹Some movers are missing data for the sale price. The HRS and AHEAD provide no imputations for missing values of the sale price. A bracketing technique is used to obtain ranges for persons unable to provide a sale price, but we have made no attempt here to convert the bracketed amounts to values. The analysis is restricted to observations that specify a sale price.

Here we consider more formally the change in home equity of movers and stayers. As mentioned above, one way to think about this is to treat movers as the treatment group and stayers as the "control" group. The home equity of stayers and movers at the beginning and at the end of the interval can be represented by:

	Beginning	End
Stayers	п	" + t
Movers	п	" + t + m

In this case, a difference-in-difference estimate yields m, the "treatment" effect. We can estimate this for all households combined, or for any subgroup, by

$$(1) \Delta E = t + mM$$

where t is a constant term--and represents a time (inflation) effect--and m is the additional effect for movers, with M a dummy variable identifying movers.

Estimates of this equation, by change in household status, are shown in Table 7. This table presents estimates for households who owned at both the beginning and at the end of the interval. Data are presented by the subsequent—at the end of the interval—status of the initial homeowners. OLS estimates are shown in the left portion of the table. Median regression estimates are shown in the right portion of the table. The median regression estimates should be less affected than the OLS estimates by reporting errors or other outliers in the data.

The key mover effect estimate, m, measures the difference between the change in the equity of stayers and the change for movers. The OLS estimates show negative mover effects in each comparison, but only the mover effects for the HRS 21 and AHEAD 11 groups are significantly different from zero at the 5 percent significance level. And, with the possible exception of the estimated mover effect for the 2 to 1 HRS households, the estimated effect is much lower than the bias suggested in Table 6. For

example, the estimated mover effect for continuing two-person households is -\$1,612. Referring back to Table 6, however, we see that the bias estimate for HRS households is between \$20,000 and \$33,000. Thus, since most families are continuing two person families, a reasonable judgment from these data is that the equity of the continuing two-person households in fact <u>increased</u> by about \$25,000. Coincidentally, this increase matches the estimated increased for such households based on selling prices, which is discussed below. For each of the other groups, with the exception of the 2 to 1 HRS families, the estimated mover effect is much less than the bias estimates shown in Table 6, suggesting rather large increases in home equity.

For the HRS households, the median regression mover effect estimates are also small and typically not significantly different from zero. And, the estimates are less then the median bias estimates Table 6. Thus, based on the estimated mover effects in conjunction with the bias estimates, we conclude that home equity likely increases substantially when families move and buy another home.

The median estimates for the AHEAD households are larger than the median HRS estimates and are more precisely measured. For the 21 and 2N groups, the estimates are greater than the bias estimates in Table 6, in particular for the 2N group. Thus these data suggest that for households in which a member dies, and for households in which a member enters a nursing home, home equity is reduced when these households move and buy again. The analysis below based on selling prices, however, suggests an increase in the median home equity of these groups as well.

5. Estimates Based on Selling Price

Each home owner re-interviewed in the HRS and AHEAD is asked whether the home was sold since the previous interview. For many of these households, the selling

price is reported.¹² In this section, we estimate the change in the home equity of families who sell and buy another home, and the change in equity of those who sell and then choose another tenure. Table 8 shows summary data on home equity for adjacent waves of HRS and AHEAD. The first column shows reported home equity from the first of the two waves. The second column shows the reported selling price (obtained from the second wave interview) minus the mortgage reported in the initial wave. The sale occurred sometime between the two waves, but the mortgage pertains to the data of the last interview prior to the sale. The third column shows home equity reported in the second of the two waves. For households who purchased another home (the first and third panels of the table), this is the equity in the newly purchase home. For households that did not purchase another home (the second and fourth panels), this column is zero.

Like the data in Table 6 on reported home values versus selling prices, these data show that households who sell and buy another home substantially overestimate their pre-sale housing equity. For those who sell and do not purchase another home, the over-estimation is not so apparent. For several of these groups the reported equity seems to underestimate realized equity, based on selling price minus the mortgage. We believe that the reported selling price is likely to be close to the actual selling price, unlike the pre-sale assessment of home equity. The last column shows reported home equity at the end of the interval. In principle, home equity right after a purchase should also be accurately reported. For each of the intervals, the reported new home equity at the end of the period is substantially greater than gain in home equity from the sale of the prior home, suggesting that equity in the new home is greater than equity in the prior

¹²There is more missing sale price data than home equity data, used in earlier sections of the paper. Home equity (home value and mortgage balance) is obtained from the Housing module. Information on the sale price is obtained from a module on Capital Gains that has more incomplete responses. There are no imputations for missing or incomplete (bracketed) sale price data. Partly for this reason, we do not use the weights when analyzing the sale price data.

home.

Based on the same data, Table 9 shows the estimated change in home equity for households that have sold a home and purchased another, by change in family status. These estimates are obtained from simple OLS and median regression estimates of the form

$$\Delta E = m + \varepsilon$$

where) E is equity in the new home at the end of the period minus equity from the sale of the prior home. Here, m is the estimated increase in home equity. This specification is estimated for several years separately and for several family status change groups. For all but two groups, there is a substantial increase in home equity. Many of the estimates are for small groups, however, and are not significantly different form zero.

We now consider whether the change in home equity depends on the relationship between income and housing wealth. It might be expected that persons with relatively low income and relatively high housing equity would be more likely to withdraw housing equity. And those with low equity and high income would be more likely to add to housing equity. We begin with estimates of the probability of moving and buying another home, and the probability of moving and discontinuing home ownership, thus withdrawing all housing equity. These outcomes will depend, in particular, on the level of home equity and the level of income in the initial period. Then we show estimates of the relationship between the change in equity, given a move, on the one hand, and initial income and home equity levels on the other hand.

Households that own in the initial period can either stay in the same house, move to another house, or discontinue home ownership by moving to a rental apartment or some other arrangement. The probabilities of the latter two transitions may be specified as:

$$Pr[OmO] = c(21 \text{ or } 2N \text{ or } 1N) + a11 + b22 + \alpha Y + \beta E + \gamma Y \cdot E + \varepsilon$$

$$(4)$$

$$Pr[OR] = c(21 \text{ or } 2N \text{ or } 1N) + a11 + b22 + \alpha Y + \beta E + \gamma Y \cdot E + \varepsilon$$

where OmO identifies families who sell a home, then move and buy another home (own to move to own) and OR identifies families who discontinue ownership (own to rent or other). The parameter "a" is the effect of a continuing one persons household and "b" the effect of a continuing two-person household. (The estimated parameters are of course not constrained to be the same for the OmO and OR groups.) The omitted categories, captured in the constant tern c(21, 2N, and 1N), are the 21, 2N, and 1N households. Initial period income is denoted by Y and initial home equity is denoted by E. . Here, (indicates whether the effect of Y depends on E (or, equivalently, whether the effect of E depends on Y).

Given the decision to move to another home or to discontinue ownership, we then estimate the conditional change in home equity for the two groups, given that a move occurs. The change in equity equations are in the same format, given by

$$\Delta E(OmO) = c(21 \text{ or } 2N \text{ or } 1N) + a11 + b22 + \alpha Y + \beta E + \gamma Y \cdot E + \varepsilon$$

(5)
$$\Delta E(OR) = c(21 \text{ or } 2N \text{ or } 1N) + a11 + b22 + \alpha Y + \beta E + \gamma Y \cdot E + \varepsilon$$

Given the estimated probabilities and conditional changes in housing equity, we can simulate the expected change in equity for homeowners as

$$\Delta E = \Delta E(OmO) + \Delta E(OR) =$$

$$\Pr[OmO] * E(\Delta E | OmO) +$$

$$\Pr[OR] * E(\Delta E | OR)$$

where the expected change in decomposed into it's component parts. We present below the simulation for selected quantiles on income and home equity.

Estimates of the probability of a move are shown in Table 10. Estimates for HRS households are shown on the first page of the table, estimates for AHEAD households on the second page. Parameter estimates are shown at the top of each page. The lower part of each page shows simulated mover probabilities at selected income and equity quantiles. The simulations show that initial income and home equity have little effect on the probabilities of moving, either to another home or to some other arrangement, although income has a statistically significant positive effect on the probability of moving and buying again in both the HRS and the AHEAD samples. Both income and home equity are statistically significant and positively related to the probability of discontinuing ownership for HRS households. But neither is statistically significant in the estimates for AHEAD households. In most of the simulations the difference in probabilities between "house-poor and income-rich" households and "house-rich and income-poor" households is only a few percentage points.

Estimates of the change in home equity between the survey waves for families who move and buy anther home are shown in Table 11. These estimates are based on the sample of respondents that report a sale price for the former home and report both the home value and mortgage debt for their current home.¹³ Both OLS and median

¹³Both the sale price of the old home and the value of and mortgage on the new home are reported in the same wave. The survey does not inquire about the mortgage

regression estimates are shown. The top half of the table shows results for HRS households and the bottom half for AHEAD households. In none of the estimates is there a statistically significant difference in the change for the 11 or for the 22 groups compared to the 21-2N-1N groups combined. However, there are substantial and significant effects for initial income and home equity. The greater the level of initial home equity (based on selling price minus the mortgage), the smaller the increase in equity when the family moves. And the larger initial income, the greater the increase in home equity for households that move. None of the equity-income interaction terms is significantly different from zero. The table also shows simulations of the change in home equity at the 20th, the 50th, and the 80th quantiles of equity and income.

Evaluated at the median (50th quantile) of income and home equity, the simulated change in equity is positive for all family status groups, with the exception of the simulation for the 11 group based on median regression estimates. For all family status groups the greatest simulated reduction in home equity is at the 80th equity quartile and 20th income quantile. The greatest simulated increase in home equity at the 80th income quartile and the 20th equity quantile. Thus relatively house-rich and income-poor families reduce equity and relative house-poor and income-rich households add to home equity when they move and buy another home. For example, based on the OLS estimates for the 22 HRS households, at the high-equity-low-income quantiles home equity is reduced by -\$15,422; at the low-equity-high-income quantiles home equity is increased by +\$54,778. The pattern of the simulated changes based on the median regression estimates is similar to the pattern based on OLS estimates.

The change (decrease) in the home equity of the families who discontinue home ownership is shown in Table 12. In this case, the decline in equity is simply the sale

obligation discharged on the old home. To obtain home equity for the old home we use the mortgage reported in the prior wave.

price minus the mortgage. Thus we cannot use the initial home equity to predict the change in equity, as in Table 11 for those who sell and buy again. Thus estimates of the reduction in equity are based on income only. Essentially the estimates tell us how home equity is related to income. For this selected group of households who sell and do not buy another home, home equity is negatively related to income. Based on the simulated changes, the greatest equity reductions occur in families where a household member dies or in which a household member enters a nursing home.

As a summary, the move probabilities and change in home equity results reported in Tables 10-12 are combined to calculate expected change in housing equity. These results are reported on an annual basis in Table 13. The top part of the table shows results for movers who sell and buy another house. The bottom part shows results for movers who sell and discontinue ownership. The table shows results by equity-income quantile, as in several of the tables above. But in this table, the expected change in equity is decomposed into its component parts: the probability of a move, and the change in equity given a move. For example, consider the HRS 22 households. Evaluated at the median of home equity and income, the expected increase in equity through home "upgrading" is \$815. Only 3.3 percent of families upgrade each year, but those that do add \$12,531 to home equity. Averaged over all HRS households, home equity is increased by \$823 through selling and buying a new home. Evaluated at the median of home equity and income, about 1.5 percent of AHEAD 22 households move and buy another home each year. Those that do add \$7,426 to home equity. The expected increase in home equity, averaged across all AHEAD household types, is \$399. Viewed in this way, the expected changes in the equity of HRS and AHEAD households are not very different at the median: +\$823 for the HRS group and +\$399 for the AHEAD group.

For HRS 22 households in the high-equity-low-income group, the expected

annual reduction in equity is -\$486: 3.2 percent move and, given a move, the reduction in home equity is -\$7,711. Averaged over all HRS households in this high-equity-low-income group, the expected reduction in home equity through selling and buying another home is -\$528. The AHEAD households reveal a similar pattern, although again they are less likely to move than the younger HRS households.

The estimates for persons who sell and discontinue ownership are shown in the bottom portion of the table. Again consider 22 HRS families evaluated at the median of equity and income. Only 0.7 percent of households discontinue ownership each year. Those that do reduce equity by -\$29,162 on average. Averaged over all HRS 22 families, equity is reduced by -\$379 through divesting of homes. This reduction can be compared to the +\$815 average increase through upgrading.

Overall, the average equity of all HRS households is reduced by -\$610 in this way, compared to an increase of +\$823 through upgrading. For all AHEAD households average equity is reduced by -\$1,918 by sellers who discontinue ownership between survey waves, compared to an increase of +\$399 through movers who upgrade.

Table 14 presents an accounting of the expected <u>annual</u> change in the home equity of all HRS and of all AHEAD initial homeowners.¹⁴ The first column shows the expected change in home equity for households who move and purchase another home. (Recall that the expected change is the probability of a move times the average change in home equity given a move.) In both surveys, families that move to a new home increase home equity on average. The second column is the expected change in home equity for households that discontinue ownership. These effects are negative and largest among households experiencing precipitating shocks. The third column—the sum of the first two columns—is the net annual change in home equity.

¹⁴Waves of the HRS were two years apart. In the AHEAD there were two years between wave 1 and wave 2, and three years between wave 2 and wave 3.

On average, HRS households increase home equity by \$214 per year. AHEAD households, on average, reduce home equity by \$1,918 annually. This represents a decline of about 1.76% of initial home equity. Most of this reduction is accounted for by households who experience precipitating shocks. For these households, home equity falls by almost 7.84 percent per year. For continuing one- and two-person households the changes are much smaller. Home equity of two-person households declines by 0.11 percent per year, and home equity for one-person households fell by 1.15 percent per year.

Thus in the absence of precipitating shocks there is little systematic reduction in home equity as families age. Families who move to a new home increase home equity on average. Reductions in equity come from families who sell and discontinue home ownership. And most of these moves are associated with precipitating shocks to family status. We find no systematic withdrawal of home equity to support non-housing consumption.

D. CONCLUSIONS

Home equity is the principle asset of a large fraction of elderly Americans. In this paper we have used HRS and AHEAD panel data, as well as SIPP data, to determine whether households withdraw assets from housing equity as they age. We give particular attention to the relationship between changes in home equity and changes in household structure. There are two ways for households to change home equity: by discontinuing home ownership or by selling and moving to another home. We find that households are unlikely to discontinue home ownership. Giving up home ownership most often follows the death of a spouse or entry of a family member into a nursing home. But even in these circumstances, selling the home is the exception and not the rule. In the absence of a precipitating shock, it is much more likely that a family will sell

and buy a new home than discontinue ownership. And, households who sell and buy again tend to increase rather than reduce home equity. That is, assets are transferred to housing. Overall--combining the effects of discontinuing ownership and moving to another home--we find that housing equity of HRS households increases with age, and the equity of AHEAD households declines slightly. The overall decline among the older AHEAD households is about 1.76 percent per year is primarily accounted for a 7.84 percent decline among households experiencing precipitating shocks to family status. Families that remain intact reduce housing equity very little, about 0.11 percent per year for two-person households and 1.15 percent per year for one-person households.

We use two approaches to determine whether households wish to reduce home equity as they age. One approach is to compare changes in home equity of movers to the changes of stayers. If households are withdrawing equity when they sell and move to a new home, then the change in the equity of the movers will be smaller than the change for stayer. These results, however, are confounded by the tendency of the reported self-assessment of home valuation to exceed true home value. A comparison of the selling price of a home with the prior self-assessment of the home value shows that home values reported prior to a sale far exceed realized sales prices. Taking account of the bias, we find that the comparison of equity in the old home compared to the new home (and using the change in the home value of stayers as a control) suggests that families who sell and buy a new home increase home equity on average.

The second approach is based on the comparison of the selling price of the old home (minus the mortgage on the home) with the reported equity value in the newly purchase home. We believe that these are the most reliable data on the change in home equity when an old home is sold and a new home is bought. Based on these "sale price" data, we find that on average households increase home equity when they move to a new house. We also find that equity-rich and income-poor families tend to

reduce home values when they sell and buy a new house. For continuing two-person HRS households, for example, the between-wave reduction for those at the 80th equity quantile and at the 20th income quantile is estimated to be -\$15,422. On the other hand, house-poor and income-rich families tend to increase home equity when they sell and buy a new home. For continuing two-person HRS households at the 20th equity quantile and the 80th income quantile, the increase in housing equity is estimated to be +\$54,778.

As we emphasized in our earlier paper, these results suggest that in considering whether families have saved enough to maintain their pre-retirement standard of living after retirement, housing equity should not be counted on to support general non-housing consumption. Families apparently do not intend to save to finance general retirement consumption through investment in housing, as they might through a 401(k) plan or through some other financial form of saving. Rather the findings here, as well as our earlier findings, indicate that families purchase homes to provide an environment in which to live, even as they age through retirement years. For these reasons, they are unlikely to want reverse annuity mortgages to withdraw assets from home equity. It may be appropriate, however, to think of housing as a reserve or buffer that can be used in catastrophic circumstances that result in a change in household structure. In this case, having used the home equity along the way—through a reverse mortgage for example—would defeat the purpose of saving home equity for a "rainy day."

Although these results are based largely on new HRS and AHEAD data files, and are based on different methods of analysis, the findings correspond closely to the conclusions we reached in our earlier papers, based on different data sources. These conclusions also correspond closely to the findings of a recent survey of older households sponsored by the American Association of Retired Persons (AARP), showing that the preponderance of older families agree with the statement that: "What

I'd really like to do is stay in my current residence as long as possible'."¹⁵ Like our findings, the results of the AARP survey also imply that most households do not intend to liquidate housing equity to support non-housing retirement consumption.

¹⁵More detail is presented in Venti and Wise [2001].

APPENDIX: MORTALITY CORRECTION

The analyses using the SIPP data are based on cohorts constructed from cross-section surveys. For example, the home ownership (or home equity) profile for a cohort is constructed by combining data for all households age A in the first survey year with data for households age A+T from a survey T years later. If the likelihood of survival from A to A+T is related to wealth, then these cohort profiles can be affected by differential mortality. We correct for this problem by reweighting the sample. Households are assigned an adjusted weight that is inversely related to the probability of survival from age A to age A+T.

Baseline estimates of these survival probabilities for one and two person households are obtained from waves 1 and 2 of AHEAD. A one-person household "survives" if the person is present in waves 1 and 2. A two-person household "survives" if both members are present in the second wave. Survival probabilities are estimated from the AHEAD for five year age intervals and for housing equity quartiles. Households that are older and households that have lower levels of housing wealth are less likely to survive. Since the AHEAD only includes households age 70 and over, published survival rates by age (from the NCHS) were used to extrapolate the AHEAD survival probabilities back to age 50.

The final step is to reweight the data. For each household observation of age A and housing equity quartile Q, the SIPP frequency weight is multiplied by the inverse of the cumulative survival probability. The survival probabilities are assumed to be one

for households less than age 50. Thus households that are unlikely to survive are given higher weights. For each observation the probability of surviving to age A given equity quartile Q is

$$S(A,Q) = \prod_{a=50}^{A} s(a,a+1:Q)$$

where s(a,a+1;Q) is the one-year survival rate for a household in equity quartile Q. For each household in each year the SIPP frequency weight is multiplied by the inverse of S(A,Q).

REFERENCES

American Association of Retired Persons. May 2000. "Fixing to Stay: A National Survey of Housing and Home Modification Issues."

Congressional Budget Office. 1993. Baby Boomers in Retirement: An Early Perspective. September.

Bernheim, B. Douglas. 1992. Is the Baby Boom Generation Preparing Adequately for Retirement. Technical Report. Merrill Lynch. Princeton N.J.

Engen, Eric, William Gale and Cori Uccello. 1999. "The Adequacy of Retirement Saving." *Brookings Papers on Economic Activity*. Number 2, pp. 65-165.

Feinstein, Jonathan and Daniel McFadden. 1989. "The Dynamics of Housing Demand by the Elderly: Wealth, Cash Flow, and Demographic Effects", in D. Wise (ed.) *The Economics of Aging*, University of Chicago Press.

Gustman, Alan and Thomas Steinmeier. 1999. "Effects of Pensions on Savings: Analysis with Data From the Health and Retirement Study." Carnegie-Rochester Conference Series on Public Policy. Vol. 50. June. P271-324.

Hurd, Michael. "Portfolio Holdings by the Elderly." Mimeograph. December 1999.

Juster, F. Thomas and Richard Suzman. 1995. "An Overview of the Health and Retirement Study." Journal of Human Resources. Vol. 30. PS7-S56.

Kiel, Katherine and Jeffrey Zabel. 1999. "The Accuracy of Owner-Provided House Values: The 1978-91 American Housing Survey." Real Estate Economics. Vol. 27 no. 2 p. 263-298.

Mayer, Christopher and Katerina Simons. 1994. "Reverse Mortgages and the Liquidity of Housing Wealth." Journal of the American Real Estate and Urban Economics Association. Vol. 22 No. 2 pp. 235-255.

Megbolugbe, Issac, Jarjisu Sa-Aadu, and James Shilling. 1997. "Oh, Yes, the Elderly Will Reduce Housing Equity under the Right Circumstances." Journal of Housing Research. Vol. 8 no. 1 pp53-74.

Merrill, Sally R. 1984. "Home Equity and the Elderly", in H. Aaron and G. Burtless (ed.), *Retirement and Economic Behavior*. Brookings Institution.

Merrill, Sally R, Meryl Finkel and Nadine Kutty. 1994. "Potential Beneficiaries From Reverse Mortgage Products for Elderly Homeowners: An Analysis of AHS Data." Journal of the American Real Estate and Urban Economics Association. Vol. 22 No. 2 pp. 257-299.

Moore, James F. and Olivia S. Mitchell. 1997. "Projected Retirement Wealth and Savings Adequacy in the Health and Retirement Study." NBER Working Paper No. 6240, October.

Sheiner, Louise and David Weil. 1993. "The Housing Wealth of the Aged." NBER Working Paper No. 4115.

Venti, Steven F., and David A. Wise. 1989. "Aging, Moving, and Housing Wealth", in D. Wise (ed.) *The Economics of Aging*, University of Chicago Press.

Venti, Steven F., and David A. Wise. 1990. "But They Don't Want To Reduce Housing Equity", in D. Wise (ed.) *Issues in the Economics of Aging*, University of Chicago Press.

Venti, Steven F. and David A. Wise. 1991. "Aging and the Income Value of Housing Wealth." Journal of Public Economics. Vol. 44. P. 371-397.

Venti, Steven F. and David A. Wise. 2001. "Aging and Housing Equity." NBER Working Paper No. 7882.

in the HRS All % Grow 43.75 68.3 2.97 4.6 2.06 3.2	up % Sequence 39% 222	nces in	the AHEA	
All % Groot 43.75 68.3 2.97 4.6 2.06 3.2	up % Sequence 39% 222	N		
43.75 68.3 2.97 4.6 2.06 3.2	39% 222		All %	Group %
2.97 4.6 2.06 3.2		1203		
0.13 0.2 4.06 6.3 1.73 2.7 0.62 0.9 0.13 0.2 4.98 7.7 1.53 2.4 1.24 1.9 0.2 0.3	22% 220 27% 22N 21% 22T 24% 2DD 21% 200 27% 2DT 21% 2ND 29% 2TT 20% 2D0 24% 2NN 31% Subtotal	293 133 33 27 234 112 47 26 20 19 11 2158	19.93 4.86 2.2 0.55 0.45 3.88 1.86 0.78 0.43 0.33 0.31 0.18	55.75% 13.58% 6.16% 1.53% 1.25% 10.84% 5.19% 2.18% 1.20% 0.93% 0.88% 0.51% 100.00%
1.57 4.4 0.69 1.9 0.16 0.4 0.13 0.3 2.37 6.7 0.91 2.5 0.13 0.3 4.27 12. 0.85 2.4	66% 11D 95% 11N 95% 110 97% 1DD 70% 100 98% 1ND 97% 1NN 10% Subtotal	2217 405 186 142 462 266 98 66 3842 35 6035	36.74 6.71 3.08 2.35 7.66 4.41 1.62 1.09	57.70% 10.54% 4.84% 3.70% 12.02% 6.92% 2.55% 1.72% 100.00%
	1.57 4.4 0.69 1.9 0.16 0.4 0.13 0.3 2.37 6.7 0.91 2.5 0.13 0.3 4.27 12. 0.85 2.4	1.57 4.46% 11D 0.69 1.95% 11N 0.16 0.45% 110 0.13 0.37% 1DD 2.37 6.70% 100 0.91 2.58% 1ND 0.13 0.37% 1NN 4.27 12.10% Subtotal 0.85 2.40% 100.00% Other All	1.57 4.46% 11D 405 0.69 1.95% 11N 186 0.16 0.45% 110 142 0.13 0.37% 1DD 462 2.37 6.70% 100 266 0.91 2.58% 1ND 98 0.13 0.37% 1NN 66 4.27 12.10% Subtotal 3842 0.85 2.40% Other 35 All 6035 0.74	1.57 4.46% 11D 405 6.71 0.69 1.95% 11N 186 3.08 0.16 0.45% 110 142 2.35 0.13 0.37% 1DD 462 7.66 2.37 6.70% 100 266 4.41 0.91 2.58% 1ND 98 1.62 0.13 0.37% 1NN 66 1.09 4.27 12.10% Subtotal 3842 0.85 2.40% Other 35 0.6 All 6035 100.02

Table 2. P	ercent Own, I	Rent, and O	ther By Age, of the AHEA	from Wave	1 of the HR	S and Wave
	One-F	erson House	eholds	Two-F	erson House	eholds
age	own	rent	other	own	rent	other
51-53	58.3	34.0	7.7	87.7	10.8	1.5
54-56	54.5	37.0	8.4	90.9	7.7	1.4
57-61	62.5	29.5	8.0	90.5	7.1	2.4
70-74	67.5	22.8	9.8	91.1	7.0	1.9
75-79	64.0	25.6	10.3	87.8	8.6	3.7
80-84	60.3	25.3	14.4	81.1	12.8	6.0
85+	48.4	31.8	19.9	78.7	15.1	6.2

Table 3. Ten	Table 3. Tenure transitions, by initial tenure and by change in household status, for HRS and AHEAD households, in percent.						
	Initia	al Homeowners in th	e HRS				
Change in Household Status	Tenure	Subsequent Period Status (%)	I % Move	N			
22	own	98.3	7.1	9173			
	rent or other	1.7	65.7	165			
2D	own	95.6	8.4	316			
	rent or other	4.4	55.6	13			
2N	own	88.6	18.9	12			
	rent or other	11.4	0	1			
11	own	95.2	6.1	3150			
	rent or other	4.8	54.5	169			
1N	own rent or other	100 0	0	3 0			
	Initial	Homeowners in the	AHEAD				
Change in Household Status	Tenure	Subsequent Period Status (%)	I % Move	N			
22	own	96.9	3.9	2332			
	rent or other	3.1	38.5	75			
2D	own	88.8	9.4	358			
	rent or other	11.2	76.1	51			
2N	own	75	6.4	35			
	rent or other	25	79.9	14			
11	own	91.3	4.5	2841			
	rent or other	8.7	47.2	269			
1N	own	39.9	0	57			
	rent or other	60.1	92.6	79			

	Initial	Renters-Others in	the HRS			
Change in		Subsequent Perio	od			
Household Status	Tenure	Status (%)	% Move	N		
22	own rent or other	22.3 77.7	51.3 21.1	220 822		
2D	own rent or other	12.4 87.6	46.8 40.2	8 64		
2N	own rent or other	0 100	47.5	0 5		
11	own rent or other	11.4 88.6	46.5 22.2	239 2002		
1N	own rent or other	0 100	43.6	0 3		
Initial Renters-Others in the AHEAD						
	Initial F	Renters-Others in th	ne AHEAD			
Change in	Initial F	Renters-Others in the				
Change in Household Status	Initial F			N		
Household		Subsequent Perio	od	N 31 253		
Household Status	Tenure	Subsequent Perio Status (%) 11.9	od % Move 8.8	31		
Household Status 22	Tenure own rent or other own	Subsequent Perio Status (%) 11.9 88.1 14.5	% Move 8.8 10.4 49.5	31 253 11		
Household Status 22 2D	Tenure own rent or other own rent or other own	Subsequent Perio Status (%) 11.9 88.1 14.5 85.5	% Move 8.8 10.4 49.5 22.1	31 253 11 77		

Note: Based on authors' estimates from the HRS and AHEAD. All percentages are based on weighted samples. However, the sample sizes presented in the table are unweighted.

Table 4. Mean Change in Housing Equity of Initial Owners, by change in family status and by subsequent tenure, for movers and stayers, means and medians.

Mean Changes

Change in Status		Tenure in sown	Subsequent rent or other	Period all	Number own	of Observ rent or other	rations all	Initial Home Equity
22	all stayer mover	6569 6686 5074	-54155 -54155	5855 6686 -3305	8918 8295 623	106 0 106	9024 8295 729	102310 102852 96335
2D	all stayer mover	6288 8997 -21935	-28079 -28079	5547 8997 -23169	294 266 28	7 0 7	301 266 35	83212 83939 77158
2N	all stayer mover	4203 4750 1863		4203 4750 1863	12 9 3	0 0 0	12 9 3	83650 88372 63426
11	all stayer mover	642 935 -3739	-48476 -48476	-697 935 -17549	2961 2779 182	86 0 86	3047 2779 268	95555 96012 90829
1N AHEAD	all stayer mover	-44095 -44095		-44095 -44095	2 2 0	0 0 0	0 2 0	77747 77747 0
22	all stayer mover	-4555 -4103 -15877	-73974 -73974	-5367 -4103 -29557	2309 2213 96	30 0 30	2339 2213 126	115978 115103 132706
2D	all stayer mover	-7182 -5777 -20432	-81900 -81900	-13805 -5777 -51390	354 322 32	39 0 39	393 322 71	105418 102228 120352
2N	all stayer mover	-18869 -18498 -24319	-105730 -105730	-37168 -18498 -90020	35 33 2	12 0 12	47 33 14	118825 123456 105715
11	all stayer mover	-4675 -4011 -18500	-92350 -92350	-8446 -4011 -55077	2801 2671 130	126 0 126	2927 2671 256	102764 102209 108598
g1N	all stayer mover	-13013 -13013	-73671 -73671	-48315 -13013 -73671	57 57 0	72 0 72	129 57 72	77533 82910 73671

			Me	edians				Ī
Change in Status		Tenure in S own	Subsequent rent or other	Period all	Number o	of Observ rent or other	ations all	Initial Home Equity
22	all stayer mover	693 1745 -360	-50905 -50905	1474 1745 -4946	8918 8295 623	106 0 106	9024 8295 729	81033 81326 72721
2D	all stayer mover	-1632 2217 -5481	-32530 -32530	1474 2217 -10999	294 266 28	7 0 7	301 266 35	71491 73193 42594
2N	all stayer mover	6794 -2311 15899		2450 -2311 15899	12 9 3	0 0 0	12 9 3	79994 79994 87989
11	all stayer mover	125 639 -389	-40633 -40633	222 639 -8854	2961 2779 182	86 0 86	3047 2779 268	60493 62333 49376
1N	all stayer mover	-3971 -3971		-3971 -3971	2 2 0	0 0 0	0 2 0	33971 33971
AHEAD								
22	all stayer mover	-5179 -2087 -8271	-64173 -64173	-2348 -2087 -16869	2309 2213 96	30 0 30	2339 2213 126	90242 89114 101608
2D	all stayer mover	-10008 -2303 -17712	-73322 -73322	-4869 -2303 -50761	354 322 32	39 0 39	393 322 71	80090 76706 80217
2N	all stayer mover	-26230 -9941 -42520	-90242 -90242	-13978 -9941 -54145	35 33 2	12 0 12	47 33 14	90242 95882 90242
11	all stayer mover	-2087 -1739 -2434	-73322 -73322	-2434 -1739 -37434	2801 2671 130	126 0 126	2927 2671 256	73799 73322 74869
1N	all stayer mover	-6040 -6040	-64173 -64173	-39921 -6040 -64173	57 57 0	72 0 72	129 57 72	64173 69521 64173

Table 5.	Table 5. Change in the housing equity of initial owners and initial renters, by change in family status.						
		Mea	ns	Mediar	ns		
Family Status	Change in Tenure	Change in Housing Equity	Initial Housing Equity	Change in Housing Equity	Initial Housing Equity	Number	
HRS 22							
2D	OO OR RO RR All	6565 -61073 64117 0 6192	102893 61073 0 0 92472	1695 -50905 35000 0 0	81326 50905 0 0 72721	8919 164 215 822 10120	
	OO OR RO RR All	6223 -75575 45707 0 3345	84329 75575 0 0 69176	1734 -52281 6000 0 0	72721 52281 0 0 56928	296 12 8 64 380	
2N	OO OR RO RR	4203 0	83650 0	2450 0 0	79994 0	12 1 0 5	
	All	2850	56727	ő	34854	18	
11	OO OR RO RR All	642 -50716 51883 0 1126	96874 50716 0 0 57784	621 -40663 36361 0 0	62333 40663 0 0 20897	2961 161 228 2002 5352	
1N	OO OR RO	-44095	77747	-3971	33971	2 0 0	
	RO RR All	0 -25501	0 44964	0 -3971	0 33971	3 5	

AHEAD						
22	OO	-4555	116475	-2217	90242	2309
	OR	-80472	80472	-67682	67682	74
	RO	79697	0	45000	0	31
	RR	0	0	0	0	253
	All	-5241	103938	-207	80217	2667
2D 2N	OO OR RO RR All	-7182 -80749 70915 0 -10956	107705 80749 0 0 86415	-2631 -73322 58825 0 0	80217 73322 0 0 62042	354 50 11 77 492
ZIN	OO	-18869	122320	-9941	95882	35
	OR	-97003	97003	-84602	84602	14
	RO	13369	0	13369	0	1
	RR	0	0	0	0	17
	All	-29941	90771	-9782	62042	67
11	OO	-4675	103232	-1739	74869	2801
	OR	-81412	81412	-67682	67682	266
	RO	73623	0	50269	0	128
	RR	0	0	0	0	1744
	All	-5265	64540	0	37434	4939
1N	OO OR RO RR All	-13013 -72546 57386 0 -18043	82910 72546 0 0 30229	-6040 -56401 65000 0 0	69521 56401 0 0	57 79 7 204 347

T	able 6. Compa	rison of Estimate	ed Home Values	and Sale Pri	ces
Survey	Interval and Sample Size	Estimate of Home Value in Initial Year	Reported Sale Price in Next Year	Mean Difference	Percent Difference
		Mea	ins		
HRS	1992-1994 N=250	135,607	115,665	19,942	14.7
	1994-1996 N=233	157,068	123,883	33,186	21.1
	1996-1998 N=236	162,264	138,206	24,048	14.8
AHEAD	1993-1995 N=163	101,568	81,625	19,943	19.6
	1995-1998 N=179	131,382	109,447	21,935	16.7
		Medi	ans		
HRS	1992-1994 N=250	106,151	96,208	7,117	6.7
	1994-1996 N=233	109,838	98,347	8,083	7.4
	1996-1998 N=236	140,159	122,276	8,290	5.9
AHEAD	1993-1995 N=163	83,848	69,094	5,888	7
	1995-1998 N=179	89,445	77,081	6,546	7.3

Source: Authors' calculations from the AHEAD and HRS. All figures are in 1998 dollars are use household weights.

Table 7. Estimates of the mover equity effect using stayers as the "control" group, for <u>initial homeowners</u>, for two- and one-person households, for the HRS and the AHEAD households, by estimation method.

Change in household status	time effect (t)	OLS Estir t statistic	mates mover effect (m)	t statistic	Medi time effect (t)	an Regre t statistic	ession Es mover effect (m)	timates t statistic
HRS 2 to 2 2 to 1 2 to N 1 to 1 1 to N	6686 8997 4750 935	2.26 2.62 0.26 0.45	-1612 -30931 -2887 -4674	0.15 2.67 0.07 0.57	1745 2216 -2311 639	1.66	-2104 -7698 18210 -1028	2.24 1.76 1.16 0.73
AHEAD 2 to 2 2 to 1 2 to N 1 to 1 1 to N	-4103 -5777 -18498 -4011	2.46 1.5 2.61 2.57	-11774 -14656 -5821 -14489	1.38 1.18 0.21 1.99	-2087 -2303 -9941 -1739	3 1.51 3.77	-6185 -15409 -32579 -696	2.46 3.16 4.49 0.47

Note: Too few observations to estimate 1 to N transitions

Table 8 Com	narison of initia	I reported home	equity colling n	rica minus
		e equity at the er		
Interval	Initial reported equity prior to home sale	Selling price minus mortgage	Reported equity at end of interval	Sample size
HRS	an for Household	ds that Purchase	d Another House	e
1992-1994	76518	64940	89317	181
1994-1996	112382	86599	126228	174
1996-1998	108412	89038	120220	166
AHEAD	100412	09030	120990	100
1993-1995	108821	89284	110690	71
1995-1998	154104	114388	123737	61
1995-1990	134104	114300	123737	O I
Mean f	or Households tl	nat Did Not Purc	hase Another Ho	ouse
1992-1994	61851	55697	0	55
1994-1996	52308	57226	Ö	48
1996-1998	72408	86769	Ö	38
AHEAD		00.00	· ·	
1993-1995	75857	61543	0	44
1995-1998	78005	72313	Ö	51
			-	
Med HRS	ian for Househol	ds that Purchase	ed Another Hous	se .
1992-1994	57679	49806	65903	181
1994-1996	74941	69045	88852	174
1996-1998	82636	72082	110964	166
AHEAD	02000	72002	11000-	100
1993-1995	78258	67826	79590	71
1995-1998	95013	70606	96000	61
1000 1000	00010	70000	00000	01
Media	n Households th	at Did Not Purch	ase Another Ho	use
HRS				
1992-1994	55137	39649	0	55
1994-1996	32819	42664	0	48
1996-1998	69561	85949	0	38
AHEAD				
1993-1995	72668	65244	0	44
1995-1998	79590	73213	0	51
Notos:				
Notes:	blee are used			
 No imputed varia All values are in 				
3. The data are not				
o. The data are not	woignica.			

Table 9. Estimates home, by method	of the change in hor of estimations, for HI	ne equity for movers	s who bought another rvals, in 1998 dollars.					
Interval	Estimated Change in Home Equity	t statistic	Sample size					
	OLS Estimates							
HRS								
1992-1994	24377	3.54	181					
1994-1996	39629	2.86	174					
1996-1998 AHEAD	31952	4.55	166					
1993-1995	21406	1.37	71					
1995-1998	9349	0.59	61					
HRS (pooled								
waves)	31345	6.39	373					
2 to 2								
1 to 1	40014	1.73	96 53					
other	20742	1.5	52					
AHEAD (pooled								
waves)	10007	0.04	00					
2 to 2	13887	0.91	63					
1 to 1	9052	0.45	52					
other	43794	2.01	17					
HRS	Median Regres	ssion Estimates						
1992-1994	6303	1.86	181					
			174					
1994-1996	15455	2.35						
1996-1998	19803	3.42	166					
AHEAD	1000	0.04	74					
1993-1995	1066	0.24	71					
1995-1998	9818	1.12	61					
HRS (pooled								
waves)	17152	1 01	272					
2 to 2	17153	4.01	373 96					
1 to 1	-294	0.04	86 52					
other	8856	1.11	52					
AHEAD (pooled								
waves)	2420	0.07	60					
2 to 2	3438	0.37	63					
1 to 1	0	0	52					
other	10111	0.55	17					

Table 10.	Probit estima	ates of m	nove proba	bilities, for	HRS and A	HEAD ho	useholds.		
	Buy Another Home Discontinue Ownership								
			HRS Hou	seholds					
w11 w22 E Y EY constar	Estimate -0.25623 -0.19400 0.00051 0.00822 -0.00004 at -1.35449			w11 w22 E Y EY constant	Estimate -0.00725 -0.30320 -0.00625 -0.01961 0.00006 -1.80789	t-stat 0.06 2.71 3.22 2.66 0.37 16.81			
	Selected Qua income	antiles o equity	f income a	nd initial re	ported hor income	ne equity equity			
20th 50th 80th	17871 42986 81105	30796 68192 131984		20th 50th 80th	17871 42986 81105	30796 68192 131984			
	Sim	ulated P	robabilities	at Selecte	d Quantiles	5			
		'1	2 to 2 Hou	ıseholds					
income	e 20th	equity 50th	80th	income	20th	equity 50th	80th		
20th 50th	0.063	0.065	0.063	20th 50th	0.015	0.013	0.013		
80th	0.069	0.005	0.070	80th	0.011	0.013	0.009		
			1 to 1 Hou	ıseholds					
income	equity 20th	50th	80th	income	equity 20th	50th	80th		
20th	0.055	0.050	0.056	20th	0.031	0.026	0.027		
50th 80th	0.061	0.058	0.062	50th 80th	0.023	0.026	0.020		
	.,	Othe	r Househol	ds (21, 2N,	,				
income	equity 20th	50th	80th	income	equity 20th	80th			
20th	0.090	0.004	0.091	20th	0.031	0.027	0.027		
50th 80th	0.099	0.094	0.099	80th	0.024	0.027	0.021		

AHEAD Households							
Вι	ıy Another	Home		D	iscontinue	Ownersh	nip
w11 w22 E Y EY constant	-0.11317 -0.17486 0.00909 0.01433 -0.00049 -1.69889	20.83	6 :	w11 w22 E Y EY constant	-0.90675 -1.36739 -0.00438 -0.02379 -0.00091 -0.70091	13.57 15.47 0.74 1.09 0.61 8.89	
30	•		r income a	nd initial re	-		
	income	equity			income	equity	
20th 50th 80th	10909 21433 40609	37434 74869 139042		20th 50th 80th	10909 21433 40609	37434 74869 139042	
	Sim	ulated Pi	robabilitie	s at Selecte	d Quantiles	S	
			2 to 2 Ho	useholds			
income	20th	equity 50th	80th	income	20th	equity 50th	80th
20th	0.034		0.041	20th	0.017		0.015
50th 80th	0.037	0.037	0.043	50th 80th	0.014	0.015	0.011
		'1	1 to 1 Ho	useholds	11		
income	20th	equity 50th	80th	income	equity 20th	50th	80th
20th	0.039	0.040	0.047	20th	0.049	0.044	0.044
50th 80th	0.042	0.043	0.049	50th 80th	0.041	0.044	0.035
		Othe	r Househo	lds (21, 2N,	1N)		
income	equity 20th	50th	80th	income	20th	equity 50th	80th
20th	0.049	0.054	0.059	20th	0.228	0.040	0.211
50th 80th	0.053	0.054	0.062	80th	0.204	0.212	0.182

Table 11. Estimates of the change in home equity based on income and equity levels, by estimation method, for the HRS and the AHEAD households, with simulated change in equity for selected income and equity quantiles. For households purchasing another home.

			-	
HRS	\Box	LICA	hΛ	de
III	HU	use	ш	ıus

	OLS estimate	S t-stat	Median regression estimate t-stat
11 Families 22 Families Equity Income Income* Equity Constant	48.4 2134.4 -5315.7 2593.1 10.5 47719.4	0.00 0.16 -10.91 4.40 1.20 3.64	-762.6 -0.08 9765.2 1.04 -4798.4 -8.53 2024.1 2.33 18.4 0.57 25646.6 2.60
	Selected	quantiles o	f income and initial home equity

	income	equity
20th	17871	30796
50th	42986	68192
80th	81105	131984

Simulated equity changes for selected income and equity quantiles

2 to 2						
income	20th	equity 50th	80th	20th	equity 50th	80th
Income	2011	3011	Ootii	2001	30111	oour
20th 50th	38176	25061	-15422	24353	11929	-23870
80th	54778	20001	1854	37510	11020	-9537
1 to 1						
	2046	equity	0046	2041	equity	0046
income	20th	50th	80th	20th	50th	80th
20th 50th	36090	22975	-17508	13825	1402	-34397
80th	52692	22313	-232	26982	1402	-20065
21, 2N, 1N						
:	0041	equity	0041-	0011-	equity	0041-
income	20th	50th	80th	20th	50th	80th
20th	36041	22926	-17557	14588	2164	-33635
80th	52644	22920	-280	27744	2104	-19303

		AH	EAD House	holds		
	OL estimate	.S t-stat		Median regi estimate	ession t-stat	
11 Families 22 Families Equity Income	-15713.5 -8999.6 -6234.6 5998.9	-0.49 -0.29 -5.21 1.83		-20551.8 231.9 -7619.1 2289.0	-0.80 0.01 -4.56 0.60	
Income* Equity Constant	37.5 60189.0	0.36 1.82		141.5 54972.1	0.64 1.77	
	Selected	quantiles	of income a	nd initial home e	quity	
	income	equity				
20th 50th 80th	10909 21433 40609	37434 74869 139042				
Simu	lated equity	/ changes 1	for selected	I income and equi	ity quanti	les
2 to 2		o avvitu			a au titu	
income	20th	equity 50th	80th	20th	equity 50th	80th
20th 50th 80th	34548 52781	17970	-28386 -9021	29758 38129	5337	-46091 -33449
1 to 1						
income	20th	equity 50th	80th	20th	equity 50th	80th
20th 50th	27834	11256	-35099	8974	-15447	-66874
80th	46067		-15735	17345		-54233
21, 2N, 1N income	20th	equity 50th	80th	20th	equity 50th	80th
20th	43547	26070	-19386	29526	E10E	-46323
50th 80th	61781	26970	-22	37897	5105	-33681

Table 12. Estimates of the change in home equity based on income, by estimation method, for the HRS and the AHEAD households, with simulated change in equity for selected income and equity quantiles. For households <u>not</u> purchasing another home.

		nc	ome.					
HRS Households								
	OL estimate	S t-stat		Median regre	ession t-stat			
11 Families 22 Families Equity Income Income* Equity	13084.3 18754.4 0.0 -1791.8 0.0	0.86 1.37 0.00 2.40 0.00		10552.8 18635.4 0.0 -2063.0 0.0	0.48 0.85 0.00 1.46 0.00			
Constant	-69374.6	5.16		-52943.1	2.63			
;	Selected quai	ntiles of inco	ome and ini	tial home equity				
20th 50th 80th	income 17871 42986 81105	equity 30796 68192 131984						
Simulate	ed equity cha	nges for sel	ected inco	me and equity qu	antiles			
2 to 2		-,			*1			
income	20th	equity 50th	80th	eq 20th	uity 50th	80th		
20th 50th	-53822	-58323	-53822	-37994	-43176	-37994		
80th	-65153	-30323	-65153	-51040	-43170	-51040		
1 to 1		oquity		00	uity			
income	20th	equity 50th	80th	20th	50th	80th		
20th 50th	-59492	-63993	-59492	-46077	-51258	-46077		
80th	-70823	-00000	-70823	-59122	-51250	-59122		
21, 2N, 1N		equity		ea.	uity			
income	20th	50th	80th	20th	50th	80th		
20th	-72577	-77077	-72577	-56630	-61811	-61811		
80th	-83907	-11011	-83907	-69675	01011	-69675		

		AHEAD H	louseholds			
	OLS Es estimate	timates t-stat		Median regi estimate	ression t-stat	
11 Families 22 Families Equity Income	24825.9 24737.6 0.0 -6200.7	1.81 1.66 0.00 2.47		27010.7 35495.2 0.0 -6954.9	2.30 2.47 0.00 1.43	
Income* Equity	0.0	0.00		0.0	0.00	
Constant	-72100.7	4.79	omo and init	-71111.1	6.05	
5	-		ome and init	ial home equity		
	income	equity				
20th 50th 80th	10909 21433 40609	37434 74869 139042				
Simulate	d equity cha	anges for se	lected incon	ne and equity q	uantiles	
2 to 2						
income	20th	equity 50th	80th	20th	quity 50th	80th
20th	-54127	60653	-54127	-43203	EOEOO	-43203
50th 80th	-72544	-60653	-72544	-63859	-50522	-63859
1 to 1				_		
income	20th	equity 50th	80th	20th	quity 50th	80th
20th	-54039	COECE	-54039	-51688	50007	-51688
50th 80th	-72455	-60565	-72455	-72344	-59007	-72344
21, 2N, 1N		'1				
income	20th	equity 50th	80th	20th	quity 50th	80th
20th 50th	-78865	Q5201	-78865	-78698	96017	-78698
80th	-97281	-85391	-97281	-99354	-86017	-99354

Table 13. Summary of annual change in home equity, decomposed into probability of a move times the change in equity given the move, by family status, for selected equity and income quantiles. Based on probit move probability estimates and OLS equity change

		,		Equity	-Income (Quantile	
			50-50	80-20	20-80	80-80	20-20
		For movers who					
HRS	22	Prob OmO	.033	.032	.035	.035	.032
		Change OmO	12531	-7711	27389	927	19088
		Expected Change	815	-486	1890	65	1203
	11	Prob OmO	.029	.028	.031	.031	.028
		Change OmO	11488	-8754	26346	-116	18045
		Expected Change	667	-490	1607	-7	993
	Other	Prob OmO	.047	.046	.050	.050	.045
		Change OmO	11463	-8779	26322	-140	18021
		Expected Change	1078	-799	2606	-14	1622
	All	Expected Change	823	-528	1935	42	1221
AHEAD	22	Prob OmO	.015	.017	.015	.018	.014
		Change OmO	7426	-11730	21810	-3728	14276
		Expected Change	275	-481	807	-160	486
	11	Prob OmO	.018	.019	.017	.020	.016
		Change OmO	4651	-14504	19036	-6502	11502
		Expected Change	200	-682	800	-319	449
	Other	Prob OmO	.022	.024	.022	.026	.020
		Change OmO	11145	-8011	25529	-9	17995
		Expected Change	602	-473	1353	0	882
	All	Expected Change	399	-528	1045	-130	650

For movers who sell and discontinue ownership							
Equity-Income Quantile							
			50-50	80-20	20-80	80-80	20-20
HRS	22	Prob OmR	.007	.007	.006	.005	.008
		Change OmR	-29162	-26911	-32577	-32577	-26911
		Expected Change	-379	-350	-359	-293	-404
	11	Prob OmR	.013	.014	.012	.010	.016
		Change OmR	-31997	-29746	-35412	-35412	-29746
		Expected Change	-832	-803	-815	-708	-922
	Other	Prob OmR	.014	.014	.012	.011	.016
		Change OmR	-38539	-36289	-41954	-41954	-36289
		Expected Change	-1041	-980	-1007	-881	-1125
	All	Expected Change	-610	-576	-588	-502	-662
AHEAD	22	Prob OmR	.006	.006	.006	.005	.007
		Change OmR	-25063	-22367	-29977	-29977	-22367
		Expected Change	-376	-336	-420	-330	-380
	11	Prob OmR	.018	.018	.017	.014	.020
		Change OmR	-25027	-22330	-29940	-29940	-22330
		Expected Change	-1101	-983	-1228	-1048	-1094
	Other	Prob OmR	.088	.087	.084	.075	.094
		Change OmR	-35286	-32589	-40199	-40199	-32589
		Expected Change	-7481	-6876	-8200	-7316	-7430
	All	Expected Change	-1918	-1743	-2116	-1849	-1907

Table 14. Ac	Table 14. Accounting for the change in home equity among initial homeowners in the HRS and the AHEAD						
	tile fik5 alld tile AffEAD						
	Expected A	nnual Change in H	ome Equity				
Survey and household structure	Move and purchase new home	Discontinue home ownership	All	Initial home equity	% of initial equity		
HRS							
22 11 Other	815 667 1078	-379 -832 -1041	436 -166 37	75128 81105 79858	0.58 -0.20 0.05		
All	823	-610	214	76952	0.03		
AHEAD							
22 11 Other All	275 200 602 399	-376 -1101 -7481 -1918	-101 -901 -6879 -1519	94257 78496 87777 86445	-0.11 -1.15 -7.84 -1.76		
column 2: Pr((column 3: E(D	OmO)* E(DHE O OmR)* E(DHE O OHE O) al home equity o	mR)					

Figure 1. Percent Owning for Two-Person Households
Mortality Adjusted Data from SIPP

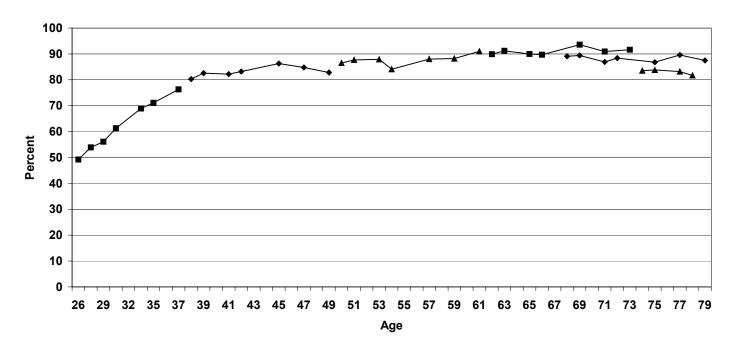


Figure 2. Percent Owning for One-Person Households
Mortality Adjusted Data from SIPP

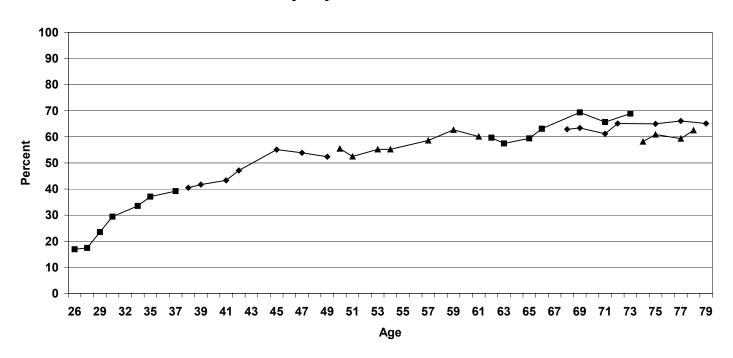


Figure 3. Home Equity for Two-Person Households
Mortality and CPI Adjusted Data from SIPP

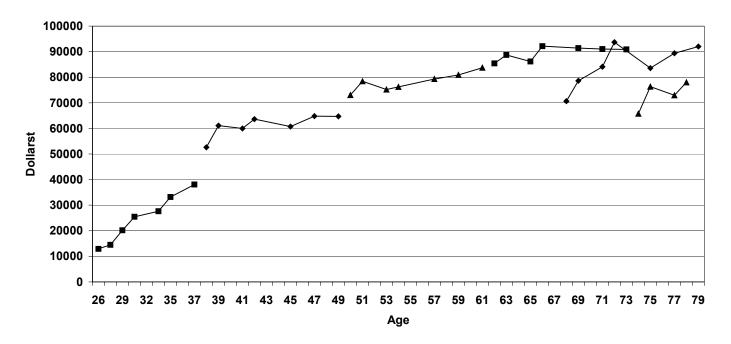


Figure 4. Home Equity for One-Person Households
Mortality and CPI Adjusted Data from SIPP

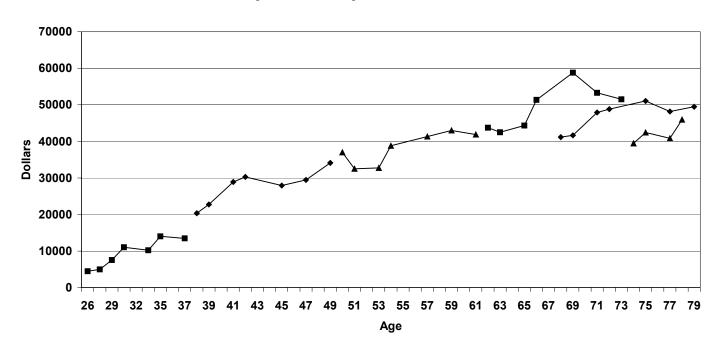


Figure 5a. Percent Owning for Two-Person Households
Data from HRS and AHEAD

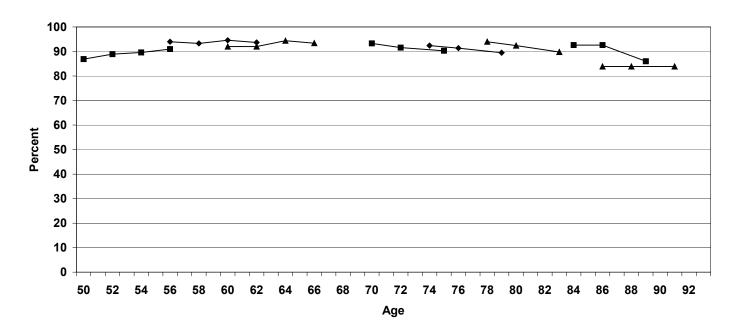


Figure 5b. Percent Owning for Two-Person Households
Data from AHEAD Only

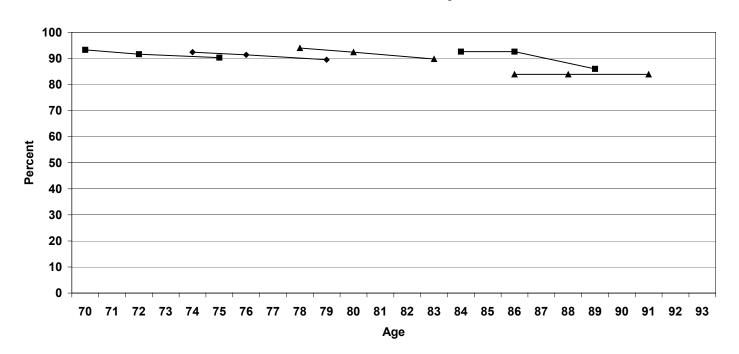


Figure 6a. Percent Owning for One-Person Households
Data from HRS and AHEAD

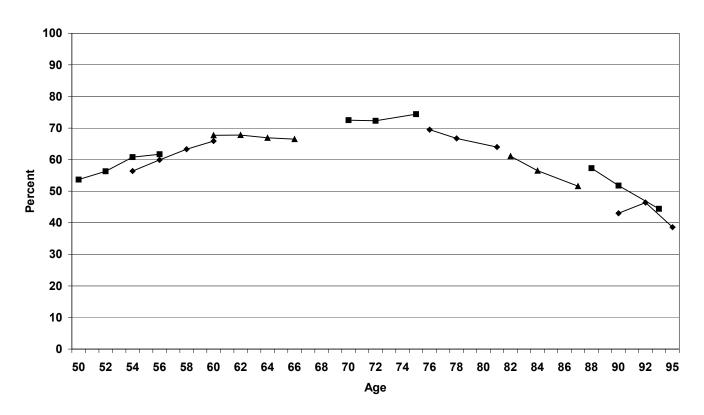


Figure 6b. Percent Owning for One-Person Households
Data from AHEAD Only

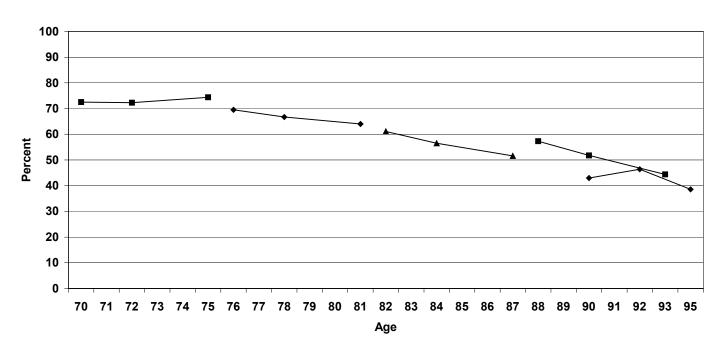


Figure 7a. Mean Home Equity for Two-Person Households
Data from HRS and AHEAD

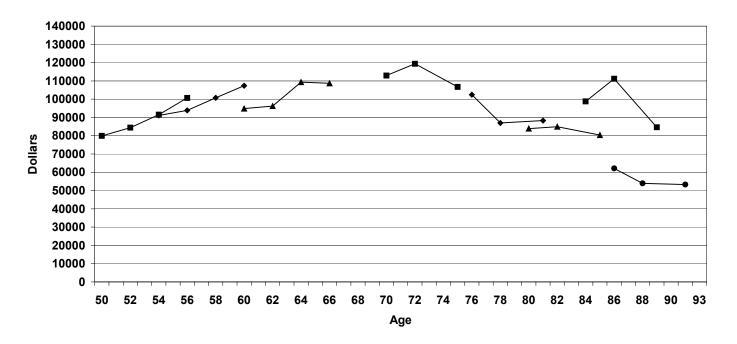


Figure 7b. Mean Home Equity for Two-Person Households
Data from AHEAD Only

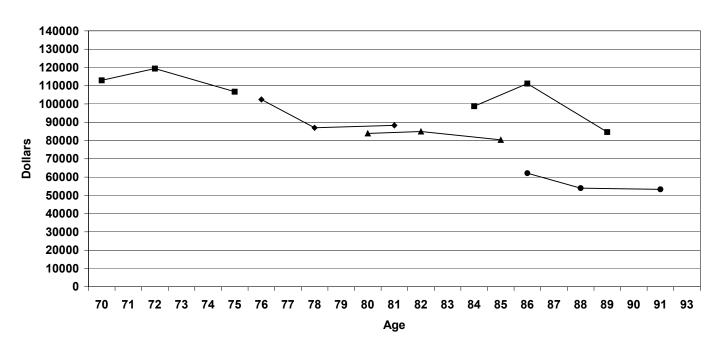


Figure 8a. Mean Home Equity for One-Person Households
Data from HRS and AHEAD

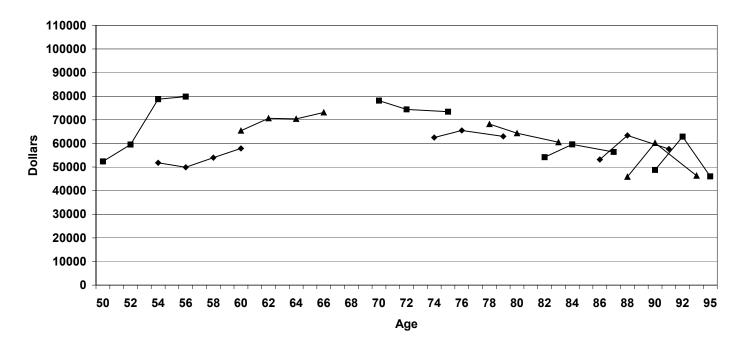


Figure 8b. Mean Home Equity for One-Person Households
Data from AHEAD Only

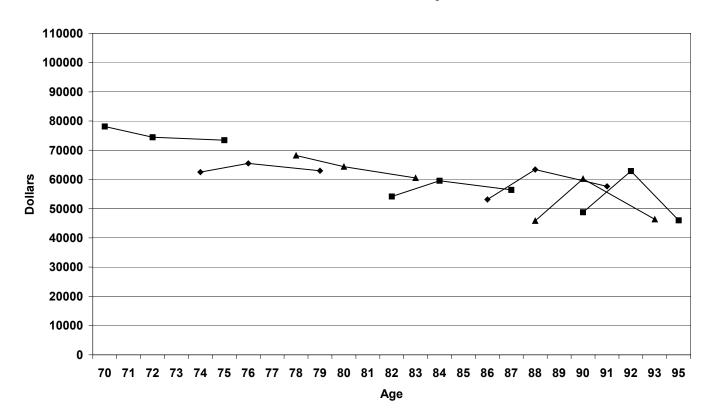


Figure 9a. Median Home Equity for Two-Person Households
Data from HRS and AHEAD

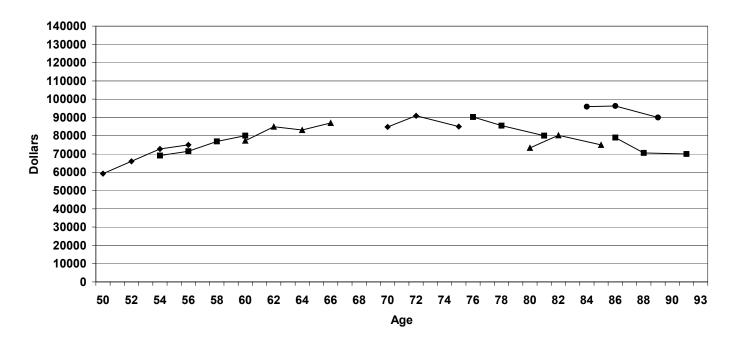


Figure 9b. Median Home Equity for Two-Person Households
Data from AHEAD Only

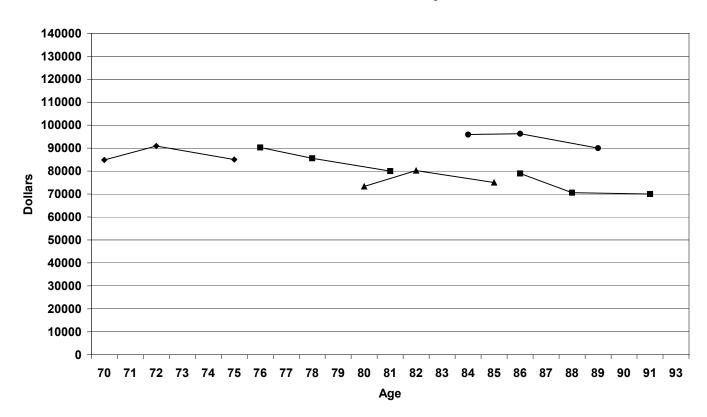


Figure 10a. Median Home Equity for One-Person Households
Data from HRS and AHEAD

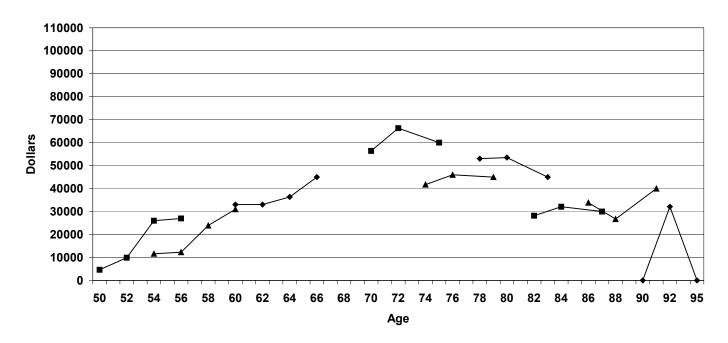


Figure 10b. Median Home Equity for One-Person Households
Data from AHEAD Only

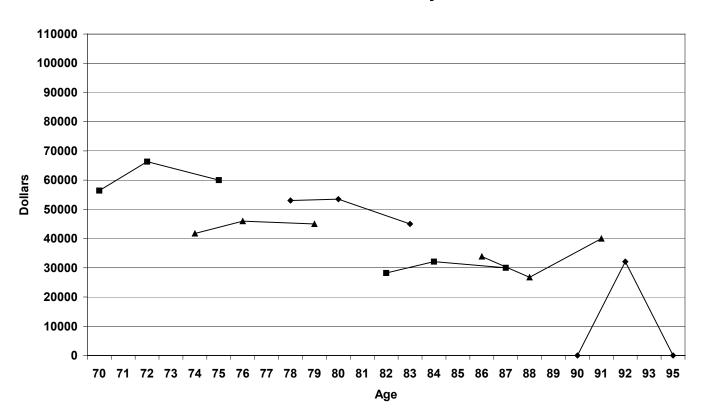


Figure 11a. Mean Non-Housing Equity for Two-Person Households
Data from HRS and AHEAD

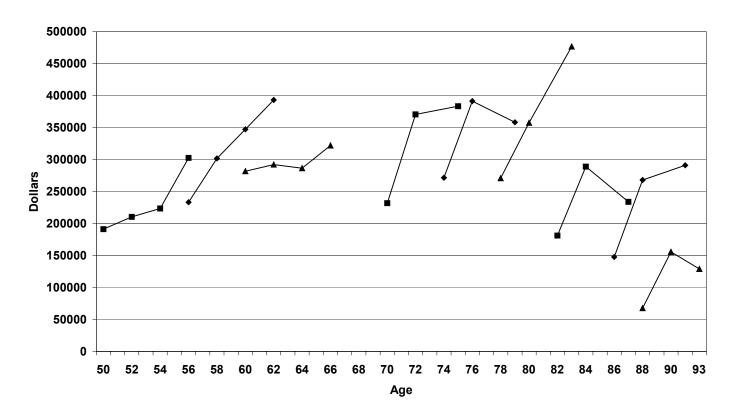


Figure 11b. Mean Non-Housing Equity for Two-Person Households
Data from AHEAD Only

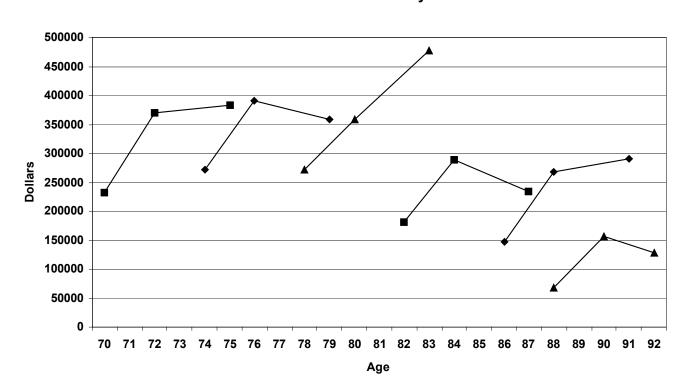


Figure 12a. Mean Non-Housing Equity for One-Person Households
Data from HRS and AHEAD

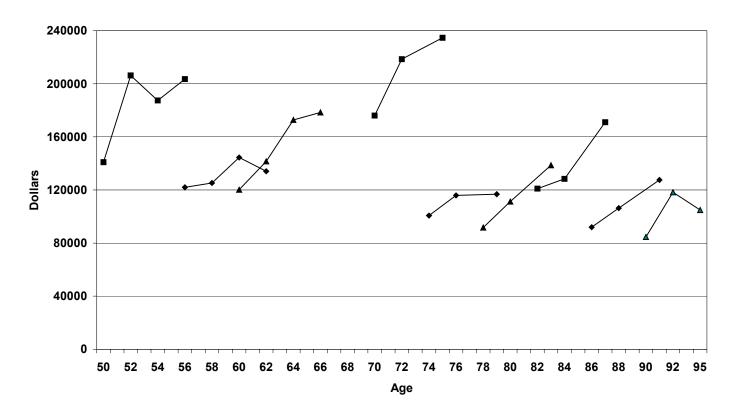


Figure 12b. Mean Non-Housing Equity for One-Person Households
Data from AHEAD Only

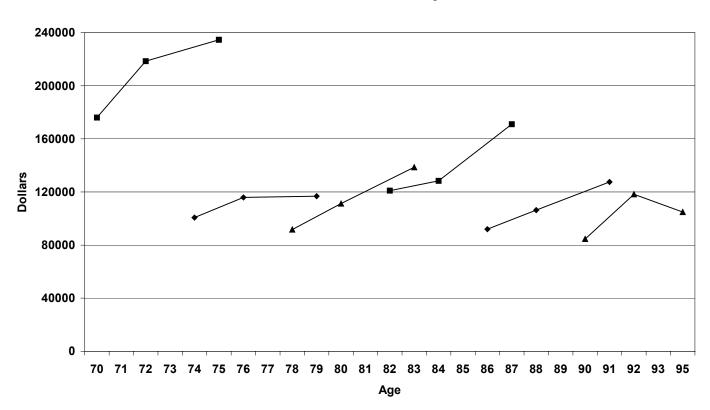


Figure 13a. Median Non-Housing Equity for Two-Person Households
Data from HRS and AHEAD

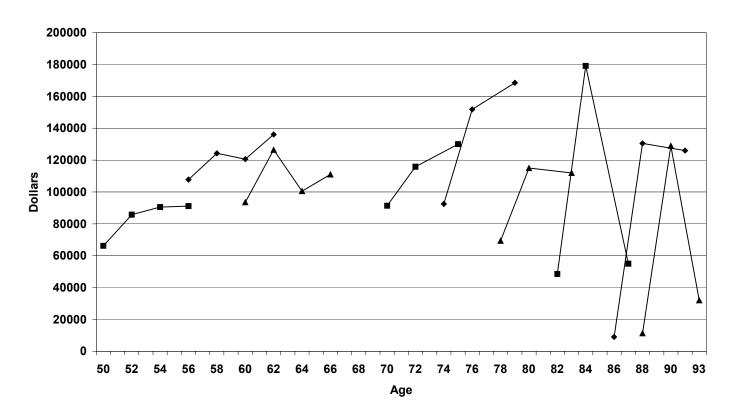


Figure 13b. Median Non-Housing Equity for Two-Person Households
Data from AHEAD Only

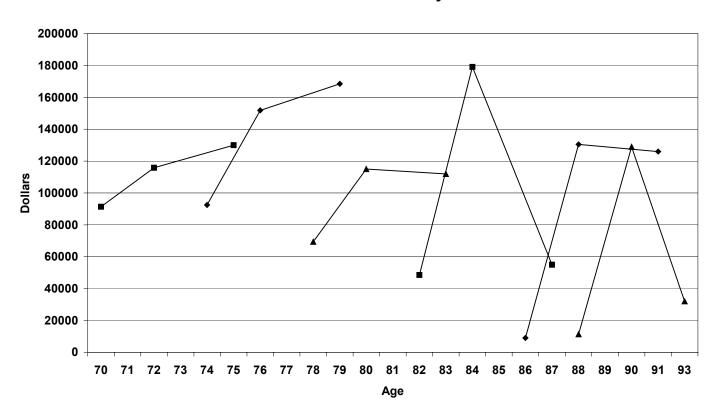


Figure 14a. Median Non-Housing Equity for One-Person Households
Data from HRS and AHEAD

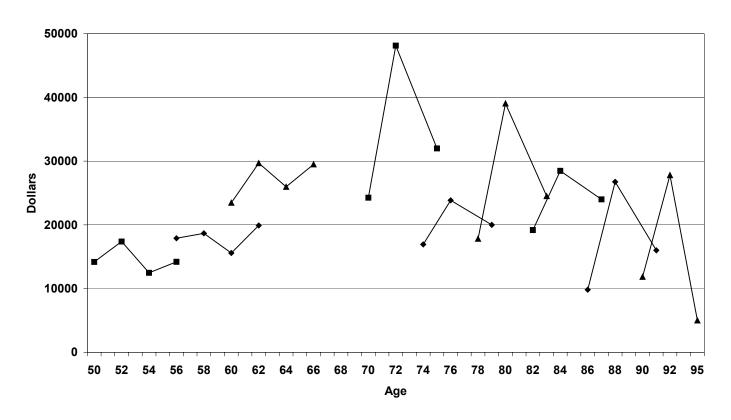


Figure 14b. Median Non-Housing Equity for One-Person Households
Data from AHEAD Only

