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Sustaining Impacts When Transfers End Women Leaders, Aspirations, and Investments in Children

Karen Macours and Renos Vakis

9.1 Introduction

The intergenerational transmission of poverty often occurs through low levels of investment in education and nutrition. Conditional cash transfer (CCT) programs and many other development interventions specifically aim to increase human capital investment by the poor. A large body of evidence shows that CCT programs have been successful in augmenting investment in education and nutrition in many settings (Fiszbein and Schady 2009; Ganimian and Murnane 2014). A key question is whether CCTs can have lasting impacts on investment behavior after households stop receiving transfers. Only a few papers study whether the impacts on households' human capital investments persist after such programs end, and the evidence is mixed (Macours, Schady, and Vakis 2012; Baird, McIntosh, and Özler 2016). Even

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less is known about the possible mechanisms underlying persistence. A better understanding is needed to derive lessons regarding optimal design of new programs and adjustments to existing ones. More generally, knowing whether and how short-term programs can result in long-term increases in human capital investment is important for policy design.

For programs to have a persistent effect on households' human capital investments, they need either to permanently lift existing liquidity constraints or to change the value households attribute to investments in education and nutrition.¹ The latter may occur if the interventions increase the perceived returns to such investments by reducing information asymmetries or changing preferences. Nguyen (2008) and Jensen (2010) show that changes in the perceived returns to schooling through information can lead to educational gains. Recent evidence also suggests the potential of external interventions to shift preferences by changing parents' aspirations for their children (Beaman et al. 2012; Bernard et al. 2014).²

Understanding how to design interventions to maximize such shifts hence becomes an important policy question. Several design features of CCT programs could be contributing to shifts in investment behavior. The CCTs typically include heavy social marketing and conditionalities enforcing attendance at regular meetings in which the nutritional, health, and educational objectives are discussed. To the extent that such messages get internalized, one could expect increased human capital investments to persist. Targeting the transfers to women in the household could shift gender norms regarding decision-making within the household, and this too could persist after the transfers stop. Often programs also assign specific roles to key women in the community to reinforce the messages, but causal evidence on their specific role is rare.³

This chapter shows that interactions with local female leaders can contribute to the persistence of a program's impacts by providing evidence for a CCT pilot program in Nicaragua. It builds on Macours and Vakis (2014), where we showed that exposure to successful and motivated female leaders substantially increased impacts on nutritional and educational investments, as well as future-oriented attitudes, while the program was operating. This chapter analyzes whether these shifts were sustained after the program ended. A priori the answer is not obvious. Increasing aspirations in the presence of many other remaining constraints may lead to only

^{1.} If some decisions are driven by habits, a program that changes habits can also have a persistent effect.

^{2.} External interventions can also change the aspirations of children themselves (Wydick, Glewwe, and Rutledge 2013) or aspirations of adults for themselves (Lybbert and Wydick 2016).

^{3.} The importance of these local female leaders has been recognized in several qualitative evaluations of CCT programs (Adato 2000; Adato and Roopnaraine 2004). In Colombia, an independent ECD intervention specifically targeted the "madre voluntarias" of the CCT program in recognition of their local leadership role (Attanasio et al. 2014).

short-term gains, and households could revert back to preprogram behavior when the transfers stop. On the other hand, if social interactions during the program changed norms and beliefs regarding human capital investments, the increased investment levels could persist even after the end of the program. Macours, Schady, and Vakis (2012) show that the Nicaraguan CCT indeed had persistent effects on parental investments in early childhood. This chapter helps explain why.

Using data collected two years after the program ended, we show that social interactions with successful and motivated leaders were crucial for the persistence of the educational and nutritional investment. Two years after the transfers stopped, former beneficiaries who live in the proximity of such leaders still show significantly higher investments in both education and nutrition of their children. Random exposure to successful leaders also led to significant shifts in parental aspirations and expectations for their children's future.⁴

We use a two-stage randomized design to identify the social interaction effects. The program combined a regular CCT with interventions aimed at increasing households' productive potential. Because it targeted the vast majority of households in each community and explicitly encouraged group formation, it is a good setting to analyze the role of social interactions. The experiment varied the nature and the size of the benefit packages leaders and other households received, and as such created random variation in whether beneficiaries lived close to leaders that received the largest package. In general all leaders have higher human capital investments and aspirations than other beneficiaries, and hence provide potential examples to follow. The leaders that received the largest package, in addition, outperformed other leaders in terms of economic outcomes and also had higher expectations for their children's future. Earlier findings also showed that leaders with the largest package communicated more with other beneficiaries during the intervention.⁵ We analyze whether proximity to these successful and motivated leaders affected human capital investments of other beneficiaries.

We follow Manski (2000) and define social interactions as interactions with agents—in this case leaders with the largest package—that affect actions of other agents through changing constraints, expectations, or preferences. We provide empirical evidence in support of those different channels by

4. We draw on a rich set of questions measuring parental expectations and aspirations. As in Beaman et al. (2012) and Bernard et al. (2014) we measure aspirations by asking parents about what they would like their children to achieve on a number of dimensions, such as the desired education level or occupation. As in the later paper, we also separately measure expectations by asking parents what they think their children realistically will achieve on those same dimensions. Both set of indicators more broadly capture future-oriented attitudes.

5. More specifically, distance between houses generally reduces communication between leaders and beneficiaries (as expected), but this was not the case for leaders with the largest package. The differences between leaders with different packages were significant (Macours and Vakis 2014).

exploiting the random variation in the type of package each of the nonleaders received and the variation in per capita expenditure levels that resulted two years after the intervention. Beneficiaries of the largest package—on average—still had higher per capita expenditures two years after the end of the transfers, and more so when they were exposed to leaders with the same package. Such effects do not exist for beneficiaries of the other packages. Hence while changes for the first group can be driven in part by relaxing spending constraints, changes for the other groups are more likely driven by shifts in expectations or aspirations (and hence preferences).

The chapter contributes to the developing literature in economics and the wider social sciences on the role and the formation of aspirations (Genicot and Ray 2014; Besley 2017). Appadurai (2004) and Ray (2006) argue that upward mobility might be difficult for the poor when they lack the capacity to aspire, that is, when their own experiences and the experiences of those that are close to them suggest that escaping poverty is not a feasible option. Yet learning about the positive experiences of others that are sufficiently "close" may help open their "aspiration window." Hence social interactions may be instrumental in changing aspirations and shaping positive attitudes toward the future, and in turn lead to investments in their children's future.⁶ Empirical evidence of such mechanisms is rare due to the "reflection problem." This chapter addresses the problem through the randomized assignment of leaders and other beneficiaries to different benefit packages.

More broadly, this chapter relates to recent work on the potential of social interactions to shift norms and behavior (Paluck and Shepherd 2012; Feigenberg, Field, and Pande 2013) and to the emerging literature about mental models and attitudinal changes (Jensen and Oster 2009; La Ferrara, Chong, and Duryea 2012; World Bank 2015; Hoff and Stiglitz 2016). By focusing on local female leaders, the chapter also relates to the literature on female reservations for local leadership positions in India (Chattopadhyay and Duflo 2004; Beaman et al. 2009), and in particular to Beaman et al. (2012), who show that a law reserving leadership positions for women affected girls' educational aspirations.

Finally, this study relates to the growing literature on longer-term impacts of CCT programs (see Molina-Millan et al. [2016] for a review). We contribute by studying the impacts of a one-year randomized pilot program after it ended, and for which the experimental control group was never phased in. This allows providing clean evidence of the sustainability of impacts on human capital investments after only a few years, avoiding selection (attrition) concerns that often hamper long-term studies. That said, because the intervention only lasted one year, it differs from many of the large CCT programs

^{6.} Appadurai (2004) describes how mobilization by social movements can expand the capacity to aspire, in part through regular social gatherings and sharing ideas and experiences about future-oriented activities among the poor.

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in Latin America, where beneficiaries often receive transfers for many years.⁷ We return to this point in the conclusion. By focusing on the impact on human capital investment, we complement other studies that analyze whether the impacts on human capital outcomes (as opposed to investments) or other welfare outcomes persist on the longer run. While some studies analyze long-term impacts of ongoing programs (Behrman, Parker, and Todd 2009, 2011; Gertler, Martinez, and Rubio-Codina 2012; Araujo, Bosch, and Schady 2016), others, like us, provide evidence on programs with short duration (Barham, Macours, and Maluccio 2013a, 2013b; Macours, Premand, and Vakis 2012; Barrera-Osorio, Linden, and Saavedra 2015; Filmer and Schady 2014).

The chapter is organized as follows: in the next section we discuss the program and the relevance of social interactions. Section 9.3 discusses the data and the empirical strategy. Section 9.4 shows that social interactions with successful leaders had persistent impacts on other beneficiaries' human capital investments. Section 9.5 shows results for per capita expenditures, parental expectations, and aspirations; section 9.6 concludes.

9.2 Program Information and Design

9.2.1 Program Description and Treatment Packages⁸

The Atención a Crisis program was a one-year pilot program implemented in 2006 by the Ministry of the Family in Nicaragua. In the treatment communities, three different treatments were randomly allocated among 3,000 eligible households. All selected households were eligible for the basic CCT, which included cash transfers conditional on children's primary school attendance and health center visits. The transfers came with a strong social marketing message reinforcing the importance of investing in children's education and in a diversified diet. Take-up of the CCT was 95 percent. In addition to the CCT, one-third of the eligible households received a scholarship for a vocational training for one adult (with take-up of 89 percent). Another third of eligible households received, in addition to the basic CCT, a US\$200 lump-sum grant to invest in a small nonagricultural business (with take-up of 99 percent). This last treatment was perceived by the beneficiaries as the most attractive and involved the largest cash amount. We call it the "largest package." Given the high take-up rates, we henceforth refer to eligible households in treatment communities as beneficiaries.

The program design aimed to change households' investment behavior through several mechanisms. The level of cash transfers was substantial, rang-

^{7.} In the large national CCT programs in Colombia or Mexico, for instance, beneficiaries only exit when their children reach a certain age or after households reach a higher income level.

^{8.} More details about the program are provided in the online appendixes of Macours, Schady, and Vakis (2012) and Macours and Vakis (2014), as well as the following website: http://go.worldbank.org/VUYJAQ3UN0.

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ing from 18 percent of average annual household income for those receiving the basic CCT package to 34 percent for those receiving the productive investment package. The conditionalities and social marketing on education, health, and nutrition aimed at changing households' perspectives about investment in long-term human capital. The program design also created many opportunities for enhanced communication between beneficiaries. More than 90 percent of the households in treatment communities were eligible for the program, increasing the opportunities for information sharing, possibly resulting in higher motivation and program ownership. Beneficiaries were also required to participate in local events ranging from discussions on nutrition and health practices to workshops on the importance of education, business development, and labor market skills. The program put in place a system of volunteer local promotoras to enhance information flows and compliance with program requirements. The promotoras met frequently with small groups of beneficiary women to talk about these requirements and the program's objectives. As such, the program created a lot of new leadership positions for women.⁹ Women self-selected into these positions and then subsequently were randomly allocated to one of the three program packages (see below). Interviews during and after the intervention showed that most of the promotoras had taken strong ownership of the messages and objectives of the program, and were committed to reminding other beneficiaries that the purpose of the cash transfers was to invest in the nutrition and education of their children. During payment days, for instance, promotoras would often organize with the beneficiaries in their group to collectively buy food products and material for their children. Among other things, this allowed beneficiaries to directly observe investments by their promotoras. Qualitative evidence further confirms that beneficiaries were very aware of investments by others, with plenty of stories about children in the village going to school well fed, with new cloths and material.

9.2.2 Program Randomization

The program targeted six municipalities in the northwest of Nicaragua, and a first lottery randomly selected fifty-six intervention and fifty control communities. Baseline data were used to define household program eligibility using proxy means methods for both treatment and control.¹⁰ In the treatment communities, the main female caregiver from each eligible household was invited to a registration assembly. If there were more than thirty eligible households in a community, several assemblies were organized at the same time, and households were assigned to one of the assemblies based on the geographic location of their house. In total, there were 134 assemblies (hence, on average, 2.4 per community).

^{9.} Before the program, leadership positions for women were limited mostly to positions as teachers and health coordinators.

^{10.} As more than 90 percent of all households were eligible, the analysis in this chapter is limited to the eligible households.

During the assemblies, the program objectives and its various components were explained and women were asked to volunteer for the promotora positions. Volunteers were approved by the assembly and each promotora became responsible for a group of approximately ten beneficiaries living close to her, with *promotoras* and beneficiaries mutually agreeing on the compositions of the group. After the groups were formed, and at the very end of each assembly, all the beneficiaries-including the promotoras-participated in a second lottery process through which the three packages described above were randomly allocated among the beneficiaries, with each of the three packages assigned to one-third of households in the treatment communities. As a result of the two lotteries, households were randomly assigned to the control group (in the control communities), or to one of three packages: the CCT, the CCT plus training, or the CCT plus productive investment grant (the largest package). Since promotoras and existing female leaders in the treatment communities were randomly allocated to one of the three treatment groups, beneficiary households were randomly exposed to leaders with a treatment package that could be different from theirs. In particular, as there are on average four leaders in each assembly, some beneficiaries will randomly live close to several leaders that got the largest package, while others may not have any leaders with that package in their registration assembly.¹¹ This is the main exogenous variation that we exploit.

9.3 Data and Empirical Strategy

9.3.1 Data

In treatment communities, data were collected from all households. In control communities, a random sample of households was selected at baseline so that the control group was of equal size as each of the three intervention groups (1,000 households). The data analyzed in this chapter was collected between August 2008 and May 2009, approximately two years after the last transfer. Individuals who had migrated out of the area were tracked to different locations in Nicaragua, resulting in a very low attrition rate (3 percent at the household level), which is uncorrelated with treatment.

The survey instrument was modeled after the Nicaraguan Living Standard Measurement Survey (LSMS), with modules on education, health, and detailed household expenditures, among others. For the main set of results, we use the same education and nutrition investment indicators as those used in Macours and Vakis (2014). Specifically, for child-level education

^{11.} While the meetings of the *promotoras* with their groups were, by design, more frequent than meetings with the larger group of beneficiaries of an assembly, we use the larger assemblies as the reference group in part because the administrative information on the composition of the small groups is less precise than the information on who participated in which assemblies. In addition, it is possible that beneficiaries reorganized the groups after the assemblies, so that the effective groups may not correspond to the administrative data on groups.

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outcomes, we consider all children between seven and eighteen years old, and use an indicator of whether the child was attending school, the number of days the child has been absent from school in the last month, and the amount spent on school expenditures since the start of the academic year. Nutrition investment is measured at the household level and is measured by the shares of food expenditures for animal products and for vegetables and fruits, reflecting the emphasis of the program's messaging on the importance of such nutrients for children. To account for multiple hypotheses testing, we also combine the education and nutrition variables in two aggregate indices, by first calculating *z*-scores for each variable using the mean and standard deviation of the control group, and then averaging over these *z*-scores, following Kling, Liebman, and Katz (2007).

We complement this analysis with indicators of investments in children from birth to seven years old, using the same indicators of investment in early childhood as Macours, Schady, and Vakis (2012). We analyze impacts on three families of outcomes by calculating average *z*-scores of a set of indicators for nutrition, education, and health.¹² These are the three early childhood risk factors for which investments on average were still higher in the treatment than in the control, two years after the end of the transfers. As this is an age group that is not yet in primary school, the stimulation index can be seen as the equivalent of the education index for the older children. We use these indices to specifically analyze social multiplier effects on investment in children born after the end of the transfer (and hence approximately from birth to two years old). This allows testing whether the change in investment behavior is also observed for children not directly affected by the intervention, which provides a strong test of a more permanent shift in investment behavior.

A specific module was added in 2008 to ask mothers about expectations and aspirations for all their children between seven and fifteen years old.¹³ Mothers were asked both what they desired (to measure aspirations) and what they realistically expected for their children in terms of final educational attainment, occupation, future monthly earnings, and living standards. To proxy for future living standards, we also asked mothers for the number of rooms they desired and expected for the house their children would live in in thirty years' time. For occupation, we consider two possible definitions. The first is a dummy indicating whether the mother expected or desired a professional

12. The nutrition index is the average of the *z*-scores for the share of food in total expenditures, the shares of animal proteins and of fruit and vegetables in total food expenditures, and the reverse of the share of staples in total food expenditures; the stimulation index is the average of the *z*-scores for variables indicating whether the household has pen and paper, has a toy, somebody tells stories or sings to the child, and the number of hours reading to the child per week; and the health index is the average of the *z*-scores for variables indicating whether the child was weighed, got vitamins or iron, got deworming drugs, and the number of days sick in bed. See Macours, Schady, and Vakis (2012) for detailed definitions.

13. The module was not asked for children younger than seven, as mothers demonstrated difficulties answering such questions for their young children during piloting.

job for her child, that is, a job for which university education is required. The second is a dummy indicating whether she expected or desired a professional or skilled salary job, that is, a job for which at least secondary education would be required. For monthly earnings, and taking into account the highly skewed nature of the distribution of this variable, we follow Athey and Imbens (2016) and use an indicator of the rank in the earnings distribution.¹⁴ Finally, to account for multiple hypotheses testing, we use an aggregate indicator for both aspirations and expectations, which is the average of the standardized measures for educational attainment, occupation, monthly earnings, and living standard (number of rooms in the house), following Kling, Liebman, and Katz (2007).¹⁵ All standardized measures were obtained by subtracting the mean and dividing by the standard deviation of the control group.

9.3.2 Outcomes for Leaders

Identification relies on the random allocation of beneficiaries to one of the three program packages or the control, and the random allocation of these same packages among leaders. We consider both the leadership positions created in the treatment communities by the program (the *promotoras*) and other women with leadership positions, since they are not mutually exclusive (many health coordinators and teachers volunteered to be *promotoras*). Female leaders tend to be younger and more educated than the average female beneficiary. While beneficiaries on average have completed three years of education, leaders have completed, on average, five years. Other indicators of socioeconomic status at baseline are similar between leaders and nonleaders.

In Macours and Vakis (2014) we show that the randomization worked well and that the short-term returns to the largest package for the leaders were higher than for the other beneficiaries. During the intervention, leaders with the largest package also had higher nonagricultural and total income than leaders with other packages, reflecting the additional cash they had received to start new activities. As the income level and the income sources of these leaders at baseline were similar to those of the other beneficiaries, it seems plausible that beneficiaries could identify with their success during the program and that this might have motivated and inspired them.

The largest package is also the only intervention that led to gains in average income and consumption levels two years after the end of the program (Macours, Schady, and Vakis 2012; Macours, Premand, and Vakis 2012). Table 9.1 shows that leaders who received this package continue to

^{14.} Results are qualitatively similar when using the absolute value of earnings, winsorized at the 95th percentile.

^{15.} As an important share of parents desire professional jobs for their children, but few expect their children to get such jobs, we use the variable for professional job in the aggregate index for aspirations, and the variable for professional or skilled wage job in the aggregate index for expectations.

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Table 9.1	Comparison of follow-up outcome	s of leaders wit	h largest pack:	age with other l	eaders and nonlea	ders		
		Leader T1	Leader T2	Leader T3	Nonleader T3	<i>P</i> -value leaders T3–T1	<i>P</i> -value leaders T3-T2	<i>P</i> -value leaders T3–nonleader T3
Economic activities (Income from nom	in córdoba, per capita) avricultural self-em plovment	489.1	546	810	557	0.04**	0.09*	0.04**
Income from com	mercial activities	190.9	156	404	222	0.05^{**}	0.02^{**}	0.05^{**}
Income from agri	cultural wages	602.8	749	679	973	0.51	0.55	0.01^{**}
Value animal stoc	k c	1,630	2,104	2,191	1,631	0.14	0.84	0.13
Total income		11,707	12,049	12,272	10,925	0.51	0.78	0.05*
Expectations for chi	dren's future							
Average index		0.15	0.17	0.37	0.00	0.01^{***}	0.02^{**}	0.00^{***}
Expected years of	education attained	9.72	9.82	10.39	8.64	0.06*	0.07*	0.00^{***}
Expected occupat	ion: professional	0.03	0.04	0.04	0.05	0.43	0.77	0.90
Expected occupat	ion: professional or skilled empl.	0.35	0.36	0.47	0.26	0.04^{**}	0.06^{*}	0.00^{***}
Expected number	of rooms in house	2.64	2.76	2.87	2.74	0.05*	0.40	0.28
Expected monthly	/ earnings	2,132	2,047	2,332	1,976	0.12	0.02^{**}	0.00^{***}
Expected monthly	/ earnings (winsorized 95%)	1,703	1,637	1,898	1,593	0.13	0.04^{**}	0.00^{***}
Aspirations for child	ren's future							
Average index		0.06	0.19	0.17	0.02	0.15	0.83	0.06*
Desired years of 6	education attained	14.05	14.39	14.49	13.56	0.12	0.78	0.00^{***}

Desired occupation: professional	0.53	0.63	0.60	0.53	0.13	0.62	0.16
Desired occupation: professional or skilled empl.	0.90	0.94	0.92	0.88	0.49	0.56	0.09*
Desired number of rooms in house	5.18	5.35	5.221	5.25	0.83	0.49	0.88
Desired monthly earnings: rank	2,140	2,239	2,273	2,037	0.30	0.79	0.06*
Desired monthly earnings (winsorized 95%)	4,532	4,655	4,831	4,274	0.38	0.64	0.10^{*}
Human capital investment							
Attending school	0.863	0.84	0.82	0.77	0.49	0.96	0.00^{***}
Number of days absent from school	4.329	5.16	5.71	6.35	0.44	0.70	0.02^{**}
School expenditures	767	683	636	518	0.51	0.67	0.00^{***}
Share of food expenditures for animal products	0.17	0.17	0.18	0.16	0.37	0.97	0.07*
Share of food expenditures for vegetables and fruit	0.07	0.07	0.08	0.07	0.47	0.51	0.16
Weter Commissional interest to trace households in two		titice Doctor		tood award:	od or o bourt	lond blodon	dete Dete en edu

Note: Sample includes intent-to-treat households in treatment communities. Economic outcomes and food expenditures are household-level data. Data on eduaspirations questions refer to children nine to fifteen years old. Education questions refer to children seven to eighteen years old. Average expectation index is average of standardized outcomes for expected years of education, professional or skilled employment, number of rooms in the house, and monthly earnings rank. Average aspiration index is average of standardized outcomes for expected years of education, professional employment, number of rooms in the house, cation, expectations, and aspirations are child-level data. Highest and lowest 0.5 percent outliers of income and expenditures data trimmed. Expectation and and monthly earnings rank. Earnings ranks are calculated by converting the absolute monthly earnings to the rank in the earnings distribution, combining answers of leaders and nonleaders. All monetary values are in córdoba ($1 \text{ USS} = \sim 20$ córdobas). *P*-values account for clustering at the community level. ***Significant at the 1 percent level.

**Significant at the 5 percent level.

*Significant at the 10 percent level.

stand out. Two years after the end of the intervention, leaders who had the largest package still have higher incomes from nonagricultural selfemployment than other leaders. And their nonagricultural income and total income is significantly higher than for other beneficiaries who received the same package, even if their income from agricultural wages is lower. This suggests they may have been better in maintaining their new commercial activities and likely continue to be seen as successful leaders in the community.

We observe the same patterns with respect to parents' expectations for their children's future. For leaders who received the largest package, expectations are significantly higher than for other leaders. They expect their children to achieve higher schooling levels and earn higher wages, and are 11 percentage points more likely to expect their children to become a professional or skilled salary earner (table 9.1). There are also large differences in the expectations of these leaders and those of other beneficiaries who received the same package, with leaders expecting their children to obtain 1.5 years more education, and 21 percentage points more likely to become a professional or skilled salary earner. These latter differences are consistent with the program identifying natural leaders through self-selection. The differences mean that these leaders may be seen as local success stories—in both current achievement and their attitudes toward the future—that others could aspire to emulate.

A similar pattern is found for differences between leaders and others in their reported aspirations for their children, although differences in aspirations are smaller than differences in expectations. Comparing mean values of aspirations and expectations shows large gaps between the two sets of outcomes, with expectations for educational attainment, for instance, five years less than aspirations and similarly large differences for earnings, occupation, and living standards. Interestingly, these gaps are smaller for leaders than nonleaders. The pattern suggests that both leaders and other households internalize their constraints when reporting their expectations, but they also suggest a capacity to aspire to a much better lives for their children.¹⁶

In line with the other results, leaders' investments in the education and nutrition of their children are higher than those of others beneficiaries. The significant differences in human capital investment between leaders and nonleaders in table 9.1 mirror similar findings from the baseline and the midline survey (Macours and Vakis 2014). Leaders with the largest package hence provided positive examples for others, in line with the program objectives, both during the program and two years after the transfers ended.

16. The questions for expectations specifically asked: "Taking into account your current situation, what do you expect . . . ?"

9.3.3 Empirical Specification

To analyze whether higher exposure to leaders with the largest package changed education and nutrition investments of other beneficiaries, we calculate the share of leaders randomly allocated to the largest package in each registration assembly, including—as before—both *promotoras* and other women with leadership positions in the community. The average number of leaders in an assembly is four so that there is substantial variation in the share of leaders that got the largest package in an assembly. There is much less variation in the share of other beneficiaries who got the largest package since the number of households in each assembly was relatively large and thus the share of nonleaders with the largest package in each assembly is close to one-third in all assemblies.¹⁷

Our general specification is

(1)
$$Y_{ia} = \delta_0 + \delta_1 T_{ia} + \delta_2 (T_{ia} * S_a) + \delta_3 S_a + \varepsilon_{ia}$$

where Y_{ia} is an outcome indicator for eligible household *i* (or a child of household *i*) who was invited to assembly *a*, T_{ia} is assignment of *i* to any of the three treatment groups, and S_a is the share of leaders in the assembly that randomly received the largest package in i's registration assembly. Given that households were invited to particular assemblies based on geographic proximity, S_a will capture the share of leaders with the largest package that live in the proximity of *i*.¹⁸ Since S_a is always 0 in the control communities, and since all eligible households in the treatment communities receive one of the three intervention packages, the term $\delta_3 S_a$ cancels out of the estimation. The coefficients of interest are δ_1 and δ_2 . A finding, for example, that δ_1 and δ_2 are both positive would imply that while assignment to the treatment group increases the outcome of interest (δ_1) , there is an additional impact of the program that comes from the social interactions (δ_2). We also explore how the share of leaders with the largest package affects impacts for beneficiaries of each of the three packages separately. All regressions are estimated on the sample of eligible households (or their children) that are not leaders themselves.

9.4 Social Interaction Effects on Human Capital Investments

9.4.1 Main Results

We first pool households across treatment packages and investigate whether there is a general relationship between program impacts and proximity to

17. The shares of peers at the 10th and 90th percentiles of the distribution are 21 and 39 percent, respectively. In contrast, for the leaders, the shares at the 10th and 90th percentiles of the distribution are 0 and 67 percent.

18. Location of one's house might be endogenous, and people living in the proximity of leaders might also be more likely to be their family members, or otherwise have similar characteristics. The identification in this chapter does not depend, however, on the proximity to the leader per se, but instead it depends on the random allocation of certain packages to those leaders. leaders who received the largest package. Table 9.2 presents in the top panel the results for 2008, the main focus of this chapter, and in the bottom panel, the findings for 2006 from our earlier work for comparison. The interaction terms in the top panel suggest that social interactions are crucial to sustain program impacts on education and nutrition investments after the end of the intervention. Indeed, the findings indicate no significant sustained impacts on human capital investments when no leader was assigned the largest package, in contrast to the findings during program implementation.

The interaction terms suggest that the higher the share of leaders with the largest package, the less likely children are absent in school and the more households invest in education, in animal proteins, and in fruit and vege-tables. The social multiplier effects are not only statistically significant but also large. For example, school expenditures increase 49 percent when all the leaders in one's assembly got the largest package, while school absences decline by 21 percent. Strikingly, the magnitude of the social multiplier effects two years after the end of the program are similar, if not larger, than those while the intervention was in place. The coefficients of the *z*-scores in table 9.2 also imply that, two years after the transfers, the impact on nutrition, respectively educational, investment was only significantly different from zero if at least 33 percent, respectively 75 percent, of leaders in one's registration assembly received the largest package.

Table 9.3 shows the social interaction impacts on human capital investments by treatment group. The effects are strongest for beneficiaries of the largest package. For instance, school expenditures more than double for beneficiaries of the largest package in the extreme case that the share of female leaders with the same package changes from 0 to 1. The impacts are about half the size for the beneficiaries of the training packages (and even smaller for those with the basic package) for most outcomes and the interaction effects for education investments are not significant. Nevertheless, as for the 2006 findings, the *p*-values indicate that we cannot reject that the social effects are the same for the three groups for most variables. And when pooling the basic and the training packages, the interaction effects for school expenditures, the nutrition index, expenditures for animal proteins, and fruit and vegetables are all significant (not shown). This suggests that the results are not only driven by complementarities between the extra cash received by beneficiaries and the leaders' package.

Note that while the coefficients of the interaction effects are large, there are on average about four leaders in a registration assembly. The estimates hence indicate that having one additional leader with the largest package in one's assembly reduces school absences by 0.4 days per month and increases school expenditures by about 16 percent. For households that have the largest package, one additional leader with the same package increases school attendance by 2.5 percentage points and increases school expenditures by 25 percent. These are not only large effects, but are similar or even larger than

		Ed	lucation			Nutrition	
	Z-score education investment	Attending school (7–18-year-olds)	Number of days absent from school (7–18-year-olds)	School expenditures (7–18-year-olds)	Z-score nutrition investment	Share of food expenditures for animal products	Share of food expenditures for fruit and vegetables
		2008 (1	two vears after progra	im ended)			
Intent-to-treat \times % leaders	0.116^{**}	0.045	-1.506*	310.9***	0.353***	0.039**	0.022^{***}
with largest package	(0.050)	(0.040)	(0.88)	(118)	(0.12)	(0.017)	(0.008)
Intent-to-treat	-0.031	-0.008	0.197	-68.80	-0.013	-0.005	0.001
	(0.027)	(0.026)	(0.58)	(62.5)	(0.066)	(0.010)	(0.004)
Mean dep. variable in control	0.001	0.777	6.341	493.4	-0.002	0.154	0.0581
Observations	5,231	5,228	5,228	5,205	3,230	3,214	3,214
		200)6 (during implement	ation)			
Intent-to-treat \times % leaders	0.127^{**}	0.062*	-1.760^{***}	191.7^{***}	0.201^{**}	0.022	0.014^{**}
with largest package	(0.054)	(0.032)	(0.669)	(70.9)	(0.093)	(0.017)	(0.006)
Intent-to-treat	0.134^{***}	0.050 * * *	-1.352^{***}	188.6^{***}	0.381^{***}	0.055^{***}	0.019^{***}
	(0.028)	(0.019)	(0.405)	(34.8)	(0.058)	(0.010)	(0.004)
Mean dep. variable in control	0.001	0.761	6.209	300.9	-0.003	0.152	0.066
Observations	5,181	5,176	5,169	5,153	3,294	3,278	3,279

household-level data for food expenditures. Excluding households with female leaders. Intent-to-treat estimators. Highest and lowest 0.5 percent of outliers in expenditures trimmed. Robust standard errors in parentheses, corrected for clustering at the community level. ***Significant at the 1 percent level.

**Significant at the 5 percent level. *Significant at the 10 percent level.

		-	`	-			
		Edu	cation			Nutrition	
	Z-score education	Attending	Number of days absent from	School	Z-score nutrition	Share of food expenditures for animal	Share of food expenditures for fruit and
	investment	(7–18-year-olds)	(7–18-year-olds)	(7–18-year-olds)	investment	products	vegetables
		2008	(two years after prog	ram ended)			
Productive investment	0.188^{**}	0.093*	-2.676^{**}	485.4**	0.498^{***}	0.050^{**}	0.034^{***}
package \times % leaders	(0.091)	(0.050)	(1.09)	(200)	(0.13)	(0.019)	(0.011)
with largest package							
Training package $\times \%$	0.096	0.029	-1.017	246.2	0.364^{**}	0.038*	0.023^{**}
leaders with largest	(0.071)	(0.061)	(1.38)	(165)	(0.15)	(0.021)	(0.011)
package							
Basic package \times % leaders	0.062	-0.001	-0.538	192.8	0.222	0.032	0.011
with largest package	(0.060)	(0.053)	(1.15)	(154)	(0.15)	(0.020)	(0.012)
Productive investment	-0.050	-0.034	0.764	-114.0	-0.022	-0.004	-0.001
package	(0.032)	(0.032)	(0.69)	(72.0)	(0.062)	(0.001)	-0.005
Training package	-0.014	0.007	-0.041	-36.92	-0.072	-0.012	-0.002
	(0.034)	(0.030)	(0.69)	(77.2)	(0.082)	(0.013)	-0.006
Basic package	-0.024	0.011	-0.299	-46.06	0.051	0.001	0.005
	(0.029)	(0.031)	(0.69)	(69.4)	(0.0)	(0.011)	-0.005
P-value test social effect	0.413	0.360	0.291	0.348	0.350	0.575	0.373
on T3 vs. T2							
P-value test social effect	0.196	0.109	0.116	0.193	0.035**	0.252	0.069*
on T3 vs. T1							
P-value test social effect	0.668	0.671	0.744	0.779	0.350	0.743	0.350
on T1 vs. T2							
Mean dependent variable	0.001	0.777	6.341	493.4	-0.002	0.154	0.0581
in control							
Observations	5,231	5.228	5.228	5,205	3,230	3.214	3.214

Social interaction effects on human capital investments by intervention group

Table 9.3

		200	6 (during implemente	ttion)			
Productive investment	0.192^{**}	0.097^{**}	-2.579***	291.6***	0.328**	0.044^{**}	0.019*
package \times % leaders	(0.076)	(0.047)	(0.975)	(102.5)	(0.13)	(0.019)	(0.011)
with largest package							
Training package $\times \%$	0.097	0.047	-1.356	145.6^{*}	0.139	0.017	0.008
leaders with largest	(0.066)	(0.041)	(0.844)	(81.9)	(0.11)	(0.021)	(0.007)
package							
Basic package \times % leaders	0.102	0.045	-1.293	149.3*	0.158	0.006	0.016
with largest package	(0.062)	(0.052)	(1.128)	(82.8)	(0.13)	(0.021)	(0.010)
Productive investment	0.129^{***}	0.045^{**}	-1.107^{**}	174.3^{***}	0.368^{***}	0.049^{***}	0.020^{***}
package	(0.031)	(0.022)	(0.458)	(39.5)	(0.067)	(0.011)	(0.005)
Training package	0.125^{***}	0.049^{**}	-1.438^{***}	181.4^{***}	0.379***	0.057***	0.018^{***}
	(0.032)	(0.023)	(0.479)	(39.4)	(0.059)	(0.011)	(0.004)
Basic package	0.149^{***}	0.057^{**}	-1.574^{***}	211.8^{***}	0.393^{***}	0.058^{***}	0.020^{***}
	(0.033)	(0.026)	(0.584)	(42.2)	(0.070)	(0.011)	(0.005)
P-value test social effect	0.226	0.434	0.360	0.151	0.170	0.174	0.327
on T3 vs. T2							
<i>P</i> -value test social effect on T3 vs T1	0.192	0.306	0.238	0.124	0.208	0.0325**	0.810
<i>P</i> -value test social effect	0.934	0.964	0.959	0.964	0.889	0.603	0.518
on T1 vs. T2							
Mean dependent variable	0.001	0.761	6.209	300.9	-0.003	0.152	0.066
in control							
Observations	5,181	5,176	5,169	5,153	3,294	3,278	3,279
Note: Coefficients for index of	family of outcom	es calculated follov	ving Kling, Liebmar	t, and Katz (2007).	The share of lead	ers measures the sl	nare of female
<i>Note</i> : Coefficients for index of	family of outcom	es calculated follov	ving Kling, Liebmar	, and Katz (2007).	The share of 1	ead	eaders measures the sh

leaders with the productive investment package over all female leaders in a beneficiary's registration assembly. Individual-level data for education and household-level data for food expenditures. Excluding households with female leaders. Intent-to-treat estimators. Highest and lowest 0.5 percent of outliers in expenditures trimmed. Robust standard errors in parentheses, corrected for clustering at the community level.

***Significant at the 1 percent level.

**Significant at the 5 percent level.

*Significant at the 10 percent level.

the effects found in 2006. Hence, interactions with leaders had a remarkably persistent impact on other households' investment behavior, and the impact is particularly important for households that themselves received the largest package.

In contrast to the impacts during the intervention, however, none of the packages had positive impacts on investments for beneficiaries that were not exposed to any leader with the largest package, and indeed the point estimates are negative for a number of indicators. Hence positive significant effects are only found for the subset of the beneficiaries with high exposure to successful leaders.

9.4.2 Robustness

The results are robust to several alternative specifications.¹⁹ A first concern could be that the results are driven by extreme values in the independent variable. While the average share of leaders with the largest package is 0.33, for 95 percent of the observations, the range is between 0 and 0.67. The first robustness check in table 9.4 excludes observations with values above 0.67. This does not substantially alter any of the results, even if, as expected, the standard errors increase. The results are also robust to clustering the standard errors at the level of the registration assembly, as opposed to the community level, and to not excluding outliers. The next two specifications show that the results are further robust to controls for the total number of people in an assembly, or the total number of peers (defined as beneficiaries that are not leaders) in an assembly. Finally, the results remain generally robust when including a community fixed effect, with the exception of the food expenditures for animal products, even if the variation in the independent and to variable is reduced.

Table 9.4 also shows alternative specifications using the number of leaders with the largest package instead of the share. These specifications separately control for the total number of leaders in the registration assembly. The coefficient on the number of leaders with the largest package is consistent with the main results in terms of sign, size, and magnitude. We can then also compare the coefficient of the number of leaders with the largest package and with the coefficient of the number of peers with the largest package (last specification in table 9.4). The results suggest that social interaction effects from peers might be more limited: the coefficients are generally not significant and smaller than the coefficients for the number of leaders, with the exception of the expenditures for animal products. The coefficients for leaders and peers are significantly different for school attendance, absences, and spending on fruit and vegetables. Note, however, that these results should be interpreted with caution, given that they could be

19. Table 9.4 presents robustness checks for the beneficiaries with the productive investment package. Results pooling all beneficiaries are similarly robust.

	Attending school (7–18-year-olds)	Number of days absent from school (7–18-year-olds)	School expenditures (7–18-year-olds)	Share of food expenditures for animal products	Share of food expenditures for fruit and vegetables
Base specification	0.093* (0.050)	-2.676** (1.09)	485.4** (200)	0.050^{**} (0.019)	0.034^{***} (0.011)
Robustness checks Excluding extreme values independent variable	0.064	-2.087*	319.5* (196)	0.061***	0.037***
SE clustered at level of assembly	(70.0) 0.093 (0.061)	(1.23) -2.676** (1.29)	(1 60) 485.4*** (177)	(0.022) 0.050** (0.022)	(0.012) 0.034*** (0.010)
Not excluding outliers			726.0**	0.051**	0.040***
Controlling for number of people in assembly	0.093*	-2.681** /1.06)	485.8**	0.051***	0.034***
Controlling for number of peers in assembly	0.095*	(1.00) -2.745**	495.4**	0.051***	0.034^{***}
With community fixed effects	(0.048) 0.096^{*}	(1.05) -2.668**	(203) 350.2	0.009	(0.011) 0.022**
Alternative specifications with number of leaders	(10.0)	(1.16)	(226)	(0.019)	(0.010)
No. leaders with largest package controlling for total no. leaders	0.019 (0.015)	-0.599* (0.33)	94.64** (47.2)	0.013^{**} (0.005)	0.009 *** (0.003)
No. leaders with largest package controlling for total no. leaders	0.033**	-0.855**	72.98	0.001	0.006**
and community i.e. Alternative specifications with number of leaders and number of peers	(010.0)	(05.0)	(8.66)	(cnn.n)	(500.0)
No. leaders with largest package controlling for total no. leaders	0.031	-0.833*	100.3*	0.009	0.007**
No. peers with largest package controlling for total no. peers and	(0.019) -0.003	(0.42) 0.028	(55.2) 32.76	(0.007) 0.012^{**}	(0.003) 0.001
community f.e.	(0.020)	(0.45)	(36.9)	(0.005)	(0.003)
P-value test social effect leader = social effect peer	0.064*	0.046**	0.265	0.567	0.100^{*}
Note: See notes to table 9.3. Every line corresponds to a separate specif in the same specification. Peers are defined as all beneficiaries with the sa cludes observations for which the value of the share is in the unper 5 perc	ication, with the exce me package that are n ent of the distribution	ption of the last specified to the last specified of leaders. Specification	cation where the num n with extreme values	ber of leaders and p of independent varia	eers are included ble excluded: ex-

Robustness checks and alternative specifications: beneficiaries of largest package two years after program

Table 9.4

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***Significant at the 1 percent level. **Significant at the 5 percent level. *Significant at the 10 percent level. driven by the fact that there is less variation to identify the social effects of peers.

9.4.3 Results for Investments in Early Childhood

Table 9.5 shows estimates of the social multiplier effects for investments during early childhood, showing estimates for all children from birth to seven years old, and separately for the cohort of children born after the end of the transfers (i.e., children approximately from birth to two years old). The results show relatively large social multiplier effects for both age groups for both nutrition and stimulation. In contrast, we do not observe a similar pattern for health investments. This result is interesting, as the health conditionalities in Atención a Crisis were never monitored due to coordination problems between the ministry of health and the ministry of the family. It seems plausible that the health component was seen as less salient by leaders and less emphasized during discussions about the program. For nutrition and stimulation, the coefficients are positive and significantly different from zero for almost all beneficiaries, except those that were not exposed to any leader with the largest package. Impacts are larger for beneficiaries with higher exposure to such leaders.

Importantly, we find similar strong and significant social multiplier effects for investments in nutrition and stimulation for children born after the end of the transfers. This is true even if the statistical power is reduced as the cohort is much smaller. There are no significant differences in the coefficient of the multiplier effects for beneficiaries with different packages. Overall, these results point to a permanent shift in investment behavior among families exposed to successful leaders that goes beyond the impacts on the children that directly benefited from the intervention. This result suggests that the results in this chapter are not only driven by lasting impacts on children directly exposed to the positive CCT shock.

9.5 Social Interaction Effects on Per Capita Expenditures, Expectations, and Aspirations

While the identification strategy allows to clearly demonstrate the importance of the social interaction effects, it does not necessarily help to understand how exactly leaders might be influencing other households' investments. Indeed, one can wonder whether interaction with leaders with the largest package may have lifted economic constraints of other households, whether the interaction effects are driven by other households mimicking the behavior of these leaders, or whether they capture actual shifts in aspirations and expectations of nonleader households for the future of their children.

We investigate this question by analyzing the data regarding mothers' expectations and aspirations for children's final educational levels, future occupation, and earnings and living standards. Table 9.6 shows results of

Children 0–	7 years old	Children	born after end of ti (0-2 years old)	ansfers
Nutrition Stimu	lation Healt	h Nutrition	Stimulation	Health
age 0.323*** 0	.143* 0.00	4 0.383**	0.243**	0.087
(0.11) (0	.078) (0.07	(0.18) (0.18)	(0.11)	(0.089)
-0.001 0	.048 0.08	3** 0.067	0.033	-0.010
(0.047) (0	.051) (0.03	5) (0.072)	(0.059)	(0.045)
3,410 3,4	05 3,410	(099	641	660
toomes calculated following Kling, vroteins, and fruit and vegetables in s/sings to child, and number of hour- and number of days sick in bed. Se	Liebman, and Katz total food expenditu s reading to child per te Macours, Schady,	(2007). Nutrition index is res: stimulation index i week; health index inclu and Vakis (2012) for de	includes share of four the share of four the share of the second state of the second state of the share of th	od in total sehold has as weighed, ne share of
age 0.323*** (0.11) -0.001 (0.047) 3,410 (0.047) 3,410 toomes calculated following <i>K</i> troteins, and fruit and vegetab visings to child, and number of and number of days sick in b	$\begin{array}{c c} & 0 \\ & 0$	0.143* 0.00 (0.078) 0.078 0.048 0.078 0.048 0.08 0.051) 0.08 3,405 3,416 Cling, Liebman, and Katz 0.03 Ling, Liebman, and Katz ed. See Macours, Schad dy;	0.143* 0.004 0.383** (0.078) (0.076) (0.18) 0.048 (0.076) (0.18) 0.048 (0.035) (0.067) 0.051 (0.035) (0.072) 3,405 3,410 660 Cling, Liebman, and Katz (2007). Nutrition index less in total food expenditures; stimulation index inclued and a mocus. Schady, and Vakis (2012) for detail the converse of the mocle and converse over all female leaders in a heneficiary.	0.143* 0.004 0.383** 0.243** (0.078) (0.076) (0.18) (0.11) 0.048 (0.076) (0.18) (0.11) 0.048 0.083** 0.067 (0.033) 0.051 0.083** 0.067 (0.033) 3.405 3.410 660 641 Cling, Liebman, and Katz (2007). Nutrition index includes share of for hour hours reading to child per week; health index includes whether child we hours reading to child per week; health index includes whether child we hours reading to child per week; nealth index includes whether child we hours reading to child per week; nealth index includes whether child we hours cover all female leaders in a hearteristic in a security in a security. That security in a securi

ņ households with female leaders. Intent-to-treat estimators. Robust standard errors in parentheses, corrected for clustering at the community level. ***Significant at the 1 percent level.

**Significant at the 5 percent level.

*Significant at the 10 percent level.

	Average expectation index	Expected years of education attained	Expected occupation: professional	Expected occupation: professional or skilled empl.	Expected number of rooms in house	Expected monthly earnings: rank	Years of education attained	Log (per capita expenditures)
Intent-to-treat × % leaders with largest package Intent-to-treat Mean dependent variable in the control Observations	0.267*** (0.095) -0.052 (0.056) 0.00 4,304	0.805 (0.50) -0.132 (0.27) 8.41 4,300	0.035** (0.017) 0.002 -0.008 0.023 4,291	0.137*** (0.048) -0.033 (0.030) 0.25 4,291	0.304** (0.12) -0.032 (0.076) 2.62 4,299	315.1** (147) -79.30 (93.4) 2,543 4,242	0.700*** (0.22) -0.171 (0.15) 3.12 4,086	0.126 (0.096) 0.008 (0.049) 9.11 3,230
	Average aspiration index	Desired years of education attained	Desired occupation: professional	Desired occupation: professional or skilled empl.	Desired number of rooms in house	Desired monthly earnings: rank		
Intent-to-treat × % leaders with largest package Intent-to-treat Mean dependent variable in the control Observations	0.205* (0.10) -0.018 (0.067) -0.001 4,302	0.946** (0.42) -0.090 (0.26) 13.29 4,300	0.115* (0.065) -0.015 (0.048) 0.50 4,299	0.028 (0.026) 0.015 (0.019) 0.872 4,299	0.146 (0.37) -0.063 (0.15) 5.19 4,300	342.2* (185) 26.48 (124) 2,485 4,289		
<i>Note:</i> The share of leaders me Individual-level data for child Intent-to-treat estimators. Role outcomes for expected years o average of standardized outcoi are calculated by converting th ***Significant at the 1 percent	asures the share of ren seven to fifteer ren standard erro of education, profe mes for expected y te absolute monthl level.	f female leaders 1 years old in 200 1 years old in 200 200 200 200 200 200 200 200 200 200	with the producti 08. Per capita exp , corrected for clu, employment, nu n, professional em rank in the earnir	ve investment pac enditure measure- ustering at the com mber of rooms in ployment, numbe igs distribution, au	kage over all fen dat the househo amunity level. An the house, and 1 r of rooms in the and combining an	nale leaders in a l latel. Excludin (erage expectation monthly earnings house, and mont swers of leaders a	beneficiary's regis beneficiary's regis in index is average arank. Average as i rank. Average as inthy earnings rank nd nonleaders.	tration assembly. h female leaders. of standardized piration index is . Earnings ranks

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Table 9.6

the main specification for these outcomes, and also shows the spillovers on the educational level attained by 2008 and on per capita expenditures levels. These questions were only asked for children less than fifteen years old, as older children are more likely to already have reached their final education levels.²⁰ The top panel shows the impacts on the expectations mothers reported for their children, while the lower panel shows impacts on their aspirations.

Table 9.6 shows that parents' expectations about their children obtaining professional jobs or skilled salary jobs are strongly affected by exposure to leaders with the largest package. Having one more such leader in one's registration assembly increases expectations of parents for their children to become (white-collar) professionals by almost 50 percent (starting from a very low level in the control group). Strong social multiplier effects are also found for expectations regarding children's future earnings and living standards. The social interaction effects for mothers' aspirations follow a similar pattern.

Averaging over the different indicators, we find that the difference between no exposure and full exposure to leaders with the largest package increases expectations regarding children's future with 0.27 standard deviations, while it increases aspirations with 0.21 standard deviations.²¹ The coefficients of the z-scores in table 9.6 also imply that, two years after the transfers, the impact on expectations and aspirations was only significantly different from zero if at least 60 percent of leaders in one's registration assembly received the largest package. Overall, these findings show that interactions with successful female leaders changed beneficiaries' expectations and aspirations for their children's educational and occupational future, consistent with the sustained higher levels of human capital investments.

The table further shows that the large spillover effects in investments are reflected in spillovers in educational attainment by 2008. Indeed, two years after the end of the intervention, being exposed to one additional leader with the largest package increases children's school attainment with 0.18 years of schooling. A comparison of this coefficient with the estimate on expectations suggests that parents expect the educational gains to persist and possibly slightly increase in the future. The estimates also imply, however, that the one-year CCT program did not significantly increase educational attainment for more than half of the children in the sample. Finally, the table shows that there is no significant social multiplier for per capita expenditures.

21. Increases in aspirations (expectations) reflect a change toward more ambitious aspirations (expectations).

^{20.} In an alternative specification, we excluded children below nine years old from the analysis, as the younger children did not directly benefit from the educational component of the CCT during the intervention. Results are broadly similar, but social multipliers on expected years of education, attained years of education, and expected earnings are slightly larger for the beneficiaries of the largest package.

To further understand the potential role of relaxing economic constraints, table 9.7A shows social multiplier effects separately for each of the three types of beneficiaries. The last column in table 9.7A shows a significant social multiplier effect on per capita expenditures for households who got the largest package, but no such effects exist for the two other packages. Hence, beneficiaries who got the largest package are still better off two years after the intervention when a sufficiently large share of leaders in their proximity received the same package. This result is in line with findings on similar spillovers for productive investments during the intervention (Macours and Vakis 2014). In contrast, the coefficients of the interaction effects of per capita expenditures for beneficiaries with the basic and training package are very small, not significantly different from zero, and significantly different from the interaction effect for the largest package. Hence for the two other groups, economic spillovers cannot explain the change in education and nutrition investment.²²

Considering then the impacts on expectations (table 9.7A), we see that social multiplier effects are significant for all three interventions, and are not significantly different from each other. Hence beneficiaries of the three packages expect a better future for their children, as long as they have sufficiently high exposure to leaders with the largest package. This is so even if only beneficiaries with the largest package are economically better off two years after the interventions. This mirrors the findings for investments, where we also found no significant differences between groups, even if the point estimates are higher for beneficiaries of the largest package. Hence the social multiplier effects do not just come from changes in economic constraints, as expectations and investments change also for the groups for whom economic constraints were not relaxed.

The results on aspirations (table 9.7B) complete the picture, as the social multiplier effects for aspirations are concentrated on beneficiaries who received the training package (and to a lesser extent those with the basic package). Possibly, for beneficiaries of the productive package, their own experience of trying to develop a nonagricultural activity may have dampened the impact of leaders' experiences on aspirations, in particular, given that average aspiration levels were already high. One could also hypothesize that the focus of this package on nonagricultural self-employment led these beneficiaries to put less weight on professional occupations or high levels of education for their children. In contrast, the results for beneficiaries of the training package are driven in particular by the aspirations for education and professional occupation, which may suggest training led to a higher

^{22.} One could have thought such economic spillovers could arrive from leaders employing other beneficiaries in their new business, or otherwise transferring economic benefits from their increased income to other beneficiaries.

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expectation	Expected	Freedad	occupation:	Expected	Expected	Vanue of	Ιοα
index	n education attained	occupation: professional	or skilled employment	rooms in house	earnings: rank	education attained	per capita (per capita expenditures)
Productive investment package $\times \%$ 0.270**	0.658	0.105^{**}	0.158***	0.315*	343.3	0.521	0.277***
leaders with largest package (0.12)	(0.74)	(0.036)	(0.058)	(0.17)	(207)	(0.44)	(0.083)
Training package \times % leaders with 0.306**	0.991	-0.010	0.157**	0.385**	267.7	0.931***	0.070
largest package (0.14)	(0.73)	(0.028)	(0.076)	(0.16)	(223)	(0.32)	(0.10)
Basic package \times % leaders with largest 0.220*	0.761	0.016	0.092	0.217	310.4	0.596^{*}	0.051
package (0.12)	(0.53)	(0.027)	(0.070)	(0.19)	(237)	(0.35)	(0.15)
Productive investment package -0.055	-0.087	-0.010	-0.041	-0.020	-112.5	-0.169	-0.010
(0.057)	(0.28)	(0.011)	(0.033)	(0.07)	(102)	(0.18)	(0.045)
Training package –0.068	-0.225	0.010	-0.036	-0.065	-87.8	-0.265	0.011
(0.072)	(0.38)	(0.014)	(0.037)	(0.10)	(108)	(0.19)	(0.057)
Basic package –0.029	-0.086	0.006	-0.021	-0.013	-22.2	-0.061	0.022
(0.066)	(0.30)	(0.010)	(0.035)	(0.10)	(119)	(0.17)	(0.064)
<i>P</i> -value test social effect on T3 vs. T2 0.830	0.737	0.004^{***}	0.996	0.757	0.743	0.490	0.026^{**}
<i>P</i> -value test social effect on T3 vs. T1 0.693	0.863	0.088*	0.428	0.629	0.911	0.885	0.061^{*}
<i>P</i> -value test social effect on T1 vs. T2 0.550	0.725	0.527	0.476	0.452	0.899	0.430	0.860
Mean dependent variable in the control 0.00	8.41	0.023	0.25	2.62	2,543	3.12	9.11
Observations 4,304	4,300	4,291	4,291	4,299	4,242	4,086	3,230

Individual-level data for children seven to fifteen years old in 2008. Per capita expenditure measured at the household level. Excluding households with female leaders. Intent-to-treat estimators. Robust standard errors in parentheses, corrected for clustering at the community level. Average expectation index is average of standardized outcomes for expected years of education, professional or skilled employment, number of rooms in the house, and monthly earnings ranks ranks are calculated by converting the absolute monthly earnings to the rank in the earnings distribution, and combining answers of leaders and nonleaders. ***Significant at the 1 percent level.

**Significant at the 5 percent level. *Significant at the 10 percent level.

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Social interaction effects on parental expectations, educational attainment, and per capita expenditures two years after program Table 9.7A

Social interaction effects on parental aspirations by intervention group two years after program ended

Table 9.7B

~				-		
N.		Desired		occupation:	Desired	Desired
A.2	Average	years of	Desired	professional	number of	monthly
asi	spiration	education	occupation:	or skilled	rooms in	earnings:
	index	attained	professional	empl.	house	rank
Productive investment package $\times \%$ –0.	0.008	0.338	-0.032	-0.038	-0.481	331.0
leaders with largest package (0.	0.12)	(0.57)	(0.091)	(0.042)	(0.34)	(304)
Training package \times % leaders with 0 .	0.406***	2.114***	0.297 * * *	0.084^{**}	0.378	266.3
largest package (0.	0.13)	(0.70)	(0.097)	(0.042)	(0.24)	(236)
Basic package \times % leaders with largest 0 .	0.233	0.379	0.081	0.039	0.603	470.6*
package (0.	0.18)	(0.52)	(0.078)	(0.046)	(0.89)	(276)
Productive investment package 0.	0.068	0.169	0.039	0.034	0.163	68.1
0)	0.072)	(0.29)	(0.051)	(0.023)	(0.17)	(154)
Training package -0.	0.147**	-0.661^{*}	-0.119^{**}	-0.016	-0.300^{**}	26.5
(0)	0.074)	(0.36)	(0.059)	(0.027)	(0.15)	(125)
Basic package 0.	0.015	0.212	0.033	0.025	-0.087	-32.3
0)	(620)	(0.29)	(0.051)	(0.026)	(0.22)	(151)
<i>P</i> -value test social effect on T3 vs. T2 0.	0.006***	0.027^{**}	0.007^{***}	0.290	0.033^{**}	0.853
<i>P</i> -value test social effect on T3 vs. T1 0.	0.206	0.958	0.309	0.832	0.211	0.701
<i>P</i> -value test social effect on T1 vs. T2 0.	.319	0.016^{**}	0.017^{**}	0.375	0.784	0.548
Mean dependent variable in the control -0.	.001	13.29	0.50	0.872	5.19	2,485
Observations	4,302	4,300	4,299	4,299	4,300	4,289

assembly. Individual-level data for children seven to fifteen years old in 2008. Excluding households with female leaders. Intent-to-treat estimators. Robust standard errors in parentheses, corrected for clustering at the community level. Average aspiration index is average of standardized outcomes for expected years of education, professional employment, number of rooms in the house, and monthly earnings rank. Earnings ranks are calculated by converting the absolute monthly earnings to the rank in the earnings distribution, and combining answers of leaders and nonleaders. ***Significant at the 1 percent level.

**Significant at the 5 percent level. *Significant at the 10 percent level.

orientation toward education.²³ Yet the results show these increased aspirations only materialized if they were exposed to a large share of leaders with the largest package. Indeed, there even is a significant negative impact on aspirations if none of the leaders in their proximity got the largest package, suggesting that some training beneficiaries in fact got demotivated.²⁴

While the differences in findings for expectations and aspirations are intriguing, they could in part be driven by measurement errors. Parents may find hypothetical questions regarding the desired future for their children hard to answer and the difference between expectations and aspirations, while theoretically important for economists, are not necessarily accurately captured by the answers mothers gave to the respective questions. A more cautious interpretation of the results in tables 9.6 and 9.7A and B, therefore, is that there was a significant social multiplier effect on parents' attitudes regarding their children's future for the three types of beneficiaries. This attitudinal change in turn is consistent with the social multiplier effect on investment behavior, and helps explain the increased investment by beneficiaries for whom per capita expenditure levels did not increase.

Finally, one can wonder whether the results after the end of the program result because leaders with the largest package are still communicating more with other beneficiaries than leaders with other packages. This does not appear to be the case, as we find no significant social multiplier effects on the probability of talking to a leader, a teacher, or health coordinators two years after the intervention (not reported). This is in line with qualitative interviews after the end of the program, which suggested that some *promotoras* may have continued to meet with beneficiaries, but that this was rather limited. Nevertheless, increased communication during the program may have played a role in shifting the local social norms toward more investment in children on the short term, which in turn may have led to persistent changes in investment behavior.

9.6 Conclusions

Many development interventions aim, through a variety of mechanisms, to shift the investment behavior of beneficiary households. Conditional cash transfer programs have an implicit or explicit objective to change households' attitudes and the social norms toward investment in the education, health, and nutrition of their children. When programs are designed to last for only a limited period, the sustainability of the impacts might crucially

24. As training did not lead to significant increases in income, this does not seem implausible.

^{23.} The vocational training may have made beneficiaries more aware of the potential gains to formal education through a number of channels: exposure to the professional staff conducting the training, increased awareness of the benefits of skilled wage employment, or awareness of the importance of education to increase returns to other training (as illiterate beneficiaries in particular were very limited in their choice of courses).

depend on whether changes in investment behavior persist after the end of the program. Yet, the mechanisms through which such change can be reached and reinforced are not always clear.

This chapter shows that social interactions with successful and motivated local leaders can change the way parents think about their children's future and result in sustainable changes in educational and nutritional investment. The evidence in this chapter hence draws attention to the positive role local leaders can play, which contrasts with the focus in many policy discussions of the negative role of leaders through elite capture. The results suggest that natural leaders living in people's close proximity can be important vehicles for change by motivating and encouraging others and by providing examples that people aspire to follow. We find these effects when both leaders and other beneficiaries received sizable transfers, and social effects are particularly large when leaders and beneficiaries received the same package. Hence the results do not suggest that interventions should be primarily targeted to leaders. Instead, it points to the importance of assuring that development program designs take into account the presence of local natural leaders and enhance their ownership of a program's objectives to help shift beneficiaries' attitudes.

The large social interactions effects found in this chapter are suggestive of the existence of multiple equilibria and can hence be interpreted in the context of dynamic poverty trap models (Barrett, Carter, and Chavas, introduction, this volume). Interventions such as the program studied in this chapter may not only affect external constraints, but also shift internal potential constraints, potentially breaking existing poverty traps, in line with Lybbert and Wydick (chapter 4, this volume). If internal constraints are partly driven by beliefs or social norms, local leaders can have an important role in helping shift these social norms and help tip communities toward the high equilibrium.

The results also have implications for the debate on the sustainability of using cash or asset transfer programs in low- and middle-income countries. The evidence in this chapter suggests that designing such programs in ways that facilitate and encourage social interactions may be important to create sustainable change. An important caveat for the interpretation of the findings is that we provide evidence of program persistence for a pilot program that only lasted one year. As such it differs from many other CCT programs in Latin America, where beneficiaries often receive transfers for many years. This could have implications for the external validity of the findings, as households do not necessarily react similarly to a one-year transitory shock than to a longer-term transfer program. The chapter hence should primarily be seen as a proof-of-concept for the role local leaders can play in sustainably shifting poor households' educational and nutritional investments. More generally, the findings highlight the potential importance for careful attention to social dynamics in the design and implementation of programs targeting human capital investments.

References

- Adato, Michele. 2000. *Final Report: The Impact of Progresa on Community Social Relationships*. Washington, DC: International Food Policy Research Institute.
- Adato, Michelle, and Terry Roopnaraine. 2004. *Final Report: A Social Analysis of the Red de Proteccion Social (RPS) in Nicaragua*. Washington, DC: International Food Policy Research Institute.
- Appadurai, Arjun. 2004. "The Capacity to Aspire." In *Culture and Public Action*, edited by V. Rao and M. Walton. Palo Alto, CA: Stanford University Press.
- Araujo, Caridad, Mariano Bosch, and Norbert Schady, 2016. "Can Cash Transfers Help Households Escape an Inter-generational Poverty Trap?" NBER Working Paper no. 22670, Cambridge, MA.
- Athey, Susan, and Guido W. Imbens. 2016. "The Econometrics of Randomized Experiments." In *Handbook of Economic Field Experiments*, vol. 1, edited by A. Banerjee and E. Duflo. Amsterdam: Elsevier.
- Attanasio, Orazio, Camila Fernandez, Emla Fitzsimons, Sally Grantham-McGregor, Costas Meghir, and Marta Rubio-Codina. 2014. "Using the Infrastructure of a Conditional Cash Transfer Program to Deliver a Scalable Integrated Early Child Development Program in Colombia: Cluster Randomised Controlled Trial." *British Medical Journal* 349:5785. https://doi.org/10.1136/bmj.g5785.
- Baird, Sarah, Craig McIntosh, and Berk Özler. 2016. "When the Money Runs Out: Do Cash Transfers Have Sustained Effects?" Policy Research Working Paper no. 7901, World Bank.
- Barham, Tania, Karen Macours, and John A. Maluccio. 2013a. "Boys' Cognitive Skill Formation and Physical Growth: Long-Term Experimental Evidence on Critical Ages for Early Childhood Interventions." *American Economic Review Papers and Proceedings* 103 (3): 467–71.
- ———. 2013b. "More Schooling and More Learning? Effects of a Three-Year Conditional Cash Transfer Program in Nicaragua after 10 Years." IADB Working Paper no. IDB-WP-432, Inter-American Development Bank.
- Barrera-Osorio, Felipe, Leigh L. Linden, and Juan E. Saavedra. 2015. "Long-Term Educational Consequences of Alternative Conditional Cash Transfer Designs: Experimental Evidence from Colombia." Unpublished manuscript.
- Beaman, Lori, Raghabendra Chattopadhyay, Esther Duflo, Rohini Pande, and Petia Topalova. 2009. "Powerful Women: Does Exposure Reduce Bias?" *Quarterly Journal of Economics* 124 (4): 1497–540.
- Beaman, Lori, Esther Duflo, Rohini Pande, and Petia Topalova. 2012. "Female Leadership Raises Aspirations and Educational Attainment for Girls: A Policy Experiment in India." Science 355:582–86.
- Behrman, Jere R., Susan W. Parker, and Petra E. Todd. 2009. "Schooling Impacts of Conditional Cash Transfers on Young Children: Evidence from Mexico." *Economic Development and Cultural Change* 57 (3): 439–77.

——. 2011. "Do Conditional Cash Transfers for Schooling Generate Lasting Benefits? Five-Year Follow-Up of PROGRESA/Oportunidades." *Journal of Human Resources* 46 (1): 93–122.

- Bernard, Tanguy, Stefan Dercon, Kate Orkin, and Alemayehu Seyoum Taffesse. 2014. "The Future in Mind: Aspirations and Forward-Looking Behaviour in Rural Ethiopia." CSAE Working Paper no. WPS/2014-16, Centre for the Study of African Economies, University of Oxford.
- Besley, Timothy. 2017. "Aspirations and the Political Economy of Inequality." *Oxford Economic Papers* 69 (1): 1–35.
- Chattopadhyay, Raghabendra, and Esther Duflo. 2004. "Women as Policy Makers: Evidence from a Randomized Policy Experiment in India." *Econometrica* 72 (5): 1409–43.

- Feigenberg, Benjamin, Erica Field, and Rohini Pande. 2013. "The Economic Returns to Social Interaction: Experimental Evidence from Microfinance." *Review of Economic Studies* 80 (4): 1459–83.
- Filmer, Deon, and Norbert R. Schady. 2014. "The Medium-Term Effects of Scholarships in a Low-Income Country." *Journal of Human Resources* 49 (3): 663–94.
- Fiszbein, Ariel, and Norbert R. Schady, 2009. "Conditional Cash Transfers: Reducing Present and Future Poverty." World Bank Policy Research Report, Washington, DC, World Bank.
- Ganimian, Alejandro J., and Richard J. Murnane, 2014. "Improving Educational Outcomes in Developing Countries: Lessons from Rigorous Evaluations." NBER Working Paper no. 20284, Cambridge, MA.
- Genicot, Garance, and Debraj Ray. 2014. "Aspirations and Inequality." NBER Working Paper no. 19976, Cambridge, MA.
- Gertler, Paul, Sebastian Martinez, and Marta Rubio-Codina. 2012. "Investing Cash Transfers to Raise Long-Term Living Standards." *American Economic Journal: Applied Economics* 4 (1): 164–92.
- Hoff, Karla, and Joseph E. Stiglitz. 2016. "Striving for Balance in Economics: Towards a Theory of the Social Determination of Behavior." *Journal of Economic Behavior & Organization* 126 (PB): 25–57.
- Jensen, Robert. 2010. "Impact of Information on the Returns to Education on the Demand for Schooling in the Dominican Republic." *Quarterly Journal of Economics* 125:515–48.
- Jensen, Robert, and Emily Oster. 2009. "The Power of TV: Cable Television and Women's Status in India." *Quarterly Journal of Economics* 124 (3): 1057–94.
- Kling, Jeffrey, R., Jeffrey B. Liebman, and Lawrence F. Katz. 2007. "Experimental Analysis of Neighborhood Effects." *Econometrica* 75 (1): 83–119.
- La Ferrara, Eliana, Alberto Chong, and Suzanne Duryea. 2012. "Soap Operas and Fertility: Evidence from Brazil." *American Economic Journal: Applied Economics* 4 (4): 1–31.
- Lybbert, Travis, and Bruce Wydick. 2016. "Poverty, Aspirations, and the Economics of Hope: A Framework for Study with Preliminary Results from the Oaxaca Hope Project." Paper presented at NBER conference "The Economics of Asset Dynamics and Poverty Traps," Washington, DC, June 28. https://basis.ucdavis .edu/sites/g/files/dgvnsk466/files/inline-files/NBER-Conference-DC_6-28-16.pdf.
- Macours, Karen, Patrick Premand, and Renos Vakis. 2012. "Transfers, Diversification and Household Risk Strategies: Experimental Evidence with Lessons for Climate Change Adaptation." CEPR Discussion Paper no. 8940, London, Center for Economic Policy Research.
- Macours, Karen, Norbert R. Schady, and Renos Vakis. 2012. "Cash Transfers, Behavioral Changes, and Cognitive Development in Early Childhood: Evidence from a Randomized Experiment." *American Economic Journal: Applied Economics* 4 (2): 247–73.
- Macours, Karen, and Renos Vakis. 2014. "Changing Households' Investment Behaviour through Social Interactions with Local Leaders: Evidence from a Randomized Transfer Program." *Economic Journal* 124:607–33.
- Manski, Charles F. 2000. "Economic Analysis of Social Interactions." Journal of Economic Perspectives 14 (3): 115–36.
- Molina-Millan, Teresa, Tania Barham, Karen Macours, John A. Maluccio, and Marco Stampini. 2016. "Long-Term Impacts of Conditional Cash Transfers in Latin America: Review of the Evidence." IAB Technical Note no. IDB-N-923, Inter-American Development Bank.

- Nguyen, Trang. 2008. "Information, Role Models and Perceived Returns to Education: Experimental Evidence from Madagascar." Working paper, Massachusetts Institute of Technology.
- Paluck, Elizabeth Levy, and Hana Shepherd. 2012. "The Salience of Social Referents: A Field Experiment on Collective Norms and Harassment Behavior in a School Social Network." *Journal of Personality and Social Psychology* 103 (6): 899–915.
- Ray, Debraj. 2006. "Aspirations, Poverty, and Economic Change." In *Understanding Poverty*, edited by Abhijit V. Banerjee, Roland Bénabou, and Dilip Mookherjee. Oxford: Oxford University Press.
- World Bank. 2015. *World Development Report 2015: Mind, Society and Behavior*. Washington, DC: World Bank.
- Wydick, Bruce, Paul Glewwe, and Laine Rutledge. 2013. "Does International Child Sponsorship Work? A Six-Country Study of Impacts on Adult Life Outcomes." *Journal of Political Economy* 121 (2): 393–436.