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The Decline of Defined Benefit Retirement Plans and Asset Flows

James M. Poterba, Steven F. Venti, and David A. Wise

Many analysts have suggested that population aging will adversely affect the assets of baby boomers when they retire. They argue that when a large population cohort is working and accumulating resources for retirement, their demand for wealth is high, and this raises the price of financial assets and other stores of wealth. Conversely, when a large cohort retires, the argument suggests that cohort members are likely to sell their assets to finance consumption and thereby to drive down asset prices. This argument suggests that the rapidly increasing population of older people in the United States and around the world might lead to lower returns in financial markets in the decades ahead.

This chapter examines the effect of population aging on the demand for financial assets in retirement saving plans, particularly defined benefit (DB) plans, in the United States. It is part of a larger project that aims to evaluate the potential empirical importance of demographic trends on financial market returns in the United States. Our analysis focuses on re-

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tirement saving programs because the future effect of demographic trends on financial asset demand is likely to be most pronounced in asset holdings in retirement saving plans. The inflows and outflows from these plans, particularly DB plans, are sensitive to demographic trends. Thus, a key stepping stone in understanding the effect of population age structure on asset returns is forecasting the effect of demographic trends on the net cash flows to retirement saving plans and the stock of assets in these plans.

Over the past two and a half decades, there has been a fundamental change in saving for retirement in the United States. There has been a rapid shift from saving through employer-managed DB pensions to defined contribution (DC) retirement saving plans that are largely controlled by employees. Just two or three decades ago, employer-provided DB plans were the primary means of saving for retirement in the United States. But since that time, 401(k) and other personal retirement accounts have become the principal form of retirement saving in the private sector. Defined benefit plans have remained an important form of retirement saving for federal employees and for state and local employees, although even for these employees, personal retirement accounts are becoming increasingly important. More than 80 percent of private retirement plan contributions in 2000 and 2001 were to 401(k) and other personal accounts. Contributions to personal retirement plans accounted for only 12 percent of total contributions to federal pension plans in 2000, but had increased to 17 percent by 2004. Thus, to understand the effect of demographic trends on the demand for retirement assets in the coming decades, it is important to evaluate the likely flows into and out of both 401(k) and DB plans.

In Poterba, Venti, and Wise (2008), we described the rise of 401(k) plans and the implications of this rise for the flow of assets into and out of 401(k) plans over the next four decades. In this chapter we describe the decline in DB plans and assess its implications for the flow of assets into and out of DB plans over the next four decades. We then bring together projections of net pension flows to both DB and DC plans. Schieber and Shoven (1997) consider the implications of population aging for private pension fund saving and project saving as a percent of payrolls. Their projection method differs from ours, and it does not consider the rising importance of self-directed DC plans.

There is a substantial and growing literature on the link between population age structure and returns in financial markets. The U.S. Government Accountability Office (2006) provides a recent review of related research. Several studies have used an overlapping-generations framework to explore the theoretical effects of a transitory increase in the population growth rate, a "baby boom," on the equilibrium rate of return. These studies, while based on stylized models, offer valuable insight on the direction of asset market effects. Other research has taken a more empirical ap-

proach and explored the reduced from relationship between summary measures of demographic structure and the returns to investors holding bonds and stocks. The existing findings, illustrated, for example, in Brooks (2002), Geanakoplos, Magill, and Quinzii (2004), and Poterba (2005), span a range of different potential outcomes. Most of the existing work adopts a closed-economy approach, either studying how a baby boom in a single economy will affect returns in that economy or examining the correlation between a nation's population structure and financial market returns in that nation. Recent analyses, however, such as Boersch-Supan, Ludwig, and Winter (2005) and Krueger and Ludwig (2006), move beyond this setting and examine how demographic trends affect international capital flows as well as domestic asset markets.

Very few studies have used household-level data on asset accumulation to project future asset demand. This chapter on DB plans and our companion paper on 401(k) plans begin with a disaggregated analysis of current asset flows in and out of the pension system and use these flows as a base to project pension assets in future years. These projections are a critical input to assessing the effect of demographically induced asset flows on market rates of return.

This chapter is divided into ten sections. In the first, we present a cohort description of the decline in the participation rate of *employed* persons in DB plans over the past two decades. Then we describe a series of analyses that provide the basis for projections of future assets in DB plans. We begin section 10.2 with a cohort description of the dollar amount of pension benefits received by persons over age fifty-five, and we develop projections of DB benefits in the future. In section 10.3, we present a parallel cohort description of the probability of receipt of DB benefits by persons over the age of fifty-five. In section 10.4, we return to the estimation of DB participation during the working years. While in the first section we considered DB participation profiles for *employed* persons, here we consider analogous profiles for all persons in the population. The estimates obtained in this section are used to supplement the estimates obtained in section 10.3 to project benefits after retirement for all persons in the population. In section 10.5, we combine the information described in the previous three sections to develop projections of the total value of DB benefits in future years. Section 10.6 presents projections of DB pension wealth for cohorts retiring between 1982 and 2040 and, for each cohort, compares DB wealth to projected 401(k) assets from Poterba, Venti, and Wise (2008). In section 10.7, we consider projections of the total value of assets in DB trust funds. Section 10.8 brings together projections for both DB and 401(k) plans to explore the flow in and out of the pension system as a whole. It is the change in the assets in DB trust funds combined with the change in the assets in personal retirement accounts that may affect the rate of return on

the investments of future generations of retirees. In section 10.9, we discuss what our projections imply for the relative change in asset demand in the future. Section 10.10 summarizes our findings.

10.1 **Participation of Employees in DB Plans**

Defined benefit participation data were obtained form the several waves of the Survey of Income and Program Participation (SIPP) for the years 1984, 1987, 1991, 1993, 1995, 1998, and 2003. Our analysis here and in subsequent sections is based on organization of the data by cohort. We sometimes define cohorts by the age of the cohort in 1984 and sometimes by the year in which the cohort attains age sixty-five. When referring to the cohort age in 1984, the age is proceeded by "C." When referring to the year the cohort attains age sixty-five, the year is proceeded by "A." Thus, C65 and A1984 identify the same cohort.

To project DB participation in the future, or to predict DB participation in earlier years, we must make projections beyond the range of the observed SIPP data. The cohorts for which data are observed, and the ages for which data are observed, are shown in table 10.1. The table also shows the cohorts, and ages, for which projections are made. Data for all years spanned by the SIPP surveys are available for cohorts C25 through C45. Data for some of the survey years are available for cohorts C9 to C24 and for cohorts C46 to C64. The cohorts for which some SIPP data are available are noted in bold in the table.

Figure 10.1 shows the data for selected cohorts. It is clear that the DB participation rate of employed persons declined consistently with successively younger cohorts. For example, at age forty-five, the participation

14010 1011	participation given employment								
Cohort defined by age in 1984	Age in 1984	Age in 2003	Cohort defined by year age 65 is attained						
C-1	-1	18	A2050						
C0	0	19	A2049						
C1	1	20	A2048						
C8	8	27	A2041						
C9	9	28	A2040						
C25	25	43	A2024						
C45	45	63	A2004						
C64	64	83	A1985						
C65	65	84	A1984						
C100	100	119	A1949						

Table 10.1 Cohorts, observed data, and projected data for defined benefit

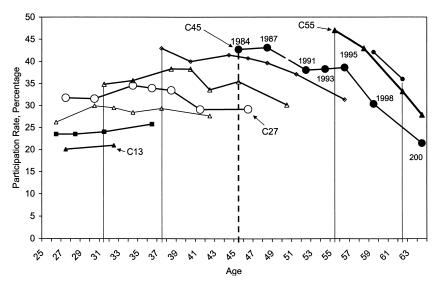


Fig. 10.1 DB participation rate of employed persons, selected cohorts

rate of those who attained age forty-five in 1984, the C45 cohort, was about 43 percent. But the participation rate of the C27 cohort, which attained age forty-five in 2002, was about 29 percent. Comparisons at other ages show similar differences.

Not only is there a cohort effect, with younger cohorts having a successively lower participation rate at all ages, but there is also a *within-cohort* decline in the participation rate with age. The within-cohort decline with age for older cohorts is likely explained in part by retirement. Defined benefit plans typically provide incentives to retire early, and DB participants, on average, retire earlier than persons without these plans; some participants may retire as early as age fifty-five. But even for younger cohorts, there is typically a within-cohort decline in DB participation rates with age.

For comparison, similar cohort data for 401(k) plans and other personal retirement plans are shown in figure 10.2. At age forty-five, the 401(k) participation rate of the cohort that attained age forty-five in 1984 was only 8 percent. But the participation rate of the cohort that attained age forty-five in 2002 was about 47 percent.

Figure 10.3 shows DB participation rates for every other cohort. Again, it is clear that with few exceptions, the data show consistently lower participation rates with successively younger cohorts.

Because we rely on the SIPP data, we have compared the SIPP participation rates by age with rates by age from the Bureau of Labor Statistics (BLS). The comparisons are shown in table 10.2. The SIPP data pertain to

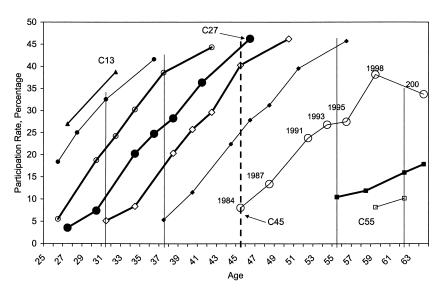


Fig. 10.2 401(k) participation rate of employed persons, selected cohorts

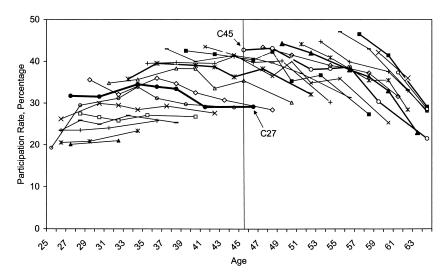


Fig. 10.3 DB participation rate for employees, every other cohort

all employees age twenty-five to sixty-four, including persons employed in private-sector firms, self-employed persons, and persons in federal, state, or local government employment. The BLS data pertain to employees of all ages, but exclude the self-employed and federal employees. Thus, the comparison is imperfect. For the four years that a direct comparison can

Table 10.2		rticipation rate by age, based on Survey of In ation (SIPP) and Bureau of Labor Statistics aged 25–64	
Vanr	RI S privata firms	RIS private and state and local	CIDD

Year	BLS private firms	BLS private and state and local	SIPP
1984			39.4
1985			
1986			
1987			36.8
1988			
1989			
1990	35	43	
1991	34	41	34.9
1992	32	40	
1993	28	37	34.2
1994	28	36	
1995		35	31.6
1996	27	35	
1997		35	
1998		30	29.9
1999	21		
2000	19		
2001			
2002			
2003	20		26.1

Source: Wiatrowski (2005).

be made, the data show that the SIPP estimates are below the BLS estimates as might be expected given the differences in the coverage of the two series. All three series show the same downward trend over time.

The cohort data shown in figures 10.1 and 10.2 allow comparisons between the participation rates of some cohorts at a given age—say forty-five—but these data alone do not allow comparisons that include very young or very old cohorts who were not represented in the SIPP data. For example, we cannot compare the participation rates at age forty-five of cohorts C55 and C13—marked in figure 10.1. To do this, we need to project forward the future participation rates of younger cohorts at older ages and project backward the participation rates of older cohorts at younger ages.

We have made these projections by fitting the cohort data and then using the estimated parameters to predict outside range of the observed data, while relying on estimated cohort effects. The detailed estimates are not shown in the chapter but are available on request. The data represented in figures 10.1 to 10.3 are based on employed persons. As explained in the following, we need to develop estimates based on the percent of all persons who are covered by a DB plan. These projections are explained in section 10.4.

10.2 DB Pension Benefits of Recipients and Projections

We now begin a series of analyses that provide the basis for projections of future assets in DB plans. We begin in this section with a cohort description of the dollar amount of the DB benefits of *persons* who received DB benefits at ages over fifty-five and then describe how we project benefit amounts in the future. In section 10.3, we present a parallel description of the percent of older persons that receives DB benefits.

Data on DB benefits received by retirees, like the participation data for employed persons shown in the preceding, are obtained from SIPP waves for the years 1984, 1987, 1991, 1993, 1995, 1998, and 2003. And, as with the participation data, we first present a cohort description of the observed data. We then fit the observed data and use the fitted model to project benefits outside the range of the observed data. Table 10.3 describes the observed data and the cohorts for which data must be projected. We obtain data on benefits received for persons aged fifty-five to eighty-five. Some SIPP data are available for cohorts C36 to C67. Benefits received for younger cohorts—C47 to C35—and for older cohorts—C68 to C102 must be projected. The numerical amounts table 10.3 show monthly pension benefits observed in the SIPP. An "X" indicates that a pension benefit amount is not observed in the SIPP and that the amount must be estimated. Partial data are available for cohorts C36 to C48. Cohort C36 was age fifty-five in 2003, and thus SIPP data are available for only one year (2003) for this cohort. For successively older cohorts through C48, data are available for more years. Beginning with cohort C55—that was age fiftyfive in 1984—data are available for all seven years that the SIPP data are available. Complete data are available for all cohorts through C66, that was age eighty-five in 2003, the last year of SIPP data. For cohorts C67 through C85, successively fewer years of data are observed in the SIPP. Cohort C85 was age eighty-five in 1984, and thus SIPP data for this cohort are only available in that year.

By fitting the observed data on the receipt of benefits, we can interpolate values for years between the years for which data are available. In addition, the model used to fit the data can be used to project benefits for cohorts outside of the range of the observed data. The cohorts and ages for which such projections must be made are indicated by "X" in table 10.3. We project benefits received back to A1947 because members of this cohort attain age 100 in 1982 and may thus be receiving benefits in the initial year (1982) of our DB asset projections. We project benefits received forward as far out as 2096 to allow calculation of employer DB pension liabilities, described in the following.

Figure 10.4 shows the actual data for the monthly level of benefits for selected cohorts. The data are for all persons age fifty-five to eighty-five and include persons who are receiving benefits from federal or state and local re-

A1947	C102						×	×	×	×	×	×	×	×	×	×	×	×;	× >	< ×	: ×	×	×	×	×	×	×	×	×	×	×	×	××	< >	<
A1947 /	CIGI						×	×	×	×	×	×	×	×	×	×	×	×;	× >	< ×	: ×	×	×	×	×	×	×	×	×	×	×	×	××	< >	<
A1965	: % ×	:×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×;	× >	< ×	: ×	: ×	×	×	×	×	×	×	×	×	×	×	>	< >	~
A1964 A		:×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×;	× >	< ×	:×	×	×	×	×	×	×	×	×	307	×	×	Þ	< >	<
41965		:×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	× >	< ×	: ×	×	×	×	×	×	×	×	303	×	×	×	>	< >	<
41972	: } ×	:×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	× >	< ×	: ×	378	×	450	467	×	×	×	×	×	×	×	>	< >	<
A1973 A	s ×	:×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	×	× >	< ×	375		470	415	×	x	×	457	×	452	×	×	>	< >	x
41982	:	:×	×	×	×	×	×	×	×	×	×	×	478	x	518	546	x	x	x	7. ×	633	×	718	x	×	724	×	×	×	×	×	×	>	< >	v
1983	§ ×	:×	×	×	×	×	×	×	×	×	×	499	×	558	295	×	×	×	288	x 9) ×	756	×	×	757	×	×	×	×	156	×	×	>	< >	x
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$\frac{1}{2} \cdots \frac{A2001}{C48} \cdots$	C48		1065	×I	1375	×I	X	1139	ΧI	X	×I	×I	1152	×	×	×	×	×	× >	< >	: ×	: ×	×	×	×		×	×	×	×		×	>		Y
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A2008 A2007 A2001	<u>41</u> <u>44</u> 2 48	1313 <u>x</u>	<u>x</u> <u>x</u> 1065	X X	\underline{x} \underline{x} $\underline{1375}$	$\frac{1592}{}$ \times \times	$X = 1555 \qquad x$	X X <u>1139</u>	X X	X X	X	X X	X X 1152	×	×	X	X	×; ×;	× > × >	< × < ×	: × : ×	: ×	X X X	×	<u>o C48</u>		2 C85 X	bserved data X	X	X X X	×	x x x	>	< >	v v v
$\frac{A2012}{C^{37}} \dots \underbrace{A2008}_{C^{41}} \underbrace{A2007}_{C^{43}} \dots \underbrace{A2001}_{C^{48}} \dots $	X 1325 X 813	<u>x</u> 1313 <u>x</u>	$X = \frac{\overline{x}}{\overline{x}} = \frac{1065}{\overline{x}}$	\overline{X} \overline{X} \overline{X} X	X x x x	X = 1592 x x	X X 1555 X	X X <u>1139</u>	X X X	x x x	X X X	x x x	X X X 1152	X X X	× × ×	X	X	×; ×;	×	< × < × < ×	: × : × : ×	: ×	X X X X	x x x	s C36 to C48	rvey years) for cohorts C55 to C66	C67 to C85	ween observed data		X X X	X X X	X X X X	>	< >	V V V V
A2008 A2007 A2001	1613 X 1325 X 813	1624 <u>x 1313</u> <u>x</u>	X imes X imes X imes X imes 1065	X X X X	$X X \underline{x} \underline{x} \underline{x} 1375$	X X 1592 x x	X X X 1555 X	X X X X 1139	x x x x	\overline{X} X X X	\mathbf{x} \mathbf{x} \mathbf{x} \mathbf{x}	X X X X	X X X 1152	X X X X	x	X	X X X X	×; ×; ×;	×	< × < × < ×	: × : × : ×	: ×	X X X X	x x x	s C36 to C48	rvey years) for cohorts C55 to C66	C67 to C85	ween observed data		x x x x x	X X X	x	>	< > < > < > < > < > < > < > < > < > < >	V V V V
A2013 A2012 A2008 A2007 A2001 A	X 1613 X 1325 X 813	X = 1624 $X = 1313$ $X = 1313$	X X X X X X 1065	\overline{X} \overline{X} \overline{X} X X X	$X X X \underline{X} \underline{X} \underline{X} \underline{1375}$	X X X 1592 X X	X X X X 1555 X	$X X X X X \frac{1139}{}$	\mathbf{x} \mathbf{x} \mathbf{x} \mathbf{x} \mathbf{x}	X X X X X	\overline{X} X X X X X	X X X X X X	X X X X X X X 1152	x	\mathbf{x} \mathbf{x} \mathbf{x} \mathbf{x} \mathbf{x} \mathbf{x}	X X X X X	X X X X X	x; x; x;	×	< × < × < ×	: × : × : ×	: ×	X X X X	x x x	Partial data for cohorts C36 to C48			ween observed data		x x x x x	X X X X	X X X X X X	>	< > < > < > < > < > < > < > < > < > < >	V V V V V V

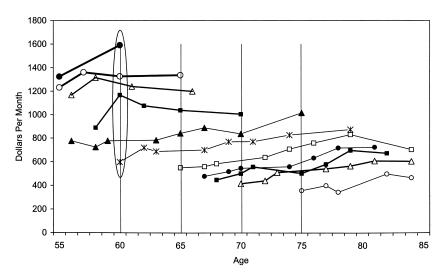


Fig. 10.4 DB benefit receipt for selected cohorts

tirement programs, as well as from private-sector pensions plans. Two features of the data stand out. One is that benefits of younger cohorts are much greater than benefits of older cohorts. For example, the benefit at age sixty for the cohort that attained that age in 1984 was about \$600 per month; the benefit for the cohort that attained age sixty in 2003 was about \$1,600 per month (circled in the figure). The other feature of the data is that within-cohort benefits increase with age. In part this results from the indexing of benefits from some DB plans, especially federal and state and local plans. In addition, private employer plans, which are typically not indexed, sometimes grant cost-of-living increases on an ad hoc basis after retirement.

We fit the cohort data on benefits with the following specification:

$$B_{ac} = \alpha_{a \le 60} A_1 + \alpha_{60 < a \le 65} A_2 + \alpha_{65 < a \le 70} A_3 + \alpha_{70 < a \le 75} A_4 + \alpha_{>75} A_5 + \sum_{c=1971}^{2013} c_c C_c$$

Here, B is the dollar amount of monthly pension benefits, and cohorts are defined by the year the cohort attains age sixty-five—the C_c indicator variable—and the variables A_1 through A_5 specify age as piecewise linear with break points at sixty, sixty-five, seventy, and seventy-five.

1. The level of DB benefits received is derived from SIPP data on receipt of income from the following sources: pension from company or union; federal civil service or other federal civilian employee pension; U.S. military retirement pay; National Guard or reserve forces retirement; state government pension; local government pension; U.S. government railroad retirement; veterans compensation or pension. We have assumed that all monthly income received from these sources is DB income although it is possible that withdrawals of DC assets may be included.

The parameter estimates are presented in table 10A.1. We use the estimated parameters from this specification to project benefits forward for the younger cohorts and to project benefits backward for the older cohorts. There are at least two years of observed SIPP data for cohorts as old as A1970 (observed at age seventy-nine in 1984 and age eighty-two in 1987). There are not enough SIPP observations to reliably estimate cohort effects for cohorts younger than A2012. We obtain benefit estimates for cohorts attaining age sixty-five prior to 1970 by shifting the A1970 benefit profile by according to the Social Security Administration's average wage index. That is, we assume that the pension benefit growth paralleled wage growth between 1949 and 1970. We only use benefits received in calendar years 1982 and later. For cohort A1970, for example, we use estimates beginning at age seventy-seven, in 1982. We use the wage index to estimate cohort effects for cohorts that retired before 1970, but we only project for older ages (after 1982) for each cohort.

The more important projections are for cohorts younger than the A2012 cohort. Which specific extrapolation to use, however, is open to question. The following figures show some of the relevant data. Figure 10.5 shows the estimated cohort effects for cohorts A1982 to A2012. An exponential fit over all of the cohort effects—with an R^2 value of 0.99—suggests that, on average, benefits increase by 4.5 percent with each successively younger cohort. But the figure also shows that the rate of change in the cohort effects may have declined over time. The estimated rate of change was only

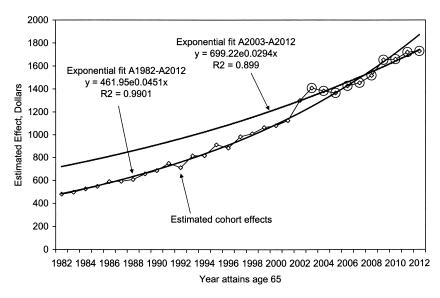


Fig. 10.5 Estimated cohort effect from benefit regression by year cohort attains age 65

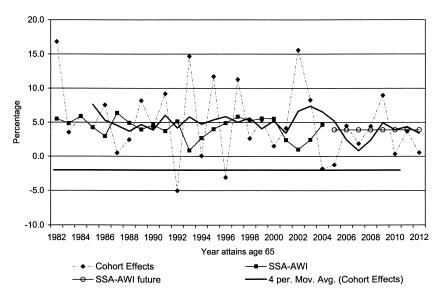


Fig. 10.6 Percent change in estimated benefit cohort effect versus the Social Security average wage index (AWI), by year cohort attains age 65

2.94 percent for the A2003 to A2012 cohorts. (The cohort effects for these years are circled in figure 10.5.)

Figure 10.6 shows a different view of the data, beginning with the change in the estimated cohort effects from one cohort to the next younger cohort. It is clear that the percent changes vary substantially from one cohort to the next. The figure also shows the four-year moving average of these changes. It is clear from the moving average that the average change from one cohort to the next younger cohort was close to 5 percent over most of the period but declined to close to 1 percent around 2006. The figure also shows the Social Security average wage index over the years 1982 to 2003. These data suggest a noticeable correspondence between the four-year moving average of the change in the estimated cohort effects and the wage index. After 2003, the SSA assumes a wage increase of 3.9 percent annually. On average, the change in the cohort effects was only about 2.9 percent over these years.

Based on these data, we have assumed that for younger cohorts benefits will increase at 3.9 percent for each successively younger cohort, the same as the SSA intermediate assumption for the average wage index. Figure 10.7 shows benefit profiles for selected cohorts, including projected profiles for cohorts A2013 to A2050 (dashed lines in the figure) and the fitted cohort profiles for cohorts A1974 to A2012 (solid lines). Profiles for cohorts older than A1974 are not shown (these cohorts are relatively unimportant because our projection of DB assets begins in 1982). All profiles are in year

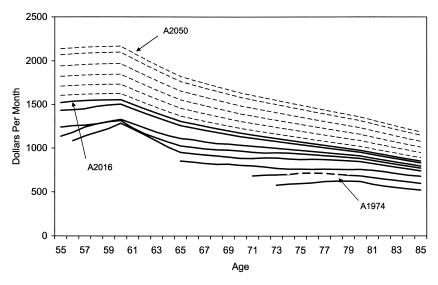


Fig. 10.7 Fitted and projected benefit profiles, selected cohorts (in year 2000 dollars)

Notes: Last fitted profile graphed is A2012. First projected profile graphed is A2016.

2000 dollars. The benefit model parameter estimates show that cohort age profiles in *current* dollars are slightly upward sloping, perhaps due to the indexing of government pensions and ad hoc cost-of-living increases for other pension benefits. The age profiles in 2000 dollars in the figure slope downward, however.

10.3 Receipt of DB Pension Benefits

In the previous section, we discussed the dollar amount of the DB benefits of persons who received DB benefits at each age. To obtain an estimate of the total dollar amount of DB benefits, we need also to determine the proportion of persons who receive benefits at each age. We now consider the probability of benefit receipt.

We begin with cohort data on the percentage of persons receiving DB benefits at each age. Figure 10.8 shows data for selected cohorts (essentially every other cohort). Two features of the data are evident. First, the cohort effects are rather small, with the exception of the older cohorts. In other words, the profile of benefit receipt by age is about the same for all the cohorts represented in the SIPP data (with the exception of the oldest cohorts). For example, among cohorts observed at age seventy, the probability of receipt of benefits ranges from about 45 percent for the youngest cohort to 49 percent for the oldest cohort. Second, the age at which the maximum percent of persons receive benefits is about seventy.

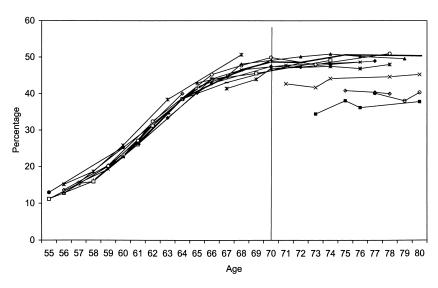


Fig. 10.8 Percent receiving DB benefits, selected cohorts

To project benefit receipt for future (younger) cohorts that are not observed in the SIPP data, we assume the age-benefit-receipt profile shown in figure 10.8 continues to apply. However, we allow this profile to shift (downward) for younger cohorts. We first estimate age and cohort effects for cohorts observed in the SIPP. Then we predict benefits for younger cohorts, assuming that—except for the cohort effects—the age-benefit-receipt profile in the same for younger cohorts as it has been for older cohorts observed in the SIPP.

We fit the benefit receipt data using a probit model and a piecewise linear specification for age like the one presented in the preceding. The sample consists of the same persons age fifty-five to eighty-five used in the previous section to estimate the level of benefits. The specification is

$$\begin{split} R_{ac} &= \alpha_{a \le 60} A_1 + \alpha_{60 < a \le 65} A_2 + \alpha_{65 < a \le 70} A_3 + \alpha_{70 < a \le 75} A_4 \\ &+ \alpha_{75 < a} A_5 + \sum_{1200}^{c=2012} c_c C_c, \end{split}$$

where *R* is the receipt of benefits, and cohorts are defined by the age the cohort attains age sixty-five. The estimated cohort effects are shown in table 10A.2. Figure 10.9 shows that the estimated cohort effects are essentially unchanged for cohorts that attained age sixty-five between 1982 and 2003. The cohort effects were smaller but increasing for cohorts that attained age sixty-five between 1969 and 1981. The larger cohort effects for cohorts that attained age sixty-five in 2004 and later years are based on very few SIPP observations and pertain only to persons who were receiving benefits at

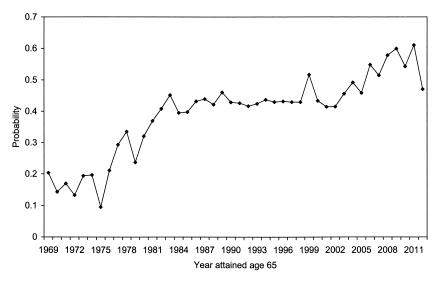


Fig. 10.9 Cohort effects for the receipt of benefits by year attained age 65

young ages. For example, data for the A2004 cohort are only observed through age sixty-four and for the A2012 cohort only for age fifty-six.

Because we cannot reliably estimate cohort effects for cohorts younger that A2003, we assume that the benefit receipt rates of successively younger cohorts follow the same pattern as the DB participation rates of these same cohorts when they were in the labor force. The assumption is that if fewer persons in a particular cohort participated when young, then fewer persons will receive benefits after retirement. We do this by first calculating the maximum DB participation rate over ages twenty-five to sixty-four for each cohort. We use this maximum rate to predict benefit receipt when the cohort is retired. We use cohort-to-cohort changes in the maximum DB participation rate, and the 2003 fitted age-benefit receipt profile, to project benefit receipt for cohorts A2004 through A2060. Before presenting these projections, we describe the population-based DB participation rates that are required to make the calculations.

10.4 Population-Based DB Participation Rates

The previous two sections have developed projections for the level of benefits (conditional on receipt) and for the probability of receiving DB benefits after retirement. As noted in the preceding, the probability of receiving benefits cannot reliably be estimated for cohorts retiring after 2003 using the SIPP data. These cohorts are not observed after age fifty-five in the data. However, these same cohorts are observed at younger ages when they are in the labor force. We infer the probability of benefit receipt after

Year	DB participation given employment	Employment to population ratio	(DB participation) × (Employment to population ratio)
1984	39.4	0.752	29.6
1987	36.8	0.767	28.2
1991	34.9	0.778	27.2
1993	34.2	0.771	26.4
1995	31.6	0.777	24.6
1998	29.9	0.778	23.3
2003	29.7	0.763	22.7

Table 10.4 Defined benefit (DB) participation and ratio of employment to population, all persons aged 25–64

Source: Authors calculations from Survey of Income and Program Participation Surveys.

retirement from the DB participation rates of these cohorts when they were in the labor force. This section develops estimates of DB participation that are closely related to those presented in section 10.1 for *employed* persons. Here we consider the participation rate of *all persons in the population* because we will use our estimates to infer benefit receipt for all retirees (not just those who were employed).

The employment-based participation rates and the population-based rates can differ substantially. To see this, note that the percent of population that participates in a DB plan at age *a* is given by

$$\left(\frac{DB}{P}\right)_{a} = \left(\frac{E}{P}\right)_{a} \Pr[DB_{a} \mid E_{a}],$$

where E is employment, and P is population, and a denotes age. Table 10.4 shows the probability that a person has a DB plan given that the person is employed and the employment to population ratio for each year of SIPP data available between 1984 and 2003. These data show that the last term, $\Pr[DB_a \mid E_a]$, has declined between 1984 and 2003, but the fraction of the population employed $(E/P)_a$ increased between 1984 and 1998 but fell in 2003. The percentage of employed persons participating in a DB plan declined from 39.4 percent to 29.7 percent over this period.

The overall effect of these trends is that cohort effects are smaller when *all* persons are used as the base than when *employed* persons are used as the base. This can be seen by comparing figure 10.1 with figure 10.10, which are shown together in the following. For example, at age forty-five, the difference in the participation rates of cohorts 27 and 45 is about 13 per-

^{2.} The employment to population ratio for persons sixteen and over reported by the BLS shows an increase from 59.5 percent in 1984 to 64.1 percent in 1998 and then a drop to 62.3 percent in 2003. The SIPP data we use are for ages twenty-five to sixty-four.

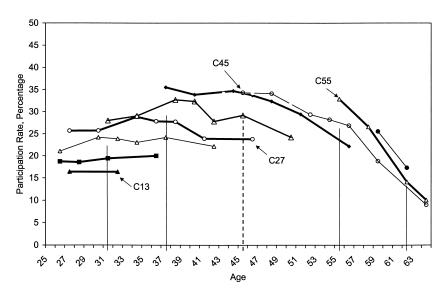


Fig. 10.10 DB participation rate for all persons, selected cohorts

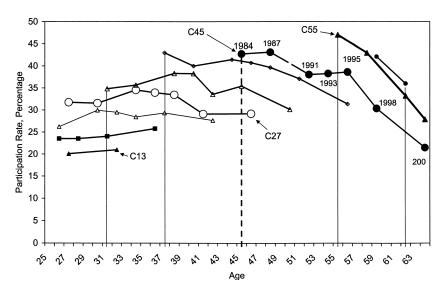


Fig. 10.1 DB participation rate of employed persons, selected cohorts

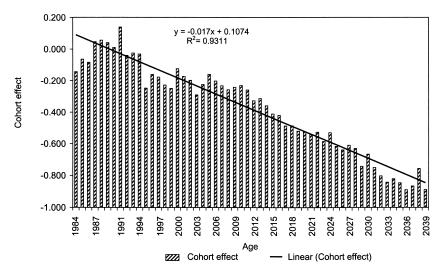


Fig. 10.11 DB participation for all persons, estimated cohort effects, by year attained age 65

centage points based on those employed but only about 10 percentage points based on all persons.

As with the participation based on employed persons, we fit the population-based participation data using a probit model with this specification:

$$\begin{split} DB(\text{Pop})_{ac} &= \alpha_{a \leq 30} A_1 + \alpha_{30 < a \leq 35} A_2 + \alpha_{35 < a \leq 40} A_3 + \alpha_{40 < a \leq 45} A_4 + \alpha_{45 < a \leq 50} A_5 \\ &+ \alpha_{50 < a \leq 55} A_6 + \alpha_{55 < a \leq 60} A_7 + \alpha_{60 < a} A_8 + \sum_{c=0}^{c=65} c_c C_c \end{split}$$

The estimated cohort effects from this specification (defined here by age in 1984) are shown in figure 10.11, while the complete estimation results are shown in table 10A.3.

Selected fitted cohort profiles together with the corresponding cohort data are shown in figure 10.12. We judge that the fitted profiles represent the data quite well. We then use the model estimates to predict population-based participation rates at all ages (twenty-five to sixty-four) for each of the cohorts for which we are able to estimate cohort effects. These predictions are shown (as solid lines) in figure 10.13 for selected cohorts. We predict at all ages for each cohort because we want to determine the age at which the participation rate is at a maximum for each of the cohorts. The use of this calculation is explained in the following. Predicting at older ages for the younger cohorts suggests that by the time these cohorts reach retirement age, their participation in DB plans will be very low. For example, based on these predictions, the DB participation rate of persons in the C11

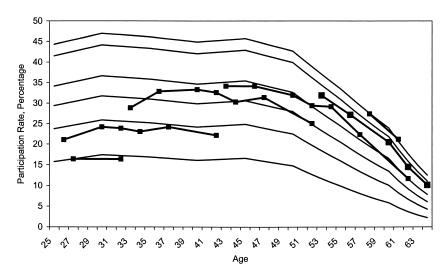


Fig. 10.12 Actual versus fitted DB participation profiles for selected cohorts (all persons)

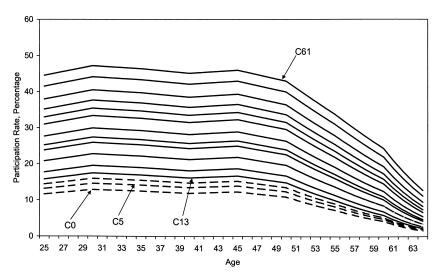


Fig. 10.13 DB participation for all persons, projections to younger cohorts (dotted), (2.1% decline for cohorts that retire 2040-2050)

cohort will be only 8.8 percent when that cohort attains age fifty-five (cohort C11 attains age sixty-five in 2038).

The youngest cohort in the SIPP data is A2040. For later use, we will need cohorts as young as A2050. Figure 10.11 shows that the decline in the estimated cohort effects, with successively younger cohorts, is close to lin-

ear—at -0.017 per cohort. Thus, we extrapolate the estimated cohort effects linearly to project the effects for younger cohorts. This extrapolation yields profiles for the younger cohorts shown as dashed lines in figure 10.13.

All of the projections we make are subject to substantial uncertainty. These projection in particular raise the prospect that the past trend may not be a good predictor of the future trend. For example, several large companies have recently announced that traditional DB pension plans would be phased out. Other companies could follow this lead more rapidly than our projections suggest. This would lead to a faster-than-projected decline in DB assets.

10.5 Projected Benefits Paid

To develop projections of the total value of benefits paid by DB plans, we combine estimates of the level of benefits in section 10.2 with estimates of the probability of benefit receipt in section 10.3. To help to forecast benefit receipt for younger cohorts (retiring after 2003), we use estimates of DB participation of these younger cohorts during their working ages. In particular, we use the population-based estimates of DB participation by cohort discussed in section 10.4 to predict pension receipt after retirement. The necessary data is set out in table 10.5. We first use the estimates of probability of benefit receipt for cohorts A1982 to A2003 to predict the percent of each of these cohorts that receives benefits at age seventy (recall from figure 10.8 that age seventy is the age at which the maximum percent of persons receive benefits). As discussed in section 10.3, for these cohorts there are a sufficient number of SIPP observations to obtain reasonably reliable estimates of benefit receipt. These probabilities are shown in column (3) of table 10.5. The probabilities are also graphed in figure 10.14 for the years 1982 through 2003.

We next need to project benefits at age seventy for younger cohorts (A2004 to A2040). To make these projections, we use the population-based DB participation rates, discussed in section 10.4. From these data, we have calculated the maximum participation rates, over ages twenty-five to sixty-four, for each cohort. These estimates are shown in column (5) of table 10.5. We use the maximum participation rate during working years to predict the probability of benefit receipt at age seventy for cohorts A2004 and younger. We assume that the year-to-year percent change in the probability of receipt of benefits at age seventy is the same as the year-to-year percent change in the maximum DB participation rate. Thus, for example, the last cohort for which the receipt probability is observed is A2003 (47.3 percent). We project the 2004 receipt probability by assuming it declines by the same percentage amount as the maximum participation probability (from 38.6 percent for A2003 to 37.7 percent in 2004). The prediction for 2004,

Table 10.5 Projections of the probability of benefit receipt at age 70, by cohort

Year cohort attains age 65 (1)	Age of cohort in 1984 (2)	Probability receive benefits at age 70 (from benefit receipt data) (3)	Projected probability receive benefits at age 70 (4)	DB participation maximum (from DB participation data) (5)
1982	67	45.5		
1983	66	47.0		
1984	65	48.8		
1985	64	46.5		39.7
1986	63	46.6		42.9
1987	62	48.0		42.1
1988	61	48.3		47.2
1989	60	47.6		45.4
1990	59	49.1		47.6
1991	58	47.9		47.0
1992	57	47.7		45.8
1993	56	47.4		50.9
1993 1994	55	47.4 47.7		43.8
1995	53 54	48.2		43.6 44.4
	53			
1996	53 52	47.9		44.2
1997		48.0		35.8
1998	51	47.9		39.1
1999	50	47.9		38.4
2000	49	51.4		36.6
2001	48	48.0		35.7
2002	47	47.3		40.5
2003	46	47.3	46.1	38.6
2004	45		46.1	37.7
2005	44		41.9	34.2
2006	43		44.9	36.6
2007	42		47.9	39.1
2008	41		45.9	37.5
2009	40		44.5	36.4
2010	39		43.4	35.4
2011	38		44.1	36.0
2012	37		44.6	36.4
2013	36		43.3	35.4
2014	35		40.3	32.9
2015	34		40.9	33.4
2016	33		38.9	31.8
2017	32		36.7	29.9
2018	31		36.3	29.6
2019	30		33.5	27.3
2020	29		33.5	27.4
2021	28		32.1	26.2
2022	27		31.9	26.0
2023	26		31.2	25.5
2024	25		31.9	26.0
2025	24		29.5	24.1

Table 10.5 (continued)

Year cohort attains age 65 (1)	Age of cohort in 1984 (2)	Probability receive benefits at age 70 (from benefit receipt data) (3)	Projected probability receive benefits at age 70 (4)	DB participation maximum (from DB participation data) (5)
2026	23		31.7	25.9
2027	22		28.4	23.1
2028	21		27.5	22.5
2029	20		28.6	23.4
2030	19		27.9	22.7
2031	18		23.9	19.5
2032	17		26.5	21.7
2033	16		23.6	19.3
2034	15		21.9	17.9
2035	14		20.7	16.9
2036	13		21.4	17.4
2037	12		20.6	16.8
2038	11		19.3	15.7
2039	10		20.0	16.3
2040	9		19.5	15.9

Note: DB = defined benefit.

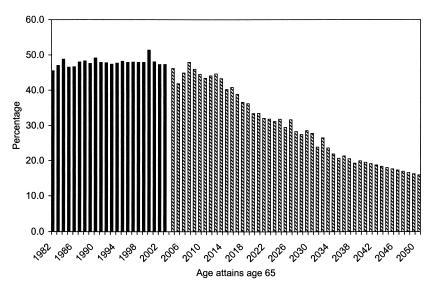


Fig. 10.14 Percent receiving benefits at age 70, by cohort (solid is fitted SIPP estimate, and striped is projected)

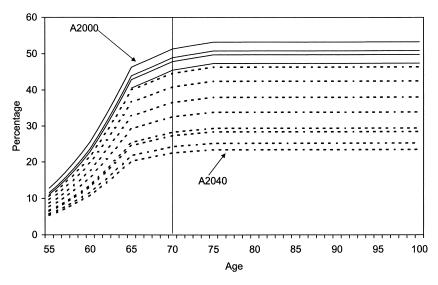


Fig. 10.15 Projected percent receiving benefits for selected cohorts attaining age 65 in years 1982 through 2040 (solid is fitted SIPP estimate, and dotted is projected)

46.2 percent (shown in column [4]) is $47.3\% \times (37.7\%/38.6\%)$. The same calculations are continued to project the probability of receiving benefits at age seventy for cohorts A2004 through A2040. For cohorts younger than A2040 (not shown in the table) a 2 percent decline is assumed. This is an extrapolation of the decline for the cohorts A2031 to A2040.

From the percent that receives benefits at age seventy, we predict the percent that receives benefits at each age, as described in section 10.3. These estimates are shown for selected cohorts in figure 10.15.

10.6 Present Value of DB Benefits at Sixty-Five

We have projected the average level of the DB benefits of recipients, by age and cohort, B_{ac} , for cohorts attaining age sixty-five in the years 1982 through 2050. We have also projected the probability of benefit receipt for each age and cohort. First, we obtain for each cohort the average present value (PV) of benefits at age sixty-five for persons who *receive* DB benefits, given by:

(1)
$$PV(recipients)_{c,65} = \sum_{a=65}^{100} \left(\prod_{t=65}^{a} S_{t,c} \right) \left(\frac{B_{a,c}}{(1+r)^{(a-65)}} \right),$$

where $S_{t,c}$ is the cohort-specific probability of survival to age t conditional on being alive at t-1, and r is the discount rate. Second, we calculate the present value of DB benefits at age 65 for *all* cohort members, both those with and those without a DB plan. The expected average benefit received

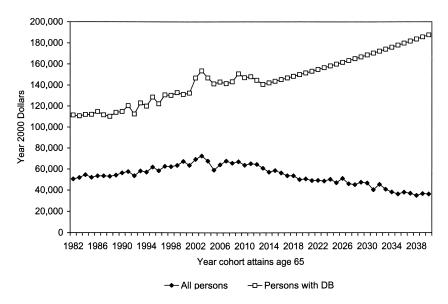


Fig. 10.16 Present value DB benefits at age 65: All persons and persons with a DB

by all persons of age a in cohort c is $\overline{B}_{ac} = \Pr[\text{benefit receipt}] \cdot [\text{benefits}_{ac}]$ benefit receipt] = $B_{ac} \cdot P_{ac}$. Thus the present value of benefits averaged over all persons is given by:

(2)
$$PV(\text{all persons})_{c,65} = \sum_{a=65}^{100} \left(\prod_{t=65}^{a} S_{t,c} \right) \left(\frac{B_{a,c} \cdot P_{ac}}{(1+r)^{(a-65)}} \right).$$

Figure 10.16 graphs both of these present value calculations. The present value amounts have been converted to constant year 2000 dollars. A real discount rate of 3 percent and average SSA age-specific survival probabilities are assumed. For these calculations, pension benefits received prior to age sixty-five have been ignored. The top profile shows that the average present value of DB benefits for persons that receive DB benefits will be greater for cohorts retiring in the future than for cohorts retiring today. This is because we assume that nominal benefits increase by 3.9 percent annually, and the inflation rate is 2.8 percent. However, the lower profile shows that the PV of DB wealth, averaged across all persons, will decline in the future as fewer persons participate. Indeed, these projections show DB wealth peaking in 2003 at about \$73,000 and falling to about \$50,000 by 2020.³

^{3.} These estimates indicate that in 2000 the average of DB benefits over all persons was \$67,386. Based on HRS data, Johnson, Burman and Kobes (2004), estimate that the mean present value of employer-sponsored pension income for persons sixty-five to sixty-nine in 2000 was \$50,203. Our estimate should be larger than theirs because we include persons of all ages. In particular, average benefits increase with age because death rates selectively leave in the sample persons with higher benefits and because of ad hoc cost-of-living increases and indexed benefits in many government plans.

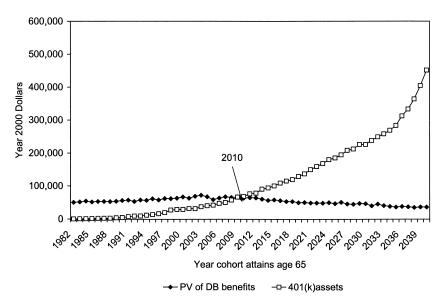


Fig. 10.17 Present value DB benefits at age 65 and 401(k) assets at age 65, all persons—Historical rates of return

For comparison, we show the average present value of 401(k) assets at age sixty-five, reported in Poterba, Venti and Wise (2008).⁴ The comparison between the average present value of DB benefits at age sixty-five and average 401(k) assets at age sixty-five is graphed in figures 10.17 and 10.18 for *all* persons. The comparisons in figures 10.19 and 10.20 pertain to persons *with* plans. Figures 10.17 and 10.19 show 401(k) assets at retirement assuming that the return on equities in the future will be equal to the historical average. Figures 10.18 and 10.20 show 401(k) assets assuming that the return on equities in the future will be 300 basis points less than historical average return.

Assuming historical rates of return, figure 10.17 shows that average 401(k) assets of *all persons* reach the average PV of DB benefits of all persons in 2010 when both are about \$67,000. Thereafter, assets in 401(k) accounts continue to grow, reaching about \$137,000 in 2020, \$226,000 in 2030, and \$452,000 in 2040. Assuming historical rates of return less 300 basis points, average 401(k) assets reach the average PV of DB benefits of all persons in 2011 when both are about \$66,000. Thereafter, assets in 401(k)

^{4.} These projections use actual annual pretax returns through 2005. Beginning in 2006, we assume that the average annual nominal return on equities is 12 percent and that the average nominal return on corporate bonds is 6 percent. Ibbotson Associates (2006) reports that the historical arithmetic mean of pretax returns on long-term corporate bonds has been 6.2 percent per year, while large-capitalization stocks have returned an average of 12.3 percent over the period 1926 to 2005. These returns are the pretax total return available on a portfolio with no management fees. We have not as yet accounted for asset management fees.

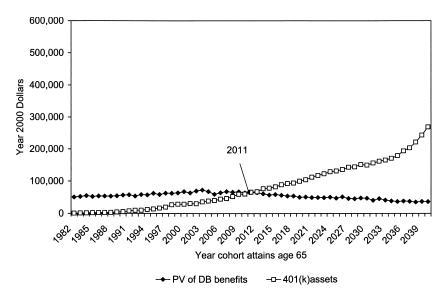


Fig. 10.18 Present value DB benefits at age 65 and 401(k) assets at age 65, all persons—Historical rates of return minus 300 basis points

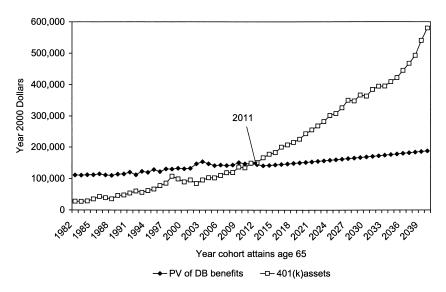


Fig. 10.19 Present value of DB benefits at age 65 for persons with a DB and 401(k) assets at age 65 for persons with a 401(k)—Historical rates of return

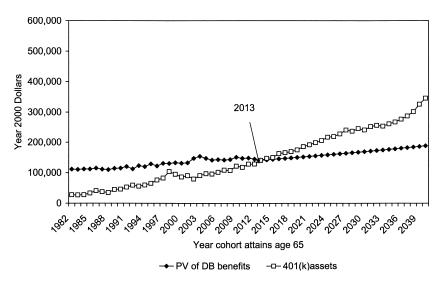


Fig. 10.20 Present value of DB benefits at age 65 for persons with a DB and 401(k) assets at age 65 for persons with a 401(k)—Historical rates of return minus 300 basis points

accounts continue to grow, reaching about \$104,000 in 2020, \$149,000 in 2030, and \$269,000 in 2040. The lower rate of return on equities substantially reduces the accumulation of assets in 401(k) plans. Even with the lower rate of return on equities, however, by 2040 the accumulation of assets in 401(k) plans would be 3.7 times as large as the historical maximum level of assets (fully funded) in DB plans, which was realized in 2003.

Figure 10.19 shows the present value of DB benefits for persons who have a DB plan and the 401(k) assets for persons who have a 401(k) plan, assuming the average historical equity return. For these persons, balances in 401(k) accounts reach the PV of DB benefits in 2011, when both are about \$148,000. Thereafter, the 401(k) assets continue to grow, reaching about \$243,000 in 2020, \$363,000 in 2030, and \$580,000 in 2040. For persons with DB plans, the average present value of benefits also continues to grow, reaching \$187,000 by 2040. By 2040, the accumulation of 401(k) assets is about 3 times the PV of DB assets. Figure 10.20 shows 401(k) assets assuming that the rate of return on equities is 300 basis points lower than the historical average. Here the assets of persons with 401(k) accounts reach the level of the PV of DB benefits for persons who receive benefits in 2014, when both are about \$145,000. Thereafter, the 401(k) assets continue to grow, reaching about \$185,000 in 2020, \$241,000 in 2030, and \$345,000 in 2040. For persons with DB plans, the average present value of benefits also continues to grow, reaching \$187,000 by 2040. Thus, for persons with plans, by 2040 the accumulation in 401(k) plans is about 1.8 times the PV of DB assets at sixty-five.

Our 401(k) asset projections do not account for legislated increases in the contribution limits between 2003 and 2007. The limit increases are large for all income groups, with the largest increases for persons with incomes between \$15,000 and \$20,000. Only a small proportion of persons are currently contributing at the maximum. However, as incomes increase, a larger and larger fraction of employees are likely to be contributing at the new limits. For this reason, it is likely that future contributions to 401(k) plans will be greater than our assumptions (projections) suggest. In addition, we have not accounted for the effects of the Pension Protection Act of 2006, which gives employers latitude to set more "saving friendly" defaults in 401(k) plans. Beshears et al. (2008) survey some of the recent evidence on how changing defaults for enrollment, contribution rates, and asset allocations can significantly increase retirement saving.

10.7 Assets in DB Trust Funds and Total DB Benefits Paid

Finally, we want to estimate the level of total assets in DB trust funds. These estimates are likely to be particularly important for assessing future changes in the demand for financial assets. We made similar calculations with respect to 401(k) and other personal retirement account assets in Poterba, Venti, and Wise (2008). We believe the DB component, together with the 401(k) component, represent a substantial fraction of the demography-induced change in asset demand.

There are at least two general ways to predict future assets in DB plans. One way is to predict total benefits paid in future years and then to suppose that assets in a year are sufficient to pay the present value of these future obligations. We take this approach here. We believe, however, that this approach should yield fully funded current assets that are greater than actual assets. A second way is to predict future assets based on an extrapolation of current assets compared to fully funded liabilities. A possible extrapolation procedure, for example, is to assume that assets will continue to be a given percent below the fully funded level. We have not pursued this approach here.

To obtain the present value of future obligations, we must first calculate total DB benefits paid each year. As described in the preceding, we have projected the average level of DB benefits of recipients by age and cohort (or, alternatively, by age and calendar year). We have also projected the conditional probability of receiving benefits for each age and year. We now combine these projections with demographic projections obtained from the Social Security Administration to obtain total DB benefits paid in each year. More precisely, we calculate the total dollar value of DB benefits paid to all persons in year *t* as:

(3)
$$DB_{t} = \sum_{a=55}^{100} N_{a,t} P_{a,t} B_{a,t},$$

where $N_{a,t}$ is the number of persons age a in year t, $P_{a,t}$ is the probability that benefits are received by a person age a in year t, and $P_{a,t}$ is the average benefit received (conditional on benefit receipt) by a person of age $P_{a,t}$ is the average benefit receipt) by a person of age $P_{a,t}$ is the average benefit receipt) by a person of age $P_{a,t}$ is the average benefit receipt.

Figure 10.21 shows these totals for the years 1982 to 2004, together with a constructed series that sums together private-sector DB benefits, benefits paid to federal employees, and benefits paid to state and local employees. The private-sector data are from the Form 5500 data and exclude benefits paid directly by insurance carriers. Federal DB benefits include payments made by the Civil Service Retirement System, the Federal Employees Retirement System, and the Military Service Retirement System. Defined benefit and DC benefits are not reported separately for state and local plans. The data used here include DC as well as DB benefits. Thus, our projected DB benefits should be somewhat smaller than sum of these reported government and private-sector benefits.

The projected totals are close to the constructed totals in the early years but are smaller than the constructed totals in the later years. The discrepancy is due in part at least to the growing importance of 401(k)-like accounts in the state and local government plans. Thus, we believe that the comparison lends credence to our estimates, based on SIPP data, for these years. Our forward projections depend on the assumptions we have made to project benefit levels and benefit receipt for future cohorts not represented in the SIPP data.

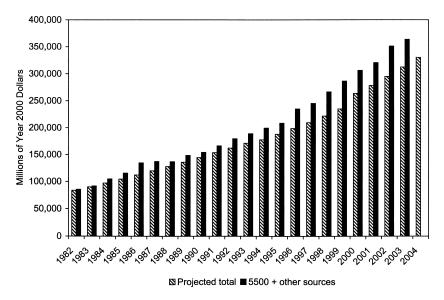


Fig. 10.21 DB benefits paid: Projected versus 5500 plus other sources, 1982 to 2004

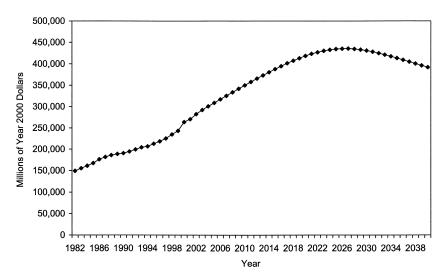


Fig. 10.22 Projected DB benefits paid: By year

Our projected estimates of total DB benefits in future years are shown in figure 10.22 in year 2000 dollars. Total benefits paid from DB plans continue to increase until 2027. The profile turns down eventually because the probability that benefits are received (at age seventy) reaches a maximum with the cohort that attains age sixty-five in 2007, although this cohort continues to receive benefits after 2007. From a peak of \$435,000 in 2027, the real value of total benefits paid declines modestly to \$392,000 by 2040. As noted in figures 10.17 to 10.20, the real level of benefits increases through the end of the projection period, so the decline in total benefits in figure 10.22 is driven by the decline in the probability of benefit receipt.

For comparison, projected benefits paid from DB plans are graphed against amounts withdrawn from 401(k) plans in figure 10.23.5 Benefits from DB plans exceed withdrawals of 401(k) assets until 2028, assuming historical rates of equity returns. After 2028, the value of DB benefits falls each year, and 401(k) withdrawals increase rapidly thereafter. If the future average rate of return on equities is assumed to be the average historical rate less 300 basis points, the benefits from DB plans exceed withdrawals of 401(k) assets until 2034.

We use our projections of benefits to be paid from DB plans to help to project future assets in DB trust funds. Suppose that in each year assets held by DB plan sponsors were equal to the present value of future obliga-

^{5.} The 401(k) withdrawals shown in this figure are projected amounts disbursed when account owners are alive and do not include balances that remain in accounts when account owners die.

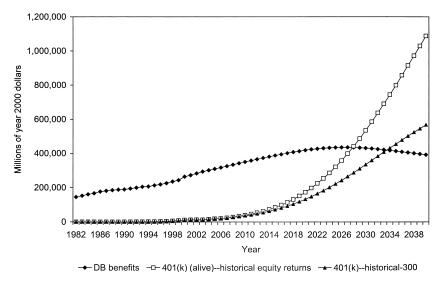


Fig. 10.23 DB benefits paid versus 401(k) withdrawals

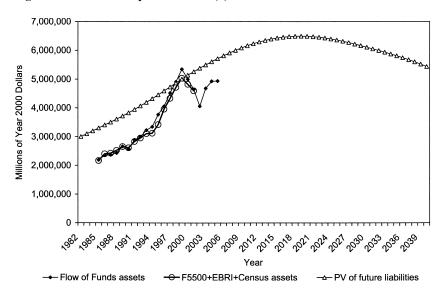


Fig. 10.24 Present value of DB liabilities versus DB assets

tions. If future liabilities are discounted at 3 percent and firms have a twenty-year planning horizon, the present value of liabilities each year is shown in figure 10.24. The figure also shows two series representing reported total assets in DB plans. The first is from the flow of funds accounts that include private-sector DB plans and the value of assets in DB and DC

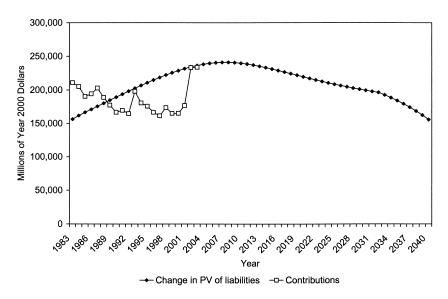


Fig. 10.25 Change in present value of future DB liabilities versus contributions

plans for government sponsors. The second series is composed of private-sector DB assets from Form 5500 reports, federal DB assets from various federal agency annual reports, and the sum of DB and DC assets for state and local governments from the Census Bureau's series on state and local government retirement systems. Actual assets, for the years they are available, are substantially below our calculation of the present value of liabilities in all years, with the possible exception of 1999. But the gap between actual assets and fully funded assets declined between 1985 and 1999. During this period, contributions to DB plans sometimes exceeded and sometimes fell short of the change in the present value of liabilities, as shown in figure 10.25. The gap closed in the late 1990s largely because of the stock market boom.

It is perhaps not surprising that actual assets are less than our estimates of the assets employers would have to hold to fully fund projected liabilities. Private-sector plan sponsors have substantial latitude in the assumption of interest rates, investment returns, when benefits will be paid, and other features that determine funding levels.⁶ The former director of the Pension Benefit Guaranty Corporation (PBGC), Bradley Belt (2005), estimated that private DB plans were underfunded by \$450 billion in 2004. There are even fewer restrictions on the funding of federal, state, and local plans, and many are thought to be substantially underfunded.

^{6.} Bergstresser, Desai, and Rauh (2006) discuss these issues in the context of earnings manipulation.

Our fully funded method yields assets that are well above actual assets in DB plans, at least in recent years. We know of no way to confidently predict future funding levels, however. Thus, we have projected future assets levels according to the fully funded method described in the preceding.

10.8 DB and 401(k) Assets Combined

What will be the change in asset levels for all pension plans—DB and 401(k) plans combined—in future years? The evolution of total assets depends on three components: contributions, withdrawals, and the internal buildup. In our prior work, we projected assets held in 401(k) plans. These 401(k) assets, DB plan liabilities (fully funded DB assets), and the sum of both series are shown in figure 10.26 and 10.27, assuming historical equity returns and historical returns minus 300 basis points, respectively. If equities in 401(k) plans earn the historical return, then 401(k) asset balances overtake DB balances in 2009. If equities earn 300 basis points less than the historical average, then 401(k) assets first exceed DB assets in 2010.

We also consider total contributions to DB and 401(k) plans and total withdrawals from DB and 401(k) plans. Figure 10.28 shows contributions to 401(k) plans, the annual change in DB liabilities (a rough measure of contributions to a fully funded DB plan), and the sum of these contributions. The sum is relatively flat between 2000 and 2040. In contrast, the sum of benefits paid grows rapidly through 2040, especially if historical equity rates of return are assumed. Withdrawals from 401(k) plans, DB benefits

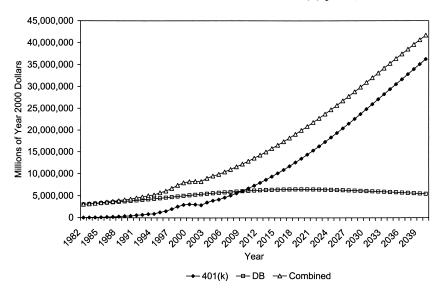


Fig. 10.26 Projected assets: 401(k), DB, and combined—Historical equity rates of return

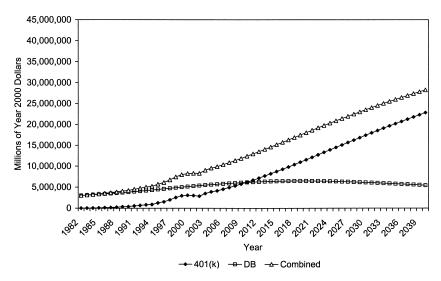


Fig. 10.27 Projected assets: 401(k), DB, and combined—Historical equity return minus 300 basis points

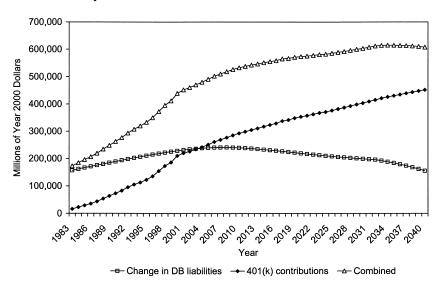


Fig. 10.28 Projected contributions: 401(k), DB (change in liabilities), and combined—Historical equity returns

paid, and the sum of the two are shown in figure 10.29 using historical equity returns and figure 10.30 using the historical return less 300 basis points.

Figure 10.31 shows the total projected contributions to and withdrawals from DB and 401(k) plans combined. Withdrawals assuming historical eq-

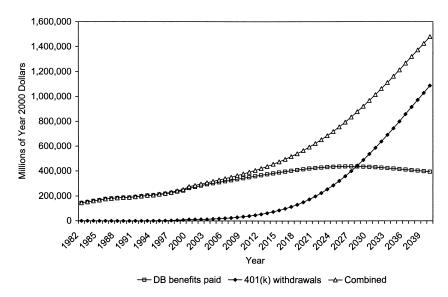


Fig. 10.29 Projected withdrawals: 401(k), DB (benefits paid), and combined—Historical equity returns

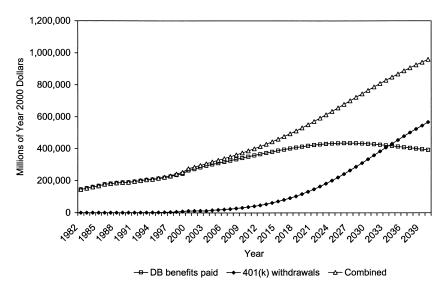


Fig. 10.30 Projected withdrawals: 401(k), DB (benefits paid), and combined—Historical equity returns minus 300 basis points

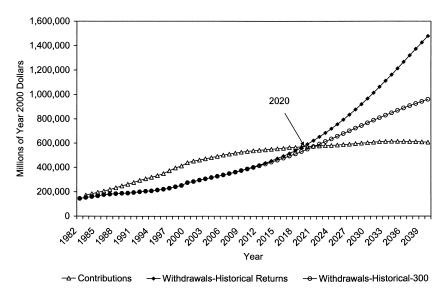


Fig. 10.31 Projected contributions and withdrawals: DB and 401(k) plans combined

uity rates of return and withdrawals assuming the historical return minus 300 basis points are shown in the figure. Withdrawals exceed contributions after 2020 and 2021 for these two withdrawal series, respectively. Figure 10.32 shows the combined withdrawals minus contributions. Again, withdrawals exceed contributions after 2020 if historical equity rates of return are assumed and after 2022 if historical rates minus 300 basis points are assumed. The excess of withdrawals over contributions reaches about \$872 billion by 2040 (in year 2000 dollars) when historical rates of return are assumed but only about \$353 billion if historical rates minus 300 basis points are assumed. Because of internal buildup, however, under either equity return scenario total assets in pension plans continue to grow through 2040. This is shown in figures 10.26 and 10.27.

10.9 Demographically Sensitive Assets and Rates of Return

Whether demographically induced changes in DB and DC assets will have an appreciable effect on the rates of return on equities and other fi-

7. Schieber and Shoven (1997) consider the implications of population aging for private pension fund saving and project saving as a percent of payrolls. Their projection method is very different from ours, and it does not focus on the rising importance of self-directed DC plans in the way that ours does. Nevertheless, their results are qualitatively similar to ours. They project that total private pension withdrawals will exceed contributions beginning in 2024, reversing the pattern of earlier years.

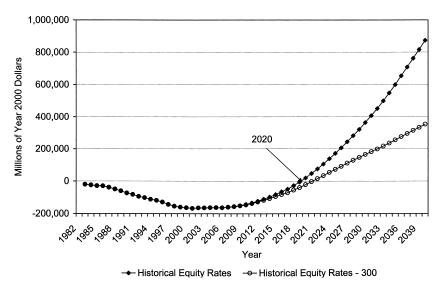


Fig. 10.32 Projected withdrawals minus contributions: DB and 401(k) plans combined

nancial assets depends on the magnitude of these changes relative to the other components of asset demand. Other changes in asset demand may reinforce or counterbalance them. Because asset markets are global, domestic demand for financial assets is not the only force determining the returns on stocks and bonds in the United States. Several recent studies have explored how demographic change in the currently developed world will affect global asset demand. Our estimates of the future demand for pension assets in the United States, if replicated in other countries, could be an important input to such studies. Whether international capital flows will reinforce or moderate the demographic pressures on asset returns that may arise from changes in the population age structure in the United States depends critically on the future rate of development of currently young economies and on the age structure of other developed nations.

When evaluating how demographic change affects the demand for financial assets in general, and specifically corporate equities, one must remember that a large fraction of financial assets are owned by a small fraction of the population. This group of high net worth investors is likely to be less sensitive to age-related changes in asset demand than other investors who have life-cycle motives for saving. Kennickell (2006) reports that in the 2004 Survey of Consumer Finances, the wealthiest 5 percent of households owned 65.6 percent of equities, including mutual funds, and 79.1 percent of equities, excluding mutual funds. Pension assets represent a smaller share of the assets of this group than of other less-affluent sectors of the population.

Retirement plan assets are one of the most, if not the most, demographically sensitive components of the household financial balance sheet. There are clear demographic effects on assets in DB plans, which are typically paid out as benefit annuities at the time of retirement. The link between demographic structure and 401(k) plan assets is less mechanical because older households have discretion over the rate at which they draw down assets in retirement. There is uncertainty both about the date at which withdrawals will begin and the rate of such withdrawals once they start. At present, many 401(k) participants do not begin to make withdrawals until they are required to do so at age 70.5. In the future, the average age of first withdrawal is likely to increase even for those who do not wait until 70.5 to begin withdrawals because the average age of retirement is likely to rise. This is likely to result both from the increase in the normal Social Security retirement age from sixty-five to sixty-seven by 2027, and from the conversion from DB plans, with strong incentives to leave the labor force early, to personal retirement accounts without these incentives. Longer working lives will probably delay withdrawals from personal retirement accounts and increase the accumulation of retirement assets.

For households that are constrained by the mandatory withdrawal requirements from personal retirement accounts, withdrawals of DC plan assets are likely to overstate the decline in asset demand at older ages. There is no requirement that households consume their minimum distributions, and households that would prefer not to make any withdrawals from their 401(k) plans or individual retirement accounts (IRAs) may simply reinvest the mandatory payouts in other investment options. It is possible that such households will bequeath a substantial portion of their tax-deferred assets to their heirs, who may continue to accumulate assets in a tax-deferred setting for many years. Recent legislation has facilitated such transfers of personal retirement account assets. While this consideration suggests that 401(k) assets may remain substantial even for households at very advanced ages, it is also possible that the financial burden of health care in retirement will be greater for future retirees than for past cohorts and that this will necessitate greater expenditures during retirement. If the cost of health care continues to increase, future retirees are likely to spend more on health care than current retirees. This could affect not only the accumulation of retirement assets, but could accelerate withdrawals from retirement accounts as well.

The many uncertainties that arise in trying to link one source of asset demand to equilibrium rates of return make us reluctant to attempt to quantify the impact of demographically induced changes in retirement asset accumulation on prospective stock or bond returns. There is also a fundamental circularity in this question. Our projections of future DC assets are based on assumptions about the rate of return that plan assets will earn in the next few decades. This underscores the importance of consid-

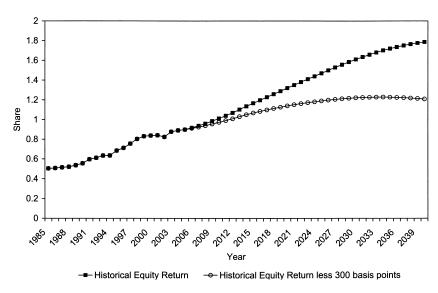


Fig. 10.33 Total pension assets as share of GDP—Historical return and historical return minus 300 basis points

ering the sensitivity of future retirement asset projections to rate of return assumptions. This concern notwithstanding, we have developed one metric to provide some indication of the quantitative importance of the demographically induced changes in retirement asset demand.

Figure 10.33 shows our projected demand for all pension plan assets, including both bonds and equities, as a percentage of the Social Security Administration intermediate gross domestic product (GDP) projection. The figure shows that pension assets grow from 50 percent of GDP in 1985 to 100 percent in 2010 to 179 percent in 2040. If the return on equity is 300 basis points lower than the historical return, then total pension assets grow to about 120 percent of GDP by 2040. Both of these projections suggest a very substantial increase retirement saving over the coming decades.

Figure 10.34 shows the projected demand for corporate equity in both DB and DC plans as a percentage of the Social Security Administration intermediate GDP projection: one based on the historical equity rate of return and the other based on the historical return less 300 basis points. Historically, pension equities grew from about 30 percent of GDP in 1985 to about 55 percent in 2005. Our projections suggest continued growth between now and 2040, to 142 percent of GDP when the projection is based on the historical equity return and 84 percent of GDP when the assumed rate of return is the historical value minus 300 basis points. The shape of

^{8.} The equity share is based on projected accumulated equity assets, which depends on the assumed rate of return on equities. We assume no rebalancing in 401(k) plans.

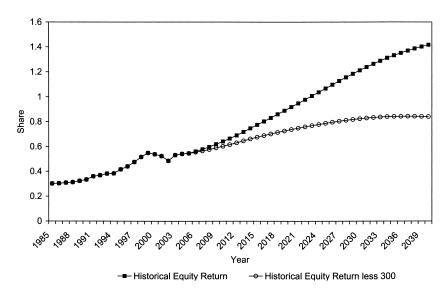


Fig. 10.34 Pension equity as share of GDP—Historical return and historical return minus 300 basis points

the projected growth path is also affected by the different rate of return assumptions. For example, the one-year growth rate of pension assets between 2039 and 2040 is 1.0 percent in the former case and -0.2 percent when we assume the lower equity return. The figure shows that even when we make the lower rate of return assumption, the stock of pension assets continues to increase for the next three decades, even relative to the aggregate economy. The projections do not suggest a sharp decline in pension assets, either in absolute dollars or relative to GDP, when the baby boom cohort reaches retirement age. If low returns depress the growth of 401(k) assets, however, there may be some years in which total retirement assets as a share of GDP are stable or decline slightly.

10.10 Summary

A key component to any effort to analyze how demographic trends may affect future returns on financial assets is a careful analysis of how these trends will affect assets in DB and DC pension plans. This chapter on DB pension plan assets and our earlier companion paper on 401(k) assets explore this issue. The dramatic decline in DB participation in the past three decades stands in contrast to the rapid expansion of 401(k)-like plans. We develop projections of the DB benefits of cohorts retiring between now and 2040 and use them to project the total assets held in DB plans through

2040. The projections are based on extrapolation of cohort data from many waves of the SIPP along with demographic projections from the Social Security Administration.

We project the present value of DB benefits at age sixty-five for cohorts who reach age sixty-five in each year from 1982 to 2040, and we compare these projections to projections of 401(k) assets. Our projections suggest that the average (over *all persons*) of the present value of real DB benefits at age sixty-five attained an historical maximum in 2003, when the value was \$72,637. Our projections also suggest that the average value of 401(k) assets at age sixty-five surpasses the average present value of DB benefits at age sixty-five in about 2010. Thereafter, the value of 401(k) assets grows rapidly, attaining levels much greater than the maximum present value of DB benefits. If equity returns between 2006 and 2040 are comparable to those observed historically, by 2040 average projected 401(k) assets will be over six times larger than the historical maximum level of DB benefits at age sixty-five, attained in 2003. Even if equity returns average 300 basis points below their historical value, we project that average 401(k) assets in 2040 would be 3.7 times as large as the value of DB benefits in 2003.

The projected growth of real 401(k) assets more than offsets the projected decline in real DB plan assets during the next three decades. Focusing on DB assets alone suggests that an aging population, in conjunction with a shift away from DB plans, will lead to a decline in the real value of pension assets averaged across all retirees in future cohorts. When we combine projected 401(k) assets with projected DB assets, however, we find that real pension assets not only increase, but increase substantially, in future decades.

Our findings underscore the need for further analysis of the factors that determine the diffusion of 401(k) plans across corporations, especially small companies with low-wage workers, as well as the contribution behavior and withdrawal behavior of 401(k) participants. The growing role of 401(k)-type plans in the retirement landscape suggests that understanding asset accumulation and draw-down in these plans is a critical component of any analysis of the effect of demographic change on financial markets.

Appendix

Table 10A.1 Regression estimates for the level of pension benefits

Cohort		Parameter estimates		
Age in 1984	Year age 65	Coefficient	error	t-statistic
86	1963	41.84	59.67	0.70
85	1964	-22.90	75.76	-0.30
83	1966	-42.08	50.39	-0.84
82	1967	27.12	55.60	0.49
81	1968	110.47	62.20	1.78
80	1969	55.35	68.64	0.81
79	1970	132.11	59.10	2.24
78	1971	206.22	55.86	3.69
77	1972	156.01	53.63	2.91
76	1973	151.60	50.57	3.00
75	1974	171.49	51.14	3.35
74	1975	246.42	48.08	5.12
73	1976	232.83	51.14	4.55
72	1977	264.87	48.86	5.42
71	1978	312.69	51.26	6.10
70	1979	328.91	50.51	6.51
69	1980	352.85	49.27	7.16
68	1981	410.67	49.83	8.24
67	1982	479.79	48.30	9.93
66	1983	496.66	49.11	10.11
65	1984	526.01	49.72	10.58
64	1985	548.35	50.30	10.90
63	1986	589.81	49.88	11.83
62	1987	592.74	49.54	11.97
61	1988	607.28	50.00	12.15
60	1989	657.08	50.77	12.94
59	1990	685.37	51.76	13.24
58	1991	748.48	53.73	13.93
57	1992	710.96	51.80	13.73
56	1993	815.37	53.83	15.15
55	1994	815.87	53.67	15.20
54	1995	911.54	55.21	16.51
53	1996	883.43	55.71	15.86
52	1997	983.28	58.73	16.74
51	1998	1,009.20	60.24	16.75
50	1999	1,063.51	60.98	17.44
49	2000	1,079.51	61.99	17.41
48	2001	1,124.29	63.02	17.84
47	2001	1,299.43	71.18	18.25
46	2002	1,407.14	71.53	19.67
45	2003	1,381.55	77.27	17.88
44	2004	1,364.24	75.67	18.03
43	2003	1,425.66	84.44	16.88
43	2000	1,423.00	04.44	10.88

Table 10A.1 (continued)

		Parameter estimates		5
Cohort Age in 1984 Year age 65		Coefficient	Standard error	t-statistic
	1001 050			, statistic
42	2007	1,452.48	83.95	17.30
41	2008	1,516.71	84.13	18.03
40	2009	1,652.97	114.61	14.42
39	2010	1,658.59	116.33	14.26
38	2011	1,720.02	123.20	13.96
37	2012	1,729.32	127.80	13.53
Intercept Age		-131.08	54.89	-2.39
55–60		52.53	7.38	7.11
60-65		-13.07	5.73	-2.28
65–70		10.55	4.29	2.46
70–75		17.93	4.20	4.27
75-80		16.91	4.85	3.49
> 80		0.92	5.84	0.16
No. of observations	32,388			
<i>F</i> (56, 32331)	69.53			
R^2	0.1112			
Root MSE	802.96			

Note: MSE = mean square error.

Table 10A.2 Probit estimates for receipt of pension benefits

		Parameter estimates		
Age in 1984	Year age 65	Coefficient	Standard error	t-statistic
80	1969	0.20	0.14	1.43
79	1970	0.14	0.13	1.11
78	1971	0.17	0.12	1.40
77	1972	0.13	0.12	1.09
76	1973	0.19	0.12	1.60
75	1974	0.20	0.12	1.65
74	1975	0.10	0.12	0.83
73	1976	0.21	0.12	1.84
72	1977	0.29	0.11	2.66
71	1978	0.34	0.11	3.04
70	1979	0.24	0.11	2.17
69	1980	0.32	0.11	2.93
68	1981	0.37	0.11	3.38
67	1982	0.41	0.11	3.77
66	1983	0.45	0.11	4.17
65	1984	0.40	0.11	3.64
				(continued)

Table 10A.2 (continued)

Cahant		Parameter estimates		
Age in 1984	Year age 65	Coefficient	Standard error	t-statistic
64	1985	0.40	0.11	3.68
63	1986	0.43	0.11	3.98
62	1987	0.44	0.11	4.09
61	1988	0.42	0.11	3.92
60	1989	0.46	0.11	4.28
59	1990	0.43	0.11	3.98
58	1991	0.43	0.11	3.95
57	1992	0.42	0.11	3.85
56	1993	0.42	0.11	3.91
55	1994	0.44	0.11	4.01
54	1995	0.43	0.11	3.94
53	1996	0.43	0.11	3.95
52	1997	0.43	0.11	3.91
51	1998	0.43	0.11	3.91
50	1999	0.52	0.11	4.71
49	2000	0.43	0.11	3.94
48	2001	0.41	0.11	3.74
47	2002	0.41	0.11	3.73
46	2003	0.46	0.11	4.07
45	2004	0.49	0.11	4.36
44	2005	0.46	0.11	4.00
43	2006	0.55	0.11	4.80
42	2007	0.51	0.11	4.51
41	2008	0.58	0.12	4.78
40	2009	0.60	0.12	4.95
39	2010	0.54	0.12	4.41
38	2011	0.61	0.12	4.94
37	2012	0.47	0.12	3.82
Intercept Age		-6.94	0.38	-18.31
55-60		0.10	0.01	15.45
60-65		0.11	0.00	23.08
65–70		0.03	0.00	5.25
70–75		0.01	0.01	1.78
75–80		0.00	0.01	0.01
No. of observations	78,686			
Pseudo R ²	0.0593			
Wald χ^2 (49)	4,943.32			

Table 10A.3 Probit estimates for participation using population base

Cohort		Parameter estimates		
Cor	nort		Standard	
Age in 1984	Year age 65	Coefficient	error	t-statistic
65	1984	-0.143	0.122	-1.17
64	1985	-0.064	0.107	-0.59
63	1986	-0.083	0.105	-0.79
62	1987	0.046	0.092	0.51
60	1989	0.056	0.091	0.62
59	1990	0.040	0.091	0.44
58	1991	0.010	0.094	0.11
57	1992	0.139	0.094	1.48
56	1993	-0.040	0.098	-0.40
55	1994	-0.025	0.095	-0.26
54	1995	-0.031	0.096	-0.32
53	1996	-0.247	0.096	-2.56
52	1997	-0.160	0.096	-1.67
51	1998	-0.177	0.090	-1.98
50	1999	-0.227	0.087	-2.62
49	2000	-0.249	0.084	-2.95
48	2001	-0.124	0.081	-1.52
47	2002	-0.173	0.080	-2.17
46	2003	-0.198	0.078	-2.52
45	2004	-0.290	0.079	-3.65
44	2005	-0.225	0.078	-2.88
43	2006	-0.160	0.078	-2.07
42	2007	-0.202	0.077	-2.63
41	2007	-0.232	0.078	-2.99
40	2009	-0.257	0.078	-3.31
39	2010	-0.237 -0.241	0.078	-3.31 -3.09
38	2010	-0.241 -0.230	0.077	-3.09 -2.97
37	2011	-0.259	0.078	-2.97 -3.31
36	2012	-0.239 -0.327	0.078	-3.31 -4.18
35	2013	-0.327 -0.313	0.078	-4.18 -4.02
34	2014	-0.313 -0.358	0.078	-4.02 -4.62
33	2015	-0.338 -0.410	0.077	-4.62 -5.29
32				-5.29 -5.44
	2017	-0.419	0.077	
31	2018	-0.486	0.077	-6.31
30	2019	-0.485	0.077	-6.31
29	2020	-0.519	0.077	-6.76
28	2021	-0.525	0.077	-6.84
27	2022	-0.543	0.077	-7.05
26	2023	-0.526	0.077	-6.82
25	2024	-0.586	0.078	-7.47
24	2025	-0.530	0.078	-6.76
23	2026	-0.618	0.079	-7.81
22	2027	-0.640	0.080	-8.03
21	2028	-0.610	0.079	-7.68
20	2029	-0.631	0.080	-7.91
				(continued)

Table 10A.3 (cor	itinued)
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		Parameter estimates		3
Cohort		Standard		
Age in 1984	Year age 65	Coefficient	error	t-statistic
19	2030	-0.743	0.081	-9.16
18	2031	-0.667	0.081	-8.24
17	2032	-0.750	0.082	-9.20
16	2033	-0.803	0.082	-9.79
15	2034	-0.842	0.082	-10.27
14	2035	-0.821	0.082	-10.02
13	2036	-0.846	0.083	-10.19
12	2037	-0.889	0.085	-10.51
11	2038	-0.866	0.092	-9.38
10	2039	-0.756	0.092	-8.23
9	2040	-0.888	0.096	-9.28
Intercept Age		-0.519	0.229	-2.27
<30		0.013	0.008	1.75
30-35		-0.004	0.006	-0.67
35-40		-0.006	0.006	-1.15
40-45		0.004	0.006	0.76
45-50		-0.015	0.006	-2.59
50-55		-0.048	0.007	-7.29
55-60		-0.056	0.008	-6.64
>60		-0.111	0.015	-7.26
No. of observations	216,969			
Pseudo R ²	0.021			
Wald χ^2 (64)	1,930.12			

References

Belt, Bradley D. 2005. Testimony before the Committee on Education and the Workforce, United States House of Representatives, March 2.

Bergstresser Daniel, Mihir A. Desai, and Joshua Rauh. 2006. Earnings manipulation and managerial investment decisions: Evidence from sponsored pension plans. *Quarterly Journal of Economics* 121 (1): 157–95.

Beshears, John, James Choi, David Laibson, and Brigitte Madrian. 2008. The importance of default options for retirement saving outcomes: Evidence from the United States. In *Lessons from pension reform in the Americas*, ed. Stephen J. Kay and Tapen Sinha, 59–87. Oxford, UK: Oxford University Press.

Boersch-Supan, Axel, Alexander Ludwig, and Joachim Winter. 2005. Aging, pension reform, and capital flows: A multi-country simulation model. NBER Working Paper no. 11850. Cambridge, MA: National Bureau of Economic Research.

Brooks, Robin. 2002. Asset market effects of the baby boom and Social Security reform. *American Economic Review* 92 (May): 402–06.

Geanakoplos, John, M. Michael Magill, and Martine Quinzii. 2004. Demography

- and the long-run predictability of the stock market. *Brookings Papers on Economic Activity*, Issue no. 1:241–325.
- Ibbotson Associates. 2006. Stocks, bonds, bills, and inflation, 2006 yearbook. Chicago: Ibbotson Associates.
- Johnson, Richard, Leonard Burman, and Deborah Kobes. 2004. *Annuitized wealth at older ages: Evidence from the Health and Retirement Study.* Final Report to the Employee Benefits Security Administration, U.S. Department of Labor. Washington, DC: Urban Institute.
- Kennickell, Arthur B. 2006. Currents and undercurrents: Changes in the distribution of wealth, 1989–2004. Federal Reserve Board of Governors, Working Paper.
- Krueger, Dirk, and Alexander Ludwig. 2006. On the consequences of demographic change for rates of return to capital and the distribution of wealth and welfare. Goethe University, Frankfurt. Mimeograph.
- Poterba, James M. 2005. The impact of population aging on financial markets in developed countries. In Global demographic Change: Economic Impact and challenges, ed. Gordon H. Sellor, 163–216. Kansas City, MO: Federal Reserve Bank of Kansas City.
- Poterba, James M., Steven F. Venti, and David A. Wise. 2008. New estimates of the future path of 401(k) assets. *Tax Policy and the Economy* 22:43–80.
- Schieber, Sylvester, and John Shoven. 1997. The consequences of population aging for private pension fund saving and asset markets. In *The economic effects of aging in the United States and Japan*, ed. M. Hurd and N. Yashiro. 111–30. Chicago: University of Chicago Press.
- U.S. Government Accountability Office. 2006. Retirement of baby boomers is unlikely to precipitate dramatic decline in market returns, but broader risks threaten retirement security. Report no. GAO-06-718. Washington, DC: GPO, July.
- Wiatrowski, William. 2005. Documenting benefits coverage for all workers. USDOL/BLS, December. http://www.bls.gov/opub/cwc/cm20040518ar01p1.htm.

Comment Jonathan Skinner

Everyone knows that the United States is embarking on a fundamental demographic shift as the baby boomers age, but there's less agreement on how it will affect the financial security of future retirees. James M. Poterba, Steven F. Venti, and David A. Wise (2007b) have provided some critical answers to this larger question by charting the course of defined benefit plans and their future inflows and outflows. The chapter is remarkable not because the results are shocking—indeed, they appear quite reasonable—but because of the incredible attention to detail in building up from the microlevel patterns of data to aggregate predictions. By harnessing millions of individual-level observations from a variety of sources and years, they not only provide a solid foundation for the aggregate estimates, but they also

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